

Not covered on final

Justify T-test, Omnibus F-test, partial F-test in Linear Regression.

$\vec{Y} = X\vec{\beta} + \vec{\epsilon}$ Typical assumption $\vec{\epsilon} \sim N_n(\vec{0}, \sigma^2 I_n) \Rightarrow \vec{Z} = \frac{1}{\sigma} \vec{\epsilon} \sim N_n(\vec{0}, I_n)$

$n \begin{bmatrix} Y \\ \vdots \\ 1 \end{bmatrix} = n \begin{bmatrix} X \\ \vdots \\ 1 \end{bmatrix} \begin{bmatrix} \beta \\ \vdots \\ 1 \end{bmatrix} + n \begin{bmatrix} \epsilon \\ \vdots \\ 1 \end{bmatrix}$ Homoskedasticity (same variance)
 $\Rightarrow \vec{Y} - X\vec{\beta} = \vec{\epsilon}$
 $\Rightarrow \frac{\vec{Y} - X\vec{\beta}}{\sigma} \sim N_n(\vec{0}_n, I_n)$

Using vector densities, the least square estimate of unknown β is:

$\hat{\beta} = (X^T X)^{-1} X^T \vec{Y} = (X^T X)^{-1} X^T (X\vec{\beta} + \vec{\epsilon}) = (X^T X)^{-1} X^T X \vec{\beta} + (X^T X)^{-1} X^T \vec{\epsilon}$

$\Rightarrow \hat{\beta} \sim N(\vec{\beta}, \underbrace{(X^T X)^{-1} X^T \sigma^2 \frac{1}{\sigma} \vec{\epsilon}}_{\vec{Z}})$
 $= N(\vec{\beta}, \sigma^2 (X^T X)^{-1})$

$\Rightarrow \beta_k \sim N(\beta_k, \sigma^2 (X^T X)^{-1}_{kk})$

$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim N(0, 1) \Rightarrow \frac{\hat{\beta}_k - \beta_k}{\sigma \sqrt{(X^T X)^{-1}_{kk}}} \sim N(0, 1)$

$\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \sim T_{n-1}$

But we have student's problem, σ is unknown.

$\frac{1}{\sigma^2} \vec{\epsilon}^T \vec{\epsilon} = \frac{1}{\sigma^2} \underbrace{\vec{\epsilon}^T P \vec{\epsilon}}_{\chi_p^2} + \frac{1}{\sigma^2} \underbrace{\vec{\epsilon}^T (I - P) \vec{\epsilon}}_{\chi_{n-p}^2}$



$\text{rank}[X] = p < n$

$P = X(X^T X)^{-1} X^T$
 orthogonal projection matrix
 $\text{rank}[P] = p$

$I - P$ orthogonal projection onto set of \mathbb{R}^n

$\text{rank}[I - P] = n - p$

$(X(X^T X)^{-1} X^T)(X(X^T X)^{-1} X^T)$

$PP = P$

$(I - P)(I - P) = I - P$

$\vec{\epsilon}^T P \vec{\epsilon} = \vec{\epsilon}^T P P \vec{\epsilon} = \vec{\epsilon}^T P^T P \vec{\epsilon} = (P \vec{\epsilon})^T P \vec{\epsilon} \xrightarrow{*(\frac{1}{\sigma^2})}$
 $P \vec{\epsilon} = P(Y - X\beta) = PY - PX\beta$
 $= X(X^T X)^{-1} X^T Y - X(X^T X)^{-1} X^T X \beta$
 $= X\hat{\beta} - X\beta = X(\hat{\beta} - \beta)$

$$\Rightarrow \frac{\frac{1}{\sigma^2} (X(\hat{\beta} - \beta))^T (X(\hat{\beta} - \beta))}{\sigma^2} = \frac{(\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)}{\sigma^2} \sim \chi_p^2$$

$$\vec{E}^T (I-p) (I-p) \vec{E} = ((I-p)\vec{E})^T ((I-p)\vec{E})$$

$$\begin{aligned} (I-p)\vec{E} &= (I-p)(\vec{Y} - X\hat{\beta}) = IY - pY - IX\hat{\beta} + pX\hat{\beta} \\ &= Y - pY - X\hat{\beta} + X(X^T X)^{-1} X^T X \hat{\beta} \\ &= Y - X(X^T X)^{-1} X^T Y = Y - X\hat{\beta} = E \end{aligned}$$

$$\begin{aligned} \frac{1}{\sigma^2} \vec{E}^T \vec{E} &= \frac{1}{\sigma^2} \vec{E}^T p \vec{E} + \frac{1}{\sigma^2} \vec{E}^T (I-p) \vec{E} \\ &= \frac{1}{\sigma^2} (\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta) + \frac{1}{\sigma^2} SSE \end{aligned}$$

$\Rightarrow \vec{E}, \hat{\beta}$ independent

$$E \left[\frac{E^T E}{\sigma^2} \right] = \frac{n-p}{\sigma^2}$$

$$\text{Let } MSE = \frac{SSE}{n-p}$$

$$E \left[\frac{MSE}{\sigma^2} \right] = E \left[\frac{SSE}{(n-p)\sigma^2} \right] = \frac{1}{\sigma^2}$$

$$\Rightarrow E[MSE] = \sigma^2$$

$$\sqrt{MSE} \approx \sigma$$

"
RMSE 'root mean square error'"

$$\text{Back} \Rightarrow \frac{\hat{\beta}_k - \beta_k}{\sigma \sqrt{(X^T X)^{-1}_{kk}}} \sim N(0, 1)$$

$$\Rightarrow \frac{\hat{\beta}_k - \beta_k}{RMSE \sqrt{(X^T X)^{-1}_{kk}}} = \frac{\hat{\beta}_k - \beta_k}{\sqrt{\frac{SSE}{n-p}} \sqrt{(X^T X)^{-1}_{kk}}} = \frac{\hat{\beta}_k - \beta_k}{\sigma \sqrt{(X^T X)^{-1}_{kk}}} \sim T_{n-p}$$

"t test in Linear regression"

$H_0: \vec{\beta} = \text{some vector of values.}$

$$\begin{aligned} &= \frac{\frac{\vec{E}^T p \vec{E}}{\sigma^2 p}}{\frac{\vec{E}^T (I-p) \vec{E}}{\sigma^2 (n-p)}} \sim F_{p, n-p} \text{ 'Omnibus F-test'} \\ &= \frac{(\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta)}{\frac{p}{\frac{SSE}{n-p}}} \text{ 'In Linear Regression'} \end{aligned}$$

$$\chi^2_{n-p}$$