

$Y = g(X)$, $f_Y(y)$ known, what is $f_X(x)$?

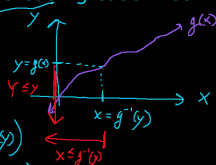
Let's consider g to be a 1:1 function (i.e. it's strictly increasing or strictly decreasing). Let's do the strictly increasing case first:

Let's try first to derive the CDF and then to get the PDF take its derivative:

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{d}{dy} [F_X(g^{-1}(y))] = f_X(g^{-1}(y)) \cdot \frac{d}{dy} [g^{-1}(y)] = f_X(g^{-1}(y)) \cdot \left| \frac{d}{dy} [g^{-1}(y)] \right|$$

$\frac{d}{dx} [g(x)] > 0 \Leftrightarrow \frac{d}{dy} [g^{-1}(y)] > 0$

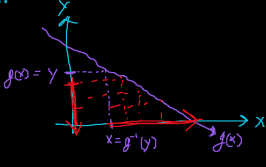


Let's consider the case where g is strictly decreasing and do the same: find the CDF and take its derivative.

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \geq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{d}{dy} [1 - F_X(g^{-1}(y))] = -f_X(g^{-1}(y)) \cdot \frac{d}{dy} [g^{-1}(y)] = f_X(g^{-1}(y)) \cdot \left| \frac{d}{dy} [g^{-1}(y)] \right|$$

$\frac{d}{dx} [g(x)] < 0 \Leftrightarrow \frac{d}{dy} [g^{-1}(y)] < 0 \Rightarrow -\frac{d}{dy} [g^{-1}(y)] = \left| \frac{d}{dy} [g^{-1}(y)] \right|$



We have the same formula for the PDF in both cases for g , thus it is the general formula.

Let's derive some simple consequences of this formula that are very powerful:

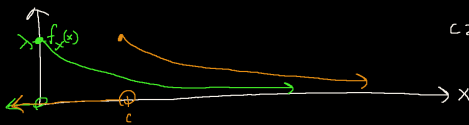
$Y = g(X) = aX + c \sim ?$ (this is called a shifted and scaled rv)

$$y = ax + c \Rightarrow y - c = ax \Rightarrow x = \frac{y - c}{a} = g^{-1}(y) \Rightarrow \left| \frac{d}{dy} [g^{-1}(y)] \right| = \left| \frac{1}{a} \right| = \frac{1}{|a|}$$

$$f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y - c}{a}\right)$$

$$Y = -X \Rightarrow f_Y(y) = f_X(-y), \quad Y = aX \Rightarrow f_Y(y) = \frac{1}{|a|} f_X\left(\frac{y}{a}\right), \quad Y = X + c \Rightarrow f_Y(y) = f_X(y - c)$$

$$X \sim \text{Exp}(\lambda) = \lambda e^{-\lambda x} \mathbb{1}_{x \in (0, \infty)}, \quad Y = X + c \sim \lambda e^{-\lambda(y - c)} \mathbb{1}_{y - c \in (0, \infty)} = e^{\lambda c} \lambda e^{-\lambda y} \mathbb{1}_{y \in (c, \infty)}$$



$$X \sim \text{Exp}(1) = e^{-x} \mathbb{1}_{x \in (0, \infty)}$$

$$Y = g(X) = -\ln\left(\frac{e^{-x}}{1 - e^{-x}}\right) = \ln\left(\frac{1 - e^{-x}}{e^{-x}}\right) = \ln(e^x - 1) \sim f_Y(y) = ?$$

$$y = \ln(e^x - 1) \Rightarrow e^y = e^x - 1 \Rightarrow e^y + 1 = e^x = \ln(e^y + 1) = x = g^{-1}(y)$$

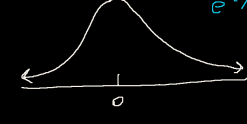
$$\left| \frac{d}{dy} [\ln(e^y + 1)] \right| = \left| \frac{e^y}{e^y + 1} \right| = \frac{e^y}{e^y + 1} = \frac{1}{1 + e^{-y}} \in (0, 1)$$

$$f_Y(y) = f_X(\ln(e^y + 1)) \cdot \frac{1}{1 + e^{-y}} = e^{-\ln(e^y + 1)} \mathbb{1}_{\ln(e^y + 1) \in (0, \infty)} \cdot \frac{e^y}{e^y + 1}$$

$$= \frac{1}{e^y + 1} \cdot \frac{e^y}{e^y + 1} = \frac{e^y}{(e^y + 1)^2} = \text{Logistic}(0, 1)$$

$\ln(e^y + 1) \in (0, \infty) \Leftrightarrow e^y + 1 \in (1, \infty) \Leftrightarrow e^y \in (0, \infty) \Leftrightarrow y \in \mathbb{R}$

no need for indicator function!



This "standard logistic" rv looks very similar to a standard normal rv but it has slightly thicker tails. Used in chess ratings (Elo), deep learning, implicitly used in "logistic regression".

$$X \sim \text{Logistic}(0, 1), \sigma > 0, Y = \mu + \sigma X \sim \text{Logistic}(\mu, \sigma) = ?$$

$$f_Y(y) = \frac{1}{\sigma} \frac{e^{-\frac{y - \mu}{\sigma}}}{(e^{-\frac{y - \mu}{\sigma}} + 1)^2}, \quad E[Y] = \mu, \quad \text{SD}[Y] = \sigma \frac{\pi}{\sqrt{3}} \approx 1.8 \sigma$$

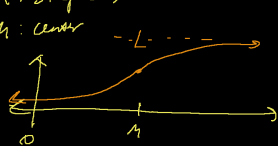
$$X \sim U(0, 1), \quad Y = -\ln\left(\frac{1}{X} - 1\right) \sim \text{Logistic}(0, 1) \quad (\text{HW})$$

Why is it called "logistic"?

$$l(x) = \frac{L}{1 + e^{-k(x - \mu)}} \in (0, L)$$

L : max value,
 k : steepness,
 μ : center

$L = 1, k = 1, \mu = 0$
 $\Rightarrow \frac{1}{1 + e^{-x}} = \frac{e^x}{1 + e^x}$



This function $l(x)$ is very useful and it's used to model population changes.

this is the std. logistic funct.

$$Y \sim \text{Logistic}(0, 1)$$

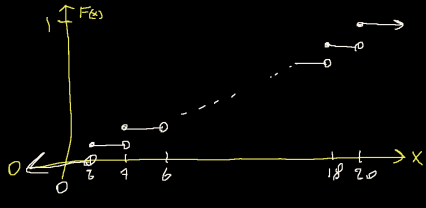
$$F_Y(y) = \int_{-\infty}^y \frac{e^t}{(1 + e^t)^2} dt = \int_{u=1}^{u=1+e^y} \frac{1}{u^2} \cdot \frac{1}{u} du = \left[-\frac{1}{u} \right]_1^{1+e^y} = 1 - \frac{1}{1+e^y} = \frac{e^y}{1+e^y} = \frac{1}{1+e^{-y}}$$

$$\text{let } u = 1 + e^t \Rightarrow \frac{du}{dt} = e^t \Rightarrow dt = \frac{1}{e^t} du = \frac{1}{u-1} du, \quad t = y \Rightarrow u = 1 + e^y, \quad t = -\infty \Rightarrow u = 1 + e^{-\infty} = 1$$

New concept. Solve for min x s.t. $q \leq P(X \leq x)$ i.e. $q \leq F(x)$. This x is called the " q th quantile" or " $100q$ th percentile". If $q = 0.5$ that x is called the "median" i.e. $\text{Med}[X]$. Otherwise the notation is $Q[X, q]$, the quantile operator.

$$X \sim U\left(\frac{1}{2}, 4, 6, \dots, 20\right) = \frac{1}{10} \mathbb{1}_{x \in \dots}$$

x	$p(x)$	$F(x)$
2	0.1	0.1
4	0.1	0.2
6	0.1	0.3
\vdots	\vdots	\vdots
18	0.1	0.9
20	0.1	1.0



$$Q[X, 0.3] = 6$$

6 is the 30%ile of the X distribution

$$Q[X, 0.9] = 18$$

$$Q[X, 0.85] = 18 \text{ is the min } x \text{ s.t. } F(x) \geq 0.85$$

$$\exists x \text{ s.t. } F(x) = 0.85$$

If X was a continuous rv with "contiguous support" meaning basically no gaps e.g. $[0, 10]$, $(0, \infty)$, \mathbb{R} but not $[0, 1] \cup [8, 9]$ then the min x s.t. $F(x) \geq q$ is where $F(x) = q$ i.e. $x = F^{-1}(q)$ which is then called the "quantile function".

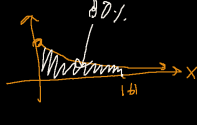
$$\text{Let } X \sim \text{Exp}(\lambda) = \lambda e^{-\lambda x} \Rightarrow F(x) = 1 - e^{-\lambda x} = q$$

$$\Rightarrow 1 - q = e^{-\lambda x} \Rightarrow \ln(1 - q) = -\lambda x \Rightarrow \ln\left(\frac{1}{1 - q}\right) = \lambda x$$

$$\Rightarrow x = \frac{1}{\lambda} \ln\left(\frac{1}{1 - q}\right) = F_X^{-1}(q) = Q[X, q]$$

$$\text{If } \lambda = 1, \quad F_X^{-1}(0.8) = \ln(5) \approx 1.61$$

$$\text{Med}[X] = F_X^{-1}\left(\frac{1}{2}\right) = \ln(2)$$



The quantile function is rarely in closed form e.g.

$T \sim \text{Erlang}(k, \lambda) \Rightarrow F(t) = P(k, \lambda t)$ whose inverse is not available in closed form. So... you use a computer to solve for t where $q = P(k, \lambda t)$.