

# Dimensionality Reduction

## Part 2: MDS

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# Dimensionality Reduction

## □ Linear methods

- PCA (Principal Component Analysis)
- **cMDS** (Classical Multidimensional Scaling)

## □ Non-linear methods

- KPCA (Kernel PCA)
- **mMDS** (Metric MDS)
- Isomap
- LLE (Locally Linear Embedding)
- Laplacian Eigenmap
- t-SNE (t-distributed Stochastic Neighbor Embedding)
- UMAP (Uniform Manifold Approximation and Projection)

# Multidimensional Scaling (MDS)

- Classical MDS (cMDS)
  - Reconstruct coordinates from **Euclidean distance** matrix
- Metric MDS (mMDS)
  - Redefined cMDS problem with loss function defined on any metric
- Non-metric MDS (nMDS)
  - When only an ordering on the distances is known
- Generalized (kernel) classical MDS

# Classical MDS (cMDS)

- Reconstruct a set of points given their **Euclidean distances**
- Given  $n \times n$  distance matrix  $D = (d_{ij})$ , reconstruct coordinates  $x_1, \dots, x_n \in \mathbb{R}^m$  with  $\|x_i - x_j\| = d_{ij}$ 
  - The solution  $X = [x_1 \ \dots \ x_n]^T \in \mathbb{R}^{n \times m}$  is not unique due to infinitely many translations, rotations, and reflections
  - A centered solution  $X = (x_{ij})$  (i.e.  $\forall k, 1 \leq k \leq m, \sum_i x_{ik} = 0$ ) can be found using cMDS
    - Note that solution is still not unique

# cMDS idea

- Given  $D = (d_{ij})$ , first note that Euclidean distance  $d_{ij}$  is related to  $X = (x_{ij})$  through
$$(d_{ij})^2 = (x_i - x_j)^T (x_i - x_j) = x_i^T x_i + x_j^T x_j - 2x_i^T x_j$$
- On the other hand, for  $X$  where  $\forall k, 1 \leq k \leq m, \sum_i x_{ik} = 0$ , we can show that

$$A = -2XX^T$$

where  $A = (d_{ij}^2)$

- Then it suffices that we compute  $-A/2$  to obtain  $XX^T$
- Finally, since  $XX^T$  can be decomposed to recover  $X$

# cMDS algorithm

- **Step 1.** Compute matrix  $CAC$

Given  $D = (d_{ij})$ , computed  $CAC$

where

$$A = \left( -\frac{1}{2} d_{ij}^2 \right)$$

$$C = I - \frac{1}{n} \mathbf{1}^T \mathbf{1}$$

- $CAC$  simultaneously centers the rows and columns of the squared distance matrix  $A$  (double centering)
- It can be shown that  $CAC = XX^T$  for centered  $X$  (proof in later slides)
  - $\Rightarrow CAC$  is positive semi-definite (proof later)
  - $\Rightarrow CAC$  decompose to non-negative values

# cMDS algorithm

- **Step 2.** Decompose  $CAC$  into orthonormal basis

Method 1: Eigendecompose  $CAC$  into  $Q\Lambda Q^T$

- Then,  $X$  can be computed as  $Q\Lambda^{1/2}$

- $Q\Lambda Q^T = Q\Lambda^{1/2}\Lambda^{1/2}Q^T = Q\Lambda^{1/2}(Q\Lambda^{1/2})^T = XX^T$

Method 2: Decompose  $CAC$  directly into  $XX^T$  using Cholesky factorization

- Only works if  $CAC$  is positive definite
- $CAC$  is (positive semi-definite and) positive definite iff all  $x_i$  are linearly independent
  - Cholesky in numpy/scipy will not execute unless the input is positive definite
  - Use one that works (e.g. pyre) or write your own with pivoting

# cMDS algorithm

- **Step 3.** Choose from the decomposed basis

Both methods face the problem that the output matrix is not of dimension  $n \times m$

- Eigendecomposition  $Q \in \mathbb{R}^{n \times n}$
- Cholesky factorization  $L \in \mathbb{R}^{n \times n}$

- If  $n < m$  (fewer datapoints than features)
  - No problem in embedding the points since  $n$  points can fit on an  $(n - 1)$ -D plane
  - Naturally suited for dimensionality reduction purpose if use all  $n - 1$  eigenvectors
    - If need fewer than  $(n - 1)$ -D space, see later slides



# cMDS algorithm

- **Step 3.** Choose from the decomposed basis

Both methods face the problem that the output matrix is not of dimension  $n \times m$

- Eigendecomposition  $Q \in \mathbb{R}^{n \times n}$
- Cholesky factorization  $L \in \mathbb{R}^{n \times n}$

- If  $n > m$  (more datapoints than features)

## Problem 1

- $CAC$  is not positive definite since there are insufficient features for linear independence
  - Bad news for Cholesky factorization

# cMDS algorithm

- **Step 3.** Choose from the decomposed basis

Both methods face the problem that the output matrix is not of dimension  $n \times m$

- Eigendecomposition  $Q \in \mathbb{R}^{n \times n}$
- Cholesky factorization  $L \in \mathbb{R}^{n \times n}$

- If  $n > m$  (more datapoints than features)

## Problem 2

- Need to deduce  $m$  (or fewer) vectors
  - In an ideal eigendecomposition there will be  $\text{rank}(XX^T) (\leq m)$  positive eigenvalues and  $n - \text{rank}(XX^T)$  zero eigenvalues
  - But eigendecomposition usually not ideal with zero eigenvalues, often resulting in complex numbers

# cMDS algorithm

- **Step 3.** Choose from the decomposed basis
  - Many implementations will output negative eigenvalues, so extra care is needed
  - For eigendecomposition
    - Remove the eigenpairs with small, negative, or complex eigenvalues, forming  $Q_1$  and  $\Lambda_1$
    - Choose the set of eigenvalues  $S$  from  $\Lambda_1$  such that  $\frac{\sum_{\lambda' \in S} \lambda'}{\sum_{\lambda \in \Lambda_1} \lambda}$  is sufficiently large
    - Finally, compute  $Q_1 \Lambda_1^{1/2}$  and retain only those in  $S$
    - Remove extraneous dimensions in  $Q_1 \Lambda_1^{1/2}$
  - For Cholesky factorization
    - Choose the vectors with the largest norms

# (Proof) $XX^T = CAC$ for centered $X$

- Will expand  $XX^T$  and  $CAC$  and show equivalence
- Given  $X \in \mathbb{R}^{m \times n}$ , denote  $XX^T$  as  $B$ , then we can write the Euclidean distance between  $x_i$  and  $x_j$  as

$$d_{ij}^2 = b_{ii} + b_{jj} - 2b_{ij} \quad (1)$$

- If  $\forall k, \sum_i x_{ik} = 0$  ( $X$  is centered), then

$$\sum_{j=1}^n b_{ij} = \sum_{j=1}^n \sum_{k=1}^m x_{ik} x_{jk} = \sum_{k=1}^m x_{ik} \left( \sum_{j=1}^n x_{jk} \right) = 0$$

$$\text{Denote } \text{tr}(B) = \sum_{i=1}^n b_{ii}, \because \sum_{j=1}^n b_{ij} = 0, (1) \Rightarrow$$

$$\left. \begin{aligned} \sum_{i=1}^n d_{ij}^2 &= \sum_{i=1}^n b_{ii} + \sum_{i=1}^n b_{jj} = \text{tr}(B) + nb_{jj} \\ \sum_{j=1}^n d_{ij}^2 &= \sum_{j=1}^n b_{ii} + \sum_{j=1}^n b_{jj} = nb_{ii} + \text{tr}(B) \\ \sum_{i,j=1}^n d_{ij}^2 &= \sum_{i,j=1}^n b_{ii} + \sum_{i,j=1}^n b_{jj} = 2n\text{tr}(B) \end{aligned} \right\} (2)$$

# (Proof) $XX^T = CAC$ for centered $X$

Rewrite (1) as  $b_{ij} = \frac{1}{2} (b_{ii} + b_{jj} - d_{ij}^2)$ , then (1)+(2)

$$\begin{aligned} b_{ij} &= \frac{1}{2} \left( \frac{1}{n} \left( \sum_{i=1}^n d_{ij}^2 - \text{tr}(B) \right) + \sum_{j=1}^n d_{ij}^2 - \text{tr}(B) \right) - d_{ij}^2 \\ &= \frac{1}{2} \left( \frac{1}{n} \left( \sum_{i=1}^n d_{ij}^2 + \frac{1}{n} \sum_{j=1}^n d_{ij}^2 - \frac{1}{n} \sum_{i,j=1}^n d_{ij}^2 \right) - d_{ij}^2 \right) \end{aligned}$$

Done, but for notation simplicity let  $a_{ij} = -\frac{1}{2} d_{ij}^2$ , then

$$b_{ij} = -\frac{1}{n} \sum_{i=1}^n a_{ij} - \frac{1}{n} \sum_{j=1}^n a_{ij} + \frac{1}{n^2} \sum_{i,j=1}^n a_{ij} + a_{ij}$$

Further make things easy to see with

$$\begin{aligned} a_{i\blacksquare} &= \frac{1}{n} \sum_{i=1}^n a_{ij}, \quad a_{\blacksquare j} = \frac{1}{n} \sum_{j=1}^n a_{ij}, \quad a_{\blacksquare\blacksquare} = \frac{1}{n^2} \sum_{i,j=1}^n a_{ij} \\ \Rightarrow b_{ij} &= a_{ij} - a_{i\blacksquare} - a_{\blacksquare j} + a_{\blacksquare\blacksquare} \end{aligned}$$

# (Proof) $XX^T = CAC$ for centered $X$

- Now expand  $CAC$  into terms consisting of  $a_{ij}$

Given  $A = (a_{ij})$ , observe that

$$\left. \begin{aligned} [1 \ 1 \ \dots \ 1]A &= n(a_{i\blacksquare}) \\ A[1 \ 1 \ \dots \ 1] &= n(a_{\blacksquare j}) \\ [1 \ 1 \ \dots \ 1]A[1 \ 1 \ \dots \ 1] &= n^2(a_{\blacksquare\blacksquare}) \end{aligned} \right\} (3)$$

On the other hand,

$$\begin{aligned} CAC &= \left(I - \frac{1}{n}J\right) A \left(I - \frac{1}{n}J\right) \\ &= A - \frac{1}{n}JA - \frac{1}{n}AJ + \frac{1}{n^2}JAJ \end{aligned} \quad (4)$$

Finally, (3)+(4) gives

$$(CAC)_{ij} = a_{ij} - a_{i\blacksquare} - a_{\blacksquare j} + a_{\blacksquare\blacksquare} = b_{ij}$$

# (Proof) $CAC$ is PSD

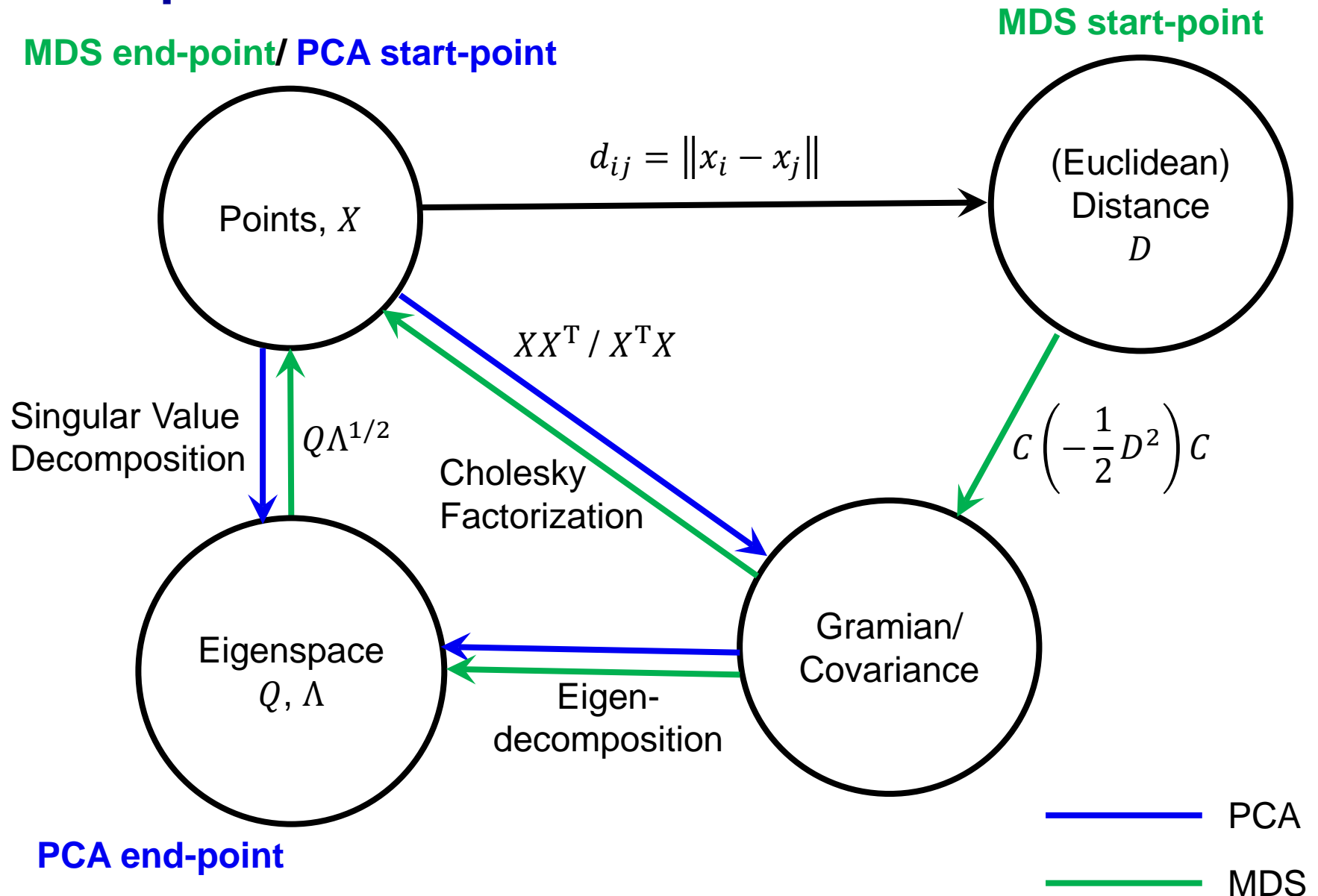
- Follows immediately from the fact that  $CAC = XX^T$ , an inner product
  - An inner product  $B = XX^T$  of any matrix  $X$  (centered or not) is called a Gram matrix, or Gramian
  - Gramians are known to be positive semi-definite (see proof in other slides)

# Comparison with PCA

	MDS	PCA
Input	$n \times n$ Euclidean distances $D$	$n \times m$ Dataset $X$
Matrix considered in theory	$n \times n$ Gramian $XX^T$	$m \times m$ Covariance matrix $X^T X$
Matrix used for decomposition	$CAC$ where $A = -\frac{1}{2}D^2$ and $C$ is the centering matrix	$XX^T$ or $X^T X$
Output	Reconstructed $X$ Alternatively, $X$ in lower dimension	Principal components of $X^T X$ (or $XX^T$ )
Decomposition Method	Cholesky factorization or Eigendecomposition	SVD or Eigendecomposition
$n < m$	Exact reconstruction of $X$ impossible	No problem
$m < n$	For exact reconstruction of $X$ , rank deficiency revealed in eigendecomposition needed to deduce $m$	No problem



# Comparison with PCA



# Limitation of cMDS

- For cMDS to work, input distances have to be Euclidean
- More precisely, the Pythagorean principle

$$(d_{ij})^2 = (x_i - x_j)^T (x_i - x_j)$$

(or, in terms of the Gramian,  $d_{ij}^2 = b_{ii} + b_{jj} - 2b_{ij}$ )

is used in establishing the relation  $XX^T = CAC$

- Such a relationship cannot be assumed for most datasets
- $XX^T = CAC$  does not hold for other metrics

# Metric MDS (mMDS)

- Given distance matrix  $(\delta_{ij})_{n \times n}$  and weights  $(w_{ij})_{n \times n}$ , find  $X = [x_1 \quad \dots \quad x_n]^T$  where  $x_i \in \mathbb{R}^r$ , which minimizes

$$\text{stress}(X) = \sum_{i,j,i < j} w_{ij} (d(x_i, x_j) - \delta_{ij})^2$$

where  $d(x_i, x_j)$  denotes the distance between  $x_i$  and  $x_j$

- The weights  $w_{ij}$  allow removing entries where  $\delta_{ij}$  is not available

# SMACOF Algorithm for mMDS

□ Minimize  $\text{stress}(X)$  through majorization

□ 
$$\text{stress}(X) = \sum_{i,j,i < j} w_{ij} (d(x_i, x_j) - \delta_{ij})^2$$
$$= \sum w_{ij} d^2(x_i, x_j) + \sum w_{ij} \delta_{ij}^2 - 2 \sum w_{ij} \delta_{ij} d(x_i, x_j)$$

Since

- $\sum w_{ij} \delta_{ij}^2$  is constant,  $C$
- $\sum w_{ij} d^2(x_i, x_j)$  is quadratic,  $\text{tr}(X' V X)$
- $\sum w_{ij} \delta_{ij} d(x_i, x_j) = \text{tr}(X' B(X) X) \geq \text{tr}(X') B(Z) Z$   
where  $B(Z) = (b_{ij})$  for

$$b_{ij} = \begin{cases} -\frac{w_{ij} \delta_{ij}}{d(x_i, x_j)} & \text{if } d(x_i, x_j) \neq 0 \text{ and } i \neq j \\ 0 & \text{if } d(x_i, x_j) = 0 \text{ and } i \neq j \end{cases}$$
$$b_{ii} = -\sum_{j=1, j \neq i}^n b_{ij}$$

# SMACOF Algorithm for mMDS

- $\text{stress}(X) = C + \text{tr}(X'VX) - 2\text{tr}(X'B(X)X)$   
which is bounded above by  
 $C + \text{tr}(X'VX) - 2\text{tr}(X'B(Z)Z) = \tau(X, Z)$
- Majorization iteratively updates  $X^k$  at the  $k^{\text{th}}$  iteration to  $\min_X \tau(X, X^{k-1})$ 
  - $\text{stress}(X)$  will decrease monotonically
  - Stops iteration when  $\text{stress}(X^k) - \text{stress}(X^{k-1})$  is below a given threshold
- Proofs for the majorization method requires too much details to provide here

# Sammon mapping

- A special case of  $\text{stress}(X)$  where weights are inversely proportional to distance  $\delta_{ij}$ 
  - Emphasize accuracy on small  $\delta_{ij}$  distances

- Given distance matrix  $(\delta_{ij})_{n \times n}$ , find  $X = [x_1 \ \dots \ x_n]^T$  where  $x_i \in \mathbb{R}^r$ , which minimizes

$$\text{stress}(X) = \frac{1}{\sum_{i,j,i < j} \delta_{ij}} \sum_{i,j,i < j} \frac{(d(x_i, x_j) - \delta_{ij})^2}{\delta_{ij}}$$

where  $d(x_i, x_j)$  denotes the distance between  $x_i$  and  $x_j$

- The simpler  $\text{stress}(X)$  allows a gradient descent optimization

# nMDS vs cMDS

## □ Similarity vs dissimilarity

- cMDS attempts to recover  $XX^T$ , a measure of the similarity between  $x_i$  and  $x_j$
- nMDS attempts to recover distances  $d(x_i, x_j)$ , a measure of the dissimilarity between  $x_i$  and  $x_j$

## □ Linear vs non-linear

- cMDS attempts to recover  $XX^T$ , a linear kernel
- nMDS, for instance Sammon mapping, can be considered as recovering a non-linear distance measure with an inverse  $(1/\delta)$  factor

## □ Closed-form vs iterative method

- cMDS is solved through a closed-form solution
- nMDS can only be approximated iteratively using gradient descent or majorization