

lec 12 math 2A3

HW: prove $H-H_1$ is idempotent

we proved last time why

$$H_1 := \frac{1}{n} J_n, \quad H = X(X^T X)^{-1} X^T \text{ also}$$

$$\frac{\| \vec{\hat{y}} - X\vec{\beta} - \bar{y} \vec{1}_n + H_1 X \vec{\beta} \|^2}{\frac{\| \vec{e} \|^2}{n-p-1}} \sim F_{p, n-p-1}$$

this allows testing of any $H_0: \vec{\beta} = \vec{\theta}$

However, we are usually interested in the omnibus/global test of $H_0: \beta_1 = \beta_2 = \dots = \beta_p = 0$ which means none of the features have any linear predictive power.

Analyzing the numerator under H_0

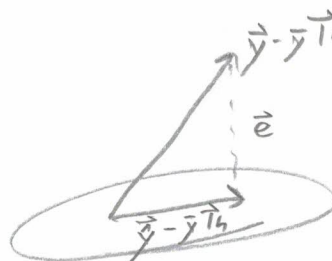
$$\begin{aligned} \left\| \vec{\hat{y}} - X \begin{bmatrix} \beta_0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} - \bar{y} \vec{1}_n + H_1 X \begin{bmatrix} \beta_0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} \right\|^2 &= \left\| \vec{\hat{y}} - (\beta_0 \vec{1}_n) - \bar{y} \vec{1}_n + H_1 (\beta_0 \vec{1}_n) \right\|^2 \\ &= \left\| \vec{\hat{y}} - \beta_0 \vec{1}_n - \bar{y} \vec{1}_n + \beta_0 \vec{1}_n \right\|^2 = \left\| \vec{\hat{y}} - \bar{y} \vec{1}_n \right\|^2. \end{aligned}$$

The estimator under H_0 is:

$$\text{let } MSR := \frac{SSR}{p}$$

$$\hat{F} = \frac{\frac{\| \vec{\hat{y}} - \bar{y} \vec{1}_n \|^2}{p}}{\frac{\| \vec{e} \|^2}{n-p-1}} \sim F_{p, n-p-1}. \quad \text{The test statistic is } \hat{F} = \frac{\frac{SSR}{p}}{\frac{SSE}{n-p-1}} = \frac{MSR}{MSE}$$

$$p_{\text{val}} = P(\hat{F} > \hat{F})$$



num of change levels of signal per feature to denominator

Relationship between omnibus F statistic and R^2 .

$$\frac{1}{F} = \frac{\frac{1}{n-p-1} \frac{SSR}{SSE}}{\frac{1}{p} \frac{SSR}{SST-SSR}} \Rightarrow \frac{1}{F} = \frac{p}{n} \frac{SST-SSR}{SSR}$$

$$\Rightarrow \frac{1}{F} = \frac{p}{n} \left(\frac{SST}{SSR} - 1 \right) \Rightarrow \frac{1}{F} = \frac{p}{n} \left(\frac{1}{R^2} - 1 \right) \Rightarrow \frac{1}{F} = \frac{p}{n R^2} - \frac{p}{n} \Rightarrow \frac{1}{n R^2} = \frac{1}{F} + \frac{p}{n}$$

$$\Rightarrow \frac{1}{R^2} = \frac{1}{p F} + 1 \Rightarrow \frac{1}{R^2} = \frac{n+p F}{p F} \Rightarrow R^2 = \frac{p F}{n+p F} \quad \text{eq. (1)}$$

$$\begin{aligned} F \rightarrow 0 &\Leftrightarrow R^2 \rightarrow 0 \\ F \rightarrow 1 &\Leftrightarrow R^2 \rightarrow 1 \end{aligned}$$

$$p+1 = n \Rightarrow R^2 = 1$$

MANUSO
transform of variable

$$\Rightarrow f_{R^2}(R^2) = \text{Beta}\left(\frac{p}{2}, \frac{n-p-1}{2}\right)$$

Cool fact!

Let's play this game again.

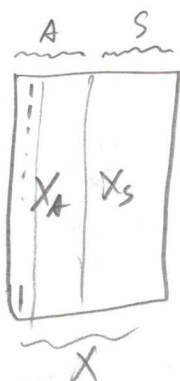
Let $S \subset \{1, 2, \dots, p\}$, let $A = \{0\} \cup S^c$
 \uparrow index of feature
 \uparrow index of feature column

Let $\vec{\beta}_S$ be the β_j 's
 s.t. $j \in S$

Let $\vec{\beta}_A$ be the β_j 's
 s.t. $j \in A$

rank $k = |S| \Rightarrow |A| = p+1-k$

Now, rearrange columns of design matrix so that:



Let $H_A = X_A(X_A^T X_A)^{-1} X_A^T$ be the proj. matrix onto $\text{colp}(X_A)$

Now $H = (H_A) + (H - H_A)$

$\vec{y} = H\vec{y} = (H_A + (H - H_A))\vec{y} = H_A\vec{y} + (H - H_A)\vec{y}$ Now show these orthogonal

$$(H_A\vec{y})^T (H - H_A)\vec{y} = \vec{y}^T H_A^T (H\vec{y} - H_A\vec{y}) = \vec{y}^T H_A^T H\vec{y} - \vec{y}^T H_A^T H_A\vec{y}$$

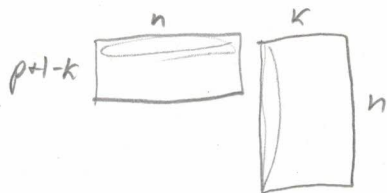
$$= \vec{y}^T H_A H\vec{y} - \|H_A\vec{y}\|^2 = \vec{y}^T H_A\vec{y} - \|H_A\vec{y}\|^2 = \vec{y}^T H_A^T H_A\vec{y} - \|H_A\vec{y}\|^2 = \|H_A\vec{y}\|^2 - \|H_A\vec{y}\|^2 = 0$$

$H_A H\vec{y} = \text{proj}_{X_A}(\text{proj}_X \vec{y}) \stackrel{\text{identity}}{=} \text{proj}_{X_A} \vec{y} = H_A\vec{y} \Rightarrow H_A H = H_A$

Proof. Let $X_A = Q_A R_A$, $X = QR$ Now $Q = [Q_A | Q_{-A}]$ by Gram-Schmidt Algorithm

$$H_A H = (Q_A Q_A^T) (Q Q^T) = Q_A Q_A^T [Q_A | Q_{-A}] \begin{bmatrix} Q_A^T \\ Q_{-A}^T \end{bmatrix} = Q_A Q_A^T (Q_A Q_A^T + Q_{-A} Q_{-A}^T)$$

$$= Q_A Q_A^T \overset{I_{p+1-k}}{Q_A Q_A^T} + Q_A Q_A^T Q_{-A} Q_{-A}^T = Q_A Q_A^T + 0_{n,n} = Q_A Q_A^T = H_A$$



Since each
 are orth.
 by GS,

$$Q_A^T Q_{-A} = 0_{p+1-k, k}$$

rank $[H_A] = p+1-k$

rank $[H - H_A] = k$

$$\frac{1}{\sigma^2} \vec{\varepsilon}^T \vec{\varepsilon} = \frac{1}{\sigma^2} \vec{\varepsilon}^T H_A \vec{\varepsilon} + \frac{1}{\sigma^2} \vec{\varepsilon}^T (I - H_A) \vec{\varepsilon} + \frac{1}{\sigma^2} \vec{\varepsilon}^T (I - H) \vec{\varepsilon}$$

by Cochran's

$$\Rightarrow \frac{1}{\sigma^2} \vec{\varepsilon}^T H_A \vec{\varepsilon} \sim \chi^2_{p+1-k} \text{ indep of } \frac{1}{\sigma^2} \vec{\varepsilon}^T (H - H_A) \vec{\varepsilon} \sim \chi^2_k \text{ indep of } \frac{1}{\sigma^2} \vec{\varepsilon}^T (I - H) \vec{\varepsilon} \sim \chi^2_{n-p+1}$$

$$\begin{aligned} \frac{1}{\sigma^2} \vec{\varepsilon}^T (H - H_A) \vec{\varepsilon} &= \frac{1}{\sigma^2} \vec{\varepsilon}^T (H - H_A)^T (H - H_A) \vec{\varepsilon} = \frac{1}{\sigma^2} \| (H - H_A) \vec{\varepsilon} \|^2 = \frac{1}{\sigma^2} \| (H - H_A) (Y - X\beta) \|^2 \\ &= \frac{1}{\sigma^2} \| HY - H_A Y - HX\beta + H_A X\beta \|^2 = \frac{1}{\sigma^2} \| \vec{y} - \vec{y}_A - X\beta + H_A X\beta \|^2 \\ &= \frac{1}{\sigma^2} \| \vec{y} - \vec{y}_A - [X_A \ X_S] \begin{bmatrix} \beta_A \\ \beta_S \end{bmatrix} + H_A [X_A \ X_S] \begin{bmatrix} \beta_A \\ \beta_S \end{bmatrix} \|^2 = \frac{1}{\sigma^2} \| \vec{y} - \vec{y}_A - \cancel{X_A \beta_A} - X_S \beta_S + \cancel{H_A X_A \beta_A} + H_A X_S \beta_S \|^2 \\ &= \frac{1}{\sigma^2} \| \vec{y} - \vec{y}_A + (I - H_A) X_S \beta_S \|^2 \end{aligned}$$

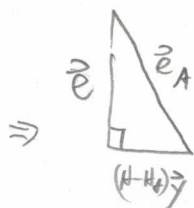
$$\Rightarrow \frac{\frac{1}{\sigma^2} \vec{\varepsilon}^T (H - H_A) \vec{\varepsilon}}{k} \sim F_{k, n-p+1} \Rightarrow \frac{\| \vec{y} - \vec{y}_A + (I - H_A) X_S \beta_S \|^2}{\frac{k}{\frac{\| \vec{\varepsilon} \|^2}{n-p+1}}} \sim F_{k, n-p+1}$$

This can be used to run tests of $H_0: \beta_S = \vec{0}$

Typically, the test of interest is $H_0: \beta_S = \vec{0}_k$

Using this H_0 , the estimator is:

$$\hat{F} = \frac{\frac{\| \vec{y} - \vec{y}_A \|^2}{k}}{\frac{\| \vec{\varepsilon} \|^2}{n-p+1}} \sim F_{k, n-p+1} \quad \uparrow \quad \frac{\frac{\| \vec{y}_A \|^2 - \| \vec{\varepsilon} \|^2}{k}}{\frac{\| \vec{\varepsilon} \|^2}{n-p+1}} \sim F_{k, n-p+1}$$



$$\Rightarrow \| \vec{y} \|^2 + \| (I - H_A) \vec{y} \|^2 = \| \vec{y}_A \|^2 \Rightarrow \| (I - H_A) \vec{y} \|^2 = \| \vec{y}_A \|^2 - \| \vec{y} \|^2$$

(Proof using algebra on H_A)

The test statistic is then:

$\Delta SSE \leftarrow$ is the addition of the K features "worth it"?

$$\hat{F} = \frac{\frac{SSE_A - SSE}{K}}{MSE} \quad \text{and} \quad p_{\text{val}} = P(\hat{F} > \hat{F})$$

SSE_A is sometimes called SSE of the "reduced model" as it has a reduced # of features $p+1-K$ relative to the "full model" which has all $p+1$ features. Note that the reduced model is nested in the full model.