

Lecture #6

multinom (2n, p)
(n, p)

prob

~~scribble~~

- Uniform discrete:

$$X_1 \& X_2 \stackrel{\text{iid}}{\sim} U(\{0, 1, 2, 3\}) := \begin{cases} 0 & \text{w.p. } \frac{1}{4} \\ 1 & \text{w.p. } \frac{1}{4} \\ 2 & \text{w.p. } \frac{1}{4} \\ 3 & \text{w.p. } \frac{1}{4} \end{cases} = \frac{1}{4} \mathbb{1}_{v \in \{0, 1, 2, 3\}}$$

$$\text{Generally, } X \sim U(A) := \frac{1}{|A|} \mathbb{1}_{v \in A}$$

$$\text{Supp}(X) = A$$

Parameter Space: $A \subset \mathbb{R}$ s.t. A is finite.

- $P_{X_1, X_2}(X_1, X_2) = \frac{1}{16} \mathbb{1}_{X_1 \in \{0, 1, 2, 3\}} \mathbb{1}_{X_2 \in \{0, 1, 2, 3\}}$

- $T = X_1 + X_2$?

$$P(t) = \sum_{X_1 \in \mathbb{R}} \sum_{X_2 \in \mathbb{R}} P_{X_1, X_2}(X_1, X_2) \mathbb{1}_{X_2 = t - X_1}$$

$$P(1) = \sum \sum p \mathbb{1}_{X_2 = 1 - X_1} = 2 \left(\frac{1}{16} \right)$$

→ (pic) taken

$$P(0.5) = 0$$

$$P(3) = 4 \cdot \left(\frac{1}{16} \right)$$

$$P(6) = \frac{1}{16}$$

→ how are they distributed?

- let $\underbrace{Y = -X}_{g(X)} \sim P_Y(Y)$

$$X \sim U(0, 1, 2, 3)$$

$$X=0 \Rightarrow Y=0 \text{ w.p. } \frac{1}{4}$$

$$X=1 \Rightarrow Y=-1 \text{ w.p. } \frac{1}{4}$$

$$X=2 \Rightarrow Y=-2 \text{ w.p. } \frac{1}{4}$$

$$X=3 \Rightarrow Y=-3 \text{ w.p. } \frac{1}{4}$$

$$\Rightarrow Y \sim U(\{0, -1, -2, -3\})$$

It's a rule.

$$X_1, X_2 \stackrel{iid}{\sim} \text{Poisson}(\lambda)$$

$$T = X_1 + X_2 \sim \text{Poisson}(2\lambda)$$

conditional
probability

$$P_{X_1|T}(x_1, t) = P(X_1 = x_1 | T = t) \neq \text{poisson}(\lambda)$$

$$\begin{aligned} P_{X_1|T}(x_1, t) &= \frac{P_{X_1, T}(x_1, t)}{P_T(t)} = \frac{P_{X_1, X_2}(x_1, t-x)}{P_T(t)} = \frac{P_{X_1}(x) P_{X_2}(t-x)}{P_T(t)} \\ &= \frac{\left(\frac{e^{-\lambda} \lambda^x}{x!} \right) \left(\frac{e^{-\lambda} \lambda^{t-x}}{(t-x)!} \right)}{e^{-2\lambda} (2\lambda)^t / t!} = \frac{\lambda^t}{(2\lambda)^t} \binom{t}{x} = \binom{t}{x} \left(\frac{1}{2} \right)^t \\ &= \text{Bin}(t, \frac{1}{2}) \end{aligned}$$

Negative transformation will be included in the midterm.

$$X_1 \& X_2 \stackrel{iid}{\sim} \text{Poisson}(\lambda)$$

$$D = X_1 - X_2$$

$$\text{SUPP}[D] = \mathbb{Z} \text{ (integers)}$$

$$= \underbrace{X_1}_x + \underbrace{(-X_2)}_y = \sum_{x \in \text{supp}(X)} P_X(x) P_Y(d-x) \mathbb{1}_{d-x \in \text{supp}(Y)}$$

$$P_X(x) = \frac{e^{-\lambda} \lambda^x}{x!}, \quad P_Y(y) = \frac{e^{-\lambda} \lambda^{-y}}{(-y)!} = P_X(-y)$$

$$\rightarrow = \sum_{x \in \{0, 1, 2, \dots\}} \left(\frac{e^{-\lambda} \lambda^x}{x!} \right) \frac{e^{-\lambda} \lambda^{-(d-x)}}{(-(d-x))!} \cdot \mathbb{1}_{d-x \in \{0, -1, -2, \dots\}}$$

$$= e^{-2\lambda} \sum_{x \in \{0, 1, 2, \dots\}} \frac{\lambda^{2x-d}}{x!(x-d)!} \mathbb{1}_{x \geq d}$$

$$= e^{-2\lambda} \sum_{x=0}^{\infty} \frac{\lambda^{2x-d}}{x!(x-d)!} \text{ if } d < 0$$

\Rightarrow PIC
taken

$$e^{-2\lambda} \sum_{x=d}^{\infty} \frac{\lambda^{2x-d}}{x!(x-d)!} \text{ if } d \geq 0$$

$$\Rightarrow \text{After that, } e^{-2\lambda} \sum_{x=0}^{\infty} \frac{\left(\frac{2\lambda}{2} \right)^{2x+|d|}}{x!(x+|d|)!} = e^{-2\lambda} I_0(2\lambda) = \text{Skellam}(\lambda, \lambda)$$

modified Bessel function of the first kind

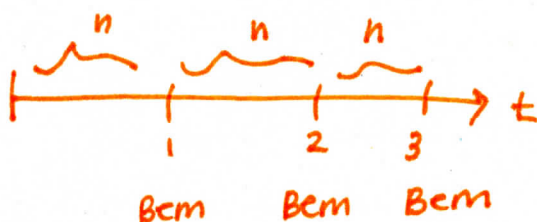


Midterm 2 material:

Let $X_1 \sim \text{Geom}(p) = \underbrace{(1-p)^x p}_{p(x)} \mathbb{1}_{x \in \{0, 1, 2, \dots\}}$

CDF

$$F(x) = P(X \leq x) = 1 - P(X > x) = 1 - P(X \geq 1) = (1 - (1-p)^{x+1}) \mathbb{1}_{x \geq 0}$$



We run n trials within each "original" period.

X_n is the # 0's until a success using the new trial def.

So, $X_n \sim (1-p)^{nx} p \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots\}}$

Every time I take a minute,
there are n in each side

Let $n \rightarrow \infty$, $p \rightarrow 0$, but $\lambda = np \Rightarrow p = \frac{\lambda}{n}$

Similar to the derivation of the Poisson from the binomial.

$$F_{X_n}(x) = 1 - (1-p)^{nx+1}$$

$$P_{X_\infty}(x) = \lim_{x \rightarrow \infty} P_{X_n}(x) = \underbrace{\left(\lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n})^n \right)^x}_{e^{-\lambda}} \underbrace{\lim_{n \rightarrow \infty} \frac{\lambda}{n}}_0 \underbrace{\lim_{n \rightarrow \infty} \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots\}}}_{\mathbb{1}_{x \in (0, \infty)}} = 0$$

$$F_{X_\infty}(x) = \lim_{n \rightarrow \infty} F_{X_n}(x) = 1 - \underbrace{\left(\lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n})^n \right)^x}_{e^{-\lambda}} \underbrace{\left(\lim_{n \rightarrow \infty} 1 - \frac{\lambda}{n} \right)}_1 \mathbb{1}_{x \geq 0} = (1 - e^{-\lambda x}) \mathbb{1}_{x \geq 0}$$

$$\text{Supp}[X_\infty] = [0, \infty)$$

$|\text{Supp}[X_\infty]| = |\mathbb{R}| \Rightarrow X_\infty$ is a continuous r.v.

Has no PMF

CDF is valid (I) $\lim_{x \rightarrow \infty} F(x) = 0$

(III) $F(x)$ monotonically increasing

(II) $\lim_{x \rightarrow \infty} F(x) = 1$

$$F'(x) = \lambda e^{-\lambda x} > 0$$