

Recall:  $X_1, \dots, X_n | \theta, \sigma^2 \stackrel{iid}{\sim} N(\theta, \sigma^2)$   
 $\theta \sim N(\mu_0, \tau^{-2})$   
 $\sigma^2 \sim \text{Invgamma}(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2})$   
 $\Rightarrow p(\theta, \sigma^2 | X) \propto N(\theta_p, \sigma_p^2) K(\sigma^2 | X)$

Recall:  $p(\theta | \sigma^2, X) = N(\theta_p, \sigma_p^2)$   
 $p(\sigma^2 | \theta, X) = \text{Invgamma}(\frac{n_0 + n}{2}, \frac{n \hat{\sigma}_{MLE}^2 + n_0 \sigma_0^2}{2})$  } Can we use these 2 conditionals to tell us about  $p(\theta, \sigma^2 | X)$ ?

•  $P(AB) = P(A|B)P(B) = P(B|A) \cdot P(A)$   
 $P(\theta, \sigma^2 | X) = P(\theta | \sigma^2, X) P(\sigma^2 | X) = P(\sigma^2 | \theta, X) P(\theta | X)$

No. You need to know one of the marginals. Not just the conditionals.

What if we used an iterative approach?

1. Assume an arbitrary  $\theta_0$

2. Draw  $\sigma_0^2$  from  $p(\sigma^2 | \theta = \theta_0, X) = \text{Invgamma}$

3. Draw  $\theta_1$  from  $p(\theta | \sigma^2 = \sigma_0^2, X)$

4. Draw  $\sigma_1^2$  from  $p(\sigma^2 | \theta = \theta_1, X)$

5. Repeat 2,3 until "convergence".

like Newton-Rafson / EM.

→ will return whole posterior.

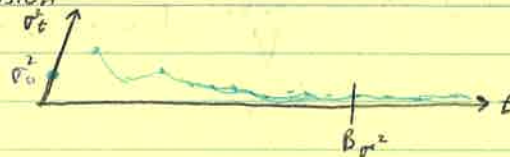
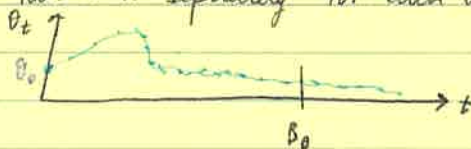
EM ≠ return whole posterior

Algorithm called Gibbs sampling or Gibbs sampler. = return  $\theta_{MLE}$ .

The output of a Gibbs sample looks like  
 $\langle [\theta_0^2, \sigma_0^2], [\theta_1^2, \sigma_1^2], \dots, [\theta_t^2, \sigma_t^2], \dots \rangle$

"Edward Teller"  
H-Bomb

look at it separately for each dimension



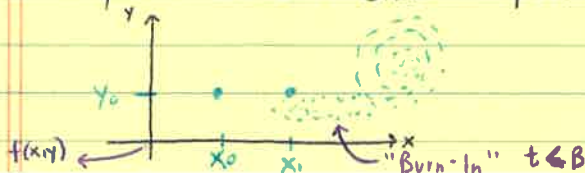
"Burn-in point"

Point where operates the way it's supposed to normally.

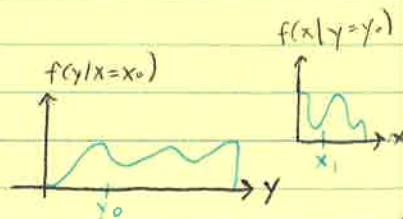
(Buying new car)

For the Gibbs ~~samples~~ <sup>chain</sup> the burn-in is the max  $B_j$  for all  $\theta_j$ 's. For  $t \geq B$ , the chain is "burned-in".

If you seek  $f(x, y)$  but only know  $f(x|y)$  and  $f(y|x)$  you can use a Gibbs sampler.



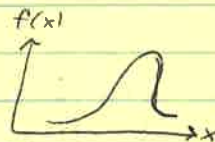
1. Begin at arbitrary  $x_0$
2. Draw  $y_0$  from  $f(y|x=x_0)$
3. Draw  $x_1$  from  $f(x|y=y_0)$



What if you want  $f(x)$ ?

$$f(x) = \int_{\text{supp}(Y)} f(x,y) dy$$

collapsing  $y$  values  
get all  $x$ -values.  
get one-dimensional hill.



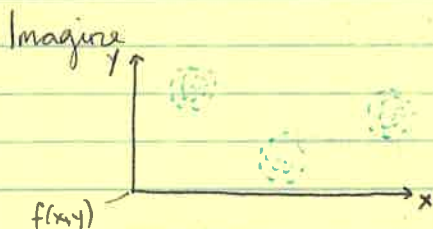
## Main Problems with Gibbs Sampler



good ✓ "Mixing"

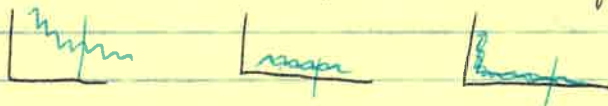


? Don't know what's going on.



⇒ Sampler gets hopelessly stuck in a piece of the distribution.

A possible solution: Use multiple Gibbs chains from different  $X_0$ 's.



A smaller problem which is more fixable is as follows:

Is  $\theta_1$  related to  $\theta_0$ ? YES.

$$\text{Correlation}[X,Y] := \frac{\text{Cov}[X,Y]}{\sqrt{\text{SE}(X)} \sqrt{\text{SE}(Y)}} = \frac{E[(X-\mu_X)(Y-\mu_Y)]}{\sqrt{\text{Var}(X)} \sqrt{\text{Var}(Y)}}$$

$$\text{Cov}[\theta_1, \theta_0] \neq 0$$

$$\text{estimator Covariance } r := \frac{s_{xy}}{s_x s_y} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum (x_i - \bar{x})^2} \sqrt{\sum (y_i - \bar{y})^2}}$$

↑  
estimators for stand. errors.

Is  $\theta_{1,000}$  related to  $\theta_{0,000}$ ? Yes. ⇒ So Burn-In doesn't help.

"Auto-correlation" for "lag" one.

$$r_{a,1} := \frac{\sum_{t=B}^{B+S-1} (\theta_t - \bar{\theta})(\theta_{t+1} - \bar{\theta})}{\sum_{t=B}^{B+S-1} (\theta_t - \bar{\theta})^2}$$

$$\bar{\theta} = \text{average of all } \theta\text{'s} \\ = \frac{1}{S} \sum_{t=B}^{B+S-1} \theta_t$$

B = Burn In

Auto = Self



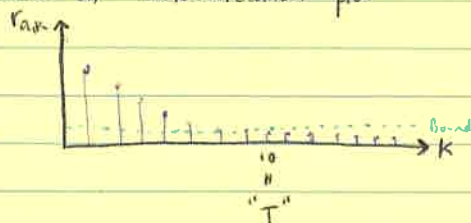
Autocorrelation of lag 2:  $r_{a,2} := \frac{\sum_{t=B}^{B+S-2} (\theta_t - \bar{\theta})(\theta_{t+2} - \bar{\theta})}{\sum_{t=B}^{B+S-1} (\theta_t - \bar{\theta})^2}$

$B+S$  = the total # of iterations.

Autocorrelation of lag  $k$ :  $r_{a,k} := \frac{\sum_{t=B}^{B+S-k} (\theta_t - \bar{\theta})(\theta_{t+k} - \bar{\theta})}{\sum_{t=B}^{B+S-1} (\theta_t - \bar{\theta})^2}$

When does correlation not matter anymore?

Look at "autocorrelation plot"



takes 10 iterations to say not correlation doesn't matter anymore. (correlation so minimal)

$\langle [\theta_B, \sigma_B^2], [\theta_{B+T}, \sigma_{B+T}^2], [\theta_{B+2T}, \sigma_{B+2T}^2], \dots \rangle$  good to have  $\geq 1000$ . length  $L$

Now your Burned & thinned chain.

$\hat{\theta}_{MSE} = E[\theta|X] \approx \bar{\theta} = \frac{1}{L} \sum_{l=1}^L \theta_l$  average out

$\hat{\theta}_{MAE} = \text{Med}[\theta|X] \approx$  order  $\theta$ 's from lowest to highest  $\theta_1, \theta_2, \dots, \theta_L$  and take  $\theta_{(\frac{L}{2})}$  middle value.

mode estimations are not easy...

$P(X^*|X) = \int P(X^*|\theta) P(\theta|X) d\theta$

explain process (in Firms?) Sample  $\theta$  from posterior Sample  $X$ .

How do we approximate sample from this Distribution?

1. Pick  $l \in \{1, \dots, L\}$
2. Sample  $X^*$  from  $P(X^*|\theta = \theta_l)$
3. Repeat 1-2.