

11/2/17

Standard normal



$$X \sim N(0, 1) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} \Rightarrow \phi(x) = e^{-\frac{x^2}{2}}$$

PDF

 $X_1, \dots, X_n$  iid?

$$\bar{Z}_n = \frac{\bar{X}_n - \mu}{\sigma} \xrightarrow{d} N(0, 1)$$

CLT

$$E(X) = 0$$

$$SE(X) = 1$$

$$Y = \mu + \delta X, \delta \in (0, \infty)$$

$$f_Y(y) = \frac{1}{\delta} f_X\left(\frac{y-\mu}{\delta}\right)$$

$$E(Y) = \mu + \delta E(X) = \mu = \frac{1}{\delta} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(y-\mu)^2}{2}}$$

$$SE(Y) = \delta SE(X) = \delta = \frac{1}{\sqrt{2\pi} \delta^2} e^{-\frac{1}{2\delta^2} (y-\mu)^2}$$

$$= N(\mu, \delta^2) \text{ "Normal"}$$

 $X_1, \dots, X_n$  iid? mean  $\mu$ , variance  $\sigma^2$ 

$$T_n = X_1 + \dots + X_n \xrightarrow{d} N(n\mu, n\sigma^2)$$

$$\bar{X}_n = \frac{X_1 + \dots + X_n}{n} \xrightarrow{d} N\left(\mu, \frac{\sigma^2}{n}\right)$$

"if  $n$  is large enough"

$$\phi_Y(t) = e^{it\mu} \phi_X(st) = e^{it\mu} e^{-\frac{s^2 t^2}{2}} = e^{it\mu - \frac{s^2 t^2}{2}}$$

$$X_1 \sim N(\mu_1, \sigma_1^2) \text{ and } X_2 \sim N(\mu_2, \sigma_2^2)$$

$$Y = X_1 + X_2 \sim \int_{\mathbb{R}} f_{X_1}(x) f_{X_2}(t-x) dx$$

$$\int_{\mathbb{R}} \frac{1}{\sqrt{2\pi}\sigma_1} e^{-\frac{1}{2\sigma_1^2}(x-\mu_1)^2} \frac{1}{\sqrt{2\pi}\sigma_2} e^{-\frac{1}{2\sigma_2^2}(t-x-\mu_2)^2} dx$$

No indicator function  
dx = -d(t-x)

$$\phi_Y(t) = \phi_{X_1}(t) \phi_{X_2}(t) = e^{it\mu_1 - \frac{\sigma_1^2 t^2}{2}} e^{it\mu_2 - \frac{\sigma_2^2 t^2}{2}}$$

$$= e^{it(\mu_1 + \mu_2) - \frac{(\sigma_1^2 + \sigma_2^2)t^2}{2}} \Rightarrow Y \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$$

$$\text{supp } f_X = (0, \infty)$$

$$X \sim N(\mu, \sigma^2), Y = e^X = g(X) \quad X = g^{-1}(Y) = \ln Y$$

$$\left| \frac{d}{dy} [g^{-1}(y)] \right| = \frac{1}{y}$$

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} [g^{-1}(y)] \right|$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \cdot \frac{1}{y} e^{-\frac{1}{2\sigma^2}(\ln(y) - \mu)^2} = \text{LogN}(\mu, \sigma^2)$$

"log normal"

you begin with  $Y_0$  of \$\$. You must return it for  $t$  time periods.  $R_t$  is the rate of return (varies)

$$Y_1 = Y_0(1 + R_1)$$

$$Y_2 = Y_1(1 + R_2) = Y_0(1 + R_1)(1 + R_2)$$

$$Y_t = Y_0 \prod_{i=1}^t (1 + R_i) = Y_0 e^{\ln \left( \prod_{i=1}^t (1 + R_i) \right)}$$

$$= Y_0 e^{\sum_{i=1}^t \ln(1 + R_i)} = Y_0 e^{\sum_{i=1}^t X_i}$$

$$\sum_{i=1}^t X_i \stackrel{d}{\sim} N(t\mu_X, t\sigma_X^2)$$

Note:  $\sum_{i=1}^t X_i \stackrel{d}{\sim} \log N(t\mu_X, t\sigma_X^2)$

$$X \sim \log N(\mu, \sigma^2), Y = aX \quad a \in (0, \infty)$$

$$Y \sim f_Y(y) = \frac{1}{a} f_X\left(\frac{y}{a}\right) = \frac{1}{a} \frac{1}{\sqrt{2\pi}\sigma} \cdot \frac{1}{a} e^{-\frac{1}{2\sigma^2} \left( \ln\left(\frac{y}{a}\right) - \mu \right)^2}$$

$$= \log N(\mu + \ln(a), \sigma^2)$$

$$X_i = \ln(1 + R_i)$$

$$\mu_R, \sigma_R^2$$

$$R = 0.03 \approx 3\%$$

$$X = \ln(1 + 0.03) \approx 0.0296$$

$$R = -0.05 \approx -5\%$$

$$X \approx \ln(1 - 0.05) \approx -0.051$$

$$\Rightarrow X \approx R$$

$$\ln(1+x) \approx x - \frac{x^2}{2} + \frac{x^3}{3} - \frac{x^4}{4} + \dots \approx x$$

if  $x$  is small

$$\Rightarrow Y_t = Y_0 e^{\sum_{i=1}^t R_i} \approx \log N(\mu_R + \ln(Y_0), t\sigma_R^2)$$

So:

$$R_t \sim N(10\%, 10\%^2)$$

start with \$1000

In 5 years what is the prob you have more than \$1650?

$$P(Y_t > 1650) = 1 - F_{Y_t}(1650)$$

$$\text{with } Y_t \approx \log N\left(\underbrace{0.5 + 6.91}_{(\ln(1000))}, \frac{5 \cdot 10^2}{.05}\right)$$

$$P(Y_t > 1650) = 1 - \Phi_{\text{norm}}(1650, 7.41, \sqrt{105})$$

If a quantity experiences normal percentage / proportional changes, then the result quantity is  $\log N$ .

$$Z \sim N(0, 1), Y = Z^2, \text{supp}(Y) = (0, \infty)$$

↑  
not 1-1

$$F_Y(y) = P(Y \leq y) = P(Z^2 \leq y) = P(Z \in [-\sqrt{y}, \sqrt{y}])$$

$$= 2P(Z \in [0, \sqrt{y}])$$

$$= 2(F_Z(\sqrt{y}) - F_Z(0)) = 2F_Z(\sqrt{y}) - 1$$

1/2

$$f_Y(y) = F_Y'(y) = 2f_Z(\sqrt{y}) \cdot \frac{1}{2\sqrt{y}} = f_Z(\sqrt{y}) \cdot y^{-1/2}$$

$$= \frac{1}{\sqrt{y}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(\sqrt{y})^2}{2}} = \frac{1}{\sqrt{y}} \frac{1}{\sqrt{2\pi}} e^{-y/2} = \chi_1^2$$

Chi-Square  
with degrees of  
freedom = 1  
Parameter

$$\Gamma(x) = \int_0^{\infty} t^{x-1} e^{-t} dt$$

$$\Gamma\left(\frac{1}{2}\right) = \int_0^{\infty} t^{-1/2} e^{-t} dt$$

$$\text{let } u = \sqrt{t}, \frac{du}{dt} = \frac{1}{2\sqrt{t}} \Rightarrow dt = 2\sqrt{t} du = 2u du$$

$$\Rightarrow t = u^2$$

$$\Gamma\left(\frac{1}{2}\right) = \int_0^{\infty} t^{-1/2} e^{-t} dt = \int_0^{\infty} \frac{1}{u} e^{-u^2} 2u du = 2 \int_0^{\infty} e^{-u^2} du = \sqrt{\pi}$$

$$= \frac{\left(\frac{1}{2}\right)! y^{-1/2} e^{-y/2}}{\Gamma\left(\frac{1}{2}\right)} = \text{Gamma}\left(\frac{1}{2}, \frac{1}{2}\right) = \chi_1^2$$

$$z_1, z_2 \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1) \quad \text{Exp}\left(\frac{1}{2}\right)$$

$$z_1^2 + z_2^2 \sim \text{conv of Gaussian}(1, 1/2) = \chi_2^2$$

convolution  $\chi_1^2$  mit  $\chi_1^2$

$$= \text{conv of Gamma}\left(\frac{1}{2}, \frac{1}{2}\right) \text{ mit Gamma}\left(\frac{1}{2}, \frac{1}{2}\right)$$

$$z_1, \dots, z_k \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1)$$

$$\sum_{i=1}^k z_i^2 \sim \text{Gamma}\left(\frac{k}{2}, \frac{1}{2}\right) = \chi_k^2$$

$$= \frac{\left(\frac{1}{2}\right)^{\frac{k}{2}} x^{\frac{k}{2}-1} e^{-\frac{x}{2}}}{\Gamma\left(\frac{k}{2}\right)}$$

$$X \sim \chi_k^2, \quad Y = \sqrt{X}, \quad g'(y) = y^2 \Rightarrow \left| \frac{d}{dy} g(y^2) \right| = 2y$$

$\text{supp}[Y] = (0, \infty)$

$$f_Y(y) = f_X(y^2) 2y = \frac{1}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} \cdot (y^2)^{\frac{k}{2}-1} e^{-\frac{y^2}{2}} \cdot 2y$$

$$= \frac{1}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} (y^2)^{\frac{k}{2}} e^{-\frac{y^2}{2}}$$

$$= \frac{1}{2^{\frac{k}{2}} \Gamma\left(\frac{k}{2}\right)} y^{k-1} e^{-\frac{y^2}{2}} \sim \chi_k$$

Chi-distribution