

①

$$D \sim \text{Laplace}(0, 1) = \frac{1}{2} e^{-|d|}$$



He published this in 1774, called it the "first law of errors".

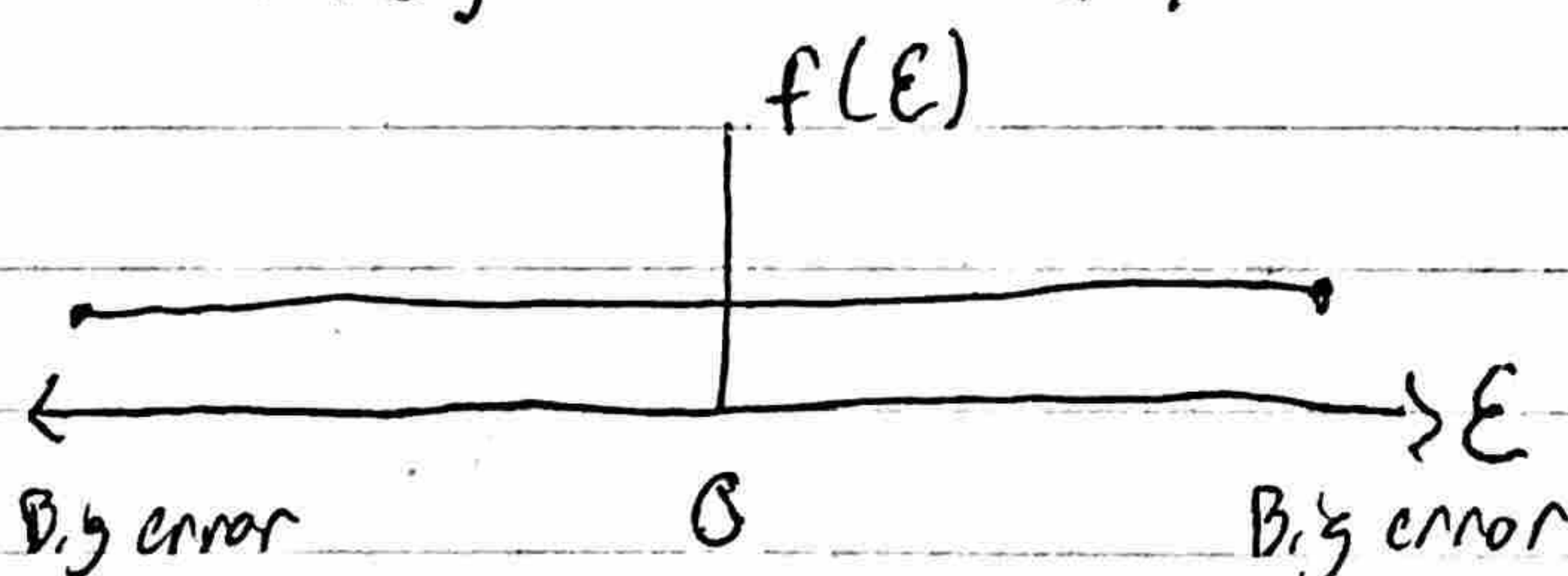
Imagine you're measuring a quantity " v " with additive random error denoted E . So the measure M is $= v + E$ which is also random

It makes sense to assume $E[E] = 0$

$\rightarrow E[M] = v$ (makes M an "unbiased estimate")

$\rightarrow \text{Med}[E] = 0 \rightarrow 50\%$ you overestimate
 50% you underestimate

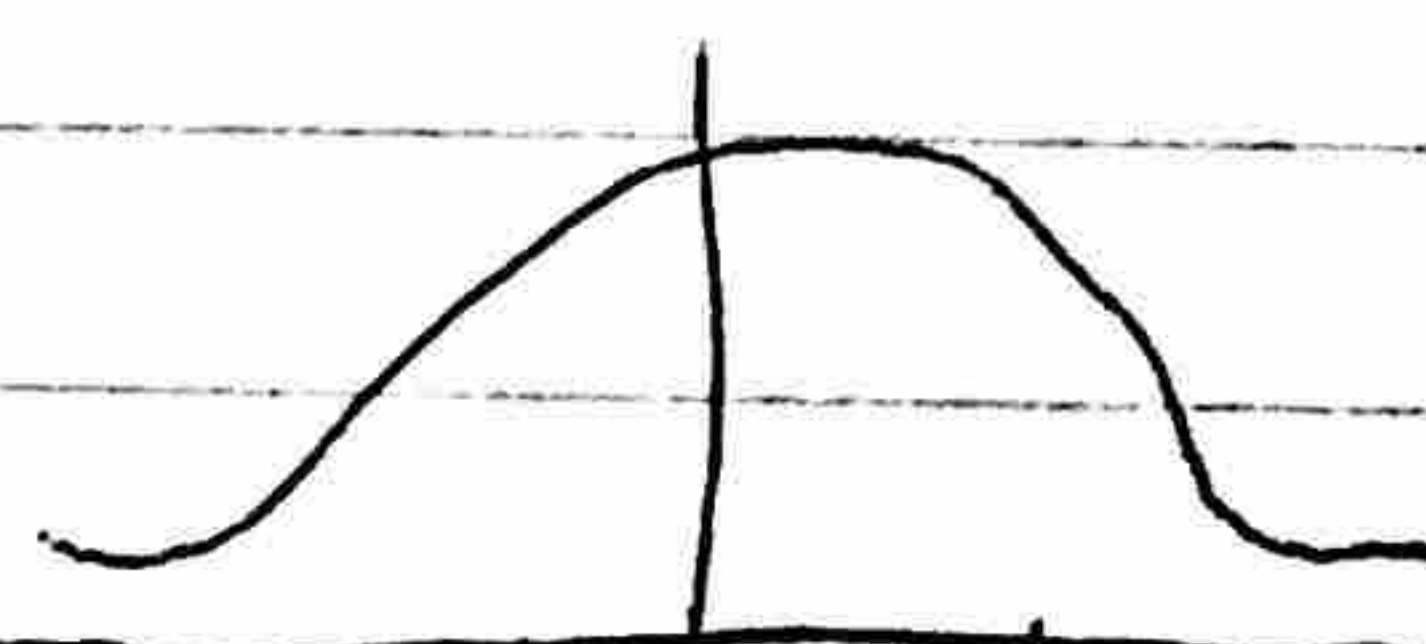
$\rightarrow f(E) = f(-E)$ (symmetric around 0)



Problem: Big error probability = small error probability

$\rightarrow f'(E) < 0$ if $E > 0$

$\rightarrow f'(E) > 0$ if $E < 0$



Why not let $f''(E) = f'(c) \rightarrow f(E) = c e^{-bE}$ ~~Laplace(0,1)~~
 $\propto \text{Laplace}(0, 1)$

(2)

Logistic also good for error.

$$X \sim \text{Exp}(1) = e^{-x} \mathbb{I}_{x \geq 0}$$

$$Y = \frac{1}{\lambda} X^{1/\kappa} \sim ? \quad \text{where } \kappa, \lambda > 0$$

- find inverse
- find derivative
- plug in

$$\lambda Y = X^{1/\kappa} \rightarrow X = (\lambda Y)^\kappa = \lambda^\kappa Y^\kappa = g^{-1}(Y)$$

$$\left| \frac{d}{dy} [g^{-1}(y)] \right| = \lambda^\kappa \kappa y^{\kappa-1}$$

$$f_Y(y) = e^{-(\lambda y)^\kappa} \mathbb{I}_{\lambda^\kappa y^\kappa \geq 0} \quad \left| \begin{array}{l} \lambda^\kappa y^\kappa \geq 0 \\ y^\kappa \geq 0 \\ y \geq 0 \end{array} \right. \begin{array}{l} \text{re-written as} \\ \lambda \cdot \lambda^{\kappa-1} \end{array} = (\kappa \lambda) (\lambda y)^{\kappa-1} e^{-(\lambda y)^\kappa} \mathbb{I}_{y \geq 0}$$

$$= \text{Weibull}(\kappa, \lambda)$$

→ famous waiting time / survival model

(3)

$$F(y) = \int_0^y (k\lambda)(\lambda t)^{k-1} e^{-(\lambda t)^k} dt$$

$$\text{let } u = (\lambda t)^k = \lambda^k t^k \quad t=0 \rightarrow u=0 \quad t=y \rightarrow u = (\lambda y)^k$$

$$\frac{du}{dt} = \lambda^k k t^{k-1} \rightarrow dt = \frac{1}{\lambda^k k t^{k-1}} du$$

$$\rightarrow = \int_0^{(\lambda y)^k} k \lambda t^{k-1} e^u \frac{1}{k \lambda^k t^{k-1}} du$$

$$= [-e^{-u}]_0^{(\lambda y)^k} = 1 - e^{-(\lambda y)^k} \rightarrow P(Y \geq y)$$

$$\rightarrow \bar{F}(y) = 1 - F(y) = e^{-(\lambda y)^k}$$

complement

$$\text{Weibull}(1, \lambda) = \lambda e^{-\lambda y} \mathbb{I}_{y \geq 0} = \text{Exp}(\lambda)$$

$$y=1, c=15 \quad P(Y \geq y+c | Y \geq c) = \frac{P(Y \geq y+c \cap Y \geq c)}{P(Y \geq c)} \\ = \frac{\bar{F}(y+c)}{\bar{F}(c)}$$

 $Y \sim \text{Weibull}$

$$\rightarrow \frac{e^{-(\lambda(y+c))^k}}{e^{-(\lambda y)^k}} = e^{-\lambda^k (y+c)^k} e^{\lambda^k y^k} \\ = e^{\lambda^k (e^k - (y+c)^k)}$$

(4)

If $\boxed{K=1} \rightarrow e^{\lambda(c-(y+c))} = e^{-\lambda y} = \bar{F}(y)$
 $= P(Y \geq y)$

$K > 1 \rightarrow P(Y \geq y+c | Y \geq c) < P(Y \geq y)$
 with inequality getting greater if c gets larger.

$K < 1 \rightarrow P(Y \geq y+c | Y \geq c) > P(Y \geq y)$

Ex] $P(Y \geq 98 \text{ yrs} | Y \geq 97 \text{ yrs}) < P(Y \geq 1 \text{ yr})$

"P that someone who is 97 lives to be 98 is less than the P that an infant who lives to be 1 yr old lives longer than 1 year"

Ex] $P(Y \geq 250,000 \text{ mi} | Y \geq 249,000 \text{ mi}) < P(Y \geq 1000)$

"P that a car makes it to 250,000 miles given it made it to 249,000 miles is less than the P that it makes it to at least 1000 miles."

$K = 1/2 \quad c^{1/2} + y^{1/2} > (y+c)^{1/2}$

$\rightarrow c + y + 2c^{1/2}y^{1/2} > y + c$

$\rightarrow 2c^{1/2}y^{1/2} > 0$

(5)

$$K=2 \quad e^{\lambda^2(c^2 - (y+c)^2)} < e^{\lambda^2 y^2}$$

$$\rightarrow \lambda^2(c^2 - (y+c)^2) < -\lambda^2 y^2$$

$$= c^2 + y^2 < (y+c)^2$$

$$= c^2 + y^2 < y^2 + c^2 + 2cy$$

$$= 0 < 2cy \quad c, y > 0$$

Order Statistics p 160-161

Consider continuous r.v's X_1, \dots, X_n

Let the order statistics $X_{(1)}, X_{(2)}, \dots, X_{(n)}$ be defined as

$X_{(1)} := \min\{X_1, \dots, X_n\}$ called minimum

$X_{(n)} := \max\{X_1, \dots, X_n\}$ called maximum

$X_{(k)} := k^{\text{th}}$ largest of $\{X_1, \dots, X_n\}$ called the k^{th} largest

Let $R := X_{(n)} - X_{(1)}$ called range

As ex, $n=4$ for realizations

$$\boxed{r = 12 - 2 = 10}$$

$$x_1 = 9 \quad x_2 = 2 \quad x_3 = 7 \quad x_4 = 12$$

$$x_{(1)} = 2 \quad x_{(2)} = 7 \quad x_{(3)} = 9 \quad x_{(4)} = 12$$

7

Let's get PDF and CDF for
the minimum, $X_{(1)}$

$$F_{X_{(1)}}(x) = P(X_{(1)} \leq x) = 1 - P(X_{(1)} > x)$$

$$= 1 - P(X_1 > x, X_2 > x, \dots, X_n > x)$$

$$\text{if i.i.d.} \rightarrow = 1 - \prod_{i=1}^n P(X_i > x)$$

$$= 1 - \prod_{i=1}^n (1 - F_{X_i}(x))$$

$$= 1 - (1 - F(x))^n$$

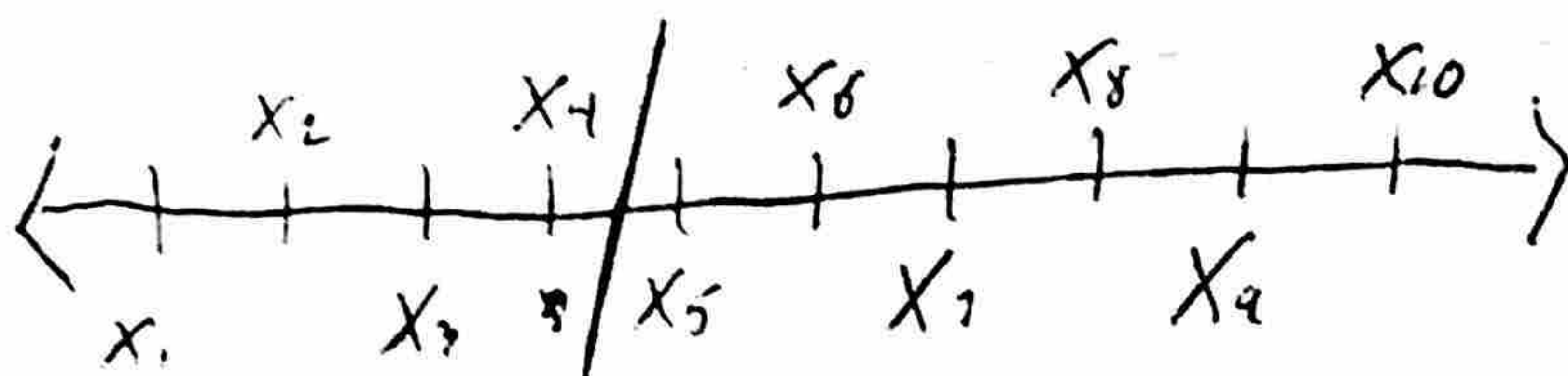
$$x := \frac{d}{dx} [F_{X_{(1)}}(x)] = -(-f(x))n(1 - F(x))^{n-1}$$

$$= n f(x) (1 - F(x))^{n-1}$$

8

Find PDF and CDF of $X_{(k)}$, the k^{th} order stat (difficult)

Let $n=10$, $k=4$



Consider $P(X_1 \leq x, X_2 \leq x, X_3 \leq x, X_4 \leq x, X_5 > x, \dots, X_{10} > x)$

$$\text{ind} \rightarrow = \prod_{i=1}^4 F_{X_i}(x) \prod_{i=5}^{10} (1 - F_{X_i}(x))$$

$$\text{iid} \rightarrow = F(x)^4 (1 - F(x))^6$$

Now consider P that 4 of the 10 $\leq x$, the other 6 of the 10 $> x$

$$= \sum_{\substack{\text{all subsets} \\ \{s_1, s_2, s_3, s_4\} \\ \subset \\ \{1, 2, \dots, 10\}}} P(X_{s_1} \leq x, \dots, X_{s_4} \leq x, X_{s_5} > x, \dots, X_{s_{10}} > x)$$

$$\text{ind} = \sum_{\text{all subsets}} \prod_{i=1}^4 F_{X_{s_i}}(x) \prod_{i=5}^{10} (1 - F_{X_{s_i}}(x))$$

$$\text{iid} = \sum_{\text{all subsets}} F(x)^4 (1 - F(x))^6 = \binom{10}{4} F(x)^4 (1 - F(x))^6$$

(9)

$$F_{X_{(k)}}(x) := P(X_{(k)} \leq x) \stackrel{\text{i.i.d.}}{=} \overset{P(U \text{ is less than } x)}{(10 \choose 4) F(x)^4 (1-F(x))^6} \\ + (10 \choose 5) F(x)^5 (1-F(x))^5 \\ + \dots \\ + (10 \choose 10) F(x)^{10} (1-F(x))^{10-10}$$

$$= \sum_{j=4}^{10} (10 \choose j) F(x)^j (1-F(x))^{10-j}$$

$$F_{X_{(k)}}(x) = \sum_{j=k}^n (n \choose j) F(x)^j (1-F(x))^{n-j}$$

$$F_{X_{(n)}}(x) = \sum_{j=n}^n (n \choose j) \sim$$

$$= (n \choose n) F(x)^n (1-F(x))^{n-n} = F(x)^n$$

$$F_{X(n)}(x) = \sum_{j=1}^n \binom{n}{j} F(x)^j (1-F(x))^{n-j}$$

$$= \left(\sum_{j=0}^n \binom{n}{j} (F(x))^j (1-F(x))^{n-j} \right) - \underbrace{\binom{n}{0}}_1 \underbrace{F(x)^0}_1 \underbrace{(1-F(x))^{n-0}}_{(1-F(x))^n}$$

$$= 1 - (1-F(x))^n$$

$$f_{X(n)}(x) = \frac{d}{dx} [F_{X(n)}(x)] =$$

next
lecture