$$f_{X_{1}}(x) = \int ... \int f_{X_{1},X_{2},...,X_{n}}(x,u_{1},u_{2},...,u_{n-1}) du_{1}...du_{n-1}$$

$$\frac{f_{X_{1}}(x)}{f_{X_{1},X_{2},...,X_{n}}(x,u_{1},u_{2},...,u_{n-1}) du_{n}} du_{1}...du_{n-1}$$

$$\frac{f_{X_{1}}(x)}{f_{X_{1},X_{2},...,X_{n}}(x,u_{1},u_{2},...,u_{n-1}) du_{n}} du_{1}...du_{n-1}$$

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$$\frac{f_{X_{1},X_{2},...,X_{n}}(x,u_{1},u_{2},...,u_{n-1}) du_{n}}{f_{X_{1},X_{2},...,X_{n}}(x,u_{1},u_{2},...,u_{n-1}) du_{n}} du_{n}...du_{n-1}$$

if
$$x < q$$
 $q \cdot 1_{x \ge q} = q(p) = 0 \le x$ being $Syp[X] \ge 0$.

if $x \ge q$ $q \cdot 1_{x \ge q} = q(0) = q \le x$ then by case assignation.

$$q \cdot 1_{X \ge q} \le X$$
Let's take the expectation of both sides:

Let's take the expectation of both sides:
$$E[\P_{X \geq q}] \leq E[X] \qquad \qquad I_{X \geq q} \qquad \qquad I_{X$$

this is called Markov's Inequality and its very famous AKA: Markov's Tail Bound.

For emple
$$X \sim \text{Exp}(1) = e^{-x} \Rightarrow P(X \neq 1) = |-F_X(1)| = e^{-x} \Rightarrow A = |$$

We see here is that the Markov Bound is very "crude" meaning

We will now prove many, many corollaries of the Markov Inequality:

Let
$$b = a_M$$
 $P(X \ge b) \le \frac{a_M}{b} \Rightarrow P(X \ge a_M) \le \frac{1}{4}$

Let h be a monotonically increasing function $Y = h(X)$
 $P(X \ge b_M) \le \frac{a_M}{b} \Rightarrow P(X \ge a_M) = \frac{1}{4}$

$$P(Y \ge h(n)) \le \frac{E(Y)}{h(n)} \Rightarrow P(h(x) \ge h(n)) \le \frac{E(h(x))}{h(n)} \Rightarrow P(Y \ge n) \le \frac{E(h(x))}{h(n)}$$
• Let X be continuous in addition to nonnegative.

$$let \ q = Quentile[X, p] = F_X(p)$$

$$P(X \ge F_X(p)) \le \frac{h}{F_X(p)} \Rightarrow |-F_X(p)| \le \frac{h}{F_X(p)}$$

$$\Rightarrow |-p \le \frac{h}{F_X(p)}| \Rightarrow F_X(p) \le \frac{h}{F_X(p)} \Rightarrow P(x \ge h(x)) = 0$$

$$| - \rho | \leq \frac{\pi}{F_{x}^{-1}(p)} \Rightarrow \frac{F_{x}^{-1}(p)}{|x|} \leq \frac{\pi}{1-p} \quad \text{e.g.} \quad \text{Mea}[X] \leq 2\pi$$

$$\mathbb{Q}[X, p]$$
• Let X be any rv => |X| is a nonnegative rv.
$$\mathbb{P}[X| \geq q) \leq \frac{\mathbb{E}[|X|]}{q}.$$

Let X be any rv with finite variance,
$$G^2$$
. Let $Y = (X - mu)^2 = Y$ is a nonnegative rv.

$$P(Y \ge b) \le \frac{|E[Y]|}{b} \Rightarrow P((X - m)^2 \ge b) \le \frac{|E[X]| - |A|}{b} \Leftrightarrow P((X - m)^2 \ge b) \le \frac{|E[X]|}{b} \Leftrightarrow P((X - m)^2 \ge a^2) \le \frac{|E[X]|}{a^2} \Leftrightarrow P((X - m)^2 \ge a^2) \le \frac{|E[X]|}{a^2} \Rightarrow P((X - m)^2 \ge a^2) \ge \frac{|E[X]$$

Let's manipulate this to get it into a more "user-friendly" form. Assume X is nonnegative:
$$\frac{disjoint coms}{disjoint coms} = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n)$$

$$= P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) + P(X-n \ge n) = P(X-n \ge n) + P(X$$

$$P(X \ge M+1) + P(X \le M-1) \stackrel{\text{REM}}{=} P(X \ge M+1) + P(X \le M \ge M)$$

$$P(X \ge M+1) + P(X \le M \ge M) \stackrel{\text{REM}}{=} P(X \ge M+1) + P(X \le M \ge M)$$

$$P(X \ge M) \le \frac{\sigma^2}{(b-M)^2}$$

Let X be an rv. Let $Y = e^{-\frac{a}{b}} > Y$ is a nonnegative $P(Y \ge b) \le \frac{\mathbb{E}^{[X]}}{b} \Rightarrow P(e^{tX} \ge b) \le \frac{\mathbb{E}^{[E^{tX}]}}{b} \leftarrow \text{the moment generating function for }$ $\Rightarrow P(e^{tX} \ge b) \le \frac{M_X(e)}{b} \Rightarrow P(e^{tX} \ge e^{tx}) \le e^{-tx} M_X(e)$

$$P(X \ge 1) \le e^{-t_1} P_X(t)$$

$$P(X \le 1) \le P(X \le 1) \le P_X(t)$$

$$P(X \le 1) \ge P_X(t)$$

$$P(X \le 1) \le P_X(t)$$

$$P(X \le 1) \le P_X(t)$$

$$P(X \le 1) \ge P_X(t)$$

$$P(X \ge 1) \ge P_X(t)$$

We first need to find the mgf for the exponential rv: $howevert_{\chi}(t) = \left[\int e^{tX} \right] = \int e^{tx} \lambda e^{-\lambda x} dx = \lambda \int e^{(t-\lambda)x} dx = \frac{\lambda}{t-\lambda} \left[e^{(t-\lambda)x} \right]^{\alpha}$ $=\frac{\lambda}{\xi-\lambda} \begin{cases} \infty - 1 & \text{if } \xi > \lambda \\ 0 - 1 & \text{if } \xi < \lambda \end{cases} = \frac{\lambda}{\lambda - t} \text{ only for } \xi < \lambda.$

Let's calculate it for $X \sim Exp(lambda)$. Warning: it's a lot of work...

If t > lambda, the mgf doesn't exist. This is why you shouldn't be using them! Chf's always exist! So they're awesome... $X \sim \mathbb{E} \times p(1) \Rightarrow M_X(t) = \frac{1}{1-t} \text{ for } t < 1.$ $P(X > a) \leq \min_{t>0} \left\{ e^{-ta} \frac{1}{1-t} \right\}$ for t < 1

 $\Rightarrow \mathcal{C}(X > A) \leq \min_{t \in \{0,1\}} \begin{cases} e^{-t_1} \frac{1}{1-t} \end{cases} = e^{-\left(\frac{1}{t}\right)1} \frac{1}{1-\left(\frac{1}{t}\right)} = \frac{e^{-1}e}{\frac{1}{t}} = \frac{4e}{e^{-1}}$ $h'(t) = \frac{(1-t)(-1)e^{-t} - e^{-t}(-1)}{(1-t)^2} = \frac{t-1) \cdot e^{-t} + e^{-t}}{(1-t)^2}$ $= e^{-t\eta \left(t\eta - \eta + 1\right)} \stackrel{\text{Set}}{=} 0 \Rightarrow t\eta - \eta + 1 = 0 \Rightarrow t_v = \frac{\eta + 1}{2} = 0$ But... the Chernoff is sometimes useless. Why? Because it requires the mgf. To get the mgf, you need to kniw the PDF or PMF. If I know the PDF or PMF, then I know analytically or can numerically compute the CDF which means I know the tail exactly or within small numerical error! So it really is only useful if you're in a situation where you only have the MGF and not the PDF/PMF.