

$Z \sim N(0, 1)$ ,  $Y = Z^2 \sim f_Y(y) = ?$  Not 1:1

$$F_Y(y) = P(Y \leq y) = P(Z^2 \leq y)$$

$$= P(Z \in [-\sqrt{y}, \sqrt{y}]) = 2P(Z \in [0, \sqrt{y}])$$

$$= 2(F_Z(\sqrt{y}) - F_Z(0)) = 2F_Z(\sqrt{y}) - 1$$

$$f_Y(y) = \frac{d}{dy} [2F_Z(\sqrt{y}) - 1] = 2 \left( \frac{1}{2} y^{-\frac{1}{2}} \right) f_Z(\sqrt{y}) = y^{-\frac{1}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(\sqrt{y})^2}{2}}$$

$$\propto y^{-\frac{1}{2}} e^{-\frac{1}{2}y} \mathbb{1}_{y \geq 0} \propto \text{Gamma}\left(\frac{1}{2}, \frac{1}{2}\right)$$

Let's say,

$$Z_1, Z_2, \dots, Z_K \stackrel{\text{iid}}{\sim} N(0, 1), Y = Z_1^2 + Z_2^2 + \dots + Z_K^2 \sim \text{Gamma}\left(\frac{K}{2}, \frac{1}{2}\right)$$

Note - The beta is always  $1/2$  and the alpha is always  $K/2$  so  $K$  is the only parameter. And because this is a common situation, we give it a special name

$\text{Gamma}\left(\frac{K}{2}, \frac{1}{2}\right) = \chi_K^2$ , the "chi squared distribution with  $K$  degrees of freedom"  $K \in \mathbb{N}$

$$E[Y] = K E[Z^2] = K$$

$$\chi_K^2 = \frac{\left(\frac{1}{2}\right)^{K/2}}{\Gamma\left(\frac{K}{2}\right)} y^{K/2-1} e^{-y/2} \mathbb{1}_{y \geq 0} \stackrel{K=1, \Gamma(1/2)=\sqrt{\pi}}{=} \frac{1}{\sqrt{2\pi}} y^{-1/2} e^{-y/2} \mathbb{1}_{y \geq 0} = \chi_1^2$$

$$\text{If } X \sim \chi_K^2, Y = \sqrt{X} \Rightarrow X = Y^2 = g^{-1}(Y), \left| \frac{d}{dy} [g^{-1}(y)] \right| = |2Y| = 2Y$$

$$f_Y(y) = f_X(Y^2) \cdot 2Y = \frac{\left(\frac{1}{2}\right)^{K/2}}{\Gamma\left(\frac{K}{2}\right)} y^{K/2-1} e^{-y/2} (2Y) \mathbb{1}_{y^2 \geq 0}$$

$$= \frac{\left(\frac{1}{2}\right)^{K/2-1}}{\Gamma\left(\frac{K}{2}\right)} y^{K/2-1} e^{-y/2} \mathbb{1}_{y \geq 0} = \chi_K^2, \text{ the chi distribution with } K \text{ degrees of freedom.}$$

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$$Z \sim N(0, 1), |Z| = \sqrt{Z^2} \sim \chi_1 = \sqrt{\frac{2}{\pi}} e^{-y^2/2} \mathbb{1}_{y \geq 0}$$

$$= 2 \left( \frac{1}{\sqrt{2\pi}} e^{-y^2/2} \right) \mathbb{1}_{y \geq 0}$$

$f_Z$

Let's,  $X \sim \text{Gamma}(\alpha, \beta)$ ,  $Y = cX$ , where  $c > 0$

$$f_Y(y) = \frac{1}{c} f_X\left(\frac{y}{c}\right) = \frac{1}{c} \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{y}{c}\right)^{\alpha-1} e^{-\beta \frac{y}{c}} \mathbb{1}_{\frac{y}{c} > 0}$$

$\frac{y^{\alpha-1}}{c^{\alpha-1}}$

$$= \frac{(\beta/c)^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-(\beta/c)y} \mathbb{1}_{y > 0} = \text{Gamma}\left(\alpha, \frac{\beta}{c}\right)$$

Here,  $X \sim \chi_K^2$ ,  $Y = \frac{X}{K} \sim \text{Gamma}\left(\frac{K}{2}, \frac{1/2}{1/K}\right) = \text{Gamma}\left(\frac{K}{2}, \frac{K}{2}\right)$

Let,  $X_1 \sim \chi_{K_1}^2$ , indep of  $X_2 \sim \chi_{K_2}^2$

Let,  $U = \frac{X_1}{K_1} \sim \text{Gamma}\left(\frac{K_1}{2}, \frac{K_1}{2}\right)$  Independent  $V = \frac{X_2}{K_2} \sim \text{Gamma}\left(\frac{K_2}{2}, \frac{K_2}{2}\right)$

Formula (last note) G-

Let,  $R = \frac{U}{V} \sim f_R(r) = \int f_U(r+t) \mathbb{1}_{r+t \in \text{Supp}[U]} f_V(t) |t| dt$

$$= \int_0^\infty \frac{a^a}{\Gamma(a)} (rt)^{a-1} e^{-art} \mathbb{1}_{rt \in [0, \infty)} \frac{b^b}{\Gamma(b)} t^{b-1} e^{-bt} dt$$

$re(0, \infty)$

$$= \frac{a^a}{\Gamma(a)} \frac{b^b}{\Gamma(b)} r^{a-1} \mathbb{1}_{r=0} \int_0^\infty t^{a+b-1} e^{-(ar+b)t} dt$$

$$= a^a b^b r^{a-1} \mathbb{1}_{r \geq 0} \frac{1}{\Gamma(a)\Gamma(b)} \Gamma(a+b) \cdot \frac{1}{(ar+b)^{a+b}} = \frac{a^a b^b}{B(a,b)} r^{a-1} (ar+b)^{-(a+b)}$$

$\frac{1}{B(a,b)}$

$\frac{b^{-(a+b)}}{b^{-a}b^{-b}} \mathbb{1}_{r \geq 0} \frac{1}{(1+\frac{a}{b}r)^{(a+b)}}$



③

$$= \frac{(a/b)^a}{\Gamma(a,b)} p^{a-1} \left(1 + \frac{a}{b} p\right)^{-(a+b)} \mathbb{1}_{p \geq 0}$$

$$= \frac{(k_1/k_2)^{k_1/2}}{\Gamma(k_1/2, k_2/2)} p^{k_1/2-1} \left(1 + \frac{k_1}{k_2} p\right)^{-\frac{k_1+k_2}{2}} \mathbb{1}_{p \geq 0}$$

$= F_{k_1, k_2}$  This is the "F distribution" or "Fisher-Snedecor distribution" with  $k_1$  numerator degrees of freedom and  $k_2$  denominator degrees of freedom  $k_1, k_2 \in \mathbb{N}$

Let,  $Z \sim N(0,1)$ ,  $X \sim \chi_K^2$ ,  $Y = \frac{Z}{\sqrt{X/K}} \sim f_Y(y) = ?$

Consider,

$$Y^2 = \frac{Z^2}{X/K} \sim F_{1, K}$$

Symmetric around 0

Consider,  $W^2 = \frac{Z^2/1}{X/K} \sim F_{1, K}$

$$F_{W^2}(w^2) = P(W^2 \leq w^2) = P(W \in [-w, w]) = F_W(w) - F_W(-w)$$

Take derivatives &

$$\frac{d}{dw} [F_{W^2}(w^2)] = \frac{d}{dw} [F_W(w)] - \frac{d}{dw} [F_W(-w)]$$

$$2w f_{W^2}(w^2) = f_W(w) - f_W(-w) = 2 f_W(w)$$

$$f_W(w) = \frac{w \left(\frac{1}{K}\right)^{\frac{1}{2}-1}}{\Gamma\left(\frac{1}{2}\right) \Gamma\left(\frac{K}{2}\right)} (w^2)^{\frac{1}{2}-1} \left(1 + \frac{w^2}{K}\right)^{-\frac{1+K}{2}} \mathbb{1}_{w^2 \geq 0}$$

$$\frac{\Gamma(K+1)}{\Gamma(K) \Gamma\left(\frac{K}{2}\right)}$$

$$\frac{\Gamma\left(\frac{K+1}{2}\right)}{\Gamma(K) \Gamma\left(\frac{K}{2}\right)} \left(1 + \frac{w^2}{K}\right)^{-\frac{K+1}{2}} = T_K$$

Student's T distribution with  $K$  degree of freedom, discovered in 1908 by William Gosset while he was working at a beer factory

$$H, K \rightarrow \alpha, R_K \rightarrow Z$$

Student's  $t$  distribution has the  $N(0, 1)$  shape but just thicker tails.

$$\text{let, } z_1, z_2 \stackrel{\text{iid}}{\sim} N(0, 1), R = \frac{z_1}{z_2} \sim \int_{-\infty}^{\infty} f(ru) f(r) |u| du$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-r^2 u^2 / 2} \frac{1}{\sqrt{2\pi}} e^{-r^2 / 2} |u| du$$

$$= \frac{1}{2\pi} \left( \int_{-\infty}^0 e^{-\frac{r^2+1}{2} u^2} (-u) du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du \right)$$

$$= \frac{1}{2\pi} \left( \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du \right)$$

$$= \frac{1}{\pi} \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du$$

U Substitution:

$$\text{let, } t = u^2 \Rightarrow \frac{dt}{du} = 2u \Rightarrow du = \frac{1}{2u} dt, \quad u=0 \Rightarrow t=0, \quad u=\infty \Rightarrow t=\infty$$

$$\text{Plug in, } \int_0^{\infty} e^{-\frac{r^2+1}{2} t} \cdot u \cdot \frac{1}{2u} dt = \frac{1}{2\pi} \cdot \frac{1}{\frac{r^2+1}{2}} \int_0^{\infty} \frac{r^2+1}{2} e^{-\frac{r^2+1}{2} t} dt$$

$$= \frac{1}{\pi} \cdot \frac{1}{1+r^2} = \text{Cauchy}(0, 1)$$

$$\text{let, } X = c + \sigma R, R \sim \text{Cauchy}(0, 1), \sigma > 0, \int_0^{\infty} \lambda e^{-\lambda x} dx = 1$$

$$X \sim \text{Cauchy}(c, \sigma) = \frac{1}{\sigma\pi} \frac{1}{1 + \left(\frac{x-c}{\sigma}\right)^2}$$

Pdf experiments  
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