

Cluster Analysis

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- Required packages

```
knitr::opts_chunk$set(echo = TRUE)
#install.packages("dplyr", "ade4", "magrittr", "cluster", "factoextra", "cluster.datasets", "xtable", "kableExtra")
knitr::opts_chunk$set(echo = TRUE)
```

Definition of a distance

- A distance function or a metric on \mathbb{R}^m , $m \geq 1$, is a function $d : \mathbb{R}^m \times \mathbb{R}^m \rightarrow \mathbb{R}$.
- A distance function must satisfy some required properties or axioms.
- There are three main axioms.
- A1. $d(\mathbf{x}, \mathbf{y}) = 0 \iff \mathbf{x} = \mathbf{y}$ (identity of indiscernibles);
- A2. $d(\mathbf{x}, \mathbf{y}) = d(\mathbf{y}, \mathbf{x})$ (symmetry);
- A3. $d(\mathbf{x}, \mathbf{z}) \leq d(\mathbf{x}, \mathbf{y}) + d(\mathbf{y}, \mathbf{z})$ (triangle inequality), where $\mathbf{x} = (x_1, \dots, x_m)$, $\mathbf{y} = (y_1, \dots, y_m)$ and $\mathbf{z} = (z_1, \dots, z_m)$ are all vectors of \mathbb{R}^m .
- We should use the term *dissimilarity* rather than *distance* when not all the three axioms A1-A3 are valid.
- Most of the time, we shall use, with some abuse of vocabulary, the term distance.

Exercice 1

- Prove that the three axioms A1-A3 imply the non-negativity condition:

$$d(\mathbf{x}, \mathbf{y}) \geq 0.$$

Euclidean distance

- It is defined by:

$$d(\mathbf{x}, \mathbf{y}) = \sqrt{\sum_{j=1}^m (x_j - y_j)^2}.$$

- A1-A2 are obvious.
- The proof of A3 is provided below.

Exercice 2

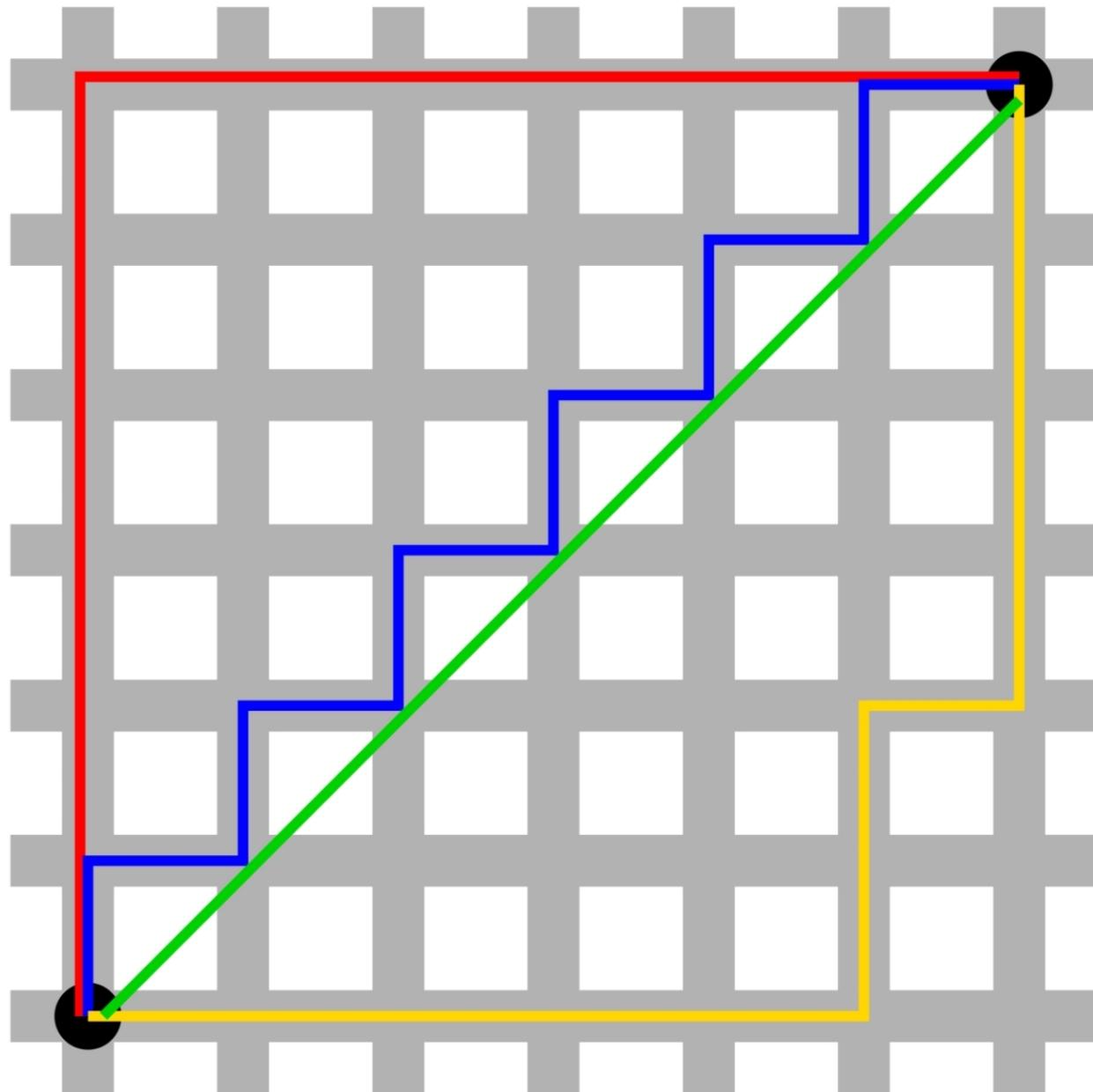
- Is the squared Euclidian distance a true distance?

Manhattan distance

- The Manhattan distance also called taxi-cab metric or city-block metric is defined by:

$$d(\mathbf{x}, \mathbf{y}) = \sum_{j=1}^m |x_j - y_j|.$$

- A1-A2 hold.
- A3 also holds using the fact that $|a + b| \leq |a| + |b|$ for any reals a, b .
- There exists also a weighted version of the Manhattan distance called the Canberra distance.



```
x = c(0, 0)
y = c(6,6)
dist(rbind(x, y), method = "euclidian")

##           x
## y 8.485281
dist(rbind(x, y), method = "euclidian", diag=T, upper=T)

##           x      y
## x 0.000000 8.485281
## y 8.485281 0.000000
6*sqrt(2)

## [1] 8.485281
```

```

dist(rbind(x, y), method = "manhattan")

##      x
## y 12
dist(rbind(x, y), method = "manhattan", diag=T, upper=T)

##      x  y
## x  0 12
## y 12  0

```

Canberra distance

- It is defined by:

$$d(\mathbf{x}, \mathbf{y}) = \sum_{j=1}^m \frac{|x_j - y_j|}{|x_j| + |y_j|}.$$

- Note that the term $|x_j - y_j|/(|x_j| + |y_j|)$ is not properly defined as: $x_j = y_j = 0$.
- By convention we set that term to be zero in that case.
- The Canberra distance is specially sensitive to small changes near zero.

```

x = c(0, 0)
y = c(6,6)
dist(rbind(x, y), method = "canberra")

##      x
## y 2
6/6+6/6

## [1] 2

```

Exercice 3

- Prove that the Canberra distance is a true distance, i.e. that it satisfies A1-A3.

Minkowski distance

- Both the Euclidian and the Manhattan distances are special cases of the Minkowski distance which is defined, for $p \geq 1$, by:

$$d(\mathbf{x}, \mathbf{y}) = \left[\sum_{j=1}^m |x_j - y_j|^p \right]^{1/p}.$$

- For $p = 1$, we get the Manhattan distance.
- For $p = 2$, we get the Euclidian distance.
- Let us also define:

$$\|\mathbf{x}\|_p \equiv \left[\sum_{j=1}^m |x_j|^p \right]^{1/p},$$

where $\|\cdot\|_p$ is known as the p -norm or Minkowski norm.

- Note that the Minkowski distance and norm are related by:

$$d(\mathbf{x}, \mathbf{y}) = \|\mathbf{x} - \mathbf{y}\|_p.$$

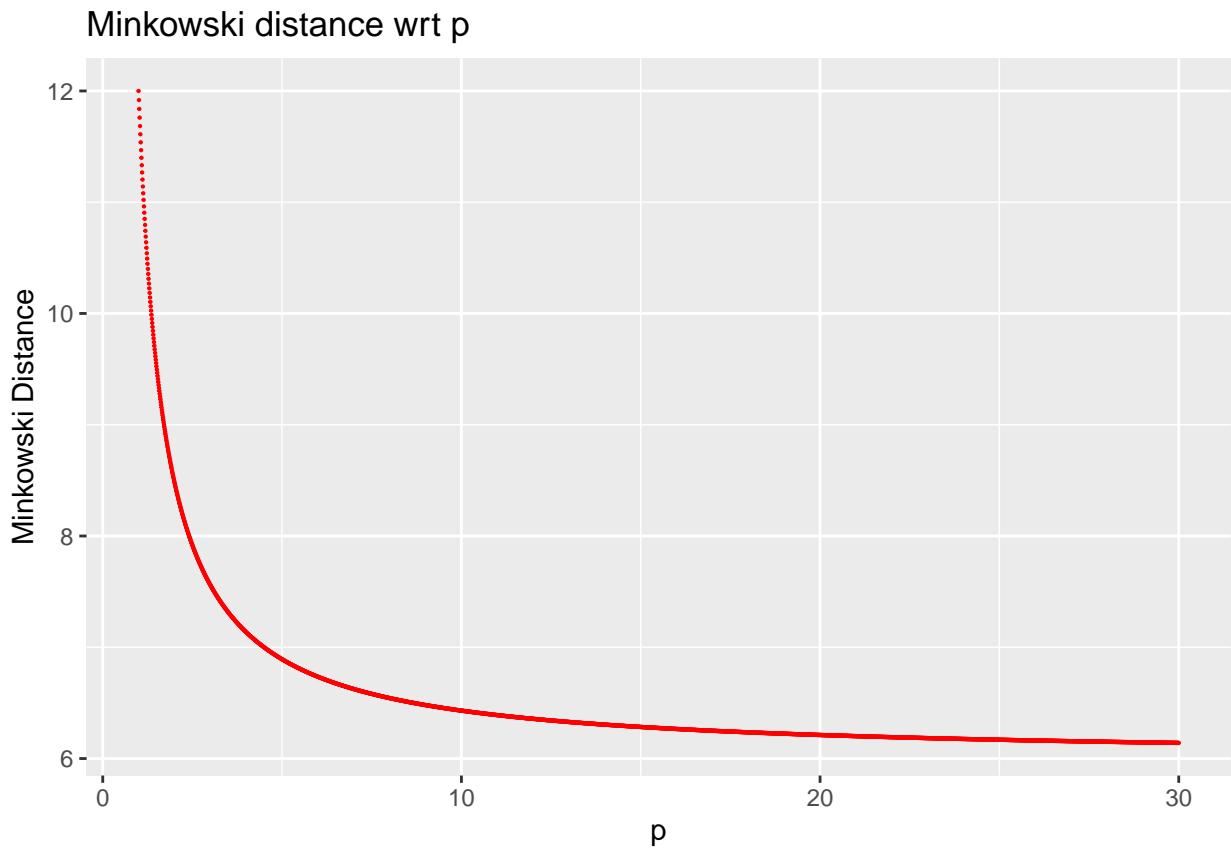
- Conversely, we have:

$$\|\mathbf{x}\|_p = d(\mathbf{x}, \mathbf{0}),$$

where $\mathbf{0}$ is the null-vetor of \mathbb{R}^m .

```
library("ggplot2")
x = c(0, 0)
y = c(6,6)
MinkowDist=c() # Initialiser à vide la liste
for (p in seq(1,30,.01))
{
  MinkowDist=c(MinkowDist,dist(rbind(x, y), method = "minkowski", p = p))
}

ggplot(data =data.frame(x = seq(1,30,.01), y=MinkowDist ) , mapping = aes( x=x, y= y))+
  geom_point(size=.1,color="red")+
  xlab("p")+ylab("Minkowski Distance")+ggtitle("Minkowski distance wrt p")
```



Exercice 4

Produce a similar graph using “The Economist” theme. Indicate on the graph the Manhattan, the Euclidian distances as well as the Chebyshev distance introduced below.

Chebyshev distance

- At the limit, we get the Chebyshev distance which is defined by:

$$d(\mathbf{x}, \mathbf{y}) = \max_{j=1, \dots, n} (|x_j - y_j|) = \lim_{p \rightarrow \infty} \left[\sum_{j=1}^n |x_j - y_j|^p \right]^{1/p}.$$

- The corresponding norm is:

$$\|\mathbf{x}\|_\infty = \max_{j=1, \dots, n} (|x_j|).$$

Minkowski inequality

- The proof of the triangular inequality A3 is based on the Minkowski inequality:
- For any nonnegative real numbers $a_1, \dots, a_m; b_1, \dots, b_m$, and for any $p \geq 1$, we have:

$$\left[\sum_{j=1}^m (a_j + b_j)^p \right]^{1/p} \leq \left[\sum_{j=1}^m a_j^p \right]^{1/p} + \left[\sum_{j=1}^m b_j^p \right]^{1/p}.$$

- To prove that the Minkowski distance satisfies A3, notice that

$$\sum_{j=1}^m |x_j - z_j|^p = \sum_{j=1}^m |(x_j - y_j) + (y_j - z_j)|^p.$$

- Since for any reals x, y , we have: $|x + y| \leq |x| + |y|$, and using the fact that x^p is increasing in $x \geq 0$, we obtain:

$$\sum_{j=1}^m |x_j - z_j|^p \leq \sum_{j=1}^m (|x_j - y_j| + |y_j - z_j|)^p.$$

- Applying the Minkowski inequality with $a_j = |x_j - y_j|$ and $b_j = |y_j - z_j|$, $j = 1, \dots, n$, we get:

$$\sum_{j=1}^m |x_j - z_j|^p \leq \left(\sum_{j=1}^m |x_j - y_j|^p \right)^{1/p} + \left(\sum_{j=1}^m |y_j - z_j|^p \right)^{1/p}.$$

Exercice 5

To illustrate the Minkowski inequality, draw 100 times two lists of 100 draws from the lognormal distribution with mean 1600 and standard-deviation 300. Illustrate with a graph the gap between the two drawn lists.

Hölder inequality

- The proof of the Minkowski inequality itself requires the Hölder inequality:
- For any nonnegative real numbers $a_1, \dots, a_m; b_1, \dots, b_m$, and any $p, q > 1$ with $1/p + 1/q = 1$, we have:

$$\sum_{j=1}^m a_j b_j \leq \left[\sum_{j=1}^m a_j^p \right]^{1/p} \left[\sum_{j=1}^m b_j^q \right]^{1/q}$$

- The proof of the Hölder inequality relies on the Young inequality:
- For any $a, b > 0$, we have

$$ab \leq \frac{a^p}{p} + \frac{b^q}{q},$$

with equality occurring iff: $a^p = b^q$.

- To prove the Young inequality, one can use the (strict) convexity of the exponential function.
- For any reals x, y , we have:

$$e^{\frac{x}{p} + \frac{y}{q}} \leq \frac{e^x}{p} + \frac{e^y}{q}.$$

- We then set: $x = p \ln a$ and $y = q \ln b$ to get the Young inequality.
- A good reference on inequalities is: Z. Cvetkovski, Inequalities: theorems, techniques and selected problems, 2012, Springer Science & Business Media.

Cauchy-Schwartz inequality

- Note that the triangular inequality for the Minkowski distance implies:

$$\sum_{j=1}^m |x_j| \leq \left[\sum_{j=1}^m |x_j|^p \right]^{1/p}.$$

- Note that for $p = 2$, we have $q = 2$. The Hölder inequality implies for that special case

$$\sum_{j=1}^m |x_j y_j| \leq \sqrt{\sum_{j=1}^m x_j^2} \sqrt{\sum_{j=1}^m y_j^2}.$$

- Since the LHS of the above inequality is greater than $|\sum_{j=1}^m x_j y_j|$, we get the Cauchy-Schwartz inequality

$$|\sum_{j=1}^m x_j y_j| \leq \sqrt{\sum_{j=1}^m x_j^2} \sqrt{\sum_{j=1}^m y_j^2}.$$

- Using the dot product notation called also scalar product notation: $\mathbf{x} \cdot \mathbf{y} = \sum_{j=1}^m x_j y_j$, and the norm notation $\|\cdot\|_2$, the Cauchy-Schwartz inequality is:

$$|\mathbf{x} \cdot \mathbf{y}| \leq \|\mathbf{x}\|_2 \|\mathbf{y}\|_2.$$

Pearson correlation distance

- The Pearson correlation coefficient is a similarity measure on \mathbb{R}^m defined by:

$$\rho(\mathbf{x}, \mathbf{y}) = \frac{\sum_{j=1}^m (x_j - \bar{x})(y_j - \bar{y})}{\sqrt{\sum_{j=1}^m (x_j - \bar{x})^2 \sum_{j=1}^m (y_j - \bar{y})^2}},$$

where $\bar{\mathbf{x}}$ is the mean of the vector \mathbf{x} defined by:

$$\bar{\mathbf{x}} = \frac{1}{n} \sum_{j=1}^m x_j,$$

- Note that the Pearson correlation coefficient satisfies P2 and is invariant to any positive linear transformation, i.e.:

$$\rho(\alpha\mathbf{x}, \mathbf{y}) = \rho(\mathbf{x}, \mathbf{y}),$$

for any $\alpha > 0$.

- The Pearson distance (or correlation distance) is defined by:

$$d(\mathbf{x}, \mathbf{y}) = 1 - \rho(\mathbf{x}, \mathbf{y}).$$

- Note that the Pearson distance does not satisfy A1 since $d(\mathbf{x}, \mathbf{x}) = 0$ for any non-zero vector \mathbf{x} . It neither satisfies the triangle inequality. However, the symmetry property is fulfilled.

Cosine correlation distance

- The cosine of the angle θ between two vectors \mathbf{x} and \mathbf{y} is a measure of similarity given by:

$$\cos(\theta) = \frac{\mathbf{x} \cdot \mathbf{y}}{\|\mathbf{x}\|_2 \|\mathbf{y}\|_2} = \frac{\sum_{j=1}^m x_j y_j}{\sqrt{\sum_{j=1}^m x_j^2 \sum_{j=1}^m y_j^2}}.$$

- Note that the cosine of the angle between the two centred vectors $\mathbf{x} - \bar{\mathbf{x}}\mathbf{1}$ and $\mathbf{y} - \bar{\mathbf{y}}\mathbf{1}$ coincides with the Pearson correlation coefficient of \mathbf{x} and \mathbf{y} , where $\mathbf{1}$ is a vector of units of \mathbb{R}^m .
- The cosine correlation distance is defined by:

$$d(\mathbf{x}, \mathbf{y}) = 1 - \cos(\theta).$$

- It shares similar properties than the Pearson correlation distance. Likewise, Axioms A1 and A3 are not satisfied.

Spearman correlation distance

- To calculate the Spearman's rank-order correlation, we need to map separately each of the vectors to ranked data values:

$$\mathbf{x} \rightarrow \text{rank}(\mathbf{x}) = (x_1^r, \dots, x_m^r).$$

- Here, x_j^r is the rank of x_j among the set of values of \mathbf{x} .
- We illustrate this transformation with a simple example:
- If $\mathbf{x} = (3, 1, 4, 15, 92)$, then the rank-order vector is $\text{rank}(\mathbf{x}) = (2, 1, 3, 4, 5)$.

```
x=c(3, 1, 4, 15, 92)
rank(x)
```

```
## [1] 2 1 3 4 5
```

- The Spearman's rank correlation of two numerical vectors \mathbf{x} and \mathbf{y} is simply the Pearson correlation of the two corresponding rank-order vectors $\text{rank}(\mathbf{x})$ and $\text{rank}(\mathbf{y})$, i.e. $\rho(\text{rank}(\mathbf{x}), \text{rank}(\mathbf{y}))$. This measure is useful because it is more robust against outliers than the Pearson correlation.

- If all the n ranks are distinct, it can be computed using the following formula:

$$\rho(\text{rank}(\mathbf{x}), \text{rank}(\mathbf{y})) = 1 - \frac{6 \sum_{j=1}^m d_j^2}{n(n^2 - 1)},$$

where $d_j = x_j^r - y_j^r$, $j = 1, \dots, n$.

- The spearman distance is then defined by:

$$d(\mathbf{x}, \mathbf{y}) = 1 - \rho(\text{rank}(\mathbf{x}), \text{rank}(\mathbf{y})).$$

- It can be shown that easaly that it is not a proper distance.
- If all the n ranks are distinct, we get:

$$d(\mathbf{x}, \mathbf{y}) = \frac{6 \sum_{j=1}^m d_j^2}{n(n^2 - 1)}.$$

```
x=c(3, 1, 4, 15, 92)
rank(x)
```

```
## [1] 2 1 3 4 5
```

```
y=c(30,2 , 9, 20, 48)
rank(y)
```

```
## [1] 4 1 2 3 5
```

```
d=rank(x)-rank(y)
d
```

```
## [1] -2  0  1  1  0
```

```
cor(rank(x),rank(y))
```

```
## [1] 0.7
```

```
1-6*sum(d^2)/(5*(5^2-1))
```

```
## [1] 0.7
```

Exercice 6

- For the two vectors $\mathbf{x} = (22, 34, 1, 12, 25, 56, 7)$ and $\mathbf{y} = (2, 64, 12, 2, 22, 5, 8)$:
- Calculate the ranks for each vector.
- Deduce the Spearman correlation distance from that ranks.
- Deduce the Spearman correlation distance from the above dispalyed alternative equation.
- Calculate the Spearman correlation distance using the **R** function.

Kendall tau distance

- The Kendall rank correlation coefficient is calculated from the number of correspondances between the rankings of \mathbf{x} and the rankings of \mathbf{y} .
- The number of pairs of observations among n observations or values is:

$$\binom{n}{2} = \frac{n(n-1)}{2}.$$

- The pairs of observations (x_i, x_j) and (y_i, y_j) are said to be *concordant* if:

$$\text{sign}(x_j - x_i) = \text{sign}(y_j - y_i),$$

and to be *discordant* if:

$$\text{sign}(x_j - x_i) = -\text{sign}(y_j - y_i),$$

where $\text{sign}(\cdot)$ returns 1 for positive numbers and -1 negative numbers and 0 otherwise.

- If $x_j = x_i$ or $y_j = y_i$ (or both), there is a tie.
- The Kendall τ coefficient is defined by (neglecting ties):

$$\tau = \frac{1}{n(n-1)} \sum_{j=1}^n \sum_{i=1}^m \text{sign}(x_j - x_i)\text{sign}(y_j - y_i).$$

- Let n_c (resp. n_d) be the number of concordant (resp. discordant) pairs, we have

$$\tau = \frac{2(n_c - n_d)}{n(n-1)}.$$

- The Kendall tau distance is then:

$$d(\mathbf{x}, \mathbf{y}) = 1 - \tau.$$

- Remark: the triangular inequality may fail in cases where there are ties.

```

x=c(3, 1, 4, 15, 92)
y=c(30, 2, 9, 20, 48)
tau=0
for (i in 1:5)
{
  tau=tau+sign(x -x[i])*%*%sign(y -y[i])
}
tau=tau/(5*4)
tau

##      [,1]
## [1,]  0.6
cor(x,y, method="kendall")

## [1] 0.6

```

Exercice 7

- For the two vectors $\mathbf{x} = (22, 34, 1, 12, 25, 56, 7)$ and $\mathbf{y} = (2, 64, 12, 2, 22, 5, 8)$:
- List all pairs of coordinates.
- How many pairs are there?
- For each pair and each cector, compute the signs of the differences in coordinates.
- Deduce the Kendall tau coefficient using the above computations.
- Calculate the Kendall tau coefficient using the R function.

Standardization

- Variables or measurements are often standardized before calculating dissimilarities.
- Standardization converts the original variables into unitless variables.

- A well known method is the z-score transformation.
- Let $\mathbf{v} \equiv (v_1, \dots, v_n)$ a vector of measurements recorded for n individuals or objects.

$$\mathbf{v} \rightarrow \left(\frac{v_1 - \bar{\mathbf{v}}}{s_{\mathbf{v}}}, \dots, \frac{v_n - \bar{\mathbf{v}}}{s_{\mathbf{v}}} \right),$$

where $\bar{\mathbf{v}}, s_{\mathbf{v}}$ are the sample mean and standard-deviation, respectively, given by:

$$\bar{\mathbf{v}} = \frac{1}{n} \sum_{i=1}^n v_i, \quad s_{\mathbf{v}} = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (v_i - \bar{\mathbf{v}})^2}.$$

- The transformed variable will have a mean of 0 and a variance of 1.
- The result obtained with Pearson correlation measures and standardized Euclidean distances are comparable.
- For other methods, see: Milligan, G. W., & Cooper, M. C. (1988). A study of standardization of variables in cluster analysis. *Journal of classification*, 5(2), 181-204

```
v=c(3, 1, 4, 15, 92)
w=c(30,2 , 9, 20, 48)
(v-mean(v))/sd(v)

## [1] -0.5134116 -0.5647527 -0.4877410 -0.2053646  1.7712699
scale(v)

##          [,1]
## [1,] -0.5134116
## [2,] -0.5647527
## [3,] -0.4877410
## [4,] -0.2053646
## [5,]  1.7712699
## attr(,"scaled:center")
## [1] 23
## attr(,"scaled:scale")
## [1] 38.9551
(w-mean(w))/sd(w)

## [1]  0.45263128 -1.09293895 -0.70654639 -0.09935809  1.44621214
scale(w)

##          [,1]
## [1,]  0.45263128
## [2,] -1.09293895
## [3,] -0.70654639
## [4,] -0.09935809
## [5,]  1.44621214
## attr(,"scaled:center")
## [1] 21.8
## attr(,"scaled:scale")
## [1] 18.11629
```

Exercice 8

Table 3 Age (in years) and Height (in centimeters) of Four

Person	Age (yr)	Height (cm)
A	35	190
B	40	190
C	35	160
D	40	160

- Consider the following example
- Plot the data using a nice scatter plot.
- Transform the Height from centimeters (cm) into feet (ft).
- Display your data in a table.
- Plot the data within a new scatter plot.
- What do you observe?
- Standardize the two variables Age and Height.
- Display your data in a table.
- Plot the standardized data within a new scatter plot.
- Conclude.

Similarity measures for binary data

- A common simple situation occurs when all information is of the presence/absence of 2-level qualitative characters.
- We assume there are n characters.
- *The presence of the character is coded by 1 and the absence by 0.
- We have have at our disposal two vectors.
- \mathbf{x} is observed for a first individual (or object).
- \mathbf{y} is observed for a second individual.
- We can then calculate the following four statistics:

$$a = \mathbf{x} \cdot \mathbf{y} = \sum_{j=1}^m x_j y_j.$$

$$b = \mathbf{x} \cdot (\mathbf{1} - \mathbf{y}) = \sum_{j=1}^m x_j (1 - y_j).$$

$$c = (\mathbf{1} - \mathbf{x}) \cdot \mathbf{y} = \sum_{j=1}^m (1 - x_j) y_j.$$

$$d = (\mathbf{1} - \mathbf{x}) \cdot (\mathbf{1} - \mathbf{y}) = \sum_{j=1}^m (1 - x_j) (1 - y_j).$$

- The counts of matches are a for (1, 1) and d for (0, 0);

- The counts of mismatches are b for $(1, 0)$ and c for $(0, 1)$.
- Note that obviously: $a + b + c + d = n$.
- This gives a very useful 2×2 association table.

		Second individual		
		1	0	Totals
First individual	1	a	b	$a + b$
	0	c	d	$c + d$
Totals		$a + c$	$b + d$	n

Table 9 Binary Variables for Eight People

Person	Sex (Male = 1, Female = 0)	Married (Yes = 1, No = 0)	Fair Hair = 1, Dark Hair = 0	Blue Eyes = 1, Brown Eyes = 0	Wears Glasses (Yes = 1, No = 0)	Round Face = 1, Oval Face = 0	Pessimist = 1, Optimist = 0	Evening Type = 1, Morning Type = 0	Is an Only Child (Yes = 1, No = 0)	Left-Handed = 1, Right-Handed = 0
Ilan	1	0	1	1	0	0	1	0	0	0
Jacqueline	0	1	0	0	1	0	0	0	0	0
Kim	0	0	1	0	0	0	1	0	0	0
Lieve	0	1	0	0	0	0	0	1	1	1
Leon	1	1	0	0	1	0	0	1	1	0
Peter	1	1	0	0	1	0	1	1	0	0
Talia	0	0	0	1	0	1	0	0	0	0
Tina	0	0	0	1	0	1	0	0	0	0

Table from Kaufman, L., & Rousseeuw, P. J. (2009). Finding groups in data: an introduction to cluster analysis (Vol. 344). John Wiley & Sons

- The data shows 8 people (individuals) and 10 binary variables:
- Sex, Married, Fair Hair, Blue Eyes, Wears Glasses, Round Face, Pessimist, Evening Type, Is an Only Child, Left-Handed.

```
data=c(
1,0,1,1,0,0,1,0,0,0,
0,1,0,0,1,0,0,0,0,0,
```

```

0,0,1,0,0,0,1,0,0,1,
0,1,0,0,0,0,0,1,1,0,
1,1,0,0,1,1,0,1,1,0,
1,1,0,0,1,0,1,1,0,0,
0,0,0,1,0,1,0,0,0,0,
0,0,0,1,0,1,0,0,0,0
)
data=data.frame(matrix(data, nrow=8, byrow=T))
row.names(data)=c("Ilan", "Jacqueline", "Kim", "Lieve", "Leon", "Peter", "Talia", "Tina")
names(data)=c("Sex", "Married", "Fair Hair", "Blue Eyes", "Wears Glasses", "Round Face", "Pessimist", "P")

```

- We are comparing the records for Ilan with Talia.

```

library(knitr)
library(xtable)
library(stargazer)
library(texreg)
library(kableExtra)
library(summarytools)

## Warning in fun(libname, pkgname): couldn't connect to display ":0"
set.seed(893)
datat<-as.data.frame(t(data))
datat=lapply(datat, as.factor)
Ilan=datat$Ilan
Talia =datat$Talia
print(ctable(Ilan, Talia, prop = 'n', style = "rmarkdown"))

```

Cross-Tabulation

Ilan * Talia

	Talia	0	1	Total
Ilan				
0		5	1	6
1		3	1	4
Total		8	2	10

- Therefore: $a = 1$, $b = 3$, $c = 1$, $d = 5$.
- Note that interchanging Ilan and Talia would permute b and c while leaving a and d unchanged.
- A good similarity or dissimilarity coefficient must treat b and c symmetrically.
- A similarity measure is denoted by: $s(\mathbf{x}, \mathbf{y})$.
- The corresponding distance is then defined as:

$$d(\mathbf{x}, \mathbf{y}) = 1 - s(\mathbf{x}, \mathbf{y}).$$

- Alternatively, we have:

$$d(\mathbf{x}, \mathbf{y}) = \sqrt{1 - s(\mathbf{x}, \mathbf{y})}.$$

- A list of some of the similarity measures $s(\mathbf{x}, \mathbf{y})$ that have been suggested for binary data is shown below.
- An more complete list can be found in: Gower, J. C., & Legendre, P. (1986). Metric and Euclidean properties of dissimilarity coefficients. *Journal of classification*, 3(1), 5-48.

Coefficient	$s(\mathbf{x}, \mathbf{y})$	$d(\mathbf{x}, \mathbf{y}) = 1 - s(\mathbf{x}, \mathbf{y})$
Simple matching	$\frac{a+d}{a+b+c+d}$	$\frac{b+c}{a+b+c+d}$
Jaccard	$\frac{a}{a+b+c}$	$\frac{b+c}{a+b+c}$
Rogers and Tanimoto (1960)	$\frac{a+d}{a+2(b+c)+d}$	$\frac{2(b+c)}{a+2(b+c)+d}$
Gower and Legendre (1986)	$\frac{2(a+d)}{2(a+d)+b+c}$	$\frac{b+c}{2(a+d)+b+c}$
Gower and Legendre (1986)	$\frac{2a}{2a+b+c}$	$\frac{b+c}{2a+b+c}$

- To calculate these coefficients, we use the function: `dist.binary()`. available in the **ade4** package.
- All the distances in the **ade4** package are of type $d(\mathbf{x}, \mathbf{y}) = \sqrt{1 - s(\mathbf{x}, \mathbf{y})}$.

```
library(ade4)
a=1
b=3
c=1
d=5
dist.binary(data[c("Ilan","Talia"),],method=2)^2
```

Ilan

Talia 0.4

```
1-(a+d)/(a+b+c+d)
```

[1] 0.4

```
dist.binary(data[c("Ilan","Talia"),],method=1)^2
```

Ilan

Talia 0.8

```
1-a/(a+b+c)
```

[1] 0.8

```
dist.binary(data[c("Ilan","Talia"),],method=4)^2
```

Ilan

Talia 0.5714286

```
1-(a+d)/(a+2*(b+c)+d)
```

[1] 0.5714286

One Gower coefficient is missing

```
dist.binary(data[c("Ilan","Talia"),],method=5)^2
```

Ilan

Talia 0.6666667

```
1-2*a/(2*a+b+c)
```

[1] 0.6666667

- The reason for such a large number of possible measures has to do with the apparent uncertainty as to how to deal with the count of zero-zero matches d .
- The measures embedding d are sometimes called symmetrical.
- The other measures are called assymmetrical.

- In some cases, of course, zero_zero matches are completely equivalent to one-one matches, and therefore should be included in the calculated similarity measure.
- An example is gender, where there is no preference as to which of the two categories should be coded zero or one.
- But in other cases the inclusion or otherwise of d is more problematic; for example, when the zero category corresponds to the genuine absence of some property, such as wings in a study of insects.

Exercice 9

- Use the data set *animals* available in the package *cluster*.
- This data set was first used in this textbook KAUFMAN, Leonard et ROUSSEEUW, Peter J. Finding groups in data: an introduction to cluster analysis. John Wiley & Sons, 2009.
- Identify the missing measurements.
- Explain the way how KAUFMAN and ROUSSEEUW, pp. 296-297 treat the missing measurements.
- Compute a distance matrix for the completed data.
- Propose a graphical way to represent that distance matrix.
- Which group of animals look close?
- Change the method of calculating and observe if it has some effect of the graph.

After the identification of missing measurements, a procedure is carried out for estimating their values. In this procedure each variable containing missing values is considered in turn. Each time the algorithm looks for the most similar complete variable and then uses the latter for filling in the missing values. In our example END has two missing values. The similarities between this variable and the complete variables are given in Figure 7.

The variable WAR has the highest similarity with END and is therefore the most appropriate for estimating the missing values of END. The two

<u>variable</u>	<u>END</u>	<u>similarity</u>								
WAR	<table border="1"> <tr> <td>1</td><td>0</td></tr> <tr> <td>1</td><td>5</td><td>4</td></tr> <tr> <td>0</td><td>1</td><td>8</td></tr> </table>	1	0	1	5	4	0	1	8	36
1	0									
1	5	4								
0	1	8								
FLY	<table border="1"> <tr> <td>1</td><td>0</td></tr> <tr> <td>1</td><td>1</td><td>3</td></tr> <tr> <td>0</td><td>5</td><td>9</td></tr> </table>	1	0	1	1	3	0	5	9	6
1	0									
1	1	3								
0	5	9								
VER	<table border="1"> <tr> <td>1</td><td>0</td></tr> <tr> <td>1</td><td>6</td><td>7</td></tr> <tr> <td>0</td><td>0</td><td>5</td></tr> </table>	1	0	1	6	7	0	0	5	30
1	0									
1	6	7								
0	0	5								
HAI	<table border="1"> <tr> <td>1</td><td>0</td></tr> <tr> <td>1</td><td>2</td><td>5</td></tr> <tr> <td>0</td><td>4</td><td>7</td></tr> </table>	1	0	1	2	5	0	4	7	6
1	0									
1	2	5								
0	4	7								

Figure 7 Similarities between a variable with missing values (END) and all variables without missing values, in the animal data set.

REVISED DATA

	W	F	V	E	G	H
	A	L	E	N	R	A
	R	Y	R	D	O	I

ant	0	0	0	0	1	0
bee	0	1	0	0	1	1
cat	1	0	1	0	0	1
cpl	0	0	0	0	0	1
chi	1	0	1	1	1	1
cow	1	0	1	0	1	1
duc	1	1	1	0	1	0
eag	1	1	1	1	0	0
ele	1	0	1	1	1	0
fly	0	1	0	0	0	0
fro	0	0	1	1	0	0
her	0	0	1	0	1	0
lio	1	0	1	1	1	1
liz	0	0	1	0	0	0
lob	0	0	0	0	0	0
man	1	0	1	1	1	1
rab	1	0	1	0	1	1
sal	0	0	1	0	0	0
spi	0	0	0	0	0	1
wha	1	0	1	1	1	0

Exercice 10

- Prove that the distances based on the Simple Matching coefficient and the Jaccard coefficient satisfy A3.
- Prove that the distances proposed by Gower and Legendre (1986) do not satisfy A3.
- Hint: Proofs and counterexamples have to be adapted from in the paper: Gower, J. C., & Legendre, P. (1986). Metric and Euclidean properties of dissimilarity coefficients. *Journal of classification*, 3(1), 5-48.

Nominal variables

- We previously studied above binary variables which can only take on two states coded 0, 1.
- We generalize this approach to nominal variables which may take on more than two states.
- Eye's color may have for example four states: blue, brown, green, grey.
- Let M be the number of states and code the outcomes as $1, \dots, M$.
- We may choose $1 = \text{blue}$, $2 = \text{brown}$, $3 = \text{green}$, and $4 = \text{grey}$.
- These states are not ordered in any way
- One strategy would be creating a new binary variable for each of the M nominal states.
- Then to put it equal to 1 if the corresponding state occurs and to 0 otherwise.
- After that, one could resort to one of the dissimilarity coefficients of the previous subsection.
- The most common way of measuring the similarity or dissimilarity between two objects through categorial variables is the simple matching approach.
- If \mathbf{x}, \mathbf{y} , are both m nominal records for two individuals,
- Let define the function:

$$\delta(x_j, y_j) \equiv \begin{cases} 0, & \text{if } x_j = y_j; \\ 1, & \text{if } x_j \neq y_j. \end{cases}$$

- Let N_{a+d} be the number of attributes of the two individuals on which the two records match:

$$N_{a+d} = \sum_{j=1}^m [1 - \delta(x_j, y_j)].$$

- Let N_{b+c} be the number of attributes on which the two records do not match:

$$N_{b+c} = \sum_{j=1}^m \delta(x_j, y_j).$$

- Let N_d be the number of attributes on which the two records match in a “not applicable” category.
- The distance corresponding to the simple matching approach is:

$$d(\mathbf{x}, \mathbf{y}) = \frac{\sum_{j=1}^m \delta(x_j, y_j)}{n} = \frac{N_{a+d}}{N_{a+d} + N_{b+c}}.$$

- Note that simple matching has exactly the same meaning as in the preceding section.
- For more details, see GAN, Guojun, MA, Chaoqun, et WU, Jianhong. Data clustering: theory, algorithms, and applications. Society for Industrial and Applied Mathematics, 2020.

Table 6.2. Some matching coefficients for nominal data.

Measure	$s(\mathbf{x}, \mathbf{y})$	Weighting of matches, mismatches
Russell and Rao	$\frac{N_{a+d} - N_d}{N_{a+d} + N_{b+c}}$	Equal weights
Simple matching	$\frac{N_{a+d}}{N_{a+d} + N_{b+c}}$	Equal weights
Jaccard	$\frac{N_{a+d} - N_d}{N_{a+d} - N_d + N_{b+c}}$	Equal weights
Unnamed	$\frac{2N_{a+d}}{2N_{a+d} + N_{b+c}}$	Double weight for matched pairs
Dice	$\frac{2N_{a+d} - 2N_d}{2N_{a+d} - 2N_d + N_{b+c}}$	Double weight for matched pairs
Rogers-Tanimoto	$\frac{N_{a+d}}{N_{a+d} + 2N_{b+c}}$	Double weight for unmatched pairs
Unnamed	$\frac{N_{a+d} - N_d}{N_{a+d} - N_d + 2N_{b+c}}$	Double weight for unmatched pairs
Kulczynski	$\frac{N_{a+d} - N_d}{N_{b+c}}$	Matched pairs excluded from denominator
Unnamed	$\frac{N_{a+d}}{N_{b+c}}$	Matched pairs excluded from denominator

Figure 1: From: GAN et al

Gower's dissimilarity

- Gower's coefficient is a dissimilarity measure specifically designed for handling mixed attribute types or variables.
- See: GOWER, John C. A general coefficient of similarity and some of its properties. *Biometrics*, 1971, p. 857-871.
- The coefficient is calculated as the weighted average of attribute contributions.
- Weights usually used only to indicate which attribute values could actually be compared meaningfully.
- The formula is:

$$d(\mathbf{x}, \mathbf{y}) = \frac{\sum_{j=1}^m w_j \delta(x_j, y_j)}{\sum_{j=1}^m w_j}.$$

- The weight w_j is put equal to 1 when both measurements x_j and y_j are nonmissing,
- The number $\delta(x_j, y_j)$ is the contribution of the j th measure or variable to the dissimilarity measure.
- If the j th measure is nominal, we take

$$\delta(x_j, y_j) \equiv \begin{cases} 0, & \text{if } x_j = y_j; \\ 1, & \text{if } x_j \neq y_j. \end{cases}$$

- If the j th measure is interval-scaled, we take instead:

$$\delta(x_j, y_j) \equiv \frac{|x_j - y_j|}{R_j},$$

where R_j is the range of variable j over the available data.

- Consider the following data set:

object	variable							
	1	2	3	4	5	6	7	8
Begonia	0	1	1	4	3	15	25	15
Broom	1	0	0	2	1	3	150	50
Camellia	0	1	0	3	3	1	150	50
Dahlia	0	0	1	4	2	16	125	50
Forget-me-not	0	1	0	5	2	2	20	15
Fuchsia	0	1	0	4	3	12	50	40
Geranium	0	0	0	4	3	13	40	20
Gladiolus	0	0	1	2	2	7	100	15
Heather	1	1	0	3	1	4	25	15
Hydrangea	1	1	0	5	2	14	100	60
Iris	1	1	1	5	3	8	45	10
Lily	1	1	1	1	2	9	90	25
Lily-of-the-valley	1	1	0	1	2	6	20	10
Peony	1	1	1	4	2	11	80	30
Pink Carnation	1	0	0	3	2	10	40	20
Red Rose	1	0	0	4	2	18	200	60
Scotch Rose	1	0	0	2	2	17	150	60
Tulip	0	0	1	2	1	5	25	10

Table 1: Flower dataset.

Data

from: Struyf, A., Hubert, M., & Rousseeuw, P. (1997). Clustering in an object-oriented environment. *Journal of Statistical Software*, 1(4), 1-30.

- The dataset contains 18 flowers and 8 characteristics:
 1. Winters: binary, indicates whether the plant may be left in the garden when it freezes.
 2. Shadow: binary, shows whether the plant needs to stand in the shadow.
 3. Tubers (Tubercale): asymmetric binary, distinguishes between plants with tubers and plants that grow in any other way.
 4. Color: nominal, specifies the flower's color (1=white, 2=yellow, 3= pink, 4=red, 5= blue).
 5. Soil: ordinal, indicates whether the plant grows in dry (1), normal (2), or wet (3) soil.
 6. Preference: ordinal, someone's preference ranking, going from 1 to 18.
 7. Height: interval scaled, the plant's height in centimeters.
 8. Distance: interval scaled, the distance in centimeters that should be left between the plants.
- The dissimilarity between Begonia and Broom (Genêt) can be calculated as follows:



Begonia



Genêt

```

library(cluster)
library(dplyr)
data <- flower %>%
  rename(Winters=V1,Shadow=V2,Tubers=V3,Color=V4,Soil=V5,Preference=V6,Height=V7,Distance=V8) %>%
  mutate(Winters=recode(Winters, "1"="Yes", "0"="No"),
         Shadow=recode(Shadow, "1"="Yes", "0"="No"),
         Tubers=recode(Tubers, "1"="Yes", "0"="No"),
         Color=recode(Color, "1"="white", "2"="yellow", "3"= "pink", "4"="red", "5"="blue"),
         Soil=recode(Soil, "1"="dry", "2"="normal", "3"= "wet")
     )
row.names(data)=c("Begonia", "Broom", "Camellia", "Dahlia", "Forget-me-not", "Fuchsia",
  "Geranium", "Gladiolus", "Heather", "Hydrangea", "Iris", "Lily", "Lily-of-the-valley",
  "Peony", "Pink Carnation", "Red Rose", "Scotch Rose", "Tulip")

res=lapply(data,class)
res=as.data.frame(res)
res[1,] %>%
knitr::kable()

```

Winters	Shadow	Tubers	Color	Soil	Preference	Height	Distance
factor	factor	factor	factor	ordered	ordered	numeric	numeric

```

flower[1:2,]

##   V1 V2 V3 V4 V5 V6 V7 V8
## 1  0  1  1  4  3 15 25 15
## 2  1  0  0  2  1  3 150 50
max(data$Height)-min(data$Height)

## [1] 180
max(data$Distance)-min(data$Distance)

## [1] 50

```

$$\frac{|1 - 0| + |0 - 1| + |0 - 1| + 1 + |1 - 3|/2 + |3 - 15|/17 + |150 - 25|/180 + |50 - 15|/50}{8} \approx 0.8875408$$

Daisy function

<code>daisy</code>	<i>Dissimilarity Matrix Calculation</i>
--------------------	---

Description

Compute all the pairwise dissimilarities (distances) between observations in the data set. The original variables may be of mixed types. In that case, or whenever `metric = "gower"` is set, a generalization of Gower's formula is used, see 'Details' below.

Usage

```

daisy(x, metric = c("euclidean", "manhattan", "gower"),
      stand = FALSE, type = list(), weights = rep.int(1, p),
      warnBin = warnType, warnAsym = warnType, warnConst = warnType,
      warnType = TRUE)

```

Cluster package description available at this link.

```

library(cluster)
(abs(1-0)+abs(0-1)+abs(0-1)+1+abs(1-3)/2+abs(3-15)/17+abs(150-25)/180+abs(50-15)/50)/8

## [1] 0.8875408

dist<-daisy(data[,1:8],metric = "Gower")
as.matrix(dist)[1:2,1:2]

##          Begonia      Broom
## Begonia 0.0000000 0.8875408
## Broom    0.8875408 0.0000000

```

More on distance matrix computation

USArrests

From [datasets v3.6.2](#)
by R-core R-core@R-project.org 99.99th
Percentile

Violent Crime Rates By US State

This data set contains statistics, in arrests per 100,000 residents for assault, murder, and rape in each of the 50 US states in 1973. Also given is the percent of the population living in urban areas.

Keywords [datasets](#)

Usage

```
USArrests
```

Note

`USArrests` contains the data as in McNeil's monograph. For the `UrbanPop` percentages, a review of the table (No. 21) in the Statistical Abstracts 1975 reveals a transcription error for Maryland (and that McNeil used the same "round to even" rule that R's `round()` uses), as found by Daniel S Coven (Arizona).

See the example below on how to correct the error and improve accuracy for the '<n>.5' percentages.

Format

A data frame with 50 observations on 4 variables.

[,1]	Murder	numeric	Murder arrests (per 100,000)
[,2]	Assault	numeric	Assault arrests (per 100,000)
[,3]	UrbanPop	numeric	Percent urban population

References

McNeil, D. R. (1977) *Interactive Data Analysis*. New York: Wiley.

- We use a subset of the data by taking 15 random rows among the 50 rows in the data set.
- We are using the function sample().
- We standardize the data using the function scale().

```
stargazer(USArrests,header=TRUE, type='html',summary=FALSE,digits=1)
```

Murder

Assault

UrbanPop

Rape

Alabama

13.2

236

58

21.2

Alaska

10

263

48

44.5

Arizona

8.1

294

80

31

Arkansas

8.8

190

50

19.5

California

9

276

91

40.6

Colorado

7.9

204

78

38.7

Connecticut

3.3

110

77

11.1

Delaware

5.9

238

72

15.8

Florida

15.4

335

80

31.9

Georgia

17.4

211

60

25.8

Hawaii

5.3

46

83

20.2

Idaho

2.6

120

54

14.2

Illinois

10.4

249

83

24

Indiana

7.2

113

65

21

Iowa

2.2

56

57

11.3

Kansas

6

115

66

18

Kentucky

9.7

109

52

16.3

Louisiana

15.4

249

66

22.2

Maine

2.1

83

51

7.8

Maryland

11.3

300

67

27.8

Massachusetts

4.4

149

85

16.3

Michigan

12.1

255

74

35.1

Minnesota

2.7

72

66

14.9

Mississippi

16.1

259

44

17.1

Missouri

9

178

70

28.2

Montana

6

109

53

16.4

Nebraska

4.3

102

62

16.5

Nevada

12.2

252
81

46

New Hampshire

2.1

57

56

9.5

New Jersey

7.4

159

89

18.8

New Mexico

11.4

285

70

32.1

New York

11.1

254

86

26.1

North Carolina

13

337

45

16.1

North Dakota

0.8

45

44

7.3

Ohio

7.3

120

75	
21.4	
Oklahoma	
6.6	
151	
68	
20	
Oregon	
4.9	
159	
67	
29.3	
Pennsylvania	
6.3	
106	
72	
14.9	
Rhode Island	
3.4	
174	
87	
8.3	
South Carolina	
14.4	
279	
48	
22.5	
South Dakota	
3.8	
86	
45	
12.8	
Tennessee	
13.2	
188	
59	

26.9

Texas

12.7

201

80

25.5

Utah

3.2

120

80

22.9

Vermont

2.2

48

32

11.2

Virginia

8.5

156

63

20.7

Washington

4

145

73

26.2

West Virginia

5.7

81

39

9.3

Wisconsin

2.6

53

66

10.8

Wyoming

6.8

161

60

15.6

```
set.seed(123)
ss <- sample(1:50, 15)
df <- USArrests[ss, ]
df.scaled <- scale(df)
stargazer(df.scaled, header=TRUE, type='html', summary=FALSE, digits=1)
```

Murder

Assault

UrbanPop

Rape

New Mexico

0.6

1.0

0.2

0.6

Iowa

-1.7

-1.5

-0.7

-1.4

Indiana

-0.5

-0.9

-0.1

-0.5

Arizona

-0.2

1.1

0.9

0.5

Tennessee

1.0

-0.1

-0.5
0.1
Texas
0.9
0.1
0.9
-0.04
Oregon
-1.0
-0.4
0.01
0.3
West Virginia
-0.8
-1.3
-2.0
-1.6
Missouri
-0.01
-0.2
0.2
0.2
Montana
-0.8
-1.0
-1.0
-0.9
Nebraska
-1.2
-1.0
-0.3
-0.9
California
-0.01
0.9
1.7

1.4
South Carolina

1.3

1.0

-1.3

-0.3

Nevada

0.8

0.7

1.0

2.0

Florida

1.6

1.6

0.9

0.6

- The R functions for computing distances.
 1. `dist()` function accepts only numeric data.
 2. `get_dist()` function [factoextra package] accepts only numeric data. it supports correlation-based distance measures.
 3. `daisy()` function [cluster package] is able to handle other variable types (nominal, ordinal, ...).
 - Remark: All these functions compute distances between rows of the data.
 - Remark: If we want to compute pairwise distances between variables, we must transpose the data to have variables in the rows.
 - We first compute Euclidian distances

```
dist.eucl <- dist(df.scaled, method = "euclidean", upper = TRUE)
```

```
stargazer(as.data.frame(as.matrix(dist.eucl)[1:3, 1:3]), header=TRUE, type='html', summary=FALSE, digits=1)
```

New Mexico

Iowa

Indiana

New Mexico

0

4.1

2.5

Iowa

4.1

```

0
1.8
Indiana
2.5
1.8
0
round(sqrt(sum((df.scaled["New Mexico",]-df.scaled["Iowa",])^2)),1)

```

```

[1] 4.1
round(sqrt(sum((df.scaled["New Mexico",]-df.scaled["Indiana",])^2)),1)

```

```

[1] 2.5
round(sqrt(sum((df.scaled["Iowa",]
-df.scaled["Indiana",])^2)),1)

```

```

[1] 1.8
• We also compute correlation based distances.

```

```

library("factoextra")
dist.cor <- get_dist(df.scaled, method = "pearson")
round(as.matrix(dist.cor)[1:3, 1:3], 1)

##           New Mexico Iowa Indiana
## New Mexico      0.0   1.7    2.0
## Iowa          1.7   0.0    0.3
## Indiana       2.0   0.3    0.0
round(1-cor(df.scaled["New Mexico",],df.scaled["Iowa",]),1)
## [1] 1.7
round(1-cor(df.scaled["New Mexico",],df.scaled["Indiana",]),1)
## [1] 2
round(1-cor(df.scaled["Iowa",],df.scaled["Indiana",]),1)
## [1] 0.3

```

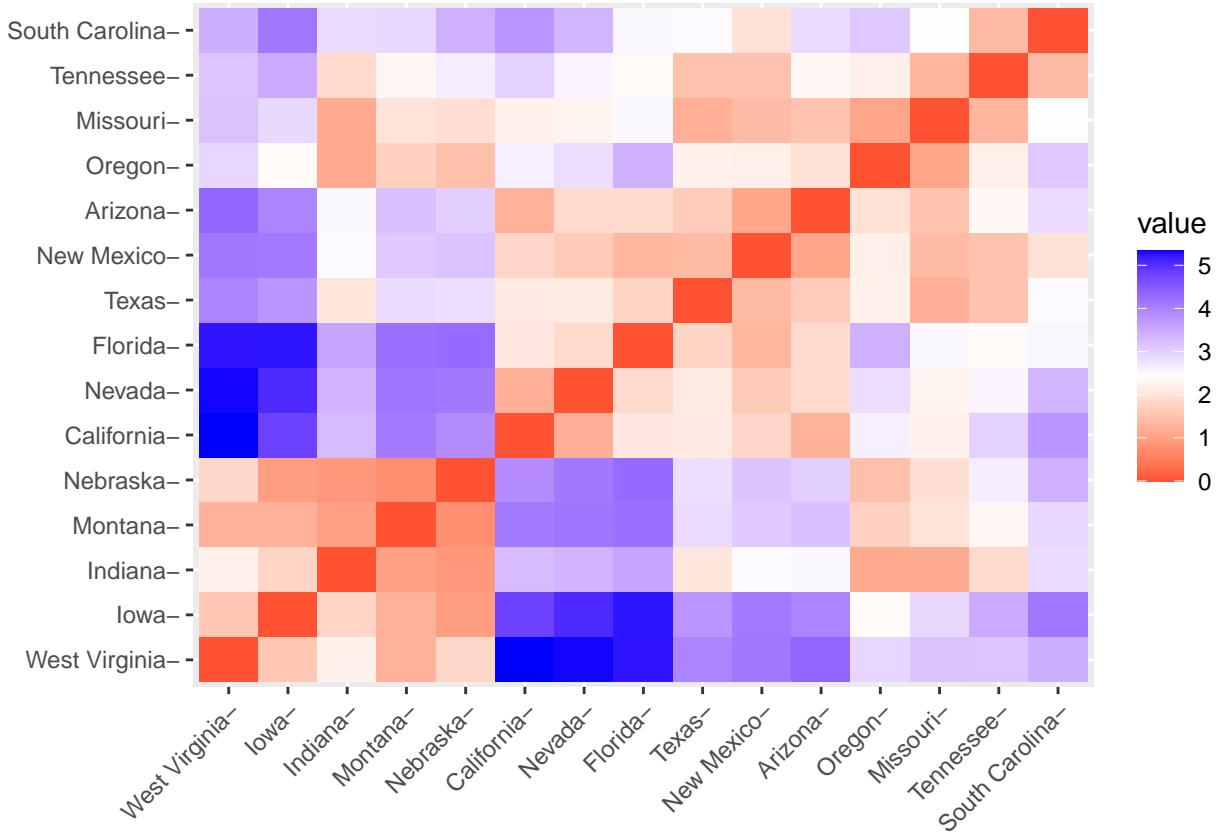
Visualizing distance matrices

- A simple solution for visualizing the distance matrices is to use the function fviz_dist() [factoextra package].
- Other specialized methods will be described later on.

```

library(factoextra)
fviz_dist(dist.eucl)

```



Partitioning Clustering

- Partitioning clustering are clustering methods used to classify observations within a data set, into multiple groups based on their similarity.
- The algorithms require the analyst to specify the number of clusters to be generated.
- This chapter describes the commonly used partitioning clustering, including:
 1. K-means clustering (MacQueen, 1967), in which, each cluster is represented by the center or means of the data points belonging to the cluster. The K-means method is sensitive to anomalous data points and outliers.
 2. K-medoids clustering or PAM (Partitioning Around Medoids, Kaufman & Rousseeuw, 1990), in which, each cluster is represented by one of the objects in the cluster. PAM is less sensitive to outliers compared to k-means.
 3. CLARA algorithm (Clustering Large Applications), which is an extension to PAM adapted for large data sets.

K-Means Clustering

- The description of the algorithm is based on:
- HARTIGAN, John A. *Clustering algorithms*. John Wiley & Sons, Inc., 1975.

Table 4.1 Nutrients in Meat, Fish, and Fowl

[*The Yearbook of Agriculture 1959* (The United States Department of Agriculture, Washington, D.C.) p. 244.] The quantity used is always 3 ounces.

		Food Energy (Calories)	Protein (Grams)	Fat (Grams)
BB	Beef, braised	340	20	28
HR	Hamburger	245	21	17
BR	Beef, roast	420	15	39
BS	Beef, steak	375	19	32
BC	Beef, canned	180	22	10
CB	Chicken, broiled	115	20	3
CC	Chicken, canned	170	25	7
BH	Beef heart	160	26	5
LL	Lamb leg, roast	265	20	20
LS	Lamb shoulder, roast	300	18	25
HS	Smoked ham	340	20	28
PR	Pork roast	340	19	29
PS	Pork simmered	355	19	30
BT	Beef tongue	205	18	14
VC	Veal cutlet	185	23	9
FB	Bluefish, baked	135	22	4
AR	Clams, raw	70	11	1
AC	Clams, canned	45	7	1
TC	Crabmeat, canned	90	14	2
HF	Haddock, fried	135	16	5
MB	Mackerel, broiled	200	19	13
MC	Mackerel, canned	155	16	9
PF	Perch, fried	195	16	11
SC	Salmon, canned	120	17	5
DC	Sardines, canned	180	22	9
UC	Tuna, canned	170	25	7
RC	Shrimp, canned	110	23	1

- The data used by the author are provided below.
- The principal nutrients in meat, fish, and fowl are listed.
- Recall that 1oz= 28.34952g.
- Estimated daily dietary allowances are: food energy (3200 cal), protein (70 g), calcium (0.8 g), and iron (10 mg).
- Table 4.2 converts the variables (with the exception of Fat) in percentage of food delivery.

Table 4.2 Nutrients in Meat, Fish, and Fowl

As a percentage of recommended daily allowances.

	Food Energy	Protein	Fat (Grams)	Calcium	Iron
Beef, braised	11	29	28	1	26
Hamburger	8	30	17	1	27
Beef, roast	13	21	39	1	20
Beef, steak	12	27	32	1	26
Beef, canned	6	31	10	2	37
Chicken, broiled	4	29	3	1	14
Chicken, canned	5	36	7	2	15
Beef, heart	5	37	5	2	59
Lamb leg, roast	8	29	20	1	26
Lamb shoulder, roast	9	26	25	1	25
Ham, smoked	11	29	28	1	25
Pork roast	11	27	29	1	25
Pork simmered	11	27	30	1	25
Beef tongue	6	26	14	1	25
Veal cutlet	6	33	9	1	27
Bluefish, baked	4	31	4	3	6
Clams, raw	2	16	1	10	60
Clams, canned	1	10	1	9	54
Crabmeat, canned	3	20	2	5	8
Haddock, fried	4	23	5	2	5
Mackerel, broiled	6	27	13	1	10
Mackerel, canned	5	23	9	20	18
Perch, fried	6	23	11	2	13
Salmon, canned	4	24	5	20	7
Sardines, canned	6	31	9	46	25
Tuna, canned	5	36	7	1	12
Shrimp, canned	3	33	1	12	26

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- For e.g., the first (BB) ligne is obtained in the following way:
- $340/3200 = 11\%$ (Food Energy).
- $20/70 = 29\%$ (Protein).
- $0.009/0.8 = 1\%$ (Calcium).

- $2.6/10 = 26\%$ (Iron).
- An argument could be made that iron is less important than calories or protein and so should be given less weight or ignored entirely.
- There are n objects and k clusters, $k \leq n$.
- Our purpose is to partition the n objects (here foods) so that objects within clusters are close and objects in different clusters are distant.
- Each cluster contains at least one object and each object belongs to only one cluster.
- There is a very large number of possible partitions.

Exercice 11

What is the number of possible partitions ?

K-Means

- The discordance between the data and a given partition is denoted by e .
- We use the technique of local optimization.
- A neighborhood of partitions is defined for each ption.
- Starting from an initial partition, search through a set of partitions at each step.
- Move from the partition to a partition in its neighborhood for which e is minim.
- If the neighborhoods are very large, it is cheaper computationally to move to the first partition discovered in the neighborhood where e is reduced from its present value.
- A number of stopping rules are possible.
 - For example, the search stops when e is not reduced by movement to the neighborhood.
 - The present partition is locally optimal in that it is the best partition in its neighborhood.
 - Consider partitions of the five ($n = 5$) beef foods {BB, BR,BS,BC,BH} into three clusters ($k = 3$).
 - Totally, there are 25 such ??.
 - A plausible neighborhood for a partition is the set of partitions obtained by transferring an object from one cluster to another.
 - For the partition (BB BR) (BS) (BC BH), the neighborhood consists of the following ten partitions:
 1. (BR) (BB BS) (BC BH)
 2. (BR) (BS) (BB BC BH)
 3. (BB) (BR BS) (BC BH)
 4. (BB) (BS) (BR BC BH)
 5. (BB BR BS) O (BC BH)
 6. (BB BR) O (BS BC BH)
 7. (BB BR BC) (BS) (BH)
 8. (BB BR) (BS BC) (BH)
 9. (BB BR BH) (BS) (BC)
 10. (BB BR) (BS BH) (BC)

K-Means Algorithm

- Let $\mathbf{x}^j \equiv (x_1^j, \dots, x_n^j)$ the vector of values for the object j , $j = 1, \dots, m$.
- The variables are assumed scaled.
- The partition has k disjoint clusters clusters C_1, \dots, C_k , which are the indices of the objects in the various clusters.
- Let m_l be the number of objects in cluster C_l .
- Each of the m objects lies in just one of the k clusters.
- Note that $\sum_{l=1}^k m_l = m$.
- The vector of means over the objects in cluster C_l is denoted by $\bar{\mathbf{x}}^l$, with

$$\bar{\mathbf{x}}^l \equiv \frac{1}{m_l} \sum_{j \in C_l} \mathbf{x}^j = (\bar{x}_1^l, \dots, \bar{x}_n^l), \quad l = 1, \dots, k,$$

where

$$\bar{x}_i^l \equiv \frac{\sum_{j \in C_l} x_i^j}{m_l}, \quad i = 1, \dots, n; \quad l = 1, \dots, k.$$

- The distance between the object j and the cluster l is $d(\mathbf{x}^j, \bar{\mathbf{x}}^l)$, where d is taken to be the Euclidian distance

$$d(\mathbf{x}^j, \bar{\mathbf{x}}^l) = \|\mathbf{x}^j - \bar{\mathbf{x}}^l\| = \left[\sum_{i=1}^n (x_i^j - \bar{x}_i^l)^2 \right]^{1/2}, \quad j = 1, \dots, m; \quad l = 1, \dots, k.$$

where $\|\cdot\|$ is the Euclidian norm.

- The error of the partition is measured taken to bye

$$e = \sum_{l=1}^k \sum_{j \in C_l} d^2(\mathbf{x}^j, \bar{\mathbf{x}}^l) = \sum_{l=1}^k \sum_{j \in C_l} \|\mathbf{x}^j - \bar{\mathbf{x}}^l\|^2.$$

- Alternatively, we have

$$e = \sum_{j=1}^m d^2(\mathbf{x}^j, \bar{\mathbf{x}}^{l(j)}) = \sum_{j=1}^m \sum_{i=1}^n \|\mathbf{x}^j - \bar{\mathbf{x}}^{l(j)}\|^2,$$

where $l(j)$ is the index of the cluster of object j .

- The general procedure is to search for a partition with a small error e by moving cases from one cluster to another.
- The search ends when no such movement reduces e .
- **STEP 1.** Assume initial clusters. Compute the cluster means $\bar{\mathbf{x}}^l$ and the initial error e .
- **STEP 2.** For the first Object, compute for every cluster l

$$\Delta e = \frac{m_l d^2(\mathbf{x}^1, \bar{\mathbf{x}}^l)}{m_l + 1} - \frac{m_{l(1)} d^2(\mathbf{x}^1, \bar{\mathbf{x}}^{l(1)})}{m_{l(1)} - 1}, \quad l = 1, \dots, k, \quad l \neq l(1).$$

- It corresponds to the error variation in transferring the first object from cluster $l(1)$ to which it belongs to cluster l .
- If the minimum of this quantity over all $l \neq l(1)$ is negative, transfer the first case from cluster $l(1)$ to this minimal l .
- Adjust the cluster means of $l(1)$ and the minimal l and add the error variation (which is negative) to e .

- **STEP 3.** Repeat STEP 2 for each object j such that $2 \leq j \leq m$.
- **STEP 4.** If no movement of an object from one cluster to another occurs for any case, stop. Otherwise, return to STEP 2.

Exercice 12

Prove that the error variation is indeed given by:

$$\Delta e = \frac{m_l d^2(\mathbf{x}^1, \bar{\mathbf{x}}^l)}{m_l + 1} - \frac{m_{l(1)} d^2(\mathbf{x}^1, \bar{\mathbf{x}}^{l(1)})}{m_{l(1)} - 1}, \quad l = 1, \dots, k, \quad l \neq l(1).$$

K-MEANS APPLIED TO FOOD NUTRIENT DATA * Only the first eight foods will be considered. * Only three variables, food energy, protein, and calcium as a percentage of recommended daily allowances are used. * The eight foods ($m = 8$) are partitioned into three clusters ($k = 3$).

Exercice 13

- Explain in details the k-means algorithm based on the following pages of Hartigan (1975).

Table 4.3 Application of K-Means Algorithm to Food Data

	Energy	Protein	Calcium	
BB	11	29	1	
HR	8	30	1	
BR	13	21	1	
BS	12	27	1	
BC	6	31	2	
CB	4	29	1	
CC	5	36	1	
BH	5	37	2	

INITIAL CLUSTER MEANS			e = 154.9
	Energy	Protein	Calcium
1. BR, CB	8.5	25	1
2. HR, BS	10	28.5	1
3. BB, BC, CC, BH	6.75	33.25	1.5

FIRST CHANGE			e = 108.2
	Energy	Protein	Calcium
1. HR, CB	8.5	25	1
2. HR, BS, BB	10.33	28.67	1
3. BC, CC, BH	5.33	34.67	1.67

SECOND CHANGE			e = 61.4
	Energy	Protein	Calcium
1. BR	13	21	1
2. HR, BS, BB	10.33	28.67	1
3. BC, CC, BH, CB	5	33.25	1.5

STEP 1. A quick initial clustering, which often works well, is based on the case sums. Suppose these are denoted by $\text{SUM}(I)$, having minimum value MIN and maximum value MAX . To obtain K initial clusters, set case I into the J th cluster, where J is the integral part of $K[\text{SUM}(I) - \text{MIN}]/(\text{MAX} - \text{MIN}) + 1$. Here the case sums are 41, 39, 35, 40, 41, 34, 42, and 44. The corresponding clusters are 3, 2, 1, 2, 3, 1, 3, and 3. Thus the initial partition is (BR CB) (HR BS) (BB BC CC BH). The values of $B(L, J)$ ($1 \leq L \leq 3, 1 \leq J \leq 3$) are next computed. For example, $B(1, 1)$, the mean of cases in the first cluster for the first variable, equals $(13 + 4)/2 = 8.5$. (See Table 4.3 for more.) The error for the initial partition is the sum of squared distances of cases from their cluster means,

$$\begin{aligned}
e[P(8, 3)] &= (11 - 6.75)^2 + (29 - 33.25)^2 + (1 - 1.5)^2 + (8 - 10)^2 \\
&\quad + (30 - 28.5)^2 + (1 - 1)^2 + (13 - 8.5)^2 + (21 - 25)^2 \\
&\quad + (1 - 1)^2 + (12 - 10)^2 + (27 - 28.5)^2 + (1 - 1)^2 \\
&\quad + (6 - 6.75)^2 + (31 - 33.25)^2 + (2 - 1.5)^2 + (8.5 - 4)^2 \\
&\quad + (29 - 25)^2 + (1 - 1)^2 + (5 - 6.75)^2 + (36 - 33.25)^2 \\
&\quad + (1 - 1.5)^2 + (5 - 6.75)^2 + (37 - 33.25)^2 + (2 - 1.5)^2 \\
&= 154.9.
\end{aligned}$$

The first three squares are the squared distance of BB from its cluster mean (6.75, 33.25, 1.5), and so on.

STEP 2. For the first case, the distances to clusters are

$$D(1, 1)^2 = (11 - 8.5)^2 + (29 - 25)^2 + (1 - 1)^2 = 22.25,$$

$$D(1, 2)^2 = 1.25,$$

$$D(1, 3)^2 = 36.4.$$

The increase in error in transferring the first case to cluster 1 is $2 \times 22.5/3 - 4 \times 36.4/3$, and that to cluster 2 is $2 \times 1.25/3 - 4 \times 36.4/3$. The cluster that is best for the first case is thus the second cluster, and the error reduction is 47.7. The new value of $e[P(8, 3)]$ is thus $154.9 - 47.7 = 108.2$.

It is necessary to update the means of clusters 2 and 3, since cluster 2 has gained the first case and cluster 3 has lost it. For example,

$$B(2, 1) = (11 + 2 \times 10)/3 = 10.33,$$

$$B(2, 2) = (29 + 2 \times 28.5)/3 = 28.67,$$

$$B(2, 3) = (1 + 2 \times 1)/3 = 1.00.$$

STEP 3. Repeating Step 1 on all cases, for case 2, cluster 2 is far closer than any other, and case 2 remains in cluster 2, with no change taking place. Continuing, no change takes place until case 6, which moves to cluster 3. No further changes occur in this pass.

STEP 4. Since some changes occurred in the last pass, another pass is necessary through all cases. No changes occur on this pass and the algorithm stops with the final cluster (BR) (HR BS BB) (BC CC BH CB). These clusters are characterized by the variables as follows: The first cluster is high in energy and low in protein, the second is high in energy and protein, and the third is low in energy and high in protein. Calcium hardly matters. The complete data set is partitioned in Table 4.4.

```
#library("cluster.datasets")
#write.csv(rda.meat.fish.fowl.1959,"Hartigandat%a1.csv")
df<-read.csv("Hartigandata1.csv")
print(df)
```

##	X		name	energy	protein	fat	calcium	iron
## 1	1	Braised beef		11	29	28	1	26
## 2	2	Hamburger		8	30	17	1	27
## 3	3	Roast beef		18	21	39	1	20
## 4	4	Beefsteak		12	27	32	1	26
## 5	5	Canned beef		6	31	10	2	37
## 6	6	Broiled chicken		8	29	3	1	14
## 7	7	Canned chicken		5	36	7	2	15
## 8	8	Beef heart		5	37	5	2	59
## 9	9	Roast lamb leg		8	29	20	1	26

```

## 10 10 Roast lamb shoulder      9    26 25      1 23
## 11 11          Smoked ham   11    29 28      1 25
## 12 12          Pork roast   11    27 29      1 25
## 13 13 Pork simmered          11    27 30      1 24
## 14 14 Beef tongue           6    26 14      1 25
## 15 15 Veal cutlet           6    33  9      1 27
## 16 16 Baked bluefish        4    31  4      3  6
## 17 17 Raw clams            2    16  1     10 60
## 18 18 Canned clams          1    10  1      9 54
## 19 19 Canned crabmeat       3    20  2      5  8
## 20 20 Fried haddock         4    23  5      2  5
## 21 21 Broiled mackerel      6    27 13      1 10
## 22 22 Canned mackerel       5    23  9     20 18
## 23 23 Fried perch           6    23 11      2 13
## 24 24 Canned salmon          4    24  5     20  7
## 25 25 Canned sardines        6    31  9     46 25
## 26 26 Canned tuna            5    36  7      1 12
## 27 27 Canned shrimp          3    33  1     12 26

```

```
df<-df[1:8,c(3,4,6)]
```

```
df
```

```

##   energy protein calcium
## 1     11      29      1
## 2      8      30      1
## 3     18      21      1
## 4     12      27      1
## 5      6      31      2
## 6      8      29      1
## 7      5      36      2
## 8      5      37      2

```

```

# The data contain some errors
df[3,1]<-13 # Error in line 3
df[6,1]<-4 # Error at line 6
df[7,3]<-1 # Error at line 7
df

```

```

##   energy protein calcium
## 1     11      29      1
## 2      8      30      1
## 3     13      21      1
## 4     12      27      1
## 5      6      31      2
## 6      4      29      1
## 7      5      36      1
## 8      5      37      2

```

```
rownames(df)<-c("BB","HR","BR","BS","BC","CB","CC","BH")
df
```

```

##   energy protein calcium
## BB     11      29      1
## HR      8      30      1
## BR     13      21      1
## BS     12      27      1
## BC      6      31      2

```

```

## CB      4      29      1
## CC      5      36      1
## BH      5      37      2
colnames(df)<-c("Energy","Protein","Calcium")
df

##      Energy Protein Calcium
## BB      11      29      1
## HR      8       30      1
## BR     13      21      1
## BS     12      27      1
## BC      6       31      2
## CB      4       29      1
## CC      5       36      1
## BH      5       37      2
kmeans(df,3)

## K-means clustering with 3 clusters of sizes 4, 1, 3
##
## Cluster means:
##      Energy Protein Calcium
## 1  5.00000 33.25000    1.5
## 2 13.00000 21.00000    1.0
## 3 10.33333 28.66667    1.0
##
## Clustering vector:
## BB HR BR BS BC CB CC BH
## 3 3 2 3 1 1 1 1
##
## Within cluster sum of squares by cluster:
## [1] 47.75000 0.00000 13.33333
##   (between_SS / total_SS =  77.2 %)
##
## Available components:
##
## [1] "cluster"      "centers"       "totss"        "withinss"      "tot.withinss"
## [6] "betweenss"    "size"          "iter"          "ifault"

```

Exercice 14

- Produce a heatmap for the data by rearranging the lines according to the found clusters (k=3).

Exercice 15

- Produce the Table 4.4 of Hartigan (1975).

Table 4.4

Clusters and cluster means from K -means algorithm applied to food data on energy, protein, and calcium expressed as percentages of daily requirements. The two clusters obtained by splitting cluster 3 are denoted by 31 and 32.

PARTITION	CLUSTERS	ENERGY	PROTEIN	CALCIUM
1	1 : BB HR BR BS BC CB CC BH LL LL HS PR PS BT VC FB AR AC TC HF MB MC PF SC DS UC RC	6.5	27.1	5.5
2	11 : BB HR BR BS BC CB CC BH LL LL HS PR PS BT VC FB AR AC TC HF MB PF UC RC 12 : MC SC DS	6.7	27.3	2.6
3	12 111 : BB HR BR BS BC CB CC BH LL LL HS PR PS BT VC FB HF MB PF UC RC 112 : AR AC TC	7.4	29.0	1.8
4	112 2 : BB HR BR BS BC CB CC BH LL LL HS PR PS BT VC FB HF MB PF UC 3 : MC SC RC 4 : DS	7.5	28.8	1.3
5	112, 3, 4 21 : HR BC CB CC BH VC FB UC 22 : BB BR BS LL LL HS PR PS BT HF MB PF	5.6	32.4	1.5
6	112, 21, 22, 4 31 : MC SC 32 : RC	9.1	25.8	1.2
7	112, 21, 31, 32, 4 221 : BB BS LL LL HS PR PS BT HF MB PF 222 : BR	4.3	23.6	19.8
8	222, 31, 32, 4 5 : BC CC BH VC FB UC 6 : AR AC 7 : BB HR BS LL LL HS PR PS MB 8 : CB BT TC HF PF	3.4	32.9	12.3
9	222, 31, 32, 4, 5, 6, 9 : BB HR BS LL LL HS PR PS 10 : CB BT HF MB PF 11 : TC	8.7 13.1	26.5 21.4	1.2 .9
		10.0	27.9	1.1
		5.3	25.4	1.2
		2.8	20.0	4.8