

Consider rv's  $X_1, X_2, \dots, X_n$  iid but PMF/PDF is unknown but we know it has expectation  $\mu$  and variance  $\sigma^2$ .

Let  $T_n = X_1 + X_2 + \dots + X_n$

Let  $\bar{X}_n = \frac{T_n}{n} = \frac{X_1 + X_2 + \dots + X_n}{n}$

From 2.1, we know  $E[\bar{X}_n] = \mu, \text{Var}[\bar{X}_n] = \frac{\sigma^2}{n}$

Let  $Z_n := \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} = \frac{\sqrt{n}}{\sigma} \bar{X}_n - \frac{\sqrt{n}}{\sigma} \mu$   $E[Z_n] = 0$   
 $\text{Var}[Z_n] = 1 = \text{sd}[Z_n]$

" $\bar{X}$  standardized"

$$\begin{aligned} \phi_{T_n}(t) &\stackrel{(P1)}{=} \phi_{X_1}(t) \phi_{X_2}(t) \dots \phi_{X_n}(t) \stackrel{(P1)}{=} \phi_X(t)^n \\ \phi_{\bar{X}_n}(t) &\stackrel{(P2)}{=} \phi_{T_n}\left(\frac{1}{n}t\right) = \phi_X\left(\frac{t}{n}\right)^n \\ \phi_{Z_n}(t) &\stackrel{(P2)}{=} e^{it\mu} \phi_{\bar{X}_n}\left(\frac{t}{n}\right) = e^{-\frac{it\mu\sqrt{n}}{\sigma} \frac{\sqrt{n}}{\sqrt{n}}} \phi_X\left(\frac{\sqrt{n}}{\sigma} \frac{t}{\sqrt{n}}\right)^n \\ &= e^{-\frac{it\mu\sqrt{n}}{\sigma\sqrt{n}}} e^{\ln\left(\phi_X\left(\frac{t}{\sigma\sqrt{n}}\right)^n\right)} = e^{-\frac{it\mu\sqrt{n}}{\sigma\sqrt{n}} + n \ln\left(\phi_X\left(\frac{t}{\sigma\sqrt{n}}\right)\right)} \\ &= e^{\frac{-\frac{it\mu}{\sigma\sqrt{n}} + \ln\left(\phi_X\left(\frac{t}{\sigma\sqrt{n}}\right)\right)}{\frac{1}{n}} \cdot \frac{\frac{t^2}{\sigma^2}}{\frac{t^2}{\sigma^2}}} = e^{\frac{t^2}{\sigma^2} \left( \frac{\ln\left(\phi_X\left(\frac{t}{\sigma\sqrt{n}}\right)\right) - \frac{it\mu}{\sigma\sqrt{n}}}{\frac{t^2}{\sigma^2 n}} \right)} = \phi_{Z_n}(t) \end{aligned}$$

We want to investigate now  $\lim_{n \rightarrow \infty} \phi_{Z_n}(t) = ?$

$$\begin{aligned} &= e^{\frac{t^2}{\sigma^2} \lim_{n \rightarrow \infty} \frac{\ln\left(\phi_X\left(\frac{t}{\sigma\sqrt{n}}\right)\right) - \frac{it\mu}{\sigma\sqrt{n}}}{\frac{t^2}{\sigma^2 n}}} = e^{\frac{t^2}{\sigma^2} \lim_{u \rightarrow 0} \frac{\ln(\phi_X(u)) - i\mu u}{u^2}} \\ &\stackrel{\text{L'Hopital's}}{\downarrow} = e^{\frac{t^2}{\sigma^2} \lim_{u \rightarrow 0} \frac{\frac{\phi_X'(u)}{\phi_X(u)} - i\mu}{u}} \stackrel{\text{let } u = \frac{t}{\sigma\sqrt{n}} \Rightarrow n \rightarrow \infty \Rightarrow u \rightarrow 0}{=} e^{\frac{t^2}{\sigma^2} \lim_{u \rightarrow 0} \frac{\phi_X(u)\phi_X''(u) - \phi_X'(u)^2}{\phi_X(u)^2}} \\ &= e^{\frac{t^2}{\sigma^2} \frac{\phi_X(0)\phi_X''(0) - \phi_X'(0)^2}{\phi_X(0)^2}} \stackrel{(P0)}{=} e^{\frac{t^2}{\sigma^2} (\phi_X''(0) - \phi_X'(0)^2)} \end{aligned}$$

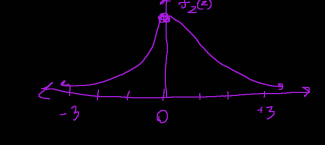
$\stackrel{(P4)}{=} e^{\frac{t^2}{\sigma^2} (i^2 E[X^2] - (i E[X])^2)} = e^{-\frac{t^2}{\sigma^2} (E[X^2] - E[X]^2)} = e^{-\frac{t^2}{2}}$

$\stackrel{(P6)}{\Rightarrow} Z_n \xrightarrow{d} Z$  where  $Z$  has chf  $\phi_Z(t) = e^{-\frac{t^2}{2}}$ .  $Z \sim f_Z(z) = ?$

Use (P6) to find PDF of  $Z$ . Check  $\phi_Z(t) \in L^1 \Rightarrow \int_{\mathbb{R}} e^{-t^2/2} dt = \sqrt{2\pi} < \infty$  YES!

$$\begin{aligned} f_Z(z) &= \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itz} \phi_Z(t) dt \\ &= \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itz} e^{-\frac{t^2}{2}} dt = \frac{1}{2\pi} \int_{\mathbb{R}} e^{-(\frac{t^2}{2} + itz)} dt \\ \frac{t^2}{2} + itz &= \left(\frac{t}{\sqrt{2}} + \frac{\sqrt{2}iz}{2}\right)^2 - \left(\frac{\sqrt{2}iz}{2}\right)^2 = \frac{t^2}{2} + 2 \frac{itiz}{\sqrt{2}} \frac{t}{\sqrt{2}} + \frac{iz^2}{2} - \frac{iz^2}{2} \\ &= \frac{1}{2\pi} \int_{\mathbb{R}} e^{-\left(\frac{t}{\sqrt{2}} + \frac{\sqrt{2}iz}{2}\right)^2} e^{-\frac{z^2}{2}} dt = \frac{1}{2\pi} e^{-\frac{z^2}{2}} \int_{\mathbb{R}} e^{-\left(\frac{t}{\sqrt{2}} + \frac{\sqrt{2}iz}{2}\right)^2} dt \\ \text{let } y &= \frac{t}{\sqrt{2}} + \frac{\sqrt{2}iz}{2} \Rightarrow \frac{dy}{dt} = \frac{1}{\sqrt{2}}, t \rightarrow \infty \Rightarrow y \rightarrow \infty, t \rightarrow -\infty \Rightarrow y \rightarrow -\infty \\ &\stackrel{\text{Gaussian Integral}}{\downarrow} \stackrel{\text{"Standard Normal"}}{\downarrow} \\ &= \frac{1}{2\pi} e^{-\frac{z^2}{2}} \int_{\mathbb{R}} e^{-y^2} \sqrt{2} dy = \frac{1}{2\pi} e^{-\frac{z^2}{2}} \sqrt{2} \sqrt{\pi} = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} = N(0, 1) \end{aligned}$$

This completes the proof of the "central limit theorem" (CLT), the crown jewel of a basic probability class, one of the most useful results that probability has given to the world at large.



AKA Laplace's Second Error Distribution. It is the most famous and widely-used error distribution on Earth.

CLT:  $X_1, \dots, X_n$  iid mean  $\mu$ , variance  $\sigma^2, \frac{\bar{X}_n - \mu}{\frac{\sigma}{\sqrt{n}}} \xrightarrow{d} N(0, 1)$

Let  $\sigma > 0, Z \sim N(0, 1), X = \mu + \sigma Z \sim f_X(x) = ?$   $\phi_X(t) \stackrel{(P1)}{=} e^{it\mu} \phi_Z\left(\frac{t}{\sigma}\right) = e^{it\mu} e^{-\frac{t^2}{2\sigma^2}}$

$f_X(x) = \frac{1}{\sigma} f_Z\left(\frac{x-\mu}{\sigma}\right) = \frac{1}{\sigma} \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2} = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2} (x-\mu)^2} = N(\mu, \sigma^2)$

$E[Z] = \frac{\phi_Z'(0)}{i} = 0, \text{Var}[Z] = E[Z^2] - E[Z]^2 = \frac{\phi_Z''(0)}{-1} = 1$  ✓

$\phi_Z'(t) = \frac{d}{dt} [e^{-t^2/2}] = -t e^{-t^2/2}, \phi_Z''(t) = \frac{d}{dt} [-t e^{-t^2/2}] = -\left(-t^2 e^{-t^2/2} + e^{-t^2/2}\right)$

$E[X] = E[\mu + \sigma Z] = \mu, \text{Var}[X] = \text{Var}[\mu + \sigma Z] = \sigma^2, \text{sd}[X] = \sigma$

$X_1 \sim N(\mu_1, \sigma_1^2)$  indep. of  $X_2 \sim N(\mu_2, \sigma_2^2), T = X_1 + X_2 \sim f_T(t) = ?$

$\phi_T(t) \stackrel{(P3)}{=} \phi_{X_1}(t) \phi_{X_2}(t) = e^{it\mu_1 - \sigma_1^2 t^2/2} e^{it\mu_2 - \sigma_2^2 t^2/2} = e^{it(\mu_1 + \mu_2) - (\sigma_1^2 + \sigma_2^2) t^2/2} \stackrel{(P1)}{\Rightarrow} T \sim N(\mu_1 + \mu_2, \sigma_1^2 + \sigma_2^2)$

$X \sim N(\mu, \sigma^2), Y = e^X \sim f_Y(y) = ?$   $g^{-1}(y) = \ln(y), \left|\frac{d}{dy} g^{-1}(y)\right| = \frac{1}{|y|}$

$f_Y(y) = f_X(\ln(y)) \frac{1}{|y|} = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2} (\ln(y) - \mu)^2} \frac{1}{|y|} = \frac{1}{\sqrt{2\pi}\sigma^2 y} e^{-\frac{1}{2\sigma^2} (\ln(y) - \mu)^2} = \text{Log-Normal distribution}$