

Bayes Factor:

$$B := \frac{P_{H_A}(X)}{P_{H_0}(X)} = \frac{\int_{\Theta_{H_A}} P_{H_A}(X | \theta) P_{H_A}(\theta) d\theta}{\int_{\Theta_{H_0}} P_{H_0}(X | \theta) P_{H_0}(\theta) d\theta}$$

Note: If  $B > 1$ ,  $H_A$  is supported. The bigger  $B$  is, the better  $H_A$  is.

Let  $H_0 : \theta = 0.5$  and  $H_A : \theta \neq 0.5$ . Assume  $\mathcal{F}$  is Binomial. For  $H_0$ :  $\theta \sim \text{Deg}(0.5)$  and for  $H_A$ :  $\theta \sim U(0, 1)$ .  $n = 100$  and  $x = 61$ . In the frequentist approach,  $H_0$  is rejected because  $p = 0.61$  which is too far from 0.5.

$$B = \frac{\int_0^1 \binom{n}{x} \theta^x (1 - \theta)^{n-x} \cdot (1) d\theta}{\int_{\{0.5\}} \binom{n}{x} 0.5^x (1 - 0.5)^{n-x} \cdot (1) d\theta} = \frac{B(x+1, n-x+1)}{0.5^n} = \frac{B(62, 98)}{0.5^{100}} = 1.39$$

Difference Conclusions:

- If  $B < 1$ , then no evidence
- If  $B \in [1 : 1.3 : 1]$ , then barely worth mentioning
- If  $B \in [3 : 1, 10 : 1]$ , then substantial
- If  $B \in [10 : 1, 30 : 1]$ , then strong
- If  $B \in [30 : 1, 100 : 1]$ , then very strong
- If  $B > 100\%$ , then decisive

Suppose  $H_0 : \theta = 0.5$  and  $H_A : \theta \neq 0.5$ . Let  $n = 104490000$ ,  $x = 52263920$  and  $\hat{\theta} = 0.50001768$ . In the frequentist approach, the p-value is 0.0003, which is less than 0.05 and thus  $H_0$  is rejected. In the Bayesian approach, assuming  $\theta \sim \text{Beta}(1, 1)$ ,

$$B = \frac{B(52263921, 104490000 - 52263920 + 1)}{0.50001768^{104490000}} = \frac{1}{12}$$

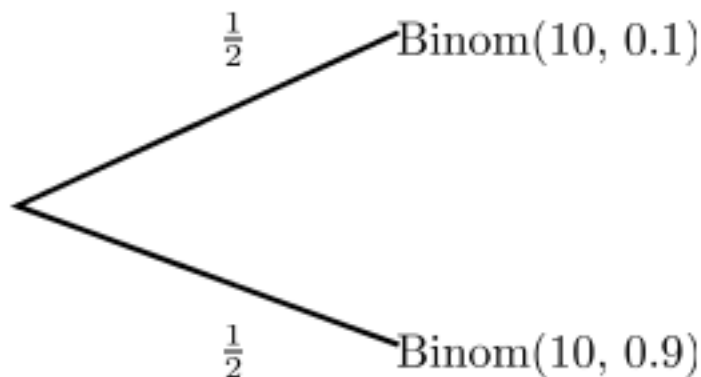
According to this, since  $B < 1$ , there is no evidence. This gives conflicting results. This happened because as  $n$  becomes large,  $H_0$  cannot be true and thus is rejected.

### End of Midterm 1 Material

Mixture Distribution: Let  $X \sim \begin{cases} N(0, 1)^2 & 0.5 \\ N(10, 1^2) & 0.5 \end{cases}$ .

$$\begin{aligned} P(X) &= \sum_{\theta \in \Theta} \mathbb{P}(X | \theta) \mathbb{P}(\theta) \\ &= \mathbb{P}(X | \theta = 0) \mathbb{P}(\theta = 0) + \mathbb{P}(X | \theta = 10) \mathbb{P}(\theta = 10) \\ &= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2} \cdot \frac{1}{2} + \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x-10)^2} \cdot \frac{1}{2} \end{aligned}$$

Suppose the following:



Then

$$\begin{aligned}
 \mathbb{P}(X) &= \sum_{\theta \in \Theta} \mathbb{P}(X \mid \theta) \mathbb{P}(\theta) \\
 &= \mathbb{P}(X \mid \theta = 0.1) \mathbb{P}(\theta = 0.1) + \mathbb{P}(X \mid \theta = 0.9) \mathbb{P}(\theta = 0.9) \\
 &= \binom{10}{x} 0.1^x (1 - 0.1)^{10-x} \cdot \frac{1}{2} + \binom{10}{x} 0.9^x (1 - 0.9)^{10-x} \cdot \frac{1}{2}
 \end{aligned}$$

What we did here is that we went from  $\theta \sim \text{Beta}(\alpha, \beta)$  to  $X \mid \theta \sim \text{Binom}(n, \theta)$ . Since  $\theta$  is continuous:

$$\begin{aligned}
 \mathbb{P}(X) &= \int_{\Theta} \mathbb{P}(X \mid \theta) \mathbb{P}(\theta) d\theta \\
 &= \int_0^1 \left( \binom{n}{x} \theta^x (1 - \theta)^{n-x} \right) \cdot \left( \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1 - \theta)^{\beta-1} \right) d\theta \\
 &= \binom{n}{x} \frac{1}{B(\alpha, \beta)} \int_0^1 \theta^{x+\alpha-1} (1 - \theta)^{n-x+\beta-1} d\theta \\
 &= \binom{n}{x} \frac{B(x + \alpha, n - x + \beta)}{B(\alpha, \beta)} \\
 &= \text{BetaBinom}(n, \alpha, \beta)
 \end{aligned}$$

This is the Beta-Binomial model. Let  $X$  is a random variable of this model; then  $X \sim \text{BetaBinom}(n, \alpha, \beta)$ .  $\text{Supp}[X] = \{0, 1, \dots, n\}$  and the parameter spaces are:  $n \in \mathbb{N}$ ,  $\alpha > 0$  and  $\beta > 0$ .

$$\begin{aligned}
 \mathbb{E}[X] &= n \frac{\alpha}{\alpha + \beta} \\
 \text{Var}[X] &= \frac{n\alpha\beta}{(\alpha + \beta)^2} \underbrace{\frac{\alpha + \beta + n}{\alpha + \beta + 1}}_{\in [1, n]}
 \end{aligned}$$

Thus the variance is an inflated binomial variance. Let  $\theta = \frac{\alpha}{\alpha+\beta}$ , then  $E[X] = n\theta$ . Let  $B = \frac{\alpha}{\theta} - \alpha$ . Then

$$\begin{aligned}
 \lim_{\alpha \rightarrow \infty} E[X] &= n\theta \\
 \lim_{\alpha \rightarrow \infty} \text{Var}[X] &= \lim_{\alpha \rightarrow \infty} n \frac{\overbrace{\alpha}^{\theta}}{\alpha + \beta} \frac{\overbrace{\beta}^{1-\theta}}{\alpha - \beta} \\
 &= \frac{\alpha + \beta + n}{\alpha + \beta + 1} \\
 &= \underbrace{n\theta(1-\theta)}_{\text{variance of binom}} \lim_{\alpha \rightarrow \infty} \frac{\alpha + \frac{\alpha}{\theta} - \alpha + n}{\alpha + \frac{\alpha}{\theta} - \alpha + 1} \\
 &= n\theta(1-\theta) \lim_{\alpha \rightarrow \infty} \frac{\alpha + n\theta}{\alpha + \theta} = n\theta(1-\theta) \cdot 1 \\
 &= n\theta(1-\theta)
 \end{aligned}$$

From this, as  $\alpha$  gets higher,  $\theta$  gets tighter and becomes degenerate and more like a binomial model.

Suppose  $X \mid \theta \sim \text{Binom}(n, \theta)$ ,  $\theta \sim \text{Beta}(\alpha, \beta)$  and  $\theta \mid X \sim \text{Beta}(\alpha + x, \beta + n - x)$ . Suppose  $X^* \mid X \sim \text{Bern}(\frac{x+\alpha}{n+\alpha+\beta})$  where  $n^* = 1$ . Then:

$$\begin{aligned}
 \mathbb{P}(X^* \mid X) &= \int_{\Theta} \underbrace{\mathbb{P}(X^* \mid \theta)}_{\text{binom}} \underbrace{\mathbb{P}(\theta \mid X)}_{\text{beta}} d\theta \\
 &= \int_0^1 \binom{n^*}{x^*} \theta^{x^*} (1-\theta)^{n^*-x^*} \cdot \frac{1}{B(\alpha+x, \beta+n-x)} \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1} d\theta \\
 &= \text{BetaBinom}(n^*, \alpha+x, \beta+n-x)
 \end{aligned}$$