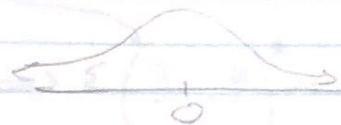


# Lecture 19

11/18/2020

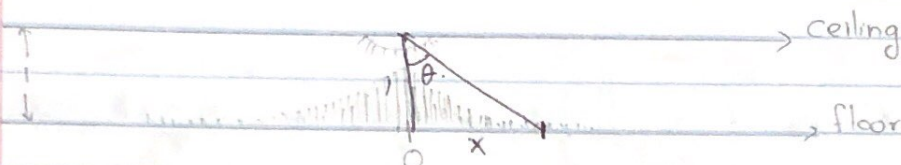
$$X \sim \text{Cauchy}(0,1) = \frac{1}{\pi} \frac{1}{x^2+1}$$

$$E[X] = \int_{\mathbb{R}} x \frac{1}{\pi} \frac{1}{x^2+1} dx = \alpha$$



$$M_X(t) = \int_{\mathbb{R}} e^{tx} \frac{1}{\pi} \frac{1}{1+x^2} dx = \alpha \text{ mgf doesn't exist.}$$

$$\phi_X(t) = \int_{\mathbb{R}} e^{itx} \frac{1}{\pi} \frac{1}{x^2+1} dx = \frac{e^{-|t|}}{|t|} \quad \text{(in fact) } \frac{1}{|t|}$$



$$\theta \sim U(-\pi/2, \pi/2) = \frac{1}{\pi} \mathbb{1}_{\theta \in [-\pi/2, \pi/2]}, \quad X = \tan(\theta) \Rightarrow \theta = \arctan(x)$$

$$\left| \frac{d}{dx} [\arctan(x)] \right| = \frac{1}{x^2+1}$$

$$f_X(x) = \int_{\theta} (\arctan(x)) \frac{1}{x^2+1} = \frac{1}{\pi} \mathbb{1}_{\arctan(x) \in [-\pi/2, \pi/2]} \frac{1}{x^2+1} = \text{Cauchy}(0,1)$$

$$\text{Let } X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$$

$$T_n \sim N(n\mu, n\sigma^2)$$

$$\bar{X}_n \sim N(\mu, \sigma^2/n) \text{ "sample mean" or "average"}$$

$$S_n^2 = \frac{1}{n-1} \sum (X_i - \bar{X})^2 \sim \chi_{n-1}^2 (s^2) = ? \text{ "sample variance"}$$

$$z_1, \dots, z_n \stackrel{iid}{\sim} N(0,1) \quad \sum z_i^2 \sim \chi_n^2 \quad \vec{z} = \begin{bmatrix} z_1 \\ \vdots \\ z_n \end{bmatrix}, \vec{z}^T \vec{z} \sim \chi_n^2$$

$$z_1 = \frac{x_1 - \mu}{\sigma}, \dots, z_n = \frac{x_n - \mu}{\sigma} \Rightarrow \sum \left( \frac{x_i - \mu}{\sigma} \right)^2 \sim \chi_n^2$$

$$(x_i - \mu)^2 = ((x_i - \bar{x}) + (\bar{x} - \mu))^2 = (x_i - \bar{x})^2 + 2(x_i - \bar{x})(\bar{x} - \mu) + (\bar{x} - \mu)^2$$

$x_i \bar{x} - \bar{x}^2 - x_i \mu + \bar{x} \mu$

$$\sum (x_i - \mu)^2 = \underbrace{\sum (x_i - \bar{x})^2}_{(n-1)S^2} + 2 \underbrace{(\bar{x} \sum x_i - n \bar{x}^2)}_{n\bar{x}} - \underbrace{\mu \sum x_i + n \bar{x} \mu}_{n\bar{x}\mu} + n(\bar{x} - \mu)^2$$

$$\sum \frac{(x_i - \mu)^2}{\sigma^2} = \underbrace{\frac{n-1}{\sigma^2} S^2}_{\text{conjecture: } \sim \chi_{n-1}^2} + \underbrace{\left( \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \right)^2}_{\sim \chi_1^2} \sim \chi_n^2$$

$\left. \begin{array}{l} U_1 \sim \chi_{k_1}^2, \text{ indep of } \\ U_2 \sim \chi_{k_2}^2 \end{array} \right\} \Rightarrow U_1 + U_2 \sim \chi_{k_1+k_2}^2$

this would be true if  $\bar{x}$  is independent of  $S^2$ .

"quadratic form"

$$\vec{z}^T \vec{z} = \vec{z}^T \mathbf{I} \vec{z} \sim \chi_n^2$$

$$\text{consider: } \vec{z}^T \begin{bmatrix} 1 & 0 & \dots & 0 \\ 0 & 0 & & \\ \vdots & & \ddots & \\ 0 & & & 0 \end{bmatrix} \vec{z} = z_1^2 \sim \chi_1^2$$

$$\vec{z}^T \begin{bmatrix} 0 & & & 0 \\ & \ddots & & \\ & & 0 & 0 \\ 0 & & & 0 \end{bmatrix} \vec{z} = z_i^2 \sim \chi_1^2 \quad \text{ranks}[B_i] = 1$$

$$\vec{z}^T \mathbf{I} \vec{z} = \vec{z}^T (B_1 + B_2 + \dots + B_n) \vec{z} = \vec{z}^T B_1 \vec{z} + \dots + \vec{z}^T B_n \vec{z} \sim \chi_n^2$$



Cochran's thm: Let  $B_1 + B_2 + \dots + B_k = I$  s.t

$$\sum_{j=1}^k \text{rank}[B_j] = n \quad \text{then (a) } \bar{Z}^T B_j \bar{Z} \sim \chi^2_{\text{rank}[B_j]}$$

and (b)  $\bar{Z}^T B_{j_1} \bar{Z}$  is independent of  $\bar{Z}^T B_{j_2} \bar{Z} \quad \forall j_1 \neq j_2$

$$\begin{aligned} \bar{Z}^T \bar{Z} &= \sum z_i^2 = \sum ((z_i - \bar{z}) + \bar{z})^2 \quad \left\{ \begin{array}{l} \sum (z_i - \bar{z}) \bar{z} \\ = n\bar{z}^2 - n\bar{z}^2 = 0 \end{array} \right. \\ &= \sum (z_i - \bar{z})^2 + 2 \sum (z_i - \bar{z}) \bar{z} + n\bar{z}^2 = \sum (z_i - \bar{z})^2 + n\bar{z}^2 \end{aligned}$$

Let  $\bar{1}_n$  is a column vector of all ones  $\Rightarrow \bar{Z} = \frac{1}{n} \bar{Z}^T \bar{1}_n$

$$n\bar{Z}^2 = n \left( \frac{1}{n} \bar{Z}^T \bar{1}_n \right)^2 = \frac{1}{n} \bar{Z}^T \bar{1}_n \frac{1}{n} \bar{Z}^T \bar{1}_n = \bar{Z}^T \left( \frac{1}{n} \bar{1}_n \bar{1}_n^T \right) \bar{Z}$$

$\text{rank}[B_2] = 1 \quad B_2 = \frac{1}{n} J_n$

Let  $J_n = \bar{1}_n \bar{1}_n^T$  which is an  $n \times n$  matrix of all entries =

$$\sum (z_i - \bar{z})^2 = \sum z_i^2 - 2 \sum z_i \bar{z} + n \bar{z}^2 = \sum z_i^2 - n \bar{z}^2 \cdot B_1 + B_2 =$$

$$= \bar{Z}^T I \bar{Z} - \bar{Z}^T \left( \frac{1}{n} J_n \right) \bar{Z} = \bar{Z}^T \left( \underbrace{I - \frac{1}{n} J_n}_{B_1} \right) \bar{Z}$$

$(I - \frac{1}{n} J) + \frac{1}{n} J = I$

Thm from math 231: If  $A$  is symmetric matrix and idempotent which means  $AA = A$  then  $\text{rank}[A] = \text{tr}[A] = \text{sum of the diagonal of } A$

$$\begin{bmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix} \begin{bmatrix} 1 & 1 & 1 \\ 1 & 1 & 1 \\ 1 & 1 & 1 \end{bmatrix} = \begin{bmatrix} 3 & 3 & 3 \\ 3 & 3 & 3 \\ 3 & 3 & 3 \end{bmatrix}$$

$$B_1^T = (I - \frac{1}{n} J)^T = I^T - \frac{1}{n} J^T = I - \frac{1}{n} J = B_1 \quad \checkmark$$

$$B_1 B_1 = (I - \frac{1}{n} J) (I - \frac{1}{n} J) = I [I - \frac{1}{n} J] - \frac{1}{n} J [I - \frac{1}{n} J]$$

$$= I - \frac{2}{n} J + \frac{1}{n} J = I - \frac{1}{n} J = B_1 \quad \checkmark \Rightarrow \text{rank}[B_1] = \text{tr}[B_1]$$

$$B_{i,i} = 1 - \frac{1}{n} = 1 - \frac{1}{n} = \sum_{b=1}^n 1 - \frac{1}{n} = n-1$$

Putting it all together, we can use Cochran's thm!

$$\sum (z_i - \bar{z})^2 \sim \chi_{n-1}^2 \text{ indep of } n\bar{z}^2 \sim \chi_1^2$$

$$\begin{aligned} \bar{z} &= \frac{z_1 + \dots + z_n}{n} = \frac{\frac{x_1 - \mu}{\sigma} + \dots + \frac{x_n - \mu}{\sigma}}{n} = \frac{\sum x_i - n\mu}{n\sigma} \\ &= \frac{\bar{x} - \mu}{\sigma} \end{aligned}$$

$$\Rightarrow n\bar{z}^2 = (n\bar{z})^2 = \left(\frac{\bar{z}}{1/\sqrt{n}}\right)^2 \stackrel{\text{above}}{=} \left(\frac{\bar{x} - \mu}{\sigma/\sqrt{n}}\right)^2 \sim \chi_1^2$$

$$\begin{aligned} \sum (z_i - \bar{z})^2 &= \sum \left( \frac{x_i - \mu}{\sigma} - \frac{\bar{x} - \mu}{\sigma} \right)^2 = \frac{1}{\sigma^2} \sum (x_i - \bar{x})^2 \\ &\text{independent of} \\ &\downarrow \\ &= \frac{n-1}{\sigma^2} s^2 \sim \chi_{n-1}^2 \end{aligned}$$

I think the first to prove this was Fisher in 1925 and then in 1936, Geary proved the iid normal rv is the \*only\* distribution that has the independent of  $\bar{x}$  and  $s^2$ .

$$\frac{\bar{x} - \mu}{\sigma/\sqrt{n}} \sim N(0,1) \text{ but what about } \frac{\bar{x} - \mu}{s/\sqrt{n}} \sim ? \text{ Not } N(0,1)$$