

Lecture 16

Recall:  $\hat{\sigma}_{MLE}^2 = \frac{SSE}{n}$

so

$$SSE = \hat{\sigma}_{MLE}^2 n$$

\* Note:

$\theta$  is fixed  
 $\sigma^2$  variable

Posterior  $P(\sigma^2 | X, \theta) \propto P(X | \sigma^2, \theta) \cdot P(\sigma^2 | \theta)$

$$X_i \sim X_1, \dots, X_n \quad \text{iid} \quad \propto \left[ \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(X_i - \theta)^2} \right] \cdot P(\sigma^2 | \theta)$$

$$= \left[ \left( \frac{1}{\sqrt{2\pi\sigma^2}} \right)^n e^{-\frac{1}{2\sigma^2} \sum (X_i - \theta)^2} \right] \cdot P(\sigma^2 | \theta)$$

SSE

$$= \left[ \left( \frac{1}{\sqrt{2\pi}} \right)^n \left( \frac{1}{\sqrt{\sigma^2}} \right)^n e^{-\frac{\hat{\sigma}_{MLE}^2 n}{2\sigma^2}} \right] \cdot P(\sigma^2 | \theta) \propto \left[ \left( \frac{1}{\sqrt{\sigma^2}} \right)^n e^{-\left( \frac{n\hat{\sigma}_{MLE}^2}{2} \right) \frac{1}{\sigma^2}} \right] \cdot K(\sigma^2 | \theta)$$

$K(X | \sigma^2, \theta)$

$$= \left[ \sigma^{\frac{n}{2}} e^{-\left( \frac{n\hat{\sigma}^2}{2} \right) \frac{1}{\sigma^2}} \right] \cdot \left[ \sigma^2 e^{-\beta \frac{1}{\sigma^2}} \right] \quad \text{conjugate prior}$$

$\sigma^2 | \theta \sim \text{Inversegamma}(\alpha, \beta)$

$$= \left[ \sigma^{\frac{n}{2} + \alpha - 1} e^{-\left( \frac{n\hat{\sigma}^2}{2} + \beta \right) \frac{1}{\sigma^2}} \right] \propto \text{Invgamma} \left( \frac{n}{2} + \alpha, \frac{n\hat{\sigma}^2}{2} + \beta \right)$$

$P(\sigma^2 | X, \theta) =$

Note:  $\sigma^2 | \theta \sim \text{Inv gamma}(\alpha, \beta)$

$$\sigma^2 | \theta, X \sim \text{Inv gamma} \left( \frac{n}{2} + \alpha, \frac{n\hat{\sigma}^2}{2} + \beta \right)$$

canonical: let  $\alpha = \frac{n_0}{2}, \beta = \frac{n_0 \hat{\sigma}_0^2}{2}$

Prior  $\left[ \sigma^2 | \theta \sim \text{Inv gamma} \left( \frac{n_0}{2}, \frac{n_0 \hat{\sigma}_0^2}{2} \right) \right] \quad \text{pseudo trials / variance}$

Posterior  $\left[ \sigma^2 | \theta, X \sim \text{Inv gamma} \left( \frac{n+n_0}{2}, \frac{n\hat{\sigma}^2 + n_0 \hat{\sigma}_0^2}{2} \right) \right]$

$\bullet X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$   
Point Estimation /  $\bullet \sigma^2 | \theta, X_i \sim \text{Inv gamma} \left( \underbrace{\frac{n+n_0}{2}}_{\alpha'}, \underbrace{\frac{n\hat{\sigma}^2 + n_0\hat{\sigma}_0^2}{2}}_{\beta'} \right)$

Note  
 $\hat{\sigma}^2 = \frac{\text{SSE}}{n}$

if  $Y \sim \text{Inv gamma}(\alpha, \beta)$   
 $\bullet E[Y] = \frac{\beta}{\alpha-1}$   
 $\sigma_{\text{MMSE}}^2 = E[\sigma^2 | \theta, X] = \frac{\beta'}{\alpha'-1} = \frac{(n\hat{\sigma}^2 + n_0\hat{\sigma}_0^2)/2}{(n+n_0)/2 - 1} \cdot \left(\frac{2}{2}\right)$

$\sigma_{\text{MMSE}}^2 = \frac{n\hat{\sigma}^2 + n_0\hat{\sigma}_0^2}{n+n_0-2}$

$\bullet \text{Mode}[Y] = \frac{\beta}{\alpha+1}$   
 $\sigma_{\text{MAP}}^2 = \text{Mode}[\sigma^2 | \theta, X] = \frac{(n\hat{\sigma}^2 + n_0\hat{\sigma}_0^2)/2}{(n+n_0)/2 + 1} \cdot \left(\frac{2}{2}\right)$

$\sigma_{\text{MAP}}^2 = \frac{n\hat{\sigma}^2 + n_0\hat{\sigma}_0^2}{n+n_0+2}$

$\bullet \text{Med}[Y] = \text{NO closed form}$   
 $\sigma_{\text{MAE}}^2 = q_{\text{inv gamma}}(0.5, \alpha', \beta')$

### uninformative priors /

• Jeffrey's:  
 $f(\sigma^2; \theta, X_i) = e^{-\frac{\text{SSE}}{2\sigma^2}} \frac{1}{\sigma^2} \cdot \sigma^2^{-\frac{n}{2}}$

$l = \ln f = -\frac{\text{SSE}}{2} \frac{1}{\sigma^2} - \frac{n}{2} \ln(\sigma^2)$

$\frac{d}{d\sigma^2} l' = \frac{\text{SSE}}{2} \frac{1}{(\sigma^2)^2} - \frac{n}{2} \cdot \frac{1}{\sigma^2}$

$\frac{d}{d\sigma^2} l'' = -\text{SSE} \frac{1}{(\sigma^2)^3} + \frac{n}{2} \cdot \frac{1}{(\sigma^2)^2}$

$I(\theta) = E_x[-l''(\sigma^2; X_i, \theta)]$

$= E_x \left[ \frac{\text{SSE}}{(\sigma^2)^3} - \frac{n}{2(\sigma^2)^2} \right]$

$= E_x[\text{SSE}] (\sigma^2)^{-3} - \frac{n}{2} (\sigma^2)^{-2}$

$E \left[ \sum_{i=1}^n (X_i - \theta)^2 \right] = \sum_{i=1}^n E[(X_i - \theta)^2] = \sum_{i=1}^n \sigma^2 = n\sigma^2$

$= n\sigma^2 (\sigma^2)^{-3} - \frac{n}{2} (\sigma^2)^{-2}$

$= n(\sigma^2)^{-2} - n/2 (\sigma^2)^{-2}$

$\bullet I(\theta) = (\sigma^2)^{-2} \left( \frac{2n}{2} - \frac{n}{2} \right) = \left[ (\sigma^2)^{-2} \left( \frac{n}{2} \right) \right]$

$\bullet P_J(\sigma^2) \propto \sqrt{I(\theta)} = \sqrt{\frac{n}{2} (\sigma^2)^{-2}} \propto \sqrt{(\sigma^2)^{-2}} = (\sigma^2)^{-1} = (\sigma^2)^{0-1} e^0$

$\bullet P_J(\sigma^2) \propto \text{Inv gamma}(0, 0)$



• Haldane:  $\sigma^2 | \theta \sim \text{Invgamma}(0, 0)$

$\sigma^2 | \theta, X_i \sim \text{Invgamma}\left(\frac{n}{2} + 0, \frac{n\hat{\sigma}^2}{2} + 0\right)$  always proper  
 $n \geq 1, \hat{\sigma}^2 > 0$

• Indifference: ?

C.  $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\theta, \sigma^2)$  let:  $\theta, \sigma^2$  are both inferred targets

$$\text{Posterior} / P(\theta, \sigma^2 | X_i) \propto P(X | \theta, \sigma^2) \cdot P(\theta, \sigma^2)$$

$$= \left[ \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(X_i - \theta)^2} \right] \cdot P(\theta, \sigma^2)$$

$$\propto \left[ \sigma^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum (X_i - \theta)^2} \right] \cdot P(\theta, \sigma^2)$$

Not invgamma b/c  $\theta$  is unknown; Kernel of Normal-Inverse gamma

$$\sum (X_i - \theta)^2 = \sum (X_i - \bar{x} - \theta + \bar{x})^2 = \sum ((X_i - \bar{x}) + (\bar{x} - \theta))^2$$

$$= \sum [(X_i - \bar{x})^2 + 2(X_i - \bar{x})(\bar{x} - \theta) + (\bar{x} - \theta)^2]$$

$$= \sum (X_i - \bar{x})^2 + 2\sum (X_i - \bar{x})(\bar{x} - \theta) + \sum (\bar{x} - \theta)^2$$

$$+ 2(-\bar{x}n\theta + n\bar{x}^2 + n\bar{x}\theta - n\bar{x}^2)$$

$$= \sum (X_i - \bar{x})^2 - 0 + \sum (\bar{x} - \theta)^2$$

$$\sum (X_i - \theta)^2 = s^2(n-1) + n(\bar{x} - \theta)^2$$

Note  
 $\bar{x} = \frac{\sum X_i}{n}$

$$\bullet \sum X_i = \bar{x}n$$

$$s^2 = \frac{\sum (X_i - \bar{x})^2}{n-1}$$

$$\bullet s^2(n-1) = \sum (X_i - \bar{x})^2$$

$s^2$ : sample variance

$$P(\theta, \sigma^2 | X_i) \propto \sigma^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum (X_i - \theta)^2} \cdot [P(\theta, \sigma^2)]$$

$$\propto \underset{\alpha}{\sigma^{-\frac{n}{2}}} e^{\underset{\beta}{-\left(\frac{s^2(n-1)}{2}\right) \frac{1}{\sigma^2}}} e^{\underset{\lambda}{-\frac{n}{2\sigma^2}} \underset{\mu}{(\bar{x} - \theta)^2}} \cdot [P(\theta, \sigma^2)]$$

$$\bullet P(\theta, \sigma^2 | X) \propto \text{NormInvgamma}\left(\mu = \bar{x}, \lambda = n, \alpha = \frac{n}{2}, \beta = \frac{(n-1)s^2}{2}\right)$$

\*\*  $P(\theta, \sigma^2)$  must be a norminvgamma also