

Math 621 Lec 4 9/7/16

$$\vec{X} \sim \text{Multinomial}(n, \vec{p}) := \binom{n}{x_1, x_2, \dots, x_k} p^{x_1} \dots p^{x_k}$$

Sometimes you see a  $k$  to indicate  $\dim(\vec{X}) = k$ ,  $\dim(\vec{p}) = ?$   $k$

No uniform function necessary due to multinomial notation. If there needed to be...

$$\mathbb{1}_{\sum x_i = n} = \prod_{i=1}^k \mathbb{1}_{x_i \in \mathbb{N}_0} \quad \text{ugh!}$$

$$\text{supp}[\vec{X}] = \{ \vec{x} : \vec{x} \in \mathbb{N}_0^k \text{ and } \vec{x} \cdot \vec{1} = n \} \quad \text{Prin space} \quad \mathcal{P} = \{ \vec{p} : \vec{p} \in (0,1)^k \text{ and } \vec{1} \cdot \vec{p} = 1 \}$$

In a drawer of 10 fruits such that  $p_1 = \frac{1}{4}$ ,  $p_2 = \frac{1}{8}$ ,  $p_3 = \frac{5}{8}$

$$\begin{aligned} P(\text{d getting 3 Apples, 2 bananas, 5 cantaloupes}) &= \binom{10}{3, 2, 5} \left(\frac{1}{4}\right)^3 \left(\frac{1}{8}\right)^2 \left(\frac{5}{8}\right)^5 \\ &= P\left(\vec{X} = \begin{pmatrix} 3 \\ 2 \\ 5 \end{pmatrix}\right) = \end{aligned}$$

e.g. If  $K=2$ ,  $\vec{p} = \begin{bmatrix} p \\ 1-p \end{bmatrix} = \begin{bmatrix} p \\ 1-p \end{bmatrix}$  since  $p_1 + p_2 = 1$  <sup>must</sup>

What is the support?

$$\text{Multinom}(n, \begin{bmatrix} p_1 \\ p_2 \end{bmatrix}) = \text{Bin}(n, p_1)$$

$$\text{Supp}[X] = \mathbb{N}_0^K \text{ s.t. } \vec{X} \cdot \vec{1} = n \text{ i.e. } \sum X_k = n \text{ why?}$$

By model assumption there's only  $n$  trials  
so there must be only  $n$  successes from all  
categories.

$$(x_1, x_2) = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \sim \text{Multinom}(n, \begin{bmatrix} p \\ 1-p \end{bmatrix}) = \binom{n}{x_1, x_2} p^{x_1} (1-p)^{x_2}$$

joint mass function (pmf) tells you the prob of  <sup>$x_1 + x_2 = n$</sup>   $\neq \text{Bin}(n, p)$  since it must be a  $p(x_1, x_2)$ !  
more than one r.v. value.

Is  $x_1, x_2$  iid? What would this mean?

$$p(x_1, x_2) = p(x_1)p(x_2) \implies p(x_1 | x_2) = p(x_1) \text{ or } p(x_2 | x_1) = p(x_2)$$

Why? Bayes Rule

Conditional pmf

If  $x_i \notin \text{Supp}(X_i) \dots$   
Conditional pmf  
undefined...

$\forall x_1 \in \text{Supp}(X_1)$   
 $\forall x_2 \in \text{Supp}(X_2)$

$$p(x_1 | x_2) = \frac{p(x_1, x_2)}{p(x_2)} \stackrel{\text{if iid}}{=} \frac{p(x_1)p(x_2)}{p(x_2)} = p(x_1)$$

$$p(x_2 | x_1) = \dots = p(x_2)$$

Are  $X_1, X_2$  ind? Instantly... no!! If you know  $x_2$ ,  
then  $X_1 = n - x_2$ ! They are the ultimate dependence...

Proof...  $\rightarrow$  the PMF

$$p(X_1 | x_2) = \frac{p(X_1, X_2)}{p(x_2)}$$

we need  $p(x_2)$   $\leftarrow$  the marginal dist.



$$p(x_2) = \sum_{x_1 \in \text{supp}(X_1)} p(x_1, x_2) = \sum_{x_1=0}^n \frac{n!}{x_1! x_2!} p^{x_1} (1-p)^{x_2} \mathbb{1}_{x_1+x_2=n}$$

any possible way  
 $x_2$  can be realized... you  
add up every  $x_1$  of  $x_2, x_1$   
together

$$= \frac{n!}{x_2!} (1-p)^{x_2} \sum_{x_1=0}^n \frac{p^{x_1}}{x_1!} \mathbb{1}_{x_1=n-x_2}$$

all of those terms are zero

except when  $x_1 = n - x_2$

$$= \frac{n!}{x_2!} (1-p)^{x_2} \frac{p^{n-x_2}}{(n-x_2)!}$$

$$= \binom{n}{x_2} (1-p)^{x_2} p^{n-x_2}$$

Symmetric exercise!

$$\Rightarrow X_2 \sim \text{Binom}(n, 1-p) \quad \text{ALSO... } X_1 \sim \text{Binom}(n, p)$$

the marginal distr. in a Multinomial is a binomial...  
makes sense...

$$P(X_1|X_2) = \frac{\frac{n!}{x_1! x_2!} p^{x_1} (1-p)^{x_2}}{\frac{n!}{x_2! (n-x_2)!} (1-p)^{x_2} p^{n-x_2}} \mathbb{1}_{x_1+x_2=n}$$

$$= \frac{(n-x_2)!}{x_1!} p^{x_1+x_2-n} \mathbb{1}_{x_1+x_2=n}$$

↑

since  $x_1 = n - x_2$

$$\text{Jth } P(X_1 = n - x_2 | x_2) = \frac{(n-x_2)!}{(n-x_2)!} p^{x_1} = 1 \quad \text{Why?} \quad \neq \text{Bin}(n, p)$$

⇒ Not independent!

Let's do genl case...

$$X \sim \text{Multin}(n, \vec{p})$$

$$P(X_{-j} | X_j) = \frac{P(X_1, \dots, X_K)}{P(X_j)} = \frac{\text{Multin}(n, \vec{p})}{\text{Binom}(n, p_j)}$$

all counts  
except  
the  $j$ 'th

$$= \frac{\frac{n!}{x_1! \dots x_K!} p_1^{x_1} \dots p_K^{x_K}}{\frac{n!}{x_j! (n-x_j)!} p_j^{x_j} (1-p_j)^{n-x_j}}$$

$$= \frac{(n-x_j)!}{x_1! \dots x_{j-1}! x_{j+1}! \dots x_K!} \frac{p_1^{x_1} \dots p_{j-1}^{x_{j-1}} p_{j+1}^{x_{j+1}} \dots p_K^{x_K}}{(1-p_j)^{n-x_j}}$$

$$\text{let } n' := n - x_j$$

$$\text{Note: } \sum_{j=1}^K x_j = n \Rightarrow x_1 + \dots + x_{j-1} + x_j + x_{j+1} + \dots + x_K = n \Rightarrow x_1 + \dots + x_{j-1} + x_{j+1} + \dots + x_K = n - x_j = n'$$

$$= \binom{n'}{x_1, \dots, x_{j-1}, x_{j+1}, \dots, x_k} \frac{p_1^{x_1} \dots p_{j-1}^{x_{j-1}} p_{j+1}^{x_{j+1}} \dots p_k^{x_k}}{(1-p_j)^{n'}}$$

$$\text{let } p_1' = \frac{p_1}{1-p_j}, p_2' = \frac{p_2}{1-p_j}, \dots, p_k' = \frac{p_k}{1-p_j} \Rightarrow p' = \begin{bmatrix} p_1' \\ \vdots \\ p_k' \end{bmatrix}$$

$$= \binom{n'}{\dots} \frac{(p_1(1-p_j))^{x_1} \dots (p_{j-1}(1-p_{j-1}))^{x_{j-1}} (p_{j+1}(1-p_{j+1}))^{x_{j+1}} \dots (p_k(1-p_k))^{x_k}}{(1-p_j)^{n'}}$$

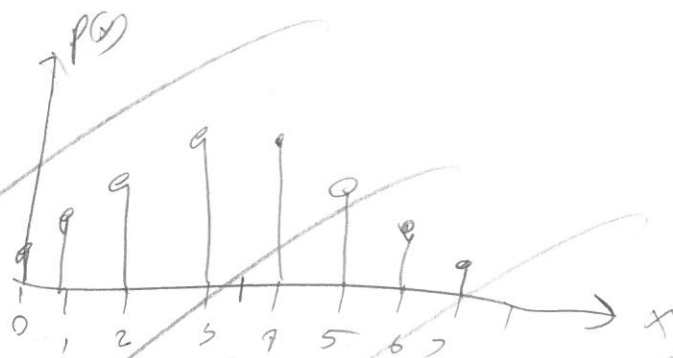
$$= \binom{n'}{\dots} \frac{(p_1')^{x_1} \dots (p_{j-1}')^{x_{j-1}} (p_{j+1}')^{x_{j+1}} \dots (p_k')^{x_k} \cancel{(1-p_j)^{x_1 + \dots + x_{j-1} + x_{j+1} + \dots + x_k}}}{(1-p_j)^{n'}}$$

$$= \text{Multinom}(n', p')$$

What is expectation? (mean?)

$$\mu := E(X) = \sum_{x \in \text{supp}(X)} x \cdot p(x)$$

$$X \sim \text{Binom}(7, \frac{1}{2})$$



$E(X)$

Balancept.

Long run average!

Recall from last lecture...  ~~$E[g(X)] = \sum_{x \in \mathcal{X}} g(x) p(x)$~~ ,  $E[g(X_1, \dots, X_n)] = \sum_{x_1 \in \mathcal{X}_1} \dots \sum_{x_n \in \mathcal{X}_n} g(x_1, \dots, x_n) p(x_1, \dots, x_n)$

$$E\left[\sum_{i=1}^n X_i\right] = \sum_{i=1}^n E[X_i] = n\mu \quad \text{if } X_i \stackrel{d}{=} X_j \forall i, j \quad (\text{proof in 2A1})$$

$$E\left[\prod_{i=1}^n X_i\right] = \prod_{i=1}^n E[X_i] \quad \text{if } X_1, \dots, X_n \stackrel{\text{ind}}{\sim}$$

$$= \mu^n \quad \text{if } X_1, \dots, X_n \stackrel{\text{iid}}{\sim}$$

$$E(g(X)) = g(E(X))$$

$$E(X+c) = E(X) + c$$

Proof:

$$E\left[\prod_{i=1}^n X_i\right] = \sum_{x_1} \dots \sum_{x_n} \underbrace{X_1 \cdot X_2 \cdot \dots \cdot X_n}_{g(x_1, \dots, x_n)} \underbrace{p(x_1, \dots, x_n)}_{p(x_1) \dots p(x_n)} = \sum_{x_1} X_1 p(x_1) \sum_{x_2} X_2 p(x_2) \dots \sum_{x_n} X_n p(x_n)$$

$$= E(X_1) E(X_2) \dots E(X_n)$$

$$\sigma^2 := \text{Var}(X) = E\left[\underbrace{(X-\mu)}_{g(X)}^2\right] = \sum_{x \in \mathcal{X}} \underbrace{g(x)}_{(x-\mu)^2} p(x) = \sum (x-\mu)^2 p(x)$$

$$= \sum x^2 p(x) + \sum (-2\mu x) p(x) + \sum \mu^2 p(x)$$

$$= E(X^2) - 2\mu^2 + \mu^2 = E(X^2) - \mu^2$$

Expected squared error  
difference from the expected

Rules

$$\text{Var}[X+a] = \text{Var}(X)$$

$$\text{Var}(cX) = c^2 \text{Var}(X)$$

$$\text{Var}(X_1 + X_2) = E\left[\left((X_1 + X_2) - (\mu_1 + \mu_2)\right)^2\right]$$

$$= E[X_1^2 + X_2^2 + \mu_1^2 + \mu_2^2 - 2\mu_1 X_1 - 2\mu_1 X_2 - 2\mu_2 X_1 - 2\mu_2 X_2 + 2X_1 X_2 + 2\mu_1 \mu_2]$$

$$= E[X_1^2] + E[X_2^2] + \cancel{\mu_1^2} + \cancel{\mu_2^2} - \cancel{2\mu_1^2} - \cancel{2\mu_1 \mu_2} - \cancel{2\mu_1 \mu_2} - \cancel{2\mu_2^2} + 2E[X_1 X_2] + \cancel{2\mu_1 \mu_2}$$

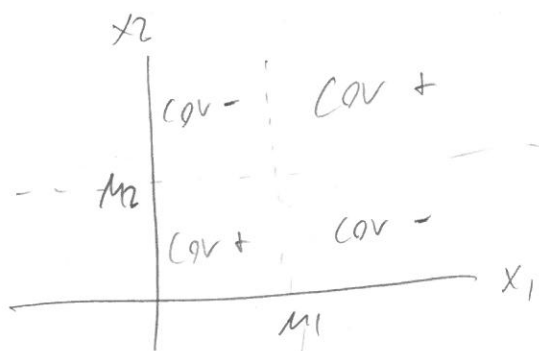
$\quad \quad \quad -\mu_1^2 \quad \quad \quad -\mu_2^2$

$$= \text{Var}(X_1) + \text{Var}(X_2) + 2(E[X_1 X_2] - \mu_1 \mu_2)$$

$$\text{Cov}(X_1, X_2) := E[X_1 X_2] - \mu_1 \mu_2$$

Same as  $E[(X_1 - \mu_1)(X_2 - \mu_2)] = E[X_1 X_2 - \mu_1 X_2 - \mu_2 X_1 + \mu_1 \mu_2]$

$$= E[X_1 X_2] - \mu_1 \mu_2 - \mu_1 \mu_2 + \mu_1 \mu_2 \quad \checkmark$$



$$\text{Corr}[X_1, X_2] = \frac{\text{Cov}[X_1, X_2]}{\text{SE}(X_1) \text{SE}(X_2)} \in [-1, 1]$$

↑  
a unitless, interpretable  
covariance

Rules  $\text{Cov}(X, X) = \text{Var}(X)$

$$\text{Cov}(a_1 X_1, a_2 X_2) = a_1 a_2 \text{Cov}(X_1, X_2)$$

$$\text{Cov}(X_1 + c_1, X_2 + c_2) = \text{Cov}(X_1, X_2)$$

$$\text{Cov}(X_2, X_1) = \text{Cov}(X_1, X_2) \quad \text{commutativity}$$

$$\text{Cov}(X+Y, Z) = \text{Cov}(X, Z) + \text{Cov}(Y, Z), \quad \text{Cov}\left(\sum_i X_i, \sum_j Y_j\right) = \sum_i \sum_j \text{Cov}(X_i, Y_j) \quad \text{Dirr?}$$

$$\begin{aligned} \text{Cov}(X+Y, Z) &= E[(X+Y - \mu_X - \mu_Y)(Z - \mu_Z)] \\ &= E[(X - \mu_X) + (Y - \mu_Y)](Z - \mu_Z) \\ &= E[(X - \mu_X)(Z - \mu_Z) + (Y - \mu_Y)(Z - \mu_Z)] \\ &= \text{Cov}(X, Z) + \text{Cov}(Y, Z) \end{aligned}$$

N.B.

$$\begin{aligned} \text{Var}(X_1 + X_2) &= \text{Var}(X_1) + \text{Var}(X_2) + 2\text{Cov}(X_1, X_2) \quad (\text{from before}) \\ &= \text{Cov}(X_1, X_1) + \text{Cov}(X_2, X_2) + \text{Cov}(X_1, X_2) + \text{Cov}(X_2, X_1) \\ &= \sum_{i=1}^2 \sum_{j=1}^2 \text{Cov}(X_i, X_j) \end{aligned}$$

$$\text{Var}(X_1 + X_2 + \dots + X_k) = \sum_{i=1}^k \sum_{j=1}^k \text{Cov}(X_i, X_j)$$

If  $\vec{X}$  is a vector of r.v. of dim  $k$ ...

$$\vec{\mu} := E[\vec{X}] = \begin{bmatrix} E[X_1] \\ E[X_2] \\ \vdots \\ E[X_k] \end{bmatrix}$$

$$\Sigma := \text{Var}(\vec{X}) = \begin{bmatrix} \text{Var}(X_1) & \text{Cov}(X_1, X_2) & \dots & \text{Cov}(X_1, X_k) \\ \text{Cov}(X_2, X_1) & \text{Var}(X_2) & & \\ \vdots & & \ddots & \\ \text{Cov}(X_k, X_1) & \dots & & \text{Var}(X_k) \end{bmatrix} = \{ \text{Cov}(X_i, X_j) \}$$

$\text{Cov}(X_i, X_j)$

symmetric matrix due to comm. prop.

$$\Sigma_0 = \text{Corr}(\vec{X}) = \begin{bmatrix} 1 & \text{Corr}(X_1, X_2) & \dots \\ \vdots & \ddots & \\ \text{Corr}(X_k, X_1) & \dots & 1 \end{bmatrix}$$

$$= \{ \text{Corr}(X_i, X_j) \}_{i=1, \dots, k, j=1, \dots, k}$$

since  $\text{Corr}(X_i, X_i) = 1$   
[mu]

$$E[\vec{1}^T \vec{X}] = \sum E(X_i)$$

let  $A = \begin{bmatrix} a_{11} & \dots & a_{1n} \\ \vdots & & \vdots \\ a_{n1} & \dots & a_{nn} \end{bmatrix} \in \mathbb{R}^{n \times n}$

what is  $\vec{1}^T A \vec{1}$ ?  $\in \mathbb{R}$   $(1 \times n) \times (n \times n) = (n \times 1)$

$$\begin{aligned} \vec{1}^T \begin{bmatrix} a_{11} + \dots + a_{1n} \\ \vdots \\ a_{n1} + \dots + a_{nn} \end{bmatrix} &= a_{11} + \dots + a_{1n} + a_{21} + \dots + a_{2n} + \dots + a_{n1} + \dots + a_{nn} \\ &= \sum_{i=1}^n \sum_{j=1}^n a_{ij} \end{aligned}$$

$$\text{Var}[\vec{1}^T \vec{X}] = \sum_{i=1}^k \sum_{j=1}^k \text{Cov}(X_i, X_j)$$