

LHS:

$$\sum_i x_i y_i - n \bar{x} \bar{y}$$

RHS:

$$\sum_i (\hat{x_i - \bar{x}})(\hat{y_i - \bar{y}})$$

$$= \sum_i (x_i y_i - \bar{x} y_i - \underline{\underline{\bar{y} x_i}} + \bar{x} \bar{y})$$

$$= \sum_i x_i y_i - \underbrace{\sum_i \bar{x} y_i}_{\text{aside}} - \underbrace{\bar{y} \sum_i x_i}_{\text{aside}} + n \bar{x} \bar{y}$$

$$\text{aside} \left\{ \begin{array}{l} - (\bar{x} y_1 + \bar{x} y_2 + \dots + \bar{x} y_n) \\ - \bar{x} (\sum_i y_i) \\ - \bar{x} (n \bar{y}) \end{array} \right.$$

$$= \sum_i x_i y_i - n \bar{x} \bar{y} - \cancel{n \bar{x} \bar{y}} + n \bar{x} \bar{y}$$

General:

$$\sum_i (x_i - \bar{x})(y_i - \bar{y}) = \sum_i x_i y_i - n \bar{x} \bar{y}$$

for any $\{(x_i, y_i)\}_{i=1}^n$

Special case:

$$x_i = y_i \quad \forall i = 1, \dots, n$$

$$\sum_i (x_i - \bar{x})(x_i - \bar{x}) = \sum_i x_i^2 - n \bar{x}^2$$

$$\sum_i (x_i - \bar{x})^2 = \sum_i x_i^2 - n \bar{x}^2$$

$$\frac{\text{Cov}(X, Y)}{\text{Var}(X)} = \frac{\text{Cov}(X, Y)}{\text{Cov}(X, X)}$$

Inference & Diagnostics for SLR

First some useful facts:

Recall if $e_i = y_i - \hat{y}_i$ then:

① $\sum_i e_i^2$ is minimized.

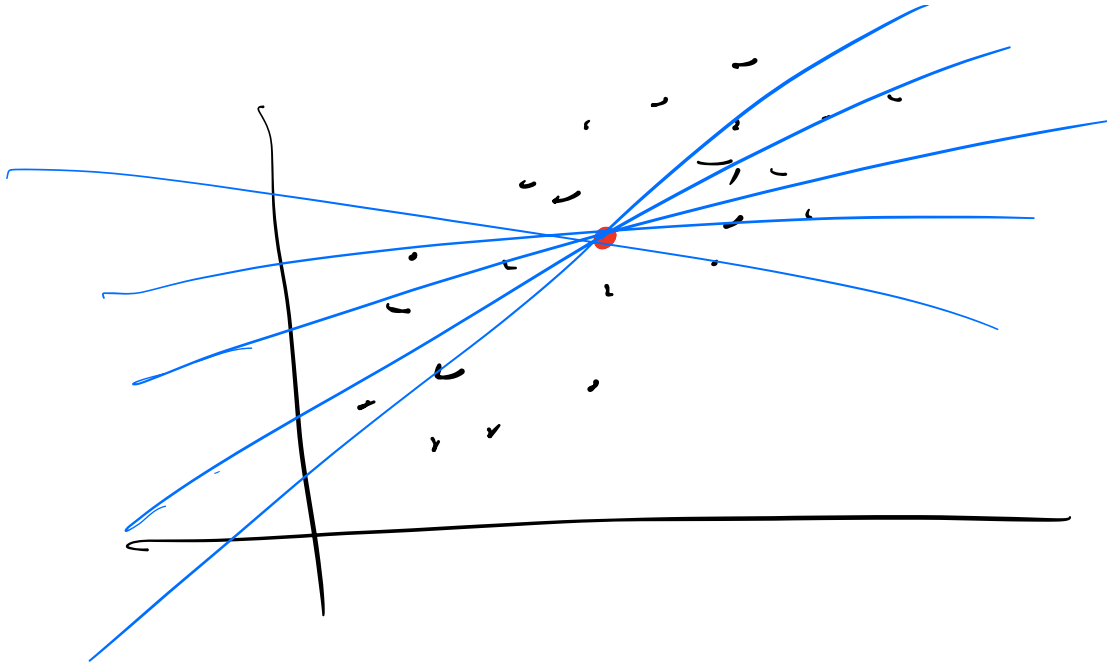
② $\sum_{i=1}^n y_i = \sum_{i=1}^n \hat{y}_i$

③ $\sum_{i=1}^n e_i = 0$

④ $\sum_{i=1}^n x_i e_i = 0$

⑤ $\sum_{i=1}^n \hat{y}_i e_i = 0$

⑥ The regression line always goes through (\bar{x}, \bar{y})



Pf of (2):

$$\sum_{i=1}^n \hat{y}_i = \sum_{i=1}^n (\hat{\beta}_0 + \hat{\beta}_1 x_i)$$

$$= \sum_{i=1}^n \hat{\beta}_0 + \sum_{i=1}^n \hat{\beta}_1 x_i$$

$$= n\hat{\beta}_0 + \hat{\beta}_1 \sum_{i=1}^n x_i$$

$$= n\hat{\beta}_0 + n\hat{\beta}_1 \bar{x}$$

$$= n(\hat{\beta}_0 + \hat{\beta}_1 \underline{\bar{x}})$$

$$= n(\bar{y} - \hat{\beta}_1 \bar{x} + \hat{\beta}_1 \bar{x})$$

$$= \sum_{i=1}^n y_i$$

We have these estimators

→ $\hat{\beta}_0$ & $\hat{\beta}_1$ which estimate

→ β_0 & β_1 .

If for a given ^{unknown} parameter θ we have an estimator $\hat{\theta}$ which satisfies:

$$E(\hat{\theta}) = \theta$$

then we say $\hat{\theta}$ is unbiased for θ .

Claim: $\hat{\beta}_0$ & $\hat{\beta}_1$ are unbiased
for β_0 & β_1 , resp.

Def:

$$\text{bias}(\hat{\theta}) = E(\hat{\theta}) - \theta$$

Pf for $\hat{\beta}_1$:

WTS:

$$E(\hat{\beta}_1) = \beta_1$$

$$E(\hat{\beta}_1) = E\left(\frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)$$

$$= E\left(\sum_{i=1}^n \left(\frac{(x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}\right)\right)$$

$$SSX = \sum_{i=1}^n (x_i - \bar{x})^2$$

$$= E \left(\sum_{i=1}^n \left(\frac{x_i - \bar{x}}{SSX} \right) (y_i - \bar{y}) \right)$$

$$= \sum_{i=1}^n \frac{(x_i - \bar{x})}{SSX} \underbrace{E(y_i - \bar{y})}_{\textcircled{A}}$$

$$\textcircled{A} E(y_i - \bar{y}) = E(y_i) - E(\bar{y})$$

$$= \beta_0 + \beta_1 x_i - E\left(\frac{1}{n} \sum_{i=1}^n y_i\right)$$

$$= \beta_0 + \beta_1 x_i - \frac{1}{n} \sum_{i=1}^n E(y_i)$$

$$= \beta_0 + \beta_1 x_i - \left(\frac{1}{n} \sum_{i=1}^n (\beta_0 + \beta_1 x_i) \right)$$

$$= \beta_0 + \beta_1 x_i - \left(\beta_0 + \beta_1 \left(\frac{1}{n} \sum_{i=1}^n x_i \right) \right)$$

$$= \cancel{\beta_0} + \beta_1 x_i - \cancel{\beta_0} - \beta_1 \bar{x}$$

$$= \beta_1 (x_i - \bar{x})$$

$$= \sum_{i=1}^n \frac{(x_i - \bar{x})}{SSX} (\beta_1 (x_i - \bar{x}))$$

$$= \sum_{i=1}^n \frac{\beta_1 (x_i - \bar{x})^2}{SSX}$$

$$= \frac{\beta_1}{SSX} (\sum_{i=1}^n (x_i - \bar{x})^2)$$

$$= \frac{\beta_1}{SSX} \cdot SSX = \beta_1$$

Ex: Show that $E(\hat{\beta}_0) = \beta_0$

What about variance?

The key is to notice that

$\hat{\beta}_0$ & $\hat{\beta}_1$ are both linear estimators (in y_i 's).

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x}) y_i}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

Note: In the original image, an orange circle highlights the numerator and denominator, and an arrow points from the label k_i to the term $(x_i - \bar{x})$ in the numerator.

Define $k_i = \frac{x_i - \bar{x}}{SS_X}$

$$\text{Var}(\hat{\beta}_1) = \text{Var}\left(\sum_{i=1}^n k_i y_i\right)$$

because $y_i \perp y_j \forall i \neq j$

$$= \sum_{i=1}^n \text{Var}(k_i y_i)$$

$$= \sum_{i=1}^n k_i^2 \text{Var}(y_i)$$

$$= \sum_{i=1}^n k_i^2 (\sigma^2)$$

$$= \sigma^2 \underbrace{\sum_{i=1}^n k_i^2}_{\rightarrow \text{algebra}} = \frac{\sigma^2}{SSX}$$

$$\boxed{= \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$

$$\sum_{i=1}^n k_i^2 = \sum_{i=1}^n \left(\frac{x_i - \bar{x}}{SSX} \right)^2$$

$$= \sum_{i=1}^n \left(\frac{(x_i - \bar{x})^2}{(SSX)^2} \right)$$

$$= \frac{1}{(SSX)^2} \boxed{\sum_i (x_i - \bar{x})^2}$$

$$= \frac{1}{(SSX)^2} \cdot SSX = \frac{1}{SSX}$$

HW:

$$\text{Var}(\hat{\beta}_0) = \sigma^2 \left(\frac{1}{n} + \frac{\bar{x}^2}{\text{SSX}} \right)$$

Strategy:

Express $\hat{\beta}_0 = \sum_{i=1}^n c_i y_i$

(Figure out what c_i is.)

Argue that

$$\text{Var}(\hat{\beta}_0) = \sigma^2 \sum_{i=1}^n c_i^2$$

Calculate $\sum_i c_i^2$.