

Lecture 3 Math 390.03-02
2/8/16

$\text{Odds}(A) = \frac{P(A)}{1-P(A)}$ $\text{Odds against } (A) = \frac{1-P(A)}{P(A)}$

Bayes Factor:

$$\frac{P(B|A)}{P(B|A^c)} = 6 \text{ very strong evidence}$$

can't c.
decide between
the scenarios

H₀: $\frac{P(B|A)}{P(B|A^c)} = \frac{.10}{.03} = 6$ technically, a Bayes Factor in favor of A

In general A_1, \dots, A_n mut. excl. coll. exh.

$$\Rightarrow P(B) = P(B, A_1) + P(B, A_2) + \dots + P(B, A_n)$$

$$P(A_2|B) = \frac{P(B|A_2) P(A_2)}{\sum_{i=1}^n P(B|A_i) P(A_i)}$$

Bayes Thm.

Likense given jmf

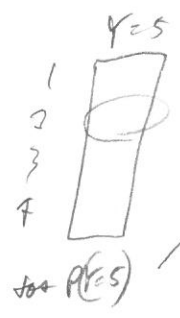
$$P(x, y)$$

$\text{Supp}(X)$

	$\text{Supp}(Y)$
1	1 2 3 4 5 6
2	
3	
4	

just like B, A

$$P(X=2|Y=5) = \frac{P(X=2, Y=5)}{P(Y=5)}$$



disjoint

$$P(Y=5) = P(X=1, Y=5) + P(X=2, Y=5) + P(X=3, Y=5) + P(X=4, Y=5)$$

Law of total prob

difficult! → a prob
 ↓ so we can allow B, A to be variable events

PMF

$$P(X|Y) = \frac{P(X, Y)}{P(Y)} = \frac{P(X) P(Y)}{P(Y)}$$

$$P(Y) = \sum_{x \in \text{supp}(X)} P(X, Y)$$

marginalization

Now $X|Y$ is given

$$f(x|y) = \frac{f(x, y)}{f(y)} = \frac{f(x) f(y)}{f(y)}$$

$$f(y) = \int_{x \in \text{supp}(X)} f(x, y) dx$$

r.v.
 $\text{supp}(X|Y) = \text{supp}(X)$?
 $E(X|Y) = E(X)$?
 $\text{Var}(X|Y) = \text{Var}(X)$?

the parameter for θ

Consider

$$P(\theta|X) = \frac{P(X|\theta) P(\theta)}{P(X)}$$

the data likelihood prior for θ prior for X "marginal likelihood"

θ under a certain model

makes no sense since θ is fact! either 0 or 1

What is $P(\theta; X) = \frac{P(X; \theta) P(\theta)}{P(X)}$

this also makes no sense since you cannot calculate $P(X)$ without knowing what θ is

incoherence in frequentist statistics

for the $P(\theta; X) = P(X; \theta)$

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We have to start at this again ~~the~~ whole sentence

Starting at this formula:

probability inference about θ

effect cause

$$P(\theta|x) = \frac{P(x|\theta) P(\theta)}{P(x)} \leftarrow \text{prior on } \theta$$

hence prob question

What does this mean? Real coin flips: $\langle 0, 1, 1 \rangle$
X

$P(\text{cause}|\text{effect})$
vs.
 $P(\text{effect}|\text{cause})$

$$P(X|\theta = 0.75) = .25 \cdot .75 \cdot .75 = .141$$

$$P(X|\theta = 0.5) = .5 \cdot .5 \cdot .5 = .125$$

generally we calculate these...

$$P(\theta = 0.25) ?? \quad P(\theta = 0.75) ??$$

this is our idea of what θ is a priori
to coinflipping coin flips

Imagine for now that we have only two theories for the prior θ
Ridiculous, yes, but go with it.

Also imagine $\underbrace{P(\theta = 0.25)}_{\text{model 1}} = \underbrace{P(\theta = 0.75)}_{\text{model 2}} = \frac{1}{2}$

Principle of indifference.

All prior hypotheses about θ are equally likely.

$$\Omega = \mathcal{X} \times \mathbb{H}_0 \rightarrow \text{supp}[\mathcal{X}]^h \rightarrow \mathbb{H}_0 \subseteq \mathbb{H}$$

[7]

3
= $\text{supp}(\mathcal{X})$
sample space
↓

$\langle 0,0,0 \rangle$			$\langle 0,0,0 \rangle$	$0,1,0$	$0,0,1$	110	011	101	111	$\theta = 0.25$
$0,1,0$	$0,1,0$	$0,1,0$	110	101	011	$\langle 1,1,1 \rangle$				$\theta = 0.75$

$$\mathcal{X} = \{ \langle 0,0,0 \rangle, \langle 0,0,1 \rangle, \langle 0,1,0 \rangle, \langle 1,0,0 \rangle, \langle 0,1,1 \rangle, \langle 1,0,1 \rangle, \langle 1,1,0 \rangle, \langle 1,1,1 \rangle \}$$

$$P(\theta=0.25) P(\langle 0,0,0 \rangle | \theta=0.25) = .75^3 = .421 \Rightarrow P(\langle 0,0,0 \rangle \& \theta=0.25)$$

$$\text{is } \langle 0,0,0 \rangle, \theta=0.25 \cap \langle 0,0,0 \rangle, \theta=0.75? \text{ YES,}$$

different pieces of the outcome!

$$P(\langle 0,0,1 \rangle | \theta=0.25) = .75^2 \cdot .25 = .141$$

$$= P(\langle 0,1,0 \rangle |) =$$

$$P(\langle 1,0,1 \rangle |) =$$

$$P(\langle 0,1,1 \rangle | \theta=0.25) = .75 \cdot .25^2 = .047$$

$$P(\langle 1,1,1 \rangle |) = .25^3 = .016$$

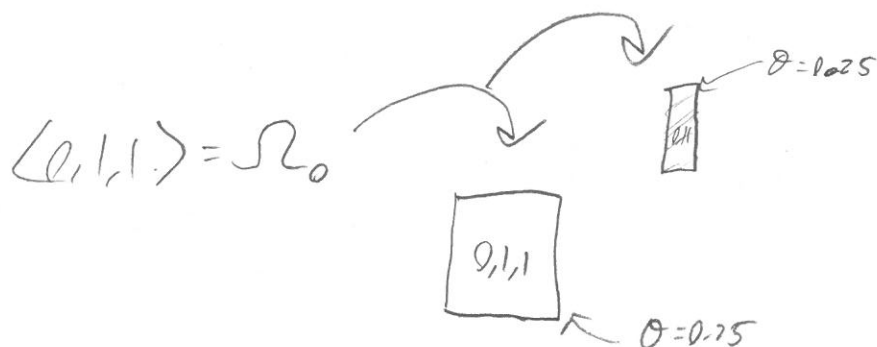
$$P(\langle 0,0,0 \rangle | \theta=0.75) = .25^3 = .016$$

$$P(\langle 0,0,0 \rangle |) = .75 \cdot .25^2 = .047$$

Now let's ask the inverse question. I've seen the effect, who's the cause. $X = \langle 0,1,1 \rangle$. I live in that world now!

What is prob $\theta=0.25$?

$P(0,1,1 | \theta = 0.25)$ ~~Not this!!!!~~ \rightarrow $P(\theta = 0.25 | 0,1,1)$
 But this



obviously...

$$\begin{aligned}
 & \frac{P(0,1,1, \theta = 0.25)}{P(0,1,1, \theta = 0.25) + P(0,1,1, \theta = 0.75)} = \frac{P(0,1,1 | \theta = 0.25) P(\theta = 0.25)}{P(0,1,1 | \theta = 0.25) P(\theta = 0.25) + P(0,1,1 | \theta = 0.75) P(\theta = 0.75)}
 \end{aligned}$$

$$P(0,1,1)$$

$$= P(\theta = 0.25 | 0,1,1) = 1 - 0.25$$

$$\Rightarrow P(\theta = 0.75 | 0,1,1) = 0.75$$

$$\begin{aligned}
 & = \frac{.047 \cdot .5}{.047 \cdot .5 + .141 \cdot .5} = \boxed{.25} \\
 & \rightarrow .094
 \end{aligned}$$

What is $P(0,1,1)$. This is the prior on the data.

It's about just seeing 0,1,1 without regard to θ .
 $\Rightarrow \Theta = \{\theta_1, \dots, \theta_n\}$, n could be ∞

$$P(X) = \sum_{i=1}^n P(X, \theta_i) \quad \text{or} = \int P(X, \theta) d\theta \quad (\text{H})$$

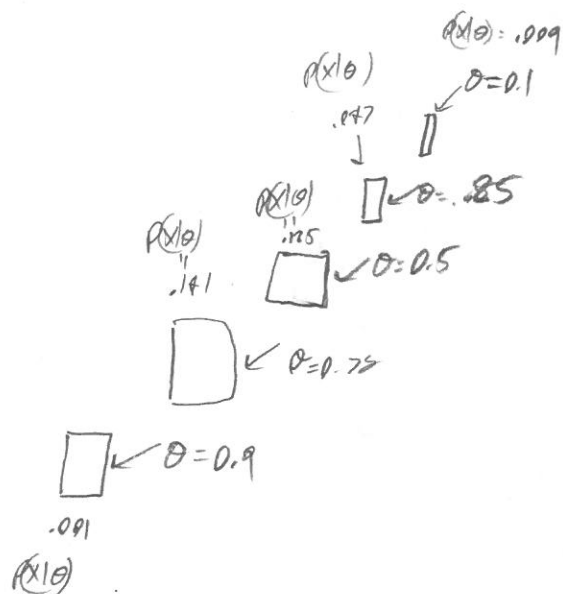
$$= \sum_{i=1}^n P(X | \theta_i) P(\theta_i) = \int P(X | \theta) P(\theta) d\theta \quad (\text{H})$$

Is X, θ independent? $P(X, \theta) = P(X) P(\theta)$? NO WAY

$$P(0,1,1, \theta = 0.25) = P(X) P(\theta = 0.25)$$

Imagine $\Theta = \{0.1, 0.25, 0.5, 0.75, 0.9\}$ s.t. $P(\theta_i) = \frac{1}{5}$ $\forall i$

observe $X = (0, 1, 1)$. want to ask $P(\Theta | X)$ again.



Note that
$$P(\theta | X) = \frac{P(X|\theta) P(\theta)}{P(X)} \propto P(X|\theta) P(\theta)$$

↑
not a $f(\theta)$

hence is a constant

$$\forall \theta \in \Theta_0$$

so if you just want to compare

$$\frac{P(\theta=0.75 | X)}{P(\theta=0.9 | X)} = \frac{P(X|\theta=0.75) P(\theta=0.75)}{P(X|\theta=0.9) P(\theta=0.9)}$$

I know this is a lot... but we're going to do it 100 times.
What is the "best" estimate of θ ?

Given the data, what is the most likely θ ?

$$\hat{\theta}_{\text{MAP}} = \underset{\theta \in \Theta_0}{\operatorname{argmax}} \{P(\theta|x)\} = \underset{\theta \in \Theta_0}{\operatorname{argmax}} \left\{ \frac{P(x|\theta) P(\theta)}{P(x)} \right\} = \underset{\theta \in \Theta_0}{\operatorname{argmax}} \{P(x|\theta) P(\theta)\}$$

maximum a posteriori (MAP)

if $P(\theta)$ is the same for all $\theta \in \Theta_0$ = $\underset{\theta \in \Theta_0}{\operatorname{argmax}} \{P(x|\theta)\}$

if $\Theta_0 = \Theta$ = $\hat{\theta}_{\text{MLE}}$ (only)

In our case, $\hat{\theta}_{\text{MAP}} = 0.75$

$\hat{\theta}_{\text{MLE}} = 0.66$ why? $\Theta_0 \neq \Theta = (0,1)$

Our choice of Θ_0 was weak. Unless you have good reason, $\Theta_0 = \Theta$.

Week 10 #1: Your prior could be bad (fruit chocolate).

$$P(\theta=0.25)=0.5 \longrightarrow P(\theta=0.25|(0,1,1))=0.25$$

"Bayesian Conditioning"

Data comes in ... update prior beliefs.

Let's see how this works more closely... One at a time...

$X_1=0, X_2=1, X_3=1$ s.t. $X_1, X_2, X_3 \stackrel{\text{iid}}{\sim} \text{Bern}(\theta)$ little bit wrong