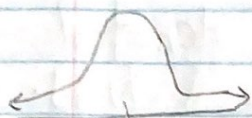


III

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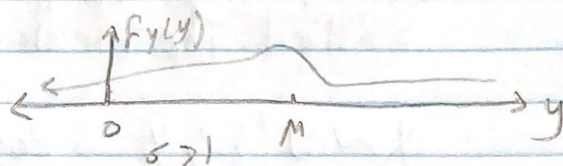
$X \sim \text{Logistic}(0,1) \approx N(0,1)$  except it has thicker tails



standard logistic  $E(X)=0, SD[X]=\frac{\pi}{\sqrt{3}}$

$$Y = \mu + \sigma X \sim \text{Logistic}(\mu, \sigma) := F_Y(y) = \frac{1}{\sigma} \frac{e^{\frac{y-\mu}{\sigma}}}{(e^{\frac{y-\mu}{\sigma}} + 1)^2}$$

$\mu \in \mathbb{R}, \sigma > 0$



Why is this called the "logistic distribution"?

There's a function called the "logistic function" and it has 3 parameters = L (Maximum Value), K (steepness), mu (center)

$$L(x) = \frac{L}{1 - e^{-K(x-\mu)}} = \frac{1}{1 + e^{-x}} \quad \text{the standard logistic function}$$

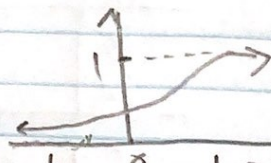
if  $L=1, K=1, \mu=0$

$$X \sim \text{Logistic}(0,1) = \frac{e^y}{(1+e^y)^2}$$

$$F_Y(y) = P(Y \leq y) = \int_{-\infty}^y \frac{e^t}{(1+e^t)^2} dt = \int_1^{1+e^y} \frac{u-1}{u^2} \frac{1}{u-1} du = \left[ -\frac{1}{u} \right]_1^{1+e^y} =$$

let  $u = 1 + e^t \Rightarrow \frac{du}{dt} = e^t \Rightarrow dt = \frac{1}{e^t} du = \frac{1}{u-1} du, t = -\infty \Rightarrow u = 1, t = y \Rightarrow u = 1 + e^y$

$$1 - \frac{1}{1+e^y} = \frac{e^y}{1+e^y}$$



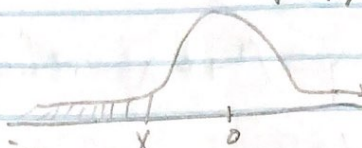


The  $q$ th "quantile" or  $100q$  "percentile" of a rv  $X$ .

distribution. Definition:

Solve for  $x$  where  $q \geq P(X \leq x) = F_X(x)$  denoted  $Q(X, q)$ .

I want the 33<sup>rd</sup> quantile



\*If  $q = 0.5$ , that quantile is called the "median",  $\text{Med}[X] = t$

$$X \sim U(\{2, 4, \dots, 20\})$$

$x$	$P(X)$	$F(x)$
2	0.1	0.1
4	0.1	0.2
6		0.3
8		0.4
10		0.5
12		0.6
14		0.7
16		0.8
18		0.9
20		1.0

$$Q(X, 0.3) = 6$$

$$Q(X, 0.9) = 18$$

$$Q(X, 0.85) = 16 \neq F(x) = 0.85$$

If  $X$  is a continuous RV with "contiguous support"

i.e. one interval with no gaps e.g.  $[0, 10]$ , the

real numbers but not e.g.  $[0, 1] \cup [2, 3]$

where there is a gap between 1 and 2, then  $F(x)$  is strictly increasing thus invertible and the minimum  $x$  s.t.  $q \geq F(x)$  would be



$$F^{-1}(q) \geq X \Rightarrow X = \underline{F^{-1}(q)} = \Phi[X, q]$$

$$X \sim \text{EXP}(\lambda) = \lambda e^{-\lambda x} \Rightarrow F(X) = 1 - e^{-\lambda x} \text{ find the } q \text{ function}$$

$$F^{-1}(q) \\ q = 1 - e^{-\lambda x} \Rightarrow 1 - q = e^{-\lambda x} \Rightarrow \ln(1 - q) = -\lambda x \Rightarrow x = -\frac{1}{\lambda} \ln(1 - q)$$

$$\Rightarrow x = \frac{1}{\lambda} \ln\left(\frac{1}{1 - q}\right) = F^{-1}(q)$$

$$X \sim \text{EXP}(1) \Rightarrow \text{Med}[X] = F^{-1}(0.5) = \ln(2)$$

$$\Phi[X, 0.8] = \ln(5)$$

It's actually rare to have a quantile function in closed form since it's rare to even have a CDF in closed form e.g.

$X \sim \text{Erlang}(k, \lambda)$ ,  $F(X) = P(k, \lambda x) = \Phi[X, q]$  can be found by solving for  $X$  in the following equation:  $q = P(k, \lambda x)$

$$X \sim \text{EXP}(\lambda), Y = Ke^{X = g(X)}, K > 0 \text{ find } F_Y(y).$$

$$\frac{y}{K} = e^x \Rightarrow x = \ln\left(\frac{y}{K}\right) = \ln(y) - \ln(K) = g^{-1}(y),$$

since  $y$  is always p.o.s.

$$\left| \frac{d}{dy} [g^{-1}(y)] \right| = \frac{1}{|y|} \stackrel{!}{=} \frac{1}{y}$$

$$F_Y(y) = F_X(g^{-1}(y)) \left| \frac{d}{dy} [g^{-1}(y)] \right| = \lambda e^{-\lambda \ln(y/K)} \frac{1}{y} \mathbb{1}_{\ln(y) - \ln(K) \in [0, \infty)} \\ = \frac{\lambda}{y} e^{\ln(y/K)} \mathbb{1}_{\ln(y) \in [\ln(K), \infty)} = \frac{\lambda}{y} \left(\frac{y}{K}\right)^{-\lambda} \mathbb{1}_{y \in [K, \infty)}$$



$$\Rightarrow \text{Pareto } I(k, n) \\ k \in (0, \infty), n \in (0, \infty)$$

$$\frac{1}{x^5} \rightarrow \frac{-1}{4x^4}$$

$$F_Y(y) = \int_k^y \frac{n k^n}{t^{n+1}} dt = n k^n \left[ -\frac{1}{n t^n} \right]_k^y = k^n \left( \frac{1}{k^n} - \frac{1}{y^n} \right) \\ = 1 - \left( \frac{k}{y} \right)^n = q$$

$$F_Y^{-1}(q) = k(1-q)^{-\frac{1}{n}}$$

$$X, Y \stackrel{iid}{\sim} \text{Exp}(1), \text{ let } D = X - Y = X + (-Y) \\ \stackrel{||}{=} e^{-X} 1_{X \in [0, \infty)}$$

$$Z \sim f_Z(z) = e^z 1_{z \in (-\infty, 0]}$$

$$f_D(d) = \int_{\text{supp}(X)} f_X^{old}(x) f_Z^{old}(d-x) 1_{d-x \in \text{supp}(Z)} dx$$

$$= \int_0^\infty e^{-x} e^{d-x} 1_{\substack{x \in [0, \infty) \\ d-x \in (-\infty, 0]}} dx = e^d \int_0^\infty e^{-2x} 1_{x \in [d, \infty)} dx$$

$$= e^d \begin{cases} \int_d^\infty e^{-2x} dx & \text{if } d \geq 0 \\ \int_0^\infty e^{-2x} dx & \text{if } d < 0 \end{cases} = e^d \begin{cases} \left[ -\frac{1}{2} e^{-2x} \right]_d^\infty & \text{if } d \geq 0 \\ \left[ -\frac{1}{2} e^{-2x} \right]_0^\infty & \text{if } d < 0 \end{cases}$$

$$= \frac{1}{2} e^d \begin{cases} e^{-2d} & \text{if } d \geq 0 \\ 1 & \text{if } d < 0 \end{cases} = \begin{cases} \frac{1}{2} e^{-d} & \text{if } d \geq 0 \\ \frac{1}{2} e^d & \text{if } d < 0 \end{cases}$$

$$= \frac{1}{2} e^{-|d|}$$

Laplace(0,1)  
Standard Laplace

MCB, 670

$$X = \mu + \sigma D \sim \text{LaPlace}(\mu, \sigma) := \frac{1}{2\sigma} e^{-\frac{|x-\mu|}{\sigma}}$$

this is also famous RV and it has another name: the "double exponential".

LaPlace published this distribution in 1774 calling it the "first law of errors".

