

Lec 18

11/16/2020

$Z \sim N(0,1)$, $Y = Z^2 \sim f_Y(y) = ?$ Not 1:1

$$F_Y(y) = P(Y \leq y) = P(Z^2 \leq y) = P(Z \in [-\sqrt{y}, \sqrt{y}]) = 2P(Z \in [0, \sqrt{y}]) \\ = 2(F_Z(\sqrt{y}) - F_Z(0)) = 2F_Z(\sqrt{y}) - 1$$

$$f_Y(y) = \frac{d}{dy} [2F_Z(\sqrt{y}) - 1] = 2 \left(\frac{1}{2} y^{-1/2} \right) f_Z(\sqrt{y}) = y^{-1/2} \frac{1}{\sqrt{2\pi}} e^{-(\sqrt{y})^2/2} \mathbb{1}_{\sqrt{y} \in \mathbb{R}} \\ \propto y^{-1/2} e^{-1/2 y} \mathbb{1}_{y \geq 0} \propto \text{Gamma}(1/2, 1/2)$$

$Z_1, Z_2, \dots, Z_k \stackrel{\text{iid}}{\sim} N(0,1)$, $Y = Z_1^2 + Z_2^2 + \dots + Z_k^2 \sim \text{Gamma}(k/2, 1/2)$

Note The beta is always $1/2$ and the α is always $k/2$ so k is the only parameter. And because this is a common situation, we give it a special name:

$\text{Gamma}(k/2, 1/2) = \chi_k^2$ The "chi squared distri. with k degrees of freedom" $k \in \mathbb{N}$

$$E[Y] = k E[Z^2] = k$$

$$k=1, \Gamma(1/2) = \sqrt{\pi}$$

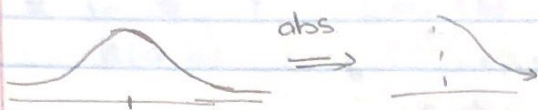
$$\chi_k^2 = \frac{(1/2)^{k/2}}{\Gamma(k/2)} y^{k/2-1} e^{-y/2} \mathbb{1}_{y \geq 0} \stackrel{k=1}{=} \frac{1}{\sqrt{2\pi}} y^{-1/2} e^{-y/2} \mathbb{1}_{y \geq 0} = \chi_1^2$$

$$X \sim \chi_k^2, Y = \sqrt{X} \Rightarrow x = y^2 = g^{-1}(y), \left| \frac{d}{dy} [g^{-1}(y)] \right| = |2y| = 2y$$

$$f_Y(y) = f_X(y^2) 2y = \frac{(1/2)^{k/2}}{\Gamma(k/2)} y^{k-2} e^{-y^2/2} (2y) \mathbb{1}_{y^2 \geq 0}$$

$$= \frac{(1/2)^{k/2-1}}{\Gamma(k/2)} y^{k-1} e^{-y^2/2} \mathbb{1}_{y \geq 0} = \chi_k \text{ the chi distribution with } k \text{ degrees of freedom}$$

$$Z \sim N(0,1), |Z| = \sqrt{Z^2} \sim \chi_1 = \sqrt{2/\pi} e^{-y^2/2} \mathbb{1}_{y \geq 0}$$



$$= 2 \left(\frac{1}{\sqrt{2\pi}} e^{-y^2/2} \right) \mathbb{1}_{y \geq 0}$$

$$X \sim \text{Gamma}(\alpha, \beta), Y = cX \text{ where } c > 0$$

$$f_Y(y) = 1/c f_X(y/c) = 1/c \frac{\beta^\alpha}{\Gamma(\alpha)} \left(\frac{y}{c}\right)^{\alpha-1} e^{-\beta y/c} \mathbb{1}_{y/c > 0}$$

$$= \frac{(\beta/c)^\alpha}{\Gamma(\alpha)} y^{\alpha-1} e^{-(\beta/c)y} \mathbb{1}_{y > 0} = \text{Gamma}(\alpha, \beta/c)$$

$$X \sim \chi_k^2, Y = X/k \sim \text{Gamma}(k/2, \frac{1}{2k}) = \text{Gamma}(k/2, k/2)$$

$$\text{Let } X_1 \sim \chi_{k_1}^2, \text{ indep. of } X_2 \sim \chi_{k_2}^2$$

$$\text{let } U = X_1/k_1 \sim \text{Gamma}(\frac{k_1}{2}, \frac{k_1}{2}) \text{ indep. of } V = X_2/k_2 \sim \text{Gamma}(\frac{k_2}{2}, \frac{k_2}{2})$$

$$R = U/V \sim f_R(r) = \int_{\text{supp}[V]} \int_U f_U(ur) \mathbb{1}_{ur \in \text{supp}[U]} f_V(t) |t| dt$$

$$= \int_0^\infty \frac{a^a}{\Gamma(a)} (rt)^{a-1} e^{-art} \mathbb{1}_{\frac{re \in (0,a)}{rt \in [0,a]}} \frac{b^b}{\Gamma(b)} t^{b-1} e^{-bt} t dt$$

$$= \frac{a^a}{\Gamma(a)} \frac{b^b}{\Gamma(b)} r^{a-1} \mathbb{1}_{r=0} \int_0^\infty t^{a+b-1} e^{-(ar+b)t} dt$$

$$= a^a b^b r^{a-1} \mathbb{1}_{r>0} \frac{1}{\Gamma(a)\Gamma(b)} \Gamma(a+b) \cdot \frac{1}{(ar+b)^{a+b}}$$

$$= \frac{a^a b^b}{B(a,b)} r^{a-1} (ar+b)^{-(a+b)} \mathbb{1}_{r>0}$$

$\frac{1}{B(a,b)} = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$

$$= \frac{(a/b)^a}{B(a,b)} r^{a-1} (1 + a/b r)^{-(a+b)} \mathbb{1}_{r>0}$$

$$= \frac{(k_1/k_2)^{k_1/2}}{B(k_1/2, k_2/2)} r^{k_1/2-1} (1 + k_1/k_2 r)^{-\frac{k_1+k_2}{2}} \mathbb{1}_{r>0}$$

$= F_{k_1, k_2}$ this is the "F distribution" or "Fisher-Snedecor Distribution" with k_1 numerator degrees of freedom and k_2 denominator degrees of freedom.
 $k_1, k_2 \in \mathbb{N}$

Let $Z \sim N(0,1)$, $X \sim \chi^2_k$, $W = Z/\sqrt{X/k} \sim f_W(w) = ?$

Consider $W^2 = Z^2/X/k \sim F_{1,k}$

symmetric around 0

$$F_{W^2}(w^2) = P(W^2 \leq w^2) = P(W \in [-w, w]) = F_W(w) - F_W(-w)$$

Take derivatives

$$d/dw [F_{W^2}(w^2)] = d/dw [F_W(w)] - d/dw [F_W(-w)]$$

$$d/dw F_{W^2}(w^2) = f_W(w) - (-f_W(-w)) = 2f_W(w)$$

$$f_W(w) = \frac{(1/k)^{1/2}}{B(1/2, k/2)} (w^2)^{1/2-1} (1 + w^2/k)^{-\frac{1+k}{2}} \mathbb{1}_{w \in \mathbb{R}}$$

$\frac{1}{B(1/2, k/2)} = \frac{\Gamma(1/2)\Gamma(k/2)}{\Gamma((k+1)/2)}$

$$= \frac{\Gamma\left(\frac{k+1}{2}\right)}{\sqrt{k\pi} \Gamma\left(\frac{k}{2}\right)} (1 + u^2/k)^{-\frac{k+1}{2}} = T_k \text{ Student's } T \text{ distrs.}$$

with k degrees of freedom, discovered in 1908 by William Gosset while he was working at a beer factory.

If $k \rightarrow \infty$ $T_k \rightarrow N(0,1)$ Student's T distribution has the $N(0,1)$ shape but just thicker tails.

$$2, 2, \text{ i.i.d } N(0,1), R = 2/2, \sim \int_{\mathbb{R}} f(u) f(u) |u| du$$

$$= \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-u^2/2} \frac{1}{\sqrt{2\pi}} e^{-u^2/2} |u| du$$

$$= \frac{1}{2\pi} \left(\int_{-\infty}^0 e^{-\frac{r^2+1}{2} u^2} |u| du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du \right)$$

$$= \frac{1}{2\pi} \left(\int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du + \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} |u| du \right)$$

$$= \frac{1}{\pi} \int_0^{\infty} e^{-\frac{r^2+1}{2} u^2} u du \quad \begin{array}{l} \text{let } t = u^2 \Rightarrow dt/du = 2u \Rightarrow du = \frac{1}{2} u dt, \\ u=0 \Rightarrow t=0, u \rightarrow \infty \Rightarrow t \rightarrow \infty \end{array}$$

$$= \frac{1}{\pi} \int_0^{\infty} e^{-\frac{r^2+1}{2} t} \frac{1}{2} dt = \frac{1}{2\pi} \frac{1}{\frac{r^2+1}{2}} \int_0^{\infty} \frac{r^2+1}{2} e^{-\frac{r^2+1}{2} t} dt$$

$$= \frac{1}{\pi} \frac{1}{1+r^2} = \text{Cauchy}(0,1)$$

$\int_0^{\infty} \lambda e^{-\lambda x} dx = 1$
PDF exponent r.v

let $X = c + \sigma R$, $R \sim \text{Cauchy}(0,1)$, $\sigma > 0$

$$X \sim \text{Cauchy}(c, \sigma) = \frac{1}{\sigma\pi} \frac{1}{1 + \left(\frac{y-c}{\sigma}\right)^2}$$