

Lecture 15.

04/02 号 [MATH341].

$\mathcal{F}$ : iid  $N(\theta, b^2)$  with  $b^2$  known.

$$P(\theta | b^2) = N(\mu_0, \tau^2) = N(\mu_0, \frac{\sigma^2}{n_0})$$

Imagine pseudodata  $y_1, \dots, y_n$  iid  $N(\mu_0, b^2)$

$$\Rightarrow \bar{Y} \sim N(\mu_0, \frac{\sigma^2}{n_0})$$

$$\Rightarrow P(\theta | X, b^2) = N(\hat{\theta}_p, b_p^2) \text{ where } \hat{\theta}_p = \frac{\frac{n\bar{X}}{b^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{b^2} + \frac{1}{\tau^2}} = \frac{n\bar{X} + n_0\mu_0}{n + n_0}$$

conjugate

$$b_p^2 = \frac{1}{\frac{n}{b^2} + \frac{1}{\tau^2}} = \frac{b^2}{n + n_0}$$

$n_* = 1$

$$P(X_* | X, b^2) = \int_{\mathbb{R}} P(X_* | \theta, b^2) P(\theta | X, b^2) d\theta$$

$$= \int_{\mathbb{R}} N(\theta, b^2) \cdot N(\hat{\theta}_p, b_p^2) d\theta$$

$$\stackrel{\text{constant}}{=} \int_{\mathbb{R}} \left( \frac{1}{\sqrt{2\pi}b^2} e^{-\frac{1}{2b^2}(X_* - \theta)^2} \right) \left( \frac{1}{\sqrt{2\pi}b_p^2} e^{-\frac{1}{2b_p^2}(\theta - \hat{\theta}_p)^2} \right) d\theta$$

$$\propto \int_{\mathbb{R}} e^{-\frac{1}{2b^2}(X_* - \theta)^2} e^{-\frac{1}{2b_p^2}(\theta - \hat{\theta}_p)^2} d\theta$$

$$= \int_{\mathbb{R}} e^{-\frac{X_*^2}{2b^2}} e^{\frac{X_*\theta}{b^2}} e^{-\frac{\theta^2}{2b^2}} e^{-\frac{\theta^2}{2b_p^2}} e^{\frac{\theta\hat{\theta}_p}{b_p^2}} e^{-\frac{\hat{\theta}_p^2}{2b_p^2}} d\theta$$

$X_*$  function  $X$

$$\propto e^{-\frac{X_*^2}{2b^2}} \int_{\mathbb{R}} e^{\frac{X_*\theta}{b^2}} e^{-\frac{\theta^2}{2b^2}} e^{-\frac{\theta^2}{2b_p^2}} e^{\frac{\theta\hat{\theta}_p}{b_p^2}} d\theta$$

$$= e^{-\frac{X_*^2}{2b^2}} \int_{\mathbb{R}} e^{\underbrace{\left(\frac{X_*}{b^2} + \frac{\hat{\theta}_p}{b_p^2}\right)\theta}_a - \underbrace{\left(\frac{1}{2b^2} + \frac{1}{2b_p^2}\right)\theta^2}_b} d\theta$$

$$\circ \theta \sim N\left(\frac{a}{b}, \frac{1}{b}\right) = \frac{1}{\sqrt{2\pi(\frac{1}{b})}} e^{-\frac{1}{2(\frac{1}{b})}\left(\theta - \frac{a}{b}\right)^2}$$

$$= \sqrt{\frac{b}{\pi}} e^{-b\left(\theta^2 - \frac{a\theta}{b} + \frac{a^2}{4b}\right)}$$

$$= \sqrt{\frac{b}{\pi}} e^{-b\theta^2 + a\theta - \frac{a^2}{4b}}$$

$$= \underbrace{\sqrt{\frac{b}{\pi}} e^{-\frac{a^2}{4b}}}_c e^{a\theta - b\theta^2}_k(\theta)$$

$$\propto e^{a\theta - b\theta^2}$$

$$= e^{-\frac{X_*^2}{2b^2}} \underbrace{\sqrt{\frac{b}{\pi}} e^{\frac{a^2}{4b}}}_c \int_{\mathbb{R}} \underbrace{\sqrt{\frac{b}{\pi}} e^{-\frac{a^2}{4b}} e^{a\theta - b\theta^2}}_k(\theta) d\theta = e^{-\frac{X_*^2}{2b^2}} \sqrt{\frac{\pi}{b}} e^{\frac{a^2}{4b}}$$

constant: not a function of  $x_t$

$$\propto e^{-\frac{x_t^2}{2b^2}} e^{\left(\frac{x_t}{b^2} + \frac{\hat{\theta}_p}{b^2}\right)^2 / 4b}$$

$$= e^{-\frac{x_t^2}{2b^2}} e^{\frac{x_t^2}{4b^2}} e^{\frac{x_t \hat{\theta}_p}{2b^2 b}} e^{\frac{\hat{\theta}_p^2}{4b^2 b}}$$

$$\propto e^{\underbrace{\frac{\hat{\theta}_p}{2b^2 b} x_t}_A - \underbrace{\left(\frac{1}{2b^2} - \frac{1}{4b^2 b}\right) x_t^2}_B}$$

$$= e^{Ax_t - Bx_t^2}$$

$$\propto N\left(\frac{A}{2B}, \frac{1}{2B}\right)$$

$$2bb^2 = 2\left(\frac{1}{2b^2} + \frac{1}{2b^2}\right) b^2 = 1 + \frac{b^2}{b^2}$$

$$\frac{1}{2B} = \frac{1}{2\left(\frac{1}{2b^2} - \frac{1}{4b^2 b}\right)} = \frac{1}{\frac{1}{b^2} - \frac{1}{2bb^2}}$$

$$= \frac{b^2}{b^2 - \frac{1}{2b}}$$

$$= \frac{2bb^2 b^2}{2bb^2 - 1}$$

$$= \frac{\left(\left(1 + \frac{b^2}{b^2}\right) b^2\right)}{\left(1 + \frac{b^2}{b^2}\right) - 1} = \frac{\left(1 + \frac{b^2}{b^2}\right) b^2}{\frac{b^2}{b^2}} = \delta_p^2 \left(1 + \frac{b^2}{b^2}\right)$$

$$= \delta_p^2 + b^2$$

$$\frac{A}{2B} = A(\delta_p^2 + b^2) = \frac{\hat{\theta}_p}{2bb^2 \delta_p^2} (\delta_p^2 + b^2) = \frac{\hat{\theta}_p}{\left(1 + \frac{b^2}{b^2}\right) b^2} (\delta_p^2 + b^2)$$

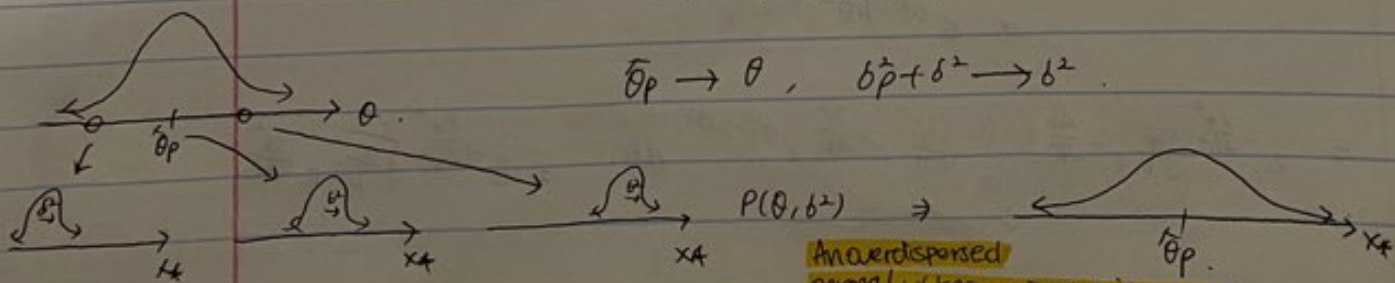
$$= \frac{\hat{\theta}_p}{(\delta_p^2 + b^2)} (\delta_p^2 + b^2)$$

$$= \hat{\theta}_p$$

$$\therefore P(x_t | x, b^2) = N(\hat{\theta}_p, \delta_p^2 + b^2)$$

$$= N(\hat{\theta}_p, \left(1 + \frac{1}{n/n_0}\right) b^2) \rightarrow N(\theta, b^2)$$

$$\hat{\theta}_p \rightarrow \theta, \quad \delta_p^2 + b^2 \rightarrow b^2$$



An overdispersed normal w/ known variance is normal



F: Find  $N(\theta, \sigma^2)$  with  $\theta$  known.

$$L(b^2; x, \theta) = (2\pi b^2)^{-n/2} e^{-\frac{1}{2b^2} \sum (\bar{x}_i - \theta)^2}$$

$$\begin{aligned} l(b^2; x, \theta) &= -\frac{n}{2} \ln(2\pi b^2) - \frac{1}{2b^2} \sum (\bar{x}_i - \theta)^2 \\ &= -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(b^2) - \frac{1}{2b^2} \sum (\bar{x}_i - \theta)^2 \end{aligned}$$

$$l'(b^2; x, \theta) = -\frac{n}{2} \frac{1}{b^2} + \frac{1}{2b^4} \sum (\bar{x}_i - \theta)^2 \stackrel{\text{set}}{=} 0$$

$$\Rightarrow -n + \frac{\sum (\bar{x}_i - \theta)^2}{b^2} = 0 \quad \rightarrow \quad \sum (\bar{x}_i - \theta)^2 = n \cdot \hat{\sigma}_{MLE}^2$$

$$\begin{aligned} \hat{\sigma}_{MLE}^2 &= \frac{\sum (\bar{x}_i - \theta)^2}{n} \quad \text{similar to } S^2 \text{ (sample variance)} \\ &= \frac{\sum (\bar{x}_i - \bar{x})^2}{n-1} \rightarrow b^2 \end{aligned}$$

Laplace Prior:  $P(b^2 | \theta) \propto 1$ , improper

Since  $b^2 \in (0, \infty)$

$$\begin{aligned} P(b^2 | x, \theta) &\propto P(x | \theta, b^2) = (2\pi b^2)^{-n/2} e^{-\frac{n \hat{\sigma}_{MLE}^2}{2b^2}} \\ &\propto (b^2)^{-n/2} e^{-\frac{n \hat{\sigma}_{MLE}^2 / 2}{b^2}} \end{aligned}$$

$$U \sim \text{Gamma}(\alpha, B) = \frac{B^\alpha}{\Gamma(\alpha)} u^{\alpha-1} e^{-Bu}$$

$$V = \frac{1}{U} \stackrel{=x(v)}{\sim} f_U(x^{-1}(v)) \left| \frac{d}{dv} [x^{-1}(v)] \right| = \frac{B^\alpha}{\Gamma(\alpha)} \left( \frac{1}{v} \right)^{\alpha-1} e^{-\frac{B}{v}} \left| -\frac{1}{v^2} \right|$$

$$t^{-1}(v) = \frac{1}{v} \quad = \frac{B^\alpha}{\Gamma(\alpha)} v^{-\alpha+1} v^{-2} e^{-\frac{B}{v}}$$

$$= \frac{B^\alpha}{\Gamma(\alpha)} v^{-\alpha-1} e^{-B/v}$$

$$= \text{InvGamma}(\alpha, B)$$

(inverse gamma r.v.)

$$\begin{aligned} \frac{B^\alpha}{\Gamma(\alpha)} v^{-\alpha-1} e^{-B/v} &\propto k(v) \\ &= v^{-\alpha-1} e^{-B/v} \end{aligned}$$

$$\begin{aligned} \Rightarrow (b^2)^{-n/2} e^{-\frac{n \hat{\sigma}_{MLE}^2 / 2}{b^2}} &= (b^2)^{-(\frac{n}{2}-1)-1} e^{-\frac{n \hat{\sigma}_{MLE}^2 / 2}{b^2}} \\ &= \text{InvGamma}\left(\frac{n}{2}-1, \frac{n \hat{\sigma}_{MLE}^2}{2}\right) \\ &= \frac{n-2}{2} \dots \end{aligned}$$

$$P(b^2 | x, \theta) \propto P(x | \theta, b^2) P(b^2 | \theta, x) \propto (b^2)^{-n/2} e^{-\frac{n \hat{\sigma}^2_{MLE}/2}{b^2}} K(b^2 | \theta).$$

$$K(b^2 | \theta) = ? \text{ to get conjugacy? } = (b^2)^a e^{-\frac{b}{b^2}} \\ = (b^2)^{-n/2+a} e^{-\frac{n \hat{\sigma}^2_{MLE}/2 + b}{b^2}}$$

$\propto \text{InvGamma}$  is the conjugate prior.

$$\text{let } P(b^2 | \theta) = \text{InvGamma}(\alpha, \beta)$$

$$\Rightarrow P(b^2 | x, \theta) \propto \left( (b^2)^{-n/2} e^{-\frac{n \hat{\sigma}^2_{MLE}/2}{b^2}} \right) \cdot \left( (b^2)^{-\alpha-1} e^{-\frac{\beta}{b^2}} \right)$$

$$= (b^2)^{-(\frac{n}{2}+\alpha)-1} e^{-\frac{n \hat{\sigma}^2_{MLE}/2 + \beta}{b^2}}$$

$$\propto \text{InvGamma} \left( \frac{n}{2} + \alpha, \frac{n \hat{\sigma}^2_{MLE}}{2} + \beta \right) = \mathcal{I}(\bar{x} - \theta)^2$$

$$\text{let } \alpha = \frac{n_0}{2}, \quad \beta = \frac{n_0 \hat{\sigma}_0^2}{2} \Rightarrow P(b^2 | \theta) = \text{InvGamma} \left( \frac{n_0}{2}, \frac{n_0 \hat{\sigma}_0^2}{2} \right)$$

$$\Rightarrow P(b^2 | \theta, x) = \text{InvGamma} \left( \frac{n+n_0}{2}, \frac{n \hat{\sigma}^2_{MLE} + n_0 \hat{\sigma}_0^2}{2} \right)$$

$$\text{Pseudodata: } Y_1, \dots, Y_{n_0} \sim \mathcal{N}(\theta, \hat{\sigma}_0^2)$$

$\uparrow$  known  $\quad \leftarrow$  belief

$$\Rightarrow n_0 \hat{\sigma}_0^2 = \mathcal{I}(\bar{x}_0 - \theta)^2$$

$$\Rightarrow \hat{\sigma}_0^2 = \mathcal{I}(\bar{x}_0 - \theta)^2 / (n_0) \quad \underline{(n_0) \text{ strength}}$$

$n_0$  small  $\Rightarrow$  uninformative

Haldane:  $n_0$  is nothing.  $n_0 = 0, \hat{\sigma}_0^2 = ?$

$$\Rightarrow P(b^2 | \theta) = \text{InvGamma}(0, 0)$$