

# Lecture 23

12/7/20

Math 621

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Convergence in prob. to a constant  $X_n \xrightarrow{P} c$  means  
 $\forall \epsilon > 0 \lim_{n \rightarrow \infty} P(|X_n - c| \geq \epsilon) = 0$ .

\* Thm - If  $X_n$  has a finite variance  $\forall n$  &  $E[X_n] = \mu$ ,  
 $\forall n$ , then.....

First consider, Chebyshev's inequality:

$$P(|X_n - \mu| \geq \epsilon) \leq \frac{\sigma_n^2}{\epsilon^2} \text{ now take limits w.r.t. } n \text{ of both sides:}$$

$$\lim_{n \rightarrow \infty} P(|X_n - \mu| \geq \epsilon) \leq \lim_{n \rightarrow \infty} \frac{\sigma_n^2}{\epsilon^2} \text{ since probabilities}$$

are in  $[0, 1]$  if the rhs is 0, then the inequality becomes an equality. Thus, if we show:

$$\lim_{n \rightarrow \infty} \frac{\sigma_n^2}{\epsilon^2} = 0 \rightarrow X_n \xrightarrow{P} \mu$$

$$\frac{1}{\epsilon^2} \lim_{n \rightarrow \infty} \sigma_n^2 = 0 \rightarrow \lim_{n \rightarrow \infty} \sigma_n^2 = 0$$

e.g.  $X_n \sim U(-\frac{1}{n}, \frac{1}{n})$  w.t.s.  $X_n \xrightarrow{P} 0$   
 $\mu = E[X_n] = 0, \sigma_n^2 = \frac{(\frac{1}{n} - (-\frac{1}{n}))^2}{12} = \frac{4}{12n^2}$

$$\lim_{n \rightarrow \infty} \sigma_n^2 = \lim_{n \rightarrow \infty} \frac{4}{12n^2} = 0 \rightarrow X_n \xrightarrow{P} 0$$

e.g.  $X_n \sim N(0, \frac{1}{n})$

$$\mu = 0, \sigma_n^2 = \frac{1}{n} \quad \lim_{n \rightarrow \infty} \sigma_n^2 = \lim_{n \rightarrow \infty} \frac{1}{n} = 0 \rightarrow X_n \xrightarrow{P} 0$$

Let  $X_1, X_2, \dots$  be iid w/ mean  $\mu$  and variance  $\sigma^2$  both finite.

$$\bar{X}_n := \frac{1}{n} \sum X_i \quad E[\bar{X}_n] = \mu, \quad \text{Var}[\bar{X}] = \frac{\sigma^2}{n}$$

$$\lim \text{Var}[\bar{X}] = \lim \frac{\sigma^2}{n} = 0 \rightarrow \bar{X} \xrightarrow{P} \mu.$$

This is called the "weak" "weak law of large numbers" (WLLN)

↓  
the average conv.  
to it's mean with  
"large number" of  
samples ( $n$ ).

↑  
Don't need  
a finite  
variance for  
this to be true.

↑ this is called  
weak b/cuz conv.  
in prob: is  
actually a weak  
type of  
conv.

The third type of convergence is called "convergence in law" or "convergence in  $L^r$ " where  $r \geq 1$ .  
As before, we will only discuss convergence in law to a constant  $c$ . So:

$X_n \xrightarrow{L^r} c$  means  $\lim E[|X_n - c|^r] = 0$ .  
e.g.  $X_n \xrightarrow{L^1} c$  means  $\lim E[|X - c|] = 0$ . "convergence in mean"  
 $X_n \xrightarrow{L^2} c$  means  $\lim E[(X - c)^2] = 0$  "means square convergence".

e.g.  $X_n \sim U(0, \frac{1}{n})$

$\lim_{n \rightarrow \infty} \mathbb{1}_{x \in [0, \frac{1}{n}]}$

WTS  $X_n \xrightarrow{L^r} 0$



$$\begin{aligned}\lim E[|X_n - 0|^r] &= \lim E[X^r] = \lim \int_{x \in [0, \frac{1}{n}]} x^r n \mathbb{1} dx \\ &= \lim n \int_0^{\frac{1}{n}} x^r dx = \lim n \left[ \frac{x^{r+1}}{r+1} \right]_0^{\frac{1}{n}} = \frac{1}{r+1} \lim n \left( \frac{1}{n} \right)^{r+1} \\ &= 0.\end{aligned}$$

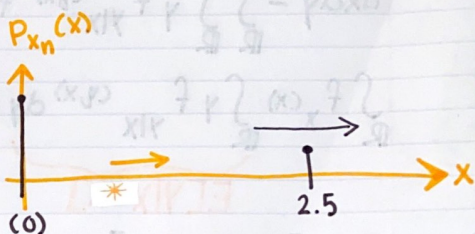
Which convergence is stronger? Law or probability?

$$\begin{aligned}X_n \xrightarrow{L^r} c &\Rightarrow X_n \xrightarrow{P} c \quad \forall r \geq 1 \\ \text{Proof: } \lim P(|X_n - c| \geq \varepsilon) &= \lim P(|X_n - c|^r \geq \varepsilon^r) \leq \\ \lim \frac{E[|X_n - c|^r]}{\varepsilon^r} &= \frac{1}{\varepsilon^r} \lim E[|X_n - c|^r] = 0 \quad \checkmark\end{aligned}$$

$$X_n \xrightarrow{P} c \not\Rightarrow X_n \xrightarrow{L^r} c \quad \forall r \geq 1$$

Counterexample:

$$X_n \sim \begin{cases} n^2 & \text{w.p. } 1/n \\ 0 & \text{w.p. } 1 - 1/n \end{cases}$$

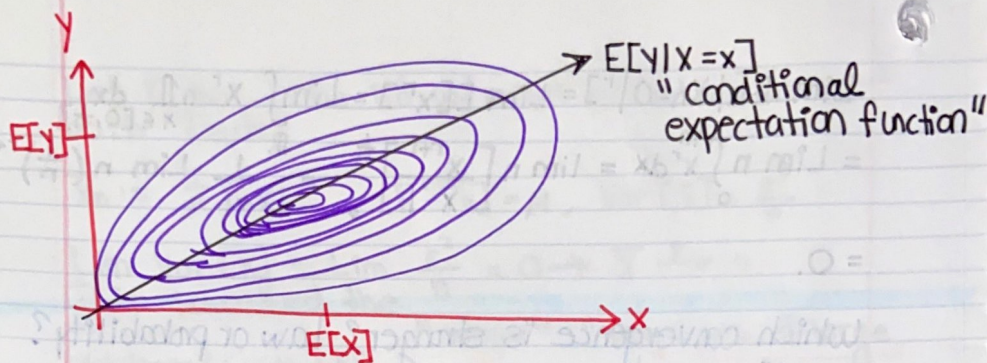


It is clear that

$$\begin{aligned}X_n &\xrightarrow{P} 0 \text{ but } \dots E[X^r] = \sum_{x \in \{0, n^2\}} x^r p(x) \\ &= 0^r (1 - \frac{1}{n}) + (n^2)^r \frac{1}{n} = n^{2r-1} \quad \begin{matrix} r=1 \\ \downarrow \\ =n \end{matrix}\end{aligned}$$

Law of Iterated Expectation

Consider two r.v.'s  $X, Y$  and their joint density  $f_{X,Y}$ .



Here's a nice identity:

$$E[Y] = \int_{\mathbb{R}} y f_Y(y) dy = \int_{\mathbb{R}} y \int_{\mathbb{R}} f_{X,Y}(x,y) dx dy = \int_{\mathbb{R}} y \int_{\mathbb{R}} f_{Y|X}(y|x) f_X(x) dx dy$$

$$dx dy = \int_{\mathbb{R}} \int_{\mathbb{R}} y f_{Y|X}(y|x) f_X(x) dy dx =$$

$$\int_{\mathbb{R}} f_X(x) \int_{\mathbb{R}} y f_{Y|X}(y|x) dy dx = \int_{\mathbb{R}} \underbrace{E[Y|X=x]}_{g(x)} f_X(x) dx$$

$$= E_x[E_Y[Y|X=x]]$$

### Law of total Variance

Here's a nice identity:  $\text{Var}[Y] = E[Y^2] - E[Y]^2$

$$= E_x[E_Y[Y^2|X]] - E_x[E_Y[Y|X]]^2 = E_x[\text{Var}_Y[Y|X] +$$

$$E_Y[Y^2|X]^2] - E_x[E_Y[Y|X]]^2 = E_x[\text{Var}_Y[Y|X]] +$$



$$\text{Let } c = E[Y|X]$$

$$E_x[E_y[Y|X]^2] - E_x[E[Y|X]]^2 = E_x[\text{Var}_y[Y|X]] + \underbrace{E_x[C^2] - E_x[C]^2}_{\text{Var}_x[C]}$$

$$\text{Var}_y[Y] = \underbrace{E_x[\text{Var}_y[Y|X]]}_I + \underbrace{\text{Var}_x[E_y[Y|X]]}_{II}$$

(I) Mean of conditional variances

(II) Variance of the conditional means. This will be large when the CDF line is tilted in places of high  $X$  density.