

Assignment ①

① Bayesian networks must be acyclic.

Suppose we have a graph $G = (V, E)$ and discrete random variables X_1, \dots, X_n and define

$$f(x_1, \dots, x_n) = \prod_{v \in V} f_v(x_v | \text{pa}_v)$$

that is, where pa_v refers to the parents of variable x_v in G , and $f_v(x_v | \text{pa}_v)$ specifies a probability distribution over x_v for every assignment

of x_v 's parents. i.e., $\sum_{x_v \in \text{Vals}(x_v)} f_v(x_v | \text{pa}_v) \leq 1$.

$$\sum_{x_v \in \text{Vals}(x_v)} f_v(x_v | \text{pa}_v) = 1. \text{ Recall}$$

that this is precisely the definition of the joint probability distribution associated with the bayesian network G , where the f_v are the conditional probability distributions.

We are to show that if G has a directed cycle,

f may no longer define a valid probability distribution. Give an example of a cyclic

graph G and distributions f_v such that

$$\sum_{x_1, \dots, x_n} f(x_1, \dots, x_n) \neq 1. \text{ (A valid probability distribution must be non-negative)}$$

(and sum to one) This is why bayesian networks must be defined on acyclic graphs.

→ Solution :-

If there's a cycle in a bayesian network, a variable can be its own ancestor, meaning a node can influence itself through a series of directed edges. This cyclicity can lead to a scenario where a variable's probability depends on itself, leading to ambiguity.

To illustrate, consider two nodes X_1, X_2 such that they form a cycle:-

$$X_1 \rightarrow X_2$$

$$X_2 \rightarrow X_1$$

Given the cycle, we have:-

① f_{X_1} - The probability distribution of X_1 (Since X_1 is also a child of X_2 due to the cycle).

② $f_{X_2|X_1}$ - The conditional distribution of X_2 given X_1 .

③ $f_{X_1|X_2}$ - The conditional distribution of X_1 given X_2 .

Assume, $f_{x_1} = 0.5$ (probability of x_1 being true)

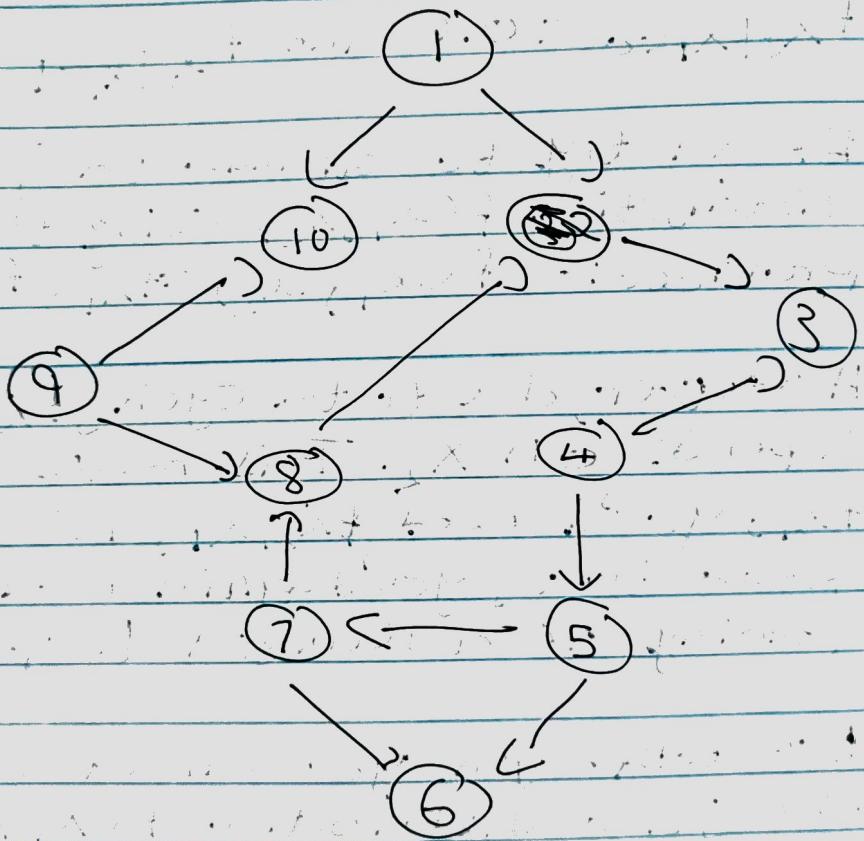
$$f_{x_2|x_1=0} = 0.6 \text{ and } f_{x_2|x_1=1} = 0.4$$

If we try to calculate the joint distribution $f(x_1, x_2)$ using the formula provided: - $f(x_1, x_2) = f_{x_1} \times f_{x_2|x_1}$

However, due to the cycle, f_{x_1} also depends on x_2 . This means to know f_{x_1} , you'd need $f_{x_2|x_1}$, and also know that, you'd again need $f_{x_1|x_2}$ leading to a recursive loop.

Due to the cycle you cannot define the joint distribution $f(x_1, x_2)$ in a way that all possible combinations of x_1 and x_2 sum to 1. This inconsistency arises from the cyclicity in the Bayesian Network.

Q(2) D-Separation. Consider the Bayesian network shown in the below figure.



Q(2) For what pair (i, j) does the statement $X_i \perp X_j$ hold? List all such pairs. And Explain all pairs for $X_i = 4$ or $X_j = 4$.

For d-separation, we have three rules to determine whether a path between two nodes is blocked.

① If we have a chain structure, A - B - C then B blocks the path unless B is observed (conditioned upon).

② For a fork structure, A ← B → C, B blocks both the paths unless B or one of its children, descendants is observed.

Based on graph,

→ 1 and 3 are connected by a chain

through 2 and 4.

→ 1 and 2 are connected by a direct link.

→ 2 and 3 are connected by a direct link

→ 3 and 4 are connected by a chain

→ 4 and 5 are connected by a chain through 7

→ 5 and 6 are connected by a direct link

→ 5 and 7 are connected by a direct link

→ 4 and 8 are connected by a direct link

→ 2 and 9 are connected by a direct link

→ 1 and 10 are connected by a direct link

Thus, the pairs that are not independent are (1,2), (1,3), (2,3), (3,4), (4,5), (5,6), (5,7)

$(4,8)$, $(2,9)$, and $(1,10)$ are FP 's

- (B) Suppose that we condition on (X_2, X_7) , shown shaded in the graph. What is the largest set A for which the statement $X_1 \perp\!\!\!\perp X_8 \mid \{X_2, X_7\}$ holds? Use the d-separation algorithm.

→ As we know, X_2 is a collider.

- ① Conditioning on a node in a chain or a fork blocks information flow.
- ② Conditioning on a node in a collider unblocks information flow, unless we also condition on a descendant of the collider.

Given conditioning on X_2 and X_7 .

X_1 and X_3 are independent since X_2 is conditioned upon.

X_1 and X_3 Any path from X_1 to X_3 goes through X_2 , which is conditioned. Thus

X_1 and X_3 are d-separated and independent given X_2 .

Ex. X_1 and X_4 : - the path from X_1 to X_4 goes through collider X_3 . Normally this would block the path but since X_3 is conditioned upon a child of X_2 ($X_3 \perp\!\!\!\perp X_2$) the path is active and X_1 and X_4 are dependent.

Ex. X_1 and X_5 : - The path goes through X_2 except X_4 and X_7 . Since X_7 is conditioned upon, this path is blocked, making X_1 and X_5 independent given the condition(s) of X_7 received.

Ex. X_1 and X_6 : - The path goes through X_5 and then X_7 (which is conditioned upon). Thus, X_1 and X_6 are d-separated and independent given X_7 .

Ex. X_1 and X_8 : - Any path from X_1 to X_8 goes through X_4 . Since X_4 and X_1 are dependent (as established earlier), X_1 and X_8 are also dependent.

Ex. X_1 and X_9 : - The direct path goes through X_{12} which is conditioned upon. Thus

x_1 and x_4 are d -separated and independent given X_2

x_1 and x_{10} , the nodes have a direct connection, making them dependent.

\Rightarrow Given the conditioning on X_2 are x_7 , the largest set A for which $x_1 \perp\!\!\!\perp A | X_2, x_7$. nodes is $A = x_3, x_5, x_6, x_9$

Q3) Consider the following distribution over 3 binary variables x, y, z :

$$\Pr(x, y, z) = \begin{cases} 1/12 & x \oplus y \oplus z = 0 \\ 1/6 & x \oplus y \oplus z = 1 \end{cases}$$

where \oplus denotes the XOR function.

(make sure your calculation of XOR function is correct by verifying

$\sum_{x,y,z} \Pr(x, y, z) = 1$) Show that there is no non-empty directed acyclic graph G such that

$$I_d\text{-sep}(G) = I(\Pr). (I_d\text{-sep}(G)$$

refers to the ~~sep~~ set of conditional

independences induced by the directed graphical structure, as obtained by the d-separation method).

→ D-sep method

(1) Verifying XOR function

$$0 \oplus 0 \oplus 0 = 0$$

$$x=0, y=0, z=0 \rightarrow 0 \oplus 0 \oplus 0 = 0, P=1/12$$

$$x=0, y=0, z=1 \rightarrow 0 \oplus 0 \oplus 1 = 1, P=1/6$$

$$x=0, y=1, z=0 \rightarrow 0 \oplus 1 \oplus 0 = 1, P=1/6$$

$$x=0, y=1, z=1 \rightarrow 0 \oplus 1 \oplus 1 = 0, P=1/12$$

$$x=1, y=0, z=0 \rightarrow 1 \oplus 0 \oplus 0 = 1, P=1/6$$

$$x=1, y=0, z=1 \rightarrow 1 \oplus 0 \oplus 1 = 0, P=1/6$$

$$x=1, y=1, z=0 \rightarrow 1 \oplus 1 \oplus 0 = 0, P=1/12$$

$$x=1, y=1, z=1 \rightarrow 1 \oplus 1 \oplus 1 = 1, P=1/6$$

$\sum = 1$, So XOR is corr.

(2) Show there's no DAG G such that $d\text{-sep}(G) = I(P_G)$.

We have to show there's no DAG. The XOR function is symmetric; the inverse of X, Y, Z on the result is equal. Hence no variable can be shown to be conditionally independent of another.

given the third.

for any structure of the graph:

→ If its $X \rightarrow Z \leftarrow Y$, this suggests Z is influenced by both X and Y . But this doesn't capture the XOR relationship properly.

→ For any other possible arrangement, we'll have one of the variables as either a parent or a child, implying conditional independence in some scenarios. XOR doesn't allow for such independence.

Conclusion: No graph G can capture the conditional independencies presented by XOR function.

Q4

Markov Property: Let $G = (V, E)$ be an undirected graph and let $X = \{X_1, \dots, X_n\}$ be a collection of random variables defined on its nodes. Recall that a

probability distribution over X is globally markov with respect to G if for any disjoint subsets A, B, C such that B separates A from C in the graph, then

the statement $x_A \perp\!\!\! \perp x_C \mid x_B$ is satisfied.

Now Consider two other notions of markovianity. A distribution is

pairwise markov wrt. to G if for any pair of nodes α, β not directly linked by an edge $e \in E$, the corresponding variables x_α and x_β are conditionally independent given all the remaining variables.

$$\forall \alpha, \beta \in V : (\alpha, \beta) \notin E, x_\alpha \perp\!\!\! \perp x_\beta \mid x_{V \setminus \{\alpha, \beta\}}$$

A distribution is locally markov wrt. G if for any α , a variable x_α when conditioned on all of its neighbours, is independent of the remaining variables.

$$\forall \alpha : x_\alpha \perp\!\!\! \perp x_{V \setminus (N(\alpha) \cup \{\alpha\})} \mid x_{N(\alpha)}$$

- (a) If a distribution is global markov, then it is locally markov. :- True
→ By definition, when a distribution is globally markov, any node x_α is independent of all nodes outside its

its neighborhood.

- (b) If a distribution is locally markov, then it is globally markov :- False
→ Being locally markov only ensures independence when considering immediate neighbours. It does not guarantee the more general condition of independence required by the global markov definition

- (c) If a distribution is locally markov then it is pairwise markov :- True
→ When a distribution is locally markov each variable is conditionally independent of all other variables given its neighbours. This implies that any pair of non-adjacent nodes are conditionally independent given all other nodes, which matches the pairwise markov.

- (d) If a distribution is pairwise markov wrt G, then it is locally markov :- true.
→ By the definition of pairwise markov, any two non-adjacent variables are conditionally independent given all other

Variables. This naturally includes the neighbors of each variable, so the distribution is also locally markov.

- (e) If a distribution is positive, then all three markov properties are equivalent (true).
- For positive distribution (where no conditional probabilities are zero), the three definitions are equivalent. This is a well-known property in the realm of graphical models. The non-zero condition allows the three properties to be interchangeable.

- (f) Markov blankets. Let (G, P) be a bayesian network. That is, $G = (V, E)$ is a DAG and P a distribution that factorizes over G . For a node X , defines its Markov Blanket $MB_G(X)$ as the set of X 's parents, its children and the parents of its children. (i.e $MB_G(X) = P_a(X) \cup Ch(X) \cup (P_a(Ch(X)) \setminus \{X\})$).

(Q) Show that for any variable X , if we define $w = \sqrt{\sum_{y \in \mathcal{Y}} p(y) MB_G(x)y}$, then X and w are d-separated given $MB_G(x)$.

→ By definition, the Markov Blanket, $MB_G(x)$, of a node X includes :-

- (1) The parents of x .
- (2) The children of x .
- (3) The parents of the children of x (other than x itself).

Given $MB_G(x)$:

- (1) X is independent of its parents because it is conditioned on them.
- (2) X is independent of the parents of its children, since it is separated from them by its children (by which we condition).
- (3) X is independent of all nodes not in its markov blanket since none of those nodes have direct or indirect paths to X .

Hence, X and w are d-separated given $MB_G(x)$.

(b) Shows that removing u only needs from the markov blanket $MB_6(x)$, violates the property you proved in the previous item. That is, for $U \in MB_6(x)$ shows that x and $w \cup \{u\}$ are not d-separated given $MB_6(x) \setminus \{u\}$.

→ ① If U is a parent of x : - By removing U , one opens up a backdoor path from x to w . Unless they are directly connected

② If U is a child of x : - Removing U opens up a path from x to the parents of U (other than x itself) through the child U .

③ If U is a parent of a child Y of x : - with its conditioning on U , there's an open path from x to w through the child Y .

In all these cases, x and $w \cup \{u\}$ are not d-separated given $MB_6(x) \setminus \{u\}$.

(C) Describe a construction that given G outputs an undirected graph $\tilde{G} = (\tilde{V}, \tilde{E})$ such that $I(\tilde{G}) \subseteq I(G)$ and removing any edge from \tilde{G} will violate this i.e. for any $e \in \tilde{E}$ define $G_{\{e\}} = (\tilde{V}, \tilde{E} \setminus \{e\})$, it holds that $I(G_{\{e\}}) \not\subseteq I(G)$.

→ Construction using a 1-1 P.D.P.

- (1) Start with G' being a copy of G without the direction on the edges
- (2) for every V-structure in G' (i.e. a structure where two nodes A and B both have directed edges into node C but there's no direct edge between A and B), add an undirected edge between A and B in G' .