

Lecture 07
Math 621

09-29-2020

from last class:

$$\text{Supp}[D] = \text{Supp}[X+Y] = \text{Supp}[X] + \text{Supp}[Y] = \mathbb{Z}$$

Convolution formula for independent discrete rv's:

$$P_D(d) = \sum_{x \in \text{Supp}[X]} P_X^{\text{odd}}(x) P_Y^{\text{odd}}(d-x) \mathbb{1}_{d-x \in \text{Supp}[Y]}$$

$$= \sum_{x \in \{0, 1, \dots\}} \frac{e^{-\lambda} \lambda^x}{x!} \frac{e^{-\lambda} \lambda^{-(d-x)}}{(-(d-x))!} \mathbb{1}_{d-x \in \{ \dots, -1, 0 \}}$$

$$= e^{-2\lambda} \sum_{x \in \{0, \dots\}} \frac{\lambda^x}{x!} \frac{\lambda^{x-d}}{(x-d)!} \mathbb{1}_{x \in \{d, d+1, \dots\}}$$

$$= e^{-2\lambda} \begin{cases} d > 0 & \sum_{x \in \{d, d+1, \dots\}} \frac{\lambda^{2x-d}}{x! (x-d)!} \quad (1) \\ d \leq 0 & \sum_{x \in \{0, 1, \dots\}} \frac{\lambda^{2x-d}}{x! (x-d)!} \quad (2) \end{cases}$$

① let $x' = x - d \Rightarrow x = x' + d$, also $d = |d|$

$$\text{so, } = \sum_{x' \in \{0, 1, \dots\}} \frac{\lambda^{2(x'+d)-d}}{(x'+d)! (x'+d-d)!} = \sum_{x' \in \{0, \dots\}} \frac{\lambda^{2x'+d}}{(x'+d)! x'!}$$

② let $d' = -d = |d|$

$$\text{so, } = \sum_{x \in \{0, \dots\}} \frac{\lambda^{2x+d'}}{x! (x+d')!} = \sum_{x \in \{0, \dots\}} \frac{\lambda^{2x+|d|}}{x! (x+|d|)!}$$

$$I_{|d|}(2\lambda) := \sum_{x=0}^{\infty} \frac{\left(\frac{2\lambda}{2}\right)^{2x+|d|}}{x! (x+|d|)!}$$

Modified Bessel function of the first kind

$$= e^{-2\lambda} I_{|d|}(2\lambda) \mathbb{1}_{d \in \mathbb{Z}} = \text{Skellam}(\lambda, \lambda) \quad (\text{comes up in diff eq's})$$

Note:
|d|, think of d ≤ 0

Skellam (λ, λ) discovered in 1946.

It's used to model point spreads in sports games, photo noise, etc.

$$X_1, X_2 \stackrel{iid}{\sim} \text{Poisson}(\lambda) \Rightarrow T = X_1 + X_2 \sim \text{Poisson}(2\lambda)$$

$$P_{X_1|T}(x, t) = \frac{P_{X_1, T}(x, t)}{P_T(t)} \in \text{JMF}$$

$$= \frac{P_{X_1, X_2}(x, t-x)}{P_T(t)}$$

$$\text{Using iid} = \frac{P_{X_1}(x) P_{X_2}(t-x)}{P_T(t)}$$

$$= \frac{e^{-\lambda} \lambda^x}{x!} \frac{e^{-\lambda} \lambda^{t-x}}{(t-x)!}$$

$$\frac{e^{-2\lambda} (2\lambda)^t}{t!}$$

$$= \frac{t!}{x!(t-x)!} \frac{\lambda^t}{(2\lambda)^t} = \binom{t}{x} \left(\frac{1}{2}\right)^t = \text{Bin}(t, \frac{1}{2})$$

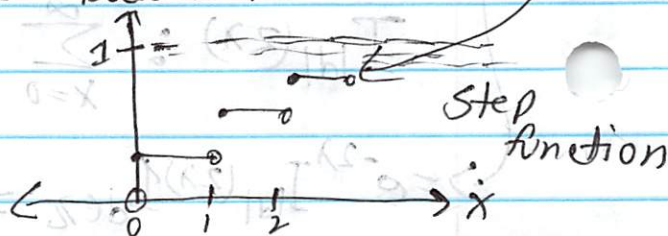
$B_1, B_2, \dots \stackrel{iid}{\sim} \text{Bern}(p)$

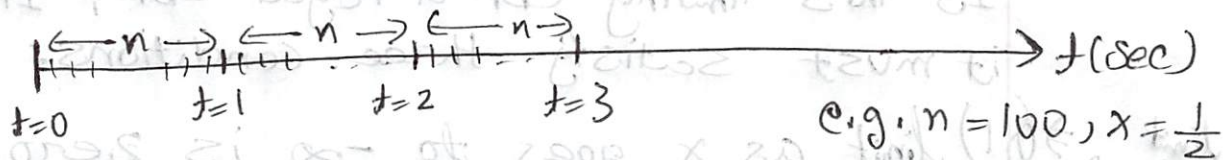
$$X_1 \sim \text{Geom}(p) = \underbrace{(1-p)^x}_{p^{\text{odd}}(x)} P \quad \forall x \in \{0, 1, \dots\}$$

$$F_{X_1}(x) = P(X_1 \leq x) = 1 - P(X_1 > x) = 1 - (1-p)^x$$

If $x=7$, then $X_1 > 7 \Rightarrow X_1 \in \{8, 9, \dots\}$

$\Rightarrow b_1, \dots, b_7 = 0$ means failed 7 times...





let there be n experiments in each second (time unit). X is in seconds...

$$P_{X_n}(x) = (1-p)^{nx} \frac{p}{n} \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots, 1, 1+\frac{1}{n}, 1+\frac{2}{n}, \dots\}}$$

$$F_{X_n}(x) = 1 - (1-p)^{nx+1}$$

let's put infinite experiments into every second (time unit), this is the limit as n goes to positive infinity. X_∞ And $p \rightarrow 0$.

$p \rightarrow 0$ but $\lambda = np \Rightarrow p = \frac{\lambda}{n}$ ala the Poisson.

$$P_{X_\infty}(x) = \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^{nx} \frac{\lambda}{n} \mathbb{1}_{x \in \{0, \frac{1}{n}, \dots\}}$$

$$= \underbrace{\left(\lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)^n\right)^x}_{e^{-\lambda}} \lim_{n \rightarrow \infty} \frac{\lambda}{n} \mathbb{1}_{x \in \{0, \frac{1}{n}, \dots\}}$$

$$= e^{-\lambda x} (0) \mathbb{1}_{x \in [0, \infty)} = 0 \quad \forall x$$

This is not a PMF

because,

$$\text{supp}[X_\infty] = [0, \infty)$$

$$\sum_{x \in \text{supp}[X_\infty]} P_{X_\infty}(x) = 0 \neq 1$$

$$F_{X_\infty}(x) = \lim_{n \rightarrow \infty} 1 - \left(1 - \frac{\lambda}{n}\right)^{nx+1}$$

$$= 1 - \lim_{n \rightarrow \infty} \underbrace{\left(1 - \frac{\lambda}{n}\right)^{nx}}_{e^{-\lambda x}} \cdot \lim_{n \rightarrow \infty} \left(1 - \frac{\lambda}{n}\right)$$

$$= (1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)}$$

intuitive proof

Is this limiting CDF a legal CDF? If so, it must satisfy three conditions:

- $\lim_{x \rightarrow -\infty} \Rightarrow$ 1) limit as x goes to $-\infty$ is zero.
 $\lim_{x \rightarrow \infty} \Rightarrow$ 2) limit as x goes to $+\infty$ is one.
3) increasing function.
i.e. its derivative is ≥ 0

$$\textcircled{1} \lim_{x \rightarrow -\infty} (1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)} = 0 \checkmark$$

$$\textcircled{2} \lim_{x \rightarrow \infty} (1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)} = 1 - \lim_{n \rightarrow \infty} \frac{1}{e^{\lambda x}} = 1 \checkmark$$

Since $\lambda > 0$

$$\textcircled{3} \frac{d}{dx} [(1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)}]$$
$$= \lambda e^{-\lambda x} \mathbb{1}_{x \in [0, \infty)} > 0 \checkmark$$



$\Rightarrow F_{X_\infty}$ is a valid CDF! of what rv??

Answer:

A continuous rv.

A continuous rv X has $\text{Supp}[X] \subseteq \mathbb{R}$

but $|\text{Supp}[X]| = |\mathbb{R}|$,

this size is known as "uncountable infinity" or the "size of the continuum".

They also have no PMF, the $P(X=x)$ is always 0 for every x . But they have a CDF (continuous) for the purposes of this class. And the derivative of the CDF is a very useful function, so it gets a

Special name which is the "probability density function" or just "density" (PDF) denoted f :

$$f_x = F'(x), \quad P(X \in (a, b)) = \underbrace{P(X \leq b)}_{F(b)} - \underbrace{P(X \leq a)}_{F(a)}$$

$$\text{FTC} \rightarrow \int_a^b f(x) dx$$

(Fundamental Thm of Calculus)

$$\rightarrow f(x) \geq 0 \quad \forall x \quad (\text{Property of the CDF})$$

$$\Leftrightarrow \text{Supp}[X] = \{x : f(x) > 0\}$$

$$\int_{\mathbb{R}} f(x) dx = 1 \Rightarrow \int_{-\infty}^{\infty} f(x) dx = \widetilde{F(\infty)} - \widetilde{F(-\infty)} = 1$$

Joint density function (Property of CDF)

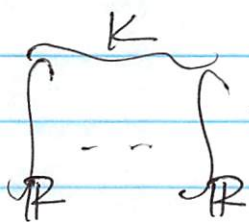
$$\vec{X} = \begin{bmatrix} X_1 \\ \vdots \\ X_k \end{bmatrix}$$

$$\sim f_{\vec{X}}(\vec{x}) = f_{X_1}(x_1) \cdot \dots \cdot f_{X_k}(x_k) = f(x_1) \cdot \dots \cdot f(x_k)$$

all components of continuous

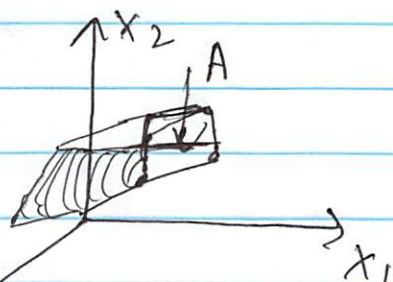
if X_1, \dots, X_k independent

independently distributed



$$\int_{\mathbb{R}} \dots \int_{\mathbb{R}} f_{\vec{X}}(\vec{x}) dx_1 \cdot \dots \cdot dx_k = 1$$

$k=2$



$$P(\vec{X} \in A) = \iint_A f_{\vec{X}}(\vec{x}) dx_1 dx_2$$

$$f_{X_1, X_2}(x_1, x_2)$$

$X \sim \text{Exp}(\lambda) = \underbrace{\lambda e^{-\lambda x}}_{f(x)} \mathbb{1}_{x \in [0, \infty)}$ Exponential RV
 Support

$F(x) = \int_{-\infty}^x f(t) dt$
 $F(x) = \int_0^x \lambda e^{-\lambda t} dt = 1 - e^{-\lambda x}$ for $x \geq 0$
 $F(x) = 0$ for $x < 0$

$f(x) \geq 0$
 $\int_{-\infty}^{\infty} f(x) dx = 1$

$\{0 \leq x < \infty\} = \{x : f(x) > 0\}$

$\int_0^{\infty} f(x) dx = 1 \Leftrightarrow \int_0^{\infty} \lambda e^{-\lambda x} dx = 1$

(CDF) $F(x) = 1 - e^{-\lambda x}$

$f(x) = \lambda e^{-\lambda x}$

probability
 density

independent

continuous

$1 = \int_0^{\infty} \lambda e^{-\lambda x} dx$

$\int_0^{\infty} \lambda e^{-\lambda x} dx = 1$

