

# Lecture 6

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02/13/20

Beta Function:  $B(\alpha, \beta) = \int_0^1 t^{\alpha-1} (1-t)^{\beta-1} dt$

Parameter Space:  $\alpha, \beta > 0$        $\text{supp}(Y) = (0, 1)$

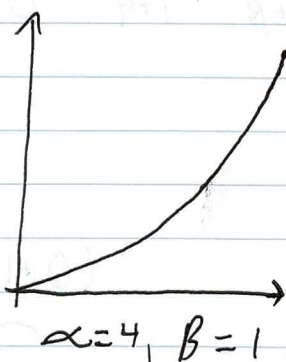
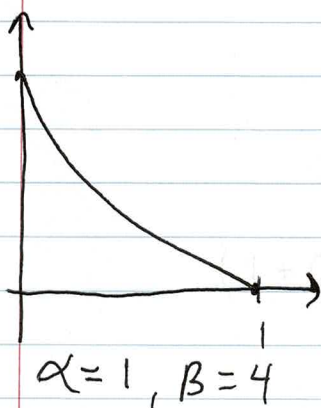
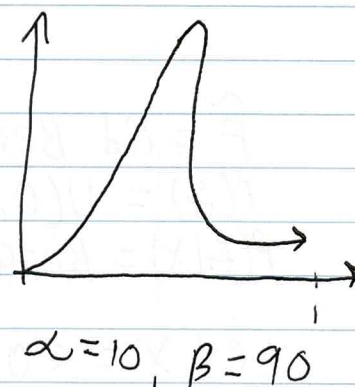
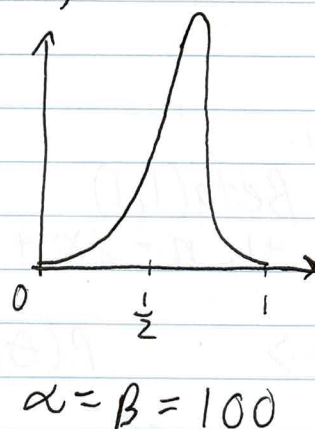
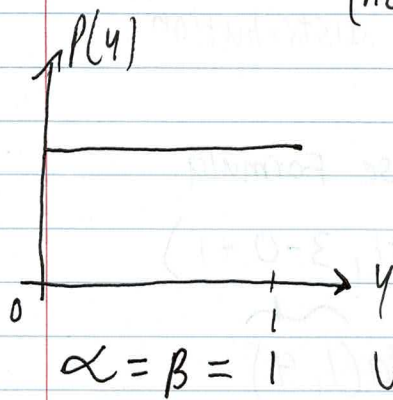
$$Y \sim \text{Beta}(\alpha, \beta) = \frac{1}{B(\alpha, \beta)} y^{\alpha-1} (1-y)^{\beta-1}$$

$$E[Y] = \frac{\alpha}{\alpha + \beta}$$

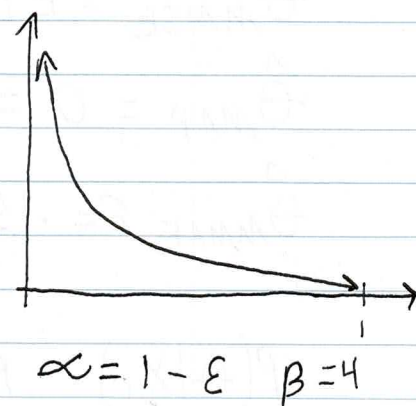
$$\text{Mode}[Y] = \frac{\alpha - 1}{\alpha + \beta - 2} \quad \alpha, \beta \geq 1$$

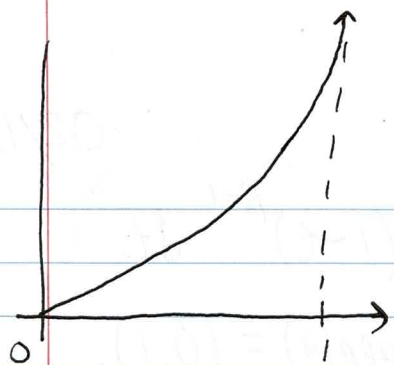
$$\text{Med}[Y] = q_{\text{beta}}(.5, \alpha, \beta)$$

(no closed form)

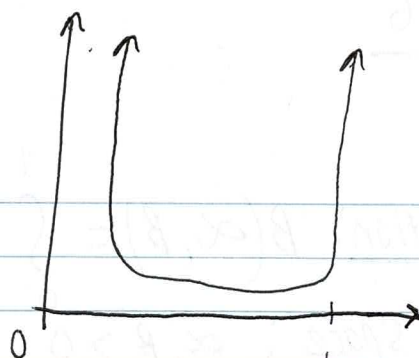


$\epsilon \approx 70$   
but small

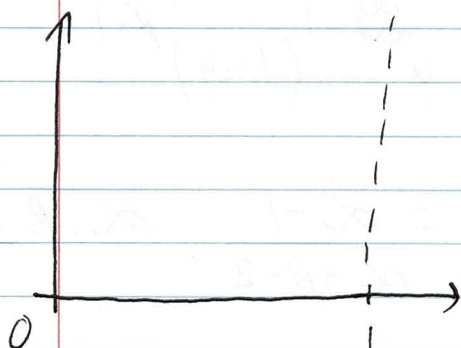




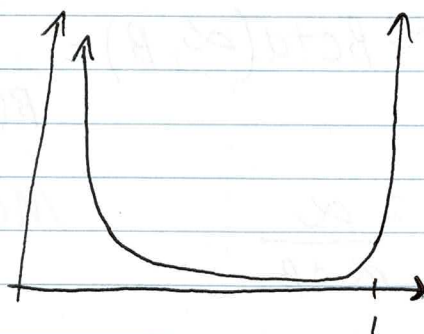
$$\alpha=4, \beta=1-\epsilon$$



$$\alpha=\beta=1-\epsilon$$



$$\alpha=\beta=\epsilon$$



$$\alpha=\beta=\frac{1}{2}$$

arcsine distribution

$\tilde{F}$  = iid Bernoulli

$$P(\theta) = U(0,1) = \text{Beta}(1,1)$$

$$P(\theta|x) = \text{Beta}(\sum x_i + 1, n - \sum x_i + 1)$$

use Formula

$$(0+1, 3-0+1)$$

$$\text{Let } X = \langle 0, 0, 0 \rangle$$

$$P(\theta|x) = \text{Beta}(1, 4)$$

$$\hat{\theta}_{\text{MMSE}} = E[\theta|x] = \frac{\alpha}{\alpha+\beta} = \frac{1}{1+4} = 0.2$$

$$\hat{\theta}_{\text{MAP}} = 0 = \hat{\theta}_{\text{MLE}}$$

$$\hat{\theta}_{\text{MMAE}} \approx .159$$

$$(0+1, 1-0+1)$$

$$P(\theta|x_1) = \frac{P(x_1|\theta)P(\theta)}{P(x_1)} = \text{Beta}(1, 2)$$

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$$P(\theta | x_1, x_2) = \frac{P(x_1, x_2 | \theta) P(\theta | x_1)}{P(x_1, x_2)} = \text{Beta}(1, 3)$$

$\xrightarrow{\text{Beta}(1,2)}$

$$P(\theta | x_1, x_2, x_3) = \frac{P(x_1, x_2, x_3 | \theta) P(\theta | x_1, x_2)}{P(x_1, x_2, x_3)} = \text{Beta}(1, 4)$$

$\xrightarrow{\text{Beta}(1,3)}$

$\tilde{F} = \text{i.i.d Bernoulli}$   $X$  is  $n$  observations

$$P(\theta) = \text{Beta}(\alpha, \beta)$$

$$\begin{aligned}
 P(\theta | x) &= \frac{P(x | \theta) P(\theta)}{\int_0^1 P(x | \theta) P(\theta) d\theta} \\
 &= \frac{\underbrace{\left( \theta^{\sum x_i} (1-\theta)^{n-\sum x_i} \right)}_{\text{likelihood}} \cancel{\frac{1}{\text{Beta}(\alpha, \beta)}} (1-\theta)^{\beta-1}}{\int_0^1 \theta^{\sum x_i} (1-\theta)^{n-\sum x_i} \cancel{\frac{1}{\text{Beta}(\alpha, \beta)}} (1-\theta)^{\beta-1} d\theta} \quad \text{cancel terms} \\
 &= \frac{\theta^{\sum x_i + \alpha - 1} (1-\theta)^{n - \sum x_i + \beta - 1}}{\int_0^1 \theta^{\sum x_i + \alpha - 1} (1-\theta)^{n - \sum x_i + \beta - 1} d\theta} \\
 &= \frac{1}{B(\sum x_i + \alpha, n - \sum x_i + \beta)} \theta^{\sum x_i + \alpha - 1} (1-\theta)^{n - \sum x_i + \beta - 1} \\
 &= \text{Beta}(\sum x_i + \alpha, n - \sum x_i + \beta)
 \end{aligned}$$



$$p(\theta) \xrightarrow{x} p(\theta|x)$$

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$$\text{Beta}(\alpha, \beta) \xrightarrow{X} \text{Beta}(\underbrace{\sum x_i + \alpha}_{\alpha'}, \underbrace{n - \sum x_i + \beta}_{\beta'})$$

Prior Parameters

Posterior Parameters

$$\hat{\theta}_{\text{MMSE}} = E[\theta|x] = \frac{\sum x_i + \alpha}{n + \alpha + \beta}$$

$$\hat{\theta}_{\text{MMAE}} = \text{med}[\theta|x] = \text{qbeta}(.5, \sum x_i + \alpha, n - \sum x_i + \beta)$$

$$\hat{\theta}_{\text{MMF}} = \text{mode}[\theta|x] = \begin{cases} \sum x_i + \alpha - 1 & \text{if } \alpha + x > 1 \\ n - \sum x_i + \beta - 2 & \text{if } \beta + n - x > 1 \end{cases}$$

Conjugacy for a given likelihood model means the prior and the posterior have the same r.v. (different parameters). Beta is the conjugate prior to the iid Bernoulli likelihood.

$$\tilde{F} = \text{Bin}(n, \theta) \quad \text{with } n \text{ fixed, known} \\ \Downarrow \quad \text{and } \theta \text{ unknown}$$

$$\binom{n}{x} \theta^x (1-\theta)^{n-x}$$

Let  $X = \sum x_i$  in the iid Bernoulli model  
 $n - X = n - \sum x_i$

# of failures

$$p(\theta) = \text{Beta}(\alpha, \beta) \Rightarrow p(\theta|x) = \text{Beta}(\underbrace{x + \alpha}_{\text{\# of successes}}, \underbrace{n - x + \beta}_{\text{\# of failures}})$$

# of successes

# of prior success or # of pseudosuccess

# of prior failure or # of pseudo failures

pseudocounts

$$p(\theta) = U(0,1) = \text{Beta}(\overset{\alpha}{1}, \overset{\beta}{1})$$

Principle of indifference  $n_0 = 2$

$$E[\theta] = \frac{1}{2}$$

The principle of indifference is "not so indifferent" because it contains information

Default Point Estimate  $\leftarrow \hat{\theta}_{MMSE} = E[\theta|x] = \frac{x+\alpha}{n+\alpha+\beta} = \frac{x}{n+\alpha+\beta} \left(\frac{n}{n}\right) + \frac{\alpha}{n+\alpha+\beta} \left(\frac{\alpha+\beta}{\alpha+\beta}\right)$

Shrinkage Estimator

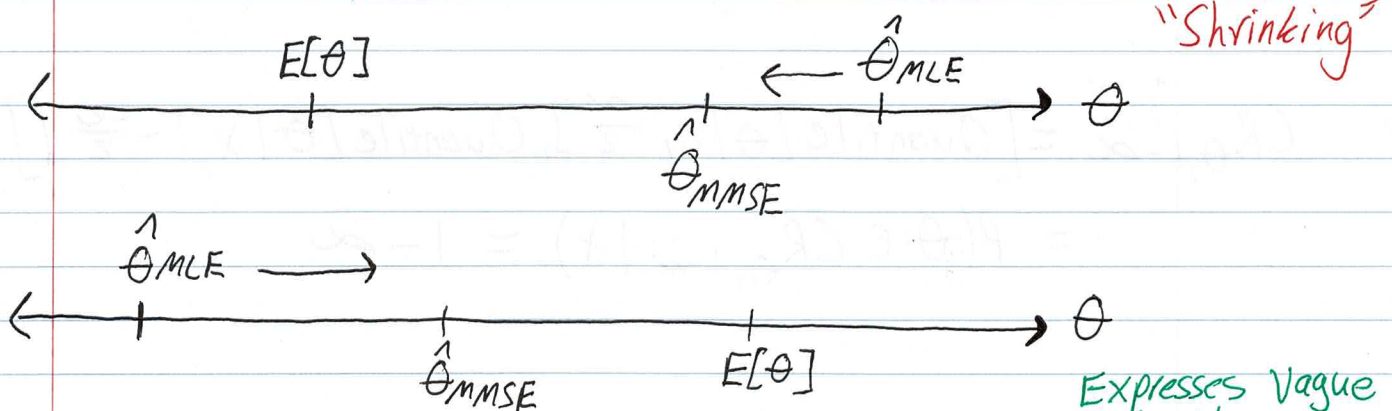
$$= \underbrace{\frac{n}{n+\alpha+\beta}}_{1-p \rightarrow \rho} \cdot \frac{x}{n} + \underbrace{\frac{\alpha+\beta}{n+\alpha+\beta}}_{p \rightarrow \rho} \cdot \frac{\alpha}{\alpha+\beta}$$

$$\hat{\theta}_{MMSE} = (1-p) \hat{\theta}_{MLE} + p E[\theta]$$

Linear Combination  $\vec{w} + \vec{x}$

If  $n$  is large  
"weigh" this more

If  $n$  is small,  
"weigh" this more



If  $\alpha, \beta$  is small  $\Rightarrow$  prior Uninformative

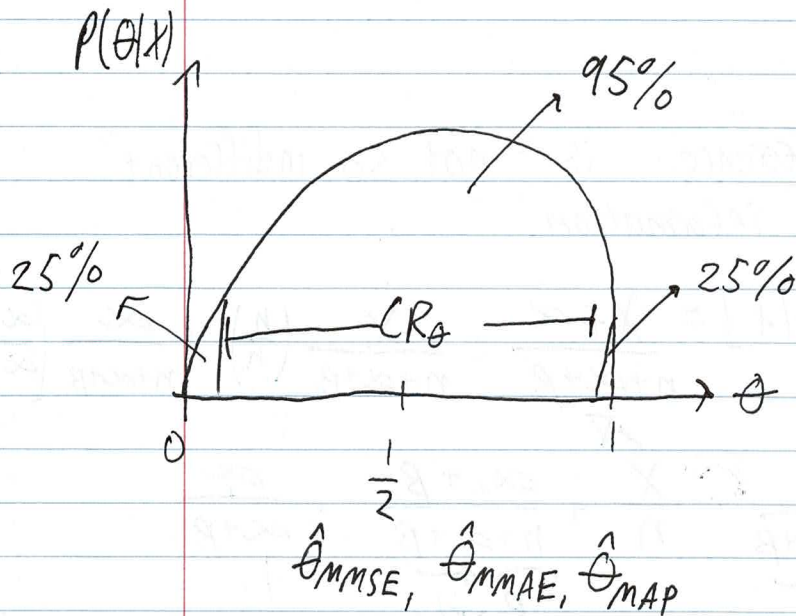
If  $\alpha, \beta$  is large  $\Rightarrow$  prior informative

Expresses Vague info about a Variable

Expresses definite info about a Variable

$\tilde{F} = \text{Binomial}$   $X=1, n=2$

$p(\theta)$  prior of indifference  $\Rightarrow p(\theta|x) = \text{Beta}(2,2)$



2<sup>nd</sup> goal of inference: Confidence set

$$CI_{\theta, 95\%} = \left[ 0.5 \pm 2\sqrt{\frac{0.5(1-0.5)}{2}} \right] = [-0.2, 1.21]$$

i.e. nonsense  $\rightarrow$  Out of range of possible  $\theta$ 's

Bayesian Credible Ranges (CR)

$$\begin{aligned} CR_{\theta, 1-\alpha} &= [\text{Quantile}[\theta|x, \frac{\alpha}{2}], \text{Quantile}[\theta|x, 1-\frac{\alpha}{2}]] \\ &= P(\theta \in CR_{\theta, 1-\alpha} | x) = 1-\alpha \end{aligned}$$