

# Lecture 15

11

04/02/20

$\tilde{F}$ : iid Normal( $\theta, \sigma^2$ ) with  $\sigma^2$  known

$$p(\theta | \sigma^2) = N(\mu_0, \tau^2) = N(\mu_0, \frac{\sigma^2}{n_0})$$

Imagine pseudodata

$$Y_1, \dots, Y_n \stackrel{iid}{\sim} N(\mu_0, \sigma^2) \Rightarrow \bar{Y} \sim N(\mu_0, \frac{\sigma^2}{n_0})$$

$$p(\theta | X, \sigma^2) = N(\hat{\theta}_p, \sigma_p^2) \text{ where ...}$$

$$\hat{\theta}_p = \frac{\frac{n\bar{X}}{\sigma^2} + \frac{\mu_0}{\tau^2}}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}} = \frac{n\bar{X} + n_0 \mu_0}{n + n_0}$$

$$\sigma_p^2 = \frac{1}{\frac{n}{\sigma^2} + \frac{1}{\tau^2}} = \frac{\sigma^2}{n + n_0}$$

$$n_X = 1 \quad p(X_* | X, \sigma^2) = \int_{\Theta} p(X_* | \theta, \sigma^2) p(\theta | X, \sigma^2) d\theta$$

$$= \int_{\mathbb{R}} N(\theta, \sigma^2) N(\hat{\theta}_p, \sigma_p^2) d\theta$$

$$= \int_{\mathbb{R}} \left( \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(X_* - \theta)^2} \right) \left( \frac{1}{\sqrt{2\pi\sigma_p^2}} e^{-\frac{1}{2\sigma_p^2}(\theta - \hat{\theta}_p)^2} \right) d\theta$$

$$\propto \int_{\mathbb{R}} e^{-\frac{1}{2\sigma^2}(X_* - \theta)^2} e^{-\frac{1}{2\sigma_p^2}(\theta - \hat{\theta}_p)^2} d\theta$$

$$= \int_{\mathbb{R}} e^{-\frac{X_*^2}{2\sigma^2}} e^{\frac{X_*\theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}} e^{-\frac{\theta^2}{2\sigma_p^2}} e^{\frac{\theta\hat{\theta}_p}{\sigma_p^2}} e^{-\frac{\hat{\theta}_p^2}{2\sigma_p^2}} d\theta$$

Also not a function of  $X_*$

[2]

$$\propto e^{-\frac{X_*^2}{2\sigma^2}} \int_{\mathbb{R}} e^{\frac{X_* \theta}{\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2}} e^{-\frac{\theta^2}{2\sigma^2 p}} e^{\frac{\theta \hat{\theta}_p}{\sigma^2 p}} d\theta$$

$$= e^{-\frac{X_*^2}{2\sigma^2}} \int_{\mathbb{R}} e^{\underbrace{\left(\frac{X_*}{\sigma^2} + \frac{\hat{\theta}_p}{\sigma^2 p}\right)\theta}_a - \underbrace{\left(\frac{1}{2\sigma^2} + \frac{1}{2\sigma^2 p}\right)\theta^2}_b} d\theta$$

$$= e^{-\frac{X_*^2}{2\sigma^2}} \int_{\mathbb{R}} e^{-\frac{1}{2\left(\frac{1}{2b}\right)}\left(\theta - \frac{a}{2b}\right)^2} d\theta$$

$$\theta \sim N\left(\frac{a}{2b}, \frac{1}{2b}\right) = \frac{1}{\sqrt{2\pi\left(\frac{1}{2b}\right)}} e$$

$$= \sqrt{\frac{b}{\pi}} e^{-b\left(\theta^2 - \frac{a\theta}{b} + \frac{a^2}{4b^2}\right)} = \sqrt{\frac{b}{\pi}} e^{-b\theta^2 + a\theta - \frac{a^2}{4b}}$$

$$= \underbrace{\sqrt{\frac{b}{\pi}} e^{-\frac{a^2}{4b}}}_C e^{\underbrace{a\theta - b\theta^2}_{k(\theta)}} \propto e^{a\theta - b\theta^2}$$

$$= e^{-\frac{X_*^2}{2\sigma^2}} \sqrt{\frac{b}{\pi}} e^{-\frac{X_*^2}{2\sigma^2}} e^{\frac{\left(\frac{X_*}{\sigma^2} + \frac{\hat{\theta}_p}{\sigma^2 p}\right)^2}{4b}}$$

$$= e^{-\frac{X_*^2}{2\sigma^2}} e^{\frac{X_*^2}{4\sigma^2 b}} e^{\frac{X_* \hat{\theta}_p}{2\sigma^2 p \sigma^2 b}} e^{\frac{\hat{\theta}_p^2}{4\sigma^2 p b}}$$

$$\propto e^{\underbrace{\frac{\hat{\theta}_p}{2\sigma^2 \sigma_p^2 b} X_*}_A - \underbrace{\left(\frac{1}{2\sigma^2} - \frac{1}{4\sigma^2 b}\right) X_*^2}_B}$$

$$= e^{AX_* - BX_*^2} \propto N\left(\frac{A}{2B}, \frac{1}{2B}\right)$$

$$2b\sigma^2 = 2 \left( \frac{1}{2\sigma^2} + \frac{1}{2\sigma_p^2} \right) \sigma^2 = 1 + \frac{\sigma^2}{\sigma_p^2}$$

$$\Rightarrow \frac{1}{2B} = \frac{1}{2 \left( \frac{1}{2\sigma^2} + \frac{1}{2\sigma_p^2} \right)} = \frac{1}{\frac{1}{\sigma^2} + \frac{1}{2b\sigma^2}}$$

$$= \frac{\sigma^2}{\sigma^2 - \frac{1}{2b}} = \frac{2b\sigma^2 \sigma^2}{2b\sigma^2 - 1} = \frac{\left(1 + \frac{\sigma^2}{\sigma_p^2}\right) \sigma^2}{\left(1 + \frac{\sigma^2}{\sigma_p^2}\right) - 1}$$

$$= \frac{\left(1 + \frac{\sigma^2}{\sigma_p^2}\right) \sigma^2}{\frac{\sigma^2}{\sigma_p^2}} = \sigma_p^2 \left(1 + \frac{\sigma^2}{\sigma_p^2}\right) = \sigma_p^2 + \sigma^2$$

$$\frac{A}{2B} = A(\sigma_p^2 + \sigma^2) = \frac{\hat{\theta}_p}{2\sigma_p^2 \sigma^2} (\sigma_p^2 + \sigma^2)$$

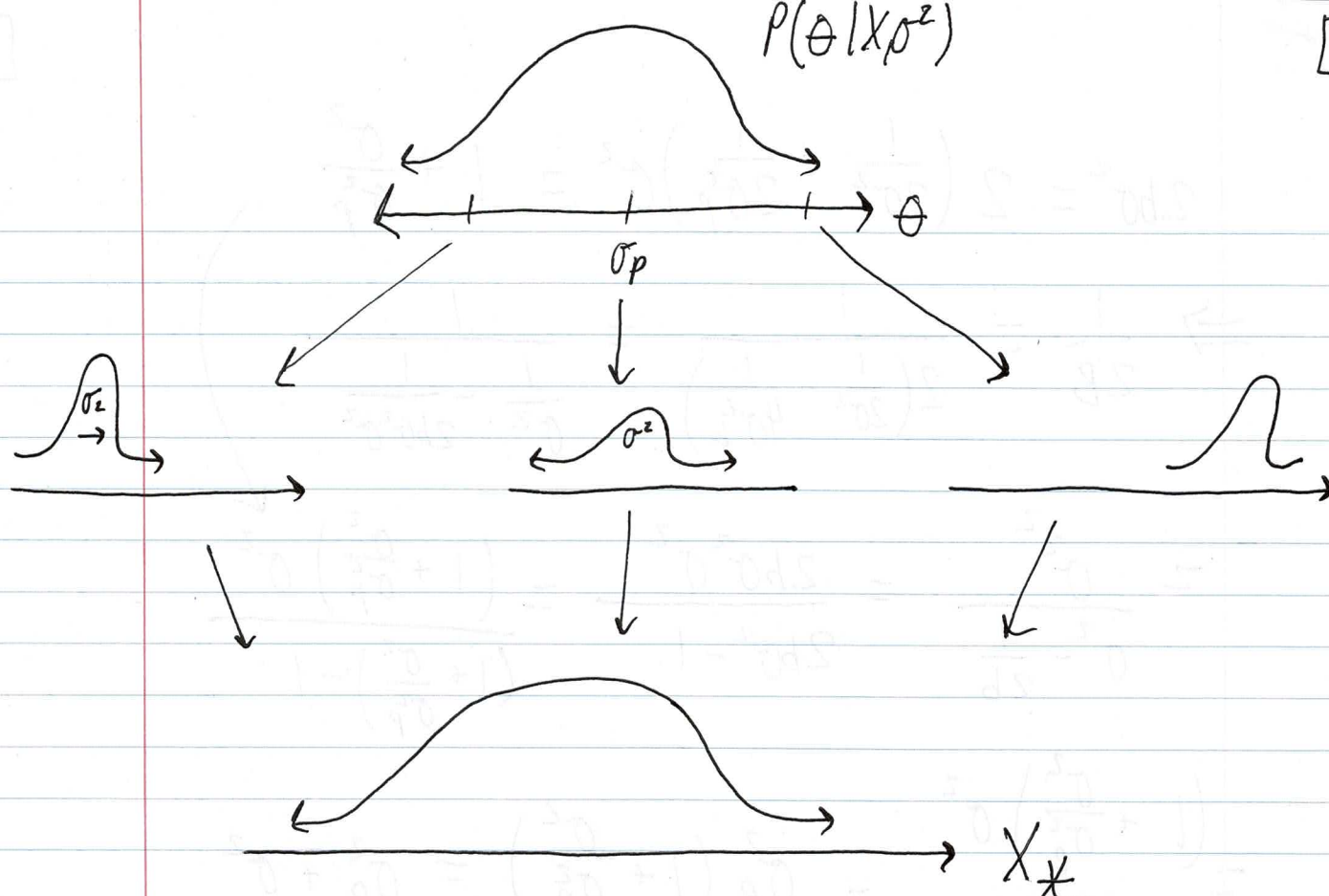
$$= \frac{\hat{\theta}_p}{\left(1 + \frac{\sigma^2}{\sigma_p^2}\right) \sigma_p^2} (\sigma_p^2 + \sigma^2) = \frac{\hat{\theta}_p}{\sigma_p^2 + \sigma^2} \sigma_p^2 + \sigma^2 = \hat{\theta}_p$$

$$P(X_* | X, \sigma^2) = \mathcal{N}(\hat{\theta}_p, \sigma_p^2 + \sigma^2)$$

$$= \mathcal{N}(\hat{\theta}_p, \left(1 + \frac{1}{n+n_0}\right) \sigma^2) \rightarrow \mathcal{N}(\theta, \sigma^2)$$

$$\hat{\theta}_p \rightarrow \theta, \quad \sigma_p^2 + \sigma^2 \rightarrow \sigma^2$$





An overdispersed normal with known Variance is normal.

$\tilde{F}$ : iid  $N(\theta, \sigma^2)$  with  $\theta$  known

$$L(\sigma^2; X, \theta) \Rightarrow (2\pi\sigma^2)^{-n/2} e^{-\frac{1}{2\sigma^2} \sum (X_i - \theta)^2}$$

$$l(\sigma^2; X, \theta) = -\frac{n}{2} \ln(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum (X_i - \theta)^2$$

$$= -\frac{n}{2} \ln(2\pi) - \frac{n}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum (X_i - \theta)^2$$

$$l'(\sigma^2; X, \theta) = -\frac{n}{2} \frac{1}{\sigma^2} + \frac{1}{2\sigma^2^2} \sum (X_i - \theta)^2 \stackrel{\text{Set}}{=} 0 \Rightarrow -n + \frac{\sum (X_i - \theta)^2}{\sigma^2} = 0$$

$$\hat{\sigma}_{MLE}^2 = \frac{\sum (X_i - \theta)^2}{n}$$

see previous

$$\sum (X_i - \sigma)^2 = n \hat{\sigma}_{MLE}^2$$

Similar to  $S^2 = \frac{\sum (X_i - \bar{X})^2}{n-1} \rightarrow \sigma^2$

Laplace Prior:  $P(\sigma^2|\theta) \propto 1$  improper since  $\sigma^2 \in (0, \infty)$

$$P(\sigma^2|X, \theta) \propto P(X|\theta, \sigma^2) = (2\pi\sigma^2)^{-n/2} e^{-\frac{n\hat{\sigma}_{MLE}^2}{2\sigma^2}}$$

$$\propto (\sigma^2)^{-\frac{n}{2}} e^{-\frac{n\hat{\sigma}_{MLE}^2/2}{\sigma^2}}$$

$$U \sim \text{Gamma}(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} U^{\alpha-1} e^{-\beta U}$$

$$V = \frac{1}{U} = t(u) \sim f_u(t^{-1}(v)) \left| \frac{d}{dv} [t^{-1}(v)] \right|$$

$$t^{-1}(v) = \frac{1}{v} = \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{1}{v}\right)^{\alpha-1} e^{-\frac{\beta}{v}} \left| -\frac{1}{v^2} \right|$$

$$= \frac{\beta^\alpha}{\Gamma(\alpha)} v^{-\alpha+1} v^{-2} e^{-\frac{\beta}{v}} = \frac{\beta^\alpha}{\Gamma(\alpha)} v^{-\alpha-1} e^{-\frac{\beta}{v}}$$

$$= \text{Inverse Gamma}(\alpha, \beta)$$

inverse gamma r.v.

$$\propto k(v) = v^{-\alpha-1} e^{-\frac{\beta}{v}}$$

$$P(\sigma^2|X, \theta) = \sigma^{2 \cdot \underbrace{-(\frac{n}{2}-1)-1}} e^{-\frac{n\hat{\sigma}_{MLE}^2/2}{\sigma^2}} \quad \leftarrow \beta$$

$$= \text{InvGamma}\left(\frac{n}{2}-1, \frac{n\hat{\sigma}_{MLE}^2}{2}\right)$$

$$= \text{InvGamma}\left(\frac{n-2}{2}, \frac{n\hat{\sigma}_{MLE}^2}{2}\right)$$

$$P(\sigma^2 | X, \sigma^2) \propto P(X | \theta, \sigma^2) P(\sigma^2 | \theta) \propto (\sigma^2)^{-\frac{n}{2}} e^{-\frac{n \hat{\sigma}_{MLE}^2 / 2}{\sigma^2}}$$

$$k(\sigma^2 | \theta) \quad k(\sigma^2 | \theta) = ? \text{ to get conjugacy?}$$

$$= (\sigma^2)^a e^{-\frac{b}{\sigma^2}} = (\sigma^2)^{-\frac{n}{2} + a} e^{-\frac{-n \hat{\sigma}_{MLE}^2 / 2 + b}{\sigma^2}}$$

$\propto$  InvGamma

is the conjugate prior

$$\text{Let } P(\sigma^2 | \theta) = \text{InvGamma}(\alpha, \beta)$$

$$\Rightarrow P(\sigma^2 | X, \theta) \propto \left( (\sigma^2)^{-n/2} e^{-\frac{n \hat{\sigma}_{MLE}^2 / 2}{\sigma^2}} \right) \left( (\sigma^2)^{-\alpha-1} e^{-\frac{\beta}{\sigma^2}} \right)$$

$$= (\sigma^2)^{-(\frac{n}{2} + \alpha) - 1} e^{-\frac{(n \hat{\sigma}_{MLE}^2 / 2) + \beta}{\sigma^2}} \propto \text{InvGamma}\left(\frac{n}{2} + \alpha, \frac{n \hat{\sigma}_{MLE}^2 / 2 + \beta}{2}\right)$$

$$\text{Let } \alpha = \frac{n_0}{2}, \quad \beta = \frac{n_0 \sigma_0^2}{2}$$

$$P(\sigma^2 | \theta) = \text{InvGamma}\left(\frac{n_0}{2}, \frac{n_0 \sigma_0^2}{2}\right)$$

$$\Rightarrow P(\sigma^2 | \theta, X) = \text{InvGamma}\left(\frac{n + n_0}{2}, \frac{n \hat{\sigma}_{MLE}^2 + n_0 \sigma_0^2}{2}\right)$$

Pseudodata:  $Y_1, \dots, Y_{n_0} \sim \mathcal{N}(\theta, \sigma_0^2)$

$\downarrow$  known       $\downarrow$  belief

$$\Rightarrow n_0 \sigma_0^2 = \sum (Y_i - \theta)^2$$

$$\Rightarrow \sigma_0^2 = \frac{\sum (Y_i - \theta)^2}{n_0}$$

$n_0$  small  $\rightarrow$  Uninformative

$$\text{Haldane: } n_0 = 0, \sigma_0^2 = ? \Rightarrow P(\sigma^2 | \theta) = \text{InvGamma}(0, 0)$$