

$$\begin{aligned} \frac{\bar{X} - \mu}{\frac{S}{\sqrt{n}}} &= \frac{\bar{X} - \mu}{\frac{1}{\sqrt{n}} \sqrt{S^2} \frac{\sigma^2}{n-1} \frac{n-1}{\sigma^2}} = \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}} \sqrt{\frac{n-1}{\sigma^2} S^2}} \\ &= \frac{\frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}}}{\sqrt{\frac{n-1}{\sigma^2} S^2}} \sim T_{n-1} \end{aligned}$$

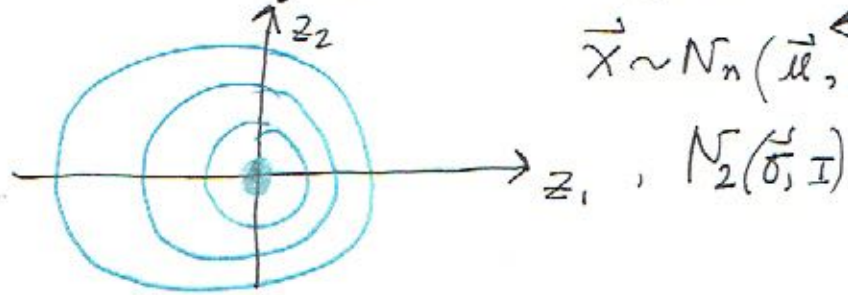
$\left\{ \frac{\bar{X} - \mu}{\frac{\sigma}{\sqrt{n}}} \right\} \sim z$
 $\left\{ \sqrt{\frac{n-1}{\sigma^2} S^2} \right\} \sim \chi^2_{n-1}$

Due to Cochran's theorem, we know \bar{X} and S^2 are independent.

Multivariate Normal Distribution (MNV)

$z_1, \dots, z_n \stackrel{iid}{\sim} N(0, 1), \quad \vec{z} = \begin{bmatrix} z_1 \\ \vdots \\ z_n \end{bmatrix}, \quad E[\vec{z}] = \vec{0},$
 $\vec{z} \sim f_{\vec{z}}(\vec{z}) = \prod_{i=1}^n f_{z_i}(z_i)$
 $\text{Var}[\vec{z}] = I_n$
 $= \prod_{i=1}^n \frac{1}{\sqrt{2\pi}} e^{-z_i^2/2} = \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} \sum_{i=1}^n z_i^2}$
 $= \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} \|\vec{z}\|^2}$

Now, $\vec{X} = \vec{Z} + \vec{\mu}, \quad \vec{\mu} \in \mathbb{R}^n,$
 $E[\vec{X}] = \vec{\mu}, \quad \text{Var}[\vec{X}] = I_n$
 $\vec{X} \sim N_n(\vec{\mu}, I)$



$$\vec{X} = A \vec{Z} = \begin{bmatrix} z_1 \\ z_1 + z_2 \\ \vdots \\ z_1 + \dots + z_n \end{bmatrix} \begin{matrix} \sim N(0,1) \\ \sim N(0,2) \\ \vdots \\ \sim N(0,n) \end{matrix}$$

but components are dependant e.g

$$\text{Cov}[\vec{X}_1, \vec{X}_2] = \text{Cov}[z_1, z_1 + z_2]$$

$$= \text{Cov}[z_1, z_1] + \text{Cov}[z_1, z_2] = 1$$

that means,

X_1, X_2 dependant

$$A = \begin{bmatrix} 1 & 0 & 0 & 0 & 0 & \dots & 0 \\ 1 & 1 & 0 & 0 & 0 & \dots & 0 \\ 1 & 1 & 1 & 0 & 0 & \dots & 0 \\ 1 & 1 & 1 & 1 & 1 & \dots & 1 \end{bmatrix}$$

Let's derive a general formula for the Variance - Covariance matrix of A (an $n \times n$ matrix of Scalars) times a random vector X of dim n :

$$\text{Var}[A\vec{X}] = E[(A\vec{X})(A\vec{X})^T] - E[A\vec{X}]E[A\vec{X}]^T$$

$$= A E[\vec{X}\vec{X}^T] A^T - A E[\vec{X}] (A E[\vec{X}])^T$$

$$= A \left(E[\vec{X}\vec{X}^T] - E[\vec{X}]E[\vec{X}^T] \right) A^T = A \Sigma A^T$$

Now,

$$\vec{X} = A\vec{Z}, \text{Var}[\vec{X}] = A I_n A^T = A A^T, \text{ conjugence: } \vec{X} \sim N(\vec{0}, A A^T)$$

$$\vec{X} = A\vec{Z} + \vec{u}, A \in \mathbb{R}^{n \times n}, \vec{u} \in \mathbb{R}^n, \vec{X} \sim f_{\vec{X}}(\vec{X}) = ?$$

$$\rightarrow g(\vec{Z}), h(\vec{X}) = \vec{Z}, \text{ where hopefully } g \text{ and } h \text{ are inverses}$$

$$h(\vec{X}) = \vec{A}(\vec{X} - \vec{u})$$

In order for the Inverse to exist...
A has to be invertible.

$$= B\vec{X} - B\vec{u} = \begin{bmatrix} \vec{b}_1 \vec{X} - \vec{b}_1 \vec{u} \\ \vec{b}_2 \vec{X} - \vec{b}_2 \vec{u} \\ \vdots \\ \vec{b}_n \vec{X} - \vec{b}_n \vec{u} \end{bmatrix} = \begin{bmatrix} h_1 \\ h_2 \\ \vdots \\ h_n \end{bmatrix}$$

$$J_n = \det \begin{bmatrix} \frac{\partial h_1}{\partial x_1} & \dots & \frac{\partial h_1}{\partial x_n} \\ \vdots & & \vdots \\ \frac{\partial h_n}{\partial x_1} & \dots & \frac{\partial h_n}{\partial x_n} \end{bmatrix} = \det \begin{bmatrix} b_{11} & b_{12} & \dots & b_{1n} \\ b_{21} & b_{22} & \dots & b_{2n} \\ \vdots & & & \vdots \\ b_{n1} & b_{n2} & \dots & b_{nn} \end{bmatrix}$$

$$f_{\vec{x}}(\vec{x}) = f_{\vec{z}}(h(\vec{x})) |J_n| = \det [A^{-1}] = \frac{1}{\det [A]}$$

$$= \frac{1}{(2\pi)^{n/2}} e^{-\frac{1}{2} \underbrace{(A^{-1}(\vec{x}-\vec{\mu}))^T}_{(\vec{x}-\vec{\mu})^T (A^{-1})^T} A^{-1}(\vec{x}-\vec{\mu})} \cdot \frac{1}{|\det[A]|}$$

Now, $DD^T = I \Rightarrow (DD^T)^T = I^T = I \Rightarrow (D^{-1})^T D^T = I \Rightarrow (D^{-1})^T = (D^T)^{-1}$

$$= \frac{1}{\sqrt{(2\pi)^n \det[A]^2}} e^{-\frac{1}{2} (\vec{x}-\vec{\mu})^T \underbrace{(A^T)^{-1} A^{-1}}_{(AA^T)^{-1}} (\vec{x}-\vec{\mu})}$$

$$= \frac{1}{\sqrt{(2\pi)^n \det(\Sigma)}} e^{-\frac{1}{2} (\vec{x}-\vec{\mu})^T \Sigma^{-1} (\vec{x}-\vec{\mu})}$$

$$\begin{aligned} \Sigma &= AA^{-1} = \text{Var}[\vec{x}] \\ \det[\Sigma] &= \det[AA^T] \\ &= \det[A] \det[A^T] \\ &= \det[A]^2 \end{aligned}$$

Multivariate characteristic function

$$\phi_{\vec{x}}(\vec{t}) = E[e^{i\vec{t} \cdot \vec{x}}] = E[e^{i(t_1 x_1 + \dots + t_n x_n)}] = E[e^{it_1 x_1} \dots e^{it_n x_n}]$$

If x_1, x_2, \dots, x_n are independent,

$$= E[e^{it_1 x_1}] \dots E[e^{it_n x_n}] = \phi_{x_1}(t_1) \phi_{x_2}(t_2) \dots \phi_{x_n}(t_n)$$

(P0) $\phi_{\vec{x}}(\vec{0}) = E[e^{i\vec{0} \cdot \vec{x}}] = 1$

(P1) If two c.f.'s are equal \Rightarrow the two r.v.'s are equal in distribution.

(P2) $\vec{y} = A\vec{x} + \vec{b}$, $A \in \mathbb{R}^{m \times n}$, $\vec{b} \in \mathbb{R}^m$, \vec{x} in $\dim n \Rightarrow \vec{y}$ is $\dim m$

$$\phi_{\vec{y}}(\vec{t}) := E[e^{i\vec{t}^T(A\vec{x} + \vec{b})}] = E[e^{i\vec{t}^T A \vec{x}} e^{i\vec{t}^T \vec{b}}]$$

$$= e^{i\vec{t}^T \vec{b}} E[e^{i(\vec{A}^T \vec{t})^T \vec{x}}]$$

$$= e^{i\vec{t}^T \vec{b}} \phi_{\vec{x}}(\vec{A}^T \vec{t})$$

Let's derive the ch's of the standard normal mvn

$$\phi_{\vec{z}}(\vec{t}) = \prod_{i=1}^n \phi_{z_i}(t_i) = \prod_{i=1}^n e^{-t_i^2/2} = e^{-\frac{1}{2} \sum t_i^2} = e^{-\frac{1}{2} \vec{t}^T \vec{z}}$$

Let's derive the ch's of the general mvn

$$\vec{x} = A\vec{z} + \vec{\mu} \sim N(\vec{\mu}, A A^T)$$

$$\phi_{\vec{x}}(\vec{t}) \stackrel{(P2)}{=} e^{i\vec{t}^T \vec{\mu}} \phi_{\vec{z}}(\vec{A}^T \vec{t}) = e^{i\vec{t}^T \vec{\mu}} e^{-\frac{1}{2} (\vec{A}^T \vec{t})^T A^T \vec{t}}$$

$$= e^{i\vec{t}^T \vec{\mu} - \frac{1}{2} \vec{t}^T \underbrace{A A^T}_{\Sigma} \vec{t}}$$

$$\vec{y} = B\vec{x} + \vec{c} \sim ? , B \in \mathbb{R}^{m \times n}, \vec{c} \in \mathbb{R}^m$$

$$\phi_{\vec{y}}(\vec{t}) \stackrel{(P2)}{=} e^{i\vec{t}^T \vec{c}} \phi_{\vec{x}}(B^T \vec{t}) = e^{i\vec{t}^T \vec{c}} \cdot e^{i(\vec{B}^T \vec{t})^T \vec{\mu} - \frac{1}{2} (\vec{B}^T \vec{t})^T \underbrace{\Sigma}_{\vec{B} \Sigma B^T} (B^T \vec{t})}$$

$$= e^{i\vec{t}^T (\vec{c} + B\vec{\mu}) - \frac{1}{2} \vec{t}^T B \Sigma B^T \vec{t}} \stackrel{(P1)}{\Rightarrow}$$

$$\vec{y} \sim N_{\mu}(B\vec{\mu} + \vec{c}, B \Sigma B^T)$$

Let, $\vec{x} \sim N_n(\vec{\mu}, \Sigma)$, consider $\text{if } (B \Sigma B^T \text{ is invertible})$

$$\text{Recall, } \vec{z} = A^{-1}(\vec{x} - \vec{\mu})$$

$$\Sigma^{-1} = (A A^T)^{-1} = (A^T)^{-1} A^{-1} = (A^{-1})^T A^{-1}$$

$$\text{Recall, } \vec{z} = A^{-1}(\vec{x} - \vec{\mu})$$

$$\Sigma^{-1} = (A A^T)^{-1} A^{-1} = (A^{-1})^T A^{-1}$$

$$= (\vec{x} - \vec{\mu})^T (A^{-1})^T A^{-1} (\vec{x} - \vec{\mu})$$

$$= (A^{-1}(\vec{x} - \vec{\mu}))^T A^{-1} (\vec{x} - \vec{\mu})$$

$$= \vec{z}^T \vec{z} \sim \chi_n^2$$

Pr Mahalanobis discovered this in 1936. He was India's founding father of statistics and founded the Indian Institute of Statistics.

This kind of like distance in R^n adjusted for all dependencies among the dimensions like a multivariate "z-score"

In one dimensions,

$$(X - \mu)(\sigma^2)^{-1}(X - \mu) = \frac{(X - \mu)^2}{\sigma^2} = \left(\frac{X - \mu}{\sigma}\right)^2 = z^2$$