

Lee 7 Anh 301 2/23/17

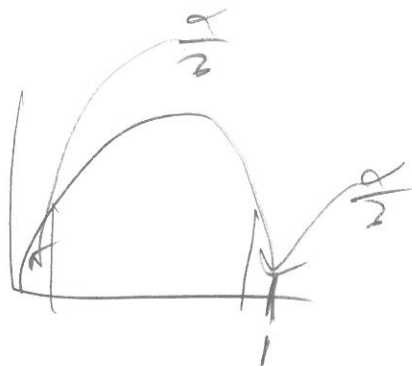
1

$\theta \sim \text{Gaussian}$

$$\theta \sim \mathcal{N}(\mu, \sigma^2) = \text{Bern}(1, 1)$$

$$n=2, x=1$$

$$\Rightarrow \theta|x \sim \text{Bern}(x+\alpha, n+x+\beta) = \text{Bern}(2, 2)$$



$$CR_{\theta, 1-\alpha} := [\text{Quantile}[\theta|x, \frac{\alpha}{2}], \text{Quantile}[\theta|x, 1-\frac{\alpha}{2}]]$$

Two-sided CR. The middle $1-\alpha$

exact value given

$$CR_{\theta, 95\%} = [.098, .908]$$

$$P(\theta \in CR) = 1-\alpha$$

$$P(\theta \in [.098, .908]) = 95\% !$$

One-sided CR \rightarrow left-sided CR lower 95%

\rightarrow right side CR higher 95%

$$CR_{L, \theta, 1-\alpha} := [-\infty, \text{Quantile}[\theta|x, 1-\alpha]] = [0, 1]$$

$$\text{in our case, } CR_{L, \theta, 95\%} = [0, \text{Quantile}[\theta|x, .95]] = [0, .865]$$

$$CR_{R, \theta, 1-\alpha} := [\text{Quantile}[\theta|x, \alpha], \infty] = [.135, 1]$$

in our case

$$CR_{R, \theta, 95\%} = [\text{Quantile}[\theta|x, .05], 1] = [.135, 1]$$

Param Hypothesis Testing (3rd goal of inference)

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→ Theory Testing

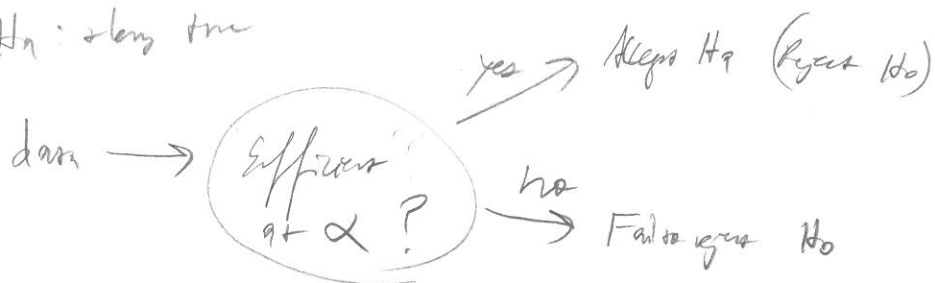
Theory: H_0 ... which I want to prove

But I want to be sure I'm right.

So I assume the ^(H₀) opposite and only accept my theory if there is overwhelming evidence.
Evidence is determined by α .

H_0 : theory not true

H_1 : theory true



$$H_0: \theta \leq 0.5 = \theta_0$$

$$H_1: \theta > 0.5 = \theta_0$$



$$\hat{\theta} \sim N\left(\theta_0, \left(\sqrt{\frac{\theta_0(1-\theta_0)}{n}}\right)^2\right) = N\left(0.5, \left(\sqrt{\frac{0.5(1-0.5)}{100}}\right)^2\right)$$

if $n=100$

$$\text{Rejection Region} = \left[0, \theta_0 + z_{\alpha} \sqrt{\frac{\theta_0(1-\theta_0)}{n}}\right] = \left[0, 0.5 + 1.64 \cdot 0.05\right] = [0, 0.58]$$

if $\hat{\theta} \in \text{Rejection Region} \Rightarrow \text{F.T.R. } H_0$

if $\hat{\theta} \notin \text{''} \Rightarrow \text{Reject } H_0$

$$p\text{-val} := P(\text{going down or more extreme} \mid H_0 \text{ true}) = P(\hat{\theta} > \theta \mid \hat{\theta} \sim N\left(0, \left(\sqrt{\frac{\theta_0(1-\theta_0)}{n}}\right)^2\right)) = \arg\min_{\alpha} \{ \hat{\theta} \in \text{Rejection Region} \}$$

$$H_0: \theta = 0.5 = \theta_0$$

$$H_1: \theta < 0.5 = \theta_0$$



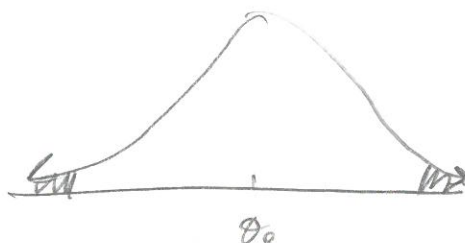
$$\text{Ret. Region} = \left[\theta_0 - z_{\alpha} \sqrt{\frac{\theta_0(1-\theta_0)}{n}}, 1 \right]$$

$$H_0: \theta = 0.5 = \theta_0$$

$$H_1: \theta \neq 0.5 = \theta_0$$

$$\theta > 0.5 \text{ or } \theta < 0.5$$

two mps to reject



$$\text{Ret. Region} = \left[\theta_0 \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\theta_0(1-\theta_0)}{n}} \right]$$

$$p_{\text{val}} = \inf_{\alpha} \{ \theta \in \text{Ret. Region} \}$$

$$p_{\text{val}} = P(\text{rejecting } H_0 \text{ or not} \mid H_0 \text{ true})$$

$$\neq P(H_0)$$

$$\neq P(H_0 \mid X)$$

$$\neq P(H_1)$$

$$\neq P(H_1 \mid X)$$

these are things you want!

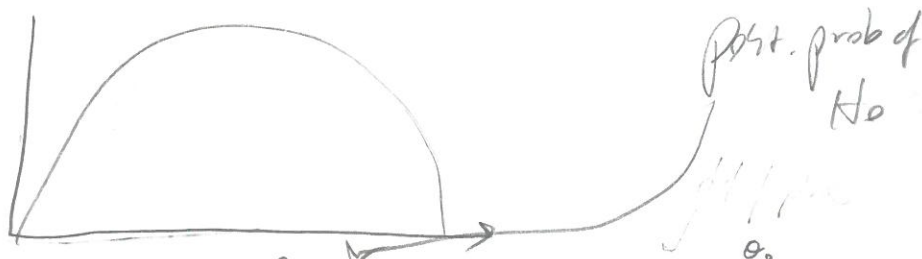
Bayesian Hyp. Testing

$$H_0: \theta \leq 0.5 = \theta_0$$

$$H_1: \theta > 0.5 = \theta_0$$

with data

$$\text{we know } P(\theta \mid X)$$



$$\text{we can calc } p_{\text{val}} = P(H_0 \mid X) = P(\theta \leq \theta_0 \mid X) = \int_{\theta_0}^{\theta_0} \frac{1}{P(\theta_0 \mid X)} \theta^{\alpha-1} (1-\theta)^{n-\alpha} d\theta$$

if low.. reject

$$= p_{\text{bet}}(\theta_0, X, \alpha, n, \beta)$$

if $p_{\text{val}} < \alpha \Rightarrow \text{Reject } H_0$
 $p_{\text{val}} \geq \alpha \Rightarrow \text{FTR } H_0$

$$P(H_0 | x) = \frac{P(x | H_0) P(H_0)}{P(x)} \quad \leftarrow \text{what is this? prior prob of } H_0$$

$$P(\theta \leq \theta_0 | x) = \frac{P(x | \theta > \theta_0) P(\theta \leq \theta_0)}{P(x | \theta > \theta_0) P(\theta \leq \theta_0) + P(x | \theta \leq \theta_0) P(\theta > \theta_0)}$$

if $\theta \sim U(0,1)$...

$$P(H_0) = P(\theta \leq \theta_0) = \theta_0 \quad \text{Prior prob, ...}$$

Rule of $\hat{\theta} = \bar{x}$? No... no longer a test statistic

Note: this scheme equivalent to $\theta_0 \in CR_{R, \theta, 1-\alpha} \Rightarrow \text{Reject } H_0$

$$H_0: \theta \geq \theta_0 = 0.5$$

$$H_1: \theta < \theta_0 = 0.5$$

$$p_{\text{val}} = P(H_0 | x) = P(\theta \geq \theta_0 | x) = \int_{\theta_0}^1 \frac{1}{B(\alpha, \beta)} e^{\alpha x - 1} (1-\theta)^{\beta + x - 1} d\theta$$

$$= 1 - \text{pbeta}(\theta_0, \alpha + x, \beta + x)$$

if $p_{\text{val}} < \alpha \Rightarrow \text{Reject } H_0$
 $p_{\text{val}} \geq \alpha \Rightarrow \text{Reject } H_0$

Note: equivalent to $\theta_0 \in CR_{L, \theta, 1-\alpha} \Rightarrow \text{Reject } H_0$

$$H_0: \theta = \theta_0$$

$$H_a: \theta \neq \theta_0$$

"point null" or "precise null"
 100% true, then H_0

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$$p_{val} = P(H_0 | X) = P(\theta = \theta_0 | X) = 0 \text{ (always) why? P-value not 24/7?}$$

why? In reality all ^{simple} point nulls are absurd...

not possible!

prob coin flip = $\frac{1}{2}$? No.. 0.500001 real or
 or sale

Two ideas ①

but 0.500001 = 0.5 for all practical purposes

$$\Rightarrow H_0: \theta \in [\theta_0 \pm f]$$

$$H_a: \theta \notin [\theta_0 \pm f]$$

margin of error
 this is usually what you need!

What is margin of error for coin flip? $f = 0.01$?

②

$$H_0: \theta = \theta_0$$

$$H_a: \theta \neq \theta_0$$

if plausible

$$\theta_0 \in CR_{\theta, 1-\alpha} \Rightarrow \text{Retain}$$

$$\theta_0 \notin CR_{\theta, 1-\alpha} \Rightarrow \text{Reject}$$

p-value criterion has not been...

This seems to be correct...

which is why we use Bayes Factors (soon)

Examples

$$H_0: \theta \leq .1$$

$$H_a: \theta > .1$$

$$\alpha = 5\%, n = 150, X = 23$$

$$\text{Res Pgn} = \left[0, .1 + 1.64 \sqrt{\frac{.1 \cdot .9}{150}} \right] = \left[0, .140 \right] \quad = .024$$

$$\hat{\theta} = .153 \notin \text{Res Pgn} \Rightarrow \text{Reject } H_0$$

$$\text{Bayesian} \dots \theta \sim U(0,1) = \text{Bern}(1,1)$$

$$\theta | x \sim \text{Bern}(\alpha + x, \beta + n - x) = \text{Bern}(24, 128)$$

$$p_{\text{val}} := P(\theta \leq .1 | x) = \int_0^1 \text{Bern}(24, 128) d\theta = p_{\text{bern}}(.1, 24, 128) = .01544 \quad \text{Similar!}$$

$$p_{\text{val}} := P(\hat{p} > .153 | \theta = .1)$$

$$P\left(\hat{Z} > \frac{.153 - .1}{.024}\right) = 2.16$$

$$= \hat{p}_{\text{norm}}(2.16)$$

$$= .01539 < \alpha \Rightarrow \text{Reject}$$

$$H_0: \theta = .5$$

$$H_a: \theta \neq .5$$

$$\alpha = 5\%$$

$$n = 100$$

$$X = 61$$

$$\text{Res Pgn} = \left[.5 \pm 2 \sqrt{\frac{.5 \cdot .5}{100}} \right] = [.4, .6]$$

$$\hat{\theta} = .61 \notin \text{Res Pgn} \Rightarrow \text{Reject}$$

$$p_{\text{val}} := P\left(\hat{p} > \frac{.61 - .5}{.05}\right) = 2P(Z > 2.2) = 2(1 - \text{pnorm}(2.2)) = .0278$$

$$\text{Bayesian} \quad \theta \sim U(0,1)$$

$$\theta | x \sim \text{Bern}(62, 100)$$

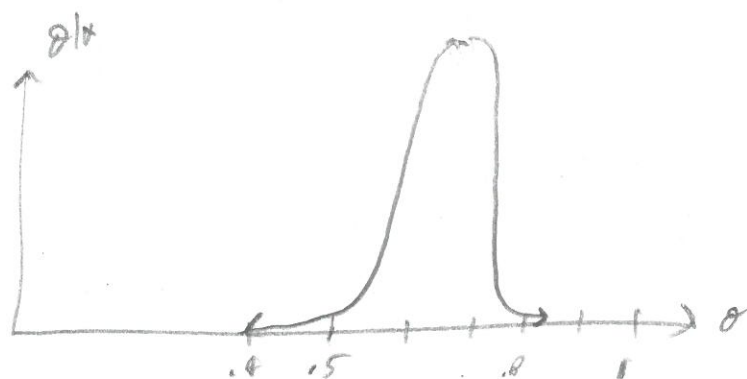
$$\textcircled{I} H_0: \theta \in [.49, .51] \quad \text{i.e. } \delta = .01 \text{ equidistant margin}$$

$$H_a: \theta \notin [.49, .51]$$

$$p_{\text{rel}} = P(\theta \in [.11, .51] | x) = \int_{.11}^{.51} \text{bern}(62, \theta) d\theta = p_{\text{bern}}(.51, 62, \theta_0) - p_{\text{bern}}(.11, 62, \theta_0) = .0197$$

$$\textcircled{\text{II}} CR_{\theta, 1-\alpha} = [p_{\text{bern}}(.025, 62, \theta_0), p_{\text{bern}}(.975, 62, \theta_0)] \\ = [.511, .700]$$

$$\theta_0 = .5 \notin CR_{\theta, 1-\alpha} \Rightarrow \text{Reject } H_0$$



Another way to "test" Bayesian like this test

$$H_0: \theta = \theta_0 = 0.5$$

$$H_a: \theta \neq \theta_0 = 0.5$$

what is this really?

maybe

$$h = 100 \Rightarrow \hat{\theta} = .61$$

$$x = 61$$

$$H_a: \theta \sim U(0, 1)$$

θ could be "anything"

Define $B := \frac{P_{H_1}(X)}{P_{H_0}(X)}$ if B big $\Rightarrow H_1$ is a better model for the data, X .

Bayes Factor

denominator in Bayes Rule!!
prob of data!

$$= \frac{\int_{\Theta_{H_1}} P(X|\theta) P_{H_1}(\theta) d\theta}{\int_{\Theta_{H_0}} P(X|\theta) P_{H_0}(\theta) d\theta}$$

oops...

$$= \frac{\int_{0.5}^1 \binom{100}{61} .5^{61} (1-.5)^{100-61} (1) d\theta}{\int_0^1 \binom{100}{61} \theta^{61} (1-\theta)^{100-61} (1) d\theta}$$

for a difference

$$= \frac{B(62, 40)}{.5^{100}} = 1.39$$

H_1 better model... but is it decisive?

Jeffreys 1961 scale of Bayes Factors interpretation for H_1