

Example 0.1. Let $x_1, \dots, x_6 \stackrel{iid}{\sim} \text{Bern}(\theta)$ be the data set $\langle 0, 0, 1, 0, 1, 0 \rangle$. Then:

$$\begin{aligned}
 l(\theta; x) &= \ln\left(\prod_{i=1}^6 \theta^{x_i} (1 - \theta)^{1-x_i}\right) \\
 &= \sum_{i=1}^6 \ln(\theta^{x_i} (1 - \theta)^{1-x_i}) \\
 &= \sum_{i=1}^6 x_i \ln(\theta) + (1 - x_i) \ln(1 - \theta) \\
 &= \ln(\theta) \sum_{i=1}^6 x_i + (6 - \sum_{i=1}^6 x_i) \ln(1 - \theta) \\
 &= \ln(\theta) 6\bar{x} + (6 - 6\bar{x}) \ln(1 - \theta) \\
 &= 6(\ln(\theta) + (1 - \bar{x}) \ln(1 - \theta))
 \end{aligned}$$

Now let's differentiate this to maximize it:

$$\frac{d}{dt} 6(\ln(\theta) + (1 - \theta) \ln(1 - \theta)) = 6\left(\frac{\bar{x}}{\theta} - \frac{1 - \bar{x}}{1 - \theta}\right)$$

If we set it equal to 0,

$$(1 - \theta)\bar{x} - \theta(1 - \bar{x}) = 0 \rightarrow \hat{\theta}_{MLE} = \bar{x}$$

Note: For our convenience, we use the natural log to differentiate \prod to \sum . It is easier to differentiate sums rather than products.

Definition 0.1. Maximum Likelihood Estimation: $\hat{\theta}_{MLE} = \bar{X}$ where \bar{X} is a random variable and has properties

Definition 0.2. Maximum Likelihood Estimate: $\hat{\theta}_{MLE} = \bar{x}$ where \bar{x} has a numerical value

Example 0.2. Let $x_1, \dots, x_n \stackrel{iid}{\sim} \text{Geom}(\theta) = (1 - \theta)^x \theta$ where x is the number of failures before stopping success. $\text{Supp}(X) = \{0, 1, \dots\} = \mathbb{N}$ and $\Theta = (0, 1)$. Then:

$$\begin{aligned}
 p(x_1, \dots, x_n) &= \mathcal{L}(\theta; x_1, \dots, x_n) \\
 &= \prod_{i=1}^n (1 - \theta)^{x_i} \theta
 \end{aligned}$$

Therefore

$$\begin{aligned}
 l(\theta; x) &= \sum \ln(1 - \theta)^{x_i} \theta \\
 &= \ln(1 - \theta) \sum x_i + n \ln(\theta)
 \end{aligned}$$

We will now differentiate this function to solve for $\hat{\theta}_{MLE}$.

$$\begin{aligned} l'(\theta; x) &= \frac{n}{\theta} - \frac{n\bar{x}}{1-\theta} = 0 \\ \frac{1}{\theta} &= \frac{\bar{x}}{1-\theta} \\ \frac{1}{\theta-1} &= \bar{x} \\ \hat{\theta}_{MLE} &= \frac{1}{\bar{x}+1} \end{aligned}$$

Properties of MLE:

1. There exists $\varepsilon > 0$ such that

$$\lim_{n \rightarrow \infty} P(|\hat{\theta}_{MLE} - \theta| \geq \varepsilon) = 0$$

2. Asymptotic Normaling: As n increases, the the parameters behave like a normal distribution

$$\hat{\theta}_{MLE} \xrightarrow{d} N(\hat{\theta}_{MLE}, SE(\hat{\theta}_{MLE})^2)$$

3. Efficiency: $\hat{\theta}_{MLE}$ has the lowest standard error theoretically possible

Inference with MLE:

- Point Estimate: $\hat{\theta}_{MLE}$
- Confidence Set: $CI_{\theta, 1-\alpha} = [\hat{\theta}_{MLE} \pm z_{\frac{\alpha}{2}} SE(\hat{\theta}_{MLE})]$
Here, θ is the parameter of interest whereas $1 - \alpha$ is the confidence level.
- Hypothesis Testing: $H_0 : \theta = \theta_0$, $H_A : \theta \neq \theta_0$ - fail to reject if $\hat{\theta}_{MLE}$ is in the region of $[\theta_0 \pm z_{\frac{\alpha}{2}} SE(\hat{\theta}_{MLE})]$

We must observe data, then pick a parametric model \mathcal{F} , do inference with MLE. The problem with this is that

1. If all data values taken are 0 and we take $\mathcal{F} = \text{Bern}(\theta)$, then $\hat{\theta}_{MLE} = \bar{x} = 0$ and $SE(\hat{\theta}_{MLE}) = \sqrt{\hat{\theta}_{MLE}(1-\hat{\theta}_{MLE})} = 0$. This gives no information and thus is a big problem. No confidence set, no hypothesis testing.
2. What if we have prior knowledge about Θ ? We can't use it because only data set can be used.
3. Frequentist Confidence Interval Interpretation: Let's say we found $CI_{\theta, 1-\alpha} = [0.42, 0.47]$. If the experiment is repeated "many" times, then a confidence level of 95% will cover θ and $1 - \alpha$ is contained in the set. But given just an interval, we can only say that a certain value will either fall in the interval or not. We can't claim that the probability that the interval contains θ is $1 - \alpha$.

4. Hypothesis testing: not satisfactory since we do not know if data values are far from being retained yet rejected or near rejection (extremeness). How good is the rejection? What is $P(H_0|x)$, or H_0 given x ?
5. Boundary Issues: Let's say $x = \langle 0, 0, 1, 0, 1, 0 \rangle$ and $\hat{\theta}_{MLE} = \frac{1}{3}$. We want a confidence set at the 95% confidence level: $CI_{\theta, 95\%} = (\frac{1}{3} \pm 2\sqrt{\frac{1}{3}\frac{2}{3}}) = (-0.6, 1.26)$. In this confidence interval, we have both a negative value and one that's greater than 1. This is no good. This happened because our data set is only composed of 6 values. Thus it cannot converge to normality. We cannot use the normal distribution to construct the interval and since we did, it came out looking wrong.

Good news: The Bayesian approach will not cause any of these issues.

Definition 0.3. Conditional Probability: $P(B|A)$, the probability of B occurring given A occurs

$$P(B|A) = \frac{P(A, B)}{P(A)}$$

Note: There is a proportionality between $P(A, B)$, the intersection of two events, and $P(B|A)$, the probability of B occurring given A occurs. Thus we can write

$$P(A, B) \propto P(B|A)$$

or

$$P(A, B) = cP(B|A)$$

Definition 0.4. Baye's Rule:

$$P(B|A) = \frac{P(A, B)}{P(A)}$$

We know from previous probability courses that $P(A, B) = P(B, A)$. We also know that $P(A, B) = P(B|A)P(A)$ and $P(B, A) = P(A|B)P(B)$. Let's set them equal to each other.

$$\begin{aligned} P(A, B) &= P(B, A) \\ P(B|A)P(A) &= P(A|B)P(B) \end{aligned}$$

This is another form of Baye's rule.

Definition 0.5. Law of Total Probability: the probability of event A occurring is sum of the probability of the intersection of event A and event B and the probability of the intersection of event A and not event B (complement of B)

$$P(A) = P(A, B) + P(A, B^C)$$

Let's combine the two equations from above.

$$\begin{aligned} P(A) &= P(A, B) + P(A, B^C) \\ &= P(A|B)P(B) + P(A|B^C)P(B^C) \\ P(B|A) &= \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|B^C)P(B^C)} \end{aligned}$$

This is another form of Baye's rule.

Note:

$$P(B|A) = \frac{P(A|B)P(B)}{P(A)}$$

The LHS is the posterior probability where B is the parameter of interest, A is the evidence/data, and $B|A$ is the targeted estimation. On the RHS, $P(A|B)$ is the likelihood or probability of data and $P(B)$ is a prior probability.

Finding $P(B|A)$ using $A(\text{data})$ and applying it to $P(B)$ is called Bayesian conditioning.