

Statistical Inference and Multivariate Analysis (MA324)

LECTURE SLIDES Lecture 02

Review of different Transformation Techniques



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Functions of Random Variables: Technique 1

In Technique 1, we try to find the JCDF of $Y = g(X)$ given the distribution of X . We will discuss this technique using examples.

Example 1: Let X_1 and X_2 be *i.i.d.* $U(0, 1)$ random variables. Find the CDF of $Y = X_1 + X_2$.

Functions of RVs: Technique 2 for DRV

Theorem: Let $\mathbf{X} = (X_1, X_2, \dots, X_n)$ be a DRV with JPMF $f_{\mathbf{X}}$ and support $S_{\mathbf{X}}$. Let $g_i : \mathbb{R}^n \rightarrow \mathbb{R}$ for all $i = 1, 2, \dots, k$. Let $Y_i = g_i(\mathbf{X})$ for $i = 1, 2, \dots, k$. Then $\mathbf{Y} = (Y_1, \dots, Y_k)$ is a DRV with JPMF

$$f_{\mathbf{Y}}(y_1, \dots, y_k) = \begin{cases} \sum_{\mathbf{x} \in A_{\mathbf{y}}} f_{\mathbf{X}}(\mathbf{x}) & \text{if } (y_1, \dots, y_k) \in S_{\mathbf{Y}} \\ 0 & \text{otherwise,} \end{cases}$$

where $A_{\mathbf{y}} = \{\mathbf{x} \in S_{\mathbf{X}} : g_i(\mathbf{x}) = y_i, i = 1, \dots, k\}$ and $S_{\mathbf{Y}} = \{(g_1(\mathbf{x}), \dots, g_k(\mathbf{x})) : \mathbf{x} \in S_{\mathbf{X}}\}$.

Functions of RVs: Technique 2 for DRV

Example 2: $X_1 \sim P(\lambda_1)$ and $X_2 \sim P(\lambda_2)$ and they are independent. Then $X_1 + X_2 \sim P(\lambda_1 + \lambda_2)$.

Example 3: $X_1 \sim \text{Bin}(n_1, p)$ and $X_2 \sim \text{Bin}(n_2, p)$ and they are independent. Then $X_1 + X_2 \sim \text{Bin}(n_1 + n_2, p)$.

Example 4: $X_i \sim \text{Bin}(n_i, p)$, $i = 1, 2, \dots, m$ and X_i 's are independent. Then $\sum_{i=1}^m X_i \sim \text{Bin}(\sum_{i=1}^m n_i, p)$.

Functions of RVs: Technique 2 for CRV

Theorem: Let $\mathbf{X} = (X_1, \dots, X_n)$ be a CRV with JPDP $f_{\mathbf{X}}$.

- 1 Let $y_i = g_i(\mathbf{x})$, $i = 1, 2, \dots, n$ be $\mathbb{R}^n \rightarrow \mathbb{R}$ functions such that $\mathbf{y} = \mathbf{g}(\mathbf{x})$ is one-to-one. That means that there exists the inverse transformation $x_i = h_i(\mathbf{y})$, $i = 1, 2, \dots, n$ defined on the range of the transformation.
- 2 Assume that both the mapping and its' inverse are continuous.
- 3 Assume that partial derivatives $\frac{\partial x_i}{\partial y_j}$, $i = 1, 2, \dots, n$, $j = 1, 2, \dots, n$, exist and are continuous.
- 4 Assume that the Jacobian of the inverse transformation

$$J \doteq \det \left(\frac{\partial x_i}{\partial y_j} \right)_{i,j=1,2,\dots,n} \neq 0$$

on the range of the transformation.

Then $\mathbf{Y} = (g_1(\mathbf{X}), \dots, g_n(\mathbf{X}))$ is a CRV with JPDP

$$f_{\mathbf{Y}}(\mathbf{y}) = f_{\mathbf{X}}(h_1(\mathbf{y}), \dots, h_n(\mathbf{y}))|J|.$$

Functions of RVs: Technique 2 for CRV

Example 5: Let X_1 and X_2 be *i.i.d.* $U(0, 1)$ random variables. Find the JPDF of $Y_1 = X_1 + X_2$ and $Y_2 = X_1 - X_2$.

Example 6: Let X_1 and X_2 be *i.i.d.* $N(0, 1)$ random variables. Find the PDF of $Y_1 = X_1/X_2$.

Remark: If X and Y are independent, then $g(X)$ and $h(Y)$ are also independent.

Moment Generating Function: Technique 3

Let $\mathbf{X} = (X_1, X_2, \dots, X_n)$ be a RV. The moment generating function (MGF) of \mathbf{X} at $\mathbf{t} = (t_1, t_2, \dots, t_n)$ is defined by

$$M_{\mathbf{X}}(\mathbf{t}) = E\left(\exp\left(\sum_{i=1}^n t_i X_i\right)\right),$$

provided the expectation exists in a neighborhood of origin $\mathbf{0} = (0, 0, \dots, 0)$.

Remark: $E(X_1^{r_1} X_2^{r_2} \cdots X_n^{r_n}) = \frac{\partial^{r_1+r_2+\dots+r_n}}{\partial t_1^{r_1} \partial t_2^{r_2} \cdots \partial t_n^{r_n}} M_{\mathbf{X}}(\mathbf{t}) \Big|_{\mathbf{t}=\mathbf{0}}.$

Def: Two RVs X and Y are said to have the same distribution, denoted by $X \stackrel{d}{=} Y$, if $F_X(\cdot) = F_Y(\cdot)$.

Theorem: Let X and Y be two RVs. Let $M_X(t) = M_Y(t)$ for all t in a neighborhood around 0, then $X \stackrel{d}{=} Y$.

Theorem: X and Y are independent iff $M_{X,Y}(t_1, t_2) = M_X(t_1)M_Y(t_2)$.

Example 7: Let $X_i, i = 1, 2, \dots, k$ be independent $Bin(n_i, p)$ RVs. Then $\sum X_i \sim Bin(\sum n_i, p)$.

Example 8: Let $X_i, i = 1, 2, \dots, k$ be iid $Exp(\lambda)$ RVs. Then $\sum X_i \sim Gamma(k, \lambda)$.

Example 9: Let $X_i, i = 1, 2, \dots, k$ be independent $N(\mu_i, \sigma_i^2)$ RVs. Then $\sum X_i \sim N(\sum \mu_i, \sum \sigma_i^2)$.