

$$\frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}} = \frac{\bar{X} - \mu}{\frac{1}{\sqrt{n}} \sqrt{S^2 \cdot \frac{n-1}{n-1} \cdot \frac{n-1}{\sigma^2}}} = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}} \sqrt{\frac{n-1}{n-1} S^2}} = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}} \sqrt{S^2}} \sim \chi_{n-1}^2$$

Due to Cochran's thm, we know \bar{X} and S^2 are independent

Multivariate Normal Distribution (MVN)

$$Z_1, \dots, Z_n \stackrel{iid}{\sim} N(0, 1), \quad \vec{Z} = \begin{bmatrix} Z_1 \\ \vdots \\ Z_n \end{bmatrix}, \quad E[\vec{Z}] = \vec{0}_n, \quad Var[\vec{Z}] = I_n$$

Standard MVN


$$\vec{Z} \sim f_{\vec{Z}}(\vec{z}) = \prod_{i=1}^n f_{Z_i}(z_i) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-z_i^2/2} = \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} \sum_{i=1}^n z_i^2} = \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} \vec{z}^T \vec{z}} = N_n(\vec{0}, I)$$

$$\vec{X} = \vec{Z} + \vec{\mu}, \quad \vec{\mu} \in \mathbb{R}^n, \quad E[\vec{X}] = \vec{\mu}, \quad Var[\vec{X}] = I_n \Rightarrow \vec{X} \sim N_n(\vec{\mu}, I)$$

$$\vec{X} = A \vec{Z} = \begin{bmatrix} Z_1 + Z_2 \\ Z_1 + Z_2 + Z_3 \\ \vdots \\ Z_1 + \dots + Z_n \end{bmatrix} \sim \begin{matrix} N(0, 1) \\ N(0, 1) \\ \vdots \\ N(0, 1) \end{matrix}$$

but the components are dependent e.g.

$$A = \begin{bmatrix} 1 & 0 & 0 & \dots & 0 \\ 1 & 1 & 0 & \dots & 0 \\ 1 & 1 & 1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & 1 & \dots & 1 \end{bmatrix}_{n \times n}$$

$$Cov[X_1, X_2] = Cov[Z_1, Z_1 + Z_2] = Cov[Z_1, Z_1] + Cov[Z_1, Z_2] = 1 + 0 = 1 \Rightarrow X_1, X_2 \text{ dep.}$$


Let's derive a general formula for the variance-covariance matrix of A (an $n \times n$ matrix of scalars) times a random vector X of dim n :

$$Var[A\vec{X}] := E[(A\vec{X})(A\vec{X})^T] - E[A\vec{X}]E[A\vec{X}]^T$$

$$= A E[\vec{X}\vec{X}^T] A^T - A E[\vec{X}] E[\vec{X}]^T A^T$$

$$= A \left(E[\vec{X}\vec{X}^T] - E[\vec{X}]E[\vec{X}]^T \right) A^T = A \Sigma A^T$$

$\Sigma = Var[\vec{X}]$ $E[\vec{X}]^T A^T$

$$\vec{X} = A\vec{Z}, \quad Var[\vec{X}] = A I_n A^T = A A^T \quad \text{Conjecture: } \vec{X} \sim N(\vec{0}, A A^T)$$

$$\vec{X} = A\vec{Z} + \vec{\mu}, \quad A \in \mathbb{R}^{m \times n}, \quad \vec{\mu} \in \mathbb{R}^m, \quad \vec{X} \sim f_{\vec{X}}(\vec{x}) = ?$$

$$g(\vec{z}), \quad h(\vec{x}) = \vec{z} \quad \text{where hopefully } g, h \text{ are inverses}$$

$$\vec{z} = h(\vec{x}) = A^{-1}(\vec{x} - \vec{\mu}) \Rightarrow \text{in order for the inverse to exist... } A \text{ has to be invertible.}$$

$$J_h = \det \begin{bmatrix} \partial h_1 / \partial x_1 & \dots & \partial h_1 / \partial x_n \\ \vdots & & \vdots \\ \partial h_m / \partial x_1 & \dots & \partial h_m / \partial x_n \end{bmatrix} = \det \begin{bmatrix} b_{11} & b_{12} & \dots & b_{1n} \\ \vdots & \vdots & \ddots & \vdots \\ b_{m1} & b_{m2} & \dots & b_{mn} \end{bmatrix} = \det[A^{-1}] = \frac{1}{\det[A]}$$

$$\det[A^{-1}] = \frac{1}{\det[A]}$$

$$f_{\vec{X}}(\vec{x}) = f_{\vec{Z}}(h(\vec{x})) |J_h| = \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} (A^{-1}(\vec{x} - \vec{\mu}))^T A^{-1}(\vec{x} - \vec{\mu})} \frac{1}{|\det[A]|}$$

$$\int \det[A^{-1}] = 1 \Rightarrow (\det[A^{-1}])^T = \det[A^{-1}] = 1 \Rightarrow (\det[A^{-1}])^T = \det[A^{-1}] = 1 \Rightarrow (\det[A^{-1}])^T = \det[A^{-1}] = 1$$

$$= \frac{1}{\sqrt{(2\pi)^n \det[A]^2}} e^{-\frac{1}{2} (\vec{x} - \vec{\mu})^T (A^T)^{-1} A^{-1} (\vec{x} - \vec{\mu})}$$

let $\Sigma = A A^T = Var[\vec{X}]$

$$\det[\Sigma] = \det[A A^T] = \det[A] \det[A^T] = \det[A]^2$$

$$= \frac{1}{\sqrt{(2\pi)^n \det[\Sigma]}} e^{-\frac{1}{2} (\vec{x} - \vec{\mu})^T \Sigma^{-1} (\vec{x} - \vec{\mu})} = N_n(\vec{\mu}, \Sigma) \quad \text{you need } \Sigma \text{ to be invertible}$$

A little bit of multivariate characteristic functions:

$$\phi_{\vec{X}}(\vec{t}) := E[e^{i \vec{t}^T \vec{X}}] = E[e^{i (t_1 X_1 + \dots + t_n X_n)}] = E[e^{i t_1 X_1} \dots e^{i t_n X_n}]$$

if X_1, \dots, X_n indep

$$\downarrow$$

$$= E[e^{i t_1 X_1}] \dots E[e^{i t_n X_n}] = \phi_{X_1}(t_1) \phi_{X_2}(t_2) \dots \phi_{X_n}(t_n)$$

$$\textcircled{P0} \quad \phi_{\vec{X}}(\vec{0}) = E[e^{i \vec{0}^T \vec{X}}] = 1$$

$\textcircled{P1}$ If two chf's are equal \Rightarrow the two rv's are equal in distribution

$\textcircled{P2}$ $\vec{Y} = A\vec{X} + \vec{b}$, $A \in \mathbb{R}^{m \times n}$, $\vec{b} \in \mathbb{R}^m$, \vec{X} is dim. n , $\Rightarrow \vec{Y}$ is dim. m

$$\phi_{\vec{Y}}(\vec{t}) := E[e^{i \vec{t}^T (A\vec{X} + \vec{b})}] = E[e^{i \vec{t}^T A\vec{X}} e^{i \vec{t}^T \vec{b}}] = e^{i \vec{t}^T \vec{b}} E[e^{i (A^T \vec{t})^T \vec{X}}]$$

$$= e^{i \vec{t}^T \vec{b}} \phi_{\vec{X}}(A^T \vec{t})$$

Let's derive the chf of the standard MVN

$$\phi_{\vec{Z}}(\vec{t}) = \prod_{i=1}^n \phi_{Z_i}(t_i) = \prod_{i=1}^n e^{-t_i^2/2} = e^{-\frac{1}{2} \sum_{i=1}^n t_i^2} = e^{-\frac{1}{2} \vec{t}^T \vec{t}}$$

Let's derive the chf of the general MVN $\vec{X} = A\vec{Z} + \vec{\mu} \sim N(\vec{\mu}, A A^T)$

$$\phi_{\vec{X}}(\vec{t}) \stackrel{\textcircled{P2}}{=} e^{i \vec{t}^T \vec{\mu}} \phi_{\vec{Z}}(A^T \vec{t}) = e^{i \vec{t}^T \vec{\mu}} e^{-\frac{1}{2} (A^T \vec{t})^T A^T \vec{t}}$$

$$= e^{i \vec{t}^T \vec{\mu} - \frac{1}{2} \vec{t}^T \Sigma \vec{t}}$$

$$\vec{Y} = B\vec{X} + \vec{c} \sim ? \quad B \in \mathbb{R}^{m \times n}, \quad \vec{c} \in \mathbb{R}^m$$

$$\phi_{\vec{Y}}(\vec{t}) \stackrel{\textcircled{P2}}{=} e^{i \vec{t}^T \vec{c}} \phi_{\vec{X}}(B^T \vec{t}) = e^{i \vec{t}^T \vec{c}} e^{i (B^T \vec{t})^T \vec{\mu} - \frac{1}{2} (B^T \vec{t})^T \Sigma (B^T \vec{t})}$$

$$= e^{i \vec{t}^T (\vec{c} + B\vec{\mu}) - \frac{1}{2} \vec{t}^T B \Sigma B^T \vec{t}} \stackrel{\textcircled{P1}}{\Rightarrow} \vec{Y} \sim N_m(B\vec{\mu} + \vec{c}, B \Sigma B^T)$$

(if $B \Sigma B^T$ is invertible)

Mahalanobis Distance

$$\text{let } \vec{X} \sim N_n(\vec{\mu}, \Sigma). \text{ Consider } (\vec{X} - \vec{\mu})^T \Sigma^{-1} (\vec{X} - \vec{\mu}) \sim ?$$

$$= (\vec{X} - \vec{\mu})^T (A^{-1})^T A^{-1} (\vec{X} - \vec{\mu})$$

$$= (A^{-1}(\vec{X} - \vec{\mu}))^T A^{-1}(\vec{X} - \vec{\mu})$$

$$\vec{Z} = A^{-1}(\vec{X} - \vec{\mu})$$

$$\Sigma^{-1} = (A A^T)^{-1} = (A^T)^{-1} A^{-1} = (A^{-1})^T A^{-1} = \vec{Z}^T \vec{Z} \sim \chi_n^2$$

PC Mahalanobis discovered this in 1936. He was India's founding father of statistics and founded the Indian Institute of Statistics.

This is kind of like distance in \mathbb{R}^n adjusted for all the dependencies among the dimensions like a multivariate "z-score"

$$\text{In one dimension} \quad (X - \mu) (\sigma^2)^{-1} (X - \mu) = \frac{(X - \mu)^2}{\sigma^2} = \left(\frac{X - \mu}{\sigma} \right)^2 = z^2$$