

41/27/16

Definition: Moment Generating Function

$$M_X(t) := E[e^{tX}], \quad e^{tX} = 1 + tX + \frac{t^2 X^2}{2!} + \frac{t^3 X^3}{3!} + \dots$$

① I. $M_X(t) = M_Y(t) \Rightarrow X \stackrel{d}{=} Y$ Taylor series expansion

② II. $E[X^k] = M_X^{(k)}(0)$

③ III. $\text{if } Y = aX + c \Rightarrow M_Y = e^{tc} M_X(at)$ if iid

④ IV. $\text{if } X, Y \text{ independent, } M_{X+Y}(t) = M_X(t) M_Y(t) = (M_Y(t))^2$

$X \sim \text{Bern}(p) \Rightarrow M_X(t) = 1 - p + pe^t$

$X \sim \text{Exp}(\lambda) \Rightarrow M_X(t) = \frac{\lambda}{\lambda - t} \quad \text{if } t < \lambda$

$X \sim \text{Binom}(n, p) \Rightarrow M_X(t) = (1 - p + pe^t)^n$

$Z \sim N(0, 1) \Rightarrow M_Z(t) = e^{t^2/2}$

$X \sim \text{Geom}(p) \Rightarrow M_X(t) = \frac{pe^t}{1 - e^t(1-p)} \quad \text{if } t < \ln\left(\frac{1}{1-p}\right)$

$X \sim N(\mu, \sigma^2) \Rightarrow M_X(t) = e^{t\mu + \frac{1}{2}\sigma^2 t^2}$

$X \sim \text{Deg}(c) \Rightarrow M_X(t) = e^{tc}$

IV. Levy's Continuity Theorem

X_1, X_2, \dots is a sequence of r.v.

$\lim_{n \rightarrow \infty} M_{X_n}(t) = M_Y(t) \Leftrightarrow X_n \xrightarrow{\text{converges}} Y$

$\text{If } n \text{ large, } M_{X_n}(t) \not\approx M_Y(t) \Rightarrow X_n \not\stackrel{d}{=} Y$

Law of Large Numbers

$X_1, \dots, X_n \stackrel{iid}{\sim}$ with mean μ

$\Rightarrow \bar{X}_n \xrightarrow{\text{w.p.1}} \mu$
 $\sim \text{Deg}(\mu)$

$M_{\bar{X}_n}(t) = M_{\frac{1}{n} \sum X_i}(t) = M_{\sum X_i}\left(\frac{t}{n}\right) = M_{X_1 + \dots + X_n}\left(\frac{t}{n}\right)$
III
 $= \left(M_{X_1}\left(\frac{t}{n}\right)\right)^n$
IV

$M_X(t) := E[e^{tX}]$

Taylor series

by def. of m.g.f. $= \left(E\left[e^{\frac{t}{n}X}\right]\right)^n = \left(E\left[1 + \frac{t}{n}X + \frac{t^2 X^2}{2! n^2} + \frac{t^3 X^3}{3! n^3} + \dots\right]\right)^n$

• We say $f(n) = o(g(n))$ if $\lim_{n \rightarrow \infty} \frac{f(n)}{g(n)} = 0$
("little o")

$g(n)$ grows "faster" than $f(n)$.

ex: $\text{is } n^2 = o(n^3)?$

$\lim_{n \rightarrow \infty} \frac{n^2}{n^3} = \lim_{n \rightarrow \infty} \frac{1}{n} = 0 \quad \checkmark$

$$\mathbb{E} \left[\frac{t^2 x^2}{2n^2} + \frac{t^3 x^3}{3!n^3} + \dots \right] = o\left(\frac{1}{n}\right)$$

$$\lim_{n \rightarrow \infty} \frac{1}{n} = \lim_{n \rightarrow \infty} \frac{t^2 x^2}{2n} + \frac{t^3 x^3}{3!n^3} + \dots = 0$$

$$= \left(\mathbb{E} \left[1 + \frac{t}{n} x + \frac{t^2 x^2}{2!n^2} + \frac{t^3 x^3}{3!n^3} + \dots \right] \right)^n = \left(\mathbb{E} \left[1 + \frac{t}{n} x + o\left(\frac{1}{n}\right) \right] \right)^n = \left(1 + \frac{t\mu}{n} + \mathbb{E} \left[o\left(\frac{1}{n}\right) \right] \right)^n$$

$$= \left(1 + \frac{t\mu}{n} + o\left(\frac{1}{n}\right) \right)^n$$

ignore expectation

$$\lim_{n \rightarrow \infty} \left(1 + \frac{t}{n} \right)^n = e$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{t}{n} + \frac{1}{n^2} \right)^n = e$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{a}{n} \right)^n = e^a$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{t}{n} + o\left(\frac{1}{n}\right) \right)^n = e \quad \text{and} \quad \lim_{n \rightarrow \infty} \left(1 + \frac{a}{n} + o\left(\frac{1}{n}\right) \right)^n = e^a$$

$$\lim_{n \rightarrow \infty} M_{\bar{X}_n}(t) = \lim_{n \rightarrow \infty} \left(1 + \frac{t\mu}{n} + o\left(\frac{1}{n}\right) \right)^n = e^{t\mu}$$

$$\xrightarrow{\bar{X} \rightarrow \mu}$$

If $X_1, X_2, \dots, X_n \stackrel{iid}{\sim}$ with mean μ and v.e. σ
 What does $\bar{X}_n - \mu$ look like as n gets large?
 $C_n = \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}}$ \bar{X} standardized to have mean 0 and v.e. 1

$$C_n = \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} = \frac{\sqrt{n} \left(\frac{X_1 + X_2 + \dots + X_n}{n} - \mu \right)}{\sigma} = \frac{X_1 + \dots + X_n - n\mu}{\sigma\sqrt{n}}$$

$$= \frac{(X_1 - \mu) + (X_2 - \mu) + \dots + (X_n - \mu)}{\sigma\sqrt{n}} = \frac{1}{\sqrt{n}} \left(\frac{X_1 - \mu}{\sigma} + \frac{X_2 - \mu}{\sigma} + \dots + \frac{X_n - \mu}{\sigma} \right)$$

$$\hookrightarrow \text{let } Z_i = \frac{X_i - \mu}{\sigma} = \frac{1}{\sqrt{n}} (Z_1 + \dots + Z_n)$$

$$\mathbb{E}[Z] = 0$$

$$\mathbb{E}[Z^2] = 1$$

$$\mathbb{E}[X^2] = 1$$

$$\text{Var}[X] = \mathbb{E}[X^2] - \mu^2$$

$$\text{if } \mu = 0$$

$$\Rightarrow \sigma^2 = \mathbb{E}[X^2]$$

$$\text{if } \sigma = 1$$

$$\Rightarrow \mathbb{E}[X^2] = 1$$

$$M_{C_n}(t) = M_{\frac{1}{\sqrt{n}}(Z_1 + \dots + Z_n)}(t) = M_{Z_1 + \dots + Z_n}\left(\frac{t}{\sqrt{n}}\right) = \left(M_Z\left(\frac{t}{\sqrt{n}}\right) \right)^n = \left(\mathbb{E} \left[e^{\frac{t}{\sqrt{n}} Z} \right] \right)^n$$

if iid

$$\mathbb{E} \left[\left(1 + \frac{t}{\sqrt{n}} Z + \frac{t^2 Z^2}{2!n} + \frac{t^3 Z^3}{3!n^{3/2}} + \frac{t^4 Z^4}{4!n^2} + \dots \right)^n \right]$$

$$\mathbb{E} \left[\left(1 + \frac{tZ}{\sqrt{n}} + \frac{t^2 Z^2}{2n} + o\left(\frac{1}{n}\right) \right)^n \right] = \left(1 + \frac{t^2}{2n} + o\left(\frac{1}{n}\right) \right)^n$$

$$\mathbb{E} \left[\frac{t^2}{2n} \right] = \frac{t^2}{2n} \mathbb{E}[Z^2] = 0, \quad \mathbb{E} \left[\frac{t^2 Z^2}{2n} \right] = \frac{t^2}{2n} = \mathbb{E}[Z^2] = \frac{t^2}{n}$$

$$\lim_{n \rightarrow \infty} M_{C_n}(t) = \lim_{n \rightarrow \infty} \left(1 + \frac{t^2}{2n} + o\left(\frac{1}{n}\right) \right)^n = e^{\frac{t^2}{2}}$$

$$\Rightarrow C_n \rightarrow Z \sim N(0, 1)$$

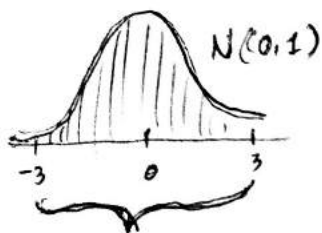
Central Limit Theorem CLT
So... if X_1, X_2, \dots, X_n iid
with mean μ and s.e. σ ,
$$C_n = \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \rightarrow Z \sim N(0, 1)$$

If n is large, $C_n \stackrel{d}{\approx} Z \sim N(0, 1)$ $\mu + \sigma Z \sim N(\mu, \sigma^2)$

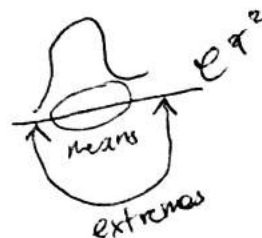
$$\text{I. } C_n \stackrel{d}{\approx} Z \sim N(0, 1)$$

$$\rightarrow \text{II. } \bar{X} \stackrel{d}{\approx} N(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2)$$

$$\rightarrow \text{III. } T \stackrel{d}{\approx} N(n\mu, (\sqrt{n}\sigma)^2)$$



Normal curve defines what normal is inside the curve



EX: Take 100 random equally likely L or R steps.
What is the probability you're more than 10 steps away from where you began?

by CLT

$$\text{Let } T = X_1 + X_2 + \dots + X_{100} \approx N(n\mu, (\sqrt{n}\sigma)^2)$$

$$X_1, X_2, \dots, X_{100} \stackrel{\text{iid}}{\sim} \begin{cases} 1 & \text{w.p. } \frac{1}{2} \\ -1 & \text{w.p. } \frac{1}{2} \end{cases}$$

$$\mu = 0$$

$$\sigma = \sqrt{\sigma^2} = \sqrt{1} = 1$$

$$N(0, 10^2)$$

$$P(|T| > 10) = P(T < -10) + P(T > 10)$$

$$= P\left(\frac{T-0}{10} < \frac{-10-0}{10}\right) + P\left(\frac{T-0}{10} > \frac{10-0}{10}\right)$$

$$\approx P(Z < -1) + P(Z > 1) = P(Z \notin [-1, 1]) \approx 32\%$$

EX2: $X_1, \dots, X_{30} \stackrel{\text{iid}}{\sim} \text{Geom}(p = \frac{1}{2})$

What is the prob. the average stopping time is more than 2.75?

$$\bar{X}_{30} \approx N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) \text{ by the CLT}$$

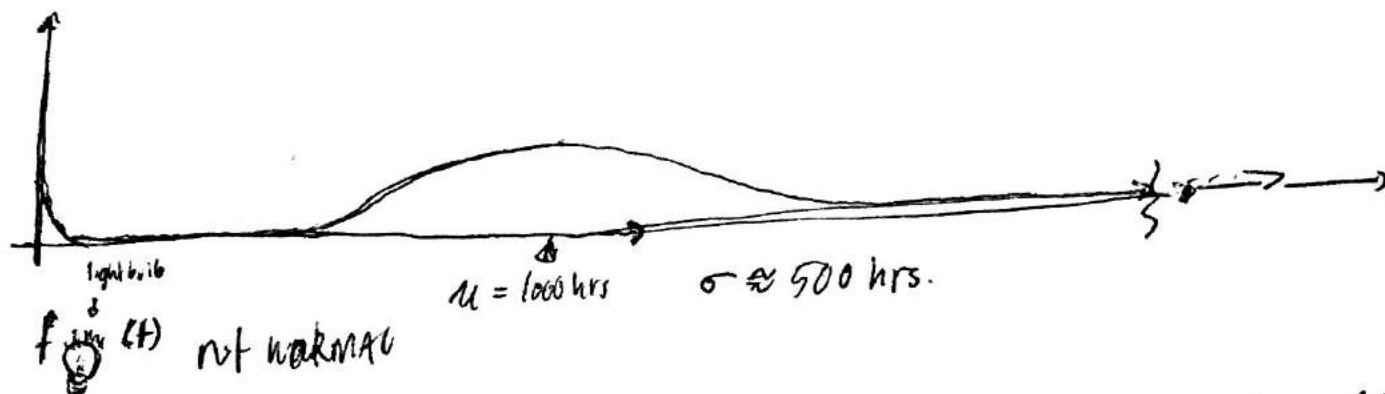
$$\mu = \frac{1}{p} = \frac{1}{\frac{1}{2}} = 2$$

$$\sigma = \sqrt{\frac{1-p}{p^2}} \approx 1.414$$

$$\frac{\sigma}{\sqrt{n}} \approx \frac{1.414}{\sqrt{30}} \approx .258$$

$$P(\bar{X} > 2.75) = P\left(\frac{\bar{X}-2}{.258} > \frac{2.75-2}{.258}\right) \stackrel{\text{by CLT}}{\approx} P(Z > 3) = .0045$$

ex 3: Light Bulb failure times



You buy 50 lightbulbs, what is probability they last on average more than 1300 hrs?

$$\hat{X} \approx N\left(\mu, \left(\frac{\sigma}{\sqrt{n}}\right)^2\right) = N\left(1000, \left(\frac{500}{\sqrt{50}}\right)^2\right)$$

by CLT

\downarrow
70.7

$$P(\bar{X} > 1300) = P\left(\frac{\bar{X} - 1000}{70.7} > \frac{1300 - 1000}{70.7}\right)$$

$$\approx P(Z > 4.24) \approx 0$$