

## Today's Agenda

- Multicollinearity Ex Discussion
- Spring Interpretation of Multicollin
- Finish up Modeling Probs
  - ↳ Non Normal Resids
  - ↳ Non linearity (vccap)
- More on Interactions

$$\hat{\beta} = VD^{-1}u^T y = V \left( \begin{array}{c} \boxed{\frac{1}{d_1}} \\ \vdots \\ \boxed{\frac{1}{d_p}} \end{array} \right)$$

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# Modeling Problems

## - Non Normality of Residuals

Let's say you fit a regression, plot the resids w/ a histogram & it looks not Normal (QQplot also).

Good

- $\hat{\beta}$  is still BLUE

Bad

- Normality was needed for inference
  - ↳ t-tests
  - ↳ F-tests
  - ↳ CIs
  - ↳ PIs

Good

↖ n is rly big  
If  $n \gg \gg 0$  then

a Central Limit Theorem type result  
gives us the fact that

$$\sqrt{n}(\hat{\beta} - \beta) \xrightarrow{D} N(0, \text{Var}(\hat{\beta}))$$

$$\hat{\beta} \sim N(\beta, \text{Var}(\hat{\beta}))$$

For this reason, if  $n$  is large this assumption  
is the "least important" to check.

How to check Non-Normality?

- Histogram of residuals
- QQ Plot

- Tests for Normality: Anderson Darling, Shapiro Wilks  
Ryan-Jenks, Kolmogorov-Smirnov

$H_0$ : residuals are Normal  $e \sim F$

$H_1$ : residuals exhibit significant deviation  
from Normality  $e \neq F$

Skew:

True parameter:  $E\left(\left(\frac{e_i}{\sigma}\right)^3\right)$

sample skewness  $\frac{1}{n} \sum_i \left(\frac{e_i}{\hat{\sigma}}\right)^3$

Kurtosis:

True parameter  $E\left(\left(\frac{e_i}{\sigma}\right)^4\right)$

sample kurtosis  $\frac{1}{n} \sum_i \left(\frac{e_i}{\hat{\sigma}}\right)^4$

Omnibus K-Squared Normality Test

$$K_{stat}^2 = \left[ Z_1(g_1) \right]^2 + \left[ Z_2(g_2) \right]^2$$

↑  
skewness  
var

↑  
kurtosis

$Z_1$  &  $Z_2$   
special  
transformations

$$K_{stat}^2 > \chi_2^2(1-\alpha) \Rightarrow \text{reject } H_0$$

# Jarque-Bera Test:

Combining skewness & kurtosis

$$JB_{stat} = \frac{n-p}{6} \left( \underset{\substack{\uparrow \\ \text{skew}}}{(g_1)^2} + \frac{1}{4} \underset{\substack{\uparrow \\ \text{kurtosis}}}{(g_2)^2} \right)$$

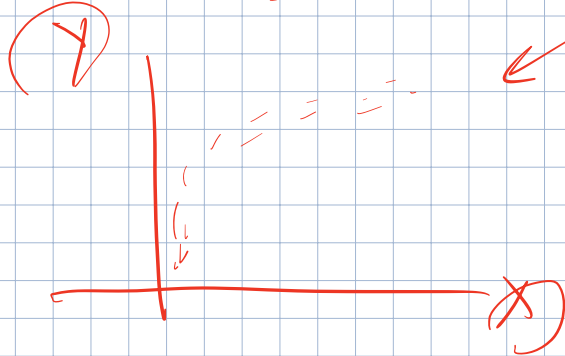
Under  $H_0$ :  $JB_{stat} \sim \chi^2_2$

Idea: If you have non Normal residuals

& you don't trust the asymptotic CI,

consider using a bootstrap CI.

False Linearity



If nonlinearity exists:

① transform  $Y$

(watch out  
→ affects  
heteroskedasticity)

② transform  $X$

③ transform  $Y$  &  $X$  (potentially log both)

$$\log Y = \beta_0 + \beta_1 \log X$$

$$e^{\log Y} = e^{\beta_0 + \beta_1 \log X}$$

$$Y = e^{\beta_0} e^{\beta_1 \log X}$$

$$= e^{\beta_0} e^{\log X^{\beta_1}}$$

$$= e^{\beta_0} X^{\beta_1}$$

$$= \hat{\beta}_0 X^{\beta_1}$$

$$\tilde{Y} = \begin{bmatrix} \log Y_1 \\ \vdots \\ \log Y_n \end{bmatrix}$$

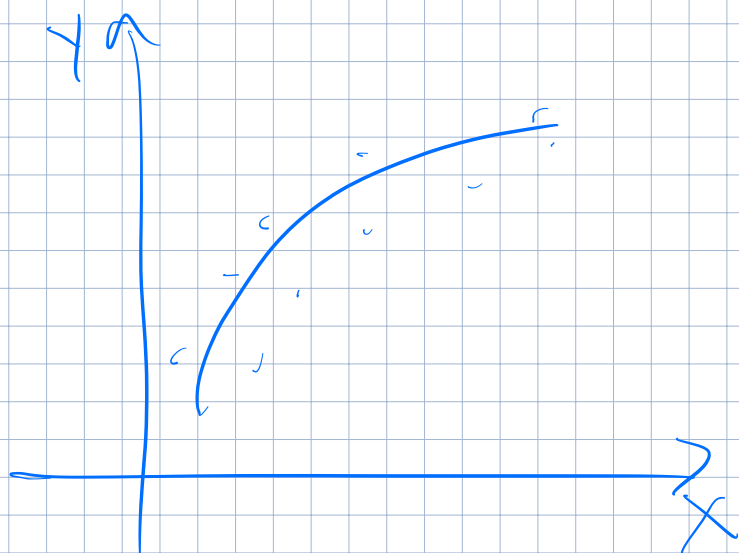
$$\tilde{X} = \begin{bmatrix} 1 & \log X_1 \\ \vdots & \vdots \\ 1 & \log X_n \end{bmatrix}$$

$$\beta = \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix}$$

$$\hat{\beta}_0 = e^{\beta_0}$$

Watch out for interpretations →

Sometimes hard after transformation



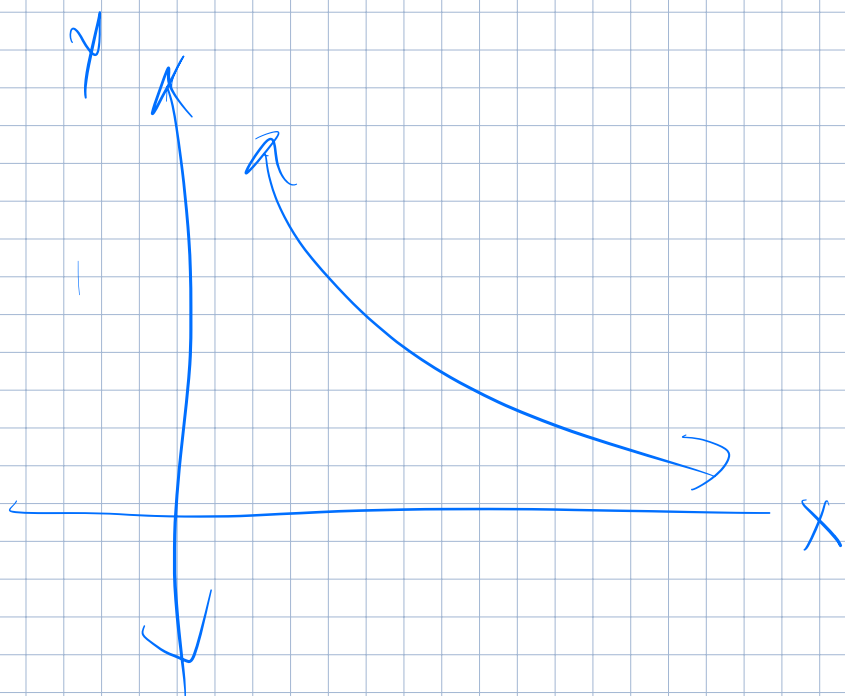
$$f(x) = \log(x)$$

$$f(x) = \sqrt{x}$$



$$f(x) = x^2$$

$$f(x) = e^x$$



$$f(x) = 1/x$$

$$f(x) = e^{-x} = \frac{1}{e^x}$$