

Q1) the joint pdf of (X, Y) is

$$f_{X,Y}(x,y) = \begin{cases} \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} x^{\alpha_1 - 1} y^{\alpha_2 - 1} e^{-\beta(x+y)} & x > 0, y > 0 \\ 0 & \text{o.w.} \end{cases}$$

Let $U = \frac{X}{X+Y}$ and $V = X+Y$

Here

~~$g(x,y) = (g_1(x,y), g_2(x,y))$~~
 ~~$(u,v) = g(x,y) = (g_1(x,y), g_2(x,y))$~~

$(u,v) = g(x,y) = (g_1(x,y), g_2(x,y))$

$= \left(\frac{x}{x+y}, x+y \right)$

and

$(x,y) = h(u,v) = (h_1(u,v), h_2(u,v))$

$= (uv, v(1-u))$

✓ clearly g is one-to-one

✓ Both g and its ^{inverse} exist and they are continuous.

partial derivatives exist and continuous as,

$\frac{\partial x}{\partial u} = v, \quad \frac{\partial x}{\partial v} = u$

$\frac{\partial y}{\partial u} = -v, \quad \frac{\partial y}{\partial v} = 1-u$

✓ Jacobian,

$J = \begin{vmatrix} v & u \\ -v & 1-u \end{vmatrix} = v > 0$

hence all the four conditions hold.

therefore, the joint PDF of (U,V) is,

~~$h(u,v)$~~

$$f_{U,V}(u,v) = \begin{cases} \frac{\beta^{\alpha_1+\alpha_2}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \cdot (uv)^{\alpha_1-1} (1-u)^{\alpha_2-1} e^{-\beta v} & |v| \\ & 0 < u < 1 \\ & v > 0 \\ 0 & \text{o.w.} \end{cases}$$

so, the marginal PDF of U is

$$f_U(u) = \int_0^\infty \frac{\beta^{\alpha_1+\alpha_2} (1-u)^{\alpha_2-1}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} u^{\alpha_1-1} v^{\alpha_1+\alpha_2-1} e^{-\beta v} dv$$

$$= \frac{u^{\alpha_1-1} (1-u)^{\alpha_2-1}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \cdot \int_0^\infty \beta (\beta v)^{\alpha_1+\alpha_2-1} e^{-\beta v} dv$$

$$= \frac{u^{\alpha_1-1} (1-u)^{\alpha_2-1}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \cdot \int_0^\infty z^{\alpha_1+\alpha_2-1} e^{-z} dz$$

$$= \frac{u^{\alpha_1-1} (1-u)^{\alpha_2-1}}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \cdot \Gamma(\alpha_1+\alpha_2)$$

$$= \frac{u^{\alpha_1-1} (1-u)^{\alpha_2-1}}{B(\alpha_1, \alpha_2)}$$

Hence, the PDF of $U = \frac{X}{X+Y}$, is

$$f_U(u) = \begin{cases} \frac{1}{B(\alpha_1, \alpha_2)} u^{\alpha_1-1} (1-u)^{\alpha_2-1}, & \text{if } 0 < u < 1 \\ 0 & \text{o.w.} \end{cases}$$

Q2) The marginal PDF of X is,

$$f_X(x) = \begin{cases} 4x \int_0^1 y \, dy & \text{if } 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

$$= \begin{cases} 2x & \text{if } 0 < x < 1 \\ 0 & \text{o.w.} \end{cases}$$

Similarly, the marginal PDF of Y is,

$$f_Y(y) = \begin{cases} 2y & \text{if } 0 < y < 1 \\ 0 & \text{o.w.} \end{cases}$$

(b) From above, it is clear that,

$$f_{X,Y}(x,y) = f_X(x) f_Y(y) \text{ for all } (x,y) \in \mathbb{R}^2.$$

Hence, X and Y are independent.

$$(c) \quad P(0 < X < 0.5, 0.25 < Y < 1)$$

$$= P(0 < X < \frac{1}{2}) P(\frac{1}{4} < Y < 1) \quad (X \text{ \& } Y \text{ are independent})$$

$$= \left(\frac{1}{2}\right)^2 \times \left(1 - \left(\frac{1}{4}\right)^2\right)$$

$$= \frac{1}{4} \times \frac{15}{16} = \frac{15}{64}$$

Also, $P(X+Y < 1)$

$$= \iint_{x+y < 1} f_{X,Y}(x,y) \, dx \, dy$$

$$= \int_0^1 \int_0^{1-y} 4xy \, dx \, dy$$

$$= 4 \int_0^1 y \int_0^{1-y} x \, dx \, dy$$

$$= 2 \cdot \frac{4}{2} \int_0^1 y (1-y)^2 \, dy$$

$$= \frac{4}{2} B(2,3)$$

$$= 2 \cdot \frac{\Gamma(2) \Gamma(3)}{\Gamma(5)}$$

$$= 2 \cdot \frac{2}{24}$$

$$= \frac{4}{24}$$

$$= \frac{1}{6}$$

Q3 (a)

$$\int_0^{\infty} P(X > x) dx$$

$$= \int_0^{\infty} \int_x^{\infty} f_X(t) dt dx, \quad 0 < x < t < \infty.$$

$$= \int_0^{\infty} f_X(t) \int_0^t x \cdot 1 \cdot dx dt$$

$$= \int_0^{\infty} t f_X(t) dt$$

$$= E(X)$$

$$\therefore E(X) = \int_0^{\infty} P(X > x) dx.$$

(Corresponding r.v. \tilde{X})

$$= \int_0^{\infty} x f_X(x) dx$$

by $x = t$

$$= E(X)$$

(b) Now, $E(X) = \int_0^{\infty} P(X > x) dx$

$$= \int_0^1 P(X > x) dx + \int_1^2 P(X > x) dx + \dots \quad \text{--- (1)}$$

From (1),

$$E(X) \leq \int_0^1 P(X > 0) dx + \int_1^2 P(X > 1) dx + \int_2^3 P(X > 2) dx$$

$$= P(X > 0) \int_0^1 dx + P(X > 1) \int_1^2 dx + \dots$$

$$= P(X > 0) + P(X > 1) + \dots$$

$$= \sum_{n=0}^{\infty} P(X > n)$$

Again from (1)

$$E(X) \geq \int_0^1 P(X > 1) dx + \int_1^2 P(X > 2) dx + \dots$$

$$= P(X > 1) + P(X > 2) + \dots$$

$$= \sum_{n=1}^{\infty} P(X > n).$$

Hence, $\sum_{n=1}^{\infty} P(X > n) \leq E(X) \leq \sum_{n=0}^{\infty} P(X > n).$

(84)

$$P(A) = \liminf_{n \rightarrow \infty} \frac{\#(A \cap \{1, 2, \dots, n\})}{n}, \quad A \subseteq \mathbb{N}.$$

If possible, let $P(A)$ is a prob. measure ~~on \mathcal{F}~~ ^{on \mathcal{F}} . ①

For $A = \{k\}$, the set $A \cap \{1, 2, \dots, n\}$ is empty if $n < k$ and has just one element k if $n \geq k$.

Thus, the sequence in the formula ①, for $P(\{k\})$ begin with $(k-1)$ zeros followed by $\frac{1}{k}, \frac{1}{k+1}, \dots$, which converges to 0.

$$\text{So, } P(\{k\}) = 0.$$

$$\text{Now, } P(\mathbb{N}) = P(\mathbb{N}) = \sum_{k=1}^{\infty} P(\{k\})$$

$$= 0$$

using countable additivity, $\{k\}$ for $k=1, 2, \dots$ are disjoint.

$$\text{But } P(\mathbb{N}) = 1.$$

So, it is a contradiction.

Hence, $P(A)$ is not a probability measure on \mathcal{F} .