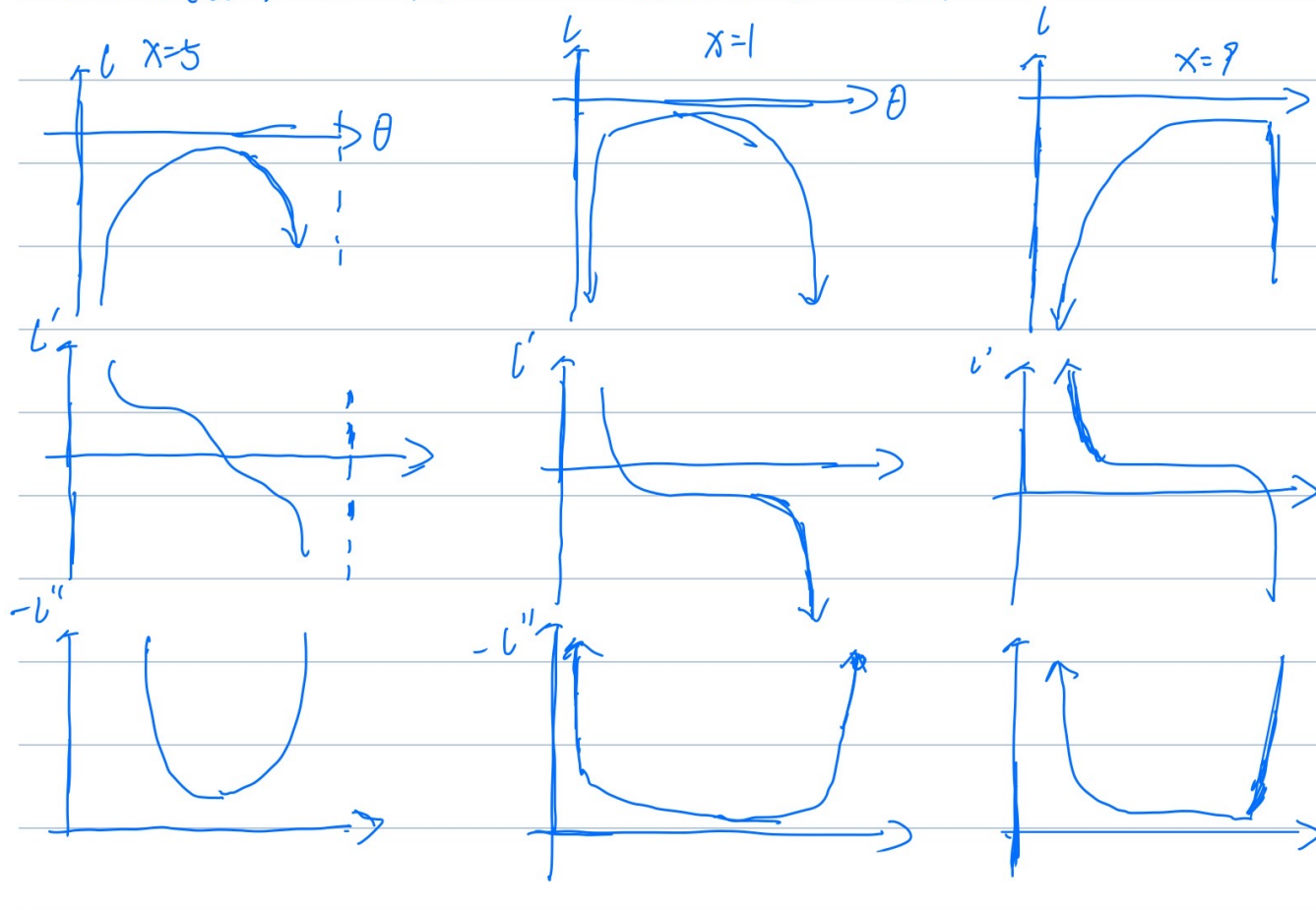


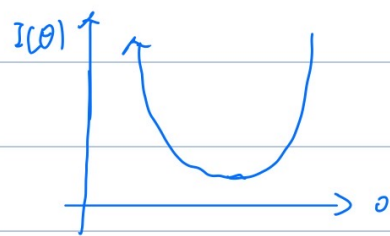
$$n=10 \quad X \sim \text{Bin}(n; \theta) := \binom{n}{x} \theta^x (1-\theta)^{n-x} \quad \ell(\theta; x) = \ln \binom{n}{x} + x \ln \theta + (n-x) \ln(1-\theta)$$

$$s(\theta; x) = \ell'(\theta; x) = \frac{x}{\theta} - \frac{n-x}{1-\theta}$$

$$-\ell''(\theta; x) = \frac{x}{\theta^2} + \frac{n-x}{(1-\theta)^2}$$



How much "informative" is in X ? $I_1(\theta = \frac{1}{2}) = 4 \quad I_1(\theta = \frac{1}{10}) \approx 1.01$



let $P(\theta) \propto \sqrt{I(\theta)}$ If $\mathcal{F}: \text{Bin}$ fixed n .

$$P(\theta) \propto \sqrt{n \cdot \frac{1}{\theta(1-\theta)}} \propto \sqrt{\frac{1}{\theta(1-\theta)}} = \theta^{\frac{1}{2}-1} (1-\theta)^{\frac{1}{2}-1} \propto \text{Beta}(\frac{1}{2}, \frac{1}{2}) \quad \text{the Jeffreys prior}$$

$$\phi = \tau(\theta) = \frac{\theta}{1-\theta}$$

$$\phi = \frac{\theta}{1-\theta} \rightarrow \theta = \frac{\phi}{1+\phi} \quad \mathcal{L}(\phi; x) = \binom{n}{x} \left(\frac{\phi}{1+\phi}\right)^x \left(\frac{1}{1+\phi}\right)^{n-x} = \binom{n}{x} \frac{\phi^x}{(1+\phi)^n}$$

$$\ell(\phi; x) = \ln \binom{n}{x} + x \ln(\phi) - n \ln(1+\phi) \quad \ell'(\phi; x) = \frac{x}{\phi} - \frac{n}{1+\phi} \quad -\ell''(\phi; x) = \frac{x}{\phi^2} + \frac{n}{(1+\phi)^2}$$

$$I(\phi) = E_X[-\ell''(\phi; x)] = \frac{n}{\phi(1+\phi)} - \frac{n}{(1+\phi)^2} = \frac{n}{\phi(1+\phi)^2}$$

$$P_T(\theta) \propto \sqrt{I(\theta)} = \sqrt{\frac{1}{\phi(\phi+1)}} \propto \phi^{-\frac{1}{2}}(\phi+1)^{-\frac{1}{2}} \stackrel{\text{via}}{\propto} \frac{1}{\pi} \phi^{-\frac{1}{2}}(\phi+1)^{-1} = \text{Beta prime}(\frac{1}{2}, \frac{1}{2}) = J_{1,1}$$

$$\begin{array}{ccc} P(X|\theta) \xrightarrow{\text{Jefferson's prior}} P_J(\theta) & \text{change of variable} & \theta = k(\phi) = \frac{\phi}{1+\phi} \\ \uparrow \phi = k(\theta) & & \uparrow \phi = k(\theta) \\ P(X|\phi) \xrightarrow{\text{Jefferson's prior}} P_J(\phi) & & P(\phi) = P_\theta(k^{-1}(\phi)) \left| \frac{d}{d\phi} [k^{-1}(\phi)] \right| \\ & & = \frac{1}{B(\frac{1}{2}, \frac{1}{2})} \left(\frac{\phi}{1+\phi} \right)^{-\frac{1}{2}} \left(\frac{1}{1+\phi} \right)^{-\frac{1}{2}} (\phi+1)^{-2} = \frac{1}{\pi} \phi^{-\frac{1}{2}}(\phi+1)^{-1} = \text{Beta prime}(\frac{1}{2}, \frac{1}{2}) \end{array}$$

you can't use prior Beta(1,1) ; is unique

Assume $P(\theta) \propto \sqrt{I(\theta)}$ prove:

$$\begin{aligned} P(\phi) &\propto \sqrt{I(\phi)} \quad \text{using } P(\phi) = P_\theta(k^{-1}(\phi)) \left| \frac{d}{d\phi} (k^{-1}(\phi)) \right| & P(\phi) = P_\theta(\theta) \left| \frac{d\theta}{d\phi} \right| &\propto \sqrt{I(\theta)} \left| \frac{d\theta}{d\phi} \right| \\ \sqrt{E_X[l'(\phi; x)^2]} &= \sqrt{I(\phi)} \propto P_J(\phi) & & = \sqrt{I(\theta) \frac{d\theta}{d\phi} \frac{d\theta}{d\phi}} \\ & & & = \sqrt{E_X[l'(\theta; x)^2] \frac{d\theta \cdot d\theta}{d\phi \cdot d\phi}} \\ & & & = \sqrt{E_X\left[\frac{d\ell}{d\theta} \cdot \frac{d\ell}{d\theta}\right] \frac{d\theta \cdot d\theta}{d\phi}} \end{aligned}$$

$P(\theta) = \text{Beta}(1,1)$ Laplace' } principle of uninformative priors

$P(\theta) = \text{Beta}(0,0)$

$P_J(\theta) = \text{Beta}(\frac{1}{2}, \frac{1}{2})$ Jefferson

$P(\theta) = \text{Beta}(d, \ell) \longrightarrow$ empirical Bayes informative prior, d, ℓ based on provided data.

new distribution: poisson

$$X \sim \text{Bin}(n, \theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x} \quad n \rightarrow \infty \quad \theta \rightarrow 0 \quad \text{but } \lambda = n\theta \quad \theta \leq \frac{\lambda}{n}$$

$$= \lim_{n \rightarrow \infty} \binom{n}{x} \left(\frac{\lambda}{n}\right)^x \left(1 - \frac{\lambda}{n}\right)^{n-x} = \lim_{n \rightarrow \infty} \frac{n!}{x!(n-x)!} \frac{\lambda^x}{n^x} \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x}$$

$$= \frac{\lambda^x}{x!} \lim_{n \rightarrow \infty} \frac{n!}{(n-x)!} \frac{1}{n^x} \left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x} = \frac{\lambda^x}{x!} \lim_{n \rightarrow \infty} \frac{n}{n} \cdot \frac{(n-1)}{n} \cdots \frac{(n-x)}{(n-x)!} \frac{1}{n^x} \cdot \underbrace{\left(1 - \frac{\lambda}{n}\right)^n \left(1 - \frac{\lambda}{n}\right)^{-x}}_{= e^{-\lambda}}$$

$$= \frac{e^{-\lambda} \lambda^x}{x!} \quad \text{poisson}(\lambda)$$

$$\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n = e^{\lim_{n \rightarrow \infty} n \ln\left(1 - \frac{\lambda}{n}\right)} = e^{\lim_{n \rightarrow \infty} \ln\left(1 - \frac{\lambda}{n}\right) + \frac{n^2}{n-\lambda} \left(\frac{\lambda}{n^2}\right)} = e^0 ? \quad \text{X}$$

\mathcal{Y} : poisson $n \geq 1$ we want to find the conjugate prior

$$p(\theta|x) = \frac{p(x|\theta)p(\theta)}{p(x)} \quad k(\theta|x) \propto p(x|\theta)p(\theta) \propto k(x|\theta)k(\theta)$$

$$k(\theta|x) \propto \left(\frac{\theta^x e^{-\theta}}{x!} \right) p(\theta) \propto \theta^x e^{-\theta} k(\theta) = \theta^x e^{-\theta} \underbrace{(\theta^a e^{-b\theta})}_{k(\theta)} = \theta^{x+a} e^{-b(\theta+1)}$$

let $\tau = b\theta$

posterior match to conjecture to what conjugate prior looks like

$$C = \int_0^\infty \theta^a e^{-b\theta} d\theta = \int_0^\infty \left(\frac{e}{b}\right)^a e^{-t} \frac{1}{b} dt = \frac{1}{b^{a+1}} \int_0^\infty t^{a+1-1} e^{-t} dt = \frac{1}{b^{a+1}} \Gamma(a+1) \Rightarrow P(\theta) = \frac{b^{a+1}}{\Gamma(a+1)} \theta^a e^{-b\theta}$$

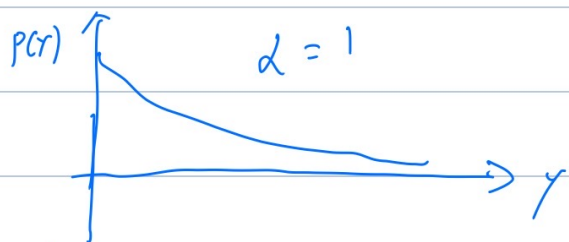
Gamma($a+1, b$)

conjugate prior in poisson model

$$p(\theta) = \text{Gamma}(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \theta^{\alpha-1} e^{-\beta\theta} \quad E[Y] = \frac{\alpha}{\beta} \quad \text{Supp}[Y] = (0, \infty)$$

$$Y \sim \text{Gamma}(\alpha, \beta) \quad \text{Var}(Y) = \frac{\alpha}{\beta^2}$$

$$\text{Mode}[Y] = \frac{\alpha-1}{\beta} ; \alpha \geq 1 \quad \text{Med}[Y] = ? \text{Gamma}(\alpha, \beta) \quad \text{no closed form.}$$



shape

