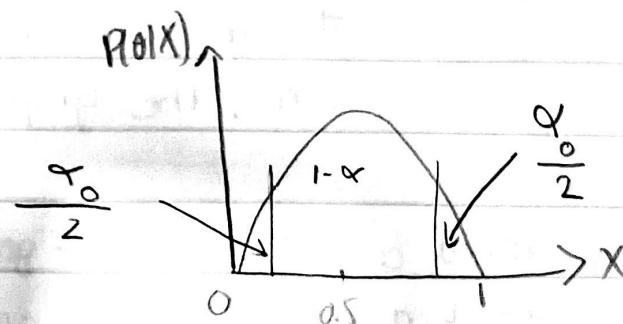


March 1.

lec 8.

$$X=1, n=2, F: \text{Bin}(n, \theta)$$

$$P(\theta) = \text{Beta}(1,1) \Rightarrow P(\theta|X) = \text{Beta}(2,2)$$



produced a 95%
credible region
for θ .

10.15 (1M)

$$CR_{\theta, 1-\alpha_0} := \left[Q\left[\theta|x, \frac{\alpha_0}{2}\right], Q\left[\theta|x, 1 - \frac{\alpha_0}{2}\right] \right]$$

$$\rightarrow = [qbeta(2.5\%, 2, 2), qbeta(97.5\%, 2, 2)] \\ = [0.094, 0.906]$$

$$P(\theta \in [0.094, 0.906] | x) = 95\%$$

This is a real probability statement! The CI approach cannot give you such a statement! The CR is highly interpretable. Also, the CR is a proper subset of θ for $\alpha_0 > 0$. This is not always true with CI's e.g. here the 95% CI for this data:

$$CI_{\theta, 95\%} = [0.5 \pm 1.96 \sqrt{\frac{0.5 - 0.5}{2}}] = [-0.21, 1.21] \notin \Theta = (0, 1)$$

The above CR is technically a two-sided CR. You can also create one-side (i.e. left-sided or right-sided) CR's.

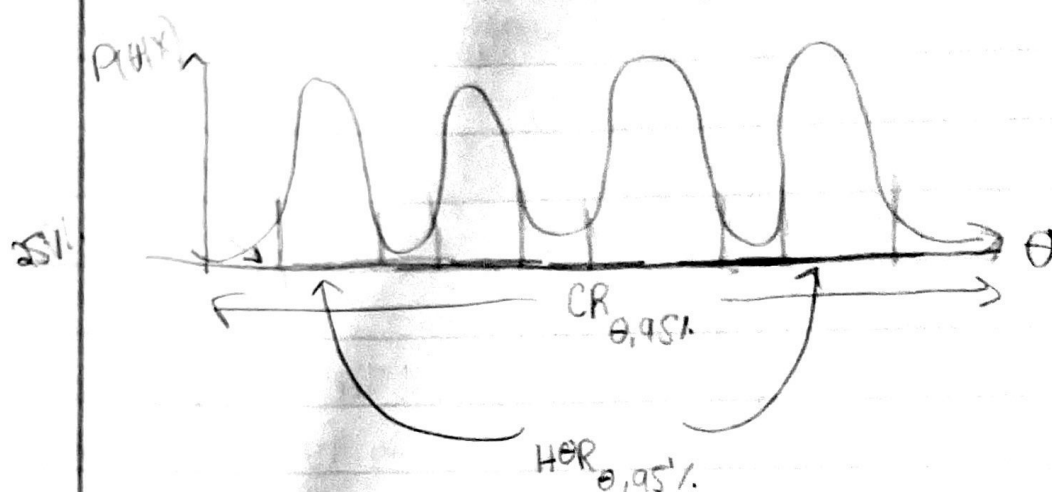
$$CR_{L, \theta, 1-\alpha_0} := [\text{smallest value in } \theta \text{ or } -\infty, Q[0|x, 1-\alpha_0]]$$

ex: in our dataset $qbeta(0.95, 2, 2)$

$$\hookrightarrow = [0, 0.865] = P(\theta < 0.865 | x) = 95\%$$

$$CR_{R, \theta, 1-\alpha_0} := [Q[\theta|x, \alpha_0], \text{large value in } \theta \text{ or } \infty] \\ \rightarrow = [qbeta(0.05, 2, 2), 1] = [0.131, 1] \Rightarrow P(\theta > 0.136 | x) = 95\%$$

Another approach (which we will see but not study further) is called the high density region (HDR) approach. Consider the following posterior for θ :



$P(\theta \in HDR_{\theta, 95\%}) = 95\%$ but it has "minimum" width of the interval $P(\theta \in \text{minimum})$

Sometimes the $CR = HDR$ (ex in unimodal posteriors)

Disadvantages of the HDR approach. (1) it can be non-contiguous ex. in pieces (2) it's computationally intense. (3) no L or R intervals.

It's a dirty solution.

Bayesian Hypothesis Testing. We can immediately compute the following quantities:

Bayesian P-value $:= P(H_0|x), P(H_a|x)$

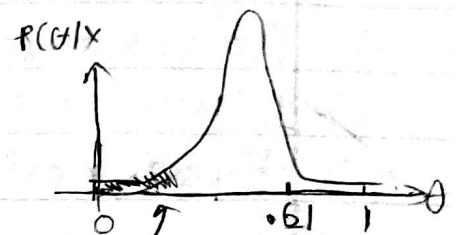
threshold of "sufficient evidence"

if $P(H_0 | x) < \alpha_0 = \text{Reject } H_0$

Let's recreate the hypothesis testing example from Lec. 10. Flips of a coin where $x=61$ were heads. Test if the coin is unfairly weighted towards heads.

$H_a: \theta > 0.5 \Rightarrow H_0: \theta \leq 0.5$. Assume $P(\theta) = \text{Beta}(1,1)$

$P(\theta | x) = \text{Beta}(\underbrace{61+1}_{\text{prior}}, \underbrace{39+1}_{\text{Beta}})$



$P(\theta | x) = \text{Beta}(62, 40)$

$P(H_0 | x) = P(\theta \leq 0.5 | x) = \int_0^{0.5} P(\theta | x) d\theta$ posterior

$= \int_0^{0.5} \frac{1}{\beta(62, 40)} \theta^{61} (1-\theta)^{39} d\theta$

$= \text{pbeta}(0.5, 62, 40)$

$= .014 = 1.4\%$

\Rightarrow Reject H_0 i.e. coin is unfairly weighted towards heads

uber driver does $n=200$ rides and got 37 non-star rating. If his true proportion of non-5 star rating is less than 25%, then uber policy is to not fire the driver. Prove he should be fired (or not) at a 5% significant level.

$$H_a: \theta > 25\%$$

$$\Rightarrow H_0: \theta \leq 25\%$$

$$P(\theta) = \text{Beta}(1,1)$$

$$F: (n, \theta)$$

$$P(\theta|x) = \text{Beta}(1+37, 1+200-37)$$

$$\text{Beta}(38, 164)$$

posterior

$$P(H_0|x) = P(\theta \leq 0.25|x) = \int_0^{0.25} d\theta = 1 - \int_{0.25}^1 d\theta$$

$$= 1 - \text{pbeta}(0.25, \underbrace{38}_{\alpha}, \underbrace{164}_{\beta})$$

$P(\theta|x)$

$$1 - \text{pbeta}(0.25, 38, 164) = 0.983 > 5\%$$

Retain H_0 . Don't fire him

Let's test the coin again. Flip $\overset{n}{\sim} 100$ times and get 43 heads. Test if the coin is unfair at 5% significance.
 \tilde{x} $H_a: \theta \neq 0.5 \Rightarrow H_0: \theta = 0.5$

$$P(\theta) = \text{Beta}(1,1) \Rightarrow P(\theta|x) = \text{Beta}(44, 58)$$

$$P(H_0|x) = P(\theta = 0.5|x) = 0 \quad \text{Reject } H_0 \text{ always ?? Yes...}$$

Using this approach, two sided test are always rejected (if the posterior is continuous). Does this make sense? Any infinitely precise theory of theta is wrong in the real world. A coin is never exactly 50.0000000... % likely to flip heads. So, we need to slightly reframe our hypothesis using a notation of "margin of equivalence" called δ (delta).

$$H_a: \theta \notin [\theta_0 \pm \delta] = H_0: \theta \in [\theta_0 \pm \delta]$$