

**IMPLEMENTATION AS INTERVENTION:  
CAN IMPROVING BASIC MANAGEMENT PRACTICES STRENGTHEN POLICING  
IN CHICAGO?**

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## ABSTRACT

Market forces push firms to operate closer to their production frontiers, but improving management practices can still increase their productivity. What about public sector agencies insulated from market forces? We study reforms to management practices adopted by the Chicago Police Department (CPD) in the wake of an unprecedented increase in that city's homicide rate. These reforms, known as Strategic Decision Support Centers (SDSCs), did not increase the number of police officers; instead, they helped CPD focus its existing resources on the places and people at highest risk of violence. We measure the SDSCs' impact using a synthetic controls design, creating for each treated police district a comparison district resembling it. Because treated districts are outliers, we encounter challenges applying existing methods: they either fail to produce a similar comparison district or do so in a way that may compromise the reliability of the estimates. We propose modifications to existing methods to address these shortcomings. We also modify existing inference procedures to handle cases where there are few comparison units. Our results suggest that the SDSC in the 7<sup>th</sup> police district, historically one of the most violent in Chicago, decreased shootings by 26 percent and Part I violent felonies by 10 percent without any increase in the number of officers. These findings suggest it is possible to increase public safety without increasing police resources by strengthening police management practices.

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## I. INTRODUCTION

Competitive markets push badly managed firms to exit and incumbent firms to improve their practices. Despite these pressures, firms do not always operate on their production frontier: providing them with management consulting can still raise their productivity (Bloom et al. 2013). What would we expect to see in the public sector? Public agencies seldom face market forces. Instead, voters must exert pressure on underperforming agencies, but can only do so indirectly by choosing a bundle of agency performance. Even determining an agency's performance is not straightforward: the outputs that we ascribe to one agency, like the crime or high school graduation rate, may actually be co-produced using inputs from several agencies. If public agencies operate far from their production frontier as a result of these frictions, then they may benefit even more from improved management practices.

We study a reform to management practices at one public agency: the Chicago Police Department (CPD). In 2016, Chicago's homicide rate per capita increased by fully 60 percent in a single year. Due to its size and the size of Chicago itself, this spike alone accounted for 20 percent of the nationwide increase in homicides that year. In response, CPD implemented a package of management changes that became known as the Strategic Decision Support Centers (SDSCs).<sup>1</sup> The SDSCs are designed to increase the use of data in operational planning, with the goal of making police deployments more precise. They accomplish this mostly through changes to personnel structure and process, rather than through a significant infusion of resources. For example, police districts with an SDSC have a full-time analyst tasked with analyzing crime data that were previously collected but underutilized and preparing a daily briefing for the Commander to inform his or her resource deployments.

This setting is important for three reasons. First, the policy goal of this intervention—reducing gun violence—is relevant not only in Chicago but cities across the country. Firearms account for three quarters of the nearly 20,000 homicides that occur annually in the U.S.<sup>2</sup> Half of these gun homicides occur in about 130 cities, and within those cities most occur in neighborhoods characterized by poverty, low levels of educational attainment, and segregation.<sup>3</sup> Due to this geographic concentration, the impact of urban gun violence is staggeringly disparate: homicide is by far the leading cause of death for black men ages 15-24, a group with a homicide rate almost 20 times greater than that of white men the same age (81 vs. 4.4 per 100,000).<sup>4</sup>

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<sup>1</sup> The SDSCs were designed and implemented by CPD in partnership with Chief Sean Malinowski of the Los Angeles Police Department and our research center, the University of Chicago Crime Lab.

<sup>2</sup> Centers for Disease Control (CDC) and Prevention. National Centers for Injury Prevention and Control, Web-based Injury Statistics Query and Reporting System (WISQARS). In 2017, the number of homicide firearm deaths was 14,542. In the same year, the overall number of homicides was 19,510.

<sup>3</sup> <https://www.theguardian.com/us-news/ng-interactive/2017/jan/09/special-report-fixing-gun-violence-in-america>

<sup>4</sup> CDC WISQARS. In 2017, 2,761 black men ages 15-24 died of homicide, more than the nine next most common causes of death combined.

Second, among public agencies, police departments may be especially good candidates to benefit from reforms to management practices. Police departments have traditionally been characterized by an insular culture that isolates officers from citizens,<sup>5</sup> and moving laterally across departments is challenging relative to other professions.<sup>6</sup> Furthermore, while some public agencies provide services for which there exist non-public analogs—private schools and hospitals, for example—there is no non-public analog providing the services of a police department. These features may limit the ability of police departments to adopt good management practices from other police departments, as well as from outside of policing altogether.

Finally, determining whether better management practices can improve police productivity expands our limited understanding about how to get more from policing. Research over the last two decades has consistently shown that *additional* police resources are effective at reducing crime.<sup>7</sup> Yet little is currently known about how to produce more public safety with the *existing* \$100 billion spent annually on policing in the U.S.<sup>8</sup> Cities facing a rapid deterioration in public safety, like the one Chicago experienced in 2016, cannot increase police resources quickly in response. And many cities with public safety challenges also face fiscal constraints and fraught police-community relations, limiting their ability to spend more on policing in the foreseeable future.

Measuring the impact of these management changes is not straightforward. In February 2017, the first SDSCs opened in the two police districts with the highest violent-crime rates—the 7<sup>th</sup> district on the South Side and the 11<sup>th</sup> district on the West Side—before expanding to four additional districts in March 2017. A randomized controlled trial was not an option given the pressing public safety crisis facing Chicago and the need to prioritize the highest-violence districts for reform. And with only 22 police districts in Chicago, a regression discontinuity design lacks adequate statistical power.

We instead use the synthetic controls method to measure the effects of the SDSCs. Unlike a standard difference-in-differences estimator, a synthetic control uses pre-treatment data to assign individual weights to each control group unit with the goal of approximating the pre-treatment patterns of the treated unit. However, applying the synthetic controls method to our setting is made complicated by the fact that the SDSC districts are significant outliers. The most commonly used synthetic controls estimator, introduced by Abadie, Diamond, and Hainmueller (2010; ADH hereafter), imposes rigid restrictions on the weights that so-called donor units can receive to limit extrapolation beyond the support of the data. A newer estimator introduced by Doudchenko and

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<sup>5</sup> <https://www.city-journal.org/html/taking-back-streets-12586.html>

<sup>6</sup> Traditionally, lateral vacancies are limited to entry-level and command positions. Mid-level officers transferring laterally often have to give up their rank and repeat training, including attending the police academy.

<sup>7</sup> See for example Levitt (1997, 2002); DiTella and Schargrodsy (2004); Klick and Tabarrok (2005); Evans and Owens (2007); Draca, Machin, and Witt (2011); Machin and Marie (2011); Owens (2013); Chalfin and McCrary (2018); Mello (2019).

<sup>8</sup> <https://www.urban.org/policy-centers/cross-center-initiatives/state-and-local-finance-initiative/state-and-local-backgrounders/police-and-corrections-expenditures#Question1Police>

Imbens (2017; DI hereafter) replaces these rigid constraints with a more flexible, data-driven approach. However, the way this new estimator chooses key parameters may lead it to assign weights to donor units on the basis of statistical noise rather than a shared relationship with the treated unit (overfitting). Doing so could undermine the ability of the resulting synthetic control to recover the counterfactual for the treated unit and lead to misestimation of the treatment effect.

We explore the issue of overfitting in synthetic controls and modify existing methods to guard against it. Two modifications concern the role of covariates. Applications of the ADH estimator often rely on the full pre-treatment series of outcome observations to estimate weights; as Kaul et al. (2018) demonstrate, this renders other covariates used in the estimation irrelevant. The DI estimator leaves out covariates altogether. First, we propose using covariates to limit the pool of donor units to those with characteristics similar to those of the treated unit before assigning any donor weights. Second, we propose a way to fit a synthetic control to multiple outcomes simultaneously, such as a primary outcome and one or more covariates or secondary outcomes. Doing this has two advantages: the likelihood of overfitting is reduced when a synthetic control is fit to multiple outcomes, and it yields a single synthetic control for understanding the role of mediating mechanisms in the observed treatment effect on the primary outcome. The last modification is to the DI estimator itself, aiming to make it more robust to overfitting by prioritizing out-of-sample prediction accuracy. Unlike the original estimator, the version we propose uses a time series cross-validation technique to choose parameters governing weight selection that minimize prediction error in one-step-ahead forecasts.

In addition to modifications to address overfitting, we also modify the standard synthetic controls inference method to better handle settings with few comparison areas. The standard inference method with synthetic controls is a placebo test: a placebo treatment effect is estimated for each comparison area, creating a distribution of placebo effects to which the observed effect in the treated unit is compared. However, when there are few comparison areas—16 police districts not receiving an SDSC, in our case—the distribution of placebo effects is sparse and the number of potential  $p$ -values is small. As a result, an observed effect that is close to a placebo effect may, under different realizations of the data, receive substantially different  $p$ -values, limiting our ability to tell whether the observed effect is an outlier or not. We address this problem using a resampling method that increases the density of the placebo effect distribution. Specifically, we increase the number of comparison areas by resampling with replacement the smaller and more numerous geographic areas (police beats) located within the larger and more limited geographic areas available (police districts). Robbins, Saunders, and Kilmer (2017) use a similar permutation technique to construct artificial comparison areas from random assortments of blocks, but must adjust their estimates to correct for the fact that the treatment unit they study is a contiguous area. In contrast, our method sidesteps this issue by using a resampling process that relies on the geographic structure of the real comparison police districts.

We find that the 7<sup>th</sup> district experienced a 26 percent reduction in shooting victims and a 10 percent reduction in Part I violent felony incidents after its SDSC opened. Stated differently, over the 11 post-intervention months of 2017, these results imply that the 7<sup>th</sup> district had 75 fewer shooting victims as a result of the SDSC. This reduction occurred despite no evident increase in the number of officers working in the 7<sup>th</sup> district. We try to detect changes in officer productivity by fitting a single synthetic control to multiple outcomes simultaneously, including measures of

officer activity. Using this method, the estimated violence reduction in the 7<sup>th</sup> district is qualitatively similar and accompanied by substantial increases in gun arrests, arrests of individuals with open warrants, and traffic stops. We also find that officers in the 7<sup>th</sup> district recorded substantially more “positive community interactions.” While many of these results are not statistically distinguishable from zero after correcting for multiple comparisons, they broadly suggest that a change in the behavior of existing 7<sup>th</sup> district officers coincided with the substantial reductions in violence there.

Results from the other SDSC districts are more mixed. The 11<sup>th</sup> district, for example, saw shooting victims decline by 14 percent and Part I violent felonies decline by 13 percent after its SDSC opened. However, these reductions are not robust when a single synthetic control is fit to multiple outcomes simultaneously. Estimates from other districts are either imprecisely estimated or not credible due to the poor fit of the synthetic control, particularly in the last pre-treatment year of 2016. Finally, one clear piece of evidence of the SDSCs’ impact on officer behavior is a significant increase in the number of arrests initiated by officers monitoring police video cameras after the introduction of the SDSCs.

Taken together, these results suggest that it is possible to increase public safety without increasing police resources by strengthening police management practices. This implies that police departments may not be maximizing output with their existing resources. However, the unevenness of these results also underscores a challenge to implementing “process improvement” interventions like the SDSCs. Improved management practices are not a piece of equipment that can be installed and turned on; they require that existing personnel change how they operate. In particular, from our observations of the SDSCs’ adoption, the willingness of a district Commander to embrace management reforms like the daily briefing and the crime analyst seemed integral to an SDSC’s success. This aligns with existing evidence on the importance of individual managers to firm performance (Bertrand and Schoar 2003; Bloom et al. 2018). Given the small share of CPD’s budget that an SDSC represents, and the large social benefit it may generate by reducing gun violence when implemented well, such an intervention is still likely to pass a benefit-cost test.

## **II. BACKGROUND ON CRIME AND CRIMINAL JUSTICE IN CHICAGO**

Chicago is a city of 2.7 million people, divided almost evenly between non-Hispanic whites (32.7 percent of total residents), non-Hispanic blacks (30.1 percent), and Hispanics (29.0 percent).<sup>9</sup> Like other American cities, Chicago remains very racially segregated, as shown in Figure 1: the African-American population is disproportionately concentrated on the city’s South and West Sides, the Hispanic population in a handful of neighborhoods southwest and northwest of the downtown central business district, and the white population in neighborhoods on the north side and in the surrounding suburbs. Segregation by income is closely related to segregation by race, as shown in Figure 2.

Like other American cities, Chicago also struggles with a gun violence problem that is concentrated in the city’s most disadvantaged neighborhoods (Figure 3). For most of the last decade, Chicago has experienced more murders than any other city in the country. However, this

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<sup>9</sup> [https://factfinder.census.gov/faces/nav/jsf/pages/community\\_facts.xhtml](https://factfinder.census.gov/faces/nav/jsf/pages/community_facts.xhtml)

is partly a function of Chicago's size; among cities with a population in excess of 250,000, Chicago's homicide rate per capita has ranked anywhere between the 9<sup>th</sup> and 20<sup>th</sup> highest since 2010, and has consistently been below cities like Baltimore, Detroit, New Orleans, Newark, and St. Louis. Among large American cities, most of the variation across in homicide rates is accounted for by gun homicides, while the rates for non-gun homicide are much more similar (Figure 4).

Chicago has long had the reputation of being a large city with a high rate of serious violence, but the long-term picture is more nuanced and revealing. Figure 5 shows the homicide rates per 100,000 residents in Chicago and two of its peer cities, New York City and Los Angeles, over the last 130 years. A striking feature of this graph is how generally similar the homicide rates for all three cities have been for most of this period; as recently as the early 1990s, they were virtually indistinguishable. But equally striking are two notable periods of divergence, during which Chicago experienced much higher homicide rates than its peers: the 1920s during the Prohibition era, and the last 25 years.

The gulf between Chicago and its peer cities reached its widest point in 2016, when Chicago experienced 765 homicides, an increase of almost 60 percent from 2015. This spike in homicides in Chicago single-handedly accounted for 20 percent of the nationwide increase in homicides in 2016. Annual increases of this enormous magnitude are exceedingly rare for a city of Chicago's size. This unprecedented surge in gun violence was the impetus for the SDSC intervention.

Why gun violence increased so substantially from 2015 to 2016 remains unclear; most existing explanations are, in our view, either wrong or incomplete.<sup>10</sup> One common hypothesis points to social conditions like poverty that are believed to be key "root causes" of crime and violence.<sup>11</sup> Such root causes are surely important for explaining variation across areas in crime rates at a point in time, but they are less helpful in understanding what happened in 2016 given how *sudden* the increase was in Chicago's gun violence. Even with the widely-reported budget problems facing the city of Chicago, Cook County, and the state of Illinois, key root causes generally change slowly over time and, as best we can tell in available data, they did not change rapidly at the end of 2015.<sup>12</sup> Moreover, the 2016 crime increase was mostly confined to gun crimes: shootings rose by 43 percent and gun robberies rose by 26 percent, while non-gun violent crimes only rose by 10 percent and property crime incidents by 6 percent. We would expect important changes in major social conditions to have a more broad-based impact across a range of crime types.

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<sup>10</sup> Much of the data and arguments we present here about the change in crime in Chicago from 2015 to 2016 are taken from Kapustin et al. (2017).

<sup>11</sup> See, for example, NPR's interview with a number of leading advocates and researchers in Chicago: <https://www.npr.org/2016/09/03/492549546/examining-the-reasons-for-chicagos-violence>

<sup>12</sup> One variant of these "root cause" arguments is that government funding for violence mediation and street outreach programming was cut in March 2015. Though gun violence increased following those cuts, those increases are of a magnitude typically seen in Chicago as the weather warms. Adjusting for seasonality, the significant increases in Chicago's gun violence began at the start of 2016.

Other explanations focus on changes in the policing environment in Chicago. For example, Cassell and Fowles (2018) point to a decline in street stops following the consent decree entered into by the American Civil Liberties Union (ACLU) and CPD as the cause of the gun violence spike. Because they focus on changes that occurred proximate to the gun violence increase and that affected one of the key “inputs” to crime control—policing—this explanation is more plausible than those focused on slowly-evolving root causes. However, because there were so many changes occurring within such a short period of time in late 2015 that affected CPD, there is no way to determine how much any single change contributed to the subsequent violence increase.<sup>13</sup>

Since the decline in Chicago’s street stops happened citywide, it is hard to isolate its causal effect on crime. But evidence from other cities suggests the relationship between street stops and violent crime may be weaker than commonly hypothesized. For example, the rate of street stops in New York City declined by over 95 percent from its peak in 2011 in response to legal challenges, with no attendant increase in the city’s homicide rate (Figure 6). Conversely, Milwaukee is another Midwestern city with a homicide trend that closely follows that of Chicago, including a 60 percent homicide rate increase that pre-dated Chicago’s by one year (Figure 7, Panel A). Yet Milwaukee experienced a much more modest decline in its rate of street stops than did Chicago (Figure 7, Panel B).

### **III. PROACTIVE POLICING: HIGH-RISK PLACES AND PEOPLE**

For much of policing’s history in the U.S., the focus was on “reactive policing”: responding quickly to the scene and successfully investigating the case following a call for service. But starting in the 1960s, departments across the country began shifting their focus to “proactive policing,” which a recent National Academy of Sciences committee on the topic defines as “all policing strategies that have as one of their goals the prevention or reduction of crime and disorder and that are not reactive in terms of focusing primarily on uncovering ongoing crime or on investigating or responding to crimes once they have occurred” (National Academies of Sciences, Engineering, and Medicine 2018, p. 1). The standard playbook for modern-day proactive policing is to concentrate police resources on high-risk places and people, in addition to improving relations with the community and emphasizing police problem-solving (National Academies of Sciences, Engineering, and Medicine 2018, p. 2).

Like most other police departments in the U.S., CPD has for years adopted a proactive policing strategy. However, as we discuss later, the Department’s implementation of that strategy left room for improvement in how it carried out two of its core elements: concentrating police resources on high-risk places and people. The SDSCs are designed to provide those improvements.

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<sup>13</sup> On November 24, 2015, the dash cam video of the shooting of Laquan McDonald by CPD officer Jason Van Dyke was released; on December 1, 2015, CPD Superintendent Garry McCarthy was fired and Mayor Rahm Emanuel created the Chicago Police Accountability Task Force; on December 7, 2015, the U.S. Department of Justice Civil Rights Division opened a civil pattern or practice investigation into CPD; on December 9, 2015, Mayor Rahm Emanuel delivered a speech criticizing decades of police corruption; on January 1, 2016, the consent decree between the ACLU and CPD went into effect, requiring an independent evaluation of CPD practices and procedures, and introducing a new Investigatory Stop Report (ISR) that officers were required to complete after every street stop; also on January 1, 2016, Illinois SB1304 went into effect, also requiring police to record information about each investigative street stop and to issue a receipt after each such stop.

In what follows, we discuss the basis for concentrating police resources on high-risk places and people, drawing on evidence from the larger research literature.

### *High-risk places*

Crime, particularly violent crime, is very geographically concentrated. As a result, it is often suggested that crime prevention resources can be used most effectively by deploying them to a handful of “high crime” areas. However, it is important to note that, while crime may be concentrated *ex post*, what is relevant for crime prevention is the degree to which the risk of crime is concentrated *ex ante*. For example, suppose that residential burglaries within a city occur at just 2 percent of addresses. If the annual burglary rate in the city is 2 per 100 households, then this concentration of burglaries is not surprising; in fact, burglaries could have occurred, at most, at just 2 percent of addresses within a year. In contrast to this *ex post* perspective, which notes the concentration of crimes after they occurred, the relevant perspective for crime prevention is an *ex ante* one: predict the likelihood that a given address or block will experience a burglary, and prioritize resources to those at highest risk.

One of the most common place-based strategies used by police departments is to concentrate extra police resources on places with elevated risk of violence, sometimes called “hot spot” policing. Departments vary with respect to whether they do this targeting using an explicit statistical prediction model of where crime will occur (e.g., using commercial software like PredPol or HunchLab), or an implicit prediction model of where crime will occur (e.g., using the locations of similar crimes in the recent past). Several randomized controlled trials (RCTs) within criminology suggest that increased police presence leads to fewer crimes in the locations that are targeted (see, e.g., Braga, Papachristos, and Hureau 2012 for a review).

What has been more difficult to determine is the degree to which hot spot policing actually prevents crime overall, rather than just displacing it to areas with a relatively smaller police presence. Underlying the displacement issue is in part the question of whether the features of a given place can make it unusually criminogenic. While some places surely have such features—such as a street corner used to sell drugs on Chicago’s west side that is adjacent to the access road running alongside I-290, the interstate known as the “heroin highway”—we have limited conceptual guidance, and no reliable empirical evidence, about which factors make some places criminogenic.<sup>14</sup>

The hot spot policing RCTs that have been carried out typically do not find evidence of displacement. But they also typically have lower statistical power to detect displacement than to

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<sup>14</sup> A literature within criminology often labeled under the umbrella of “opportunity theory” suggests the importance of some places in bringing together motivated offenders with suitable targets in situations where there is limited guardianship (see, e.g., National Academies of Sciences, Engineering, and Medicine 2018, p. 46).

detect impacts in the hot spots themselves.<sup>15</sup> One of the best recent studies from a methodological perspective is by Blattman et al. (2019), who find evidence of displacement for property crimes, though their experiment is carried out in Bogotá, Colombia, so the applicability of this finding to hot spot policing in the U.S. is unclear.

A related place-based strategy uses technology, such as closed-circuit television (CCTV) cameras, to help decide where to concentrate police resources in something closer to real-time, as opposed to using statistical models that rely on data recorded with some lag. Cameras can also have other effects on crime as well, such as through potential deterrence and assistance with investigations. The available research suggests the effect of CCTV cameras is a “modest but significant desirable effect on crime” (Welsh and Farrington 2008, p. 2), with impacts particularly concentrated on motor vehicle thefts, though efficacy may also depend on the other crime-control strategies that are in place. As we discuss below, the use of CCTV cameras is increasingly widespread, including in Chicago.

### *High-risk people*

Criminologists have known since the birth cohort studies of Marvin Wolfgang in Philadelphia that the *ex post* (realized) risk of violence involvement is very concentrated within the population (Wolfgang et al. 1972). He found, for example, that about half of all offenses (arrests) were accounted for by 18 percent of the birth cohort. Similar results have subsequently been replicated in a number of other settings. However, as with the burglary example earlier, this is an *ex post* calculation of concentration; it is distinct from, and arguably less useful for crime prevention than, being able to determine which specific youth are at elevated *ex ante* risk for engaging in future violence.

An example of how police departments operationalize this insight is the practice of stepped up enforcement and surveillance of individuals on probation or parole. Because these individuals are often considered to be at higher risk of committing future crimes, and the terms of their release require their compliance with certain conditions, they may receive greater attention from the police. For example, if suspected gang members are on probation and must observe a curfew or refrain from having drugs and guns at home, the police may use this information to monitor them more closely, with the goal of deterring them from engaging in crime.

A different approach to identifying high-risk people is to develop an explicit statistical prediction model of who will engage in crime, the person-based analog to commercially available place-based prediction tools. The Crime and Victimization Risk Model (CVRM), formerly known as the Strategic Subjects List (SSL), is an example of such a model that was developed by Professor Miles Wernick of the Illinois Institute of Technology for CPD. The CVRM predicts an individual’s risk of arrest or victimization for a future shooting or a homicide, relying on predictors such as a

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<sup>15</sup> We thank Justin McCrary for this observation. As a way to see the issue, imagine we have one street block that is identified as a hot spot surrounded by ten adjacent street blocks. If our prior is that crime will be displaced uniformly into the ten adjacent blocks when extra police resources are deployed to the hot spot, then we would need to be able to detect an increase in crime rates in these adjacent areas that is one-tenth as large as the reduction in crime in the hot spot. While the sample size for detecting that displacement effect is ten times as large as for detecting the impact in the hot spot itself, the statistical power for that displacement impact increases only in the square root of the number of displacement areas.

person's age at the time of their most recent arrest, incidents in which they were the victim of a shooting, and incidents in which they were arrested for illegal possession of a gun.<sup>16</sup>

The CVRM, and previously the SSL, is used by CPD to help identify people for its "custom notification" program and "call-in" sessions. During a custom notification, an officer and a community group representative refer a person to social services and warn them about the dangers of continuing to engage in risky behavior.<sup>17</sup> Similarly, call-in sessions involve delivering a message of focused deterrence—heightened penalties for continuing to engage in violence—to a group of individuals, alongside offers of social services.<sup>18</sup> A quasi-experimental study of an earlier version of CPD's call-ins, implemented as part of the local U.S. Attorney's Project Safe Neighborhoods initiative, suggests they may reduce crime in Chicago (Papachristos, Fagan, and Meares 2007). However, it is unclear whether the research design is capable of isolating the effects of the intervention itself due to the difficulty of finding a comparable comparison group, a point we return to below. A different quasi-experimental evaluation of CPD's use of an earlier version of the SSL finds that being identified by the model is not associated with a change in the likelihood of being a homicide or shooting victim, though subjects are more likely to be arrested for a shooting (Saunders, Hunt, and Hollywood 2016). Here, too, it is difficult to know whether the design is able to credibly isolate the effect of being identified as high-risk by the SSL (e.g., being above some threshold). In assessing the effectiveness of any person-based approach to crime prevention, the results of an evaluation may say as much about the intervention being delivered as they do about the method of identifying the people to receive it.

More work is necessary to understand the potential and limitations of using statistical models that predict individuals' risk to target crime-prevention efforts. For example, one study currently underway in Chicago uses, in part, a statistical model developed with data from CPD to identify men at the highest risk for gun violence involvement and refer them to an intensive social service intervention designed to reduce this risk (Bertrand et al., forthcoming). Among the issues that researchers and practitioners must contend with are the limitations of the data used to predict risk. For example, in Chicago, the arrest clearance rate for a serious offense like homicide is below the national average, and the arrest clearance rate for non-fatal shootings is in the single digits. This makes it challenging to measure, and therefore predict, alleged perpetration of serious violence. Finally, and relatedly, the risk of offending is likely to be more concentrated than the risk of victimization due to the small number of people willing to become high-volume shooters, based on anecdotal reports from officers at CPD. In contrast, a shooter may be willing to target anyone in an opposing gang and could also end up injuring innocent bystanders in the process, resulting in a diffusion of victimization risk relative to offending risk.

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<sup>16</sup> <https://home.chicagopolice.org/wp-content/uploads/2019/01/FACT-SHEET-Crime-and-Victimization-Risk-Model-1.pdf>

<sup>17</sup> <http://directives.chicagopolice.org/directives/data/a7a57bf0-1456faf9-bfa14-570a-a2deebf33c56ae59.html>

<sup>18</sup> <http://directives.chicagopolice.org/directives/data/a7a57bf0-136d1d31-16513-6d1d-382b311ddf65fd3a.pdf>

## **IV. STRENGTHENING PROACTIVE POLICING IN CHICAGO THROUGH BETTER MANAGEMENT**

Implementing a proactive policing strategy that concentrates resources on places and people at high *ex ante* risk requires three components: data, analytical capacity, and a process for incorporating analysis into decision-making. In this section, we describe how the SDSCs improved CPD's implementation of a proactive policing strategy by making improvements in these three areas. Because the Department had invested significantly in its data infrastructure for many years, the contributions made by the SDSCs in this area were comparatively more minor, despite their most outwardly visible aspect being a dedicated room within a police station outfitted with computers and large monitors. Instead, where we believe the SDSCs made the largest improvements is to CPD's *management process*: how data are analyzed, by whom, and for what purpose. Though less salient than the technology investments that accompanied the SDSCs, these improvements may represent a more substantial change to the status quo at CPD.

In this section, we first provide necessary background on how the patrol function at CPD is structured, since it is patrol that is ultimately tasked with executing any proactive policing strategy. Then, we discuss each area the SDSCs seek to improve, including their status quo prior to the implementation of the SDSCs and how the SDSCs changed CPD's practices. Finally, we describe how the SDSCs were rolled out across the Department's highest-violence police districts.

### *Structure of patrol at the Chicago Police Department*

Before describing how the SDSCs changed practices at CPD, it is helpful to first understand how the Department's patrol function is structured. CPD has approximately 13,000 officers, over half of whom are within the Bureau of Patrol. To manage the task of patrolling Chicago's roughly 230 square miles, the city is divided into 22 police districts (analogous to precincts in other cities) that each fall within one of three areas: North, Central, and South (Figure 8).<sup>19</sup>

Each police district is overseen by a Commander who decides how to task the patrol resources at his or her disposal. These resources include beat officers, discretionary officers, tactical officers, and intelligence officers. Beat officers patrol the smallest unit of geography—a beat, usually less than 1/10<sup>th</sup> the area of the district—and respond to calls for service. Discretionary officers are similar to beat officers, except they can be deployed anywhere in the district; Commanders usually have between 3 and 9 discretionary officers to work with per shift. Tactical officers typically do not respond to calls for service except those that are highest priority (e.g., shots fired), wear plainclothes, and focus primarily on gangs, executing search warrants, and making gun and other high-profile arrests. Finally, intelligence officers gather information from local sources to provide the Commander a more complete picture of crime in the district; they wear plainclothes and typically do not make arrests.

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<sup>19</sup> In addition to encompassing the city's north side, Area North includes downtown Chicago and much of what is considered the city's West Side. Area Central includes part of the near south side, and Area South covers the remainder of the city's South Side.

In addition to units operating within a district that are overseen by the Commander, each Area is assigned its own units that can be deployed across the districts in that Area. Overseen by the Area's Deputy Chief, these include saturation teams, gun teams, and gang enforcement teams. Saturation teams are deployed as-needed to parts of the Area experiencing heightened activity, acting as a rapid-response force. Gun teams are similar to districts' tactical units and focus on gun recoveries and arrests. Finally, gang teams are deployed to areas of gang activity.

Due to their varying size and the nature of crime that occurs there, the 22 districts function like 22 different small towns and cities, with Commanders as their police chiefs. A smaller district like the 15<sup>th</sup>, covering the West Side neighborhood of Austin, is under four square miles and contains 60,000 residents. A larger district like the 8<sup>th</sup>, covering several South Side neighborhoods including Chicago Lawn, is roughly 24 square miles and contains almost 250,000 residents (about the population of Buffalo, New York). Some districts experience little violent crime, while others have homicide rates rivaling some of the most dangerous cities on earth. Even among the districts experiencing serious violent crime, its nature can vary significantly: gun violence in West Side districts is reputed to be tied to the narcotics trade operating there, while in South Side districts it is thought to be driven more by interpersonal disputes. As a result, each Commander must develop his or her own strategy to address the concerns in their own district, using the information and patrol resources available to them.

#### *Available data*

Concentrating resources on places and people at high *ex ante* risk requires having access to historical (or, if available, real-time) data on the factors that predict where crime will occur and whom it will involve. Prior to the SDSCs, CPD's infrastructure for recording administrative data was already very robust. The Department's data systems track detailed information about reported crimes, victims, and calls for service, in addition to measures of officer activity ranging from arrests, investigatory street stops (Terry stops), traffic stops, and information on officers' locations from GPS units in their vehicles. These data are relatively clean and well-structured, and in most cases go back several years.

The SDSCs expanded CPD's data collection in two ways. First, the Department introduced ShotSpotter's acoustic gunshot detection sensors to each SDSC district. Without ShotSpotter, officers typically rely on calls for service to alert them to shootings. According to officers, it is not uncommon for there to be a 5 to 10-minute delay between when a shooting takes place and when someone calls 911 to report shots fired, often with an incorrect address or location for where they think it occurred. Other times, no call is received at all. After ShotSpotter detects a shooting and a human validates the audio recording, an alert is sent to the SDSC room and to the mobile phones of officers containing the incident's precise location as well as the associated audio. In addition to allowing officers to respond more quickly to the scene and gather evidence, ShotSpotter also provides the Department with an additional and more complete source of data about patterns of gun violence in the city.

Second, SDSC districts expanded their use of CCTV cameras, called Police Observation Devices (PODs). Prior to the SDSCs, monitoring of PODs was very limited due to the cumbersome software available for doing so at the time. The SDSCs introduced additional PODs, which now

number in excess of 30,000; upgraded many cameras from standard definition to high definition; and, perhaps most important, integrated all of the Department’s PODs into an easy-to-use platform called Genetec. Officers in the SDSC room use Genetec to quickly toggle between cameras, rotate their field of vision, and review stored video footage. When officers in the field respond to a situation, or when a ShotSpotter alert goes off, officers in the SDSC room use the nearest PODs to monitor it and provide information to field units.

In contrast to administrative data, information collected from human sources has historically been more difficult to record and make available. District intelligence officers (DIOs), the primary source of human intelligence within a district, often record information on paper that is stored in binders, making it difficult to access. Beat and tactical officers have no system for logging information gleaned while on patrol or from interactions with residents. Detectives, who also gather information pertinent to understanding crime in the district while investigating cases, are not assigned to districts and do not routinely share their insights with DIOs or patrol officers.

The SDSCs appear to have helped facilitate some additional information collection and sharing, though have probably not changed matters dramatically. The SDSC room itself can serve as a focal point in a district. For example, DIOs or patrol officers sometimes drop by and share pertinent information they have obtained, which is then incorporated into analyses and disseminated within the district. The process of developing a daily briefing, described below, is another mechanism to elicit information about crime from officers. Finally, though not explicitly tied to the SDSCs, the Department has prioritized the collection of data on the frequency of officers’ interactions with the community. These data are captured through a call category called a positive community interaction (PCI), which an officer calls in after engaging in such an interaction. The PCI call category has existed since 2015, however it was prioritized significantly, and tracked as part of the Department’s CompStat process, beginning in 2017.

#### *Analytical capacity*

Despite having a well-developed data infrastructure, CPD’s ability to translate data into analyses useful to Commanders for directing resources was very limited prior to the SDSCs’ adoption. Crime analysis was carried out by a handful of officers at headquarters who were responsible for collecting intelligence and identifying crime patterns and trends across all 22 districts. As a result, analysis products tailored to each district’s unique needs were not available to Commanders. Furthermore, although CPD made available several home-grown software tools for accessing and summarizing the vast quantity of data it collects, including a powerful mapping tool, these were seldom used by officers in police districts. This underutilization of the available data—either due to the cumbersome user interface of the available software, the lack of specific training on crime analysis, the absence of a role dedicated to performing this function, or some combination of all three—created a type of “last mile” problem that made local crime analysis the exception rather than the rule at CPD.

The introduction of the SDSCs significantly increased the analytical capacity available to district Commanders. For the first time in the Department’s history, a civilian role specifically devoted to crime analysis was created. The crime analyst works alongside officers and the Commander to develop analytical products that describe recent patterns of criminal activity in the

district. The analyst is capable of using all of CPD's existing software tools—as well as several new ones created specifically to automate common tasks—to make detailed maps of reported crimes or calls for service, or profiles detailing where individuals with open warrants had recently been stopped. Along with the officers staffing the SDSC room, the analyst helps draft a daily briefing for the Commander, described in further detail below.

An example of the type of work performed by crime analysts is presented in Figure 9. In 2017, then-Commander of the 7<sup>th</sup> district, Kenneth Johnson, noting that stolen vehicles were often used as a platform for armed robberies and shootings, asked the district's analyst to examine whether there was an underlying pattern to these thefts. The analyst gathered data on the locations from which cars were stolen and the locations at which they were recovered. This identified a cluster of 18 cars that were all recovered near the same intersection in the 7<sup>th</sup> district and that had been stolen over a six-month period from commercial and residential areas in the adjacent district. The Commander ordered increased patrols in the area where the cars were recovered, and patrol officers were provided information about the pattern. Shortly thereafter, an individual was arrested who had both an extensive history of motor vehicle theft and a close connection to the victims of a quadruple homicide in a neighboring district. The motor vehicle theft pattern subsided after the arrest.

In addition to crime analysts, the SDSCs also introduced HunchLab, a place-based predictive policing software. While the analytical products developed by the analysts make implicit predictions about where crime is likely to occur based on recent data and intelligence gathered from officers, HunchLab applies a statistical model to data on prior crimes and calls for service, as well as geographic and other features, in order to make explicit predictions about where crime is likely to occur. HunchLab divides each district into approximately 500-foot by 500-foot cells to which it assigns predicted probabilities that certain crimes will occur, separately for each of the day's three shifts (watches). Beat officers can access a map on their Department-issued smartphones on which several "mission boxes" are placed on cells with high predicted probabilities; in the highest-violence districts, these mission boxes indicate a high risk of shootings or robberies. Officers are asked to patrol each HunchLab mission box in their beat in three 15-minute intervals, totaling 45 minutes during their shift. In addition, officers in the SDSC room monitor HunchLab mission boxes using POD cameras accessed through Genetec.

### *Decision-making processes*

Prior to the SDSCs, each Commander had a different process for incorporating information into their decisions about which resources to deploy and how. However, because they lacked access to high-quality analysis of local crime patterns, the resulting deployment decisions were often haphazard. Or, as Commander Kenneth Johnson put it, officers were "just patrolling randomly" and "riding around rubber-necking on the street waiting for something to happen."<sup>20</sup> Commanders relied on informal systems to consult with their officers, tactical teams, and other specialized units about deployment plans for that day or that week; these conversations often happened in passing, and rarely offered the Commander a systematic view into the district's recent criminal activity. To the extent that there was a process in place for shaping patrol's response to

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<sup>20</sup> <https://chicago.suntimes.com/2017/3/29/18358635/violent-crime-falls-in-2-districts-run-by-the-johnson-brothers>

recent events, it most commonly took the form of “post-shooting missions”: increased patrols within a very small perimeter around the location of a recent shooting that lasted between 3-5 days. This type of mission appeared to generate little buy-in from officers, who estimated the likelihood of retaliation occurring in such a small area to be low. As a result, and aided by a lack of follow-up and accountability, mission fulfillment may have been lacking.

The SDSCs introduced a structure designed to give Commanders comprehensive, high-quality information in a consistent format each day for use in making patrol decisions. This structure primarily takes the form of the Commander’s daily briefing, which usually occurs at 1:00PM. Prepared by the SDSC officers, and featuring the work of the crime analyst, the presentation covers:

- recent crime trends and high-profile arrests
- high-priority open warrants
- deeper analyses into areas of interest, including those raised at previous briefings
- a comprehensive overview of available discretionary resources, including Area units, as well as their current assigned deployment locations
- the locations of HunchLab mission boxes.

In addition to the Commander, SDSC officers, and the crime analyst, others who attend the briefing include district intelligence and tactical officers, as well as non-district personnel such as Area units and detectives, whose input is valued by the Commander and who can relay information back to their respective teams.

The output of the daily briefing is a set of missions ordered by the Commander and information for dissemination to field units. Missions can vary in their complexity. For example, in response to an active conflict between two gang factions, the Commander may order traffic missions on a busy thoroughfare connecting the rival areas in order to interdict firearms that could be used for retaliatory shootings. An upcoming anniversary of a slain gang member’s death might result in heightened patrol activity in areas associated with the gang member’s rivals. Recent shootings in the vicinity of a gas station or other private business whose owner filed a criminal trespass affidavit with the Department could result in a mission to enforce the affidavit by entering the premises and dispersing trespassers and others engaged in criminal activity.<sup>21</sup> The information produced by the SDSC room is shared with patrol and tactical units during roll calls at the start of each watch. For example, in one district, SDSC officers prepare a binder with the high-priority open warrants discussed in the briefing for tactical officers to review.

In addition to daily briefings, some districts participate in weekly briefings focused specifically on sharing information about recent shootings. In preparation for these weekly briefings, SDSC staff review, in detail, every shooting and homicide that occurred in the district during the past week. Attendees at these briefings include key district personnel, Area units (gang, detectives), and representatives from law enforcement agencies at the County-level (State’s Attorney’s Office, Probation), State-level (Illinois State Police), and federal-level (U.S. Attorney’s

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<sup>21</sup> <http://directives.chicagopolice.org/directives/data/a7a57b38-14a39f6a-cc814-a39f-8ee270ffb8fdf8e8c.html?ownapi=1>

Office, FBI, ATF). The goal of these briefings is to assess the likelihood of retaliation following a given shooting, and to focus prevention efforts on those with the highest likelihood of retaliation.

### *Rollout of the SDSCs*

The impetus for developing and introducing the SDSC model was the unprecedented increase in gun violence that Chicago experienced in 2016. In September 2016, at the request of the CPD Superintendent Eddie Johnson, the Bureau of Justice Assistance (BJA) at the U.S. Department of Justice engaged Chief Sean Malinowski, then the Chief of Staff to Los Angeles Police Chief Charlie Beck, to lead a team of law enforcement experts to assess CPD's crime fighting strategy. After several months of site visits and interviews, the BJA team concluded its work and produced a set of recommendations for how the Department could improve its crime fighting strategy, particularly at the district level. These recommendations called for a combination of improvements to physical infrastructure, targeted investments in technology, the systematic use of data to inform deployment decisions, and a streamlined intelligence gathering process. Collectively, these reforms were adopted by the Department and came to be known as the SDSCs.

The decision was made to first pilot the SDSCs in the 7<sup>th</sup> and 11<sup>th</sup> districts, located on the South and West sides, respectively. The 7<sup>th</sup> and 11<sup>th</sup> districts have long experienced some of the highest levels of violence in Chicago. In 2016, each district saw its homicide rate roughly double, and collectively these two districts accounted for almost a quarter of the city's homicides that year, making them the Department's and the City's top priority. Building out the physical SDSC rooms began near the start of 2017 and was complete by February. Within weeks of becoming operational, the Department and the City decided to expand SDSCs to the four remaining so-called Tier 1 districts: the 6<sup>th</sup>, 9<sup>th</sup>, 10<sup>th</sup>, and 15<sup>th</sup> districts, in addition to the 7<sup>th</sup> and 11<sup>th</sup>. By mid-March, the SDSC rooms in all Tier 1 districts were operational.

Despite their rapid build-out of the SDSCs, the Department struggled to quickly implement one component of the model: crime analysts. No such role previously existed at CPD, and there was no precedent for how to create and fill it. Further, the City estimated that the earliest it could hire analysts for the Tier 1 districts was June 2017. Because of the crucial role they play in the SDSC model, and because, as late as December 2016, the surge in Chicago's violence showed no signs of abating, CPD asked the University of Chicago Crime Lab to provide civilian analyst support while the City's hiring process ran its course. The Crime Lab agreed to help and identified two of its analysts to work with CPD in the 7<sup>th</sup> and 11<sup>th</sup> districts. At the invitation of Chief Malinowski, these two analysts visited LAPD to observe their implementation of a similar data-driven policing model.

The Crime Lab analysts spent the month of January 2017 embedded in the 7<sup>th</sup> and 11<sup>th</sup> districts, speaking with officers, participating in ride-alongs, attending roll calls, meeting with Commanders, and familiarizing themselves with the Department's data infrastructure. As the first civilian analysts to work alongside officers at CPD, the Crime Lab analysts created prototypes of analysis products and daily briefings for Commanders, and developed new tools to automate common tasks and make possible others that were too tedious to do previously. By February 2017, when the SDSCs in the 7<sup>th</sup> and 11<sup>th</sup> districts were fully operational, the Crime Lab analysts were working with district leadership to use the new technology and existing data sources to monitor

changes in criminal activity and gang conflicts in close to real-time. When the SDSCs in the remaining Tier 1 districts opened in mid-March 2017, the Crime Lab deployed more of its existing analysts and hired additional ones to help staff these new SDSCs while the City continued to ramp up its own hiring. Similarly, when the Tier 2 districts' SDSCs began coming online in early 2018, in the 2<sup>nd</sup>, 3<sup>rd</sup>, 4<sup>th</sup>, 5<sup>th</sup>, 8<sup>th</sup>, 12<sup>th</sup>, and 25<sup>th</sup> districts, the Crime Lab deployed its analysts to staff these as well until the City's hiring process was able to catch up.

Soon after the SDSC initiative got underway, the Department and the Crime Lab recognized a need for analysis support at the Area level. Approximately 60 percent of CPD's officers work under the direction of district Commanders; the majority of remaining officers are assigned to specialized units, like gang and narcotics teams, managed by Area Deputy Chiefs. In September 2017, the Crime Lab embedded one analyst in each of the three Areas to work closely with those Deputy Chiefs, providing analytical capacity similar to what was being provided to district Commanders. This was designed to reinforce the district-level SDSCs, with Crime Lab Area analysts looking across district boundaries to spot emerging crime patterns and advising Deputy Chiefs' potential resource deployments.

In addition to providing stop-gap crime analysis capacity to CPD, the Crime Lab analysts provided extensive training to SDSC staff. Initially, this training focused on familiarizing SDSC officers with Microsoft Office software (Word, Excel, PowerPoint) to help them with the preparation of the daily briefing. Over time, however, it came to include training in the Department's existing, but underutilized, data infrastructure and the software designed to access it. When CPD's own crime analysts were hired in June 2017, Crime Lab analysts worked closely with them to share best practices learned during the first few months of the SDSCs' existence. Since then, the Crime Lab has organized several half-day training sessions for CPD's crime analysts, covering such topics as place-based analysis techniques and visualizing geospatial data, and convenes CPD's analysts twice a month to discuss new methods of data analysis, updates to tools and dashboards, and to generate ideas for future trainings.

The Crime Lab's role in the expansion of the SDSCs is unique: it provided the organization (and this research team) unprecedented insight into the functioning of the Department, both prior to and after the intervention's launch, while also shaping the intervention itself. The outsize role played by Crime Lab analysts in training their CPD counterparts, the lasting remnants of which include numerous training materials and input incorporated into CPD's standard operating procedures, will continue to have an impact on CPD's SDSCs even after their support ends.

## V. SYNTHETIC CONTROLS AND THE SDSCS

A randomized controlled trial to measure the impact of the SDSCs was not an option due to the public safety crisis facing Chicago in early 2017, which resulted in the City prioritizing for reform the police districts experiencing the most gun violence. Nor is it likely that SDSCs' impact could be measured using a regression discontinuity design because Chicago is divided into only 22 police districts.<sup>22</sup> Instead, we turn to a non-experimental panel data evaluation method to estimate the SDSCs' effect on gun violence: synthetic controls.

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<sup>22</sup> As noted by Schochet (2009), Wing and Cook (2013), and others, regression discontinuity designs have much lower statistical power to detect treatment effects than randomized controlled trials.

Unlike a standard difference-in-differences estimator, which implicitly assigns equal weights to all control group units,<sup>23</sup> the synthetic controls method uses pre-treatment data to assign individual weights to each control group unit, such that their weighted sum—the synthetic control—resembles the treated unit in its pre-treatment characteristics.<sup>24</sup> Due to the variability in how districts implemented the SDSC intervention, we apply this method separately to each of the six Tier 1 police districts that received an SDSC in 2017. We return to this implementation variability in our discussion of the results, as it may provide a source of leverage to better understand what specific police activities are most useful in reducing crime.

### *Synthetic controls: intuition and pitfalls*

Before describing how we apply the synthetic controls method to measure the SDSCs' impact on gun violence, we first outline the intuition behind the method. We pay particular attention to the issue of overfitting, which we describe below, and which we suspect will become more common as this method is applied to increasingly large datasets with many observed and untreated units.

The synthetic controls method, introduced by Abadie and Gardeazabal (2003) and ADH (2010), relies on pre-period observations of the outcome of interest for the treated unit and a set of untreated donor units. Based on these data, the method chooses a weight for each donor unit such that the weighted sum of donor units' observations in each pre-period is close to that of the treated unit. For one treated unit ( $j = 0$ ) and  $J$  donor units ( $j = 1, \dots, J$ ), using the observed outcome of interest ( $Y_{jt}$ ) in the pre-periods ( $t \leq T_0$ ), a synthetic controls estimator chooses a vector of weights  $\mathbf{w} = (w_1, \dots, w_J)$  to solve an objective function like the following one:

$$\mathbf{w}^* = \arg \min_{\mathbf{w}} \sum_{t=1}^{T_0} \left( Y_{0t} - \sum_{j=1}^J w_j Y_{jt} \right)^2 \quad (1)$$

The synthetic control tries to mimic, as closely as possible, the pre-period observations of the treated unit. The implicit assumption is that, by ensuring a good “pre-period fit,” the synthetic control’s post-period observations ( $t > T_0$ ) provide a good counterfactual for the missing potential outcome observations of the treated unit in the absence of treatment (Rubin 1974; Imbens and Rubin 2015). This assumption is crucial for the validity of the synthetic controls method, but it is impossible to test directly. Instead, when it is applied in practice, the validity of a synthetic control is typically judged by how closely it mimics the pre-period observations of the treated unit, rather

<sup>23</sup> In settings where treatment occurs at different times in different units, difference-in-differences may not assign equal weight to all control group units (see, e.g., Goodman-Bacon 2019).

<sup>24</sup> In most applications of the synthetic controls method, the pre-treatment characteristics of interest are the observed values of the primary outcome. We return to this point later in the section.

than how reliably it is able to recover the missing potential outcome observations of the treated unit (ADH 2015).<sup>25</sup>

Good pre-period fit is an important but potentially misleading indicator of a synthetic control's ability to reliably recover the missing potential outcome observations of the treated unit. A good fit may arise when a synthetic control places weight on donor units that share no underlying relationship with the treated unit but have similar pre-period observations due to idiosyncratic errors. Such donors are chosen on the basis of noise and may not provide any useful signal about the future potential outcome observations of the treated unit, threatening the assumption underlying the validity of the synthetic control. This risk of overfitting is likely to grow with the size of the candidate donor pool, which increases the odds that one or more donors will appear to be a good match for the treated unit due to noise and be assigned weight as a result.

As a stylized example, consider a treated unit with an outcome of interest determined by a vector of unit-specific characteristics,  $\mathbf{X}_0$ , in each period; a treatment effect,  $\alpha$ , in post-periods; and an idiosyncratic error:

$$Y_{0t} = f(\mathbf{X}_0, t) + \mathbf{1}[t > T_0]\alpha + \varepsilon_{0t}$$

The objective of the synthetic control is to recover the treated unit's missing potential outcome observations,  $f(\mathbf{X}_0, t)$  for  $t > T_0$ , using a weighted sum of donors. Suppose there is one signal donor unit available ( $j = 1$ ) with shared unit-specific characteristics ( $\mathbf{X}_1 = \mathbf{X}_0$ ), and  $J - 1$  noise donor units available with different unit-specific characteristics ( $\mathbf{X}_j \neq \mathbf{X}_0, j > 1$ ). Let the outcome of interest for all donor units be determined by:

$$Y_{jt} = f(\mathbf{X}_j, t) + \varepsilon_{jt}$$

The pre-period observations of the noise donors will likely differ from those of the treated unit due to their different unit-specific characteristics, but some may be similar due to their realizations of the stochastic error term.<sup>26</sup> In contrast, the pre-period observations of the signal donor are likely to be similar to those of the treated unit, but for different realizations of its stochastic error term. Further, the signal donor's post-period observations are, in expectation, equivalent to the missing potential outcome observations of the treated unit. Therefore, the optimal vector of donor weights,  $\mathbf{w} = (w_1, w_2, \dots, w_J)$ , is  $\mathbf{w} = (1, 0, \dots, 0)$ . However, there is no guarantee that a synthetic controls estimator will recover this weight vector, particularly if the number of noise donors is large, the

<sup>25</sup> “The credibility of a synthetic control depends upon how well it tracks the treated unit’s characteristics and outcomes over an extended period of time prior to the treatment” (ADH 2015, p. 500).

<sup>26</sup> A donor may also have outcome observations similar to those of the treated unit for reasons other than realizations of the stochastic error term. For example, consider two neighborhoods experiencing low rates of crime: one is socioeconomically advantaged and has low police presence, the other is disadvantaged but has high police presence that deters criminal activity. Both may appear to be plausible candidate donor units for a treated unit also experiencing low rates of crime. However, depending on the socioeconomic conditions and police presence in the treated unit, one donor may be better able to recover the counterfactual than the other, underscoring the potential importance of considering not only pre-period outcome observations but also covariates in assigning donor weights.

length of the pre-period is short, or the variance of the stochastic error distribution is high.<sup>27</sup> In trying to match the pre-period observations of the treated unit using those of the available donors, the synthetic control may overfit by assigning weight to noise donors, undermining its ability to predict the missing potential outcome observations of the treated unit in the post-period.

Although we are not the first to point out this issue, guidance for applied researchers on how to guard against it is lacking. For example, ADH (2010, 2015) warn about the potential for “interpolation” bias if the candidate donor pool contains units with characteristics different from those of the treated unit. However, they offer no recommendation on how to choose candidate donors for the pool. Instead, the estimator that ADH propose seems to incorporate this insight: part of how it chooses weights is to minimize the distance between a vector of characteristics for the treated unit and a vector of characteristics for the synthetic control.<sup>28</sup> But this vector of characteristics can (and in practice often does) include the full set of pre-period observations of the outcome of interest, in addition to other covariates. As demonstrated by Kaul et al. (2018), when the full set of pre-period outcome observations is included, the addition of other covariates proves to be irrelevant; the ADH estimator, by ultimately seeking to minimize the mean squared error between the outcome observations of the treated unit and the synthetic control during the pre-period, ignores the covariates.<sup>29</sup> In other words, in many applications of synthetic controls, the donor pool may contain units that differ substantially from the treated unit, yet the estimator discards covariate information when selecting weights.

### *Synthetic controls: choices that may affect overfitting*

In selecting a synthetic control that can reliably recover the counterfactual, a researcher faces choices in several areas about how best to incorporate the available data, both pre-period outcome observations and covariates. We focus on three areas where the decisions made can affect overfitting: the choice of the candidate donor pool, the choice of the estimator, and the choice of

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<sup>27</sup> For example, holding other factors constant, as  $J$  grows very large, so too does the probability of a noise donor with pre-period outcomes that are, by chance, closer to those of the treated unit than the signal donor:  $\sum_{t=1}^{T_0} (Y_{0t} - Y_{jt})^2 < \sum_{t=1}^{T_0} (Y_{0t} - Y_{1t})^2$  for  $j > 1$ .

<sup>28</sup> The ADH estimator relies on a nested optimization for choosing weights. The inner optimization chooses a donor weight vector,  $W$ , to minimize the distance between the treated unit and the synthetic control in  $X$ , a covariate vector containing candidate predictors, and potentially pre-period observations of  $Y$ , the outcome of interest. The outer optimization chooses a matrix,  $V$ , that assigns predictor weights to each element of  $X$  when choosing  $W$ . For a given value of  $V$ ,  $W^*(V)$  is the optimal donor weight vector chosen by the inner optimization. The outer optimization chooses a  $V^*$  that minimizes the distance between the treated unit and the synthetic control in  $Y$ , using the donor weight vector  $W^*(V^*)$ .

<sup>29</sup> As Kaul et al. (2018) note, there are alternatives to including the full set of pre-period outcome observations in the vector of characteristics, such as including their average or the last pre-period observation. These alternatives increase the likelihood that the synthetic control assigns weight to donor units with similar covariates to the treated unit, in addition to those with similar pre-period outcome observations. However, it is unclear how to choose from among these ad hoc approaches, each of which discard potentially useful pre-period outcome data.

how many outcome variables to use when fitting a synthetic control. Each of these is described in greater detail below.<sup>30</sup>

Although mentioned as a possible strategy by ADH, limiting the candidate donor pool to units with characteristics similar to the treated unit is a technique that appears to seldom be used in practice.<sup>31</sup> Instead, researchers commonly include all available donor units in the pool, perhaps in part because synthetic controls is often applied in settings with few available donor units to begin with. Rather than limiting the donor pool, researchers typically rely on the inclusion of predictor covariates when applying the ADH estimator to guide its choice of donor unit weights—which, as noted above, Kaul et al. (2018) find to be irrelevant when paired with all available pre-period observations of the outcome of interest, as is often done.

An alternative approach is to limit the candidate donor pool prior to choosing the synthetic control weights. For example, techniques from the matching literature, like propensity score estimation, exist to address a similar problem: finding control observations comparable to treatment observations when there are multiple dimensions on which their similarity can be assessed. Using available baseline covariate predictors, one could estimate a propensity score for each donor unit and limit the candidate donor pool to those with scores within some bandwidth around the treated unit. If sufficient pre-period data are available, they can be used to inform the bandwidth choice. In addition to guarding against overfitting, limiting the size of the candidate donor pool, particularly when it is large relative to the number of available pre-period outcome observations, increases the likelihood of there being a unique solution to the synthetic control estimation.

Another choice that can affect the degree of overfitting is that of the estimator used to generate weights. Recent developments in panel data causal inference methods provide researchers with alternatives to the canonical ADH synthetic controls estimator.<sup>32</sup> We consider one such alternative: the estimator proposed by DI (2017). The choice of estimator bears on the issue of overfitting in several ways, but we focus here on how they differ in the restrictions they place on the choice of weights. Restrictions impose a ‘cost’ to assigning non-zero weight to a donor. In the stylized example of signal and noise donors, when assigning weight is costless and the goal is to produce a good pre-period fit, the estimator will be unconstrained; it will assign weight to noise

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<sup>30</sup> A fourth factor that affects the likelihood of overfitting is the length of the pre-period. All things being equal, a longer pre-period reduces the likelihood of noise donors being assigned weight by a given estimator. However, in most settings, researchers already use all the pre-periods in the data they have available. It may be possible in some contexts to obtain additional outcome observations measured at shorter time intervals, but doing so can result in a noisier outcome measure, potentially offsetting any benefit from this approach.

<sup>31</sup> One example is Billmeier and Nannicini (2013), who estimate the impact of economic liberalization on a country’s real GDP per capita. Because their study includes cases of economic liberalization from countries all over the world, they consider two approaches to choosing the candidate donor pool for each treated unit: limiting donors to other countries in the same region as the treated country (type A), or not limiting donors to only countries in the same region (type B). While the authors acknowledge the trade-offs between these two approaches, they nevertheless rely on the type B donor pool to generate their treatment effect estimates in cases where the pre-period fit using the type A donor pool is poor.

<sup>32</sup> See, e.g., Athey et al. (2018); Ben-Michael, Feller, and Rothstein (2018); Powell (2018); Arkhangelsky et al. (2019).

donors if they can improve pre-period fit, even marginally.<sup>33</sup> When assigning weight is costly, however, the estimator will first assign weight to donors that are similar to the treated unit across the full series of pre-period outcome observations, which are more likely to be signal donors.

The ADH estimator imposes weight restrictions that make it unsuitable when the treated unit is an outlier relative to the candidate donor pool, as is the case in our setting. The two restrictions imposed by ADH are that weights must be non-negative ( $w_j \geq 0$ ) and sum to one ( $\sum w_j = 1$ ). Together, these conditions restrict the synthetic control to fall within the convex hull of donor units and limit extrapolation outside the support of the data. As a result, if the treated unit is an outlier and falls outside the convex hull of the donor units, the ADH synthetic control will exhibit poor pre-period fit.

In contrast, the DI estimator produces a synthetic control with better pre-period fit when the treated unit is an outlier by relaxing the ADH constraints. However, it may be more susceptible to overfitting because it does not consider covariates. Instead of requiring that weights be non-negative and sum to one, the DI estimator uses a flexible, data-driven approach—elastic net regularization—to assign smaller weights to each donor unit and to assign non-zero weight to fewer donor units. Specifically, the DI estimator chooses both a vector of weights and an intercept to solve:

$$(\mathbf{w}^*, \mu^*) = \arg \min_{\mathbf{w}, \mu} \sum_{t=1}^{T_0} \left( Y_{0t} - \mu - \sum_{j=1}^J w_j Y_{jt} \right)^2 + \lambda \left( \frac{1-\alpha}{2} \|\mathbf{w}\|_2^2 + \alpha \|\mathbf{w}\|_1 \right) \quad (2)$$

Unlike the ADH estimator, the DI estimator imposes no direct constraint on the weights, which can be negative and sum to any value; instead, their value is indirectly constrained by an elastic net penalty term that prioritizes weight vectors with fewer non-zero ( $\|\mathbf{w}\|_1 = \sum |\omega_j|$ ) and smaller ( $\|\mathbf{w}\|_2^2 = \sum \omega_j^2$ ) entries.<sup>34</sup> Furthermore, the DI estimator does not incorporate covariates in any way, which allows for the choice of an intercept term but may make it particularly susceptible to

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<sup>33</sup> If the number of donor units exceeds the number of pre-periods, the resulting fit from an unconstrained estimator will be perfect.

<sup>34</sup> The magnitude of the elastic net penalty, and the relative weight placed on the Lasso and ridge regression terms, are determined by a pair of hyperparameters  $(\alpha^*, \lambda^*)$  chosen through a cross-validation procedure. In this procedure, for a given pair of  $(\alpha, \lambda)$ , first estimate a weight vector and intercept for each donor unit  $j$ :  $(\hat{\mathbf{w}}(j; \alpha, \lambda), \hat{\mu}(j; \alpha, \lambda))$ . Then, using the estimated weight vector and intercept, calculate the predicted post-period ( $t > T_0$ ) outcome series for each donor unit:  $\hat{Y}_{jt}(\hat{\mathbf{w}}, \mu) = \hat{\mu}(j; \alpha, \lambda) + \sum_{i \neq j}^J \hat{w}_i(\alpha, \lambda) Y_{it}$ . Finally, calculate the cross-validation error across all donor units as the average difference between the observed and predicted post-period outcome series:  $CV(\alpha, \lambda) = \frac{1}{J} \sum_{j=1}^J (\hat{Y}_{jt} - Y_{jt})^2$ . The values of  $(\alpha^*, \lambda^*)$  minimize this cross-validation error.

overfitting.<sup>35</sup> We return to this point below when we discuss how we apply the DI estimator, including a modification we make intended to guard against overfitting.

Finally, the choice of how many outcome variables to use when fitting a synthetic control can also affect overfitting. Until now, we have assumed a single outcome of interest, as is the case in most applications of synthetic controls. However, fitting a synthetic control to multiple outcomes simultaneously reduces the likelihood of assigning weight to noise donors, if the idiosyncratic errors across outcomes are uncorrelated or weakly correlated. Intuitively, if a donor unit's realizations of one outcome are similar to those of the treated unit due to idiosyncratic errors, it is less likely that realizations of a different outcome are also similar to those of the treated unit.

Fitting a synthetic control to multiple outcomes simultaneously has a second benefit as well. When the researcher can observe both a primary outcome of interest and secondary outcomes representing potential mediating mechanisms through which the treatment may affect the primary outcome, obtaining a single vector of synthetic control weights facilitates interpretation of those mediating mechanisms. For example, consider an intervention intended to reduce mortality through the provision of health insurance. The synthetic control used to estimate the treatment effect on mortality (primary outcome) will likely differ from the synthetic control used to estimate the treatment effect on health insurance receipt (secondary outcome). This makes it difficult to know whether a reduction in mortality is due to greater health insurance receipt, since the two effects are derived from comparisons with two different synthetic controls.

Fitting a synthetic control to multiple outcomes simultaneously turns out to be relatively straightforward. Using the DI estimator as an example, for outcomes  $n = 1, \dots, N$ , the estimator chooses a  $1 \times J$  vector of weights and a  $1 \times N$  vector of intercepts to solve:

$$(\mathbf{w}^*, \boldsymbol{\mu}^*) = \arg \min_{\mathbf{w}, \boldsymbol{\mu}} \sum_{n=1}^N \frac{1}{T_{0n}} \sum_{t=1}^{T_{0n}} \left( \frac{Y_{0nt} - \mu_n - \sum_{j=1}^J w_j Y_{jnt}}{\frac{1}{J} \sum_{j=0}^J Y_{jnt}} \right)^2 + \lambda \left( \frac{1-\alpha}{2} \|\mathbf{w}\|_2^2 + \alpha \|\mathbf{w}\|_1 \right) \quad (3)$$

Relative to the original DI estimator in equation (2), this version replaces the mean squared error (MSE) term with a mean squared percentage error, to reflect the fact that the different outcomes  $Y_{jnt}$  may be scaled differently. Failing to account for these scale differences will implicitly give greater priority to weight vectors that minimize the errors of outcomes with larger absolute deviations.<sup>36</sup> In practice, one can operationalize this approach by providing the DI estimator with

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<sup>35</sup> As noted by DI (2017), the ADH estimator does not permit the choice of an intercept because the inner optimization it uses to choose a donor weight vector,  $W$ , seeks to minimize the distance between the treated unit and the synthetic control in  $X$ , a covariate vector containing candidate predictors that may be qualitatively different from each other. An intercept would create a fixed difference between the treated unit and the synthetic control for all covariates, violating scale invariance. In contrast, the DI estimator only seeks to minimize the distance between the treated unit and the synthetic control in pre-period observations of a single variable,  $Y$ , the outcome of interest.

<sup>36</sup> An adjustment is also made to account for the possibility that each outcome  $Y_{jnt}$  may be observed for a different number of pre-periods,  $T_{0n}$ .

$N * T_{0n}$  pre-period observations, with each set of  $T_{0n}$  observations corresponding to a different outcome, a technique similar to one used by Robbins, Saunders, and Kilmer (2017).<sup>37</sup>

### *Synthetic controls: application to the SDSCs*

We start our evaluation by applying the ADH estimator to each SDSC district. Because our focus is their impact on gun violence, we use a count of shooting victims as our primary outcome, scaling by a district's population to account for differences in size.<sup>38</sup> Reliable data on shooting victims are available going back to 2010, providing us with 14 biannual pre-period observations through the beginning of 2017 when the first SDSCs were implemented in the 7<sup>th</sup> and 11<sup>th</sup> districts.<sup>39</sup> We use all of the available pre-period outcome observations when applying the ADH estimator and, consequently, do not incorporate any covariates; we revisit the role of covariates in our application below.

The ADH synthetic control is unable to produce a good pre-period fit (Figure 10, using the 7<sup>th</sup> district as an example). This is not surprising: the SDSC districts are outliers compared to the 16 donor districts, the ADH estimator constrains the synthetic control to be within the convex hull of the donor districts, and it does not permit an intercept. Improving this situation will require either turning to a more flexible estimator or finding a candidate donor pool with donor units experiencing levels of gun violence more comparable to those of the SDSC districts.

Before turning to estimators that may be better equipped to handle outlier treated units, we first consider whether a more diverse candidate donor pool is available. Relying on the fact that Chicago's 22 police districts are subdivided into approximately a dozen beats each (Figure 11), and that some of the 16 donor districts contain beats with high rates of gun violence, we redefine our pool as the 196 donor beats within the 16 donor districts. Comparing the shootings per capita of SDSC districts to those of donor districts and donor beats, respectively, we see that the SDSC districts are not outliers to the same degree relative to the donor beat distribution (Figure 12). Applying the ADH estimator with the pool of donor beats yields synthetic controls with much better pre-period fit, though still with significant deviations (Figure 13). Though a more diverse candidate donor pool improves the performance of the ADH estimator, it remains true that the SDSC districts are outliers relative to the donor beats in their rates of gun violence.

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<sup>37</sup> Though motivated here, in part, by the desire for a single synthetic control to facilitate comparisons of treatment effects across primary and secondary/mediating outcomes, this approach can be used to include any kind of covariate into the DI estimator, helping to address the weakness described earlier.

<sup>38</sup> We have access to data on each shooting incident in the city, which provides us flexibility in how we aggregate them. However, we have found that the synthetic controls method performs poorly when time periods are shorter and, correspondingly, many of them contain no shooting incidents for the treated district.

<sup>39</sup> The biannual periods are from March through August and September through February. In all calculations and figures, we report monthly averages within each period. Because the first SDSCs opened in February 2017, and expanded to the control districts beginning in January 2018, the last pre-period is from September 2016 through January 2017 (5 months), the first post-period is from February 2017 through August 2017 (7 months), and the last post-period is from September 2017 through December 2017 (4 months).

We next turn to the DI estimator and apply it to each SDSC district, using the pool of candidate donor beats.<sup>40</sup> By relaxing the constraints imposed by the ADH estimator, the DI synthetic control achieves a much better pre-period fit (Figure 14). However, we might be worried that, with weaker constraints and so many donor beats from which to choose, the DI estimator achieves this impressive pre-period fit by overfitting, which would compromise the reliability of the synthetic control in the post-period.

This concern is exacerbated in our settings, with multiple treated units that are outliers and may be differentially susceptible to overfitting, because of how the DI estimator chooses two key parameters that govern how weights are determined. As proposed by DI, the two hyperparameters,  $(\alpha, \lambda)$ , are chosen using a cross-validation procedure to minimize error in the *post-period* among the donor units.<sup>41</sup> This procedure yields a single pair of hyperparameters to be used in determining weights for all the treated units, relying only on data from the donor units. This may be particularly problematic here, where our six treated units are outlier police districts on the South and West Sides of Chicago. Furthermore, due to differences in the availability of local non-treated beats in the donor pool, we suspect that some of our treated districts may be more susceptible to overfitting than others.<sup>42</sup>

We propose a modification to the DI estimator that aims to address these concerns. This modification departs from the original DI estimator by choosing hyperparameters separately for each treated unit using a time-series cross-validation technique.<sup>43</sup> Unlike the original DI estimator, which chooses hyperparameters that minimize error in the *post-period* among the donor units, the modified version chooses hyperparameters that minimize error in the *pre-period* for just the treated unit in question. The implicit assumption in the modified approach (that treatment has not yet occurred in the pre-period) is considerably weaker than the one in the original DI approach (that

<sup>40</sup> The cross-validation procedure outlined by DI (2017) requires estimating a synthetic control for each donor unit to choose the hyperparameters  $(\alpha^*, \lambda^*)$ . Due to difficulty estimating a synthetic control for each donor beat, we instead estimate one for each of the 16 donor districts as part of the cross-validation procedure.

<sup>41</sup> This approach relies on the assumption that, absent any treatment effect, synthetic controls chosen for the donor units should closely track their observed outcomes in the post-period. This is a strong assumption, particularly in settings where it is plausible that treatment may result in spillover effects on donor units, as is the case here.

<sup>42</sup> Gun violence on the West Side is thought to be driven by the narcotics trade that occurs along the I-290 corridor, while gun violence on the South Side is considered more idiosyncratic and less purposeful in nature, often revolving around personal disputes. This characterization is a gross oversimplification, but it highlights the fact that West Side donor beats may be poor candidates for a South Side treated unit, and South Side donor beats may be poor candidates for a West Side treated unit. Furthermore, three of the SDSC districts—the 15<sup>th</sup>, 11<sup>th</sup>, and 10<sup>th</sup>—encompass most of the gun violence on the West Side, leaving relatively few non-treated West Side districts from which donor beats can be drawn compared to the South Side, where the three SDSC districts—the 6<sup>th</sup>, 7<sup>th</sup>, and 9<sup>th</sup>—have a comparatively larger pool of non-treated South Side donors in the candidate pool.

<sup>43</sup> For each treated unit, use a subset of sequential pre-periods as a training set, with the one pre-period following as a test set. For a given value of  $(\alpha, \lambda)$ , estimate a weight vector and intercept using the training set, and use these to predict the outcome in the single-period test set, recording the MSE. Then, lengthen the training set by one period, using the next available pre-period as the new test set. Repeat this process until the last pre-period is the test set. Calculate the cross-validation error as the average MSE across all test sets. The values of  $(\alpha^*, \lambda^*)$  minimize this cross-validation error.

the donor units are unaffected by treatment in the post-period). Furthermore, by choosing hyperparameters that minimize prediction error within the treated unit across a series of one-step-ahead forecasts, the estimator builds in some additional protection against overfitting.

We apply the modified DI estimator along with two of the additional safeguards against overfitting described earlier: limiting the candidate donor pool and fitting a synthetic control to multiple outcomes simultaneously. To limit the candidate donor pool, we estimate a propensity score for each candidate donor beat using a vector of characteristics drawn from Census data.<sup>44</sup> In the modified DI estimator cross-validation step, in addition to searching over possible values of the hyperparameters ( $\alpha, \lambda$ ), we also search over potential bandwidths that determine how many of the donor beats that are “closest” in characteristics-space to include in the estimation. After choosing hyperparameters and a donor pool to minimize out-of-sample prediction error, we fit a synthetic control to six outcomes simultaneously: two primary outcomes of interest (shooting victims per capita, Part I violent felony incidents per capita) and four secondary outcomes that measure potential mediating mechanisms (warrant arrests per capita, gun arrests per capita, overall arrests per capita, and traffic stops per capita). For a given outcome  $n$  of a treated unit,  $Y_{0nt}$ , and the associated synthetic control,  $\hat{Y}_{0nt} = \sum_{j=1}^{J^*} w_j^* Y_{jnt}$ , we estimate the treatment effect in period  $t$  as:

$$\hat{\tau}_{nt} = Y_{0nt} - \hat{Y}_{0nt} \quad (4)$$

Having outlined how we choose our synthetic control and generate treatment effect estimates, we now describe how we assess the significance of those estimates. We start with the placebo test suggested by ADH (2010, 2015), in which a synthetic control is calculated for each donor unit, along with a placebo treatment effect. Under the null hypothesis that the donor units are unexposed to any treatment, the estimated treatment effect for the treated unit is compared to this distribution of placebo treatment effects. The fraction of placebo treatment effects greater than or equal to the estimated treatment effect is the  $p$ -value.

Due to difficulty estimating a synthetic control for each donor beat, we instead estimate one for each of the 16 donor districts. But this yields a distribution of placebo treatment effects with only 16 values, limiting our ability to assess precisely how extreme the estimated treatment effect is. For example, if the estimated treatment effect in district A is just slightly larger than the largest placebo treatment effect, and the estimated treatment effect in district B is much larger, both will receive  $p$ -values of zero. Under different realizations of the placebo treatment effects distribution, this  $p$ -value for district B may hold, but the one for district A may not.

We resolve the inference challenge posed by having too few donor districts through the use of a resampling method. Using the fact that each police district is comprised of individual beats, we create resampled donor police districts by sampling beats with replacement from within each actual donor district. These resampled police districts are slightly perturbed versions of the original

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<sup>44</sup> Propensity scores are estimated using a logistic regression on the sample of candidate donor beats ( $D = 0$ ) and beats in the treated district ( $D = 1$ ). Predictors include total population, fraction of households with annual income under \$40,000, fraction African American, and fraction Hispanic.

donor districts. As a result, in the example above, the estimated treatment effect in district A may no longer be larger than the largest placebo treatment effect, increasing its *p*-value. Using resampled police districts dramatically increases the size the donor district pool for the inference procedure, while preserving the geographic clustering of beats by resampling within districts rather than across them, which may be important in environments in which statistical noise has a strong patterning by place and time.<sup>45</sup>

### *Spillovers*

One concern worth noting is the potential for bias due to treatment spillover from the introduction of the SDSCs, particularly given the close geographic proximity of some candidate donor units. This potential spillover could operate through several channels. For example, if treatment induces a reallocation of resources across districts, then the outcomes of donor districts may be affected by treatment even if they did not receive an SDSC.<sup>46</sup> Even if resources across districts remain fixed, individuals deterred from engaging in crime in SDSC districts may do so elsewhere.

Though important to take seriously, we think the magnitude of any spillover effect is likely to be small for the outcomes of interest, for two reasons. First, the SDSC intervention is designed to change within-district resource allocations—which district units are deployed where and when—rather than across-district allocations. As discussed in Section IV, the SDSC provides a district Commander with information, analytical capacity, and a management process with which to implement a proactive policing strategy. The district Commander does not have command over the Area units operating in his or her district, nor do data from the GPS units within each vehicle in CPD’s fleet and attendance records suggest there was an influx of resources into districts, either from the Area or from elsewhere in the city, following the SDSCs’ introduction (see, e.g., 7<sup>th</sup> district, Figure 15).

Second, the primary outcome of interest—violent crime, and particularly gun violence—is among the crimes least likely to be susceptible to displacement by enhanced policing. Gun violence is often tied to a feature of a specific location that is immobile in the short run, such as a lucrative drug-selling corner or the block of a rival street organization member, or it occurs following an altercation. In neither case is such gun violence likely to move several blocks away if deterred from occurring in its original location. This is consistent with the best available evidence on crime

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<sup>45</sup> This is related to but slightly different from the approach employed by Robbins, Saunders, and Kilmer (2017), who generate placebo areas using a permutation technique that groups together many comparison areas that are smaller (block-level) than their treatment area (neighborhood-level). As a result, their placebo areas are random assortments of small comparison areas that lack the structure of the treatment area, requiring them to standardize their effect estimates to guard against the resulting bias. In contrast, we avoid this issue by creating placebo districts using a resampling procedure, wherein we resample beats from existing comparison districts with replacement, preserving the structure of those comparison districts in the process.

<sup>46</sup> The sign of any bias resulting from such a reallocation is ambiguous. If the SDSC reduces violence in a treated district, then resources may be shifted from there to districts without an SDSC. On the other hand, the Department may have an incentive to bolster the SDSC effort by allocating additional officers to those districts.

displacement, which detects adverse spillovers for property crime but not violent crime (Blattman et al. 2019).

## VI. THE IMPACT OF AN SDSC ON SERIOUS CRIME

We turn to estimating the impact of the SDSCs on serious crime in each of the six Tier 1 districts using the modified DI estimator described in the previous section. Our focus is on two primary measures of serious crime: shooting victims and Part I violent felonies per capita.<sup>47</sup> We first estimate synthetic controls separately for each of these outcomes in each of the SDSC districts. Then, we estimate a single synthetic control for each SDSC district that fits to these primary outcomes along with four secondary outcomes: gun arrests, warrant arrests, overall arrests, and traffic stops per capita.<sup>48</sup> The first set of estimates serve as our main results for the impact of the SDSCs on serious crime. The second set of estimates serve as both a robustness check and as a way to understand more about the potential mechanisms driving the primary results.

Because we conduct multiple hypothesis tests throughout this analysis, we must correct for multiple comparisons. We do this in two ways. First, we control for the familywise error rate (FWER), or the probability of falsely rejecting one or more true null hypotheses (type I error). Second, we control for the false discovery rate (FDR), or the proportion of true null hypotheses that are falsely rejected among the total number of rejected hypotheses (“discoveries”). Controlling for the FWER is generally the more conservative multiple comparison procedure, as it limits the probability of making any type I error. In contrast, controlling for the FDR is generally the less conservative multiple comparison procedure, as it limits the fraction of false rejections among the set of discoveries. We correct for multiple comparisons across districts and outcomes separately for the main and secondary sets of results, reporting multiple comparisons-corrected  $q$ -values alongside uncorrected  $p$ -values.<sup>49</sup>

Table 1 reports estimates of the SDSCs’ impacts across several outcomes in the 7<sup>th</sup> district. The 7<sup>th</sup> district experienced a 26 percent decline in shooting victims and a 10 percent decline in Part I violent felony incidents following the introduction of the SDSC. This implies that there were 75 fewer shooting victims in the 7<sup>th</sup> district during the 11 months of 2017 after the SDSC was introduced than there would have been otherwise. For both outcomes, the synthetic controls are able to closely match the observed pre-treatment patterns in the 7<sup>th</sup> district, and the placebo tests

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<sup>47</sup> Part I violent felonies include homicide, criminal sexual assault, robbery, aggravated assault, and aggravated battery.

<sup>48</sup> A warrant arrest is an arrest made on the basis of an outstanding warrant being located in the Law Enforcement Agencies Data System (LEADS) database. A gun arrest is an arrest of a person found to illegally possess a firearm or after the use of a firearm in a shooting that did not result in injury to a person.

<sup>49</sup> To control for the FWER, we use the Holm-Bonferroni correction (Holm 1979), which is uniformly more powerful than the classic Bonferroni correction and does not rely on assumptions about the independence of the tests being performed, as does the more powerful procedure suggested by Hochberg (1988). To control for the FDR, we use the Benjamini-Yekutieli procedure, which similarly does not rely on assumptions about the independence of the tests being performed (Benjamini and Yekutieli 2001). Both corrections are implemented using the Stata `qvalue` package. For the main set of results, these corrections account for 12 comparisons (6 districts, 2 primary outcomes per district). For the secondary set of results, these corrections account for 36 comparisons (6 districts, 2 primary and 4 secondary outcomes per district).

make clear that these reductions are outliers relative to their respective placebo effect distributions (Figures 16 and 17), though neither effect is statistically distinguishable from zero after correcting for multiple comparisons. We attempt to detect changes in officer productivity by estimating a single synthetic control jointly across both the primary and secondary outcomes. Using this method, the treatment effects for the primary outcomes remain qualitatively similar and the fit of the synthetic control remains good (Figure 18). The treatment effect estimates for the secondary outcomes suggest substantial increases in several measures of officer activity: gun arrests (28 percent), warrant arrests (36 percent), overall arrests (20 percent), and traffic stops (77 percent). Some of these increases appear to predate the introduction of the SDSC, others exhibit a poor quality fit for the synthetic control, and many are not statistically distinguishable from zero after correcting for multiple comparisons. Nevertheless, their co-movement broadly suggests a change in the behavior of 7<sup>th</sup> district officers that coincided with the district's substantial reductions in violence.

Table 2 reports analogous estimates for the 11<sup>th</sup> district. Relative to the 7<sup>th</sup> district, the 11<sup>th</sup> district experienced a more modest 14 percent reduction in shooting victims and a slightly larger 13 percent reduction in Part I violent felony incidents; only the latter estimate is statistically distinguishable from zero after adjusting for multiple comparisons. More worryingly, neither of these estimates are robust to estimation of the synthetic control jointly across both primary and secondary outcomes, with the estimated treatment effects changing sign. Similarly, the quality of the synthetic control fit, while high when estimated separately for the primary outcomes (Figures 19 and 20), degrades when estimated jointly across primary and secondary outcomes (Figure 21).

Tables 3-6 report estimates for the remaining SDSC districts: the 6<sup>th</sup>, 9<sup>th</sup>, 10<sup>th</sup>, and 15<sup>th</sup>. Most of the treatment effect estimates for these districts, across both primary and secondary outcomes, are very imprecisely estimated. The notable exception is the 15<sup>th</sup> district, where the treatment effect estimate suggests that shooting victims *increased* following the introduction of the SDSC. To assess the credibility of this result, we must examine the quality of the pre-period fit to see whether the synthetic control is able to approximately match the outcome series for shooting victims, particularly immediately prior to the treatment. When the synthetic control is estimated separately for the shooting victim outcome, it appears to underestimate the observed pattern in the 15<sup>th</sup> district in 2016, just prior to the SDSC's introduction (Figure 22). When the synthetic control is estimated jointly for the primary and secondary outcomes, this prediction error increases substantially (Figure 23). These deviations cause the synthetic control to significantly underestimate the true series, partially explaining why we estimate an adverse treatment effect in the post-period.

Perhaps the clearest change in police behavior induced by the SDSCs is in arrests initiated by officers monitoring the Department's POD cameras. As described in Section IV, POD cameras predated the SDSCs, though their number and quality—from standard definition to high definition—both expanded as part of the intervention. Furthermore, the cumbersome software previously available for monitoring the PODs was replaced by an easy-to-use platform that made them more accessible to officers. As a result, with each wave of new SDSCs—in Tier 1 districts in early 2017, and in Tier 2 districts in early 2018—the number of POD-initiated arrests increased substantially (Figure 24).

## VII. CONCLUSION

We study the impact of SDSCs: a package of reforms to basic management practices within several high-violence police districts in Chicago. These reforms are intended to help CPD execute a proactive policing strategy that calls for focusing resources on the places and people at highest risk for serious crime. Though outwardly an intervention focused on data and technology enhancements, at their core the SDSCs mostly layer management improvements—tasking one person to oversee data analysis, setting up a process for analysis to inform decision-making—atop a data infrastructure that predates them.

Our results suggest that any reductions in serious crime stemming from the SDSCs were not universally shared across the districts that received them. The 7<sup>th</sup> district experienced substantial decreases in shooting victims and Part I violent felonies that, while rendered imprecise by adjustment for multiple comparisons, are nevertheless robust to joint estimation with secondary outcomes. In contrast, the estimated reductions in the 11<sup>th</sup> district were not similarly robust. Results in the other five SDSC districts were either imprecisely estimated or inconclusive. This underscores an important point about the SDSC intervention: as a reform meant to improve implementation of an existing policing strategy, it, too, may be subject to imperfect implementation.

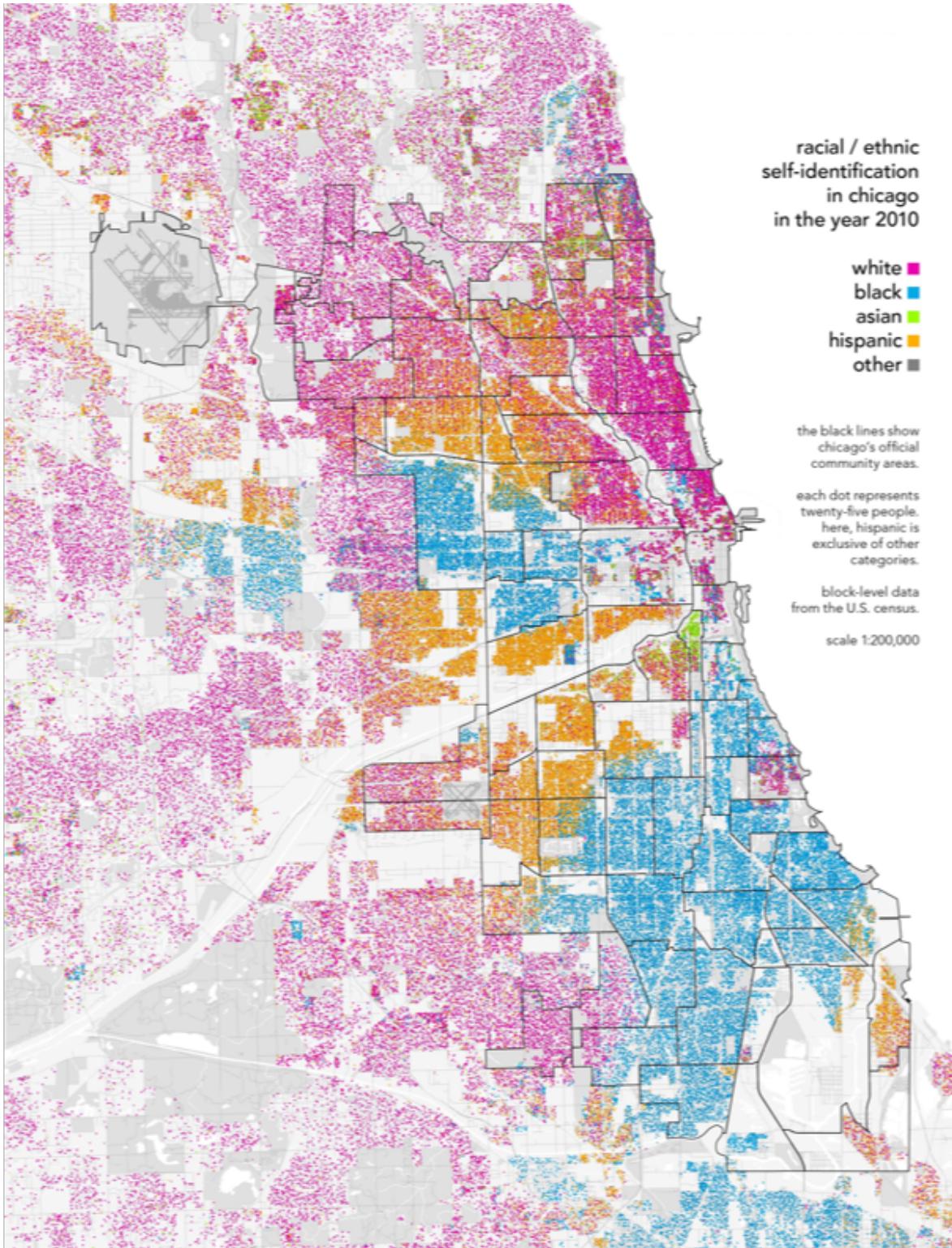
We will continue to study which aspects of the SDSC appear to drive the crime improvements experienced in the 7<sup>th</sup> district. One potentially important factor is community engagement, the third pillar of the proactive policing strategy described earlier. Although not directly a focus of the SDSC intervention, some district Commanders, particularly then-Commander Kenneth Johnson of the 7<sup>th</sup> district, made a point of emphasizing the importance of officers engaging in positive interactions with members of the community. Commander Johnson’s emphasis that his officers do this may be reflected, in part, in the staggering increase in the volume of community interaction calls, or PCIs, they made following the SDSC’s implementation (Figure 25). Of course, we do not mean to suggest that this increase is entirely or even mostly reflective of greater engagement with community members by officers. However, it is noteworthy that this increase was by far the largest in the 7<sup>th</sup> district, suggesting either a real change in officer behavior, a greater willingness to follow through on the Commander’s orders, or some combination of the two that merits further study to understand what role, if any, they played in the serious crime reduction there.

These results also highlight an important dimension of potential policing reform that seldom receives attention in the academic literature. Many studies on policing and crime control focus on estimating how increasing the *size* of the former may affect the latter. A separate literature is concerned with the efficacy of the specific tactics and strategies employed by police (Weisburd and Eck 2004). Both of these literatures offer practical guidance to policymakers, but they are of limited use when the level of policing resources in the short run is fixed, and when existing strategy and tactics already reflect best practices. In this scenario, which Chicago and other cities find themselves in, the available margin for improvement is to strengthen the implementation of the existing strategy, such as through reforms to management practices that can improve productivity.

Perhaps the literature most relevant for understanding which interventions might improve policing and reduce crime on the implementation margin is not the literature about policing and crime at all, but rather the literature about the relationship between management practices and productivity (Bloom and Van Reenen 2007; Bloom et al. 2013). This literature, broadly speaking, suggests that adopting superior management practices can have a sizeable effect on firms' productivity. It also suggests that these practices may not be adopted on their own due to information barriers, a phenomenon that is common not only in policing but other municipal agencies as well.

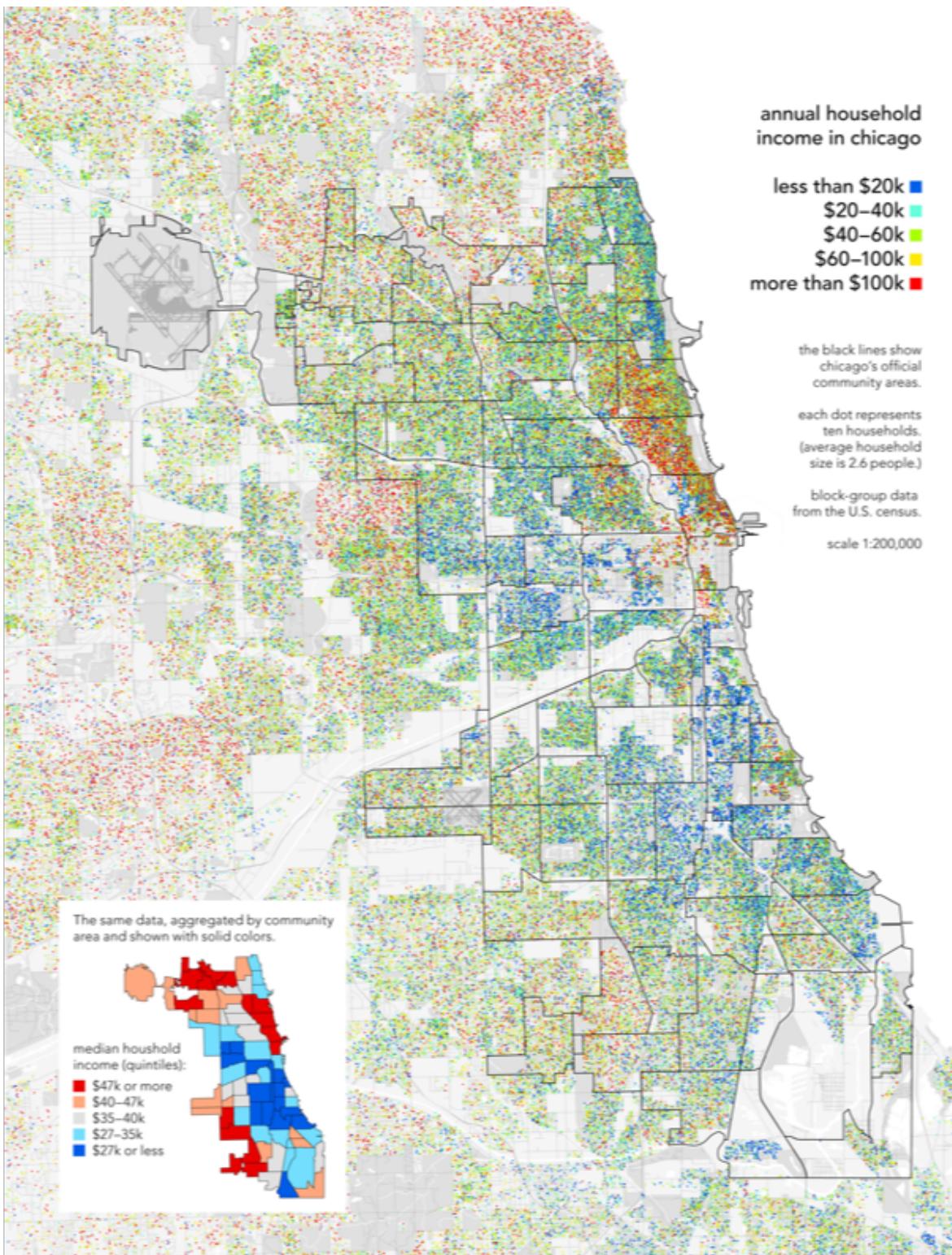
Finally, it is worth noting that, even if the SDSCs' implementation was uneven, they may still represent a worthwhile investment. The cost to build an SDSC is approximately \$2 million, with ongoing operating costs likely to be substantially less than this figure. If the SDSC in the 7<sup>th</sup> district was able to prevent 94 shooting victimizations from occurring, then using an estimated cost per gunshot injury derived from contingent valuation methods of \$1.5 million in 2019 dollars (Cook and Ludwig 2000), that one SDSC generated savings of \$141 million in a single year. Compared to CPD's annual budget of \$1.6 billion in 2018, such investments in better management are modest and seem likely to pay for themselves.

**Figure 1.** Racial segregation across Chicago's community areas



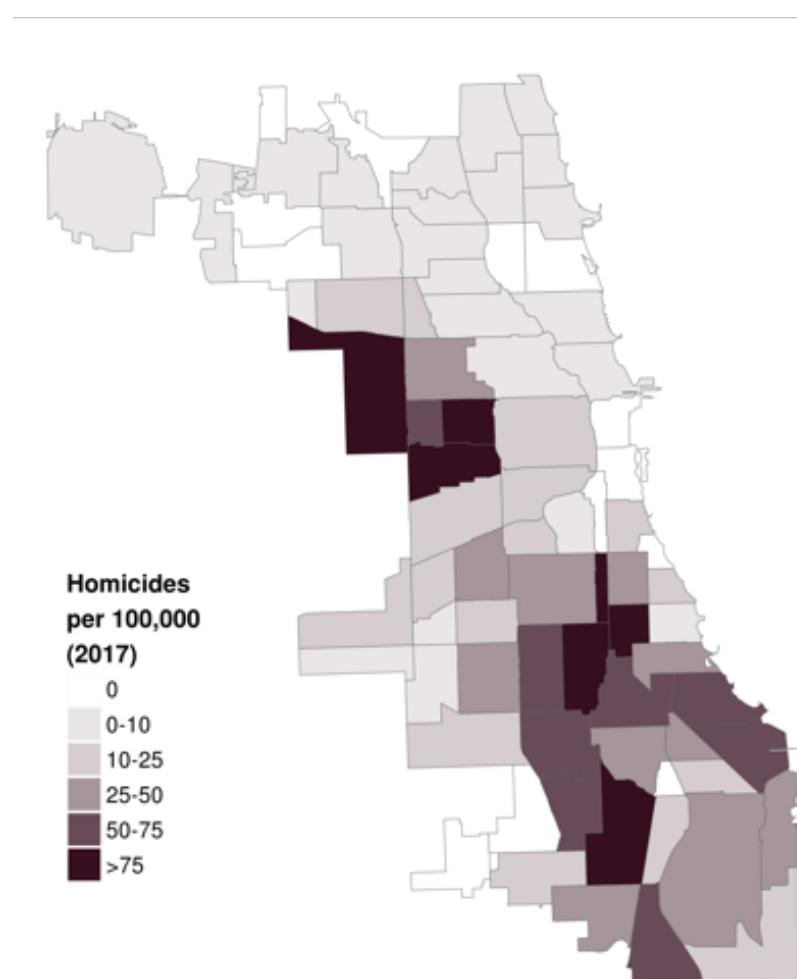
**Notes:** Bill Rankin (2009). <http://www.radicalcartography.net/index.html?chicagodots>

**Figure 2.** Income segregation across Chicago's community areas



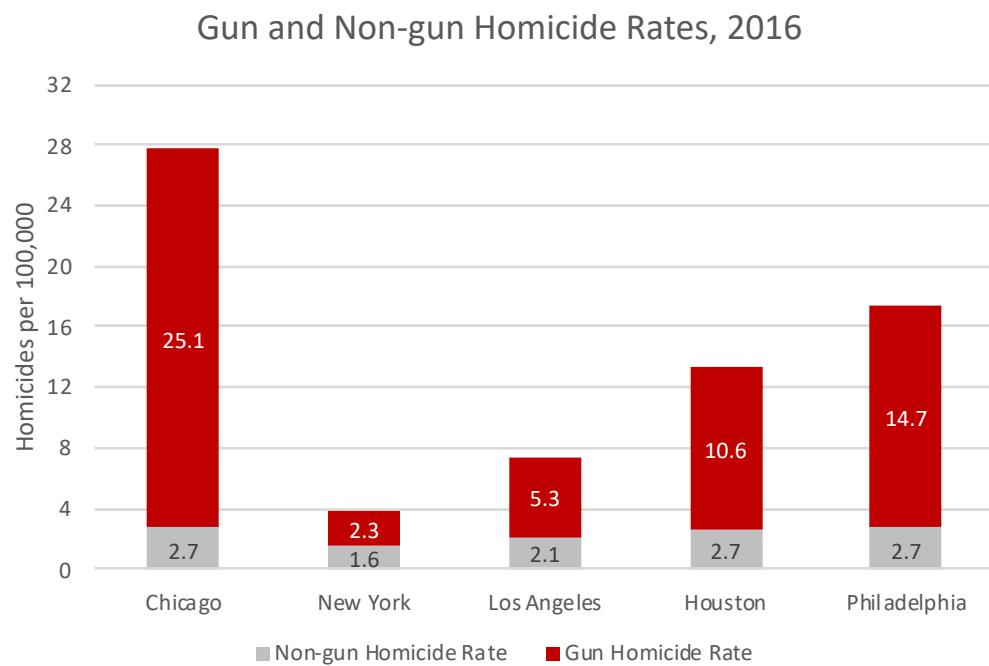
**Notes:** Bill Rankin (2009). <http://www.radicalcartography.net/index.html?chicagodots>

**Figure 3.** Homicide victims per 100,000 across Chicago's community areas, 2017



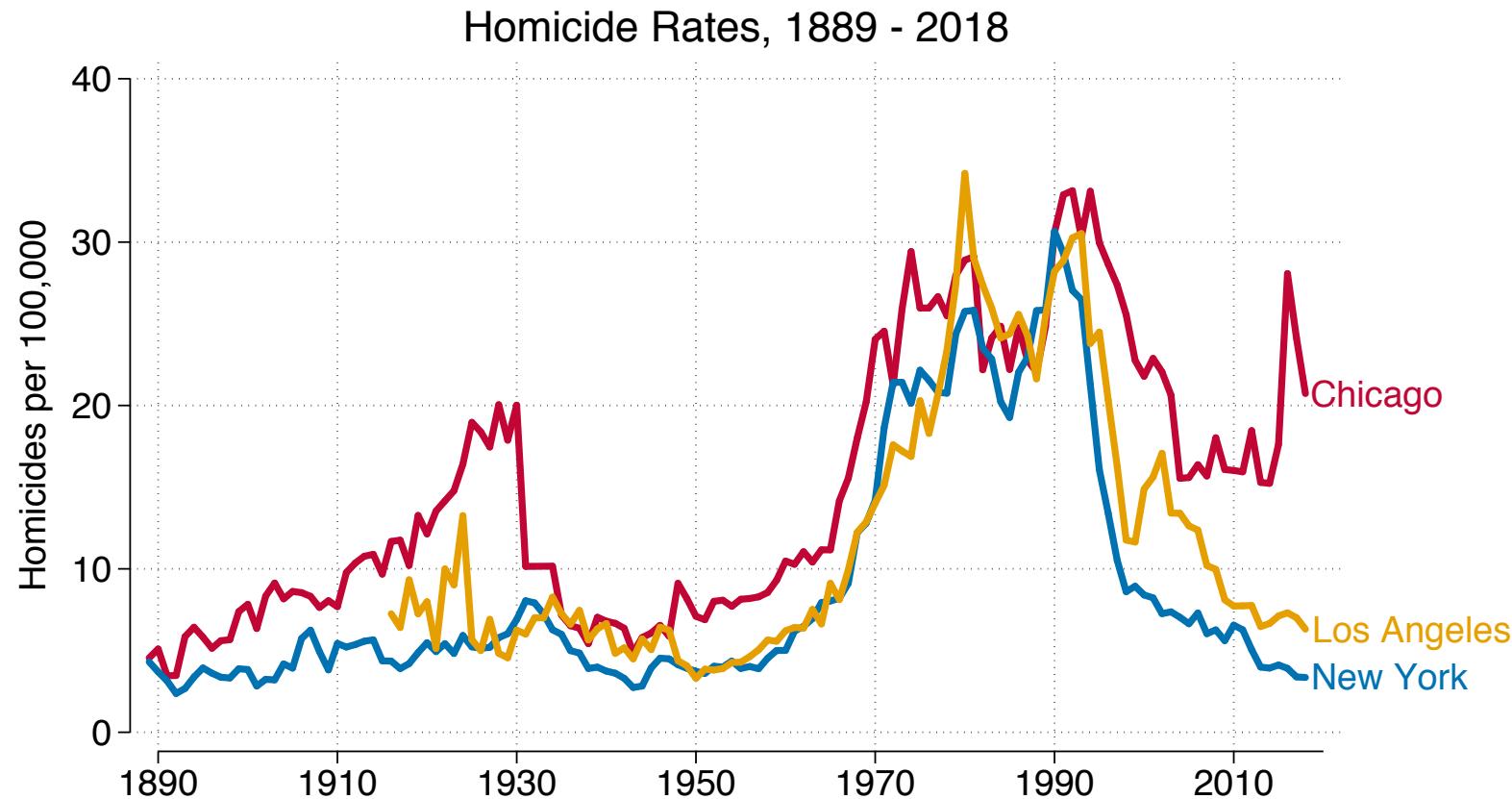
**Notes:** CPD data.

**Figure 4.** Gun and non-gun homicide rates among large American cities, 2016



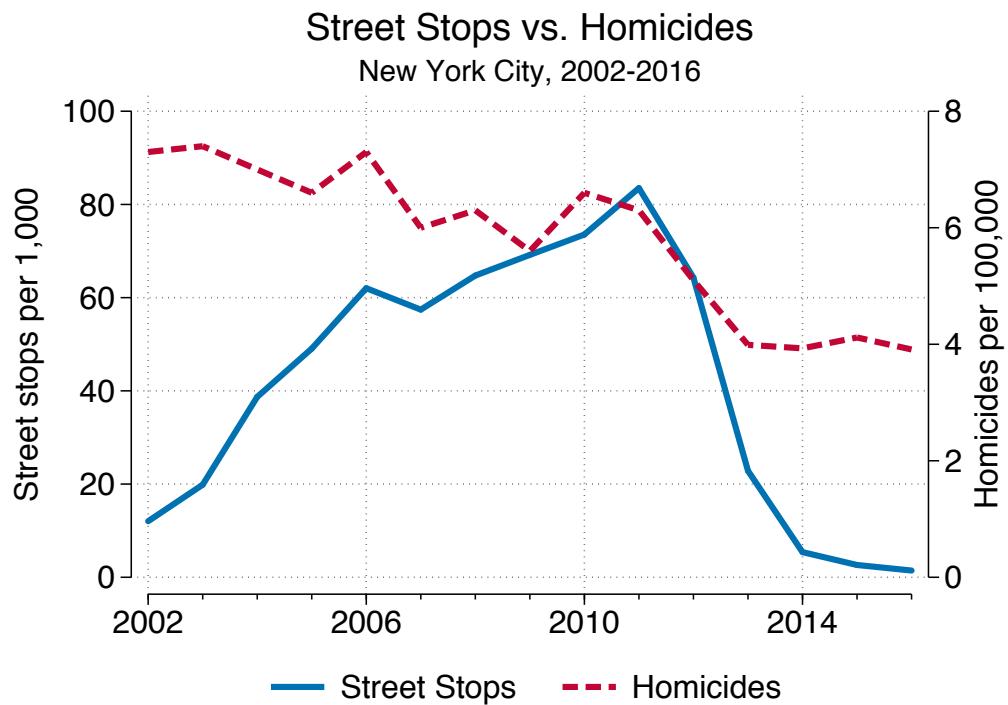
**Notes:** Records provided by the police departments of Chicago, New York City, Los Angeles, Houston, and Philadelphia.

**Figure 5.** Long-run homicide rates in New York City, Los Angeles, and Chicago, 1889 – 2018



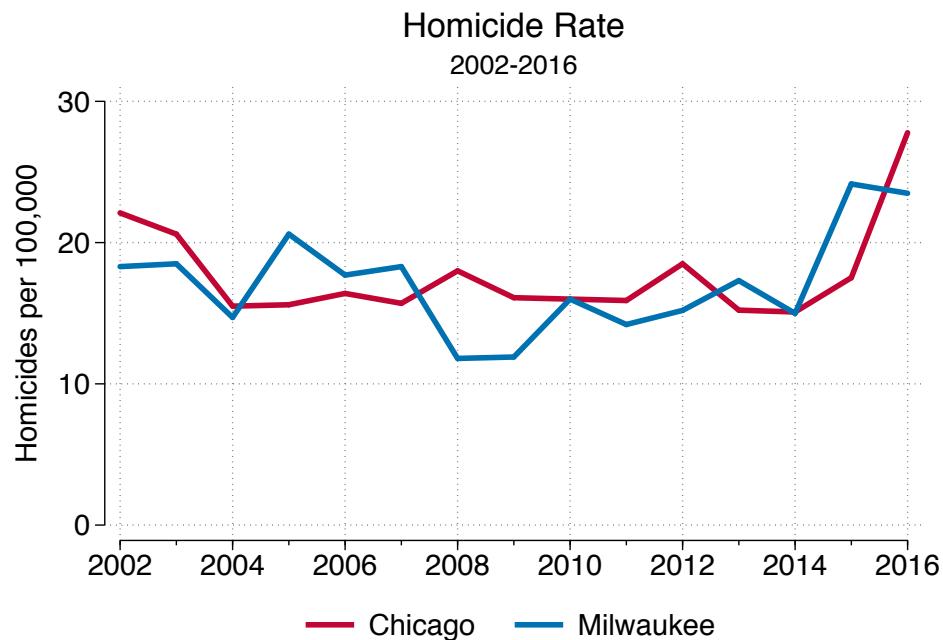
**Notes:** Chicago homicide data for 1889 through 1930 from the Chicago Historical Homicide Project at Northwestern University (<http://homicide.northwestern.edu/>). Chicago homicide data for 1930 through 1959 from the FBI's Uniform Crime Reports ([ICPSR 3666](#)). Los Angeles homicide data for 1916 through 1959 from the Historical Violence Database at the Criminal Justice Research Center, the Ohio State University (<https://cjrc.osu.edu/research/interdisciplinary/hvd>). New York City homicide data for 1890 through 1959 from the National Institute of Justice ([ICPSR 3226](#)). Homicide data for 1960 through 2017 from the FBI's Uniform Crime Reports ([Open ICPSR](#)). Homicide data for 2018 from the police departments of Chicago, New York City, and Los Angeles.

**Figure 6.** Street stops and homicides in New York City, 2002 – 2016



**Notes:** Homicide data from the FBI's Uniform Crime Reports. Stop, question, and frisk data from the New York Police Department.

**Figure 7.** Comparison of homicide rates in Chicago and Milwaukee, and street stop rates in Milwaukee, 2002 – 2016



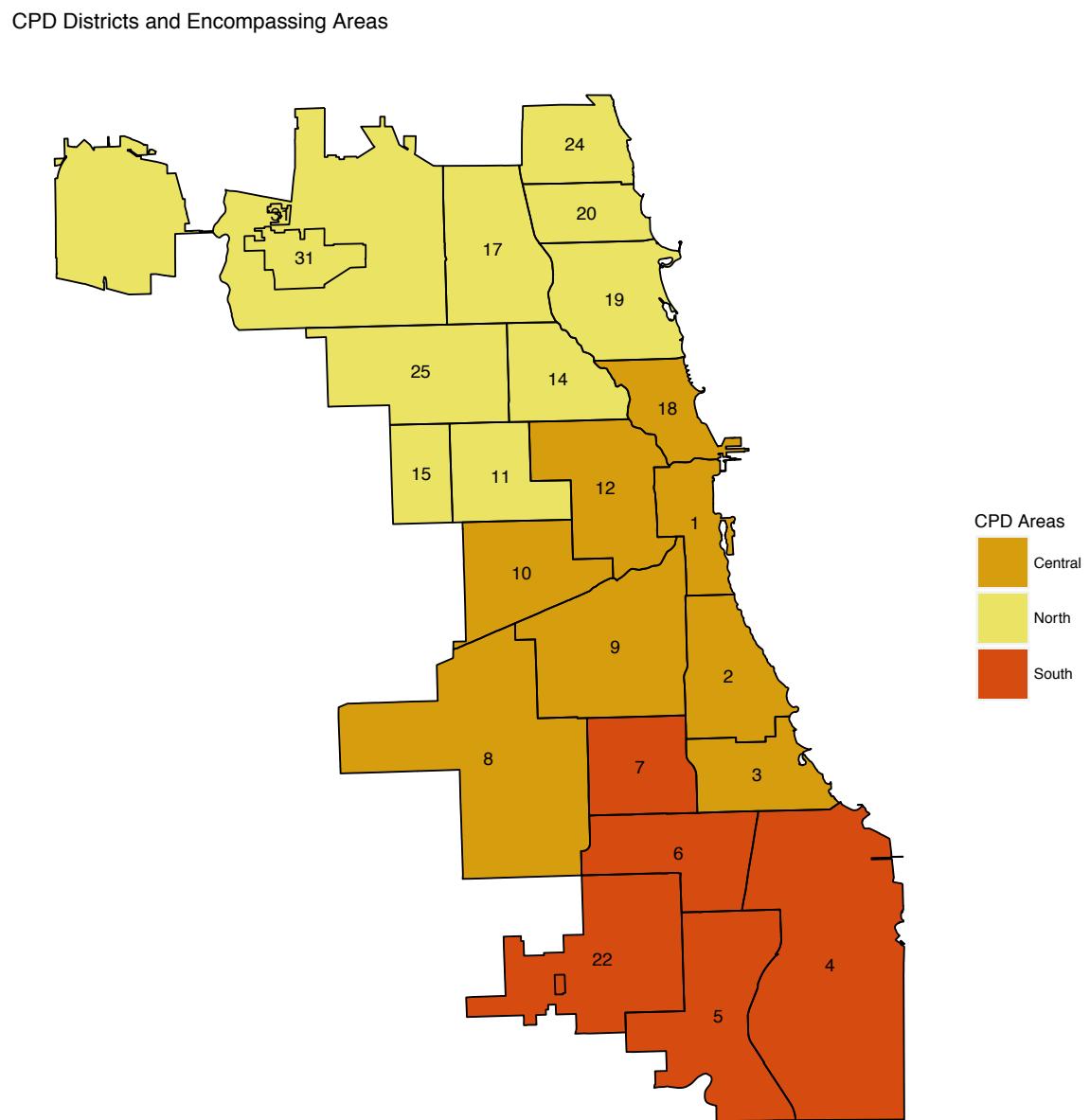
*Panel A*



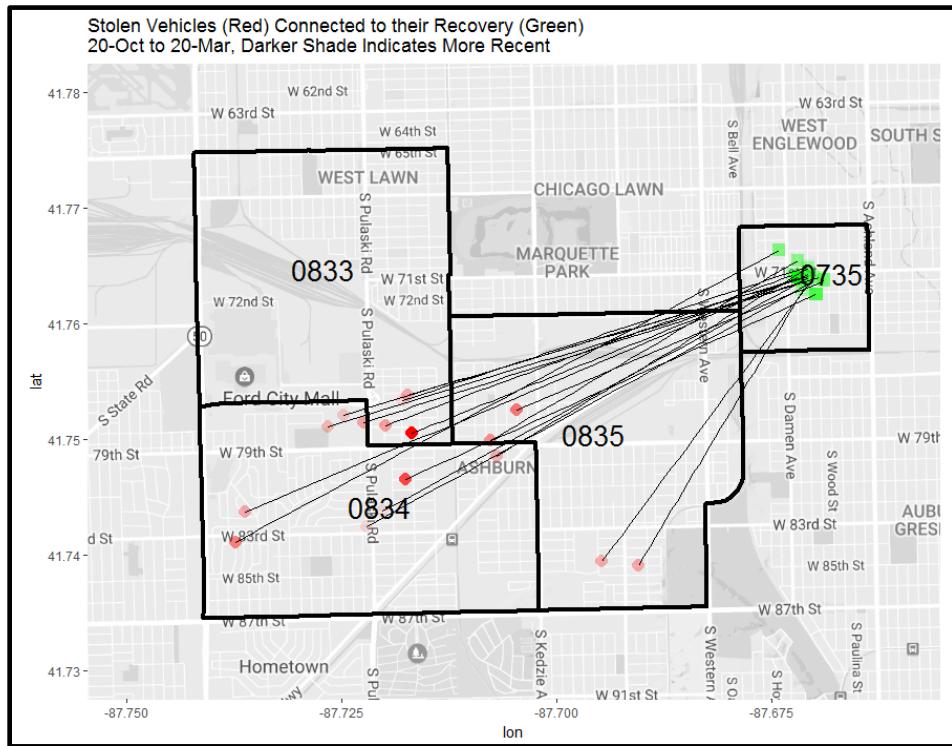
*Panel B*

**Notes:** Homicide data from the FBI's Uniform Crime Reports. Stop data from Milwaukee PD.

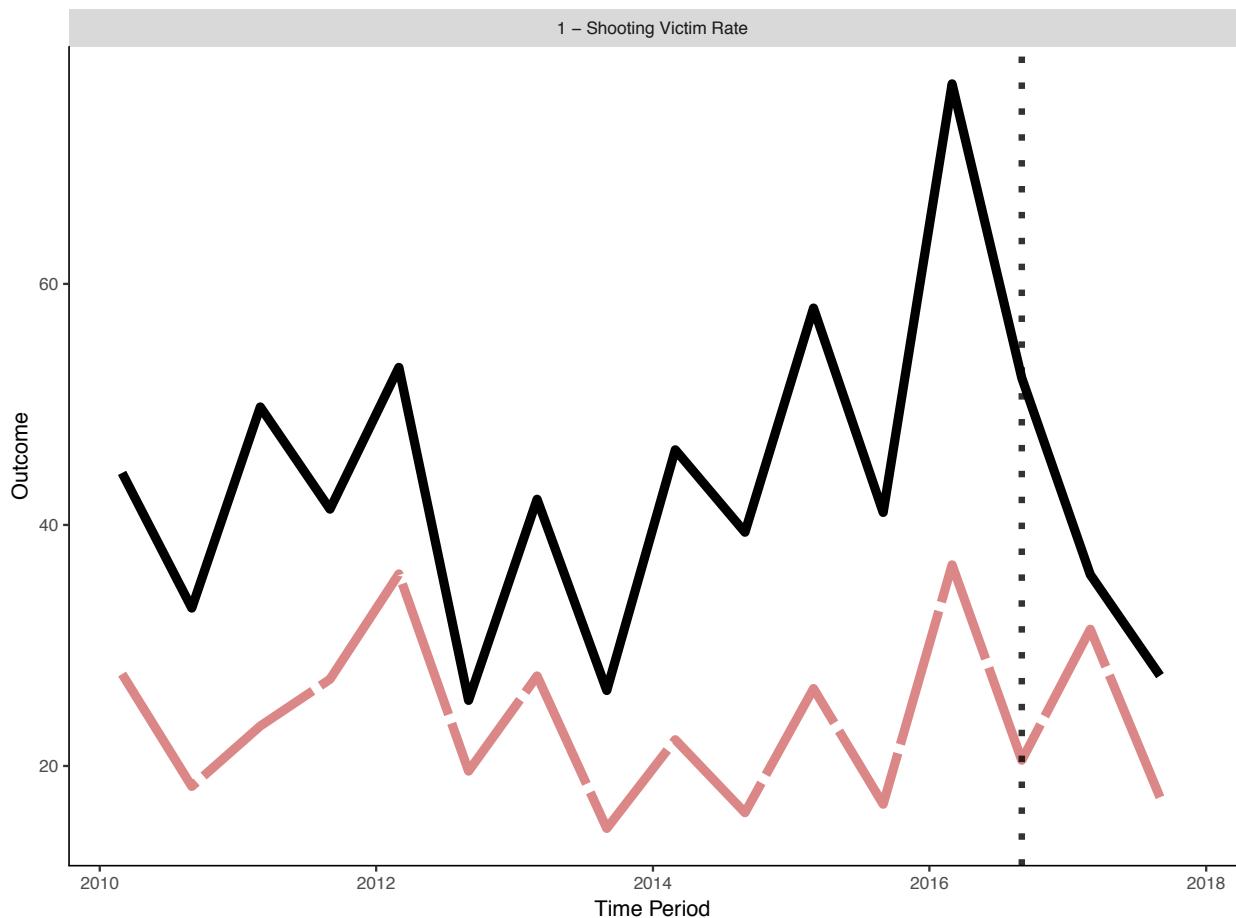
**Figure 8.** Chicago Police Department patrol areas and districts



**Figure 9.** Sample analysis involving motor vehicle thefts from the 7<sup>th</sup> district

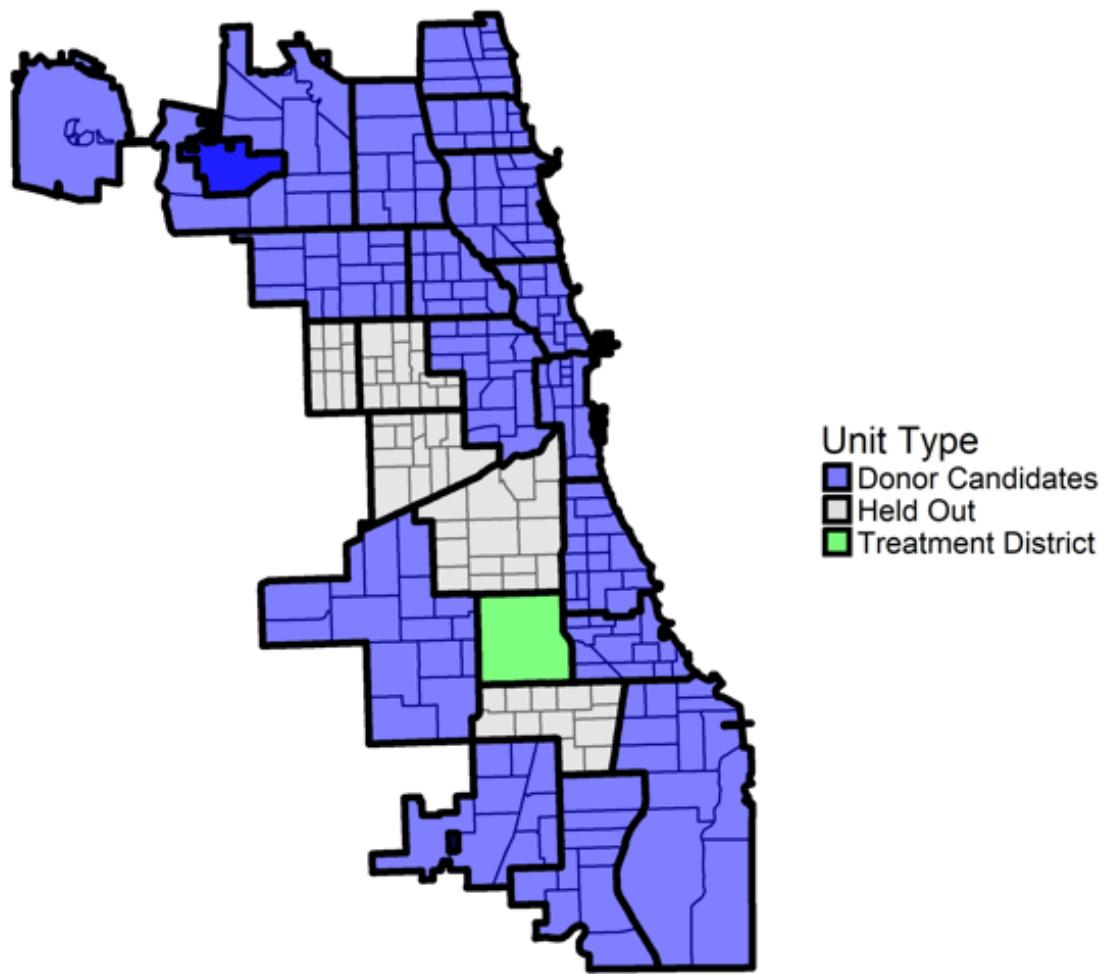


**Figure 10.** Shootings per capita in 7<sup>th</sup> district and ADH synthetic control, districts as donors



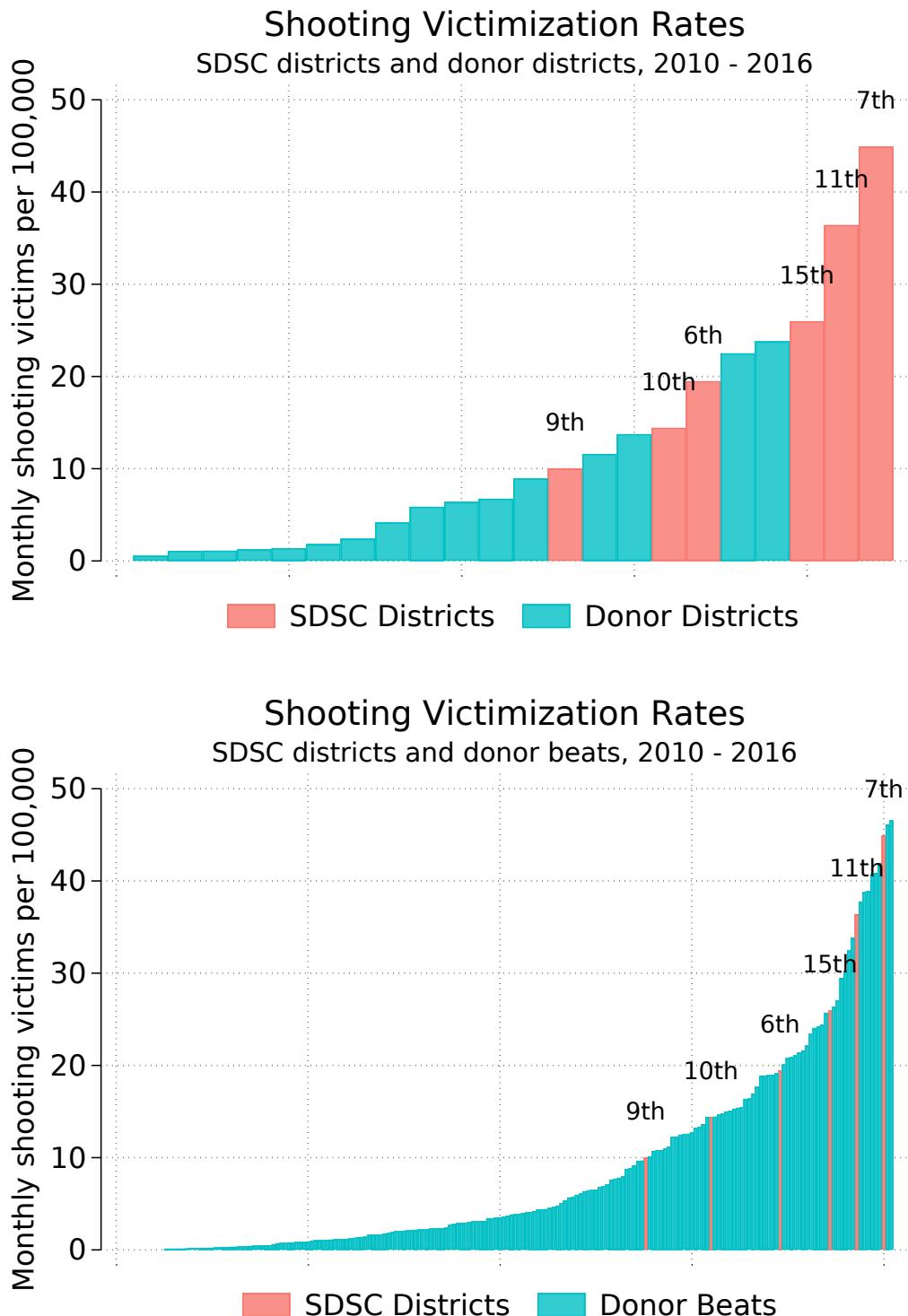
**Notes:** Black line denotes biannual monthly average shootings per capita in the 7<sup>th</sup> district. Red line denotes synthetic control using the ADH estimator and 16 non-SDSC districts as donors. Dotted vertical line represents last pre-period.

**Figure 11.** Chicago Police Department districts and beats



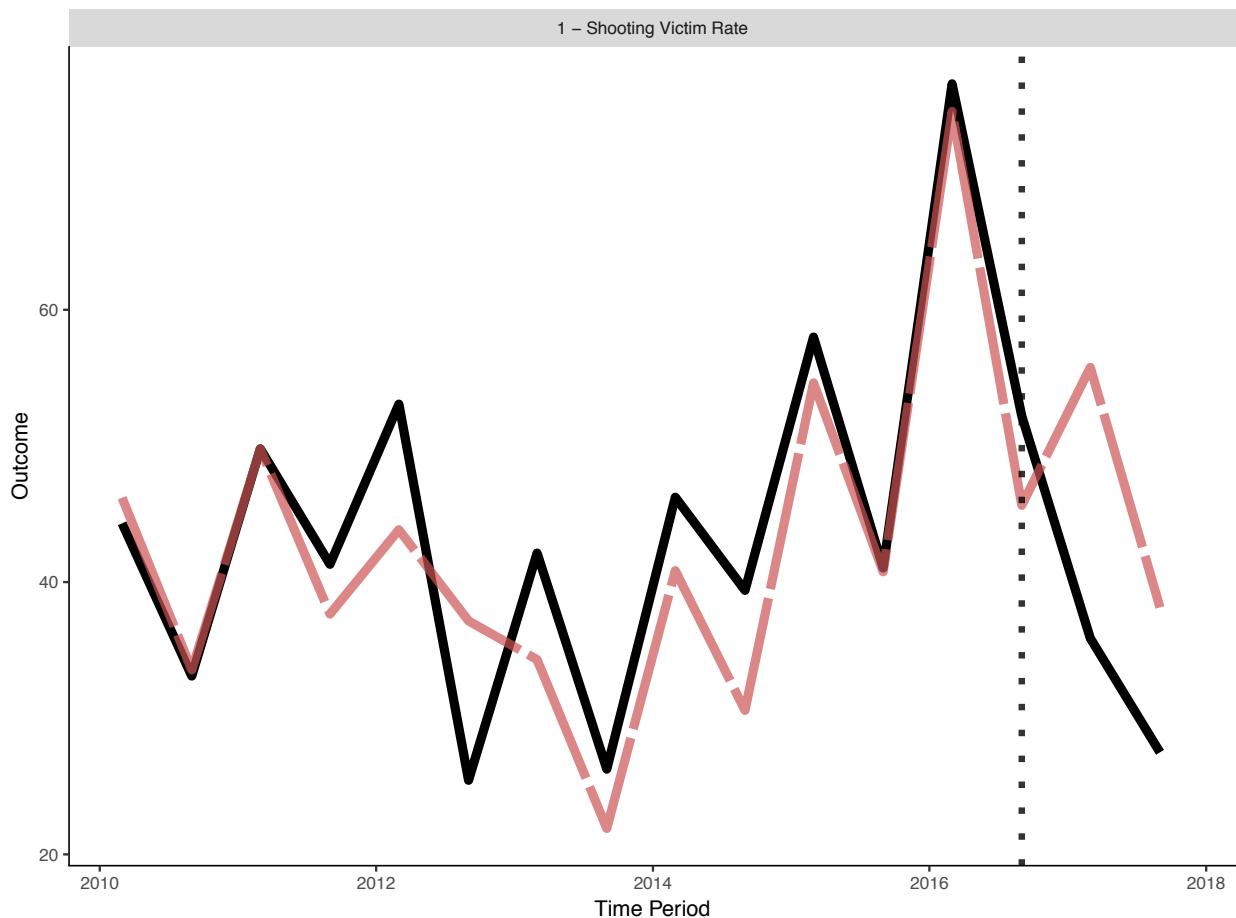
**Notes:** Chicago Police Department district boundaries (bold) and beat boundaries. In this example, the 7<sup>th</sup> district is shaded green; the other five SDSC districts are shaded gray and excluded from the donor pool; and the 196 beats from among the 16 non-SDSC donor districts shaded purple comprise the donor pool.

**Figure 12.** Distributions of monthly shooting victims per capita



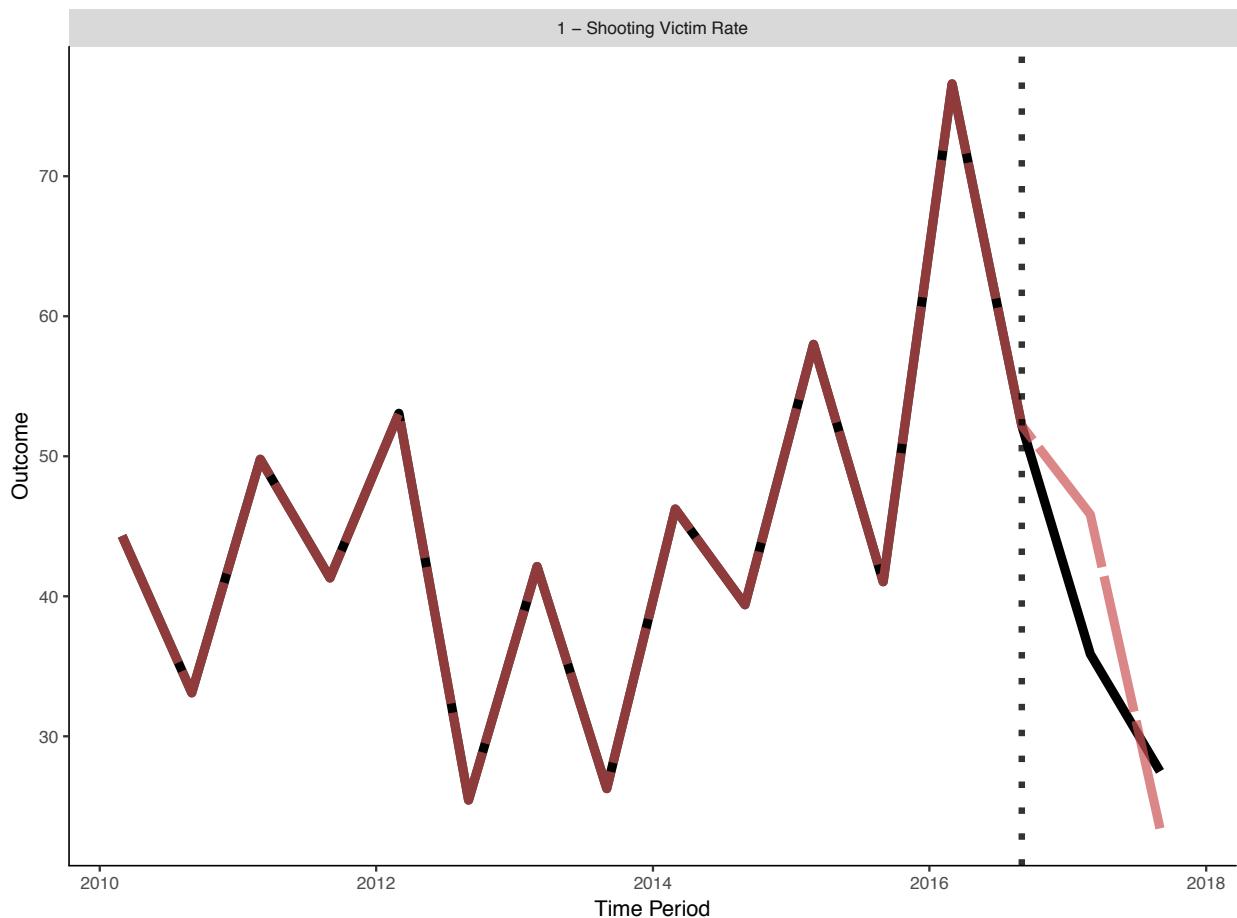
**Notes:** Figure shows distributions of average monthly shooting victims per 100,000 in the pre-treatment period for SDSC districts and donor districts (top panel) or donor beats (bottom panel).

**Figure 13.** Shootings per capita in 7<sup>th</sup> district and ADH synthetic control, beats as donors



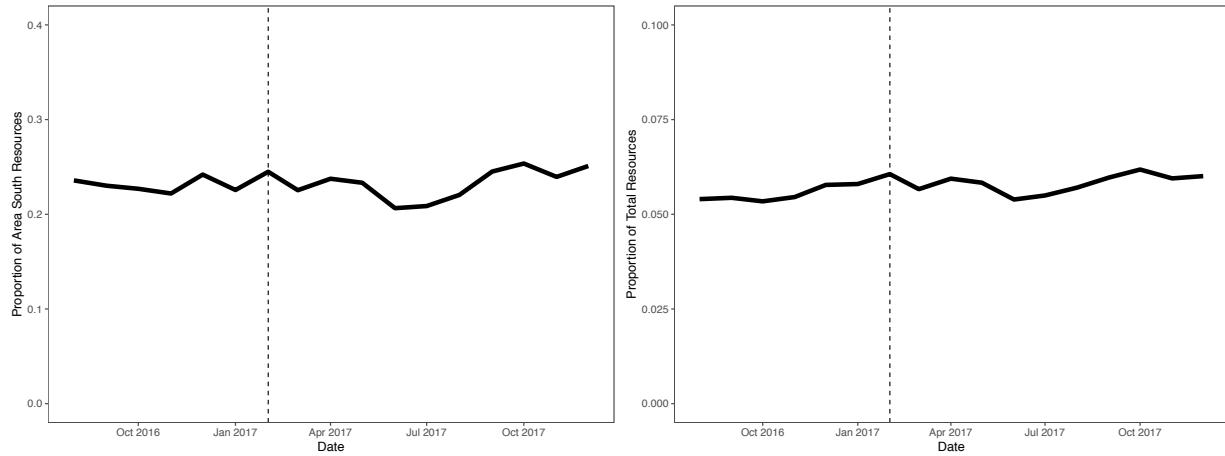
**Notes:** Black line denotes biannual monthly average shootings per capita in the 7<sup>th</sup> district. Red line denotes synthetic control using the ADH estimator and 196 non-SDSC beats as donors.  
Dotted vertical line represents last pre-period.

**Figure 14.** Shootings per capita in 7<sup>th</sup> district and DI synthetic control, beats as donors

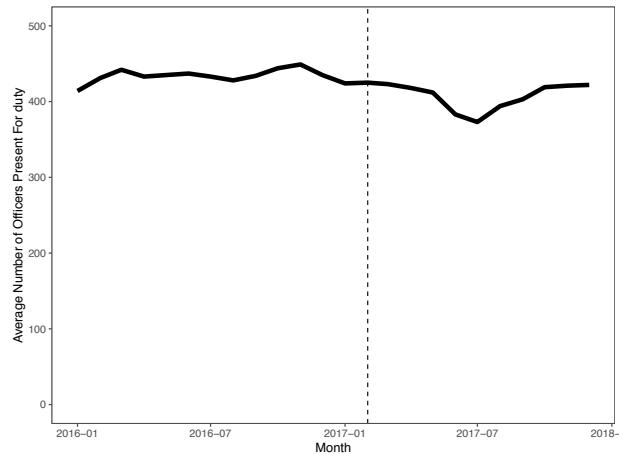


**Notes:** Black line denotes biannual monthly average shootings per capita in the 7<sup>th</sup> district. Red line denotes synthetic control using the DI estimator and 196 non-SDSC beats as donors. Dotted vertical line represents last pre-period.

**Figure 15.** Officer working in the 7<sup>th</sup> district



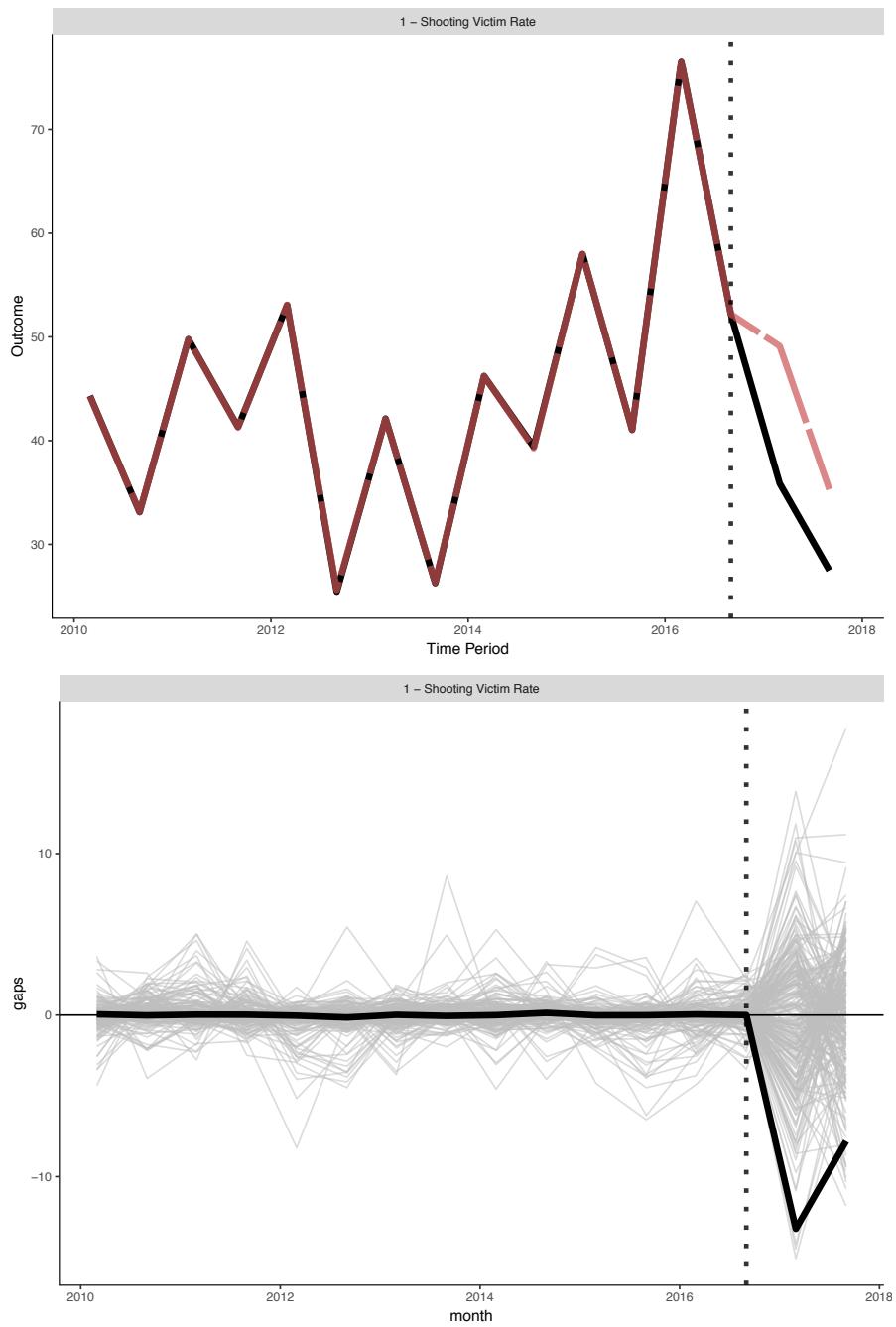
*Panel A. Share of all officers in Area South (left) and citywide (right) in the 7<sup>th</sup> district, GPS data*



*Panel B. Number of officers present for duty in the 7<sup>th</sup> district, attendance roster data*

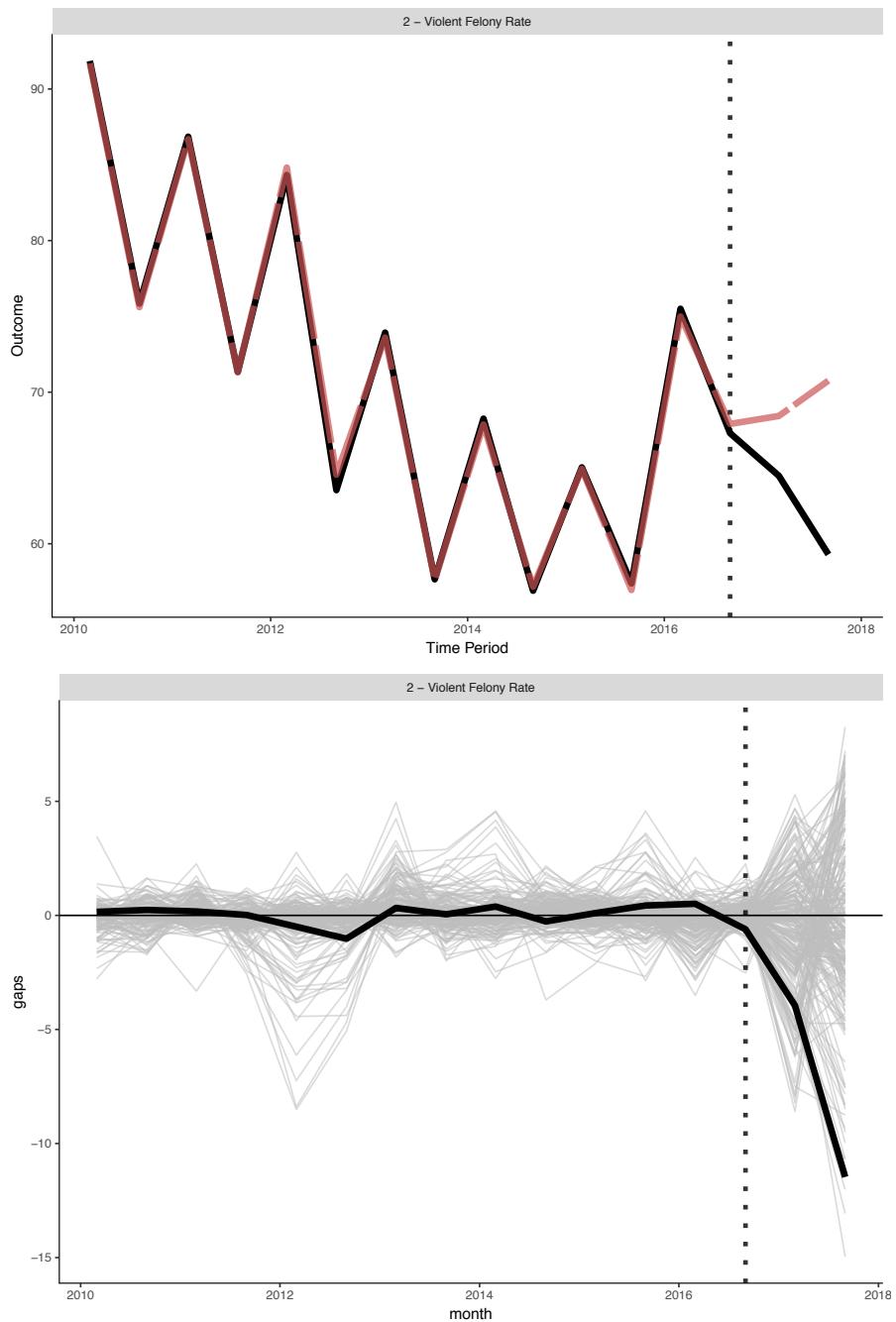
**Notes:** Figures in Panel A show the share of officers (at the Area level and citywide) operating in the 7<sup>th</sup> district each month, based on data from GPS units located in their vehicles. As a robustness check, the figure in Panel B relies on data from attendance rosters to record the number of unique CPD employees reporting to work in the 7<sup>th</sup> district.

**Figure 16.** Synthetic control results for the 7<sup>th</sup> district: shooting victims



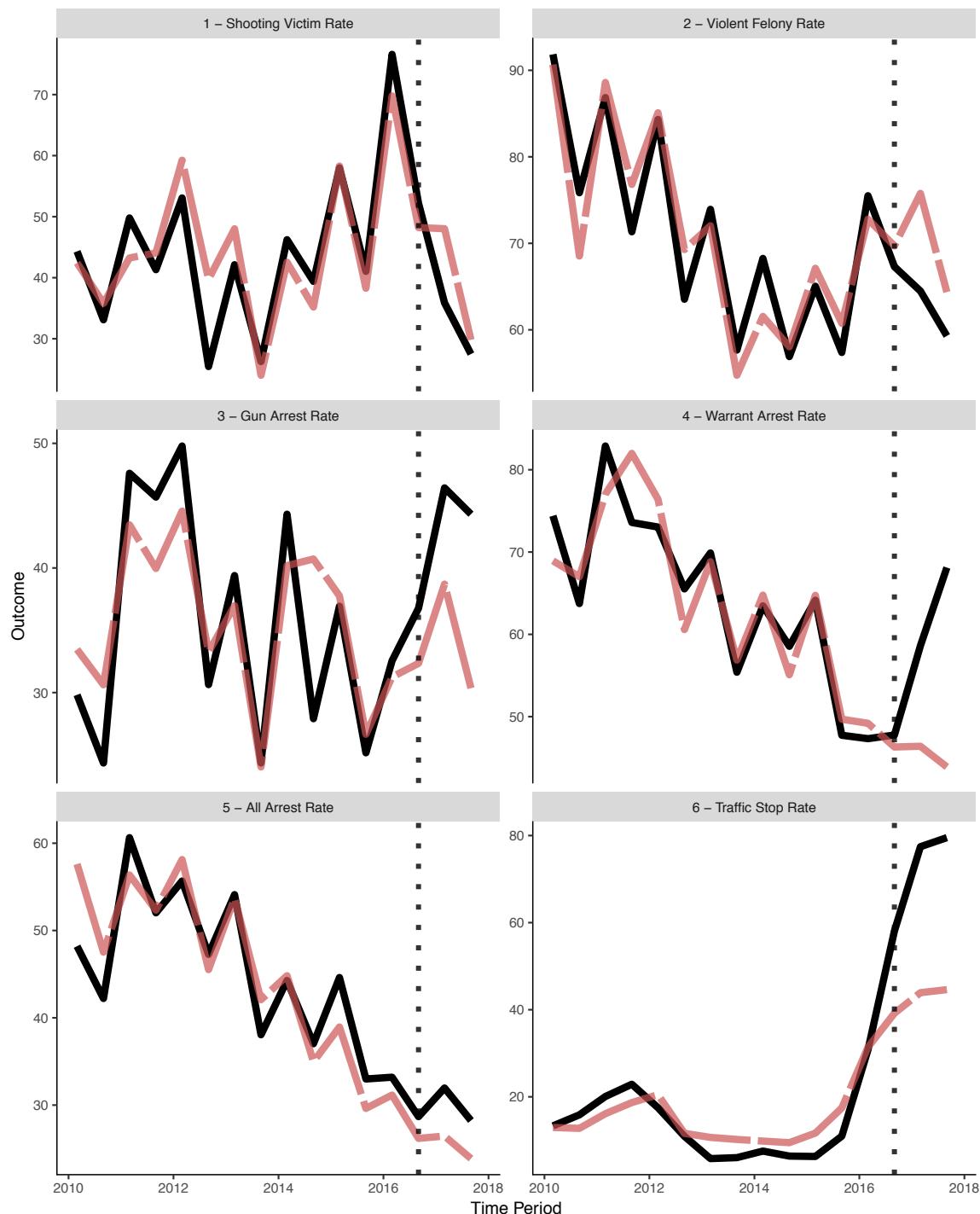
**Notes:** Synthetic control for the 7<sup>th</sup> district, using modified DI estimator and beats as donors, in red. Placebo test results plot differences between the actual and synthetic control unit, for the 7<sup>th</sup> district (bold line) and the resampled donor districts (light gray lines).

**Figure 17.** Synthetic control results for the 7<sup>th</sup> district: Part I violent felonies



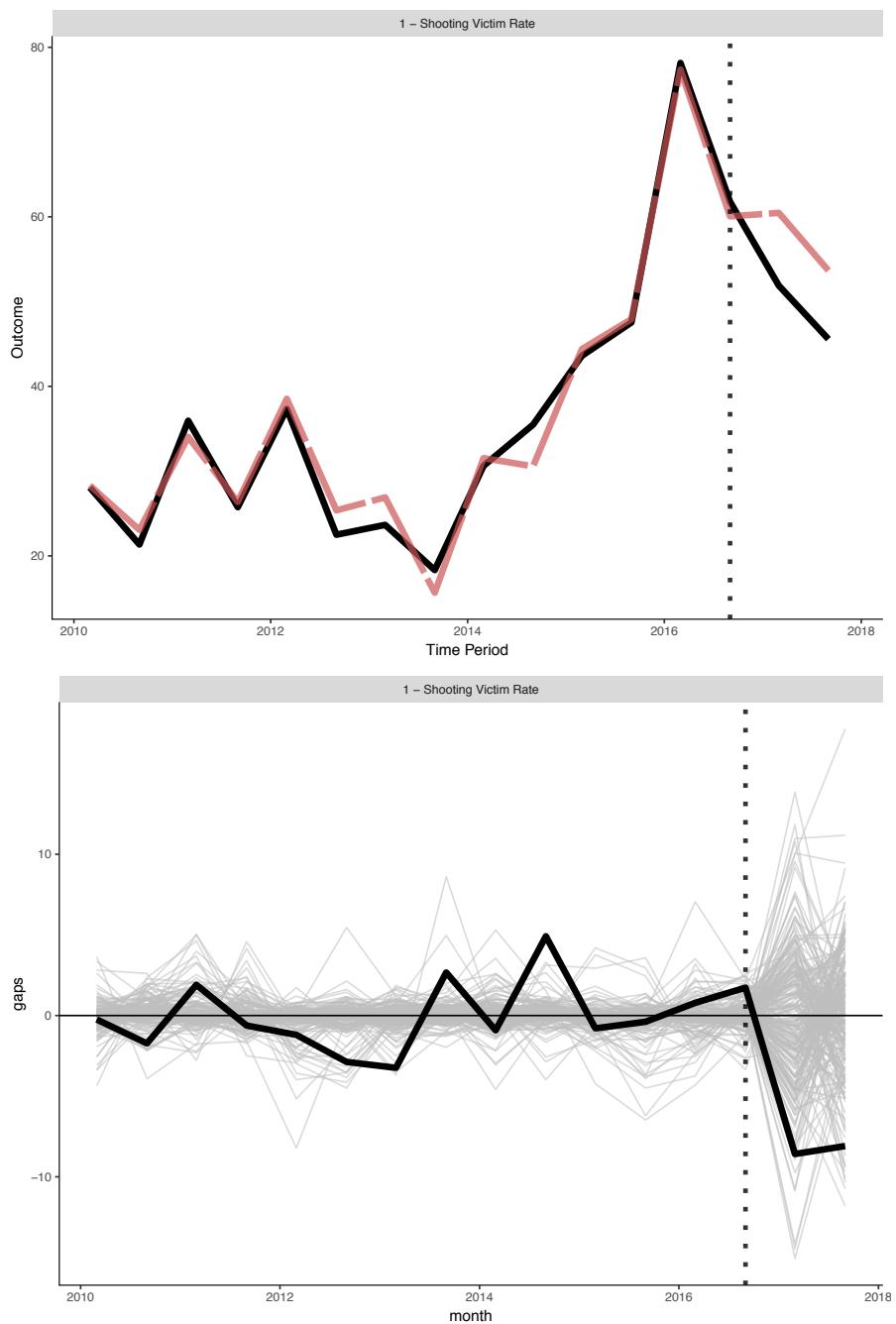
**Notes:** Synthetic control for the 7<sup>th</sup> district, using modified DI estimator and beats as donors, in red. Placebo test results plot differences between the actual and synthetic control unit, for the 7<sup>th</sup> district (bold line) and the resampled donor districts (light gray lines).

**Figure 18.** Synthetic control results for the 7<sup>th</sup> district: primary and secondary outcomes



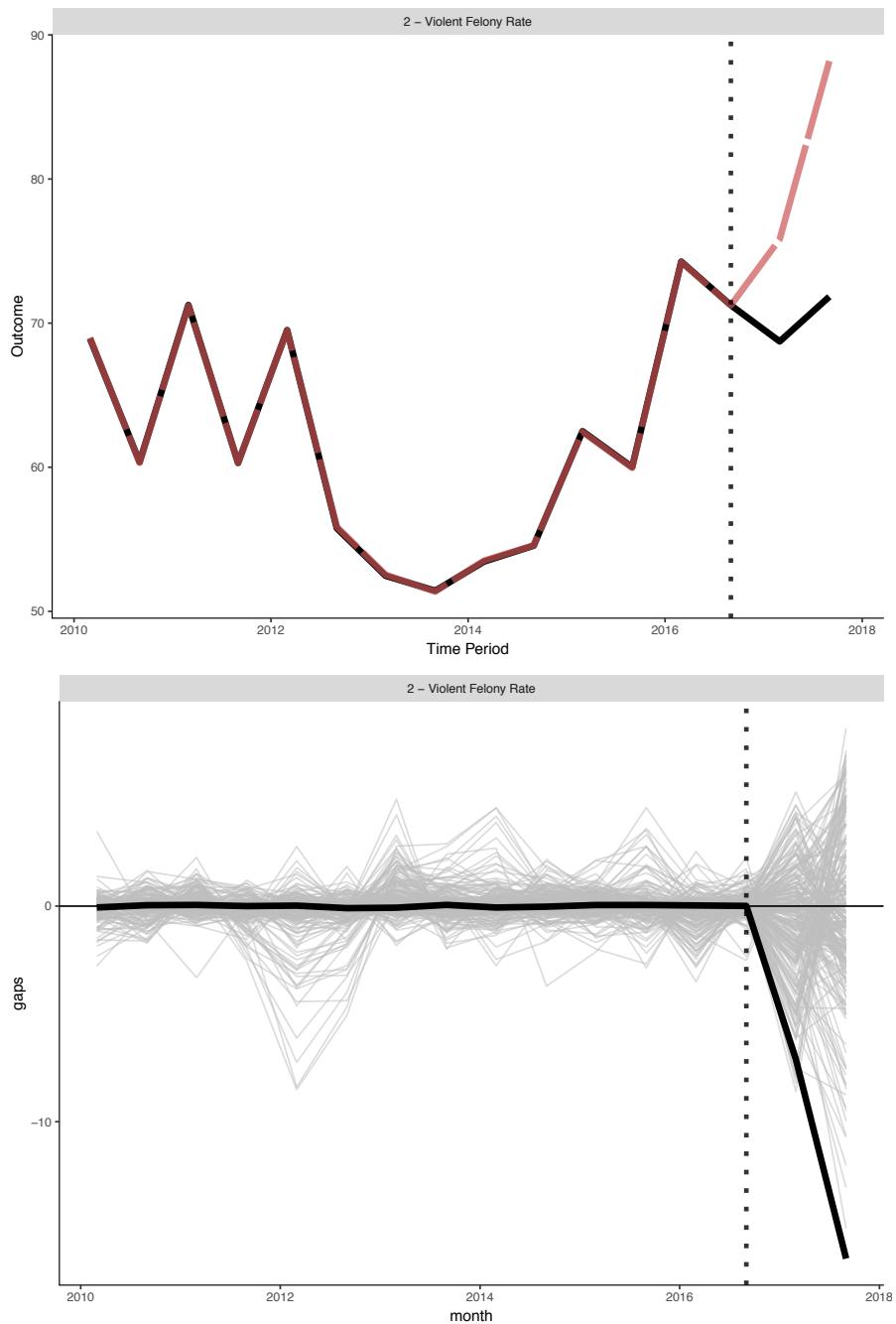
**Notes:** Synthetic control for the 7<sup>th</sup> district, using modified DI estimator fit to five outcomes (shooting victims per capita, Part I violent felony incidents per capita, gun arrests per capita, warrant arrests per capita, overall arrests per capita, traffic stops per capita) and beats as donors, in red.

**Figure 19.** Synthetic control results for the 11<sup>th</sup> district: shooting victims



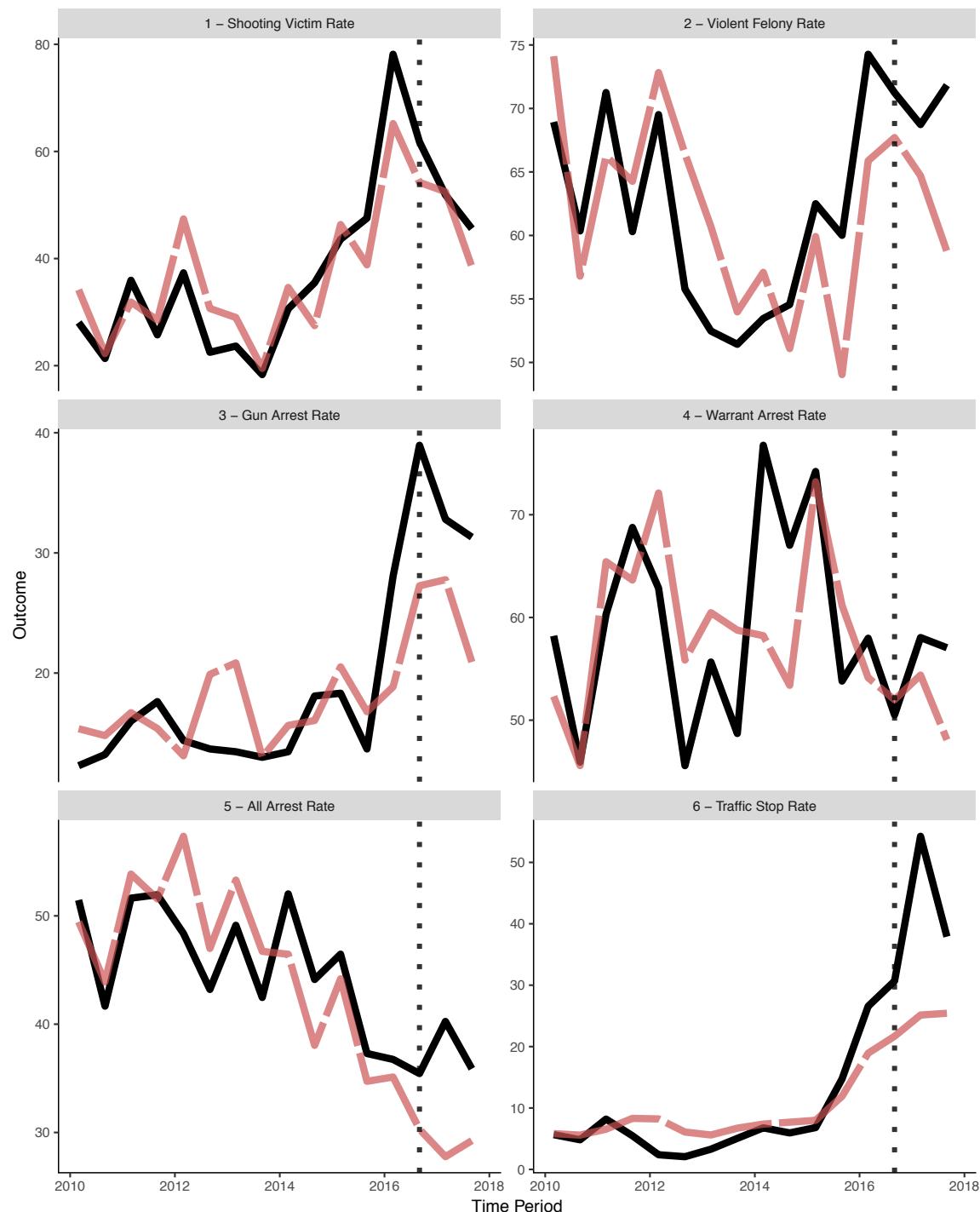
**Notes:** Synthetic control for the 11<sup>th</sup> district, using modified DI estimator and beats as donors, in red. Placebo test results plot differences between the actual and synthetic control unit, for the 11<sup>th</sup> district (bold line) and the resampled donor districts (light gray lines).

**Figure 20.** Synthetic control results for the 11<sup>th</sup> district: Part I violent felonies



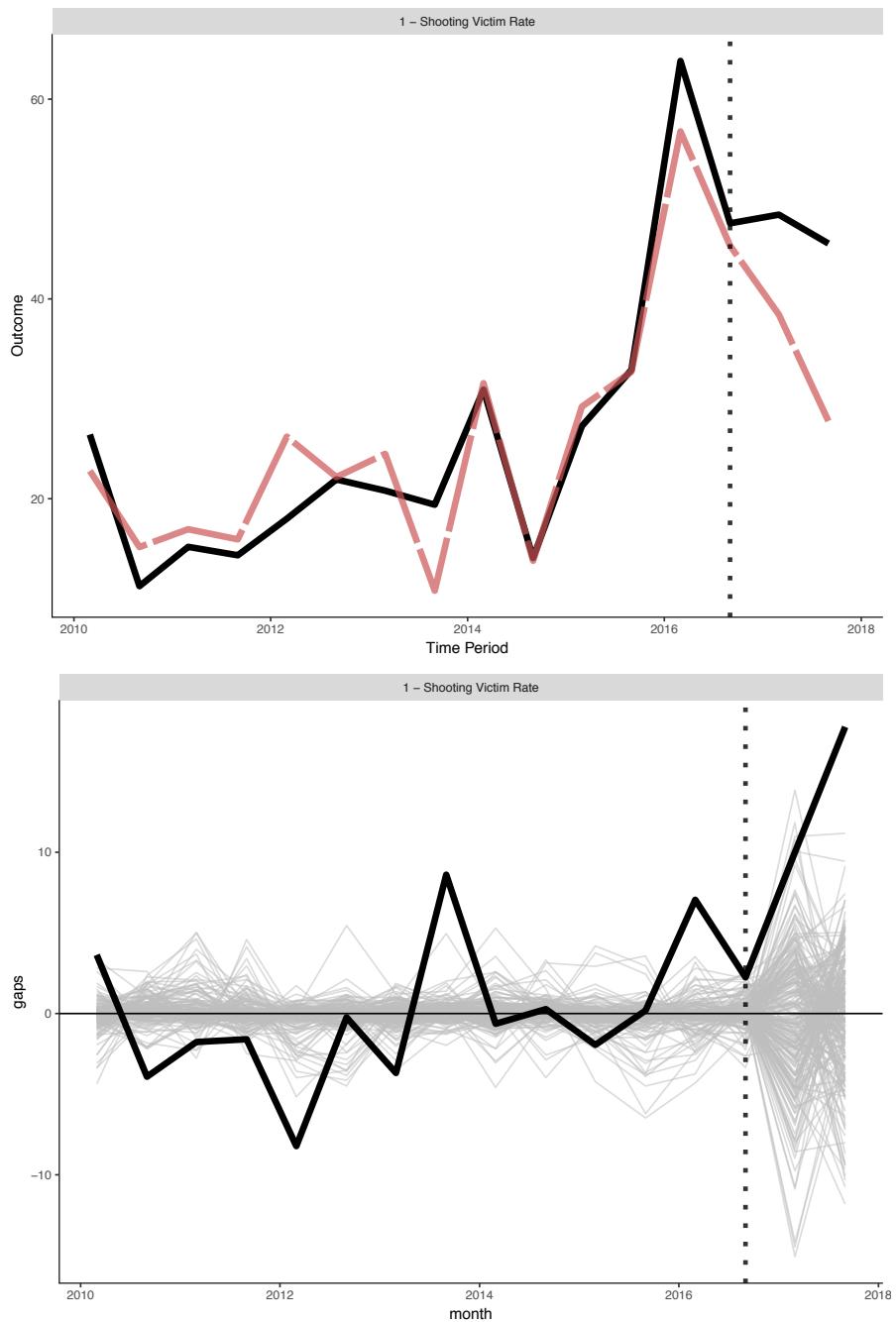
**Notes:** Synthetic control for the 11<sup>th</sup> district, using modified DI estimator and beats as donors, in red. Placebo test results plot differences between the actual and synthetic control unit, for the 11<sup>th</sup> district (bold line) and the resampled donor districts (light gray lines).

**Figure 21.** Synthetic control results for the 11<sup>th</sup> district: primary and secondary outcomes



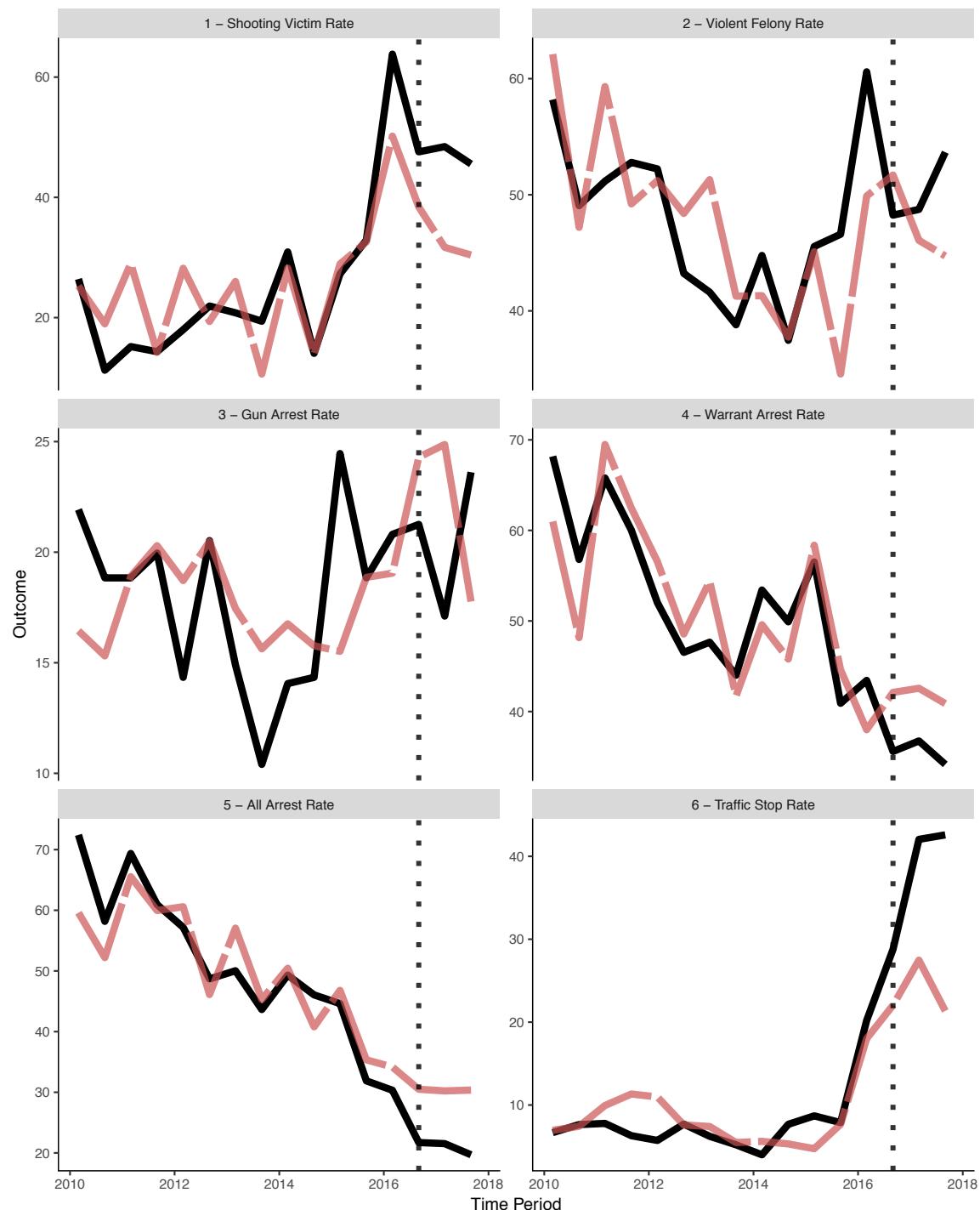
**Notes:** Synthetic control for the 11<sup>th</sup> district, using modified DI estimator fit to five outcomes (shooting victims per capita, Part I violent felony incidents per capita, gun arrests per capita, warrant arrests per capita, overall arrests per capita, traffic stops per capita) and beats as donors, in red.

**Figure 22.** Synthetic control results for the 15<sup>th</sup> district: shooting victims



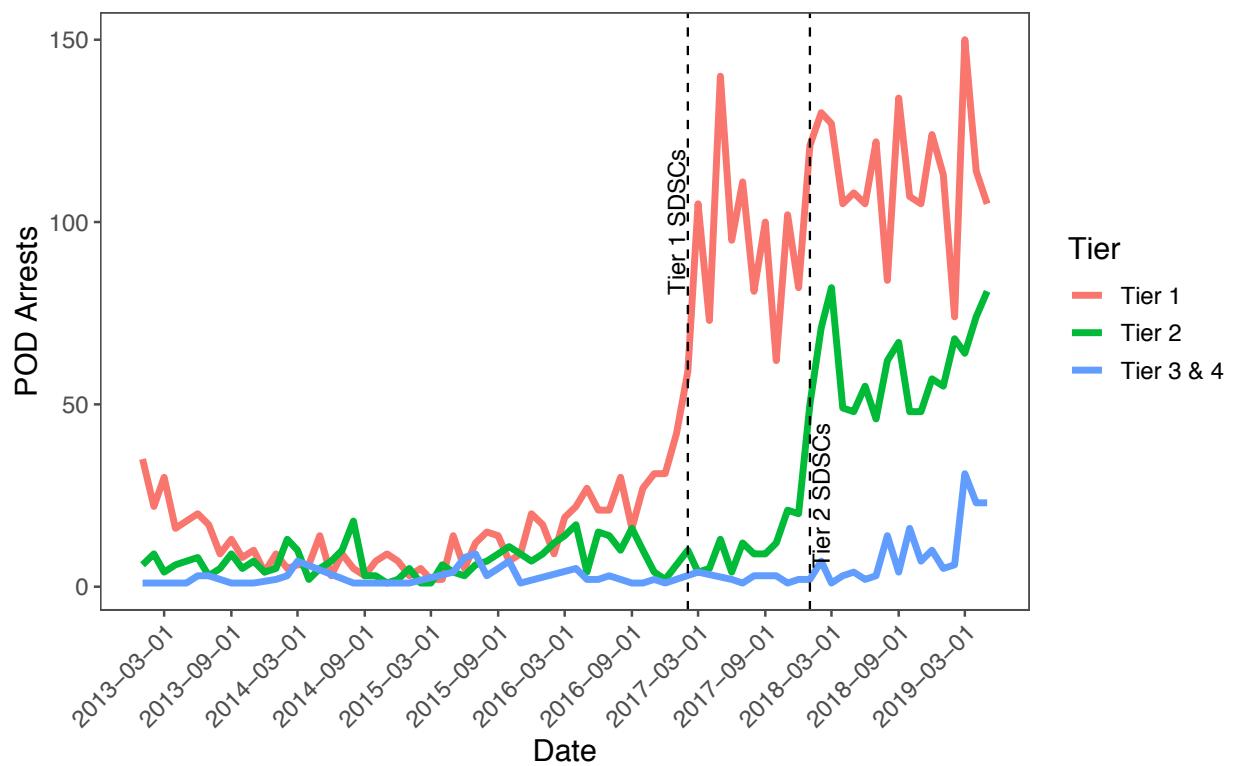
**Notes:** Synthetic control for the 15<sup>th</sup> district, using modified DI estimator and beats as donors, in red. Placebo test results plot differences between the actual and synthetic control unit, for the 15<sup>th</sup> district (bold line) and the resampled donor districts (light gray lines).

**Figure 23.** Synthetic control results for the 15<sup>th</sup> district: primary and secondary outcomes



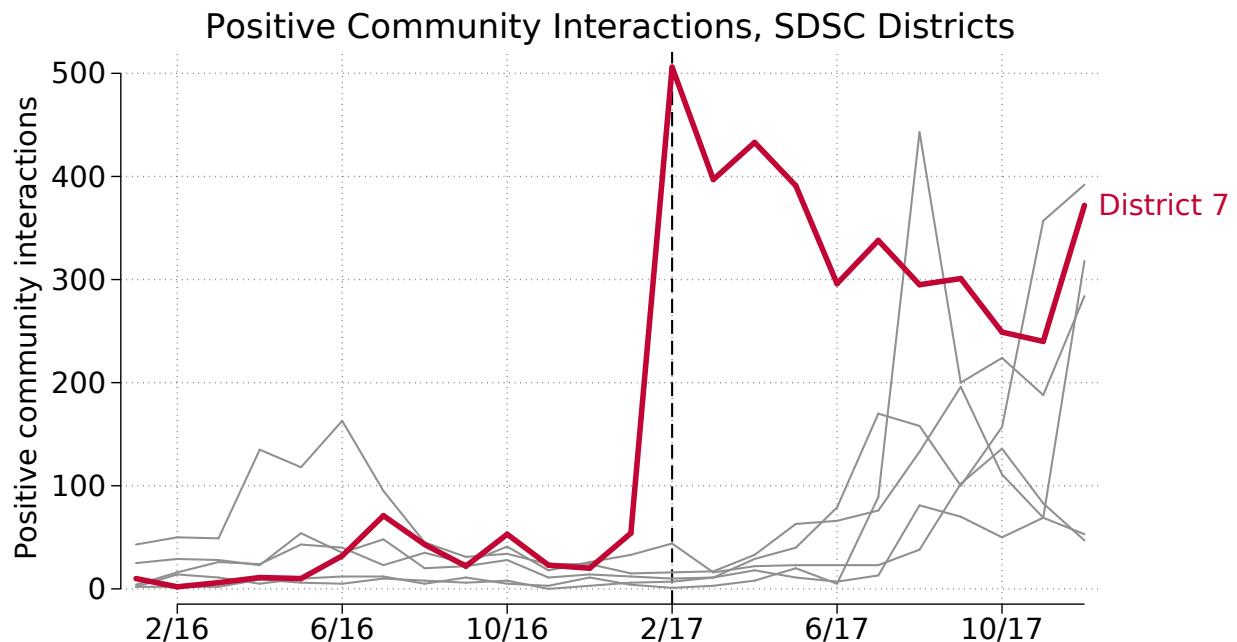
**Notes:** Synthetic control for the 15<sup>th</sup> district, using modified DI estimator fit to five outcomes (shooting victims per capita, Part I violent felony incidents per capita, gun arrests per capita, warrant arrests per capita, overall arrests per capita, traffic stops per capita) and beats as donors, in red.

**Figure 24.** POD camera-initiated arrests



**Notes:** Arrests initiated by officers monitoring POD cameras, by district Tier. SDSCs were first introduced in early 2017 in the six Tier 1 districts, followed by the seven Tier 2 districts in early 2018.

**Figure 25.** Positive community interactions: 7<sup>th</sup> district vs. other SDSC districts



**Notes:** Monthly volume of positive community interaction (PCI) calls by officers in Tier 1 districts.

**Table 1.** Synthetic control results for the 7<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	-26%	0.012	0.146	0.184
Part I Violent Felonies	-10%	0.028	0.312	0.323
<i>Synthetic control estimated jointly</i>				
Shooting Victims	-21%	0.008	0.316	0.294
Part I Violent Felonies	-13%	0.008	0.316	0.294
Gun Arrests	28%	0.089	1.000	1.000
Warrant Arrests	36%	0.040	1.000	0.809
Overall Arrests	20%	0.142	1.000	1.000
Traffic Stops	77%	0.000	0.000	0.000

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

**Table 2.** Synthetic control results for the 11<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	-14%	0.049	0.437	0.369
Part I Violent Felonies	-13%	0.000	0.000	0.000
<i>Synthetic control estimated jointly</i>				
Shooting Victims	5%	0.223	1.000	1.000
Part I Violent Felonies	12%	0.028	0.992	0.644
Gun Arrests	28%	0.239	1.000	1.000
Warrant Arrests	11%	0.287	1.000	1.000
Overall Arrests	37%	0.020	0.729	0.526
Traffic Stops	91%	0.004	0.162	0.245

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

**Table 3.** Synthetic control results for the 6<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	2%	0.543	1.000	1.000
Part I Violent Felonies	2%	0.243	1.000	1.000
<i>Synthetic control estimated jointly</i>				
Shooting Victims	13%	0.401	1.000	1.000
Part I Violent Felonies	-1%	0.818	1.000	1.000
Gun Arrests	8%	0.275	1.000	1.000
Warrant Arrests	0%	0.506	1.000	1.000
Overall Arrests	12%	0.247	1.000	1.000
Traffic Stops	34%	0.093	1.000	1.000

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

**Table 4.** Synthetic control results for the 9<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	-9%	0.348	1.000	1.000
Part I Violent Felonies	-24%	0.040	0.405	0.369
<i>Synthetic control estimated jointly</i>				
Shooting Victims	-16%	0.470	1.000	1.000
Part I Violent Felonies	-7%	0.777	1.000	1.000
Gun Arrests	-46%	0.344	1.000	1.000
Warrant Arrests	11%	0.672	1.000	1.000
Overall Arrests	22%	0.389	1.000	1.000
Traffic Stops	-41%	0.121	1.000	1.000

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

**Table 5.** Synthetic control results for the 10<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	39%	0.081	0.583	0.461
Part I Violent Felonies	-5%	0.591	1.000	1.000
<i>Synthetic control estimated jointly</i>				
Shooting Victims	27%	0.101	1.000	1.000
Part I Violent Felonies	-1%	0.988	1.000	1.000
Gun Arrests	60%	0.134	1.000	1.000
Warrant Arrests	-5%	0.789	1.000	1.000
Overall Arrests	9%	0.377	1.000	1.000
Traffic Stops	-28%	0.324	1.000	1.000

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

**Table 6.** Synthetic control results for the 15<sup>th</sup> district

Outcome	Effect Estimate (1)	<i>p</i> -value (2)	<i>q</i> -value	
			FWER Control (3)	FDR Control (4)
<i>Synthetic controls estimated separately</i>				
Shooting Victims	37%	0.000	0.000	0.000
Part I Violent Felonies	-1%	0.073	0.583	0.461
<i>Synthetic control estimated jointly</i>				
Shooting Victims	52%	0.000	0.000	0.000
Part I Violent Felonies	11%	0.121	1.000	1.000
Gun Arrests	-13%	0.259	1.000	1.000
Warrant Arrests	-15%	0.283	1.000	1.000
Overall Arrests	-31%	0.045	1.000	0.809
Traffic Stops	67%	0.016	0.599	0.490

**Notes:** Synthetic control treatment effects using modified DI estimator and beats as donors. FWER control *q*-values calculated using the Holm-Bonferroni correction. FDR control *q*-values calculated using the Benjamini-Yekutieli correction.

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