

# Statistical Inference and Multivariate Analysis (MA324)

LECTURE SLIDES  
Lecture 24

## Interval Estimation



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# Two-sample Problems

**Example 1:** Let  $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu_1, \sigma^2)$  and  $Y_1, Y_2, \dots, Y_m \stackrel{i.i.d.}{\sim} N(\mu_2, \sigma^2)$ . Also, assume that  $X_i$ 's and  $Y_j$ 's are independent. Here,  $\mu_1, \mu_2$ , and  $\sigma$  are assumed to be unknown and we are interested to construct a  $100(1 - \alpha)\%$  CI for  $\mu_1 - \mu_2$ . Let us first try to construct a pivot based on minimal sufficient statistic  $(\bar{X}, \bar{Y}, S^2)$ , where the pooled sample variance  $S^2$  is defined by

$$S^2 = \frac{1}{n + m - 2} \left[ \sum_{i=1}^n (X_i - \bar{X})^2 + \sum_{i=1}^m (Y_i - \bar{Y})^2 \right].$$

Now, notice that

$$\bar{X} - \bar{Y} \sim N \left( \mu_1 - \mu_2, \sigma^2 \left( \frac{1}{n} + \frac{1}{m} \right) \right) \implies T = \frac{(\bar{X} - \bar{Y}) - (\mu_1 - \mu_2)}{\sigma \sqrt{\frac{1}{n} + \frac{1}{m}}} \sim N(0, 1)$$

Of course,  $T$  is a pivot, but we cannot use it to construct the required confidence interval due to the presence of unknown  $\sigma$  in  $T$ .

**Example 2:** Let  $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} N(\mu_1, \sigma_1^2)$  and  $Y_1, Y_2, \dots, Y_m \stackrel{i.i.d.}{\sim} N(\mu_2, \sigma_2^2)$ . Also, assume that  $X_i$ 's and  $Y_j$ 's are independent. Here,  $\mu_1, \mu_2, \sigma_1$ , and  $\sigma_2$  are assumed to be unknown and we are interested to construct a  $100(1 - \alpha)\%$  CI for  $\frac{\sigma_2^2}{\sigma_1^2}$ . In this case minimal sufficient statistic is  $(\bar{X}, \bar{Y}, S_1^2, S_2^2)$ , where  $S_1^2$  and  $S_2^2$  are sample variances based on the samples  $X_i$ 's and  $Y_j$ 's, respectively. In this case,

$$T = \frac{\sigma_2^2 S_1^2}{\sigma_1^2 S_2^2} = \frac{\frac{(n-1)S_1^2}{(n-1)\sigma_1^2}}{\frac{(m-1)S_2^2}{(m-1)\sigma_2^2}} \sim F_{n-1, m-1}.$$

Thus,

$$P\left(F_{n-1, m-1, 1-\frac{\alpha}{2}} \leq \frac{\sigma_2^2 S_1^2}{\sigma_1^2 S_2^2} \leq F_{n-1, m-1, \frac{\alpha}{2}}\right) = 1 - \alpha$$

so that a  $100(1 - \alpha)\%$  CI for  $\frac{\sigma_2^2}{\sigma_1^2}$  is

$$C(\mathbf{X}) = \left[ \frac{S_2^2}{S_1^2} F_{n-1, m-1, 1-\frac{\alpha}{2}}, \frac{S_2^2}{S_1^2} F_{n-1, m-1, \frac{\alpha}{2}} \right].$$

# Asymptotic CI

- In many cases it is very **difficult** to **find pivot** for a **small sample**. For example, it is difficult to find a pivot to construct CI for successes probability of a Bernoulli distribution.
- However, we may be able to find CI quite easily if the **sample size** is sufficiently **large**. This CI is called **asymptotic confidence interval**.
- For this purpose, **convergence in distribution** (mainly CLT or large sample distribution of MLE) and **convergence in probability** (consistent estimator) are handy tools.

# Distribution Free Population Mean

Let  $X_1, X_2, \dots$  be i.i.d. random variables with mean  $\mu$  and finite variance  $\sigma^2$ . Then, using CLT

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{\mathcal{L}} Z \sim N(0, 1).$$

Thus, if we have a RS with large sample size  $n$ , we can approximate the distribution of  $\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}$  using a standard normal distribution. Hence,

$$P\left(-z_{\alpha/2} \leq \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \leq z_{\alpha/2}\right) \approx 1 - \alpha.$$

If  $\sigma$  is **known** and  $n$  is sufficiently large, we can use the last statement to find an asymptotic CI for  $\mu$  and it is given by

$$\left[\bar{X}_n - \frac{\sigma}{\sqrt{n}}z_{\alpha/2}, \bar{X}_n + \frac{\sigma}{\sqrt{n}}z_{\alpha/2}\right].$$

If  $\sigma$  is **unknown**, we can proceed as follows. Using WLLN, we have  $\frac{S_n}{\sigma} \xrightarrow{P} 1$ , and hence,

$$\frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \xrightarrow{\mathcal{L}} Z \sim N(0, 1).$$

Hence,  $P\left(-z_{\alpha/2} \leq \frac{\sqrt{n}(\bar{X}_n - \mu)}{S_n} \leq z_{\alpha/2}\right) \approx 1 - \alpha$ . An asymptotic CI for  $\mu$  is given by

$$\left[\bar{X}_n - \frac{S_n}{\sqrt{n}}z_{\alpha/2}, \bar{X}_n + \frac{S_n}{\sqrt{n}}z_{\alpha/2}\right].$$

Note that this **method can be used for any distribution** of  $X_1, X_2, \dots, X_n$ , as long as the **conditions of CLT hold true**. Therefore, it is called **distribution free**.

# Using MLE

Let  $\hat{\theta}_n$  be a **consistent estimator** of  $\theta$  and  $\sqrt{n}(\hat{\theta}_n - \theta) \xrightarrow{\mathcal{L}} N(0, b^2(\theta))$ , where  $b(\theta) > 0$  for all  $\theta \in \Theta$ . Assume that  $b(\cdot)$  is a **continuous function**. Then,  $\frac{b(\hat{\theta}_n)}{b(\theta)} \xrightarrow{P} 1$ , and hence,  $\frac{\sqrt{n}(\hat{\theta}_n - \theta)}{b(\hat{\theta}_n)} \xrightarrow{\mathcal{L}} N(0, 1)$ . A  $100(1 - \alpha)\%$  asymptotic CI for  $\theta$  is given by

$$\left[ \hat{\theta}_n - \frac{b(\hat{\theta}_n)}{\sqrt{n}} z_{\alpha/2}, \hat{\theta}_n + \frac{b(\hat{\theta}_n)}{\sqrt{n}} z_{\alpha/2} \right].$$

Under some regularity conditions, we may use MLE of  $\theta$  in place of  $\hat{\theta}_n$ .

**Example 3:** Let  $X_1, X_2, \dots, X_n \stackrel{i.i.d.}{\sim} \text{Bernoulli}(p)$ , where  $p \in (0, 1)$ . We are interested to construct asymptotic CI for  $p$ . We know that  $\hat{p}_n = \bar{X}_n \xrightarrow{P} p$  and  $\frac{\sqrt{n}(\bar{X}_n - p)}{\sqrt{p(1-p)}} \xrightarrow{\mathcal{L}} N(0, 1)$ . Here,  $b(p) = \sqrt{p(1-p)}$ , which is a continuous function in  $p \in (0, 1)$ . Hence,  $\frac{\sqrt{n}(\bar{X}_n - p)}{\sqrt{\bar{X}_n(1-\bar{X}_n)}} \xrightarrow{\mathcal{L}} N(0, 1)$ . A  $100(1 - \alpha)\%$  asymptotic CI for  $p$  is

$$\left[ \bar{X}_n - \sqrt{\frac{\bar{X}_n(1 - \bar{X}_n)}{n}} z_{\alpha/2}, \bar{X}_n + \sqrt{\frac{\bar{X}_n(1 - \bar{X}_n)}{n}} z_{\alpha/2} \right].$$