

Math 380
Lecture 4 2/10/16

" $X_1, X_2 \sim \text{iid Bern}(\theta)$ " correct? [1]

one at
a time
 $\langle 0, 1, 1 \rangle$

$$P(\theta=0.25 | X_1=0) = \frac{P(X_1=0 | \theta=0.25) P(\theta=0.25)}{P(X_1=0)}$$

$$P(X_1=0) = P(X_1=0 | \theta=0.25) P(\theta=0.25) + P(X_1=0 | \theta=0.75) P(\theta=0.75)$$

$$= \frac{.75 \cdot .5}{.75 \cdot .5 + .25 \cdot .5} = .75$$

$P(\theta=0.25) = 0.5 \longrightarrow P(\theta=0.25 | X_1) = 0.75$

$\theta=0.25$ is looking better!

Now pretend this is the prior...

$$P(\theta=0.25 | X_2=1) = \frac{P(X_2=1 | \theta=0.25) P(\theta=0.25)}{P(X_2=1)}$$

$$P(X_2=1) = P(X_2=1 | \theta=0.25) P(\theta=0.25) + P(X_2=1 | \theta=0.75) P(\theta=0.75)$$

$$= \frac{.25 \cdot .75}{.25 \cdot .75 + .75 \cdot .25} = .5$$

$P(\theta=0.25) = 0.5 \longrightarrow P(\theta=0.25 | X_1, X_2) = .5$

Our information didn't tell us anything!

Now pretend this is the prior

$$P(\theta = 0.25 | X_3 = \phi) = \frac{P(X_3 = \phi | \theta = 0.25) P(\theta = 0.25)}{P(X_3 = \phi)} = .25$$

$$P(X_3 = \phi) = P(X_3 = \phi | \theta = .5) P(\theta = .5) + P(X_3 = \phi | \theta = .75) P(\theta = .75) = .75 \cdot .5 + .25 \cdot .5$$

ALSO recall...

$$P(\theta = 0.25 | X_1 = 0, X_2 = 1, X_3 = 1) = .25$$

Is it possible that each piece of data comes in, you update your prior to the new posterior?? Etc.?

Let's see...

$$P(\theta | X_1, \dots, X_n) = \frac{P(X_1, \dots, X_n | \theta) P(\theta)}{P(X_1, \dots, X_n)}$$

$$= \frac{P(X_n | \theta) \dots P(X_2 | \theta) P(X_1 | \theta) P(\theta)}{P(X_n, \dots, X_2 | X_1) P(X_1)}$$

$$= \frac{P(X_n | \theta) \dots P(X_2 | \theta) P(X_1 | \theta) P(\theta)}{P(X_n, \dots, X_2 | X_1) P(X_1)}$$

$$= \frac{P(x_1|\theta) \dots P(x_2|\theta)}{P(x_1, \dots, x_2|x_1)} P(\theta|x_1)$$

\uparrow
 prior of x_1, \dots, x_2 now seeing x_1

\nwarrow this now takes the place of the prior

do one more...

$$\frac{P(x_3, \dots, x_n|\theta)}{P(x_3, \dots, x_n|x_1, x_2)} \frac{P(x_2|\theta)}{P(x_2)} P(\theta|x_1)$$

$$\frac{P(x_2|\theta)}{P(x_2)} \frac{P(x_1|\theta) P(\theta)}{P(x_1)}$$

$$\frac{P(x_1, x_2|\theta) P(\theta)}{P(x_1, x_2)}$$

$$P(\theta|x_1, x_2)$$

etc...

Explain $X \perp \theta \Rightarrow$ posterior = prior
 explain why this is conceptually
 explain why $X \not\perp \theta$ given F

how the prior
 when considering x_3, x_4, \dots, x_n

Another concept... we have x_1, x_2, x_3 . What does x_4 look like?

How would you do this before?

$$\hat{\theta}_{MLE} = \frac{0+1+1}{3} = 0.66 \Rightarrow x_4 \sim \text{Bern}(0.66)$$

you wouldn't guess yesterday

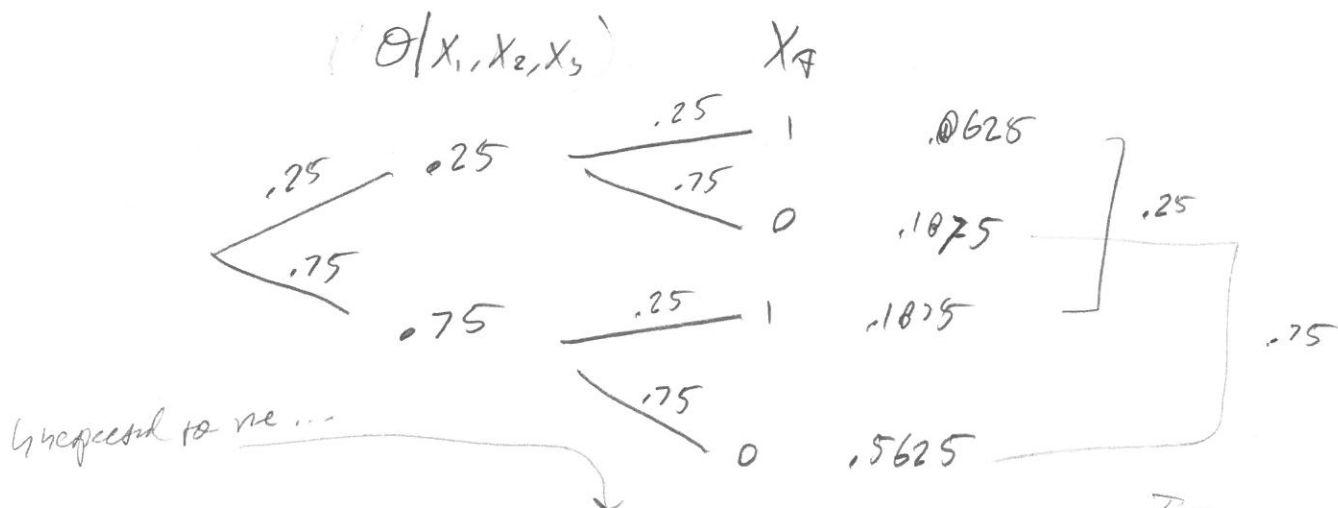
Or you can say $H_0: \theta = 0.5$ $H_1: \theta \neq 0.5$ $X = 0, 1, 1 \Rightarrow$ Reject H_0

$\Rightarrow X_4 \sim \text{Bern}(0.5)$

\uparrow
Yours guess H_0 !

In Bayesian Palace, we seek $P(X_4 | X_1, X_2, X_3)$

AKA the "prior predictive distribution". The seen X_1, X_2, X_3 , what does X_4 look like?



$X_4 | X_1, X_2, X_3 \sim \text{Bern}(0.25)$

I can condition on A, B where I must eliminate

\downarrow

$\sum_{X \in \mathcal{X}(Y)} P(X, Y) | A, B$

$= \sum P(X|Y) P(Y)$

How to find this?

Recall

$$P(Y) = \sum_{X \in \mathcal{X}(Y)} P(X, Y)$$

$$P(X_4 | X_1, X_2, X_3) = \sum_{\theta \in \Theta_0} P(X_4, \theta | X_1, X_2, X_3)$$

$$= \sum P(X_4 | \theta, X_1, X_2, X_3) P(\theta | X_1, X_2, X_3) = \sum P(X_4 | \theta) P(\theta | X_1, X_2, X_3)$$

\parallel
 $P(X_4 | \theta)$

Why?

generally $P(X^* | X) = \sum_{\theta \in \Theta_0} P(X^* | \theta) P(\theta | X) = \int P(X^* | \theta) P(\theta | X) d\theta$

\parallel
 Θ_0

$$\begin{aligned}
 p(Y_4 | \theta, X_1, X_2, X_3) &= \frac{p(\theta, X_1, X_2, X_3, Y_4)}{p(\theta, X_1, X_2, X_3)} \\
 &= \frac{p(X_1, X_2, X_3, Y_4 | \theta) p(\theta)}{p(X_1, X_2, X_3 | \theta) p(\theta)} \\
 &= \frac{p(X_1, X_2, X_3 | \theta) p(Y_4 | \theta)}{p(X_1, X_2, X_3 | \theta)}
 \end{aligned}$$

Why?? Once you know θ , the previous info is not relevant... you don't get anything out of it.

Does $p(X_1, X_2, X_3) \stackrel{?}{=} p(X_4)$??? Am I the idiot? $p(X|Y) = p(X)$ if $X \perp Y$?

Probably, $X_1, X_2, X_3 \stackrel{iid}{\sim} \text{Bern}(\theta)$? In frequentist world this is true since the underlying assumption is θ is fixed

Now X_1, X_2, X_3 Not iid $\text{Bern}(\theta)$

Since if I know X_1 , my θ has changed!

$$p(\theta) \rightarrow p(\theta | X_1)$$

So $X_1 | \theta, X_2 | \theta, X_3 | \theta \stackrel{iid}{\sim} \text{Bern}(\theta)$

You need the cond'l r.v.'s now!

Let's return to the beginning

$$\theta \sim \begin{cases} 0.25 & \text{up } \frac{1}{2} \\ 0.75 & \text{up } \frac{1}{2} \end{cases} \Rightarrow \theta | X \sim \begin{cases} 0.25 & \text{up } 0.25 \\ 0.75 & \text{up } 0.25 \end{cases}$$

with I would be chosen based on #1's
↓

We said our $\hat{\theta}_{\text{Bayes}} = \hat{\theta}_{\text{MAP}} = \arg\max_{\theta} \{P(\theta|X)\} = 0.75$

What about $\hat{\theta}_{\text{Bayes}} = E[\theta|X] = \sum_{\theta \in \Theta} \theta P(\theta|X)$
 AKA the "posterior mean"
 -OR-

$$= \int \theta f(\theta|X) d\theta$$

$$\theta \in \Theta_0$$

14th Jan
 Calc $= 0.25 \cdot 0.25 + 0.75 \cdot 0.75 = .625$

Note $.625 \notin \Theta_0$

Without getting carried away, $E[\theta|X] = \hat{\theta}_{\text{MMSE}}$ Not in scope of course
 \uparrow
 min mean sq err $\rightarrow \ell(\hat{\theta}, \theta)$

$$= E_{\theta}[(\hat{\theta}_{\text{MMSE}} - \theta)^2] = \int (\hat{\theta} - \theta)^2 P(\theta) d\theta$$

$$\theta \in \Theta_0$$

 Sometimes this is called a "Bayes Estimator"

So... this is a good pt. estimator if you care about sq. error loss

$\hat{\theta}_{\text{MAP}}$ is not a Bayes Estimator except for a weird loss function and still isn't unless θ is discrete. Most priors are not $p, \mu, \sigma^2, \lambda$, etc

So... .625 $\notin \Theta_0$. Maybe it's time to use $\Theta_0 = \Theta = (0,1)$

How about $\theta \sim U(0,1)$. This is non continuous... every value in $(0,1)$ is represented and given equal prob. It's principle of indifference on steroids!

$$P(\theta | x_1, x_2, x_3) = \frac{P(x_1, x_2, x_3 | \theta) P(\theta)}{P(x_1, x_2, x_3)} = 1 \quad \text{since } f(x; a=0, b=1) = 1 \mathbb{1}_{x \in [0,1]}$$

Why $h_{1/2}$????

$$\int_{\Theta} P(x_1, x_2, x_3 | \theta) P(\theta) d\theta = \int_{\Theta} P(x_1=0 | \theta) P(x_2=1 | \theta) P(x_3=1 | \theta) P(\theta) d\theta$$

$$= \frac{P(x_1=0 | \theta) P(x_2=1 | \theta) P(x_3=1 | \theta)}{\int_0^1 P(\theta) d\theta}$$

What are these??
"params of prior" =
"hyperparameters"

$$= \frac{(1-\theta) \cdot \theta \cdot \theta}{\int_0^1 (1-\theta) \theta^2 d\theta} = \frac{(1-\theta) \theta^2}{\left[\frac{\theta^3}{3} - \frac{\theta^4}{4} \right]_0^1} = \frac{(1-\theta) \theta^2}{\frac{1}{12}} = 12(1-\theta) \theta^2$$

$$\hat{\theta}_{MAP} = \underset{\theta \in \Theta}{\operatorname{argmax}} \{ \cdot \} = \underset{\theta \in (0,1)}{\operatorname{argmax}} \{ \theta^2 - \theta^3 \} \quad 2\theta - 3\theta^2 = 0$$

$$\Rightarrow 2 - 3\theta = 0$$

$$\Rightarrow \theta = \frac{2}{3} = \hat{\theta}_{MLE} \text{ cool!!}$$

$$\hat{\theta}_{MSE} = E[\theta | x] = \int_0^1 \theta (12(1-\theta) \theta^2) d\theta = 12 \int_0^1 (\theta^3 - \theta^4) d\theta = 12 \left[\frac{\theta^4}{4} - \frac{\theta^5}{5} \right]_0^1 = 12 \left(\frac{1}{4} - \frac{1}{5} \right) = \frac{12}{20} = 0.6$$

Or more Bayes Estimator... if $l(\theta, \hat{\theta}) = a|\theta - \hat{\theta}|$ is absolute loss,

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$$\Rightarrow \hat{\theta}_{MAE} = \text{Median}[\theta|x]$$

$$= \left\{ \theta : \int_0^{\theta} p(\theta|x) d\theta = 0.5 \right\}$$

$$= \int_0^y 12(1-\theta)\theta^2 d\theta = 0.5$$

$$\left[\frac{\theta^3}{3} - \frac{\theta^4}{4} \right]_0^y = \frac{1}{24}$$

$$\Rightarrow \frac{y^3}{3} - \frac{y^4}{4} = \frac{1}{24}$$

$$\Rightarrow 4y^3 - 3y^4 = \frac{1}{2}$$

~~conjecture 4 results, 2, 1, 1, 1, 1~~

$$y \in \{0.614272, 1.24798, -0.26 + 0.30i, -0.26 + 0.30i\}$$

we have $\hat{\theta}_{MAP}, \hat{\theta}_{MSE}, \hat{\theta}_{MAE}$... these estimates...

$$P(X_4 | X_1, X_2, X_3) = \int_{\theta \in \Theta} P(X_4 | \theta) P(\theta | X_1, X_2, X_3) d\theta$$

$$= \int_0^1 \theta^{x_4} (1-\theta)^{1-x_4} 12(1-\theta)\theta^2 d\theta$$

$$= 12 \int_0^1 \theta^{x_4+2} (1-\theta)^{2-x_4} d\theta$$

$$\underline{\underline{B(x_4+3, 3-x_4)}}$$

$$\Gamma(x) = \int_0^{\infty} x^{t-1} e^{-x} dx$$

$= (x-1)!$ if $x \in \mathbb{N}$
 example!
 to all \mathbb{R}

Beta Function

$$B(x, y) = \frac{\Gamma(x)\Gamma(y)}{\Gamma(x+y)}$$