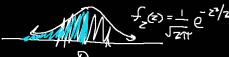


$$Z \sim N(0,1), Y = Z^2 \sim f_Y(y) = ? \text{ Not 1:1}$$



$$F_Y(y) = P(Y \leq y) = P(Z^2 \leq y) = P(Z \in [-\sqrt{y}, \sqrt{y}]) = 2P(Z \in [0, \sqrt{y}])$$

$$= 2(F_Z(\sqrt{y}) - F_Z(0)) = 2F_Z(\sqrt{y}) - 1$$

$$f_Y(y) = \frac{d}{dy} [2F_Z(\sqrt{y}) - 1] = 2 \left(\frac{1}{2} y^{-1/2} \right) f_Z(\sqrt{y}) = y^{-1/2} \frac{1}{\sqrt{2\pi}} e^{-(\sqrt{y})^2/2}$$

$$\propto y^{-1/2} e^{-1/2 y} \mathbb{1}_{y \geq 0} \propto \text{Gamma}\left(\frac{1}{2}, \frac{1}{2}\right)$$

$$Z_1, Z_2, \dots, Z_k \stackrel{\text{iid}}{\sim} N(0,1), Y = Z_1^2 + Z_2^2 + \dots + Z_k^2 \sim \text{Gamma}\left(\frac{k}{2}, \frac{1}{2}\right)$$

Note the beta is always 1/2 and the alpha is always k/2 so k is the only parameter. And because this is a common situation, we give it a special name:

$$\text{Gamma}\left(\frac{k}{2}, \frac{1}{2}\right) = \chi_k^2 \text{ The "chi squared distribution with k degrees of freedom" } k \in \mathbb{N}.$$

$$E[Y] = k E[Z^2] = k \quad k=1, \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

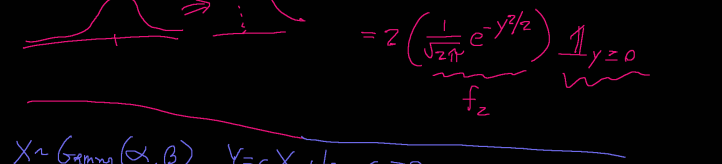
$$\chi_k^2 = \frac{(\frac{1}{2})^{k/2}}{\Gamma(\frac{k}{2})} y^{\frac{k}{2}-1} e^{-y/2} \mathbb{1}_{y \geq 0} \stackrel{k=1}{=} \frac{1}{\sqrt{2\pi}} y^{-1/2} e^{-y/2} \mathbb{1}_{y \geq 0} = \chi_1^2$$

$$X \sim \chi_k^2, Y = \sqrt{X} \Rightarrow x = y^2 = g^{-1}(y), \left| \frac{d}{dy} [g^{-1}(y)] \right| = |2y| = 2y$$

$$f_Y(y) = f_X(y^2) 2y = \frac{(\frac{1}{2})^{k/2}}{\Gamma(\frac{k}{2})} y^{k-2} e^{-y^2/2} (2y) \mathbb{1}_{y^2 \geq 0}$$

$$= \frac{(\frac{1}{2})^{k/2-1}}{\Gamma(\frac{k}{2})} y^{k-1} e^{-y^2/2} \mathbb{1}_{y \geq 0} = \chi_k \text{ the chi distribution with k degrees of freedom}$$

$$Z \sim N(0,1), |Z| = \sqrt{Z^2} \sim \chi_1 = \sqrt{\frac{2}{\pi}} e^{-y^2/2} \mathbb{1}_{y \geq 0}$$



$$X \sim \text{Gamma}(\alpha, \beta), Y = cX \text{ where } c > 0$$

$$f_Y(y) = \frac{1}{c} f_X\left(\frac{y}{c}\right) = \frac{1}{c} \left(\frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{y}{c}\right)^{\alpha-1} e^{-\beta y/c} \mathbb{1}_{\frac{y}{c} > 0} \right)$$

$$= \frac{(\beta/c)^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-(\beta/c)y} \mathbb{1}_{y \geq 0} = \text{Gamma}\left(\alpha, \frac{\beta}{c}\right)$$

$$X \sim \chi_k^2, Y = \frac{X}{k} \sim \text{Gamma}\left(\frac{k}{2}, \frac{1}{k}\right) = \text{Gamma}\left(\frac{k}{2}, \frac{k}{2}\right)$$

$$\text{let } X_1 \sim \chi_{k_1}^2 \text{ indep. of } X_2 \sim \chi_{k_2}^2$$

$$\text{let } U = \frac{X_1}{k_1} \sim \text{Gamma}\left(\frac{k_1}{2}, \frac{k_1}{2}\right) \text{ indep. of } V = \frac{X_2}{k_2} \sim \text{Gamma}\left(\frac{k_2}{2}, \frac{k_2}{2}\right)$$

$$R = \frac{U}{V} \sim f_R(r) = \int_{\text{sup}[V]} f_U(rv) \mathbb{1}_{rv \in \text{sup}[U]} f_V(v) |v| dv$$

$$= \int_0^\infty \frac{a^q}{\Gamma(a)} (rv)^{q-1} e^{-qrv} \mathbb{1}_{rv \in [0, \infty)} \frac{b^b}{\Gamma(b)} v^{b-1} e^{-bv} dv$$

$$= \frac{a^q}{\Gamma(a)} \frac{b^b}{\Gamma(b)} r^{q-1} \mathbb{1}_{r \geq 0} \int_0^\infty t^{q+b-1} e^{-(ar+b)t} dt$$

$$= a^q b^b r^{q-1} \mathbb{1}_{r \geq 0} \frac{1}{\Gamma(a)\Gamma(b)} \Gamma(a+b) \cdot \frac{1}{(ar+b)^{a+b}} = \frac{a^q b^b}{B(a,b)} r^{q-1} \frac{1}{(ar+b)^{a+b}}$$

$$= \frac{(a/b)^q}{B(a,b)} r^{q-1} \left(1 + \frac{a}{b}r\right)^{-(a+b)} \mathbb{1}_{r \geq 0} = \frac{(k_1/k_2)^{k_1/2}}{B(k_1/2, k_2/2)} r^{\frac{k_1}{2}-1} \left(1 + \frac{k_1}{k_2}r\right)^{-\frac{k_1+k_2}{2}} \mathbb{1}_{r \geq 0}$$

$$= F_{k_1, k_2} \text{ this is the "F distribution" or "Fisher-Snedecor Distribution" with } k_1 \text{ numerator degrees of freedom and } k_2 \text{ denominator degrees of freedom. } k_1, k_2 \in \mathbb{N}$$

$$\text{let } Z \sim N(0,1), X \sim \chi_k^2, W = \frac{Z^2}{X/k} \sim f_W(w) = ?$$

$$\text{Consider } W^2 = \frac{Z^2/1}{X/k} \sim F_{1,k} \quad \text{Symmetric around 0}$$

$$F_{W^2}(w^2) = P(W^2 \leq w^2) = P(W \in [-w, w]) = F_W(w) - F_W(-w)$$

$$\text{Take deriv...}$$

$$\frac{d}{dw} [F_{W^2}(w^2)] = \frac{d}{dw} [F_W(w)] - \frac{d}{dw} [F_W(-w)]$$

$$2w f_{W^2}(w^2) = f_W(w) - (-f_W(-w)) = 2f_W(w)$$

$$f_W(w) = \frac{1}{\sqrt{k\pi}} \frac{(\frac{1}{k})^{\frac{k}{2}}}{B(\frac{1}{2}, \frac{k}{2})} (w^2)^{\frac{k}{2}-1} \left(1 + \frac{w^2}{k}\right)^{-\frac{1+k}{2}} \mathbb{1}_{w^2 \geq 0}$$

$$= \frac{\Gamma(\frac{k+1}{2})}{\sqrt{k\pi} \Gamma(\frac{k}{2})} \left(1 + \frac{w^2}{k}\right)^{-\frac{k+1}{2}} = T_k$$

Student's T distribution with k degrees of freedom, discovered in 1908 by William Gosset while he was working at a beer factory

$$\text{If } k \rightarrow \infty \quad T_k \rightarrow Z$$

Student's T distribution has the N(0,1) shape but just thicker tails.

$$Z_1, Z_2 \stackrel{\text{iid}}{\sim} N(0,1), R = \frac{Z_1}{Z_2} \sim \int_{\mathbb{R}} f(r) f(u) |u| du$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-r^2 u^2/2} \frac{1}{\sqrt{2\pi}} e^{-u^2/2} |u| du = \frac{1}{2\pi} \left(\int_{-\infty}^0 e^{-\frac{r^2+1}{2} u^2} |u| du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du \right)$$

$$= \frac{1}{2\pi} \left(\int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du \right) = \frac{1}{\pi} \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du$$

$$\text{let } t = u^2 \Rightarrow \frac{dt}{du} = 2u \Rightarrow du = \frac{1}{2u} dt, u=0 \Rightarrow t=0, u \rightarrow \infty \Rightarrow t \rightarrow \infty$$

$$= \frac{1}{\pi} \int_0^{\infty} e^{-\frac{r^2+1}{2} t} \frac{1}{2u} dt = \frac{1}{2\pi} \frac{1}{\frac{r^2+1}{2}} \int_0^{\infty} \frac{r^2+1}{2} e^{-\frac{r^2+1}{2} t} dt$$

$$= \frac{1}{\pi} \frac{1}{1+r^2} = \text{Cauchy}(0,1)$$

$$\text{let } X = c + \sigma R, R \sim \text{Cauchy}(0,1), \sigma > 0$$

$$X \sim \text{Cauchy}(c, \sigma) = \frac{1}{\sigma\pi} \frac{1}{1 + \left(\frac{x-c}{\sigma}\right)^2}$$