

$$H_0: \theta = \theta_0$$

$$H_a: \theta \neq \theta_0 = 0.5 \quad \text{two-sided test}$$

$$\text{if } P_{\text{val}} := P(H_0 | x) < \alpha \Rightarrow \text{Reject } H_0 / \text{Accept } H_a$$

$$= P(\theta = \theta_0 | x) \stackrel{P(\theta) = U(0,1)}{\downarrow} 0$$

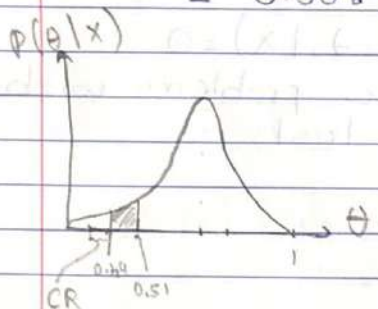
### Two Ideas

I Declare  $\delta$  e.g.  $\delta = 0.01$ ; a "margin of equivalence"  
then you modify the hypotheses

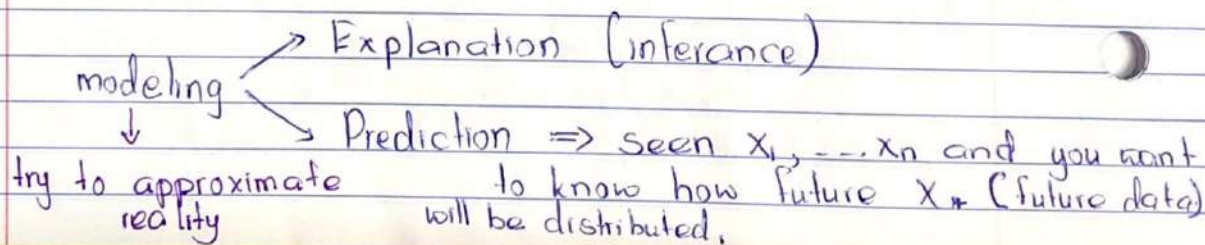
$$H_0: \theta \in [\theta_0 \pm \delta] \quad \text{e.g. } [0.49, 0.51]$$

$$H_a: \theta \notin [\theta_0 \pm \delta]$$

$$\begin{aligned} P_{\text{val}} &= P(H_0 | x) = P(\theta \in [\theta_0 \pm \delta] | x) \\ &= \text{qbeta}(0.51, b2, h0) - \text{qbeta}(0.49, b2, h0) \\ &= (\text{small}) \quad 0.609 - 0.607 \\ &= 0.002 < \alpha = 5\% \Rightarrow \text{Reject } H_0 \end{aligned}$$



II If  $\theta_0 \in CR_{0,1-\alpha} \Rightarrow$  Retain  $H_0$  else Reject  
Downside: no pvalue!

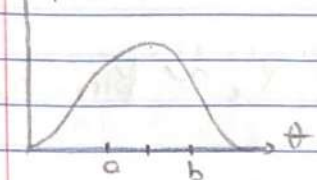


$P(X^* | x)$ ; the posterior predictive distribution.

If  $\theta$  was known, what is the posterior predictive distribution?

$P(X^* | \theta)$  the best you can do, but it is not possible since  $\theta$  is unknown.

$P(\theta | x)$



$$P(X^* | \theta) = P(X^* | \theta = a) P(\theta = a) + P(X^* | \theta = b) P(\theta = b)$$

weighted average

continuous posterior

$$P(X^* | \theta) = \int_{\mathcal{H}} P(X^*, \theta | x) d\theta$$

$$= \int_{\mathcal{H}} p(x^* | \theta, x) p(\theta | x) d\theta$$

If  $\theta$  is unknown;  $x$

doesn't any more information.

$$= \int_{\mathcal{H}} \overset{\text{likelihood}}{p(x^* | \theta)} \overset{\text{posterior}}{p(\theta | x)} d\theta$$

discrete posterior  $\downarrow$

$$= \sum_{\theta \in \mathcal{H}} p(x^* | \theta) p(\theta | x)$$

$$\mathcal{H}_0 = [0.5, 0.75]; x = \langle 0, 1, 1 \rangle$$

$X^* \sim ?$  One future observation

$$P(X^* | x) = \text{Bern}(\cdot)$$

$$\text{Supp}[X^* | x] = \{0, 1\}$$



Consider Rich's theory.  
 $X^* | x \sim \text{Bern}(\hat{\theta}_{MLE} = 2/3)$

Problems

- ①  $\hat{\theta}_{MLE}$  may not be  $\in \mathcal{H}_0$ .
- ②  $\hat{\theta}_{MLE}$  could be 0 or 1.
- ③ If multiple future observation  $X_* \sim \text{Bin}(n_*, \theta_{MLE})$ ,  
BAD IDEA.

Assume prior of indifference posterior.

$$P(\theta = 0.75 | x) = 0.53, \quad P(\theta = 0.5 | x) = 0.47$$

$$P(X^* | x) = \frac{P(X^* | \theta = 0.75) P(\theta = 0.75 | x) + P(X^* | \theta = 0.5) P(\theta = 0.5 | x)}{P(\theta = 0.75 | x) + P(\theta = 0.5 | x)}$$

$$= P(X^* | \theta = 0.75) \cdot (0.53) + P(X^* | \theta = 0.5) \cdot (0.47)$$

$$= (0.75)^{x^*} (0.25)^{1-x^*} (0.53) + (0.5)^{x^*} (0.5)^{1-x^*} (0.47)$$

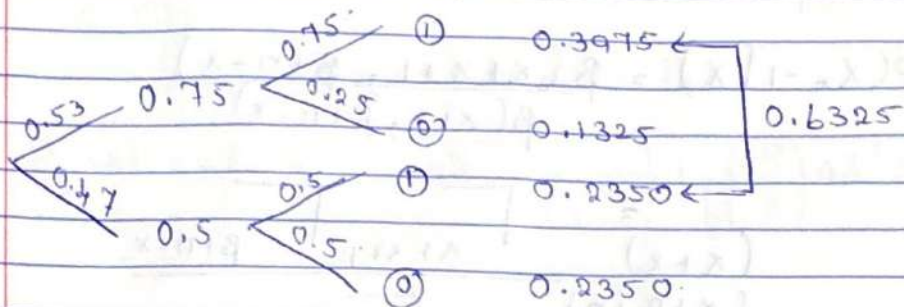
$$P(X^* | x) = \text{Bern}(\cdot) = \text{Bern}(0.6325)$$

trick: compute  $P(X^* = 1 | x) = (0.75)^1 (0.25)^0 (0.53) + (0.5)^1 (0.5)^0 (0.47)$   
 $= 0.6325$

Posterior after see the data.

$P(\theta | x)$

$P(x^* | \theta)$



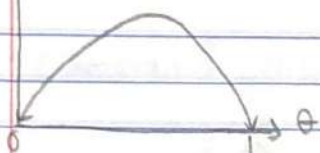
$\bar{F}$ : Binomial Fixed  $n$

$P(\theta) = \text{Beta}(\alpha, \beta)$

$\Rightarrow P(\theta | x) = \text{Beta}(\alpha + x, \beta + n - x)$

What is the posterior predictive distribution for  $n_* = n$ ?

$P(\theta | x)$



$$P(x_* | x) = \int_{\theta} P(x_* | \theta) P(\theta | x) d\theta.$$

$$= \int_0^1 (\theta^{x_*} (1-\theta)^{1-x_*}) \left( \frac{1}{\beta(\alpha+x, \beta+n-x)} \theta^{\alpha+x-1} (1-\theta)^{\beta+n-x-1} \right) d\theta$$

$$= \frac{1}{\beta(\alpha+x, \beta+n-x)} \int_0^1 \theta^{\alpha+x+x_*-1} (1-\theta)^{\beta+n-x-x_*+1-1} d\theta$$

$$= \frac{\beta(\alpha+x+x_*, \beta+n-x-x_*+1)}{\beta(\alpha+x, \beta+n-x)}$$

$$= \text{Bern} \left( \frac{\alpha+x}{\alpha+\beta+n} \right) \Rightarrow \hat{\theta}_{\text{MMSE}} = E[\theta | x].$$



Trick compute.

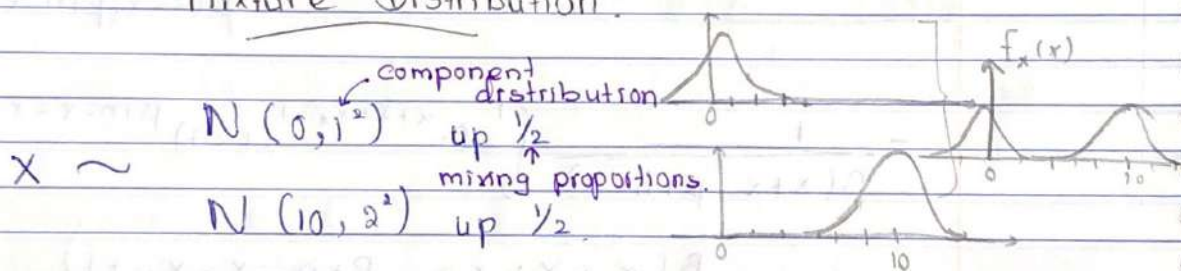
$$P(X_{n+1}=1 | x) = \frac{\beta(x+x+1, \beta+n-x)}{\beta(x+x, \beta+n-x)}$$

$$= \frac{\frac{\Gamma(x+x) \Gamma(\alpha+x) \Gamma(\beta+n-x)}{\Gamma(x+\beta+n)}}{\frac{\Gamma(x+x) \Gamma(\alpha+x) \Gamma(\beta+n-x)}{\Gamma(\alpha+\beta+n) \Gamma(\alpha+\beta+n-x) \Gamma(\alpha+\beta+n)}}$$

$$= \frac{(\alpha+x)}{(\alpha+\beta+n)} \frac{\beta(x+x, \beta+n-x)}{\beta(\alpha+n, \beta+n-x)}$$

$$= \frac{(\alpha+x)}{(\alpha+\beta+n)}$$

Mixture Distribution.



$$f_X(x) = ? \quad p(x) = \sum_{\theta \in \Theta} P(x, \bar{\theta})$$

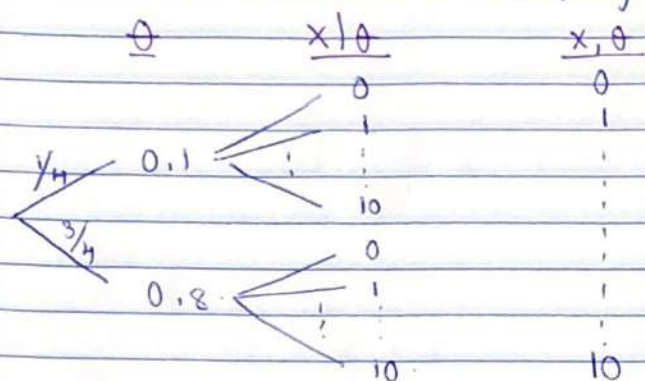
$$= \sum_{\theta \in \Theta} P(x | \bar{\theta}) P(\bar{\theta})$$

$$= \left( \frac{1}{\sqrt{2\pi(1^2)}} e^{-\frac{1}{2 \cdot 1^2} (x-0)^2} \right) \left( \frac{1}{2} \right) + \left( \frac{1}{\sqrt{2\pi(2^2)}} e^{-\frac{1}{2 \cdot 2^2} (x-10)^2} \right) \left( \frac{1}{2} \right)$$

$$x \sim \begin{cases} \text{Bin}(10, 0.1) & \text{up } 1/4 \\ \text{Bin}(10, 0.8) & \text{up } 3/4 \end{cases}$$



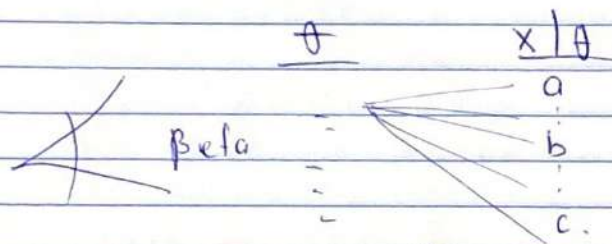
$$P_x(x) = \left[ \binom{10}{x} 0.1^x 0.9^{10-x} \right] \left( \frac{1}{4} \right) + \left[ \binom{10}{x} 0.8^x 0.2^{10-x} \right] \left( \frac{3}{4} \right)$$



≠ Binomial  
can't have two  
bump

Mixture distribution's has a discrete # of components.  
Compound distribution's do not.

$$x \sim \begin{cases} \text{Bin}(n, \theta) \\ \text{Bin}(n, \theta) \end{cases} \quad \left\{ \begin{array}{l} \theta \text{'s come from beta.} \end{array} \right.$$



$$P(\theta | x) = \frac{P(x | \theta) P(\theta)}{P(x)}$$

$$= \frac{P(x | \theta) P(\theta)}{\int_{\mathcal{H}} P(x | \theta) P(\theta) d\theta}$$

$$\int_{\mathcal{H}} P(x | \theta) P(\theta) d\theta$$

compound distribution.