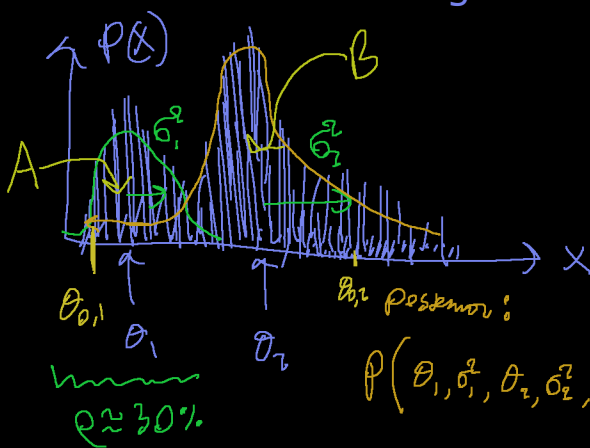


Consider the following data

theory



$$f(x) = p N(\theta_1, \sigma_1^2) + (1-p) N(\theta_2, \sigma_2^2)$$

How many parameters are in this model?

$$P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p | x) \propto P(x | \theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p) P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p)$$

$$P(x_1, \dots, x_n | \theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p) = \prod_{i=1}^n \left(e^{\frac{1}{\sqrt{2\pi}\sigma_1}} e^{-\frac{1}{2\sigma_1^2}(x_i - \theta_1)^2} + (1-p) e^{\frac{1}{\sqrt{2\pi}\sigma_2}} e^{-\frac{1}{2\sigma_2^2}(x_i - \theta_2)^2} \right)$$

integral B

we're in trouble because we can't simplify this... product of (a_i + b_i)

$$P(A|x) = \int P(A, B|x) dB$$

"Data Augmentation" or "parameter augmentation". We introduce parameters:

$$I_1 := \mathbb{1}_{x_1 \text{ comes from distribution A}} = \begin{cases} 1 & \text{if } x_1 \text{ from distr. A} \\ 0 & \text{if } x_1 \text{ from distr. B} \end{cases}$$

$$I_2 := \mathbb{1}_{x_2}$$

$$\vdots$$

$$I_n := \mathbb{1}_{x_n}$$

New Bayesian Setup:

$$P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p, I_1, \dots, I_n | x) \propto P(x | \theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p, I_1, \dots, I_n) P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p, I_1, \dots, I_n)$$

$$P(x) \rightarrow \prod_{i=1}^n \left(\frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{1}{2\sigma_1^2}(x_i - \theta_1)^2} \right)^{I_i} \left(\frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2\sigma_2^2}(x_i - \theta_2)^2} \right)^{1-I_i}$$

$$= \left(\frac{1}{\sqrt{2\pi}\sigma_1} \right)^{\sum I_i} e^{-\frac{1}{2\sigma_1^2} \sum I_i (x_i - \theta_1)^2} \left(\frac{1}{\sqrt{2\pi}\sigma_2} \right)^{n - \sum I_i} e^{-\frac{1}{2\sigma_2^2} \sum (1-I_i)(x_i - \theta_2)^2}$$

$$P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p | I_1, \dots, I_n) = P(I_1, \dots, I_n | p, \theta_1, \sigma_1^2, \theta_2, \sigma_2^2) P(p, \theta_1, \sigma_1^2, \theta_2, \sigma_2^2)$$

$$= P(I_1, \dots, I_n | p) P(p) P(\theta_1 | \sigma_1^2) P(\theta_2 | \sigma_2^2) P(\sigma_1^2) P(\sigma_2^2)$$

$$= \prod_{i=1}^n P(I_i | p) P(p) \propto 1 \propto \frac{1}{\sigma_1^2} \propto 1 \propto \frac{1}{\sigma_2^2}$$

Jeffreys priors

clearly not a kernel that we know

$$P(\theta_1, \sigma_1^2, \theta_2, \sigma_2^2, p, I_1, \dots, I_n | x) \propto e^{\sum I_i} (1-p)^{n - \sum I_i} \left(\frac{1}{\sigma_1^2} \right)^{-\frac{1}{2} \sum I_i - 1} \left(\frac{1}{\sigma_2^2} \right)^{-\frac{1}{2} (n - \sum I_i) - 1} e^{-\frac{1}{2\sigma_1^2} \sum I_i (x_i - \theta_1)^2} e^{-\frac{1}{2\sigma_2^2} \sum (1-I_i)(x_i - \theta_2)^2}$$

Create a Gibbs Sampler:

$$P(p | \dots) \propto e^{\sum I_i} (1-p)^{n - \sum I_i} \propto \text{Beta}(\sum I_i + 1, n - \sum I_i + 1)$$

$$P(\theta_1 | \dots) \propto e^{-\frac{1}{2\sigma_1^2} \sum I_i (x_i - \theta_1)^2} \propto e^{-\frac{\sum I_i x_i^2}{2\sigma_1^2} + \frac{\sum I_i x_i \theta_1}{\sigma_1^2} - \frac{\sum I_i \theta_1^2}{2\sigma_1^2}} \propto e^{\frac{\sum I_i x_i}{\sigma_1^2} \theta_1 - \frac{\sum I_i}{2\sigma_1^2} \theta_1^2}$$

$$\propto N\left(\frac{\sum I_i x_i}{\sum I_i}, \frac{\sigma_1^2}{\sum I_i}\right) = N(\bar{x}_A, \frac{\sigma_1^2}{n_A}) \quad \text{let } n_A = \sum I_i$$

$$P(\theta_2 | \dots) \propto N\left(\frac{\sum (1-I_i) x_i}{n - \sum I_i}, \frac{\sigma_2^2}{n - \sum I_i}\right) = N(\bar{x}_B, \frac{\sigma_2^2}{n_B}) \quad \text{let } n_B = n - \sum I_i$$

$$P(\sigma_1^2 | \dots) \propto \left(\frac{1}{\sigma_1^2} \right)^{\frac{1}{2} \sum I_i - 1} e^{-\frac{\sum I_i (x_i - \theta_1)^2}{2\sigma_1^2}} \propto \text{InvGamma}\left(\frac{\sum I_i}{2}, \frac{\sum I_i (x_i - \theta_1)^2}{2}\right)$$

$$= \text{InvGamma}\left(\frac{n_A}{2}, \frac{n_A \hat{\sigma}_1^2}{2}\right)$$

$$P(\sigma_2^2 | \dots) \propto \text{InvGamma}\left(\frac{n - \sum I_i}{2}, \frac{\sum (1-I_i) (x_i - \theta_2)^2}{2}\right) = \text{InvGamma}\left(\frac{n_B}{2}, \frac{n_B \hat{\sigma}_2^2}{2}\right)$$

$$P(I_i | \dots) \propto \left(e^{\frac{1}{\sqrt{2\pi}\sigma_1}} e^{-\frac{1}{2\sigma_1^2}(x_i - \theta_1)^2} \right)^{I_i} \left((1-p) \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2\sigma_2^2}(x_i - \theta_2)^2} \right)^{1-I_i}$$

$$\propto A^{I_i} B^{1-I_i} \propto \left(\frac{A}{A+B} \right)^{I_i} \left(\frac{B}{A+B} \right)^{1-I_i} \propto \text{Bern}(p) = \text{Bern}\left(\frac{A}{A+B}\right)$$

We can now build a Gibbs Sampler to have inference for all 5 + n parameters. We do need to be careful to specify a good starting location though.

Set $\theta_{0,1} = 25\%$ ile of data $\sigma_{0,1}^2, \sigma_{0,2}^2 = 1$ (both meters)
 $\theta_{0,2} = 75\%$ ile of data

$I_1, \dots, I_{n/2} = 1$ and $I_{n/2+1}, \dots, I_n = 0$.

This will force the lower group to be centered at theta1 and the upper group to be centered at theta2.

Bayes Factors (B) AKA "ratio of evidences". We want to compare two models:

$M_1 = \langle \mathcal{F}_1, P_1(\theta) \rangle$ model 1 is likelihood 1 and prior 1

$M_2 = \langle \mathcal{F}_2, P_2(\theta) \rangle$ model 2 is likelihood 2 and prior 2

$$B = \frac{P_{M_1}(x)}{P_{M_2}(x)} = \frac{\int P_1(x|\theta) P_1(\theta) d\theta}{\int P_2(x|\theta) P_2(\theta) d\theta}$$

if $> 1 \Rightarrow M_1$ better
 if $< 1 \Rightarrow M_2$ better

marginal probs. avg'd over all θ for $P_j(x|\theta)$ and $P_j(\theta)$ for both models.

Often Bayes Factors compare H_0 vs H_A . H_0 and H_A differ on θ .

Imagine $n=100$ coin flips and $x=61$ heads. Want to test if coin is unfair:

$H_0: \theta = 0.5$

$H_A: \theta \neq 0.5$

$\alpha = 5\%$

$\hat{\theta} = 61/100 = 0.61$

Frequency Test

$$\text{Retainment Region} = [\theta_0 \pm z_{\frac{\alpha}{2}} \sqrt{\frac{\theta_0(1-\theta_0)}{n}}] = [0.5 \pm 2 \sqrt{\frac{0.5(1-0.5)}{100}}] = [0.49, 0.51]$$

$\hat{\theta} \notin \text{Ret. Region} \Rightarrow \text{Reject } H_0$.

Bayesian Test using CR method let $\theta \sim U(0,1)$

$$CR_{0.95,1} = [p_{0.025}(0.025, 61+1, 39+1), p_{0.975}(0.975, 61+1, 39+1)] = [0.511, 0.700]$$

$\theta_0 = 0.5 \notin CR_{0.95,1} \Rightarrow \text{Reject } H_0$.