

Lecture 14 3/30/16

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$$\propto P(X|\theta, \sigma^2) P(\theta) P(\sigma^2)$$

if $P(\theta) \propto 1$, $P(\sigma^2) \propto \frac{1}{\sigma^2}$

Recall $P(\theta, \sigma^2 | X) = P(\theta | X, \sigma^2) P(\sigma^2 | X) \ll$

$$N(\bar{x}, \frac{\sigma^2}{n}) \text{ Invgamma}(\frac{n}{2}, \frac{(n-1)s^2}{2})$$

$\propto \sum_{i=1}^n (x_i - \bar{x})^2$
 unbiased estimator for variance "Sample var"
 N-Invgamma dist

$$P(\sigma^2 | X) = \int_{\text{supp}(\theta)} P(\theta, \sigma^2 | X) d\theta = \int_{\mathbb{R}} (\sigma^2)^{-\frac{n}{2}+1} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n(\bar{x}-\theta)^2)} d\theta$$

why

$$= (\sigma^2)^{-\frac{n}{2}+1} e^{-\frac{1}{2\sigma^2}(n-1)s^2} \int_{\mathbb{R}} e^{-\frac{n}{2\sigma^2}(\bar{x}-\theta)^2} d\theta$$

$\propto 1$
 Kernel of Norm

$$f(x|z) = \int_y f(x,y|z) dy = \int_y f(x|y,z) f(y|z) dy$$

"Marginalization"

$$\propto \text{Invgamma}(\frac{n}{2}, \frac{(n-1)s^2}{2})$$

if this is $P(y)$
 then we get $\Rightarrow \text{AVG!}$

This is not proportional to 1, it is proportional to $\text{sig}^2(1/2)$ because the constant of integration is a $f(\text{sig}^2)$. The alpha of the resulting inverse gamma is then $(n-1)/2$ but otherwise this is correct.

Per $z_0, z_1, \dots, z_n \stackrel{\text{iid}}{\sim} N(0,1)$

$$\sum_{i=1}^n z_i^2 \sim \chi^2_n := \frac{2^{-n/2}}{\Gamma(n/2)} \theta^{\frac{n}{2}-1} e^{-\frac{\theta}{2}} = \text{Gamma}(\frac{n}{2}, \frac{1}{2})$$

free variable

$$\frac{z_0}{\sqrt{\sum_{i=1}^n z_i^2}} \sim T_2 := \frac{\Gamma(\frac{n+1}{2})}{\Gamma(\frac{n}{2}) \sqrt{\pi}} (1 + \frac{1}{2}\theta^2)^{-\frac{n+1}{2}} \rightarrow \bar{X} \sim T_{n-1}(n, \sqrt{\frac{s^2}{n}}) := \text{PDF}$$

Student's central and scaled T

Classic randoms (Mark 242)

$$\frac{n-1}{\sigma^2} s^2 \sim \chi^2_{n-1}, \quad \frac{\bar{X} - \mu}{\sqrt{\frac{s^2}{n}}} \sim T_{n-1}$$

"Standard"

T_1 d. Cauchy

$s^2 \rightarrow \sigma^2$ so $T_{n-1} \rightarrow Z$ None with same tails

5 lectures to prove these

two dist's and two random

$$\Rightarrow P(\theta, \sigma^2 | x) = \underbrace{N(\bar{x}, \frac{\sigma^2}{n})}_{P(\theta | \sigma^2, x)} \underbrace{\text{InvGamma}(\frac{n}{2}, \frac{(n-1)S^2}{2})}_{P(\sigma^2 | x)}$$

ditto over here from page 1

Process called N-InverseGamma distr.

Expectation? Variance? Not covered...

How to sample?

Step 1: Draw σ^2 from $\text{InvGamma}(\frac{n}{2}, \frac{(n-1)S^2}{2})$

Step 2: Draw θ from $N(\bar{x}, \frac{\sigma^2}{n})$ using σ^2 from Step 1

Repeat many, many times

Recall, $P(\theta, \sigma^2) \propto \frac{1}{\sigma^2}$ Uniform...

How to make it informative? (General model can incorporate prior info)

Since we had $V(0,1)$ but we use $\text{Beta}(\alpha, \beta)$

or $\text{Gamma}(1,0)$ but we use $\text{Gamma}(\alpha, \beta)$

or $N(0, \infty)$ but we use $N(\mu_0, \tau^2)$

before we assumed $P(\sigma^2 | \theta) = P(\sigma^2)$

$\theta | x, \sigma^2 \sim N(\frac{\bar{x}n + m\mu_0}{n+m}, \frac{\sigma^2}{n+m})$
 \uparrow
 if $\theta | \sigma^2 \sim N(\mu_0, \frac{\sigma^2}{m})$ what does prior mean? in regressions
 $\tau^2 = f(\sigma^2)$
 why??

Look at the above

It seems prior of θ should be dependent on prior of σ^2
 τ^2 is a form of σ^2 now directly

$$\Rightarrow \theta | \sigma^2 \sim N(\mu_0, \frac{\sigma^2}{m}), \sigma^2 \sim \text{InvGamma}(\frac{\nu_0}{2}, \frac{\nu_0 \sigma_0^2}{2})$$

$\Rightarrow \theta, \sigma^2 \sim \text{NInvGamma}$... conjugate for Normal with params θ, σ^2

For HW...

$$P(\theta, \sigma^2 | x) = N\left(\frac{\frac{n}{n+m} \bar{x} + \frac{m}{n+m} \mu_0, \frac{\sigma^2}{n+m}\right)$$

$$\text{e. Inubama} \left(\frac{n_0 + n}{2}, \frac{n_0 \sigma_0^2 + (n-1) S^2 + \frac{n n_0}{n+m} (\bar{x} - \mu_0)^2}{2} \right)$$

To draw samples

Step 1: σ^2 from J.G.

Step 2: θ from normal

↳ Step 1: σ^2 from

Step 1

HW for everyone — prove by conjugate...

Bayes

⇒ Multiparameter models are really hard!!

θ, σ^2 unknown but only care about θ ...

WANT:

$P(\theta | x)$ How does this differ from $P(\theta | x, \sigma^2)$
which we did before??

we don't know this

and θ , but σ^2 "goes in the way". σ^2 is called a "nuisance parameter". What to do?

Recall $f(x) = \int_{\text{supp}(f)} f(x|y) dy$, $f(x|z) = \int_{\text{supp}(f)} f(x|y|z) dy$

⇒ $P(\theta | x) = \int_{\text{supp}[\sigma^2]} P(\theta, \sigma^2 | x) d\sigma^2$ marginal out σ^2 ... Avg. your idea about θ over all possible σ^2 's.

Use non-refuse prior function first...

LA

$$p(\theta|x) \propto \int_0^\infty (\sigma^2)^{-\frac{n}{2}-1} e^{-\frac{1}{2\sigma^2}((n-1)s^2 + n(\bar{x}-\theta)^2)} d\sigma^2$$

let $A := (n-1)s^2 + n(\bar{x}-\theta)^2$ ✓ not a function of σ^2

$$t := \frac{A}{2\sigma^2} \Rightarrow \sigma^2 = \frac{A}{2t} \Rightarrow \frac{d\sigma^2}{dt} = -\frac{A}{2t^2}$$

$$= \int_0^\infty \left(\frac{A}{2t}\right)^{-\frac{n}{2}-1} e^{-t} \left(-\frac{A}{2t^2} dt\right)$$

$\Rightarrow d\sigma^2 = -\frac{A}{2t^2} dt$
 $\sigma^2 \in (0, \infty)$
 $\Rightarrow t \in (0, \infty)$

$$= - \int_0^\infty \frac{A^{-\frac{n}{2}-1}}{2^{-\frac{n}{2}-1} t^{-\frac{n}{2}+1+2}} e^{-t} dt$$

$t^{-(-\frac{n}{2}+1)} = t^{\frac{n}{2}-1} = t^{\frac{n-2}{2}}$

$$\propto A^{-\frac{n}{2}} \int_0^\infty t^{\frac{n-2}{2}} e^{-t} dt$$

Kernel of gamma

\Rightarrow constant

$$\propto A^{-\frac{n}{2}}$$

$$= \left((n-1)s^2 + n(\bar{x}-\theta)^2\right)^{-\frac{n}{2}}$$

$$\propto \left(1 + \frac{n(\bar{x}-\theta)^2}{(n-1)s^2}\right)^{-\frac{n}{2}}$$

$$\propto T(n-1, \bar{x}, \frac{s^2}{n})$$

the T is the

overdispersed Normal

when σ^2 is allowed to vary...

high-central scale

Student t distr.

$$X \sim T(\nu, \mu, \sigma^2)$$

$$:= \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\frac{\nu}{2}) \sqrt{2\pi\sigma^2}} \left(1 + \frac{1}{\nu} \left(\frac{x-\mu}{\sigma}\right)^2\right)^{-\frac{\nu+1}{2}}$$

$$\propto \left(1 + \frac{1}{\nu} \left(\frac{x-\mu}{\sigma}\right)^2\right)^{-\frac{\nu+1}{2}}$$

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Traditionally, the non-central scaled data is centered and scaled (standardized) to yield:

$$\Rightarrow X \sim T(\nu, \mu, \sigma^2) \Rightarrow \frac{X - \mu}{\sigma} \sim T_\nu$$

↑
Student's T distr.

Note: $T_\nu \stackrel{d}{=} \frac{Z}{\sqrt{\chi_\nu^2}}$ when $Z \sim N(0,1)$

$$\chi_\nu^2 \text{ or } \chi_\nu^2 := \sum_{i=1}^{\nu} Q_i^2 \quad Q_i \sim N(0,1)$$

this is the test statistic in classic statistics for Hypothesis tests of μ when σ^2 unknown

$$\Rightarrow P\left(\frac{\theta - \bar{x}}{\frac{s}{\sqrt{n}}} \mid X\right) = T_{n-1}$$

By a similar exercise... If $P(\theta | \sigma^2) = N(\mu_0, \frac{\sigma_0^2}{n})$
and $P(\sigma^2) = \text{Invgamma}(\frac{\nu_0}{2}, \frac{\nu_0 \sigma_0^2}{2})$

$$\Rightarrow P\left(\frac{\theta - \mu_1}{\frac{\sigma_1}{\sqrt{m}}} \mid X\right) = T_{\nu_0 + n}$$

where $\mu_1 = \frac{n}{n+m} \bar{x} + \frac{m}{n+m} \mu_0$

$$\sigma_1 = \sqrt{\frac{\nu_0 \sigma_0^2}{n+m} + \frac{m}{(n+m)^2} (n-1) s^2}$$

} I think...
(HW)

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$$= \int_{\mathbb{R}} P(\theta, \sigma^2 | x) d\theta$$

$$P(\sigma^2 | x) = \text{Inv Gamma} \left(\frac{n_0 + n}{2}, \frac{1}{2} \left(n_0 \sigma_0^2 + (n-1)S^2 + \frac{nm}{n+m} (\bar{x} - \mu_0)^2 \right) \right)$$

\Rightarrow Mcbray ... Bayesian Models very complicated!

This is only $N(\theta, \sigma^2)$ with conj. prior!!!

What about bigger models ?? WORSE!!!

Need better ways....

Under conj. priors...

$$P(x^* | x) = \int \int P(x^* | \theta, \sigma^2) P(\theta, \sigma^2 | x) d\sigma^2 d\theta$$

$$= \int \int \underbrace{P(x^* | \theta, \sigma^2)}_{N(\theta, \sigma^2)} \cdot \underbrace{P(\theta | x, \sigma^2)}_{N\left(\frac{n}{n+m} \bar{x} + \frac{m}{n+m} \mu_0, \frac{\sigma^2}{n+m}\right)} \cdot \underbrace{P(\sigma^2 | x)}_{\text{Inv Gamma}\left(\frac{n_0+n}{2}, \frac{n_0 \sigma_0^2 + (n-1)S^2 + \frac{nm}{n+m} (\bar{x} - \mu_0)^2}{2}\right)} d\sigma^2 d\theta$$

$$= T\left(n-1, \bar{x}, S \sqrt{\frac{n}{n-1}}\right) \text{ super Herculean effort!} \quad (\text{EC})$$

OR: How to draw samples?

Step 1: Draw σ^2 from Inv Gamma.

Step 2: Draw θ from Normal using σ^2 from step 1.

Step 3: Draw x^* from Normal using θ, σ^2 from steps 1, 2.

DO EX NOW