

Transformations of densities (more dim)

we've done these transformations in 1 Dimension:

X is a continuous random variable

g is a one-to-one function

and $Y = g(X)$

then:

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} [g^{-1}(y)] \right|$$

controls stretching
or compression

Let \vec{X} be a vector random variable

with dimension n $\vec{X} = (X_1, X_2, \dots, X_n)$

let its density $f_{\vec{X}}(\vec{X})$ be known $f_{\vec{X}}(x_1, x_2, \dots, x_n)$

Let $\vec{Y} = g(\vec{X})$ where $g: \mathbb{R}^n \rightarrow \mathbb{R}^n$ $g(x_1, x_2, \dots, x_n) = (y_1, y_2, \dots, y_n)$

with an inverse function h (i.e. $\vec{X} = h(\vec{Y})$)

then,

$$f_{\vec{Y}}(\vec{y}) = f_{\vec{X}}(h(\vec{y})) |J_h(\vec{y})|$$

where J_h , the Jacobian Determinant is defined

$$J_h = \det \begin{pmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_1}{\partial y_2} & \dots & \frac{\partial h_1}{\partial y_n} \\ \frac{\partial h_2}{\partial y_1} & \frac{\partial h_2}{\partial y_2} & \dots & \frac{\partial h_2}{\partial y_n} \\ \vdots & \vdots & & \vdots \\ \frac{\partial h_n}{\partial y_1} & \frac{\partial h_n}{\partial y_2} & \dots & \frac{\partial h_n}{\partial y_n} \end{pmatrix}$$

$$\det \begin{pmatrix} a & b \\ c & d \end{pmatrix} = ad - bc$$

note that $g(x_1, x_2, \dots, x_n) = (y_1, y_2, \dots, y_n)$ and $h(y_1, y_2, \dots, y_n) = (x_1, \dots, x_n)$

so $Y_1 = g_1(X_1, X_2, \dots, X_n)$

and $X_1 = h_1(Y_1, Y_2, \dots, Y_n)$

$Y_2 = g_2(X_1, X_2, \dots, X_n)$ etc.

$X_2 = h_2(Y_1, Y_2, \dots, Y_n)$ etc.

Steps to take:

for

$T = X_1 + X_2$ want to find pdf of T : $f_T(t) = ?$

① Find a "clever" g so that we can...

② Find h (inverse of g)

③ Compute J_h (Jacobian)

④ Substitute into change of variables formula

$$f_{\vec{Y}}(\vec{y}) = f_{\vec{X}}(h(\vec{y})) |J_h(\vec{y})| = f_{\vec{X}}(h_1(y_1), h_2(y_2)) |J_h(\vec{y})|$$

⑤ Integrate out the nuisance dimensions.

notice
 $T = Y_1$

① $Y_1 = X_1 + X_2$ so $g_1(X_1, X_2) = X_1 + X_2$
 $Y_2 = X_2$ so $g_2(X_1, X_2) = X_2$

$Y_1 = X_1 + X_2$
 $Y_2 =$

can use
lowercase
for all
of these

② $X_1 = Y_1 - X_2 = Y_1 - Y_2$ so $h_1(Y_1, Y_2) = Y_1 - Y_2$
 $X_2 = Y_2$ so $h_2(Y_1, Y_2) = Y_2$

③

$$J_h = \det \begin{pmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_1}{\partial y_2} \\ \frac{\partial h_2}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \end{pmatrix} = \det \begin{pmatrix} 1 & -1 \\ 0 & 1 \end{pmatrix} = (1)(1) - (-1)(0)$$

$$= (1)(1) - (-1)(0) = 1$$

$$J_h = 1$$

$$\det \begin{pmatrix} a & b \\ c & d \end{pmatrix} = ad - bc$$

④ $f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}(y_1 - y_2, y_2) |1| = f_{X_1, X_2}(y_1 - y_2, y_2)$
↑
joint density

⑤ Want marginal density of Y_1 (since $T = Y_1$ here)

$$f_{Y_1}(y) = \int_{\mathbb{R}} f_{Y_1, Y_2}(y, y_2) dy_2$$

"marginizing"

if X_1, X_2 indep.

$$\Rightarrow f_T(y) = \int_{\mathbb{R}} f_{X_1, X_2}(t-u, u) du = \int_{\mathbb{R}} f_{X_1}(t-u) f_{X_2}(u) du$$

$t = X_1 + X_2$
 $X_2 = t - X_1$
 Let $X_1 = u$
 $X_2 = t - u$

If X_1, X_2 are i.i.d

$$f_T(y) = \int_{\text{supp}[X_1]} f_{X_1}(t-u) f_{X_2}(u) \mathbb{1}_{t-u \in \text{supp}[X_2]} du$$

ratio $\rightarrow R = \frac{X_1}{X_2} \sim f_R(r) = ?$ (Find density of R)

density = PDF

① Find a clever g :

$$\left. \begin{array}{l} Y_1 = \frac{X_1}{X_2} \quad \text{so } g_1(X_1, X_2) = \frac{X_1}{X_2} \\ Y_2 = X_2 \quad \text{so } g_2(X_1, X_2) = X_2 \end{array} \right\} g(X_1, X_2) = \left(\frac{X_1}{X_2}, X_2 \right)$$

② Find h , the inverse of G

$$X_1 = Y_1 X_2 = Y_1 Y_2 \quad \text{so } h_1(Y_1, Y_2) = Y_1 Y_2$$

$$X_2 = Y_2 \quad \text{so } h_2(Y_1, Y_2) = Y_2$$

$$\textcircled{3} J_h = \det \left(\begin{bmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_1}{\partial y_2} \\ \frac{\partial h_2}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \end{bmatrix} \right) = \det \left(\begin{bmatrix} y_2 & y_1 \\ 0 & 1 \end{bmatrix} \right)$$

$$= y_2 \cdot 1 - y_1 \cdot 0 = y_2$$

Substitute into change of variables

$$\textcircled{4} f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}(y_1 y_2, y_2) |y_2|$$

$$\textcircled{5} f_{Y_1}(y_1) = \int_{\mathbb{R}} f_{Y_1, Y_2}(y_1, y_2) dy_2 \quad \text{Integrate out extra dimensions}$$

$$= \int_{\mathbb{R}} f_{X_1, X_2}(ru, u) |u| du \quad \leftarrow \begin{array}{l} \text{these} \\ \text{work} \\ \text{for any} \\ \text{ratio} \end{array}$$

if are indep.

$$f_{Y_1}(y_1) = \int_{\text{supp}[X_1]} f_{X_1}(ru) f_{X_2}(u) \mathbb{1}_{u \in \text{supp}[X_2]} |u| du$$

$$R = \frac{X_1}{X_1 + X_2} \sim f_R(r) = ? \quad (\text{find this density})$$

$$(1) Y_1 = g_1(X_1, X_2) = \frac{X_1}{X_1 + X_2}$$

$$Y_2 = g_2(X_1, X_2) = X_1 + X_2$$

$$(2) X_1 = Y_1(X_1 + X_2) = Y_1 Y_2 \quad \text{so } h_1(Y_1, Y_2) = Y_1 Y_2$$

$$X_2 = Y_2 - X_1 = Y_2 - Y_1 Y_2 \quad \text{so } h_2(Y_1, Y_2) = Y_2 - Y_1 Y_2$$

$$(3) J_h = \det \begin{pmatrix} \frac{\partial h_1}{\partial y_1} & \frac{\partial h_1}{\partial y_2} \\ \frac{\partial h_2}{\partial y_1} & \frac{\partial h_2}{\partial y_2} \end{pmatrix} = \det \begin{pmatrix} y_2 & y_1 \\ y_2 & 1 - y_1 \end{pmatrix}$$

$$J_h = y_2(1 - y_1) - y_1(-y_2) = y_2 - y_2 y_1 + y_1 y_2 = y_2$$

$$J_h = y_2$$

$$(4) f_{Y_1, Y_2}(y_1, y_2) = f_{X_1, X_2}(y_1 y_2, y_2 - y_1 y_2) |y_2|$$

$$(5) f_R(r) = \int_{\mathbb{R}} f_{X_1, X_2}(ru, u - ru) |u| du$$

$$\text{if } X_1, X_2 \text{ indep.} = \int_{\mathbb{R}} f_{X_1}(ru) f_{X_2}(u - ru) |u| du$$

$$= \int_{\text{Supp}[X_1]} f_{X_1}(ru) f_{X_2}(u - ru) \mathbb{1}_{u - ru \in \text{Supp}[X_2]} |u| du$$

$$X_1 \sim \text{Gamma}(\alpha_1, \beta) \quad X_2 \sim \text{Gamma}(\alpha_2, \beta) \quad \Rightarrow X_1, X_2 \text{ are indep.}$$

$$R = \frac{X_1}{X_1 + X_2} \sim f_R(r) = ?$$

$$\text{Supp}[R] = (0, 1)$$

$$\text{Supp}[X_1] = (0, \infty)$$

$$\text{Supp}[X_2] = (0, \infty)$$

$$\text{Supp}[X_1 + X_2] = (0, \infty)$$

proportion
of
a Gamma
rel. to
total of
Gamma's

Using previous formula...

$$\text{PDF } f_R(r) = \int_0^\infty \left(\frac{\beta^{\alpha_1}}{\Gamma(\alpha_1)} (ru)^{\alpha_1-1} e^{-\beta ru} \right) \left(\frac{\beta^{\alpha_2}}{\Gamma(\alpha_2)} (u - ru)^{\alpha_2-1} e^{-\beta(u - ru)} \right) \cdot \mathbb{1}_{u - ru \in (0, \infty)} u du$$

same as
 $u(1-r) \in (0, \infty)$
same as
 $u \in (0, \infty)$
since $(1-r) \in (0, 1)$

$$= \int_0^\infty \left(\frac{\beta^{\alpha_1}}{\Gamma(\alpha_1)} r^{\alpha_1-1} u^{\alpha_1-1} \right) \left(\frac{\beta^{\alpha_2}}{\Gamma(\alpha_2)} u^{\alpha_2-1} (1-r)^{\alpha_2-1} e^{-\beta u} \right) u du$$

$$= \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} r^{\alpha_1-1} (1-r)^{\alpha_2-1} \int_0^\infty u^{\alpha_1 + \alpha_2 - 1} e^{-\beta u} du = \frac{\Gamma(\alpha_1 + \alpha_2)}{\beta^{\alpha_1 + \alpha_2}}$$

$$= \frac{\Gamma(\alpha_1 + \alpha_2)}{\Gamma(\alpha_1) \Gamma(\alpha_2)} r^{\alpha_1-1} (1-r)^{\alpha_2-1}$$

$$= \frac{1}{\mathcal{B}(\alpha_1, \alpha_2)} = \text{Beta}(\alpha_1, \alpha_2)$$

$$X_1 \sim \text{Gamma}(\alpha_1, \beta) \quad X_2 \sim \text{Gamma}(\alpha_2, \beta) \Rightarrow X_1, X_2 \text{ are indep.}$$

$$R = \frac{X_1}{X_2} \sim \int_{\text{Supp}[X_1]} f_{X_1}(ru) f_{X_2}(u) \mathbb{1}_{u \in \text{Supp}[X_2]} |u| du$$

$\uparrow u \in (0, \infty)$
 so $|u| = u$

Supp[R]
is $(0, \infty)$

$$= \int_0^\infty \left(\frac{\beta^{\alpha_1} (ru)^{\alpha_1-1} e^{-\beta ru}}{\Gamma(\alpha_1)} \right) \left(\frac{\beta^{\alpha_2} u^{\alpha_2-1} e^{-\beta u}}{\Gamma(\alpha_2)} \mathbb{1}_{u \in (0, \infty)} \right) u du$$

Support
does not
match Beta
dist.
so this
is not
Beta \rightarrow

$$= \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} r^{\alpha_1-1} \int_0^\infty u^{\alpha_1 + \alpha_2 - 1} e^{-\beta(r+u)} du$$

$$= \frac{\beta^{\alpha_1 + \alpha_2}}{\Gamma(\alpha_1) \Gamma(\alpha_2)} \cdot \frac{\Gamma(\alpha_1 + \alpha_2)}{(\beta(r+1))^{\alpha_1 + \alpha_2}}$$

$\underbrace{\beta^{\alpha_1 + \alpha_2} (r+1)^{-\alpha_1 - \alpha_2}}_{\beta^{\alpha_1 + \alpha_2} (r+1)^{-\alpha_1 - \alpha_2}}$

$$= \frac{1}{B(\alpha_1, \alpha_2)} \frac{r^{\alpha_1-1}}{(1+r)^{\alpha_1 + \alpha_2}} \mathbb{1}_{r \in (0, \infty)}$$

$$= \text{BetaPrime}(\alpha_1, \alpha_2)$$

called a BetaPrime distribution

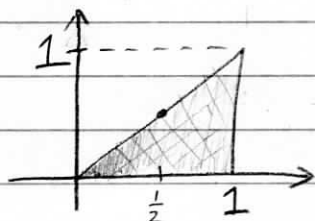
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Conditional Densities

Consider $\overset{\text{r.v.}}{X} \sim \overset{\text{uniform}}{U}(0,1)$

$$Y | X=x \sim U(0,x)$$

r.v. X is \uparrow parameter needed for this conditional density



$$f_{Y|X=1}(y) = 1 \quad \text{also written } f_{Y|X}(y;1) = 1$$

$$f_{Y|X=0.1}(y) = 10 \quad \text{also written } f_{Y|X}(y;0.1) = 10$$

$$f_{X,Y} = ? \quad f_Y = ? \quad f_{X|Y} = ?$$

\uparrow
will use Bayes Theorem
to do this