

Some function of the data x_1, \dots, x_n

$$\hat{\theta} = S(\tilde{x}_1, \dots, \tilde{x}_n) = \bar{x} \in \mathbb{R}, \text{ e.g. } 0.198$$

\uparrow
 $F: \text{iid Bern}(\theta)$

Some function of the data x_1, \dots, x_n

by property z

$F: \text{iid Bern}(\theta)$

$$\hat{\theta}_{MLE} = S(x_1, x_2, \dots, x_n) = \bar{x} \sim N(\theta, SE[\hat{\theta}_{MLE}])$$

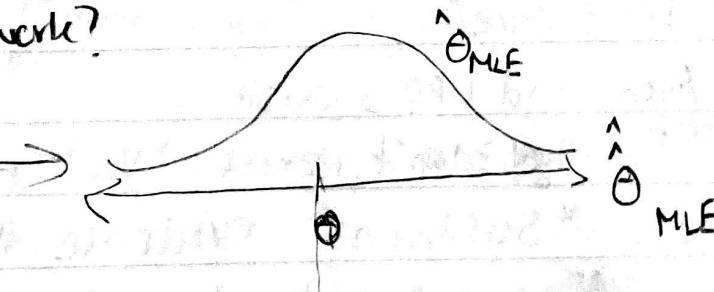
the same
function except
of the random
variable

a random variable, not a value
random variable have distribution

we use this normality to create the confident interval (CI)

Why do CI work?

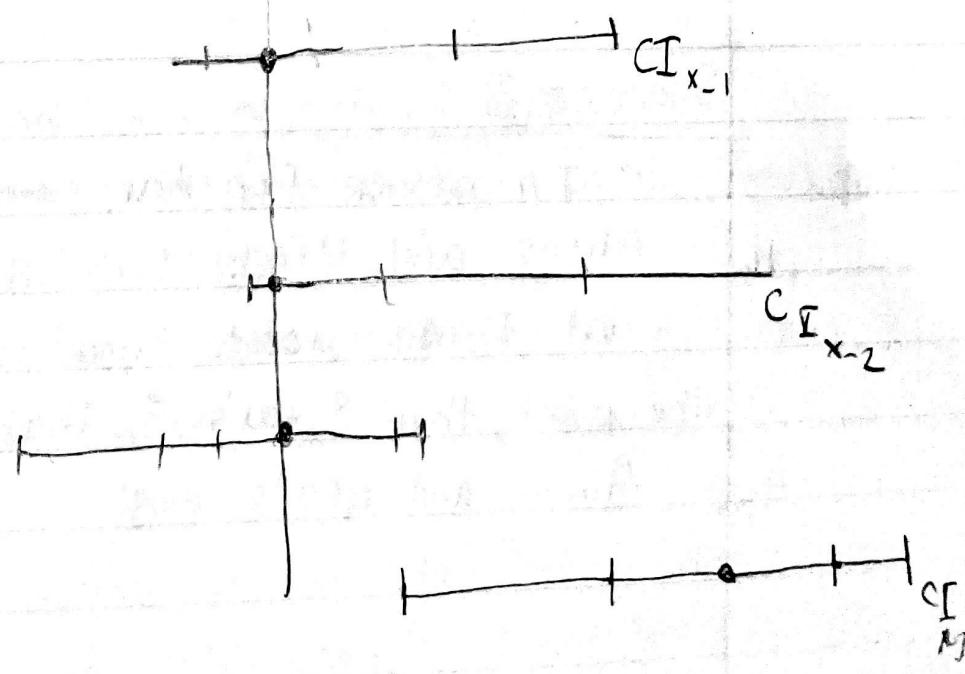
If the property is
write there is
a distribution



$$\alpha = 5\%$$

$$Z = 1.96$$

$\frac{\alpha}{2}$



CI's have 95% probability (before they're realized) of capturing (i.e. including) the real value of theta

Inference goal #3: "hypothesis testing" also called "theory testing"

Consider a situation where "I" am trying to convince "you" of something

Scenario I:

- * I declare

H_a = Alien and UFO's exist

If they don't exist *you* need to provide *me* *Sufficient* evidence that:

H_o = Alien and UFO's do not exist.

Scenario II:

- * I'll assume for the moment that

H_o = Aliens and UFO's don't exist

and I will provide *you* sufficient evidence to the point that *you're* that

H_a = Alien and UFO's exist

Scenario II is more convincing and it is how science generally works the theory I'm trying to demonstrate is called the "alternative hypothesis" (H_a) since it's alternative to maybe business-as-usual. In scenario II, you assume the opposite of the theory which is called the "null hypothesis" (H_0). This is the "hypothesis testing" procedure.

In our context, theories are phrases as mathematical statements about theta the unknown parameter. We will study three types of H_a 's:

$$H_a: \theta \neq \theta_0 \xleftarrow[\text{value}]{\text{constant}} \text{vs } H_0: \theta = \theta_0 \rightarrow \begin{array}{l} \text{two sided test/} \\ \text{two tail test} \end{array}$$

$$H_a: \theta > \theta_0 \quad H_0: \theta \leq \theta_0 \rightarrow \begin{array}{l} \text{right-side test/} \\ \text{right-side tail} \end{array}$$

$$H_a: \theta < \theta_0 \quad H_0: \theta \geq \theta_0 \rightarrow \begin{array}{l} \text{left-sided test/} \\ \text{left-tailed test} \end{array}$$

There are two outcome of a hypothesis test

You were not shown sufficient evidence of H_a . Thus, you "fail to reject H_0 " or "retain H_0 ".

You were indeed shown sufficient evidence of H_a . thus, you "reject H_0 " or "accept H_a ".

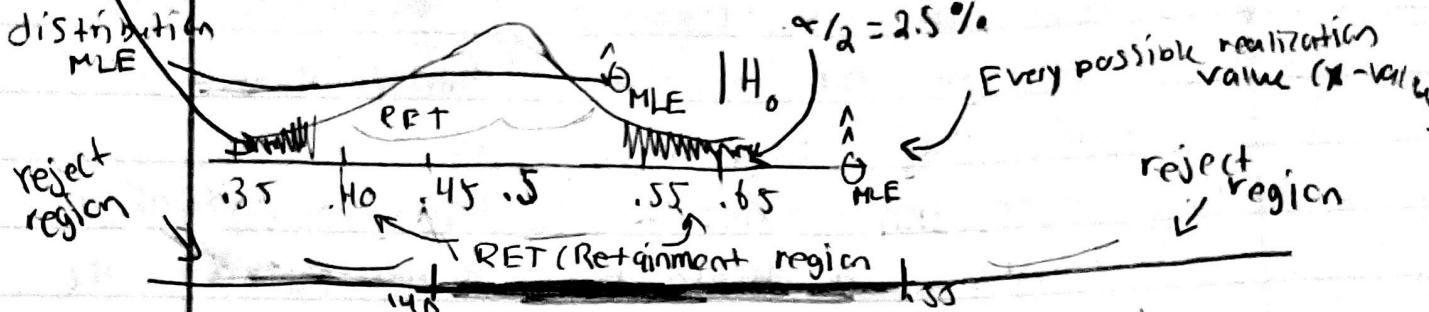
Imagine you're flipping a coin n times and you're counting the number of HEADS, then $\text{F} \sim \text{iid Bern}(\theta)$. You want prove the coin is unfair. \rightarrow

$$H_a: \theta \neq 0.5$$

$$H_0: \theta = 0.5$$

$$\alpha/2 = 2.5\%$$

$$\text{If } H_0 \text{ is true, } \theta = 0.5 \Rightarrow \hat{\theta}_{MLE} \sim N(\theta, SE[\hat{\theta}_{MLE}]^2)$$



What constitutes "sufficient evidence". It's a probability of rejecting when H_0 is true. (Denoted alpha). Everyone is different.

If alpha = 5% in a 2-tailed test, we put 1/2 the probability in each tail. 5% is the most common scientific standard. Thus, we retain the H_0 for the most non-weird 95% of the theta values and reject H_0 for the 5% most weird theta values.

$$\text{RET} = [\theta_0 \pm z_{\alpha/2} SE[\hat{\theta}_{MLE}]] = [0.5 \pm 1.96 \sqrt{\frac{0.5(1-0.5)}{100}}] \\ = [0.402, 0.590]$$

ex. if $\bar{x} = \frac{61}{100} = .61 \quad \alpha = 5\%$



$\hat{\theta}_{MLE} = .61 \notin RET \Rightarrow$ reject H_0 and conclude coin is unfair

e.g. if $\bar{x} = \frac{59}{100} = .59$

$\hat{\theta}_{MLE} = .59 \in RET \Rightarrow$ fail to reject H_0 and conclude there's not enough evidence of coin being unfair

We've covered the "frequentist" approach to statistical inference. But there are problems w/ it ...

① $F = \text{iid } \text{Bem}(0)$, $X = \langle 0, 0, 0 \rangle$

\hat{x} it's a vector; it has magnitude

$$\hat{\theta}_{MLE} = \bar{x} = 0$$

Is that a good point estimate? No. You shouldn't be able to say something is absolutely impossible after $n=3$ trials.

$$CI_{\theta, 1-\alpha} = \left[\theta \pm 1.96 \sqrt{\frac{\theta(1-\theta)}{3}} \right] = \{0\}$$

Is this good confidence set? No. This is not a good set of "reasonable values".

- (2) what if you had prior knowledge that θ was restricted to e.g. $[0, 1, 0.2]$ and not the full $(0, 1)$. Ans. \Rightarrow you can't "enter that into" your inference, by using MLE
- (3) Consider the frequentist interpretation of a CI.
- Before you do the experiment, you have a 95% probability of capturing theta. But this doesn't tell you anything about after your experiment. After your experiment you have an interval e.g. $[0.37, 0.43]$ and can't say:
- $$P(\theta \in [0.37, 0.43]) = 0.95$$
- can't say = \rightarrow no randomness; no probability
- 95% of CI's will cover theta. But again, I only make one. So this interpretation doesn't help! In conclusion, any specific CI means NOTHING.
- (4) Hypothesis tests result in a binary outcome: either you reject H_0 or you fail to reject H_0 . What if you want to know

$$\text{data } P(H_0 | X) \text{ or } P(H_a | X) \text{ data}$$

You cannot!!! One thing you can do:

$$P_{\text{val}} := P(\text{Seeing } \hat{\theta} \text{ or more extreme} | H_0) + P(H_0 | X)$$

⑤ \tilde{F} : iid $Bern(\theta)$, $x \in \{0, 1\}$, $\hat{\theta} = .33$

$$CI_{\theta, 95\%} = \left[.33 \pm 1.96 \sqrt{\frac{.33 \cdot .67}{3}} \right] = [-0.20, 0.87]$$

This is not legal Confident set, because it's outside of the legal parameter space which is $(0, 1)$

$$\alpha = 5\%$$

$$H_0: \theta = 0.5\% = RET = \left[0.5 \pm 1.96 \sqrt{\frac{.5 \cdot .5}{3}} \right] = [-.066, 1.066]$$

This is a good hypothesis test because it's never going to reject.

The problem in #5 is b/c the asymptotic normality of the MLE doesn't "kick in" until n is large (MLE property 2 is not true yet).