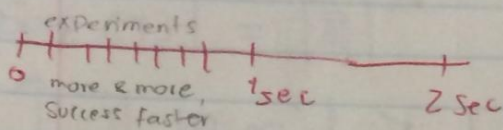


$$T \sim \text{Geom}(p) := (1-p)^{t-1} p, \quad F(t) = 1 - (1-p)^t, \quad E[t] = \frac{1}{p}$$



What can t be?

$$(0, \infty) \neq \text{Supp}(T)$$

experiments

Within each second, n experiments

$$P(t) = (1-p)^{nt-1} p, \quad F(t) = 1 - (1-p)^{nt}, \quad E[t] = \frac{1}{pn}$$

$$= \frac{1}{p} \exp. \frac{1}{n} \frac{\text{sec}}{\exp}$$

$$\lambda := pn$$

\nearrow small \nearrow big

$$\lim_{n \rightarrow \infty} P(t) = \left(1 - \frac{\lambda}{n}\right)^{nt-1} \frac{\lambda}{n} = \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{nt} \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{-1} \lim_{n \rightarrow \infty} \frac{\lambda}{n}$$

$$= \left(\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n \right)^t \underbrace{\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{-1}}_1 \underbrace{\lim_{n \rightarrow \infty} \frac{\lambda}{n}}_0$$

\nearrow

is this a PMF? this doesn't have a PMF

$$\sum_T P^x(t) = 0 \neq 1$$

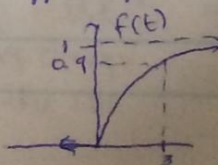
Not a PMF

$$F^x(t) = \lim_{n \rightarrow \infty} 1 - \left(1 - \frac{\lambda}{n}\right)^{nt}$$

$$= 1 - \left(\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n \right)^t$$

$$= 1 - e^{-\lambda t}$$

yes, it is a CDF



$$e := \lim_{n \rightarrow \infty} \left(1 + \frac{1}{n}\right)^n$$

$$e := \sum_{i=0}^{\infty} \frac{1}{i!}$$

$$e := \int_1^e \frac{1}{x} dx = 1$$

n	$f(n)$
10	2.519
100	2.705
1000	2.717
10,000	2.718

$$\lim_{n \rightarrow \infty} \left(1 + \frac{a}{n}\right)^n = e^a, \underline{a \in \mathbb{R}}$$

$$\text{let } m = \frac{n}{a} \Rightarrow m = ma$$

$$\Rightarrow \frac{1}{n} = \frac{a}{n}$$

$$\left(\lim_{n \rightarrow \infty} \left(1 + \frac{1}{n}\right)^m \right)^a = e^a$$

$$F(0) = 1 - e^{-\lambda(0)} = 1 - 1 = 0$$

$$\lim_{t \rightarrow \infty} F(t) = 1 - \lim_{t \rightarrow \infty} \frac{1}{(e^\lambda)^t} = 1 - 0 = 1 \checkmark$$

$$\int 1 - e^{-\lambda t}$$

Why it should be 1?

*Prove through the geometric series...

exam

$$\begin{cases} p \in (0, 1) \\ n \in \mathbb{N} \end{cases}$$

$$F'(t) = \lambda e^{-\lambda t} = \frac{\lambda}{e^{\lambda t}} > 0$$

$$\lambda > 0$$

monotonically increasing function.

$$F(x) := P(X \leq x) \leftarrow \text{CDF}$$

$$\frac{1 - e^{-\lambda t}}{\text{CDF}}$$

$$|\text{Supp}[X]| \leq |N|$$

$$|\text{Supp}[T]| = |\mathbb{R}|^\wedge$$

$$| (0, \infty) | > | \{1, 2, 3, \dots\} |$$

Definition of a continuous r.v.

They do not have PMF's

$$P^*(t) = 0$$

$$P(3) = 0$$

at 3sec when experiments are happening continuously
Probability you stop a $\underbrace{3.0000}_{\infty \text{ 0's}}$ Sec

$$P(3.0000\dots) = 0$$

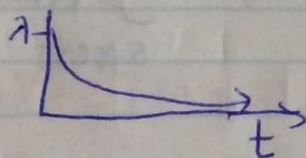
$$P(3.0000)$$

↙ btw those #'s

$$P(T \in [2.999995, 3.00004])$$

$$= F(3.00004) - F(2.99995) > 0$$

$$f(t) := f'(t) = \lambda e^{-\lambda t}$$



Abstract Metric

Density of Probability
in the r.v.

Prob. density function (PDF) rate change $\frac{\text{time}}{\text{support}}$
CDF is the King

$$\lambda = 1$$

$$f(3) = 0.05 \neq P(3) = 0$$

rate of change
of the CDF

$$P(T \in [a, b])$$

$$a, b \in \text{Supp}(T) \\ b > a$$

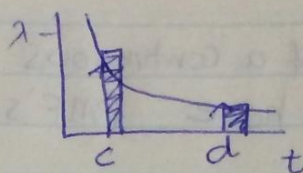
$$F(b) - F(a) = \int_a^b f(t) dt$$

↑
antiderivative

Fundamental theory of calculus

$$\frac{f(c)}{f(d)}$$

how much likely
is c or d to
happen...



$$P(T \in (-\infty, \infty)) = \lim_{t \rightarrow \infty} F(t) - \lim_{t \rightarrow -\infty} F(t) \\ = \int_{-\infty}^{\infty} f(t) dt = 1$$

* why from $-\infty$ if
not negative time?

Def:

X is a continuous r.v. if

- 1 CDF exists
- 2 PMF does not exist
- 3 $\text{Supp}[X] \subseteq \mathbb{R}$
- 4 $f(x)$ exists and
- 5 $|\text{Supp}[X]| = |\mathbb{R}|$

$$\int_{\text{Supp}[X]} f(x) dx$$

$$X \stackrel{d}{=} Y$$

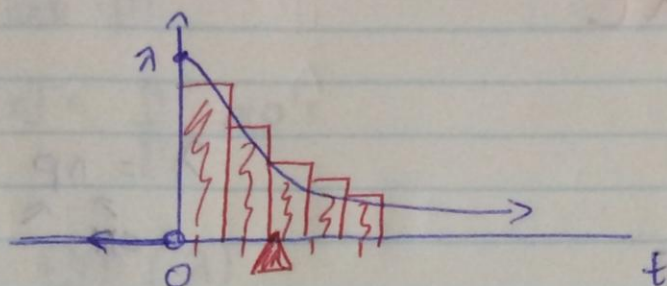
$$\text{If } F_X(x) = F_Y(y)$$

$$\text{or } F_X(x) = F_Y(y)$$

$$E[X]$$

$$= \sum x p(x)$$

probability
weighted avg.



$$\sum x \text{rect}(x)$$

:

$$E[X] = \int x f(x) dx$$

Supp[X]

\int = sum over uncountable things

$$E[g(x)] = \int g(x) f(x) dx$$

Supp[X]

$$\text{Var}[X] = \int (x - \mu)^2 f(x) dx$$

$x \in \text{Supp}[X]$

$$E[aX + c] = a E[X] + c$$

$$\text{Var}[aX + c] = a^2 \sigma^2$$

$$\text{SE}[aX + c] = |a| \sigma$$

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = n\mu$$

\hookrightarrow if identically distributed

$$\text{Var}\left[\sum X_i\right] = \sum \text{Var}[X_i] = n\sigma^2 \text{ if iid}$$

\uparrow

if X_1, \dots, X_n are independent

no longer PMF
only CDF

$$X \sim \text{Exp}(\lambda) \quad \text{pdf} \quad f(x) = \lambda e^{-\lambda x}$$

$$\text{Supp}[X] = (0, \infty)$$

Parameter space

$$\lambda \in (0, \infty)$$

$$\lambda = np$$

$n \in \mathbb{N}$ $p \in (0,1)$

$$\int_0^\infty 1 = \infty$$

idea

Solve to see when the mean is

$$E[X] = \int_0^\infty x \lambda e^{-\lambda x} dx$$

$$u = x$$

$$du = dx$$

$$dv = e^{-\lambda x} dx$$

$$V = \int dv = -\frac{1}{\lambda} e^{-\lambda x}$$

$$\int u dv = uv - \int v du$$

$$\int u dv = \int -\frac{1}{\lambda} e^{-\lambda x} dx = -\frac{1}{\lambda^2} e^{-\lambda x}$$

$$E[X] = \lambda \int_0^\infty x e^{-\lambda x} dx = \lambda \left[-\frac{x}{\lambda} e^{-\lambda x} - \frac{1}{\lambda^2} e^{-\lambda x} \right]_0^\infty$$

$$= \left[\frac{1}{\lambda} e^{-\lambda x} + x e^{-\lambda x} \right]_0^\infty = 0 + \left[\frac{1}{\lambda} \right]$$

$$X \sim \text{Geo}(p)$$

$$E[X] = \frac{1}{p} \text{ exp's}$$

$$X \sim \text{Exp}(\lambda)$$

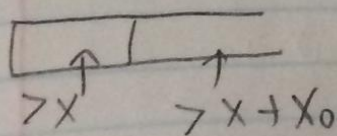
$$E[X] = \frac{1}{\lambda} \text{ Sec}$$

	Single	multiple
Discrete	Geo	Neg Bin
Continuous	Exp	<u>Erlang</u>

	PMF	PDF	CDF
Discrete	✓		✓
Continuous		✓	✓

$$P(X > x_0 + x \mid X > x)$$

$$= \frac{P(X > x_0 + x)}{P(X > x)} = \frac{e^{-\lambda(x_0 + x)}}{e^{-\lambda x}} = e^{-\lambda x_0}$$



Memoryless
same idea as
geometric