$$X \sim \text{Multi}_{2}(n, \hat{p})$$

$$\overline{b}_{k=2}$$

$$P_{x,|x_{1}}(x_{1},x_{1}) := P(X_{1}=x_{1}|X_{1}=x_{1}) = \frac{P(x_{1},x_{1})}{P(x_{1})}$$

$$Deg(n-x_{1})$$
Last time $p(x_{1}) = Bin(n, p_{1}) = Bin(n, 1-p_{1})$

$$A = \frac{(x_{1},x_{1})}{(x_{1})} \frac{p^{x_{1}}}{p^{x_{1}}}$$
Define $J_{n} := \{0,1,$

Let n'= n-x;

 $= \frac{\begin{pmatrix} x^r \\ 0 \end{pmatrix} b_{x^r}^1 \begin{pmatrix} 1 - b^r \\ 0 \end{pmatrix} b_{x^{-r}}^2}{\begin{pmatrix} x^r \\ 0 \end{pmatrix} b_{x^{-r}}^1 b_{x^{-r}}^2}$ $=\frac{\frac{\lambda^{4}}{x^{\prime}}\frac{1}{x^{\prime}}\frac{$

 $=\underbrace{\frac{x' \cdot u \cdot x'}{X' \cdot i}}_{\text{$(u-x')_i$}} \mathbb{T}^{x' = u \cdot x'} \underbrace{\mathbb{T}^{x' \cdot e_2 u}}_{\text{$(u-x')$}} \underbrace{\mathbb{T}^{x' \cdot e_2 u}}_{\text{$(u-x')$}} \underbrace{\mathbb{T}^{x' \cdot e_2 u}}_{\text{$(u-x')$}} \mathbb{T}^{x' \cdot e_2 u}_{\text{$(u-x')$}} = \text{Ded}(u-x') \mathbb{T}^{x' \cdot e_2 u}_{\text{$(u-x')$}}$

Define Jn:= {0,1,...,n3

 $\rho(A|B) = \frac{\rho(A,B)}{\rho(B)}$; if $\rho(B) = 0 \Rightarrow \rho(A|B)$ is undefined

 $= \frac{\frac{\alpha!}{x_1! \cdots x_j! \cdots x_k!} \underbrace{1}_{x_i + \cdots + x_j + 1 + x_k = n} \underbrace{1}_{x_i + \cdots + x_j + \cdots + x_k = n} \underbrace{1}_{x_i \in S_n} \cdots \underbrace{1}_{x_i \in S_n} \cdots \underbrace{1}_{x_k \in S_n} \rho_i^{X_1} \cdots \rho_i^{X_j} \cdots \rho_k^{X_k}}_{x_j \in T_n} \underbrace{1}_{x_j \in T_n} \underbrace{1}_{x_j \in T_n} \underbrace{1}_{x_j \in S_n} \cdots \underbrace{1}_{x_k \in S_n} \rho_i^{X_1} \cdots \rho_i^{X_k} \cdots \rho_k^{X_k}$

Note $p_i+...+p_{\kappa}=1=$ $p_i+...+p_{i-1}+p_{i+1}+...+p_{\kappa}=1-p_i$, divide both sides by $1-p_i$

 $= > \frac{\rho_i}{1-\rho_i} + \dots + \frac{\rho_{i-1}}{1-\rho_i} + \frac{\rho_{i-1}}{1-\rho_i} + \dots + \frac{\rho_k}{1-\rho_i} = |$

Let's generalize this conditional probability a little bit:

 $\vec{X} \sim \text{Multi}_{k}(n_{i}\vec{\rho})$ $\rho_{\underline{x}_{i}|X_{i}}(\vec{x}_{\cdot i}, x_{i}) = \frac{\rho_{x_{i}}(x_{i})}{\rho_{x_{i}}(x_{i})} = \text{Multi}_{k-i}(n_{\cdot}-x_{i},?)$

this is the vector component w/o the jth component.

 $=\frac{\mathsf{Multi}_{k}(\mathsf{n}_{1}\widehat{\mathsf{p}})}{\mathsf{Bin}\left(\mathsf{n}_{1}\mathsf{p}_{1}\right)}=\frac{\left(\mathsf{x}_{1},\ldots,\mathsf{x}_{k-1}^{\mathsf{n}_{1}},\mathsf{x}_{k}\right)\mathsf{p}_{1}^{\mathsf{x}_{1}}\ldots\mathsf{p}_{k}^{\mathsf{x}_{k}}}{\left(\mathsf{n}_{1}^{\mathsf{n}_{1}}\right)\mathsf{p}_{1}^{\mathsf{x}_{1}}\left(\mathsf{l}-\mathsf{p}_{1}^{\mathsf{n}_{1}}\right)^{\mathsf{n}^{\mathsf{x}_{2}}}}$

Note: n-x; = x,+... +xj-1+xj+1+..+xx 0/+ prob. O.

= Multix-1 (n' p') 1 x: 630

 $=\frac{n'!}{x_{*}!\cdots x_{j+1}! x_{j+1}!\cdots x_{k}!} \frac{1}{1}_{x_{j+1}+x_{j+1}+x_{j+1}+x_{k}=n'} 1_{x_{i}\in S_{n}}\cdots 1_{x_{i}\in S_{n}} 1_{x_{j}\in S_{n}}\cdots 1_{x_{k}\in S_{n}$

Define: $I_A^{"} = \frac{I_A}{I_A} = \begin{cases} 1 & \text{if } A \\ \text{undefined if } A^c \end{cases}$

$$\begin{split} \sigma^2 &= \text{Var}[X] := \text{E}[\{X - \mu\}^2], \ \sigma := \text{SD}[X] := \{\text{Var}[X] \\ &= \text{E}[X^2] - \mu^2 \\ \text{Standard deviation} \end{split}$$

$$\begin{aligned} &\text{Vor}[X, + X_*] := \text{E}[\{(X, + X_*) - (\mu_* + \mu_*)\}^2] \\ &= \text{E}[X^1 + X_*^2 + \mu_*^2 + \mu_*^2 + 2X \cdot X_* - 2X \cdot \mu_* - 2X \cdot \mu_* - 2X \cdot \mu_* - 2X \cdot \mu_* + 2\mu_* \mu_*] \\ &= \text{E}[X^1] + \text{E}[X^1] + \mu_*^2 + \mu_*^2 + 2\text{E}[X \cdot X_*] - 2\mu_* - 2\mu_* \mu_* - 2\mu_*^2 + 2\mu_* \mu_* \\ &= \sigma_*^1 + \sigma_*^2 + \sigma_*^2 + 2\mu_* + \mu_*^2 + 2\text{E}[X \cdot X_*] - 2\mu_*^2 - 2\mu_* \mu_* - 2\mu_*^2 + 2\mu_* \mu_* \\ &= \sigma_*^2 + \sigma_*^2 + 2\left(\text{E}[X \cdot X_*] - \mu_* \mu_*\right) \\ &= \sigma_*^2 + \sigma_*^2 + 2\left(\text{E}[X \cdot X_*] - \mu_* \mu_*\right) \\ &= \sigma_*^2 + \sigma_*^2 + 2\left(\text{E}[X \cdot X_*] - \mu_* \mu_*\right) \\ &= \sigma_*^2 + \sigma_*^2 + 2\left(\text{E}[X \cdot X_*] - \mu_* \mu_*\right) \end{aligned}$$

$$\text{HN: Cov}[X_1, X_*] = \text{E}[\{X \cdot \mu_*\} | X_* - \mu_*\} \\ &= \frac{\sigma_*^2 + \sigma_*^2 + 2\left(\text{E}[X \cdot X_*] - \mu_* \mu_*\right)}{\sigma_*} \end{aligned}$$

$$\text{Covariance Rules: Cov}[X_1, X_2] = \sigma_*^2 \\ \text{Cov}[X_1, X_2] = \text{Cov}[X_1, X_2] + \text{Cov}[X_1, X_3] \\ \text{Cov}[X_1, X_2] = \text{Cov}[X_1, X_2] + \text{Cov}[X_1, X_3] \\ \text{Cov}[X_1, X_2] = \sigma_*^2 \\ \text{Cov}[X_1$$

 $\vec{\mu} := E[\vec{X}] := \begin{bmatrix} E[X_1] \\ \vdots \\ E[X_n] \end{bmatrix} = \begin{bmatrix} M_1 \\ \vdots \\ M_n \end{bmatrix} \quad [\text{et } m = \begin{bmatrix} X_1, \dots, X_{1m} \\ \vdots & \vdots \\ X_{0}, \dots, X_{nm} \end{bmatrix} \quad \rightarrow \quad E[m] := \begin{bmatrix} M_1, \dots, M_{1m} \\ \vdots & \vdots \\ M_{0}, \dots, M_{nm} \end{bmatrix}$

X-Multik (n, p), what is E[X]? Var[X]?

(eview trul) = aE[X] + C \Rightarrow identically distributed

 $E[X_i] = \overline{ZE[X_i]} = n\mu$ identically distribute $E[TX_i] = \overline{TE[X_i]} = independent$

Var[X,+...+Xn]= = = Cov[X,, X.]

Review from 241. Let Xi,..., Xn be r.v.'s & a, cER

If
$$X_1, ..., X_k$$
 are independent, what is the varcov matrix?

Z = diag &o,2,..., ox23 -> diagonal matrix

E[aTX]=E[a,X,+...+a,Xx]=a, µ,+...+a, µx = aTp

Rules about vector r.v. expectations
$$E[aX+\vec{c}] = \begin{bmatrix} a\mu_1 + c_2 \\ a\mu_2 + c_k \end{bmatrix} = a\vec{\mu} + \vec{c}$$

 $(\overrightarrow{\nabla}, \overrightarrow{\nabla},)^{\mathsf{T}} = \overrightarrow{\nabla}, \overrightarrow{\nabla}, = \overrightarrow{\nabla}, \overrightarrow{\nabla},$