

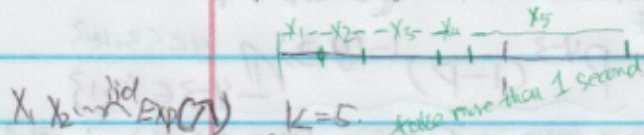
Lecture 10

$T_k \sim \text{Erlang}(k, \lambda)$

$N \sim \text{Poisson}(\lambda)$

$$P(T_k > 1) = 1 - F_{T_k}(1) = Q(k, \lambda)$$

$$F_N(x) = Q(x+1, \lambda)$$



$$\{T_5 > 1\} = \{X_1 + X_2 + X_3 + X_4 < 1\} \cup \{X_1 + X_2 + X_3 < 1\} \cup \{X_1 + X_2 < 1\} \cup \{X_1 < 1\} \cup \{X_1 > 1\}$$

$$\Rightarrow P\{T_5 > 1\} = P(N \leq 4) = F_N(4)$$

$$\Rightarrow 1 - F_{T_k}(1) = F_N(k-1) = Q(k, \lambda) \quad \text{"Poisson Process"}$$

Not event before 1 second

Why is this poisson discrete?

Geometric: experiment $\rightarrow \infty$ $p \rightarrow 0$ $\lambda = np$

$$T \sim \text{Erlang}(k, \lambda) = \frac{\lambda^k}{(k-1)!} t^{k-1} e^{-\lambda t} \mathbb{1}_{t \geq 0} \quad k \in \mathbb{N}, \lambda \in (0, \infty)$$

$$T \sim \text{NegBin}(k, p) = \binom{k+t-1}{k-1} (1-p)^t p^k \mathbb{1}_{t \in \mathbb{N}_0} \quad k \in \mathbb{N}, p \in (0, 1)$$

For Both, what if $k \in (0, \infty)$? Are Both rv's still "legal"?

$$\text{Yes, we can show that Both } \int_0^\infty \frac{\lambda^k}{(k-1)!} t^{k-1} e^{-\lambda t} dt = 1$$

We just derived two new famous rv's

$$X \sim \text{Gamma}(\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\beta x} \mathbb{1}_{x \geq 0}$$

$$X \sim \text{Ext NegBin}(k, p) = \frac{\Gamma(k+t)}{\Gamma(k)!} (1-p)^t p^k \mathbb{1}_{t \in \mathbb{N}_0}$$

the "extended Negative Binomial"

Transformations of Discrete RV's

$$X \sim \text{Bern}(p) = p^x (1-p)^{1-x} \mathbb{1}_{x \in \{0,1\}} = \begin{cases} 1 \text{ w.p. } p \\ 0 \text{ w.p. } 1-p \end{cases}$$

$$Y = X+3 = g(X) \quad \begin{cases} 4 \text{ w.p. } p \\ 3 \text{ w.p. } 1-p \end{cases} \quad \text{PMF}$$

$$X = Y-3 = g^{-1}(Y) = p^{y-3} (1-p)^{1-(y-3)} \mathbb{1}_{y-3 \in \{0,1\}} \quad \begin{cases} y \in \{3,4\} \\ P_X(y-3) \end{cases}$$

I want to find PMF of Y using PMF of X :

$$Y = g(X) \sim P_Y(y) = P_X(g^{-1}(y))$$

What assumption did I make when I "derived" this formula?

I assumed an inverse function exists, i.e. g is invertible

(if not...)

$$X \sim U(\{1, 2, \dots, 10\}) = \begin{cases} 1 \text{ w.p. } \frac{1}{10} \\ 2 \text{ w.p. } \frac{1}{10} \\ \vdots \\ 10 \text{ w.p. } \frac{1}{10} \end{cases} \quad Y = g(X) = \min\{X, 3\} \sim$$

$$\sim \begin{cases} 1 \text{ w.p. } \frac{1}{10} \\ 2 \text{ w.p. } \frac{1}{10} \\ 3 \text{ w.p. } \frac{8}{10} \end{cases} \quad \begin{aligned} & \frac{8}{10} = \frac{P(X=3) + P(X=4) + P(X=5) + \dots + P(X=10)}{1} \end{aligned}$$

$$Y = g(X) \wedge P_Y(y) = \sum_{\{x: y=g(x)\}} P_X(x) \quad \begin{aligned} & \text{if } g \text{ is invertible on support} \\ & \text{one element only} \\ & = g(x) \end{aligned} = \sum_{\{x: x=g^{-1}(y)\}} P_X(x) = P_X(g^{-1}(y))$$

$$X \sim \text{Binom}(n, p) \quad Y = X^3 \sim P_X^{(3X)} = \binom{n}{3Y} p^{3Y} (1-p)^{n-3Y} \mathbb{1}_{3Y \in \{0, \dots, n\}}$$

$$= g^{-1}(y) = 3Y$$

$$Y = X^2 \sim P_X^{(Y)} = \binom{n}{Y} p^Y (1-p)^{n-Y} \mathbb{1}_{Y \in \{0, 1, \dots, n\}}$$

Transformations for continuous RV's

for g invertible $f_Y(y) \stackrel{?}{=} f_X(g^{-1}(y))$ NO!

Let $X \sim U(0,1) = \mathbb{1}_{x \in [0,1]} = f_X(x)$

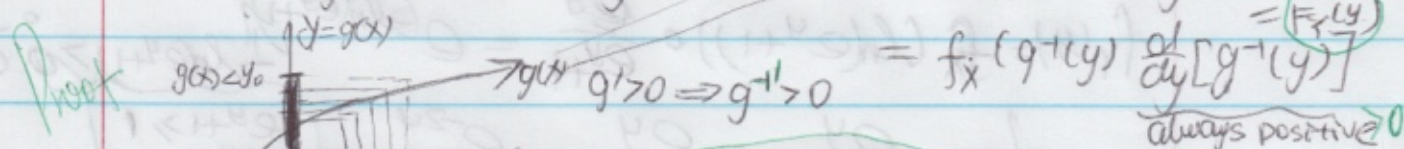
$$Y = 2X \sim f_X(g^{-1}(y)) = f_X\left(\frac{y}{2}\right) = \mathbb{1}_{\frac{y}{2} \in [0,1]} = \mathbb{1}_{y \in [0,2]}$$

$$\int_{\mathbb{R}} f_Y(y) dy = 2 \quad \text{Game over.}$$

PDF are not probabilities! So this was bound to fail because we used them as probabilities. However, CDF's are probabilities

CDF $F_Y(y) := P(Y \leq y) = P(g(X) \leq y) \stackrel{g \text{ is invertible}}{=} P(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$

$\Rightarrow \frac{d}{dy} [F_Y(y)] = \frac{d}{dy} [F_X(g^{-1}(y))] = F'_X(g^{-1}(y)) \frac{d}{dy} [g^{-1}(y)]$



$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) \stackrel{g \text{ is invertible and } g' < 0}{=} P(X \geq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$

$P(X \geq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$

$\frac{d}{dy} [F_Y(y)] = \frac{d}{dy} [1 - F_X(g^{-1}(y))]$

$= f_X(g^{-1}(y)) \left(- \frac{d}{dy} [g^{-1}(y)] \right)$

always negative

~~Rules~~ $F_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} [g^{-1}(y)] \right|$ general Rules

We can derive a less general but very useful corollary rule:

$Y = aX + c \wedge a, c \in \mathbb{R} \quad F_Y(y) = ?$ shift and scale (shift by c, scale by a)

$\stackrel{g(x) \text{ is invertible}}{\Rightarrow} g^{-1}(y) = \frac{y-c}{a} \quad \left| \frac{d}{dy} [g^{-1}(y)] \right| = \left| \frac{1}{a} \right| = \frac{1}{|a|}$

$f_Y(y) = f_X\left(\frac{y-c}{a}\right) \cdot \frac{1}{|a|}$

$Y = aX \sim f_X\left(\frac{y}{a}\right) \frac{1}{|a|}, \quad Y = X + c \sim f_X(y - c)$

$$X \sim \text{Exp}(1) = e^{-x} \mathbb{1}_{x \geq 0}$$

$$Y = g(X) = -\ln\left(\frac{e^{-X}}{1-e^{-X}}\right) = \ln\left(\frac{1-e^{-X}}{e^{-X}}\right) = \ln(e^X - 1) = g(X)$$

invertible
map
 $\wedge f_Y(y)$

$$y = \ln(e^x - 1) \Rightarrow e^y = e^x - 1 \Rightarrow e^{y+1} = e^x = x - \ln(e^y + 1)$$

$$\left| \frac{d}{dy} [g^{-1}(y)] \right| = \left| \frac{d}{dy} [\ln(e^y + 1)] \right| = \left| \frac{e^y}{e^y + 1} \right| = \frac{e^y}{e^y + 1} (= g^{-1}(y))$$

$$f_Y(y) = f_X(\ln(e^y + 1)) \cdot \frac{e^y}{e^y + 1} = e^{-\ln(e^y + 1)} \cdot \frac{e^y}{e^y + 1}$$

$$= \frac{1}{e^y + 1} \cdot \frac{e^y}{e^y + 1} = \frac{e^y}{(e^y + 1)^2} \cdot \frac{e^{-2y}}{e^{-2y}} \quad \begin{matrix} e^y + 1 > 1 \\ e^y > 0 \\ y \in \mathbb{R} \end{matrix}$$

$$= \frac{e^{-y}}{(e^{-y} + 1)^2} = \text{Logistic}(0, 1)$$

standard logistic