

## Lecture 13

03/24 et [MATH 341]

 $\mathbb{T}$ : Poisson( $\theta$ ),  $\theta \sim \text{Gamma}(\alpha, \beta)$  (continuous)  $\Rightarrow \theta|X \sim \text{Gamma}(\mathbb{I}\alpha + \alpha, n + \beta)$ 

$$n_* = 1 \quad P(X_*(x)) = \int_{\mathbb{R}^+} P(X_*(\theta)) P(\theta|x) d\theta$$

$$= \int_0^\infty \left( \frac{e^{-\theta} \theta^{x_*}}{x_*!} \right) \left( \frac{(n+\beta)^{\mathbb{I}\alpha + \alpha}}{\Gamma(\mathbb{I}\alpha + \alpha)} \theta^{\mathbb{I}\alpha + \alpha - 1} e^{-(n+\beta)\theta} \right) d\theta$$

$$= \frac{(n+\beta)^{\mathbb{I}\alpha + \alpha}}{x_*! \Gamma(\mathbb{I}\alpha + \alpha)} \int_0^\infty \theta^{\mathbb{I}\alpha + x_* + \alpha - 1} e^{-(n+\beta+1)\theta} d\theta$$

$$\text{let } t = (n+\beta+1)\theta \Rightarrow \frac{dt}{d\theta} = n+\beta+1, \quad \theta = \frac{t}{n+\beta+1}, \quad d\theta = \frac{dt}{n+\beta+1}$$

$$= \frac{(n+\beta)^{\mathbb{I}\alpha + \alpha}}{x_*! \Gamma(\mathbb{I}\alpha + \alpha)} \int_0^\infty \left( \frac{t}{(n+\beta+1)} \right)^{\mathbb{I}\alpha + x_* + \alpha - 1} e^{-t} \frac{dt}{n+\beta+1}$$

$$= \frac{(n+\beta)^{\mathbb{I}\alpha + \alpha}}{x_*! \Gamma(\mathbb{I}\alpha + \alpha)} \cdot \frac{1}{(n+\beta+1)^{\mathbb{I}\alpha + x_* + \alpha - 1}} \cdot \frac{1}{(n+\beta+1)} \left( \int_0^\infty t^{\mathbb{I}\alpha + x_* + \alpha - 1} e^{-t} dt \right)$$

$$= \frac{(n+\beta)^{\mathbb{I}\alpha + \alpha}}{x_*! \Gamma(\mathbb{I}\alpha + \alpha)} \cdot \frac{\Gamma(\mathbb{I}\alpha + x_* + \alpha)}{(n+\beta+1)^{\mathbb{I}\alpha + x_* + \alpha}} = \Gamma(\mathbb{I}\alpha + x_* + \alpha)$$

$$= (n+\beta+1)^{\mathbb{I}\alpha + \alpha} (n+\beta+1)^{x_*}$$

$$= \left( \frac{n+\beta}{n+\beta+1} \right)^{\mathbb{I}\alpha + \alpha} \left( \frac{1}{n+\beta+1} \right)^{x_*} \frac{\Gamma(\mathbb{I}\alpha + x_* + \alpha)}{x_*! \Gamma(\mathbb{I}\alpha + \alpha)}$$

$$\text{let } p := \frac{n+\beta}{n+\beta+1} \in (0, 1) \quad \text{let } r := \mathbb{I}\alpha + \alpha$$

$$\Rightarrow 1-p = \frac{1}{n+\beta+1} \in (0, 1)$$

$$= \frac{\Gamma(x_* + r)}{x_*! \Gamma(r)} p^r (1-p)^{x_*} = \text{ExtNegBin}(r, p)$$

[extended negative binomial model]

• If  $r = \mathbb{I}\alpha + \alpha \in \mathbb{N}_0 \Rightarrow \alpha \in \mathbb{N}_0$ 

$$\binom{x_* + r - 1}{x_*} p^r (1-p)^{x_*} = \text{NegBin}(r, p)$$

recall  $\Gamma(x) = (x-1)!$  if  $x \in \mathbb{N}$

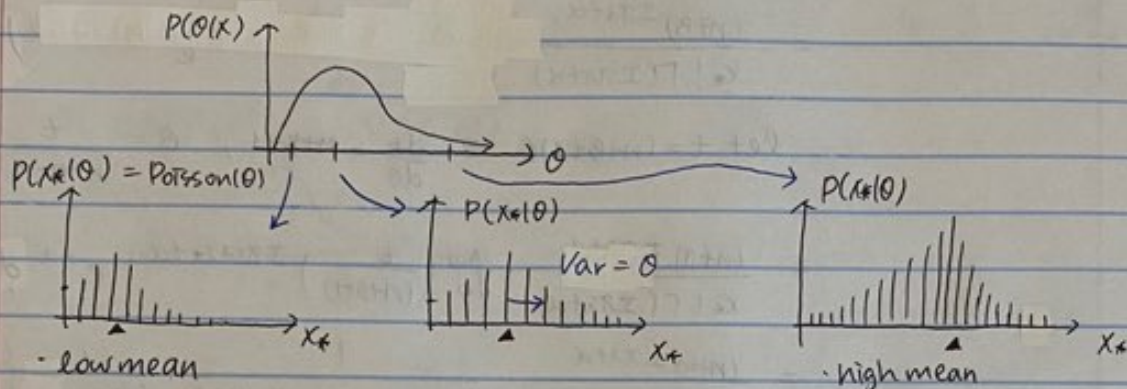


$$X_1, \dots, X_r; p \stackrel{\text{iid}}{\sim} \text{geom}(p) := (1-p)^x p \quad E[X_i] = \frac{p}{1-p}$$

$$\sum_{i=1}^r X_i \sim \text{NegBin}(r, p)$$

$$E[\sum X_i] = r \cdot \frac{p}{1-p}$$

$X_* | X \sim \text{ExtNegBin}(r, p) : \text{Overdispersed Poisson.}$



$$E[X_* | X] = r \cdot \frac{p}{1-p} = \mu$$

$$\text{Var}[X_* | X] = \frac{p \cdot r}{(1-p)^2} = \frac{1}{1-p} \mu$$

$$\frac{1}{1-p} \in (1, \infty)$$

Normal Model.  $X \sim N(\theta, b^2) = N(\theta_1, \theta_2)$

$$\dim[\theta] = 2$$

• pretend we know  $b^2$ , we want inference for the  $\theta$ .

kernel practice...

$$N(\theta, b^2) = P(X|\theta, b^2) = \frac{1}{\sqrt{2\pi b^2}} e^{-\frac{1}{2b^2} (x-\theta)^2} \propto e^{-\frac{1}{2b^2} (x-\theta)^2}$$

$$= e^{-\frac{x^2}{2b^2}} e^{\frac{x\theta}{b^2}} e^{-\frac{\theta^2}{2b^2}}$$

$$\propto e^{-\frac{x^2}{2b^2}} e^{\frac{x\theta}{b^2}}$$

$$= e^{ax - bx^2} \text{ where } a = \frac{\theta}{b^2} \text{ and } b = \frac{1}{2b^2} > 0$$

$$\Rightarrow b^2 = \frac{1}{2b}, \quad a = \frac{a}{2b}$$

$$\propto N\left(\frac{a}{2b}, \frac{1}{2b}\right)$$

" $\theta$ " " $b$ "

"Posterior"

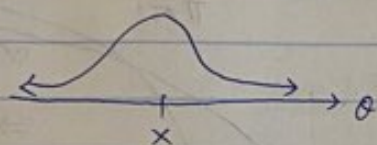
$$P(\theta | x, b^2) = \frac{1}{\sqrt{2\pi b^2}} e^{-\frac{1}{2b^2}(x-\theta)^2}$$

$$\propto e^{-\frac{1}{2b^2}(x-\theta)^2} = e^{-\frac{x^2}{2b^2}} e^{\frac{x\theta}{b^2}} e^{-\frac{\theta^2}{2b^2}}$$

$$\propto e^{\frac{x\theta}{b^2} - \frac{\theta^2}{2b^2}} = e^{a\theta - b\theta^2}, \text{ where } a = \frac{x}{b^2}, b = \frac{1}{2b^2}$$

$$\propto N\left(-\frac{a}{2b}, \frac{1}{2b}\right) = N(x, b^2) \quad \text{where } \frac{a}{2b} = \frac{\frac{x}{b^2}}{2 \cdot \frac{1}{2b^2}} = x$$

$$\Rightarrow P(\theta | x, b^2) = N(x, b^2)$$



$$\tilde{F}: x_1, \dots, x_n; \theta, b^2 \sim N(\theta, b^2)$$

$$P(x | \theta, b^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi b^2}} e^{-\frac{1}{2b^2}(x_i - \theta)^2}$$

$$= \left(\frac{1}{\sqrt{2\pi b^2}}\right)^n e^{-\frac{1}{2b^2} \sum_{i=1}^n (x_i - \theta)^2}$$

$$= \left(\frac{1}{\sqrt{2\pi b^2}}\right)^n e^{-\frac{1}{2b^2} (\sum x_i^2 - 2\theta \sum x_i + n\theta^2)}$$

$$= \sum x_i^2 - 2\theta \sum x_i + n\theta^2$$

$$= \sum x_i^2 - 2\theta n\bar{x} + n\theta^2$$

$$= \left(\frac{1}{\sqrt{2\pi b^2}}\right)^n e^{-\frac{\sum x_i^2}{2b^2}} e^{\frac{\theta n\bar{x}}{b^2}} e^{-\frac{n\theta^2}{2b^2}}$$

$$P(\theta | x, b^2) \propto e^{-\frac{\frac{a}{n\bar{x}}}{b^2} \theta - \frac{\frac{b}{n}}{2b^2} \theta^2}$$

$$\propto N\left(\bar{x}, \frac{b^2}{n}\right)$$

$$\text{Var} = \frac{1}{2b} = \frac{1}{2\left(\frac{n}{2b^2}\right)} = \frac{1}{\frac{n}{b^2}} = \frac{b^2}{n}$$

$$\text{mean} = \frac{a}{2b} = \frac{\frac{n\bar{x}}{b^2}}{2\left(\frac{n}{2b^2}\right)} = \bar{x}$$



Laplace uninformative prior  $P(\theta) = U(\theta)$  if  $\theta = (0, 1)$

$$P(\theta) \propto 1$$

$$\Rightarrow P(\theta) = 1$$

if  $\theta = (0, 100)$

$$\Rightarrow P(\theta) = \frac{1}{100} \propto 1$$

Poisson-Gamma

$$P(\theta|x) = \frac{P(x|\theta)P(\theta)}{P(x)}$$

$$\propto P(x|\theta)P(\theta)$$

amazing trick.

$$\propto P(x|\theta)$$

$$= \frac{e^{-n\theta} \theta^{I\bar{x}}}{\pi \bar{x}!} \propto e^{-n\theta} \theta^{I\bar{x}+1-1}$$

$$\propto \text{Gamma}(I\bar{x}+1, n)$$

$$\Rightarrow \theta \sim \text{Gamma}(1, 0)$$

Generally  $P(\theta) = \text{Gamma}(\alpha, B)$

$$\Rightarrow P(\theta|x) = \text{Gamma}(I\bar{x}+\alpha, n+B)$$

$\mathcal{F}: X_1 \dots X_n; \theta, b^2 \stackrel{\text{ind}}{\sim} N(\theta, b^2), b^2 \text{ known}$

prior on  $\theta$  is Laplace  $\Rightarrow P(\theta) \propto 1$  improper!

$$P(\theta|x, b^2) \propto P(x|\theta, b^2) \propto N(\bar{x}, \frac{b^2}{n})$$