

$X_1, X_2 \stackrel{iid}{\sim} \text{Poisson}(\lambda)$  From previous class  $T = X_1 + X_2 \sim \text{Poisson}(2\lambda)$   
 $D = X_1 - X_2 \sim f_D(d) = ?$   
 $D = X_1 + (-X_2) \sim ?$   
 $\text{supp}[D] = \text{supp}[X+Y] = \text{supp}[X] + \text{supp}[Y] = \mathbb{Z}$  all integers  
 convolution formula for independent discrete rv's

$$P_D^{(d)} = \sum_{x \in \text{supp}[X]} P_X^{(x)} P_Y^{(d-x)} \mathbb{1}_{d-x \in \text{supp}[Y]}$$

$$= \sum_{x \in \{0, 1, \dots\}} \frac{e^{-\lambda} \lambda^x}{x!} \frac{e^{-\lambda} \lambda^{-(d-x)}}{(d-x)!} \mathbb{1}_{d-x \in \{0, 1, \dots\}}$$

$$= e^{-2\lambda} \sum_{x \in \{0, 1, \dots\}} \frac{\lambda^x}{x!} \frac{\lambda^{x-d}}{(x-d)!} \mathbb{1}_{x \in \{d, d+1, \dots\}}$$

$$= e^{-2\lambda} \begin{cases} d > 0 & \sum_{x \in \{d, d+1, \dots\}} \frac{\lambda^{2x-d}}{x! (x-d)!} = \sum_{x' \in \{0, 1, \dots\}} \frac{\lambda^{2(x'+d)-d}}{(x'+d)! (x'+d-d)!} = \sum_{x' \in \{0, 1, \dots\}} \frac{\lambda^{2x'+d}}{(x'+d)! x'!} \\ d \leq 0 & \sum_{x \in \{0, 1, \dots\}} \frac{\lambda^{2x-d}}{x! (x-d)!} = \sum_{x \in \{0, 1, \dots\}} \frac{\lambda^{2x+d}}{x! (x+d)!} = \sum_{x \in \{0, 1, \dots\}} \frac{\lambda^{2x+|d|}}{x! (x+|d|)!} \end{cases}$$

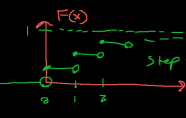
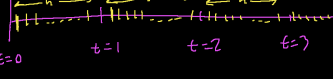
$$I_{|d|}(2\lambda) := \sum_{x=0}^{\infty} \frac{\left(\frac{2\lambda}{2}\right)^{2x+|d|}}{x! (x+|d|)!}$$
 Modified Bessel Function of the First kind (comes up in diff eq's)  

$$= e^{-2\lambda} I_{|d|}(2\lambda) \stackrel{d \in \mathbb{Z}}{=} \text{Skellern}(\lambda, \lambda)$$
 discovered in 1946

It's used to model point spreads in sports games, photo noise, etc

$X_1, X_2 \stackrel{iid}{\sim} \text{Poisson}(\lambda) \Rightarrow T = X_1 + X_2 \sim \text{Poisson}(2\lambda)$   
 $P_{X_1|T}(x, t) = \frac{P_{X_1, T}(x, t)}{P_T(t)} = \frac{P_{X_1, X_2}(x, t-x)}{P_T(t)} = \frac{P_{X_1}(x) P_{X_2}(t-x)}{P_T(t)}$   

$$= \frac{\frac{e^{-\lambda} \lambda^x}{x!} \frac{e^{-\lambda} \lambda^{t-x}}{(t-x)!}}{\frac{e^{-2\lambda} (2\lambda)^t}{t!}} = \frac{t!}{x! (t-x)!} \frac{\lambda^t}{(2\lambda)^t} = \binom{t}{x} \left(\frac{1}{2}\right)^t = \text{Bin}\left(t, \frac{1}{2}\right)$$

$\rightarrow b_1, b_2, \dots \stackrel{iid}{\sim} \text{Bernoulli}(p)$   
 $X_1 \sim \text{Bern}(p) = (1-p)^x p \mathbb{1}_{x \in \{0, 1, \dots\}}$   
 $F_{X_1}(x) := P(X_1 \leq x) = 1 - P(X_1 > x) = 1 - (1-p)^{x+1}$   
  


Let there be n experiments in each second (time unit). X is in seconds ...

$P_{X_n}(x) = (1-p)^{nx} p \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots, 1, 1+\frac{1}{n}, 1+\frac{2}{n}, \dots\}}$   
 $F_{X_n}(x) = 1 - (1-p)^{nx+1}$

Let's put infinite experiments into every second (time unit), this is the limit as n goes to positive infinity.  $X_{\infty}$

And  $p \rightarrow 0$  but  $\lambda = \lim_{n \rightarrow \infty} np = \lambda > 0 \Rightarrow p = \frac{\lambda}{n}$  a la the Poisson,  
 $P_{X_{\infty}}(x) = \lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n})^{nx} \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots\}} = \left( \lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n})^n \right)^x \lim_{n \rightarrow \infty} \frac{\lambda}{n} \lim_{n \rightarrow \infty} \mathbb{1}_{x \in \{0, \frac{1}{n}, \frac{2}{n}, \dots\}} = e^{-\lambda} \mathbb{1}_{x \in [0, \infty)}$   

$$= e^{-\lambda x} (0) \mathbb{1}_{x \in [0, \infty)} = 0 \quad \forall x$$

This is not a PMF because  $\sum_{x \in \text{supp}[X_{\infty}]} P_{X_{\infty}}(x) = 0 \neq 1$   
 $F_{X_{\infty}}(x) = \lim_{n \rightarrow \infty} 1 - (1 - \frac{\lambda}{n})^{nx+1} = 1 - \lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n})^{nx} \cdot \lim_{n \rightarrow \infty} (1 - \frac{\lambda}{n}) = 1 - e^{-\lambda x} \mathbb{1}_{x \in [0, \infty)}$

Is this limiting CDF a legal CDF? If so, it must satisfy three conditions

- (1) limit as x goes to negative infinity is zero
- (2) limit as x goes to positive infinity is one
- (3) increasing function i.e. its derivative is  $\geq 0$

$(1) \lim_{x \rightarrow -\infty} (1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)} = 0 \checkmark$  since  $\lambda > 0$   
 $(2) \lim_{x \rightarrow \infty} (1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)} = 1 - \lim_{x \rightarrow \infty} \frac{1}{e^{\lambda x}} = 1 \checkmark$   
 $(3) \frac{d}{dx} [(1 - e^{-\lambda x}) \mathbb{1}_{x \in [0, \infty)}] = \lambda e^{-\lambda x} \mathbb{1}_{x \in [0, \infty)} > 0 \checkmark$   
 $\Rightarrow F_{X_{\infty}}$  is a valid CDF... but of what rv?? A "continuous rv".

A continuous rv  $X$  has  $\text{supp}[X] \subseteq \mathbb{R}$  but  $|\text{supp}[X]| = |\mathbb{R}|$ ,

this size is known as "uncountable infinity" or the "size of the continuum". They also have no PMF, the  $P(X=x)$  is always zero for every x. But they have a CDF (continuous for the purposes of this class). And the derivative of the CDF is a very useful function, so it gets a special name which is the "probability density function" or just "density" (PDF) denoted f:

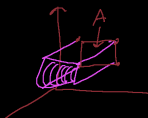
$f(x) := F'(x), P(X \in (a, b)) = P(X \leq b) - P(X \leq a) = \int_a^b f(x) dx$   
 fundamental theorem of calculus

$f(x) \geq 0 \quad \forall x$  (property of the CDF)  $\Leftrightarrow \text{supp}[X] = \{x : f(x) > 0\}$

$\int_{\mathbb{R}} f(x) dx = 1 \quad \int_{-\infty}^{\infty} f(x) dx = \lim_{b \rightarrow \infty} F(b) - \lim_{a \rightarrow -\infty} F(a) = 1$  (property of CDF)  
 joint density function (JDF)

$\vec{X} = \begin{bmatrix} X_1 \\ \vdots \\ X_K \end{bmatrix} \sim f_{\vec{X}}(\vec{x}) = f_{X_1}(x_1) \cdots f_{X_K}(x_K) = f(x_1) \cdots f(x_K)$   
 all components continuous if  $X_1, \dots, X_K$  independent i.i.d.

$\int_{\mathbb{R}} \cdots \int_{\mathbb{R}} f_{\vec{X}}(\vec{x}) dx_1 \cdots dx_K = 1$

$K=2$   
  
 $P(\vec{X} \in A) = \iint_A f_{\vec{X}}(\vec{x}) dx_1 dx_2$

$f_{X_1, X_2}(x_1, x_2)$   
 $X \sim \text{Exp}(\lambda) := \lambda e^{-\lambda x} \mathbb{1}_{x \in [0, \infty)}$   
 exponential rv  
 $\lambda \in (0, \infty)$