

Exercise 2

Maximum likelihood estimation.

Gaussian distribution expressed in terms of the parameters μ and Σ .

μ : nx1 vector

Σ : nxn symmetric matrix

We have the following form for the density function:

$$p(x|\mu, \Sigma) = \frac{1}{(2\pi)^{\frac{n}{2}}(\Sigma)^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu) \right\}.$$

Where x is a vector in \mathcal{R}^n .

In our problem we want to estimate the parameter Σ . Let us suppose that D is our dataset.

$$l(\mu, \Sigma|D) = -\frac{N}{2} \log|\Sigma| - \frac{1}{2} \sum_{i=1}^N (x_i - \mu)^T \Sigma^{-1} (x_i - \mu).$$

Taking the derivative with respect to μ is straightforward:

$$\frac{\partial l}{\partial \mu} = \sum_{i=1}^N (x_i - \mu)^T \Sigma^{-1}.$$

and setting to zero we obtain a pleasant result:

$$\hat{\mu} = \frac{1}{N} \sum_{i=1}^N x_i.$$

We use this form in order to calculate the mean of the pdf.

We now turn to the problem of computing the maximum likelihood estimate of the covariance matrix. Letting $l(\Sigma|D)$ denote the terms in the likelihood that are a function of Σ , we obtain:

$$l(\Sigma|D) = -\frac{N}{2} \log|\Sigma| - \frac{1}{2} \sum_n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu).$$

We need to take the derivative with respect to Σ and set to zero. Rewriting the log-likelihood using "track trace"

$$l(\Sigma|D) = -\frac{N}{2} \log|\Sigma| - \frac{1}{2} \sum_n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

$$\begin{aligned}
&= \frac{N}{2} \log |\Sigma^{-1}| - \frac{1}{2} \sum_n \text{tr}[(x_i - \mu)^T \Sigma^{-1} (x_i - \mu)] \\
&= \frac{N}{2} \log |\Sigma^{-1}| - \frac{1}{2} \sum_n \text{tr}[(x_i - \mu)^T (x_i - \mu) \Sigma^{-1}]
\end{aligned}$$

We now take the derivative with respect to the matrix Σ^{-1} and using 1) $\frac{\partial}{\partial A} \log |A| = A^{-T}$, 2)

$$\frac{\partial}{\partial A} \text{Tr}[AB] = \frac{\partial}{\partial A} \text{Tr}[BA] = B^T.$$

$$\frac{\partial l}{\partial \Sigma^{-1}} = \frac{N}{2} \Sigma - \frac{1}{2} \sum_n (x_i - \mu)^T (x_i - \mu).$$

Finally, setting to zero yields the maximum likelihood estimator:

$$\hat{\Sigma}_{ML} = \frac{1}{N} \sum_n (x_n - \hat{\mu}_{ML})^T (x_n - \hat{\mu}_{ML}).$$

In our case, the covariance matrices are diagonal, with all diagonal elements equal.

$$\Sigma = \begin{bmatrix} \sigma^2 & & \\ & \ddots & \\ & & \sigma^2 \end{bmatrix}$$

$$\begin{aligned}
l(\Sigma|D) &= -\frac{N}{2} \log |\Sigma| - \frac{1}{2} \sum_n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) + \text{const} \\
&= -\frac{N}{2} \log \sigma^{2k} - \frac{1}{2} \sum_n (x_i - \mu)^T \Sigma^{-1} (x_i - \mu) + \text{const} \\
&= -\frac{N}{2} \log \sigma^{2k} - \frac{1}{2} \sum_n (x_i - \mu)^T I \sigma^{-1} (x_i - \mu) + \text{const} \\
&= -Nk \log \sigma - \frac{\sigma^{-2}}{2} \sum_n (x_i - \mu)^T (x_i - \mu) + \text{const} \\
\frac{dl}{d\sigma} &= -Nk - \frac{1}{\sigma^2} \sum_n (x_i - \mu)^T (x_i - \mu) = 0 \\
\sigma_{ML}^2 &= \frac{1}{Nk} \sum_n (\vec{x}_i - \vec{\mu})^T (\vec{x}_i - \vec{\mu})
\end{aligned}$$

