

**Exercise Sheet 2 for
Design and Analysis of Algorithms
Autumn 2022**

Due 18 Oct 2022 at 23:59

Exercise 1

Suppose each CPU can execute at most one process at any time. Consider the following experiment, which proceeds in a sequence of rounds. For the first round, we have n processes, which are assigned independently and uniformly at random to n CPUs. For any $i \geq 1$, after round i , we first find all the processes p such that p has been assigned to a CPU C by itself, i.e., p is the *unique* process that has been assigned to C ; then we remove all such processes (as they would be executed) in round i . The remaining processes are retained for round $i + 1$, in which they are assigned independently and uniformly at random to the n CPUs.

- (a) If there are b processes at the start of a round, what is the expected number of processes at the start of the next round?
- (b) Suppose that every round the number of removed processes was exactly the expected number of removed processes. Show that all the processes will be removed in $O(\log \log n)$ rounds.

Hint: If x_j is the expected number of processes left after j rounds, show and use that $x_{j+1} \leq x_j^2/n$. You can use the fact that $1 - kx \leq (1 - x)^k$ for $0 < x < 1$ and $k \leq \frac{1}{x}$.

Solution:

- (a) $Pr(\text{the } i\text{th bin has exactly 1 ball}) = \binom{b}{1} \frac{1}{n} (1 - \frac{1}{n})^{(b-1)} = \frac{b}{n} (1 - \frac{1}{n})^{(b-1)}$
 $E(\text{numbers of 1 ball bins}) = \sum_{i=1}^n Pr(\text{the } i\text{th bin has exactly 1 ball}) = b(1 - \frac{1}{n})^{(b-1)}$
so, the expected number of processes at the start of the next round is $b - E(\text{numbers of 1 ball bins}) = b - b(1 - \frac{1}{n})^{(b-1)}$
- (b) If x_j is the expected number of processes left after j rounds, $x_{j+1} = x_j(1 - (1 - \frac{1}{n})^{x_j-1}) \leq \frac{x_j(x_j-1)}{n} \leq \frac{x_j^2}{n}$
 $x_1 = n(1 - (1 - \frac{1}{n})^{n-1}) = n(1 - \frac{1}{e})$, when $n \rightarrow \infty$. Set $k = (1 - \frac{1}{e})^{-1} > 1$ which is a small constant.
So, $x_1 = \frac{n}{k}$, $x_2 \leq x_1^2/n = n^2/(nk^2) = n/k^2$, $x_3 = n/k^4$, ... , $x(j+1) = n/k^{2^j} = 1$, which suppose on the $j+1$ ground, all the process are removed.
 $\therefore n = k^{2^j}$, $j = \log_2 \log_k n$, $\therefore j+1 = O(\log \log n)$.

Exercise 2

Suppose you are given a biased coin that has $Pr[\text{HEADS}] = p \geq a$, for some fixed a , without being given any other information about p .

- (a) Devise a procedure that outputs a value \tilde{p} such that you can guarantee that $Pr[|p - \tilde{p}| \geq \varepsilon p] \leq \delta$, for any choice of the constants $0 < a, \varepsilon, \delta < 1$. (The value \tilde{p} is often called the estimate of p .)
- (b) Let N be the number of times you need to flip the biased coin to obtain the estimate. What is the smallest value of N for which you can still give the above guarantee?

Hint: flip the coin a few times and consider the fraction of times seeing HEADS.

Solution:

Algorithm 1: COINSALG

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1: Initialization:  $X \leftarrow 0$ ;  $N \leftarrow$  a selected number;  
2: for  $i \leftarrow 1$  to  $N$  do do  
3:   Flip the biased coin;  
4:   if seeing HEADS then  
5:      $X \leftarrow X + 1$ ;  
6:   end if  
7: end for  
8:  $\tilde{p} \leftarrow \frac{X}{N}$ ;  
9: return  $\tilde{p}$ 
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- (a) As the Algorithm 1, I devise a procedure that return a value \tilde{p} that can guarantee that $\Pr[|p - \tilde{p}| \geq \varepsilon p] \leq \delta$, for any choice of the constants $0 < a, \varepsilon, \delta < 1$, and I will prove it as bellow.
- (b) Set $X_i = 1$ if and only if at the i th flip time, seeing the HEADS, otherwise $X_i = 0$. $X = \sum_{i=1}^N X_i$, $E(\text{number of all HEADS in } N \text{ times}) = EX = \sum_{i=1}^N EX_i = Np$, $\tilde{p} = \frac{X}{N}$
 $\Pr(|p - \tilde{p}| \geq \varepsilon p) = \Pr(|p - \frac{X}{N}| \geq \varepsilon p) = \Pr(|X - Np| \geq \varepsilon Np) = \Pr(|X - EX| \geq \varepsilon EX)$, $\because X_i \in [0, 1]$, we can use the Chernoff bound, $\therefore \Pr(|X - EX| \geq \varepsilon EX) \leq 2\exp(-\frac{\varepsilon^2 EX^2}{3}) \leq \delta$, $\therefore 2\exp(-\frac{Np\varepsilon^2}{3}) \leq \delta$, and we can solve the N .
 $N \geq \frac{-3\ln\delta}{2p\varepsilon^2}$, $\because p \geq a$, we get the smallest N value that can still give the above guarantee is $N = \frac{-3\ln\delta}{2a\varepsilon^2}$
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Exercise 3

Let X and Y be finite sets and let Y^X denote the set of all functions from X to Y . We will think of these functions as “hash” functions. A family $\mathcal{H} \subseteq Y^X$ is said to be strongly 2-universal if the following property holds, with $h \in \mathcal{H}$ picked uniformly at random:

$$\forall x, x' \in X \forall y, y' \in Y \left(x \neq x' \Rightarrow \Pr_h[h(x) = y \wedge h(x') = y'] = \frac{1}{|Y|^2} \right).$$

We are give a a stream \mathcal{S} of elements of X , and suppose that \mathcal{S} contains at most s distinct elements. Let $\mathcal{H} \subseteq Y^X$ be a strongly 2-universal hash family with $|Y| = cs^2$ for some constant $c > 0$. Suppose we use a random function $h \in \mathcal{H}$ to hash.

Prove that the probability of a collision (i.e., the event that two distinct elements of \mathcal{S} hash to the same location) is at most $1/(2c)$.

Solution:

$$\forall x, x' \in S \forall y, y' \in Y \left(x \neq x' \Rightarrow \Pr_h[h(x) = y \wedge h(x') = y'] = \frac{1}{|Y|^2} = \frac{1}{c^2 s^4} \right).$$

$$\Pr(\text{the } i, j\text{th distinct elements of } S \text{ hash to the } k\text{th location}) = \binom{s}{2} \binom{|Y|}{1} \frac{1}{|Y|^2} = \frac{s(s-1)cs^2}{2c^2 s^4} \leq \frac{1}{2c}.$$
