

## Lecture 18

Note  $X_1, \dots, X_n \stackrel{iid}{\sim} N(\theta, \sigma^2)$   $\theta, \sigma^2$  unknown

$$P(\theta, \sigma^2) \propto 1/\sigma^2 \quad (\text{Jeffrey's Prior}) \quad - \text{Prior}$$

$$P(\theta, \sigma^2 | X) = \text{Norm Inv Gamma}(\mu, \lambda, \kappa, \beta) \quad - \text{Posterior}$$

$$P(\theta | X, \sigma^2) = N(\bar{x}, \frac{\sigma^2}{n})$$

$$P(\sigma^2 | X, \theta) = \text{Inv Gamma}(\frac{n}{2}, \frac{n\hat{\sigma}^2}{2})$$

$$P(\sigma^2 | X) = \text{Inv Gamma}(\frac{n-1}{2}, \frac{(n-1)S^2}{2})$$

$$P(\theta | X) = T_{n-1}(\bar{x}, S/\sqrt{n})$$

} Marginal  
Posteriors

$$CR_{95\%, \theta} = [qt(0.025, n-1, \bar{x}, \frac{S}{\sqrt{n}}), qt(0.975, n-1, \bar{x}, \frac{S}{\sqrt{n}})]$$

• Posterior Predictive :

$$X^* | \theta, \sigma^2 \sim N(\theta, \sigma^2)$$

$$X^* | X \sim ?$$

$$P(X^* | X) = \int_{\sigma^2=0}^{\infty} \int_{\theta=-\infty}^{\infty} P(X^* | \theta, \sigma^2) P(\theta, \sigma^2 | X) d\theta d\sigma^2$$

Bayes Rule  
easier one

$$= \int_{\sigma^2} \int_{\theta} \left\{ \underbrace{P(X^* | \theta, \sigma^2)}_{\text{Norm}} \underbrace{P(\theta | X, \sigma^2)}_{\text{Norm}} \underbrace{P(\sigma^2 | X)}_{\text{Inv Gamma}} \right\} \text{ sampling from here}$$

$$= \int_{\sigma^2} \int_{\theta} \left( \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(X^* - \theta)^2} \right) \left( \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(\theta - \bar{x})^2} \right) \left( \frac{((n-1)S^2)^{\frac{n-1}{2}}}{\Gamma(\frac{n-1}{2})} (\sigma^2)^{-\frac{(n-1)}{2}-1} e^{-\frac{(n-1)S^2}{2\sigma^2}} \right) d\theta d\sigma^2$$

$$\propto \int_{\sigma^2} \int_{\theta} \frac{1}{\sigma^2} \frac{1}{\sigma^2} \frac{1}{\sigma^2} e^{-\frac{1}{2\sigma^2}(X^* - \theta)^2} e^{-\frac{n}{2\sigma^2}(\theta - \bar{x})^2} e^{-\frac{1}{2\sigma^2}(n-1)S^2} d\theta d\sigma^2$$

∴

$$= \int \sigma^2^{-\frac{(n+1)}{2}-1} e^{-\frac{n}{2\sigma^2}} \int e^{\frac{X^* + n\bar{x}}{\sigma^2} \theta} e^{-\frac{(n+1)}{2\sigma^2} \theta^2} d\theta d\sigma^2 \propto$$

Kernel of normal



$$P(x^*|x) \propto T_{n-1}\left(\bar{x}, \sqrt{s^2 \frac{n+1}{n}}\right)$$

Note

As  $n \rightarrow \infty$ 

$$P(x^*|x) \Rightarrow N(\theta, s^2)$$

Sampling  $x^*|x$ :

- 1) draw  $\sigma_0^2$  from  $\text{InvGamma}\left(\frac{n-1}{2}, \frac{(n-1)s^2}{2}\right)$   $P(\sigma^2|x)$
- 2) draw  $\theta_0$  from  $N(\bar{x}, \sigma_0^2/n)$   $P(\theta|\sigma^2 x)$
- 3) draw  $x^*$  from  $N(\theta_0, \sigma_0^2)$
- 4) Return  $x^*$

Note:  $P(\theta, \sigma^2|x) = P(x|\theta, \sigma^2) \cdot P(\theta, \sigma^2)$  ← Jeffrey's prior must ALSO be NormInvGamma

$\uparrow$  NormInvGamma       $\uparrow$

$$P(\theta, \sigma^2) = P(\theta|\sigma^2) P(\sigma^2)$$

$$\uparrow \quad \uparrow$$

$$N\left(\mu_0, \tau^2 = \left(\frac{\sigma^2}{\sqrt{n_0}}\right)^2\right) \quad \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right)$$

- Consider  $\theta$  and  $\sigma^2$  are disjoint such that:

$$P(\sigma^2) = \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right)$$

but

$$P(\theta|\sigma^2) = P(\theta) = N\left(\mu_0, \tau^2 = \frac{\sigma^2}{n_0}\right) \quad * \text{Not dependent on } \sigma^2$$

$$P(\theta, \sigma^2) = P(\theta) \cdot P(\sigma^2) \neq \text{NormInvGamma} \quad \text{so,}$$

$$P(\theta, \sigma^2|x) \neq \text{NormInvGamma}$$

$$P(\theta, \sigma^2|x) \propto P(x|\theta, \sigma^2) P(\theta, \sigma^2) \quad \left[ \text{AND} \right]$$

$$\propto P(x|\theta, \sigma^2) P(\theta) P(\sigma^2) \quad \left[ \propto P(\theta|\sigma^2 x) P(\sigma^2|x) \right]$$

$$\begin{array}{ccc} \uparrow & \uparrow & \uparrow \\ x_1, \dots, x_n \sim N(\theta, \sigma^2) & N(\mu_0, \tau^2) & \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right) \\ \text{NormInvGamma} & & \\ \left( \begin{array}{l} \theta \text{ variable} \\ \sigma^2 \text{ variable} \end{array} \right) & \left( \begin{array}{l} \theta \text{ variable} \\ \sigma^2 \text{ fixed} \end{array} \right) & \left( \begin{array}{l} \propto P(\sigma^2|\theta) \\ \theta \text{ fixed} \\ \sigma^2 \text{ variable} \end{array} \right) \end{array}$$



Note:  $\theta, \sigma^2$  are variables

$$= \left( \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x_i - \theta)^2} \right) \left( \frac{1}{\sqrt{2\pi\tau^2}} e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2} \right) \left( \frac{\left(\frac{n_0\sigma_0^2}{2}\right)^{\frac{n_0}{2}}}{\Gamma\left(\frac{n_0}{2}\right)} \sigma^2^{-\left(\frac{n_0}{2}+1\right)} e^{-\left(\frac{n_0\sigma_0^2}{2}\right)\frac{1}{\sigma^2}} \right)$$

$P(X|\theta, \sigma^2)$                        $P(\theta)$                        $P(\sigma^2)$

$$\propto \left( \sigma^2^{-\frac{n}{2}} e^{-\frac{1}{2\sigma^2} \sum (x_i - \theta)^2} \right) \left( e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2} \right) \left( \sigma^2^{-\left(\frac{n_0}{2}+1\right)} e^{-\left(\frac{n_0\sigma_0^2}{2}\right)\frac{1}{\sigma^2}} \right)$$

$$\propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n(\bar{x} - \theta)^2 + n_0\sigma_0^2 \right)} e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2}$$

SSE

$$\propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n_0\sigma_0^2 \right)} e^{-\frac{1}{2\sigma^2} n(\bar{x} - \theta)^2} e^{-\frac{1}{2\tau^2}(\theta - \mu_0)^2}$$

(fall out)                      (fall out)

$$\propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n_0\sigma_0^2 \right)} e^{-\frac{n\bar{x}^2}{2\sigma^2} + \frac{2n\bar{x}\theta}{2\sigma^2} - \frac{n\theta^2}{2\sigma^2} - \frac{\theta^2}{2\tau^2} + \frac{2\theta\mu_0}{2\tau^2} - \frac{\mu_0^2}{2\tau^2}}$$

$$\propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n_0\sigma_0^2 - n\bar{x}^2 \right)} e^{-\frac{\mu_0^2}{2\tau^2}} \left[ e^{-\left(\frac{n}{2\sigma^2} + \frac{1}{2\tau^2}\right)\theta^2} e^{\left(\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2}\right)\theta} \right]$$

(constant)

$$\propto N(\theta_p, \sigma_p^2)$$

$$\propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n_0\sigma_0^2 - n\bar{x}^2 \right)} \left[ \frac{1}{\sqrt{2\pi\sigma_p^2}} e^{-\frac{1}{2\sigma_p^2}(\theta - \theta_p)^2} \right]$$

$$P(\theta, \sigma^2 | X) \propto \sigma^2^{-\frac{n}{2} - \left(\frac{n_0}{2}+1\right)} e^{-\frac{1}{2\sigma^2} \left( (n-1)S^2 + n_0\sigma_0^2 - n\bar{x}^2 \right)} \left[ \left( \sigma_p^2 \right)^{-\frac{1}{2}} e^{-\frac{\theta_p^2}{2\sigma_p^2}} \right] \cdot \theta_p = \frac{\frac{n\bar{x}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

Kernel of some r.v.

$$\sigma_p^2 = \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}}$$

$$\text{recall: } P(\theta, \sigma^2 | X) \propto P(\theta | X, \sigma^2) P(\sigma^2 | X)$$

$$\propto P(\sigma^2 | X, \theta) P(\theta | X)$$

\* One part of kernel is Normal, only  $P(\theta | X, \sigma^2)$  fits, but  $P(\sigma^2 | X)$  is now an unknown kernel

\*\* disjoint  $\theta, \sigma^2$  creates a Semi-conjugate Model

• Sampling from  $P(\theta, \sigma^2 | X)$  when  $\theta, \sigma^2$  are disjoint

- 1) choose  $\sigma_0^2$  from  $K(\sigma^2 | X)$  <sup>\*\*</sup> unknown portion
- 2) choose  $\theta_0$  from  $N(\theta_p, \sigma_p^2)$
- 3) return  $\langle \theta_0, \sigma_0^2 \rangle$

\*\* Sampling from  $K(\sigma^2 | X)$  ; note  $P(\sigma^2 | X) = c K(\sigma^2 | X)$

Grid Sampling

- 1) pick Grid:  $\sigma_{\min}^2, \sigma_{\max}^2, g, \Delta\sigma^2$

$$G = \langle \sigma_{\min}^2, \sigma_{\min}^2 + \Delta\sigma^2, \dots, \sigma_{\min}^2 + (g-1)\Delta\sigma^2, \sigma_{\max}^2 \rangle$$

- 2) compute  $K(\sigma_g^2 | X) \quad \forall \sigma_g^2 \in G$

- 3) Approximate  $c$  ;  $c \approx \frac{1}{\sum_{\sigma_g^2 \in G} K(\sigma_g^2 | X)}$

- 4) CDF:  $F(\sigma_0^2 | X) \approx \sum_{\sigma^2 \in G : \sigma^2 < \sigma_0^2} c K(\sigma^2 | X)$

- 5) draw  $y$  from  $U(0,1)$

$$\text{compute } \sigma_0^2 = \min_{\sigma^2 \in G} F(\sigma^2) \geq y$$