

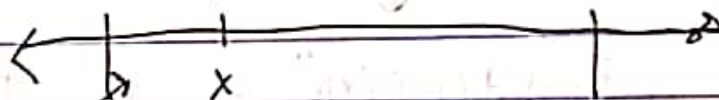
10/28

$X_1, X_2, \dots, X_n \stackrel{iid}{\sim} f(x)$ w.h. CDF = $F(x)$.

$$\Rightarrow F_{X_{(k)}}(x) = \sum_{j=k}^n \underbrace{\binom{n}{j} F(x)^j (1-F(x))^{n-j}}_{\text{PMF Binomial}}$$

comp CDF Binomial

let $U \sim \text{Binomial}(n, p = F(x))$



As x gets close to min, CDF $\rightarrow 0$

how many X_i 's land $\leq x$

The available space to fit in realizations is shrinking.

$$F_{X_{(k)}}(x) = \frac{d}{dx} \left[\sum_{j=k}^n \binom{n}{j} F(x)^j (1-F(x))^{n-j} \right]$$

$$= \sum_{j=k}^n \binom{n}{j} \frac{d}{dx} \left[F(x)^j (1-F(x))^{n-j} \right]$$

$$F(x)^j (-f(x))(n-j)(1-F(x))^{n-j-1} + (1-F(x))^{n-j} f(x) F(x)^{j-1}$$

10/28

$$= f(x) \left(j F(x)^{j-1} (1-F(x))^{n-j} - (n-j) F(x)^j (1-F(x))^{n-j-1} \right)$$

LHS \rightarrow

$$f(x) \left(\sum_{j=k}^n \frac{n!}{j!(n-j)!} j F(x)^{j-1} (1-F(x))^{n-j} - \sum_{j=k}^n \frac{n!}{j!(n-j)!} (n-j) F(x)^j (1-F(x))^{n-j-1} \right)$$

cancel $j/j!$

when $j=n$ the whole thing goes to zero.

only evaluate up to $n-1$.

$$- \sum_{j=k}^{n-1} \frac{n!}{j!(n-j+1)!} F(x)^j (1-F(x))^{n-j+1}$$

Let $L = j+1 \Rightarrow j = L-1$

At $j=k \Rightarrow L=k+1$

$$\sum_{L=k+1}^n \frac{n!}{(L-1)!(n-L+1)!} F(x)^{L-1} (1-F(x))^{n-L+1}$$

$L=k+1$

looks exactly like LHS.

everything cancels except when $j=k$.

So we end up with \Rightarrow

10/28

$$= \frac{n!}{(k-1)!(n-k)!} f(x) F(x)^{k-1} (1-F(x))^{n-k} = f_{X_{(k)}}(x)$$

Want formulas for min and max.

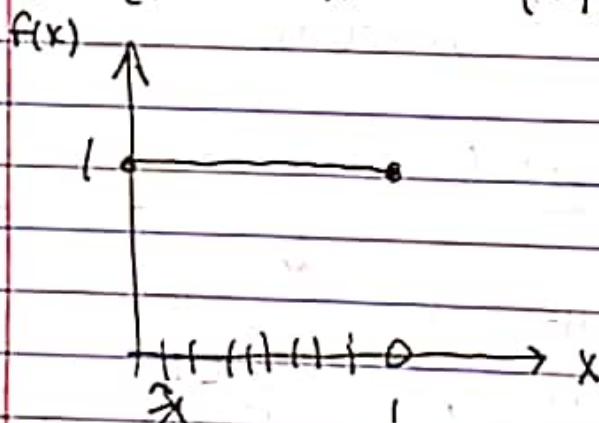
min:

$$\rightarrow f_{X_{(1)}}(x) = n f(x) (1-F(x))^{n-1}$$

max:

$$\rightarrow f_{X_{(n)}}(x) = n f(x) F(x)^{n-1}$$

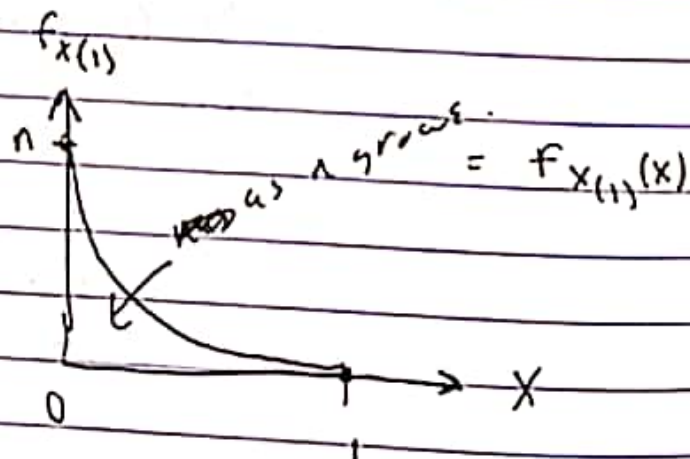
$$X_1, X_2, \dots, X_n \stackrel{iid}{\sim} U(0, 1) = 1, \quad F(x) = x$$



realizations

$X_{(1)}$ = minimum.

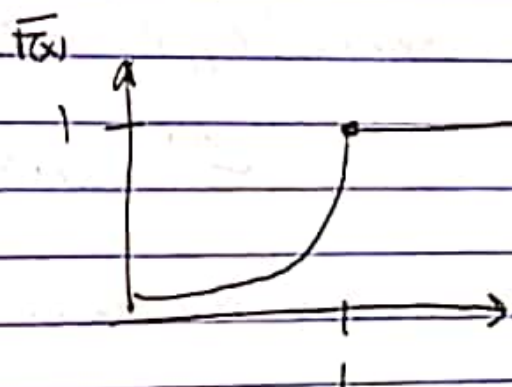
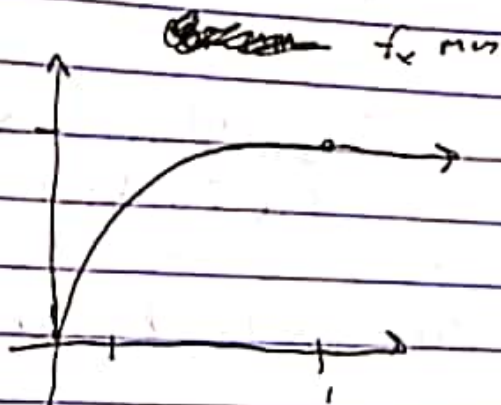
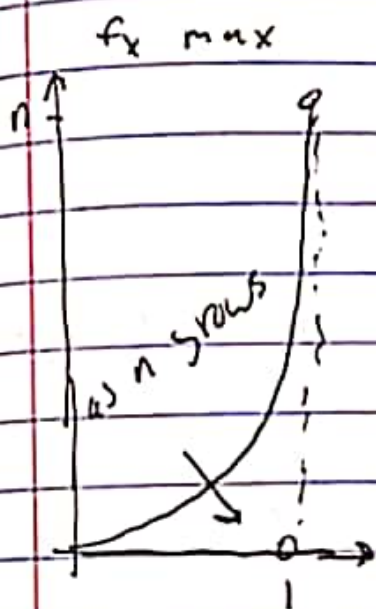
order statistics have the same support.



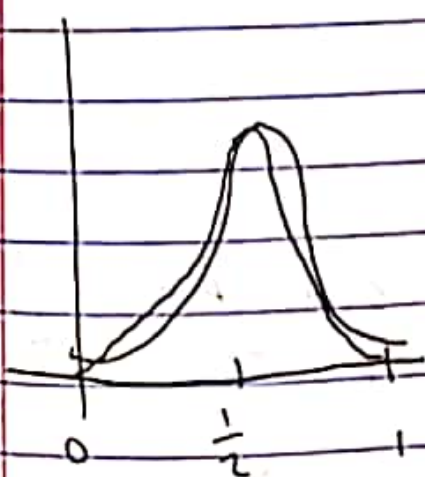
p/28

$$f_{X(1)}(x) = n(1-x)^{n-1}$$

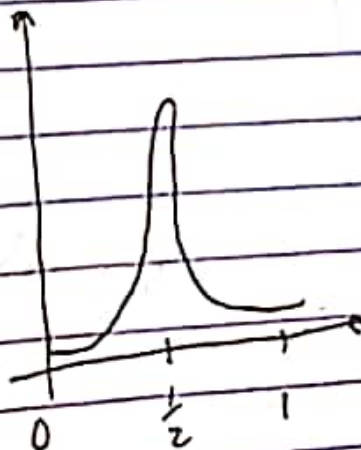
$$F_{X(1)}(x) = 1 - (1-x)^n$$



Say $K \approx \frac{1}{2}n$



As n
grows \rightarrow
layer



10/28

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) x^{k-1} (1-x)^{n-k} \mathbb{1}_{x \in [0,1]}$$

$$= \frac{\Gamma(n+1)}{\Gamma(k)\Gamma(n-k+1)} x^{k-1} (1-x)^{n-k} \mathbb{1}_{x \in [0,1]}$$

$$X_{(k)} \sim \text{Beta}(k, n-k+1).$$

order statistic

let $X \sim \text{Gamma}(\alpha_1, \beta)$ independent of $Y \sim \text{Gamma}(\alpha_2, \beta)$

$$T = X + Y \sim \text{Gamma}(\alpha_1 + \alpha_2, \beta).$$

we will prove this now...

$$f_T(t) = \int_{-\infty}^{\infty} f_X(x) f_Y(t-x) \mathbb{1}_{t-x \in \text{supp}(Y)} dx$$

$$= \int_0^t \left(\frac{\beta^{\alpha_1}}{\Gamma(\alpha_1)} x^{\alpha_1-1} e^{-\beta x} \right) \left(\frac{\beta^{\alpha_2}}{\Gamma(\alpha_2)} (t-x)^{\alpha_2-1} e^{-\beta(t-x)} \right) dx$$

$t-x \in \text{supp}(Y)$
 $x \in [0, t]$

indicator into the limit.

10/28

$$= \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} e^{-\beta t} \int_0^t x^{\alpha_1 - 1} (t-x)^{\alpha_2 - 1} dx$$

let $u = \frac{x}{t} \Rightarrow x = ut \Rightarrow u=0 \Rightarrow x=0,$

$x=t \Rightarrow u=1, \frac{du}{dx} = \frac{1}{t} \Rightarrow dx = t du$

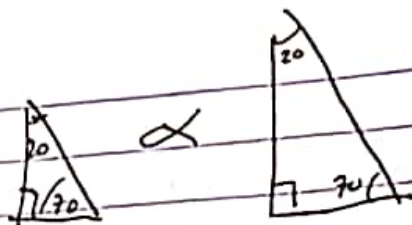
$$\frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} e^{-\beta t} \int_0^1 (ut)^{\alpha_1 - 1} (t-ut)^{\alpha_2 - 1} t du$$

can factor out t now.

$$= \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \left(\int_0^1 u^{\alpha_1 - 1} (1-u)^{\alpha_2 - 1} du \right) t^{\alpha_1 + \alpha_2 - 1} e^{-\beta t}$$

we can decompose
 $p(x) = c k(x)$ or $r(x) = c k(x)$ where c is not a function of x (normalized constant)
 $\propto k(x)$ $\propto k(x)$ and k is a function of x (kernel)

10/28



$$\sum_{\text{all } x} p(x) = 1 \Rightarrow \sum c k(x) = 1 \Rightarrow c = \frac{1}{\sum k(x)}$$

$$\int f(x) dx = 1 \Rightarrow \int c f(x) dx = 1 \Rightarrow c = \frac{1}{\int f(x) dx}$$

$k(x)$ specifies $p(x)$ or $f(x)$. The kernel tells you what random variable you have.

Example.

$$X \sim \text{Bin}(n, p) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$= \underbrace{n! (1-p)^n}_c \underbrace{\frac{1}{x! (n-x)!} \left(\frac{p}{1-p}\right)^x}_{k(x)}$$

(note for space fr/d.)

$k(x) = \text{kernel}$

$$X \sim \text{weibull}(k, \lambda) = \left(\frac{k}{\lambda}\right) (\lambda x)^{k-1} e^{-(\lambda x)^k}$$

$$= \underbrace{\left(\frac{k}{\lambda}\right) \lambda^k}_{c} \underbrace{x^{k-1} e^{-(\lambda x)^k}}_{k(x)}$$

10/28

$$X \sim \text{Gamma}(\alpha, \beta) := \underbrace{\frac{\beta^\alpha}{\Gamma(\alpha)}}_c \underbrace{x^{\alpha-1} e^{-\beta x}}_{K(x)} \quad \text{Gamma}(\alpha, \beta)$$

Back to the original problem.

$$\left(\frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \int_0^1 u^{\alpha_1 - 1} (1-u)^{\alpha_2 - 1} du \right) e^{-(\alpha_1 + \alpha_2)x}$$

" $K(x)$ of gamma

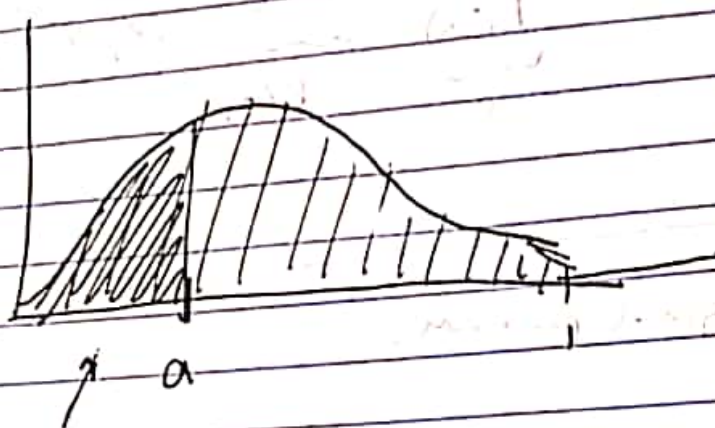
$$\text{Gamma}(\alpha_1 + \alpha_2, \beta)$$

$$\Rightarrow \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1 + \alpha_2)} = \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \int_0^1 u^{\alpha_1 - 1} (1-u)^{\alpha_2 - 1} du \Rightarrow$$

$$\int_0^1 u^{\alpha_1 - 1} (1-u)^{\alpha_2 - 1} du = \frac{\Gamma(\alpha_1) \Gamma(\alpha_2)}{\Gamma(\alpha_1 + \alpha_2)} =: B(\alpha_1, \alpha_2) \quad \text{Beta function.}$$

$$\int_0^\infty x^{\alpha_2 - 1} e^{-x} dx$$

10/28



$$B(a, \alpha_1, \alpha_2) := \int_0^a u^{\alpha_1-1} (1-u)^{\alpha_2-1} du$$

↳ incomplete Beta function.

I may be interested in the proportion of the shaded area vs. the whole.

$$I_a(\alpha_1, \alpha_2) := \frac{B(a, \alpha_1, \alpha_2)}{B(\alpha_1, \alpha_2)}$$

regularized incomplete Beta function

$$X \sim \text{Beta}(\alpha, \beta) := \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} \mathbb{I}_{x \in [0,1]}$$

$$\Gamma(\alpha + \beta)$$

parameter space

$$\alpha, \beta > 0$$

$$\Gamma(\alpha) \Gamma(\beta)$$

if $\alpha, \beta \in \mathbb{N}$

then \Rightarrow

$$(\alpha + \beta + 1)!$$

$$(\alpha-1)! (\beta-1)!$$

10/28

CDF:

$$F(x) = \int_0^x \frac{1}{\beta(\alpha, \beta)} t^{\alpha-1} (1-t)^{\beta-1} dt$$

$$= \frac{\int_0^x t^{\alpha-1} (1-t)^{\beta-1} dt}{\beta(\alpha, \beta)} = \frac{P(x, \alpha, \beta)}{\beta(\alpha, \beta)}$$

$$= I_x(\alpha, \beta).$$