

# Lec 10 MATH 343

$$\vec{y} = X\vec{\beta} + \vec{\epsilon}$$



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Assume  $\vec{\epsilon} \sim N_n(\vec{0}, \sigma^2 I_n)$

$\Rightarrow \epsilon_1, \dots, \epsilon_n$  iid  $N(0,1)$

ObP for mispec error + ignore em

random through  $\vec{\epsilon}$

fixed covariates  
unknown parameter

$$\begin{aligned} \Rightarrow \vec{\beta} &= (X^T X)^{-1} X^T \vec{y} \sim N_{p+1}(\vec{\beta}, \sigma^2 (X^T X)^{-1}) \Rightarrow \beta_j \sim N(\beta_j, \sigma^2 (X^T X)^{-1}_{jj}) \\ \Rightarrow \vec{y} &= X\vec{\beta} \\ \Rightarrow \vec{y} &= X\vec{\beta} + \vec{\epsilon} = \vec{\hat{y}} + \vec{\epsilon} \end{aligned}$$

$$\Rightarrow \frac{\beta_j - \hat{\beta}_j}{\sigma \sqrt{(X^T X)^{-1}_{jj}}} \sim N(0,1) \forall j$$

But  $\sigma$  is unknown!

ObP for  $\vec{\epsilon}$ , the residuals in the regression

left over from lec 9:

$$\Rightarrow \vec{\epsilon} = \vec{y} - X\vec{\hat{\beta}} = \underbrace{(I - H)}_{\text{geometrically}} \vec{y} = \underbrace{(I - H)(X\vec{\beta} + \vec{\epsilon})}_{\vec{0}_{(p+1) \times (p+1)}} = \underbrace{(I - H)X\vec{\beta}}_{\vec{0}_{(p+1) \times (p+1)}} + (I - H)\vec{\epsilon}$$

$$= (I - H)\vec{\epsilon} \sim N_n(\vec{0}, (I - H)\sigma^2 I_n (I - H)^T) = N_n(\vec{0}, \sigma^2 (I - H))$$

$$\Rightarrow \epsilon_i \sim N(0, \sigma^2 (1 - H_{ii}))$$

From MATH 340:  $\vec{U} \sim N_n(\vec{0}, \Sigma)$ ,  $\vec{A} \in \mathbb{R}^{n \times m}$ ,  $\vec{A} \in \mathbb{R}^{m \times n} \Rightarrow \vec{A}\vec{U} \sim N_m(\vec{0}, \vec{A}\Sigma\vec{A}^T)$

$$\vec{\hat{y}} = X\vec{\hat{\beta}} = H\vec{y} = H(X\vec{\beta} + \vec{\epsilon}) = HX\vec{\beta} + H\vec{\epsilon} = X\vec{\beta} + H\vec{\epsilon} \sim N_n(X\vec{\beta}, H\sigma^2 I_n H) = N_n(X\vec{\beta}, \sigma^2 H) \Rightarrow \hat{y}_i \sim N(\hat{y}_i, \sigma^2 H_{ii})$$

In 340 we showed that if  $\vec{Z} \sim N_n(\vec{0}, \sigma^2 I_n)$

$$\Rightarrow \vec{Z}^T \vec{Z} \sim \chi^2_n$$

$$\text{From last class: } \vec{Z}^T \vec{Z} = \underbrace{\frac{1}{\sigma^2} \vec{Z}^T H \vec{Z}}_{B_1} + \underbrace{\frac{1}{\sigma^2} \vec{Z}^T (I - H) \vec{Z}}_{B_2}$$

Since  $B_1 + B_2 = I_n$  and  $\text{rank}(B_1) + \text{rank}(B_2) = (p+1) + (n - (p+1)) = n$ , we can use Cochran's thm

$$\Rightarrow \frac{1}{\sigma^2} \vec{Z}^T H \vec{Z} \sim \chi^2_{p+1} \text{ indep of } \frac{1}{\sigma^2} \vec{Z}^T (I - H) \vec{Z} \sim \chi^2_{n - (p+1)}$$

last class

nearest all

$$\frac{1}{\sigma^2} \vec{\varepsilon}^T (\mathbf{I} - \mathbf{H}) \vec{\varepsilon} = \dots = \frac{1}{\sigma^2} \|\vec{\varepsilon}\|^2 \sim \chi^2_{n-(p+1)}$$

$$\frac{1}{\sigma^2} \vec{\varepsilon}^T \mathbf{H} \vec{\varepsilon} = \frac{1}{\sigma^2} \vec{\varepsilon}^T \mathbf{H} \mathbf{H} \vec{\varepsilon} = \frac{1}{\sigma^2} \vec{\varepsilon}^T \mathbf{H}^T \mathbf{H} \vec{\varepsilon} = \frac{1}{\sigma^2} (\mathbf{H} \vec{\varepsilon})^T (\mathbf{H} \vec{\varepsilon}) = \frac{1}{\sigma^2} \|\mathbf{H} \vec{\varepsilon}\|^2$$

$$= \frac{1}{\sigma^2} \|\mathbf{H}(\vec{y} - \mathbf{X}\vec{\beta})\|^2 = \frac{1}{\sigma^2} \|\mathbf{H}\vec{y} - \mathbf{H}\mathbf{X}\vec{\beta}\|^2 = \frac{1}{\sigma^2} \|\vec{y} - \mathbf{X}\vec{\beta}\|^2$$

$$= \frac{1}{\sigma^2} \|\mathbf{X}\vec{\beta} - \mathbf{X}\vec{\beta}\|^2 = \frac{1}{\sigma^2} \|\mathbf{X}(\vec{\beta} - \vec{\beta})\|^2 \sim \chi^2_{p+1}$$

matrix estimator error

$$\text{As } n \rightarrow \infty \quad \vec{\varepsilon} \rightarrow \vec{\varepsilon}, \quad \vec{\beta} \rightarrow \vec{\beta} \Rightarrow \|\mathbf{X}(\vec{\beta} - \vec{\beta})\|^2 \rightarrow 0$$

independent + ignore error

estimator error dropping

$$\Rightarrow \vec{\varepsilon} \text{ and } \vec{\beta} \text{ are independent} \Rightarrow \|\vec{\varepsilon}\|^2 \text{ and } \vec{\beta} \text{ are independent}$$

Consider the following estimator

from last class

$$\frac{\frac{b_j - \beta_j}{\sigma \sqrt{(\mathbf{X}^T \mathbf{X})^{-1}_{jj}}}}{\sqrt{\frac{\frac{1}{\sigma^2} \|\vec{\varepsilon}\|^2}{n-(p+1)}}} \sim T_{n-(p+1)}$$

we just proved this

Student's T test

$$= \frac{\frac{b_j - \beta_j}{\sigma \sqrt{(\mathbf{X}^T \mathbf{X})^{-1}_{jj}}}}{\frac{1}{\sigma} \sqrt{\frac{\|\vec{\varepsilon}\|^2}{n-(p+1)}}} \sim T_{n-(p+1)}$$

estimate  
the  $\sigma$  standard is then

$$\frac{b_j - \beta_j}{\sqrt{(\mathbf{X}^T \mathbf{X})^{-1}_{jj}}} = \frac{b_j - \beta_j}{s_e \sqrt{(\mathbf{X}^T \mathbf{X})^{-1}_{jj}}}$$

$$\text{AME} \rightarrow \left( \frac{\text{SSE}}{n-(p+1)} \right)$$

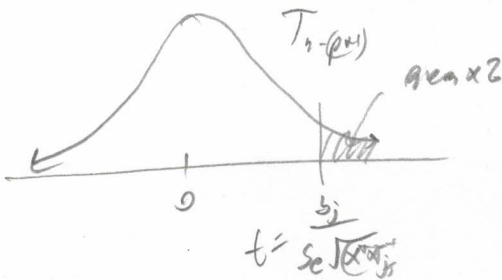
Assume level of test is  $\alpha$ .  
 Usually the inference is for  $H_0: \beta_j = 0 \Leftrightarrow H_1: \beta_j \neq 0$  why?  
 Because if  $\beta_j = 0$ , the job feature doesn't affect the response (marginally)  
 (which usually means the causal relationship is not too interesting).

In this case,

$$\frac{\hat{\beta}_j}{\text{se} \sqrt{X'X}^{-1}_{jj}} \in \left[ -t_{1-\frac{\alpha}{2}, n-p-1}, t_{1-\frac{\alpha}{2}, n-p-1} \right]$$

interesting  
the test  $\Rightarrow CI_{\beta_j, 1-\alpha} = \left[ \hat{\beta}_j \pm t_{1-\frac{\alpha}{2}, n-p-1} \cdot \text{se} \sqrt{X'X}^{-1}_{jj} \right]$

How to get p-value?



p-value =  $2P(|T_{n-p-1}| > \left| \frac{\hat{\beta}_j}{\text{se} \sqrt{X'X}^{-1}_{jj}} \right|)$

Consider a new  $\vec{x}_*$  and you want to predict its expected response  $h^*(\vec{x}_*) = \vec{x}_* \vec{\beta}$ . This is the average of many predictions of  $\vec{x}_*$  since the noise mean is zero.

Let  $\mu_* := \vec{x}_* \vec{\beta}$

$$\hat{Y}_* = \vec{x}_* \vec{B} \Rightarrow \hat{Y}_* \sim N(\underbrace{\vec{x}_* \vec{\beta}}_{\mu_*}, \sigma^2 \vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T)$$

$$\vec{B} \sim N(\vec{\beta}, \sigma^2 (\mathbf{X}^T \mathbf{X})^{-1}) \Rightarrow$$

$\Downarrow$

$$\frac{\hat{Y}_* - \mu_*}{\sigma \sqrt{\vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T}} \sim N(0, 1)$$

Again,  $\sigma$  is unknown so consider

$$\frac{\hat{Y}_* - \mu_*}{\sigma \sqrt{\vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T}} \sim N(0, 1) \quad \text{since } \hat{Y}_* \text{ is a function of } \vec{B}$$

$\uparrow$   
indep

$$\frac{\hat{Y}_* - \mu_*}{\sqrt{\frac{\frac{1}{n-1} \|\mathbf{E}\|^2}{n-1}}} \sim T_{n-(p+1)}$$

$$\frac{\hat{Y}_* - \mu_*}{\sqrt{\frac{\vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T}{\frac{1}{n-1} \|\mathbf{E}\|^2}}} \sim T_{n-(p+1)}$$

the test statistic for  $H_0: \mu_* = 0$  is

$$T = \frac{\hat{Y}_* - 0}{s_e \sqrt{\vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T}} \in [\pm t_{1-\frac{\alpha}{2}, n-(p+1)}] \Rightarrow \text{Reject } H_0$$

which can be used to test. Inverting the test yields

$$CI_{\mu_*, 1-\alpha} = \left[ \hat{Y}_* \pm t_{1-\frac{\alpha}{2}, n-(p+1)} \cdot s_e \sqrt{\vec{x}_* (\mathbf{X}^T \mathbf{X})^{-1} \vec{x}_*^T} \right]$$