

Math 341 / 641 Fall 2023

Midterm Examination One Solutions

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Full Name _____

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Instructions

This exam is 110 minutes (variable time per question) and closed-book. You are allowed **one** page (front and back) of a “cheat sheet”, blank scrap paper (provided by the proctor) and a graphing calculator (which is not your smartphone). Please read the questions carefully. Within each problem, I recommend considering the questions that are easy first and then circling back to evaluate the harder ones. No food is allowed, only drinks.

Problem 1 We are trying to infer the mean number of daily (non-holiday) phone calls in a call center that operates Monday-Friday. Our data consists of $n = 10$ non-holiday business days of phone calls. Here are the number of calls on those days with some common summary statistics already computed for you:

$\mathbf{x} = \langle 14, 22, 22, 18, 22, 20, 19, 15, 23, 18 \rangle$ where $\bar{x} = 19.3$ and $\hat{\sigma}^2 = 8.61$ for this dataset.

- (a) [4 pt / 4 pts] Which sampling assumption makes more sense (I) the sample data was measured on sampled units from a population of size $N \gg n$ or (II) the sample data is realized from an iid DGP. Justify your answer.

“Number of phone calls in a one day period” don’t seem to be sampled from a population of one day periods in the same way that people are sampled from a population. It makes much more sense that “number of phone calls” are realized from a random process which may be identically distributed day-to-day and independent day-to-day. Thus, (II)

However, if you wish to say this call center was open for N business days, then the answer could just as easily be (I). The points were given solely based on if the assumption’s explanation made sense.

- (b) [4 pt / 8 pts] If the assumption was (I), describe under what circumstances the sample would be an SRS.

If all the days in the population of days were identified, then the order of these days was shuffled, then the first n days were drawn.

- (c) [4 pt / 12 pts] If the assumption was (I) and the sample was not an SRS, what problems would that cause eventually with inference?

Inference would not be possible for the parameters of the overall population, only the subpopulation that the sampling scheme was representative of.

- (d) [4 pt / 16 pts] Regardless of your previous answers, let's assume (II) and that the data was realized from the following DGP:

$$X_1, \dots, X_n \stackrel{iid}{\sim} \text{Poisson}(\theta) := \frac{\theta^x e^{-\theta}}{x!} \mathbb{1}_{x \in \mathbb{N}_0}$$

whose parameter space is $\Theta = (0, \infty)$ and $\mathbb{E}[X] = \theta$ and $\text{Var}[X] = \theta$.

For this DGP, find the method of moments estimator for the expectation. Your answer must be a function of X_1, \dots, X_n and n .

$$\hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n X_i$$

- (e) [5 pt / 21 pts] For this DGP, find the method of moments estimator for the variance. Your answer must be a function of X_1, \dots, X_n and n .

$$\hat{\mu}_1 = \frac{1}{n} \sum_{i=1}^n X_i \quad \text{or} \quad \hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$$

- (f) [3 pt / 24 pts] For general sample size n , find the likelihood function \mathcal{L} .

$$\mathcal{L}(\theta; X_1, \dots, X_n) = \prod_{i=1}^n \frac{\theta^{X_i} e^{-\theta}}{X_i!} = \frac{\theta^{\sum_{i=1}^n X_i} e^{-n\theta}}{\prod_{i=1}^n X_i!}$$

- (g) [3 pt / 27 pts] For general sample size n , find the log likelihood function ℓ .

$$\ell(\theta; X_1, \dots, X_n) = \ln(\theta) \sum_{i=1}^n X_i - n\theta - \sum_{i=1}^n \ln(X_i!)$$

- (h) [3 pt / 30 pts] For general sample size n , find the score function s .

$$s(\theta; X_1, \dots, X_n) := \ell'(\theta; X_1, \dots, X_n) = \frac{\sum_{i=1}^n X_i}{\theta} - n$$

- (i) [4 pt / 34 pts] For general sample size n , find the maximum likelihood estimator $\hat{\theta}^{\text{MLE}}$.

$$s(\theta; X_1, \dots, X_n) \stackrel{\text{set}}{=} 0 \Rightarrow \frac{\sum_{i=1}^n X_i}{\theta} - n = 0 \Rightarrow \frac{\sum_{i=1}^n X_i}{\theta} = n \Rightarrow \hat{\theta}^{\text{MLE}} = \bar{X}$$

- (j) [5 pt / 39 pts] Prove that $\hat{\theta}^{\text{MLE}}$ is unbiased.

$$\mathbb{E}[\bar{X}] = \frac{1}{n} \mathbb{E} \left[\sum_{i=1}^n X_i \right] = \frac{1}{n} n \mathbb{E}[X] = \frac{1}{n} n \mu = \mu$$

- (k) [4 pt / 43 pts] Provide a point estimate for θ and justify why you used that estimate.

$\bar{x} = 19.3$ since \bar{X} is the MLE (or MM or unbiased)

- (l) [3 pt / 46 pts] Find the variance of $\hat{\theta}^{\text{MLE}}$.

$$\mathbb{V}\text{ar}[\bar{X}] = \frac{1}{n^2} \mathbb{V}\text{ar} \left[\sum_{i=1}^n X_i \right] = \frac{1}{n^2} n \mathbb{V}\text{ar}[X] = \frac{1}{n^2} n \theta = \frac{\theta}{n}$$

- (m) [4 pt / 50 pts] Find $I_n(\theta)$ for this iid Poisson DGP.

$$I_n(\theta) = \mathbb{E}[-\ell''(\theta; X_1, \dots, X_n)] = \mathbb{E} \left[\frac{\sum_{i=1}^n X_i}{\theta^2} \right] = \frac{n \mathbb{E}[X]}{\theta^2} = \frac{n\theta}{\theta^2} = \frac{n}{\theta}$$

- (n) [3 pt / 53 pts] Prove that $\hat{\theta}^{\text{MLE}}$ is a UMVUE.

The CRLB result is that any unbiased estimator must have variance at least equal to $1/I_n(\theta)$. Since the variance of the MLE is θ/n by part (l) and $1/I_n(\theta) = n/\theta$ by part (m), the MLE achieves the CRLB and thus it is a UMVUE.

- (o) [7 pt / 60 pts] Provide an integral expression that computes the risk under L1 loss of $\hat{\theta}^{\text{MLE}}$ *approximately*. This is a challenging problem. You may want to leave it until you're done with the rest of the test. Justify each step for partial credit.

$$R(\hat{\theta}^{\text{MLE}}, \theta) = \mathbb{E} \left[\left| \hat{\theta}^{\text{MLE}} - \theta \right| \right] = \sum_{\bar{x} \in \mathbb{R}} |\bar{x} - \theta| p_{\bar{X}}(\bar{x}) \approx \int_{\mathbb{R}} |\bar{x} - \theta| \frac{1}{\sqrt{2\pi\theta/n}} e^{-\frac{1}{2\theta/n}(\bar{x}-\theta)^2} d\bar{x}$$

The approximation follows from the CLT which inserts the density of $\mathcal{N}(\theta, \theta/n)$ for the density of the maximum likelihood estimator. Without this approximation, the true PMF of \bar{X} would need to be determined (which is actually possible, but is much more difficult).

- (p) [4 pt / 64 pts] If the true mean number of daily calls is 19 or more, the company would like to hire more representatives. So it is important that the effect be real. Attempt to prove there are 19 or more mean daily calls. What are the two hypotheses?

$$H_a : \theta > \theta_0 = 19, \quad H_0 : \theta \leq \theta_0 = 19$$

- (q) [9 pt / 73 pts] Run this hypothesis test at $\alpha = 5\%$. Render a decision and provide a “conclusion sentence” but do not calculate a p-val. You can calculate the RET and test statistic on either the original scale or the standardized scale. Hint: you'll need the variance expression from part (l).

$$\begin{aligned} \bar{X} \mid H_0 &\sim \mathcal{N}(\theta_0, \sqrt{\theta_0/n^2}) = \mathcal{N}(19, 1.38^2) \\ RET &\approx \left(-\infty, \theta_0 + z_{1-\alpha} \times \sqrt{\theta_0/n} \right] = \left(-\infty, 19 + 1.65 \times 1.38 \right] = \left(-\infty, 21.28 \right] \end{aligned}$$

Since $\bar{x} = 19.3 \in RET \Rightarrow$ retain H_0 . Conclusion: there is not enough evidence to support the theory that there are more than 19 calls per day on average at this calling center.

- (r) [2 pt / 75 pts] Is the result statistically significant? Yes / no
- (s) [2 pt / 77 pts] Is it possible you made a Type I error? Yes / no
- (t) [2 pt / 79 pts] Is it possible you made a Type II error? Yes / no
- (u) [2 pt / 81 pts] Was this test an “exact test”? Yes / no

- (v) [3 pt / 84 pts] Regardless of the test's decision, explain how this test could have had more statistical power.

Increase the sample size n .

- (w) [2 pt / 86 pts] If n was very, very large, would this test have been rejected given the same point estimate? [Yes](#) / no
- (x) [2 pt / 88 pts] If (1) n was very, very large and (2) you had the same point estimate, and (3) you rejected the null hypothesis, would this result be clinically or practically significant? Yes / [no](#)
- (y) [6 pt / 94 pts] Calculate the p-val of this test as a function of Φ .

$$\hat{z} = (19.3 - 19)/1.38 = 0.22 \Rightarrow p_{val} \approx \mathbb{P}(Z > \hat{z}) = 1 - \mathbb{P}(Z < \hat{z}) = 1 - \Phi(0.22)$$

(For those curious, the answer is ≈ 0.59 but there was no need to compute it).

- (z) [6 pt / 100 pts] Construct a 95% confidence interval (CI) estimate for θ to two decimal places. Use $\hat{\sigma}^2$ as the best guess for the variance of X . Provide one interpretation of the CI you constructed.

$$\hat{CI}_{\theta,95\%} = [\bar{x} \pm z_{97.5\%} \times \text{SE}[\bar{X}]] \approx \left[19.3 \pm 2 \times \frac{\sqrt{8.61}}{\sqrt{10}} \right] = [17.44, 21.16]$$

Possible explanations:

- (I) Before this interval was constructed, $\mathbb{P}(\theta \in \hat{CI}_{\theta,95\%}) = 95\%$.
- (II) If many such datasets were collected and $\hat{CI}_{\theta,95\%}$ were constructed, approximately 95% of the constructed $\hat{CI}_{\theta,95\%}$'s would contain θ .
- (III) The probability $\hat{CI}_{\theta,95\%}$ contains θ is zero or one.