

Lecture 21

11/25/20

Math 621

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$$\phi_{\vec{x}}(\vec{t}) := E[e^{i\vec{t}^T \vec{x}}]$$

Consider a vector rv X w/ dimension n . Consider the following operation:

$$\phi_{\vec{x}}\left(\begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix}\right) := E[e^{i[1 \ 0 \dots 0] \vec{x}}] = E[e^{i x_1}] = \phi_{x_1}(t) \Rightarrow$$

$$x_1 \sim f_{x_1}(x)$$

The bottom line is we can use multivariable chf's to immediately get marginal distributions.

$$\begin{aligned} \vec{X} \sim N(\vec{\mu}, \Sigma) &\Rightarrow \phi_{\vec{x}}\left(\begin{bmatrix} t \\ 0 \\ \vdots \\ 0 \end{bmatrix}\right) = e^{i[t \ 0 \dots 0] \vec{\mu} - \frac{1}{2}[t \ 0 \dots 0] \Sigma \begin{bmatrix} t \\ 0 \\ \vdots \\ 0 \end{bmatrix}} \\ &= e^{i t \mu_1 - \frac{t^2}{2} [1 \ 0 \dots 0] \begin{bmatrix} \sigma_{11} \\ \sigma_{21} \\ \vdots \\ \sigma_{n1} \end{bmatrix} \begin{bmatrix} 1 \\ 0 \\ \vdots \\ 0 \end{bmatrix}} \rightarrow 1 \end{aligned}$$

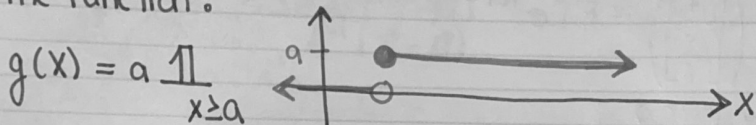
$$= e^{itM_1 - t^2 \frac{\sigma^2}{2}}$$

$$\Rightarrow X_1 \sim N(\mu, \sigma^2)$$

(p1)

We now begin the unit on the "pure math" part of prob. beginning w/ famous inequality.

Let X be a rv w/ non-neg. support i.e. $\text{Supp}[X] \geq 0$. Let a be a cts. > 0 . Consider the function:



Is $a \cdot 1_{x \geq a} \leq x \quad \forall x$?

• if $x < a \rightarrow a \cdot 1_{x \geq a} = 0 \leq x$ b/cuz $\text{Supp}[X] \geq 0$. ✓

• if $x \geq a \rightarrow a \cdot 1_{x \geq a} = a \leq x$ b/cuz we assume $x \geq a$. ✓

$$\Rightarrow a \cdot 1_{x \geq a} \leq x$$

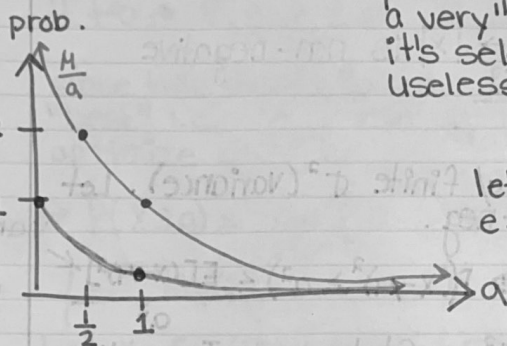
Now let's take the expectation of both sides:

$$E[a \cdot 1_{x \geq a}] \leq E[X] \Rightarrow a E[1_{x \geq a}] \leq \mu \Rightarrow a P(X \geq a) \leq \mu$$

this rv has $\text{Supp}\{0, 1\} \Rightarrow$ it's a bern. 2p.

$\Rightarrow P(X \geq a) \leq \frac{\mu}{a}$ this is called a Markov's inequality

It's a tail bdd. b/cuz it gives you an upper bdd. on what the probability of the "tail" is. It is a very "crude" bdd. which means it's seldom so useful & useless if $a < \mu$.



let $\mu = 1 = E[X]$

e.g. $X \sim \text{Exp}(1)$

$$P(X \geq a) = 1 - F_X(a) = e^{-a}$$

$$a=1 \Rightarrow P(X \geq a) = \frac{1}{e} \approx .37$$

a	$P(X \geq a)$	Markov Bdd.	Chebyshev	Chernoff
2	.1353	.5	.1	.73576
5	.0067	.2	.0635	.09158
10	.00004	.1	.012	.00123

The Markov inequalities has tons of collaries:

• let $b = a\mu \rightarrow P(X \geq b) \leq \frac{\mu}{b} \rightarrow P(X \geq a\mu) \leq \frac{1}{a}$

• let $h(x)$ be a monotonically increasing fn. (so it's 1:1)
 $P(h(x) \geq h(a)) \leq \frac{E[h(x)]}{h(a)} \rightarrow P(X \geq a) \leq \frac{E[h(x)]}{h(a)}$
 if x cts.

• let $a = \text{Quantile}[X, p] = F_X^{-1}(p)$

$$P(X \geq a) \leq \frac{M}{a} \rightarrow 1 - F_X(a) \leq \frac{M}{a} \Rightarrow (0 \leq X) \Leftarrow$$

$$1 - F_X(F_X^{-1}(p)) \leq \frac{M}{F_X^{-1}(p)} \Rightarrow 1 - p \leq \frac{M}{F_X^{-1}(p)} \Rightarrow F_X^{-1}(p) \leq \frac{M}{1-p}.$$

$$\text{e.g. Med}[X] \leq 2M$$

• Let X be any r.v. $\rightarrow |X|$ is non-negative
 $\rightarrow P(|X| \geq a) \leq \frac{E[|X|]}{a}$

• Let X be any r.v. w/ finite σ^2 (variance). Let $Y = (X - \mu)^2 \Rightarrow Y$ is non-neg.

$$P(Y \geq a^2) \leq \frac{E[Y]}{a^2} \Rightarrow P((X - \mu)^2 \geq a^2) \leq \frac{E[(X - \mu)^2]}{a^2} \leftarrow \text{variance}$$

$$\Rightarrow P(|X - \mu| \geq a) \leq \frac{\sigma^2}{a^2} \quad \text{Chebyshev's Inequality.}$$

So let's assume X is nonneg. & let's get this bdd. in a more user-friendly form.

$$P(|X - \mu| \geq a) = P(X - \mu \geq a \cup -(X - \mu) \geq a) = P(X - \mu \geq a) + P(-(X - \mu) \geq a) \\ = P(X - \mu \geq a) + P(X \leq \mu - a) \rightarrow \text{second term is zero if } a \geq \mu \text{ since } X \text{ is assumed non-neg.}$$

$$\Rightarrow P(|X - \mu| \geq a) = P(X - \mu \geq a) = P(X \geq a + \mu) \leq \frac{\sigma^2}{a^2}$$

$$\Rightarrow P(X \geq b) \leq \frac{\sigma^2}{(b - \mu)^2}$$

$$\text{let } b = \mu + a.$$

• Let X be any rv & $Y = e^{tX} \Rightarrow Y$ is non-neg. $\forall t$. mgf.
 $\Rightarrow P(Y \geq c) \leq \frac{E[Y]}{c} \Rightarrow P(e^{tX} \geq c) \leq \frac{E[e^{tX}]}{c}$

$$\begin{aligned} \text{let } c &= e^{ta} \\ \Rightarrow P(e^{tx} \geq e^{ta}) &\leq e^{-ta} M_X(t) \\ \Rightarrow P(tx \geq ta) &\leq e^{-ta} M_X(t) \end{aligned}$$

$$\text{if } t > 0 \rightarrow P(X \geq a) \leq e^{-ta} M_X(t) \quad \forall t > 0$$

$$\text{if } t < 0 \rightarrow P(X \leq a) \leq e^{-ta} M_X(t)$$

since this works for all t & we are looking for the "best" i.e. the lowest upper bdd., then just optimize over t :

$$\Rightarrow P(X \geq a) \leq \min_{t > 0} \{ e^{-ta} M_X(t) \} \quad \underline{\text{AND}} \quad P(X \leq a) \leq M_X(t)_{t < 0}$$

$$\begin{aligned} \text{Let } X \sim \text{Exp}(\lambda) \rightarrow M_X(t) &:= E[e^{tx}] = \int_0^\infty e^{tx} \lambda e^{-\lambda x} dx = \\ \lambda \int_0^\infty e^{(t-\lambda)x} dx &= \lambda \frac{1}{t-\lambda} \left[e^{(t-\lambda)x} \right]_0^\infty = \frac{\lambda}{t-\lambda} \begin{cases} \infty - 1 & \text{if } t > \lambda \\ 0 - 1 & \text{if } t < \lambda \end{cases} \\ &= \frac{\lambda}{\lambda - t} \quad \text{for } t < \lambda \quad \text{otherwise the mgf does not exist!} \end{aligned}$$

For $X \sim \text{Exp}(1)$, the chernoff bdd is.....

$$P(X \geq a) \leq \min_{t > 0} \{ e^{-ta} M_X(t) \} = \min_{t > 0} \left\{ e^{-ta} \frac{1}{1-t} \right\} \quad \text{if } t < 1$$

$$= \min_{t \in [0, 1)} \left\{ \frac{e^{-ta}}{1-t} \right\} = \frac{ae}{e^a}$$

$$h'(t) = \frac{h(t) \cdot (-a) e^{-ta} - (e^{-ta})(-1)}{(1-t)^2} =$$

→

$$\frac{e^{-ta}}{(1-t)^2} \quad \begin{matrix} \text{set} \\ \downarrow \\ = 0 \end{matrix}$$

$$\rightarrow a + -a + 1 = 0 \rightarrow t_* = \frac{a-1}{a} = 1 - \frac{1}{a} \in (0, 1) \text{ if } a > 1.$$

Let me tell you why the Chernoff bdd. is "seldom" useful. It requires the MGF. The MGF means you have the $f(x)/p(x)$ and if you know these you may have $F(x)$; which means you can calculate tail probabilities explicitly using numerical integration.