

# MF731 Homework 3

## 1. The "Exponential Premium Principal" Risk Measure.

(a) For  $\rho_2(L) = \frac{1}{2} \log(E[e^{2L}])$ , obviously it's a monotonic function.

Then for cash additivity, we have:

$$\rho_2(L - c) = \frac{1}{2} \log(E[e^{2(L-c)}]) = \frac{1}{2} \log(E[e^{2L}]) - c = \rho_2(L) - c$$

Therefore, for the convexity, for any  $\lambda \in (0, 1)$ , we'll have

$$\rho_2(\lambda L_1 + (1-\lambda)L_2) = \frac{1}{2} \log(E[e^{2(\lambda L_1 + (1-\lambda)L_2)}])$$

Hölder's inequality told us that: for  $\frac{1}{p} + \frac{1}{q} = 1 \Rightarrow E[X^p]^{\frac{1}{p}} E[Y^q]^{\frac{1}{q}} \geq E[XY]$

Here we set  $p = \frac{1}{\lambda}$ ,  $q = \frac{1}{1-\lambda}$ ,

$$\begin{aligned} \rho_2(\lambda L_1 + (1-\lambda)L_2) &\leq \frac{1}{2} \log(E[e^{2L_1}]^\lambda \cdot E[e^{2L_2}]^{1-\lambda}) \\ &= \lambda \rho_2(L_1) + (1-\lambda) \rho_2(L_2) \end{aligned}$$

So  $\rho_2$  is a convex risk measure.

(b) For  $L \sim N(0, 1)$

$$\begin{aligned} \rho_2(\lambda L) &= \frac{1}{2} \log(E[e^{2\lambda L}]) \\ &= \frac{1}{2} \log(e^{0 + \frac{1}{2}(2\lambda)^2}) \\ &= \frac{1}{2} (\frac{1}{2} 2\lambda^2) = \frac{1}{2} \lambda^2 \end{aligned}$$

obviously, it's not homogenous.

2. For the detail, see the attached code file.

3. (a) To prove same distribution, we need to show the characteristic function.

$$\text{For } a^T Z \Rightarrow E[e^{i a^T Z}] = E[E[e^{i a^T W a} | W]] = E[e^{-\frac{1}{2} a^T W a}]$$

$$\text{For } |a| Z^{(1)} \Rightarrow E[E[e^{i |a| \sqrt{W} Z^{(1)}} | W]] = E[e^{-\frac{1}{2} |a|^2 W}]$$

Since  $a^T Z$  and  $|a| Z^{(1)}$  have same characteristic function, they have same distribution.

$\therefore \tilde{Z}$  is spherical

(b) (i)  $L^{\text{lin}} = -\theta^T X = -\theta^T \mu - \theta^T A Z$ ,  $X$  is elliptically distributed  $\rightarrow Z$  is spherical.

$\therefore a^T Z$  has same distribution with  $|a| Z^{(1)}$ , set  $-A^T \theta = a$

$$\therefore |a| = \sqrt{a^T a} = \sqrt{a^T A A^T a} = \sqrt{\theta^T A A^T \theta} = \sqrt{\theta^T \Sigma \theta}$$

$\therefore L^{\text{lin}}$  has the same distribution of  $-\theta^T \mu + \sqrt{\theta^T \Sigma \theta} Z^{(1)}$

hence for any cash-additive, positively homogenous risk measure  $\rho$ .

$$\rho(L^{\text{lin}}) = -\theta^T \mu + \sqrt{\theta^T \Sigma \theta} \cdot \rho(Z^{(1)})$$

(ii) from (i), it's obvious that  $R_{\mu}^{\text{lin}^{(1)}} = -\mu^{(1)} + \frac{\Sigma \theta^T}{\sqrt{\theta^T \Sigma \theta}} \cdot \rho(Z^{(1)})$

$$\text{Thus, } R_{\mu}^{\text{lin}} = \int_0^\infty x \left( \frac{-\theta^T \mu + \frac{\theta^T \Sigma \theta}{\sqrt{\theta^T \Sigma \theta}} \rho(Z^{(1)})}{-\theta^T \mu + \sqrt{\theta^T \Sigma \theta} \rho(Z^2)} \right)$$

Therefore, we can make the same conclusion connecting  $R_C^1$  to the percentage contribution to variance, as we did when  $X \sim N(\mu, \Sigma)$