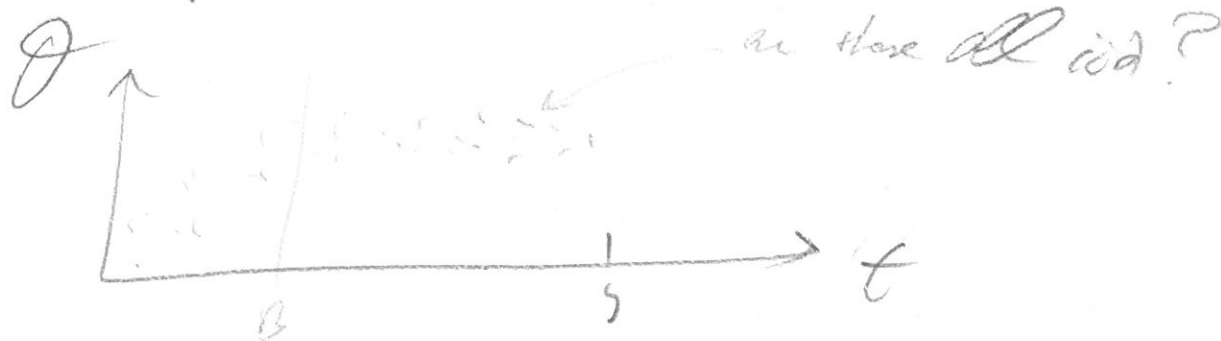


Lee 22 5/7/19 Math 381

Another problem (this is solvable)



Logic not independent since

Q_t depends on Q_{t-1} . The Q_t also depends on Q_{t-2} and Q_{t-3} ...

But at some points the dependence is so low it's negligible.

How can we measure it?

Back to basic stat:

Two r.v.'s X_1, X_2

$$\sigma_{12} = \text{Cov}[X_1, X_2] := E[(X_1 - \mu_1)(X_2 - \mu_2)] \quad \text{degree of linear dependence}$$

$$\rho = \text{Corr}[X_1, X_2] = \frac{\text{Cov}[X_1, X_2]}{\text{SE}(X_1) \text{SE}(X_2)} \quad \text{Measure size this but } \in (-1, 1) \text{ and unitless}$$

to estimate these parameters,

$$s_{12} = \frac{1}{n-1} \sum (x_{1i} - \bar{x}_1)(x_{2i} - \bar{x}_2)$$

$$s_1 = \sqrt{\frac{1}{n-1} \sum (x_{1i} - \bar{x}_1)^2}, \quad s_2 = \sqrt{\frac{1}{n-1} \sum (x_{2i} - \bar{x}_2)^2}$$

$$r = \frac{s_{12}}{s_1 s_2} = \frac{\sum (x_{1i} - \bar{x}_1)(x_{2i} - \bar{x}_2)}{\sqrt{\sum (x_{1i} - \bar{x}_1)^2} \sqrt{\sum (x_{2i} - \bar{x}_2)^2}}$$

Back to our problem... how does θ_t depend on θ_{t-1} ?
 Estimate its correlation...

[2]

$$r_{\theta 1} := \frac{\sum_{t=b+2}^S (\theta_t - \bar{\theta})(\theta_{t-1} - \bar{\theta})}{\sum_{t=b+1}^S (\theta_t - \bar{\theta})^2}$$

$$\bar{\theta} := \frac{1}{S-b} \sum_{t=b+1}^S \theta_t$$

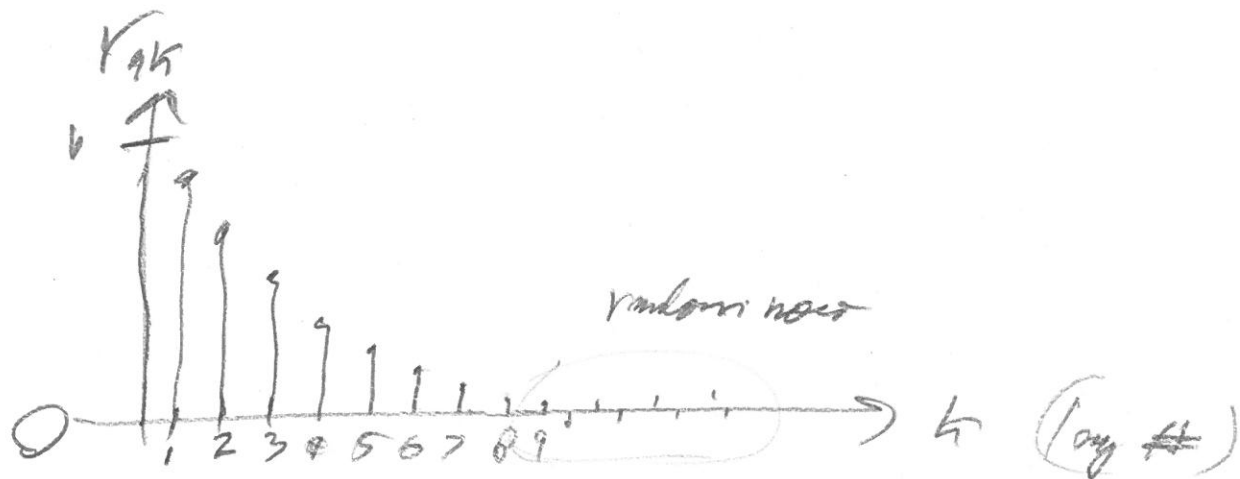
Autocorrelation: correlation with itself, lag = 1: one step back

$r_{\theta 1}$? LARGE & Positive

$$r_{\theta 2} = \frac{\sum_{t=b+3}^S (\theta_t - \bar{\theta})(\theta_{t-2} - \bar{\theta})}{\sum_{t=b+1}^S (\theta_t - \bar{\theta})^2}$$

$$r_{\theta k} = \frac{\sum_{t=b+k+1}^S (\theta_t - \bar{\theta})(\theta_{t-k} - \bar{\theta})}{\sum_{t=b+1}^S (\theta_t - \bar{\theta})^2}$$

For some lag k , $r_{\theta k} \approx 0$. Why? Events simpler are \approx independent. How to check? Create autocorrelation plot.



At $k=9$ samples look independent enough.

Now, if we take every $T=9^{\text{th}}$ sample, we have iid draws from $P(\theta_1, \dots, \theta_p | x)$. Selecting k & samples in between is called "striding".

After "striding" and "thinning" we have N samples:

$$D = \left\{ \begin{bmatrix} \theta_{1,b+1} \\ \theta_{2,b+1} \\ \vdots \end{bmatrix}, \begin{bmatrix} \theta_{1,b+1+T} \\ \theta_{2,b+1+T} \\ \vdots \end{bmatrix}, \begin{bmatrix} \theta_{1,b+1+2T} \\ \theta_{2,b+1+2T} \\ \vdots \end{bmatrix}, \dots, \begin{bmatrix} \theta_{1,b+1+NT} \\ \theta_{2,b+1+NT} \\ \vdots \end{bmatrix} \right\}$$

↑

bracket means no need to order these samples anymore, they're all iid!!

[7]

How to do Bayesian inference? EASY!

$$\hat{\theta}_{j, \text{mean}} = E(\theta_j | x) \approx \frac{1}{N} \sum_{\theta_j \in \mathcal{M}} \theta_j$$

$$\hat{\theta}_{j, \text{median}} = \text{Med}(\theta_j | x) \approx \text{SampleMedian}(\{\theta_j \in \mathcal{M}\})$$

$$CR_{\theta_j, 1-\alpha} \approx [\text{SampleQuantile}[\frac{\alpha}{2}, \{\theta_j \in \mathcal{M}\}],$$

$$H_0: \theta_j \in \Theta; \quad \text{SampleQuantile}[1-\frac{\alpha}{2}, \{\theta_j \in \mathcal{M}\}]]$$

$$p_{\text{val}} = P(H_0 | x) \approx \frac{1}{N} \sum_{\theta_j \in \mathcal{M}} \mathbb{1}_{\theta_j \in \Theta}; \quad \text{approximate integral}$$

$$P(x^* | x) = \int_{\Theta} P(x^* | \vec{\theta}) P(\vec{\theta} | x) d\vec{\theta} \approx \text{the following sampling algorithm:}$$

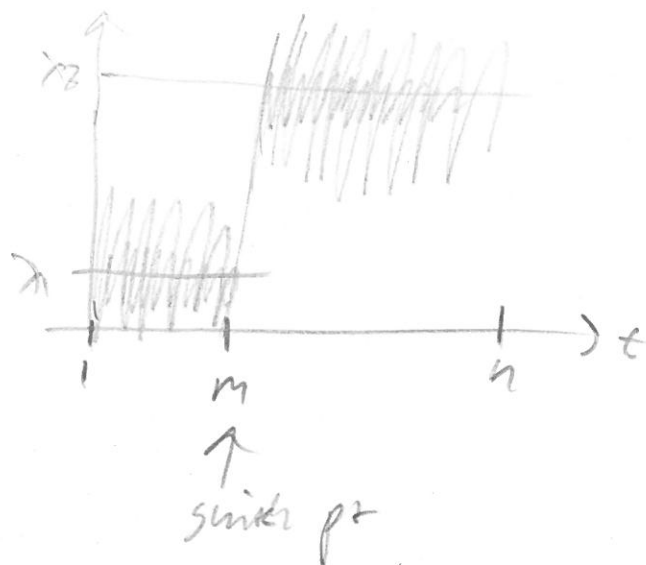
- I select $\vec{\theta}$ from \mathcal{M} at random
- II Draw x_{sim}^* from $P(x^* | \vec{\theta})$ and ship x_{sim}^*
- III Repeat steps I-II S times

1.12.16

Change pt. model

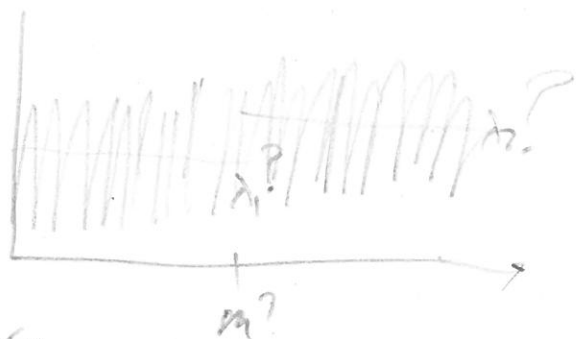
of phone calls modeled as n

for a period of time, you have a Poisson iid model with λ_1
and then it switches to a Poisson iid model with λ_2



3 params!
 λ_1, λ_2, m

this is known! But what if



You would like to capture the uncertainty in all these unknown parameters
what is data? X_1, X_2, \dots, X_n is # of phone calls at each time $t=1, 2, \dots, n$

$$P(\lambda_1, \lambda_2, m | X_1, \dots, X_n) \propto P(X_1, \dots, X_n | \lambda_1, \lambda_2, m) P(\lambda_1, \lambda_2, m)$$

What does the likelihood look like? *Conditional assumes priors known* $P(X|D)$ (6)

$$X_1, \dots, X_m | \lambda_1 \sim \text{Poisson}(\lambda_1)$$

$$X_{m+1}, \dots, X_n | \lambda_2 \sim \text{Poisson}(\lambda_2) \quad X_1, \dots, X_m \text{ iid of } X_{m+1}, \dots, X_n$$

$$P(X_1, \dots, X_n | \lambda_1, \lambda_2, m) = P(X_1, \dots, X_m | \lambda_1) P(X_{m+1}, \dots, X_n | \lambda_2)$$

$$= \prod_{t=1}^m \frac{e^{-\lambda_1} \lambda_1^{x_t}}{x_t!} \prod_{t=m+1}^n \frac{e^{-\lambda_2} \lambda_2^{x_t}}{x_t!}$$

prior

$$P(\lambda_1, \lambda_2, m) = P(\lambda_1) P(\lambda_2) P(m)$$

any reason not to?? *prob. m*

Should we use uninformative priors? Sure...

$$P(\lambda_1) = \text{Gamma}(\frac{1}{2}, \frac{1}{2}) \propto 1$$

$$P(\lambda_2) = \text{Gamma}(\frac{1}{2}, \frac{1}{2}) \propto 1 \quad \text{Laplace!}$$

"

$$m \in \{1, \dots, n\} \quad P(m) = \frac{1}{n} \quad \text{discrete uniform from Lec 2 or 3!}$$

put it all together...

$$P(\lambda_1, \lambda_2, m | X_1, \dots, X_n) \propto \prod_{t=1}^m \frac{e^{-\lambda_1} \lambda_1^{x_t}}{x_t!} \prod_{t=m+1}^n \frac{e^{-\lambda_2} \lambda_2^{x_t}}{x_t!} (1)(1)(\frac{1}{n})$$

$$\propto \prod_{t=1}^m e^{-\lambda_1} \lambda_1^{x_t} \prod_{t=m+1}^n e^{-\lambda_2} \lambda_2^{x_t}$$

$$\frac{\prod_{t=1}^m x_t!}{\prod_{t=1}^m x_t!} \propto e^{-m\lambda_1} \lambda_1^{\sum_{t=1}^m x_t} e^{-(n-m)\lambda_2} \lambda_2^{\sum_{t=m+1}^n x_t}$$

? known
 \propto distr?? No!

Ed... 3-d good sample? No... Gibbs? let's see...

$$P(\lambda_1 | x_1, \dots, x_n, \lambda_2, m) = \lambda_1^{\left(\sum_{t=1}^m x_t + 1\right) - 1} e^{-m\lambda_1} \propto \text{Gamma}\left(\sum_{t=1}^m x_t + 1, m\right)$$

$$P(\lambda_2 | x_1, \dots, x_n, \lambda_1, m) = \lambda_2^{\left(\sum_{t=1}^n x_t + 1\right) - 1} e^{-(n-m)\lambda_2} \propto \text{Gamma}\left(\sum_{t=1}^n x_t + 1, n-m\right)$$

$$P(m | x_1, \dots, x_n, \lambda_1, \lambda_2) = e^{m(\lambda_2 - \lambda_1)} \lambda_1^{\sum_{t=1}^m x_t} \lambda_2^{\sum_{t=m+1}^n x_t} \propto ? \text{ More...}$$

$\frac{1}{n} \sum_{m=1}^n k(m | -)$
 $\frac{1}{n} \sum_{m=1}^n k(m | -)$

Easy to find sample! Only n pts in support!
 No need to select $x_{\text{init}}, x_{\text{new}}$ or Δ and only 1 den
 necessary \Rightarrow No disadvantages...