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NOT ON

EXAM FROM THIS PART: ~~2nd~~ 2nd PDF. Start.

Derive F test in linear regression.

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$$\vec{Y} = X\vec{\beta} + \vec{\epsilon}$$

$$n \begin{bmatrix} Y \\ \vdots \end{bmatrix} = n \begin{bmatrix} X \\ \vdots \end{bmatrix} + n \begin{bmatrix} \epsilon \\ \vdots \end{bmatrix}$$

p full rank

columns are not duplicate
Span = p dimensions.

Standard assumption:

$$\vec{\epsilon} \sim N_n(\vec{0}, \sigma^2 I)$$

Independent, normal, homoskedastic.

 $\vec{\beta}, \sigma^2$ are unknown parameters.

$$\frac{1}{\sigma} \vec{\epsilon} \sim N_n(\vec{0}, I) \quad \frac{1}{\sigma^2} \vec{\epsilon}^T \vec{\epsilon} \sim \chi_n^2$$

least squares minimal estimate for $\vec{\beta}$ as:

Estimator

for $\vec{\beta}$

$$\begin{aligned} \hat{\vec{\beta}} &= (X^T X)^{-1} X^T \vec{Y} = (X^T X)^{-1} X^T (X\vec{\beta} + \vec{\epsilon}) \\ &= (X^T X)^{-1} X^T X \vec{\beta} + (X^T X)^{-1} X^T \vec{\epsilon} \\ &= \vec{\beta} + (X^T X)^{-1} X^T \vec{\epsilon} \end{aligned}$$

$$11/04 \quad \hat{\beta} \sim N, \left(\vec{\beta}, \underbrace{(X^T X)^{-1} X^T (\sigma^2 I) X (X^T X)^{-1}}_{\sigma^2 (X^T X)^{-1}} \right)$$

$$= N_p(\vec{\beta}, \sigma^2 (X^T X)^{-1}) \Rightarrow E[\hat{\beta}] = \vec{\beta}$$

unbiased estimator.

Based on marginalizing: $\hat{\beta}_k \sim N(\beta_k, \sigma^2 (X^T X)^{-1}_{kk})$

$$\Rightarrow \frac{\hat{\beta}_k - \beta_k}{\sigma \sqrt{(X^T X)^{-1}_{kk}}} \sim N(0, 1)$$

"Student's" problem. σ is unknown. Use estimate instead.

$$\frac{1}{\sigma^2} \hat{\Sigma}^T \hat{\Sigma} = \frac{1}{\sigma^2} \hat{\Sigma}^T (P) \hat{\Sigma} + \frac{1}{\sigma^2} \hat{\Sigma}^T (I-P) \hat{\Sigma}$$

χ^2_n

$$P := X(X^T X)^{-1} X^T \quad \text{orthogonal projection matrix.}$$

$n \times n$ $\text{rank}[P] = p$

$I-P :=$ Another orthogonal projection onto $n-p$ "missing dimensions"

$$PP = X(X^T X)^{-1} X^T X(X^T X)^{-1} X^T = P$$

$$(I-P)(I-P) = II - PI - IP + PP = I - P - P + P = I - P$$

both idempotent

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Cochran's Theorem applies.

$$\frac{1}{\sigma^2} \vec{\hat{\epsilon}}^T \vec{\hat{\epsilon}} = \chi_p^2 + \chi_{n-p}^2$$

independent

$$\vec{\hat{\epsilon}}^T P \vec{\hat{\epsilon}} = \vec{\hat{\epsilon}}^T P P \vec{\hat{\epsilon}} = (P \vec{\hat{\epsilon}})^T P \vec{\hat{\epsilon}} = (\vec{\hat{\beta}} - \vec{\beta})^T X^T X (\vec{\hat{\beta}} - \vec{\beta})$$

$$\vec{\hat{\epsilon}}^T (I - P) \vec{\hat{\epsilon}} = \vec{\hat{\epsilon}}^T (I - P)(I - P) \vec{\hat{\epsilon}} = ((I - P) \vec{\hat{\epsilon}})^T (I - P) \vec{\hat{\epsilon}} = E^T E = SSE$$

$$P \vec{\hat{\epsilon}} = P(\vec{Y} - X \vec{\beta}) = P \vec{Y} - P X \vec{\beta} = \cancel{X(X^T X)^{-1} X^T X \vec{\beta}}$$

$$= \underbrace{X(X^T X)^{-1} X^T}_{\hat{\beta}} \vec{Y} - X(X^T X)^{-1} X^T X \vec{\beta} = X(\vec{\hat{\beta}} - \vec{\beta})$$

$$(I - P) \vec{\hat{\epsilon}} = (I - P)(\vec{Y} - X \vec{\beta}) = I \vec{Y} - P \vec{Y} - I X \vec{\beta} + P X \vec{\beta}$$

$$= \vec{Y} - P \vec{Y} - X \vec{\beta} + X \vec{\beta}$$

$$= \vec{Y} - X \hat{\beta} = \vec{\hat{\epsilon}} \text{ residuals}$$

$$E\left[\frac{1}{\sigma^2} SSE\right] = n - p$$

$$E\left[\frac{SSE}{n - p}\right] = \sigma^2$$

$$E\left[\frac{1}{\sigma^2} \frac{SSE}{n - p}\right] = 1$$

$$\Rightarrow MSE \approx \sigma^2$$

$$RMSE = \sqrt{MSE} \approx \sigma$$

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SSE is independent of $(\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta) \Rightarrow \text{RMSE } \hat{\beta} \text{ indep}$

$$\text{RMSE} \frac{\hat{\beta}_k - \beta_k}{\sqrt{(X^T X)^{-1}_{kk}}} = \frac{\hat{\beta}_k - \beta_k}{\sqrt{\text{MSE} (X^T X)^{-1}_{kk} \frac{\sigma^2}{\sigma^2}}} = \frac{\hat{\beta}_k - \beta_k}{\sigma \sqrt{(X^T X)^{-1}_{kk}}} = \frac{\frac{\hat{\beta}_k - \beta_k}{\sigma}}{\sqrt{(X^T X)^{-1}_{kk}}}$$

I

$$\frac{\frac{1}{\sigma^2} \sum \hat{\epsilon}^2}{\frac{1}{\sigma^2} \sum (I - P) \hat{\epsilon}^2 / n-p} \sim F_{p, n-p}$$

$\sim \chi^2_p / p$

$\parallel \sim \chi^2_{n-p}$

$(\hat{\beta} - \beta)^T X^T X (\hat{\beta} - \beta) / n-p$

$H_0: \vec{\beta} = \text{values.}$

$\text{SSE} / n-p$ justification for
omnibus F-test.

$$\sqrt{\frac{\text{MSE} \cdot \frac{n-p}{n-p}}{\frac{\text{SSE}}{\sigma^2} / n-p}} \sim \sqrt{\frac{\chi^2_{n-p}}{n-p}}$$

$\sim T_{n-p}$ justification for student
T-test in
linear regression.