

Informative Priors.

Let  $\theta$  be the career prob. of getting a hit for a batter in baseball.

$$\hat{\theta}_{MLE} = \frac{X}{n} \quad \begin{array}{l} \leftarrow \# \text{ of hits} \\ \leftarrow \# \text{ of bats} \end{array} \quad \text{bad estimator (b/c small } n)$$

Bern  $\theta$  in history is 0.366

Average is 0.260

$$n=3, X=2$$

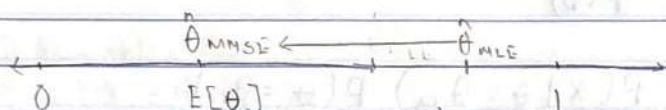
$$\hat{\theta}_{MLE} = \hat{\theta}_{MMSE} = 0.667 \text{ if } \theta \sim \text{Beta}(0,0)$$

$$\text{If } \theta \sim \text{Beta}(1,1)$$

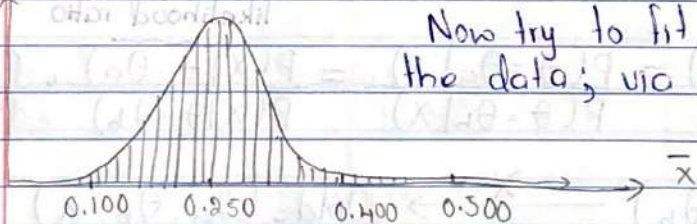
$$\hat{\theta}_{MMSE} = \frac{1+2}{7+1+1} = \frac{3}{9} = \frac{1}{3} = 0.333$$

Design a prior s.t. pick  $\alpha, \beta$

$$E[\theta] = 0.260$$



Look at previous data, e.g. all players  $\geq 500$  at bats and we examine  $\bar{x}$ 's



Now try to fit a beta distribution to the data; via maximum likelihood,

$$\hat{\alpha}_{MLE} = 78.7 \text{ \& } \hat{\beta}_{MLE} = 224.8$$

$$\Rightarrow E[\theta] = 0.260$$

$$\Rightarrow n_0 = 303.5$$

This prior is called "empirical Bayes"

$$\Rightarrow p = \frac{303.5}{303.5+3} = 99\%$$

$$\hat{\theta}_{NMSE} = \frac{(1\%)}{(1-p)} (0.667) + \frac{(99\%)}{p} (0.260) = 0.263$$

$\theta \sim \text{Beta}(\frac{1}{2}, \frac{1}{2})$  is called the Jeffery's prior.  
(uninformative).

$$\text{Odds}(A) := \frac{P(A)}{P(A^c)} = \frac{P(A)}{1-P(A)} \in [0, \infty)$$

$$\text{Odds Against}(A) = \text{Odds}(A)^{-1} = \frac{1-P(A)}{P(A)} \in [0, \infty)$$

Example

$$x = \begin{cases} 5 \text{ up } \frac{1}{6} \\ -1 \text{ up } \frac{5}{6} \end{cases}$$

$$\begin{aligned} E[\theta] &= 5(\frac{1}{6}) + (-1)(\frac{5}{6}) \\ &= \frac{5}{6} - \frac{5}{6} = 0. \end{aligned}$$

$$\text{Odds}(A, B) = \frac{P(A)}{P(B)}$$

$$P(\theta = \theta_a | x) = \frac{P(x | \theta = \theta_a) P(\theta = \theta_a)}{P(x)}$$

$$P(\theta = \theta_b | x) = \frac{P(x | \theta = \theta_b) P(\theta = \theta_b)}{P(x)}$$

$$\text{Odds}(\theta_a, \theta_b | x) = \frac{P(\theta = \theta_a | x)}{P(\theta = \theta_b | x)} = \frac{P(x | \theta = \theta_a)}{P(x | \theta = \theta_b)} \cdot \frac{P(\theta = \theta_a)}{P(\theta = \theta_b)}$$

likelihood ratio      Odds( $\theta_a, \theta_b$ )  
Prior odds

$$\Rightarrow \text{Odds}(\theta_a, \theta_b) \xrightarrow{x} \text{Odds}(\theta_a, \theta_b, x)$$

posterior odds

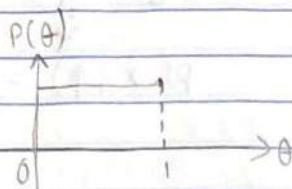
$$\text{e.g. } 1:1 \xrightarrow{x} 5:1$$



$T$ : Binomial; fixed  $n$   
 Let  $\phi(\theta)$  be odds  $\theta$ .

$$\phi(\theta) = \frac{\theta}{1-\theta}$$

$$P(\theta) = U(0,1)$$



Fisher's  
idea

→ What is prior of indifferent on  $\phi$ ?  
 $P(\phi) \equiv U(0, \infty) = 0 \neq \text{not a valid PDF}$   $\int_0^\infty 0 d\phi \neq 1$

If  $P(\theta) \sim U(0,1)$ ; What is  $P(\phi)$ ?

For a continuous random variable  $x$ ,  
 If  $y = t(x)$  where  $t$  is invertible and  $f_x(x)$  known  
 $\Rightarrow f_y(y) = f_x(t^{-1}(y)) \left| \frac{d}{dy} [t^{-1}(y)] \right|$  (derived in Math 368)

$$\phi = \phi(\theta) = \frac{\theta}{1-\theta} = t(\theta) \quad f_\phi(\phi) = f_\theta\left(\frac{\phi}{1+\phi}\right) \left| \frac{1}{(1+\phi)^2} \right|$$

$$\phi(1-\theta) = \theta$$

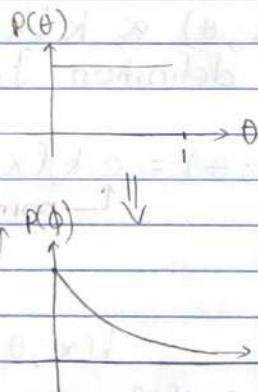
$$\phi - \theta\phi = \theta$$

$$\phi = \theta + \theta\phi$$

$$\phi = \theta(1+\phi)$$

$$\theta = \frac{\phi}{(1+\phi)} = t^{-1}(\phi)$$

$$\begin{aligned} \frac{d}{d\phi} [t^{-1}(\phi)] &= \frac{(1+\phi)(1) - \phi(1)}{(1+\phi)^2} \\ &= \frac{1}{(1+\phi)^2} \end{aligned}$$



$$\tilde{F}: p(x|\theta)$$

$$\downarrow t$$

$$p(x|\phi)$$

$$p(x|\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$$

$$\uparrow \uparrow$$

$$p(x|\phi) = \binom{n}{x} \left(\frac{\phi}{1+\phi}\right)^x \left(1 - \frac{\phi}{1+\phi}\right)^{n-x}$$

$$= \binom{n}{x} \frac{\phi^x}{(1+\phi)^n}$$

Binomial parametrization with odds

$$\phi = \phi(\theta) = \frac{\theta}{1-\theta} = t(\theta)$$

$$\Rightarrow \theta = \frac{\phi}{1+\phi} = t^{-1}(\phi)$$

$$\tilde{F}: p(x|\theta) \xrightarrow{\text{Jeffrey's protocol}} p(\theta)$$

$$\downarrow t$$

$$p(x|\phi) \xrightarrow{\text{Jeffrey's protocol}} p(\phi)$$

transformation of variables  
formula works

Let  $X$  be continuous with density  $f(x;\theta)$

$f(x;\theta) \propto k(x;\theta)$  is unique  
by definition  $\exists c > 0$  not a function of  $x$

$$f(x;\theta) = c \cdot k(x;\theta)$$

↑ normalization constant

def:

$$\left(1 = \int_{\text{supp}[x]} f(x;\theta) dx\right) = \int_{\text{supp}[x]} c k(x;\theta) dx$$

$$\Rightarrow \left( \int_{\text{supp}[x]} k(x;\theta) dx \right)^{-1} = c$$



$$P(x; \theta) \propto k(x; \theta)$$

which means  $c > 0$ .

$$P(x; \theta) = ck(x; \theta)$$

$$c = \left( \sum_{x \in \text{supp}[x]} k(x; \theta) \right)^{-1} \propto k(y; \alpha, \beta)$$

$$y \sim \text{Beta}(\alpha, \beta) = \frac{1}{B(\alpha, \beta)} y^{\alpha-1} (1-y)^{\beta-1}$$

$f_y(y)$

$$\propto y^{\alpha-1} (1-y)^{\beta-1}$$

$T$ : Binomial,  $P(\theta) = \text{Beta}(\alpha, \beta)$

$$\Rightarrow P(\theta|x) = \text{Beta}(x+\alpha, n-x+\beta)$$

$$P(\theta|x) = \frac{P(x|\theta) P(\theta)}{P(x)}$$

$$\propto P(x|\theta) P(\theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x} \left( \frac{1}{B(\alpha, \beta)} \theta^{\alpha-1} (1-\theta)^{\beta-1} \right)$$

$$\propto \theta^{x+\alpha-1} (1-\theta)^{n-x+\beta-1}$$

$$\propto \text{Beta}(x+\alpha, n-x+\beta)$$

$$y \sim N(\theta, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(y-\theta)^2}$$

$$\propto e^{-\frac{1}{2\sigma^2}(y-\theta)^2}$$

$$= e^{-\frac{1}{2\sigma^2}(y^2 - 2\theta y + \theta^2)}$$

$$= e^{-\frac{y^2}{2\sigma^2} + \frac{\theta y}{\sigma^2} - \frac{\theta^2}{2\sigma^2}}$$

$$= e$$

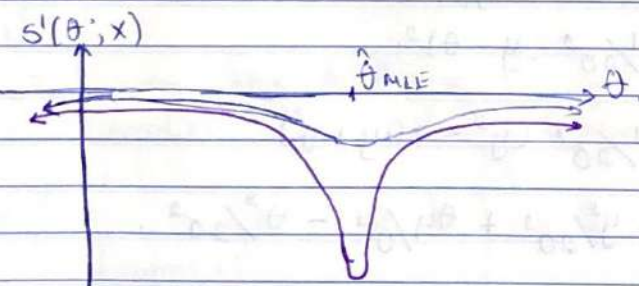
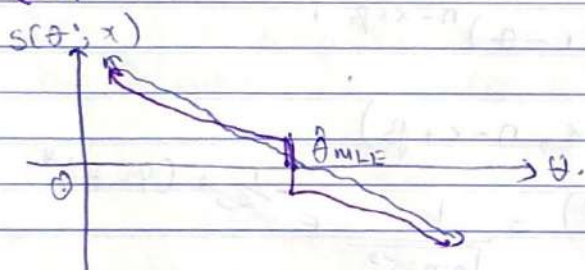
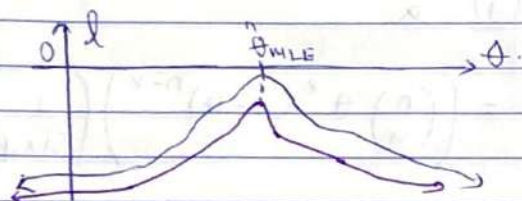
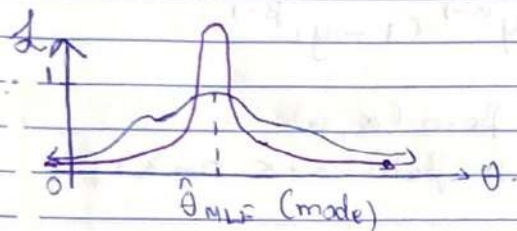
$$= e^{-\frac{y^2}{2\sigma^2} + \theta y/\sigma^2} e^{-\frac{\theta^2}{2\sigma^2}}$$

$$\propto e^{\theta y/\sigma^2 - y^2/2\sigma^2} = k(y; \theta, \sigma^2)$$

$$c = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}}$$

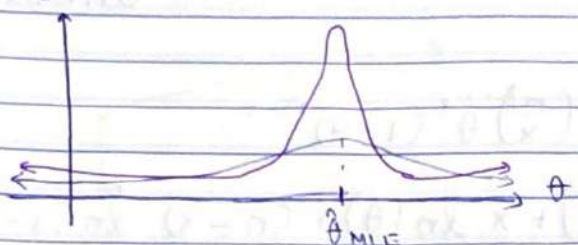
Remark:  $l(\theta; x) = \ln(L)$   $\log$   
 Score function  $s(\theta, x) = l'(\theta; x)$

fisher informative  $I(\theta) = \text{Var}_x [s(\theta; x)] = -E_x [l''(\theta; x)]$   
 HA 369  $s'(\theta; x)$





$-S'(\theta; x)$



$$x \sim \text{Bin}(n, \theta)$$

$$P(X=x; \theta) = \binom{n}{x} \theta^x (1-\theta)^{n-x}$$

$$l(\theta; x) = \ln \left( \binom{n}{x} \right) + x \ln(\theta) + (n-x) \ln(1-\theta)$$

$$S(\theta; x) = l'(\theta; x) = \frac{x}{\theta} - \frac{n-x}{1-\theta}$$

$$S'(\theta; x) = l''(\theta; x) = -\frac{x}{\theta^2} - \frac{n-x}{(1-\theta)^2}$$

$$-l''(\theta; x) = \frac{x}{\theta^2} + \frac{n-x}{(1-\theta)^2}$$

$$\Rightarrow \mathbb{E}(\theta) = \mathbb{E}_x \left[ \frac{x/\theta^2 + n-x/(1-\theta)^2}{n(1-\theta)} \right]$$

$$= \frac{n\theta}{\theta^2} + \frac{n-n\theta}{(1-\theta)^2}$$

$$= \frac{n}{\theta} + \frac{n}{1-\theta} = n \left( \frac{1}{\theta(1-\theta)} \right) = \mathbb{E}(\theta)$$