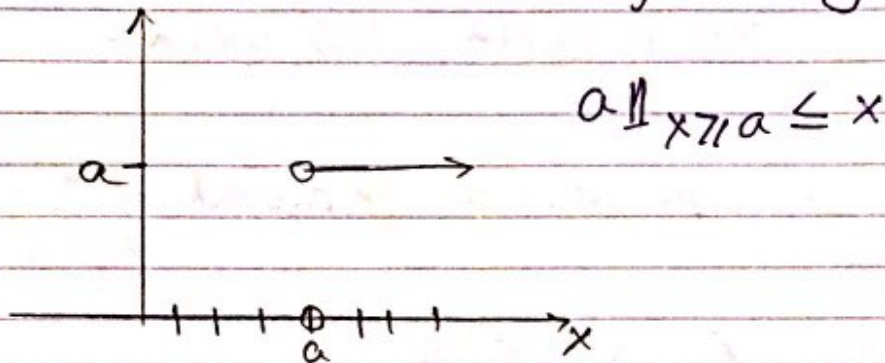


12/02/19

Let X be a non-negative r.v. (ie - $\text{supp}[X] \geq 0$) and has finite expectation μ . Let $a > 0$ (a is constant) and consider the following inequality



verify the inequality:

If $X \geq a$ then $a(1) \leq X \Rightarrow X \geq a$

if $X < a$ then $a(0) \leq X \Rightarrow X \geq 0$ (this because X is non-negative)

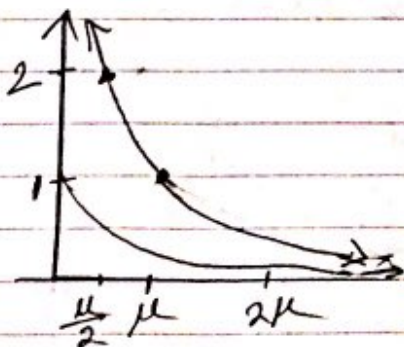
Let's take $E[\cdot]$ on both sides

$$E[a \mathbb{1}_{X \geq a}] \leq E[X]$$

$$\Rightarrow a E[\mathbb{1}_{X \geq a}] \leq \mu \Rightarrow E[\mathbb{1}_{X \geq a}] \leq \frac{\mu}{a}$$

$$\Rightarrow P(X \geq a) \leq \frac{\mu}{a} \Rightarrow 1 - F(a) \leq \frac{\mu}{a}$$

\nwarrow CDF



Tail Markov's Inequality
"Crude Bound"

(2)

$$X \sim \text{Exp}(1) \Rightarrow \mu = 1, \sigma^2 = 1 \Rightarrow P(X \geq a) = e^{-a}$$

a	$P(X \geq a)$	markov	chebyshev	chernoff
2	0.1353	0.5	1	0.73526
5	0.0067	0.2	0.0635	0.09158
10	0.0004	0.1	0.0123	0.00123

Corollaries of markov's inequality:

* $b = a\mu$
 Note $b > 0$, $P(X \geq b) \leq \frac{\mu}{b}$
 $\Rightarrow P(X \geq a\mu) \leq \frac{1}{a}$

* let $h(x)$ be 1:1 increasing
 $P(h(X) \geq h(a)) \leq \frac{E[h(X)]}{h(a)}$
 $\Rightarrow P(X \geq a) \leq \frac{E[h(X)]}{h(a)}$ if X is continuous with connected support

* Let $a = \text{quantile } [X, p] = F_X^{-1}(p)$

$$P(X \geq a) \leq \frac{\mu}{a} \Rightarrow 1 - F_X(a) \leq \frac{\mu}{a}$$

$$\Rightarrow 1 - F_X[F_X^{-1}(p)] \leq \frac{\mu}{F_X^{-1}(p)}$$

$$\Rightarrow 1 - p \leq \frac{\mu}{F_X^{-1}(p)} = F_X^{-1}(p) \leq \frac{\mu}{1-p}$$

e.g. $p = \frac{1}{2} \Rightarrow \text{med}[X] = F_X^{-1}(\frac{1}{2}) \leq 2\mu$

(3)

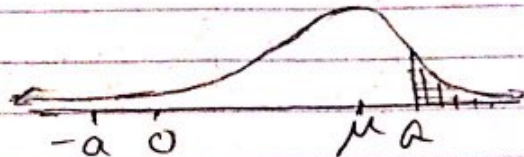
* Let X be any r.v

Note $|X|$ is non-neg

$$P(|X| \geq a) \leq \frac{E[|X|]}{a}$$

right and left
tail

Note $E[|X|] < \infty$



* Let X be any r.v mean μ , variance σ^2
let $Y = (X - \mu)^2$ Note Y is non-negative

$$P(Y \geq a^2) \leq \frac{E[Y]}{a^2}$$

$$P((X - \mu)^2 \geq a^2) \leq \frac{E(X - \mu)^2}{a^2}$$

$$\Rightarrow P(|X - \mu| \geq a) \leq \frac{\sigma^2}{a^2}$$

Chebyshev's Inequality

Assume X is non-negative

$$\Rightarrow P((X - \mu) \geq a \text{ or } (X - \mu) \leq -a)$$

$$= P(X - \mu \geq a) + P((X - \mu) \leq -a)$$

$$= P(X \geq \mu + a) + P(X \leq \mu - a)$$

$$\text{let } a > \mu \Rightarrow P(X \geq \mu + a) + P(X \leq 0)$$

$$\text{let } b = \mu + a \Rightarrow b > 2\mu$$

$$\Rightarrow P(X \geq b) \leq \frac{\sigma^2}{(b - \mu)^2}$$

(4)

corollaries of Markov's Inequality:

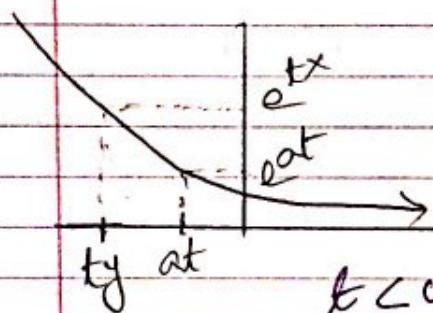
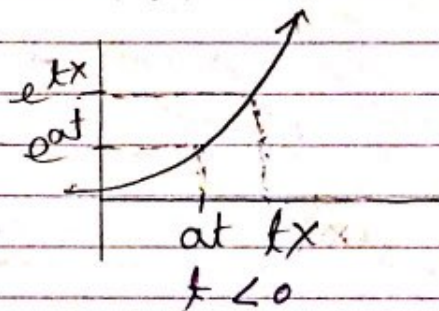
* let x be any r.v

$y = e^{tx}$ where $t \neq 0$ note y is non-neg
 $\forall t$

$$P(y \geq c) \leq \frac{E[y]}{c} \quad \text{let } c = e^{at} > 0$$

$$\Rightarrow P(e^{tx} \geq e^{at}) \leq \frac{E(t e^{tx})}{e^{at}} = e^{-at} m_x(t)$$

$$P(e^{tx} \geq e^{at}) \leq m_x(t) \quad \text{moment-generating function}$$



$$t < 0, P(x \leq a) \leq e^{-at} m_x(t)$$

Inequality for both right and left tails
 note valid for all $t \neq 0$

$$\text{if } t > 0 \quad P(x \geq a) \leq \min_{t > 0} \{ e^{-at} m_x(t) \}$$

$$t < 0 \quad P(x \leq a) = \min_{t < 0} \{ e^{-at} m_x(t) \}$$

ex $X \sim \text{Exp}(\lambda)$

$$m_x(t) = \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx = \lambda \int_0^{\infty} e^{(t-\lambda)x} dx$$

$$= \frac{\lambda}{t-\lambda} \left[e^{(t-\lambda)x} \right]_0^\infty \Rightarrow \text{if } t-\lambda < 0 \quad \frac{\lambda}{t-\lambda} (0-1) = \frac{\lambda}{\lambda-t} \quad (3)$$

if $\lambda=1$ (ie. $x \sim \text{EXP}(1)$) then, $m_x(t) = \frac{1}{1-t}$

$$\text{let } g(t) = \frac{e^{-at}}{1-t}$$

$$g'(t) = \frac{(1-t)(-a)e^{-at} - e^{-at}(-1)}{(1-t)^2}$$

$$= e^{-at}(a(t-1)+1) = 0$$

$$at - a + 1 = 0 \Rightarrow at = a - 1 \Rightarrow t = \frac{a-1}{a}$$

$$\Rightarrow t = 1 - \frac{1}{a} \in (0,1) \text{ if } a > 1$$

$$P(X \leq a) \leq e^{-a(1-\frac{1}{a})} \frac{1}{1-(1-\frac{1}{a})}$$

$$= ae^{-a+1} = \frac{e \cdot a}{e^a}$$

x, y are r.v.'s with means m_x, m_y and variance σ_x^2, σ_y^2

$$\text{cov}[x, y] := E[xy] - m_x m_y$$

measure of "linear dependence" in the units of x and y

$$\text{corr}[x, y] = \frac{\text{cov}[x, y]}{SE[x] SE[y]}$$

"correlation" which is unitless and bounded $\in (-1, 1)$

6

let $y = cx$ where $c \neq 0 \rightarrow$ constant

$$\text{corr}[X, Y] = \frac{\text{cov}[X, cX]}{\text{SE}[X] |c| \text{SE}[X]} = \frac{c \text{Cov}[X, X]}{\text{SE}[X] |c| \text{SE}[X]}$$

$$= \frac{c}{|c|} \frac{\sigma_x^2}{\sigma_x \sigma_x} = \begin{cases} 1 & \text{if } c > 0 \\ -1 & \text{if } c < 0 \end{cases}$$

— — —