

Paradox trials

Lecture 7

9/15/16

W	1
L	1 2 3

$$P(\text{Ace in a deck of cards}) = \frac{4}{52} = \frac{1}{13}$$

$$P(\text{Ace given that we have a heart}) = P(A|\heartsuit) = \frac{1}{13}$$

of 13 hearts total one is an Ace

	A
♥	2
	3
	⋮
	K

Information Known  
This information is "irrelevant"

$$P(\text{IBM stock } \uparrow \text{ in a day}) = \text{let's say } \frac{1}{2}$$

$$= P(\text{IBM stock } \uparrow \text{ in a day} \mid \text{rains in Buenos Aires}) = \frac{1}{2}$$

↑ irrelevant information

Def. of A, B being "independent" (probabilistic independence)

$$P(A) = P(A|B)$$

equivalent to

$$P(B) = P(B|A)$$

Under A, B being independent

$$P(A) = P(A|B) = \frac{P(A \cap B)}{P(B)}$$

 $A_1, A_2, \dots$  independent

$$\Rightarrow P(A \cap B) = P(A) \cdot P(B)$$

Multiplication rule

$$P(\bigcap_{i=1}^{\infty} A_i) = \prod_{i=1}^{\infty} P(A_i)$$

$$P(H_2 | H_1) = P(H_2) = \frac{1}{2} = \frac{1}{2} \prod_{i=1}^5 P(H_i)$$

$$P(H_1, H_2, H_3, H_4, H_5) = \left(\frac{1}{2}\right)^5 = \frac{1}{2^5} = \frac{1}{32}$$

$$\begin{aligned}
 & P(\{ \geq 1 \text{ 6-6 in 24 rolls of 2 dice} \}) < \frac{1}{2} \\
 & = P(1 \text{ 6-6}) + P(2 \text{ 6-6}) + \dots + P(24 \text{ 6-6}) \\
 & \quad \text{P of 1 double 6's} + \text{P of 2 double 6's} + \dots + \text{P of 24 double 6's} \\
 & = 1 - P(\text{zero double 6 in 24 rolls}) \quad \rightarrow \quad = 1 - P(\text{zero 6-6}) = \\
 & \quad \text{Complement of this} \\
 & = 1 - P(\text{Not 6-6 1st roll} \cap \text{Not 6-6 2nd roll} \cap \dots \cap \text{Not 6-6 24th roll}) \\
 & \quad \rightarrow P(\text{Not 6-6 1st}) \cdot P(\text{Not 6-6 2nd}) \cdot \dots \cdot P(\text{Not 6-6 24th}) \\
 & = P(\text{Not 6-6})^{24} = (1 - P(6-6))^{24} = (1 - P(6)P(6))^{24} \\
 & \quad = (1 - (\frac{1}{6})^2)^{24} \\
 & \quad = 0.4914039
 \end{aligned}$$

If  $P(A|B) \neq P(A)$   
 or  $P(B|A) \neq P(B)$   
 or  $P(AB) \neq P(A)P(B)$   
 $\Rightarrow A, B$  Not independent i.e. "dependent"

Marginal Probability

$$\begin{aligned}
 P(\text{Q64 bus is late}) & < P(\text{Q64 bus is late} \mid \text{rain}) \\
 & > P(\text{Q64 bus is late} \mid \text{Sun and no traffic})
 \end{aligned}$$

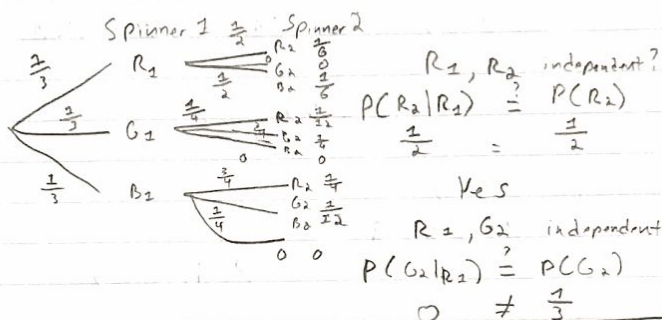
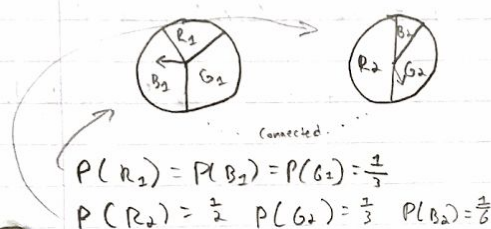
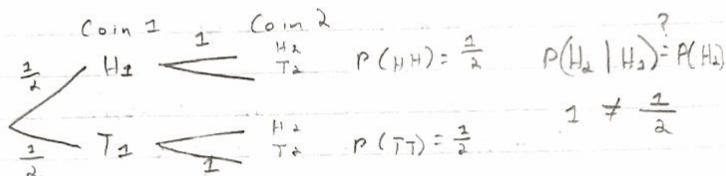
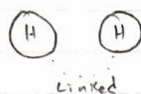
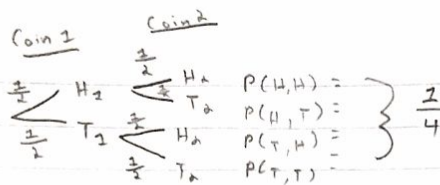
$$\begin{aligned}
 & P(\text{lung cancer} \mid \text{Smoking}) > P(\text{lung cancer}) \\
 & \quad \downarrow \\
 & \quad \text{dependent events}
 \end{aligned}$$

$A, B$  disjoint  $\stackrel{?}{\Rightarrow} A, B$  independent? No  
 (does this imply)

$$\begin{aligned}
 P(A|B) & \stackrel{?}{=} P(A) \\
 0 & \neq P(A)
 \end{aligned}$$

$$\begin{aligned}
 P(H \mid \text{coin loaded tails}) & \neq P(H) \\
 0 & \neq P(H)
 \end{aligned}$$

9/15/16



$$P(\text{Shared Birth day}) = P(\geq 1 \text{ Shared Birth day})$$

$$= P(1 \text{ shared Bday}) + P(2 \text{ shared Bday}) + \dots + P(42 \text{ shared bday})$$

$$= 1 - P(\text{No shared Bdays})$$

$$= 1 - 0.04 = \boxed{96\%}$$

$$= \frac{365 \cdot 364 \cdot 363 \cdot \dots \cdot 365 - 42 + 1}{365^{42}} = \frac{365 P_{42}}{365^{42}} = 0.039$$

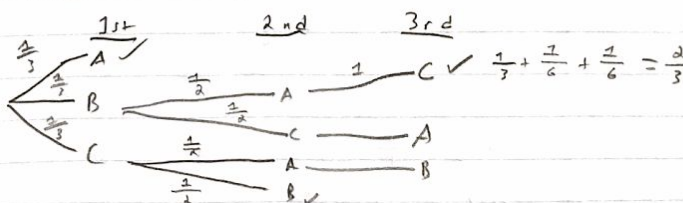
N people randomly administer their hats

$$P = P(\text{zero people get their hat})$$

$$1 - P = P(\text{at least one person gets hat})$$

$$= P(1 \text{ p gets hat}) + P(2 \text{ p get hat}) + \dots + P(n \text{ people get hat back})$$

3 people A, B, C



Let  $A_1$ : 1st person gets hat  
 $A_2$ : 2nd person gets hat  
 $\vdots$   
 $A_n$ : nth person gets hat

$$P(A \cup B) = P(A) + P(B) - P(AB)$$

$$P(A_1 \cup A_2) = \sum_{i=1}^2 P(A_i) - P(\bigcap_{i=1}^2 A_i)$$

$$P(\bigcup_{i=1}^n A_i) = \sum_{i=1}^n P(A_i) - \sum_{i \neq j} P(A_i \cap A_j) + \sum_{i \neq j \neq k} P(A_i \cap A_j \cap A_k) - \dots + (-1)^{n+1} P(\bigcap_{i=1}^n A_i)$$

Probability  
1st person  
gets hat

$$P(A_1) = \frac{1}{n}$$

$$P(A_2) = \frac{1}{n}$$

$$\sum_{i=1}^n P(A_i) = 1$$

get hat

$$P(A_1 \cap A_2) = \frac{1}{n!} \frac{n-2}{n-1} \frac{n-3}{n-2} \dots \frac{1}{2} = \frac{(n-2)!}{n!}$$

$$P(A_1 \cap A_3) = \frac{1}{n!} \frac{n-2}{n-1} \frac{n-3}{n-2} \dots \frac{1}{2} = \frac{(n-2)!}{n!}$$

$$\sum_{i \neq j} P(A_i \cap A_j) = \sum_{i \neq j} \frac{(n-2)!}{n!} = n C_2 \frac{(n-2)!}{n!} = \binom{n}{2} \frac{(n-2)!}{n!} = \frac{n!}{(n-2)! 2!} \frac{(n-2)!}{n!} = \frac{1}{2!}$$

$$P(A_1 \cap A_2 \cap A_3) = \frac{1}{n!} \frac{n-3}{n-1} \frac{n-4}{n-2} \dots \frac{1}{2} = \frac{(n-3)!}{n!}$$

$$\sum_{i \neq j \neq k} \frac{(n-3)!}{n!} = \binom{n}{3} \frac{(n-3)!}{n!} = \frac{n!}{(n-3)! 3!} \frac{(n-3)!}{n!} = \frac{1}{3!}$$

$$f(x) = \sum_{i=0}^{\infty} \frac{f^{(i)}(c)}{i!} (x-c)^i \quad \forall c \in \mathbb{R}, f \text{ cont.}$$

if  $x \approx 0 \Rightarrow c=0$

$$f(x) = \sum_{i=0}^{\infty} \frac{f^{(i)}(c_0)}{i!} x^i \approx f(c_0) + f'(c_0)x + \frac{f''(c_0)}{2!}x^2$$

$$e^x = e^0 + e^0 x + \frac{e^0 x^2}{2!} + \dots$$

$$= 1 + x + \frac{x^2}{2!} + \dots$$

$$e^{-1} = 1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} - \dots$$

$$1 - e^{-1} = 1 - ( ) = 1 - \frac{1}{2!} + \frac{1}{3!} - \dots$$

$$1 - e^{-1} \approx 0.368$$

$$p(\text{Zero ppl get hat}) = e^{-1} \approx 0.368 \approx \frac{1}{e}$$

$\uparrow$   $p(\text{at least one person gets hat})$