PLSC 503 – Spring 2020 Multiplicative Interactions, I

March 5, 2020

Primitives

$$Y = X\beta + u$$

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + u_i$$

means:

$$E(Y|D_1=1,D_2=0)-E(Y|D_1=0,D_2=0)=E(Y|D_1=1,D_2=1)-E(Y|D_1=0,D_2=1)[\equiv\beta_1]$$

$$\textit{E}(Y|D_1=0,D_2=1) - \textit{E}(Y|D_1=0,D_2=0) = \textit{E}(Y|D_1=1,D_2=1) - \textit{E}(Y|D_1=1,D_2=0) [\equiv \beta_2].$$

Interaction Effects

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} X_{2i} + u_i$$

$$E(Y_i) = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} X_{2i}$$

$$= \beta_0 + \beta_2 X_{2i} + (\beta_1 + \beta_3 X_{2i}) X_{1i}$$

$$= \beta_0 + \beta_2 X_{2i} + \psi_1 X_{1i}$$

where $\psi_1 = \beta_1 + \beta_3 X_{2i}$. This means:

$$\frac{\partial \mathsf{E}(Y_i)}{\partial X_1} = \beta_1 + \beta_3 X_{2i}.$$

Interaction Effects

Similarly:

$$E(Y_i) = \beta_0 + \beta_1 X_{1i} + (\beta_2 + \beta_3 X_{1i}) X_{2i}$$

= $\beta_0 + \beta_1 X_{1i} + \psi_2 X_{2i}$

which implies:

$$\frac{\partial \mathsf{E}(Y_i)}{\partial X_2} = \beta_2 + \beta_3 X_{1i}.$$

"Direct Effects"

If $X_2 = 0$, then:

$$E(Y_i) = \beta_0 + \beta_1 X_{1i} + \beta_2(0) + \beta_3 X_{1i}(0)$$

= $\beta_0 + \beta_1 X_{1i}$.

Similarly, for $X_1 = 0$:

$$E(Y_i) = \beta_0 + \beta_1(0) + \beta_2 X_{2i} + \beta_3(0) X_{2i} = \beta_0 + \beta_2 X_{2i}$$

Key Point

In most instances, the quantities we care about are not β_1 and β_2 , but rather ψ_1 and ψ_2 .

Inference

Point estimates:

$$\hat{\psi}_1 = \hat{\beta}_1 + \hat{\beta}_3 X_2$$

and

$$\hat{\psi}_2 = \hat{\beta}_2 + \hat{\beta}_3 X_1.$$

For variance, recall that:

$$Var(a + bZ) = Var(a) + Z^{2}Var(b) + 2ZCov(a, b)$$

Inference

Means that:

$$\widehat{\mathsf{Var}(\hat{\psi}_1)} = \widehat{\mathsf{Var}(\hat{\beta}_1)} + X_2^2 \widehat{\mathsf{Var}(\hat{\beta}_3)} + 2X_2 \widehat{\mathsf{Cov}(\hat{\beta}_1, \hat{\beta}_3)}.$$

and

$$\widehat{\mathsf{Var}(\hat{\psi}_2)} = \widehat{\mathsf{Var}(\hat{\beta}_2)} + X_1^2 \widehat{\mathsf{Var}(\hat{\beta}_3)} + 2X_1 \widehat{\mathsf{Cov}(\hat{\beta}_2, \hat{\beta}_3)}.$$

Types of Interactions: Dichotomous Xs

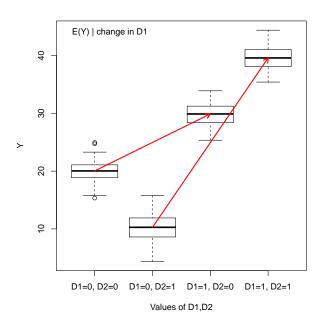
For

$$Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 D_{1i} D_{2i} + u_i$$

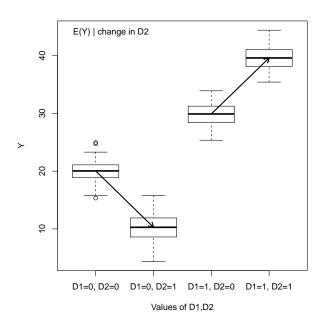
we have:

$$\begin{split} &\mathsf{E}(Y|D_1=0,D_2=0) &= \beta_0 \\ &\mathsf{E}(Y|D_1=1,D_2=0) &= \beta_0+\beta_1 \\ &\mathsf{E}(Y|D_1=0,D_2=1) &= \beta_0+\beta_2 \\ &\mathsf{E}(Y|D_1=1,D_2=1) &= \beta_0+\beta_1+\beta_2+\beta_3 \end{split}$$

Values of E(Y) for Changes in D_1



Values of E(Y) for Changes in D_2



Dichotomous and Continuous Xs

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 X_i D_i + u_i$$

gives:

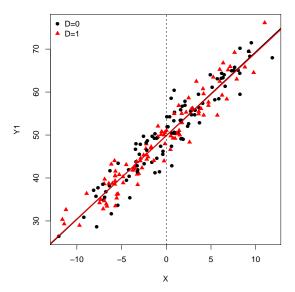
$$E(Y|X, D = 0) = \beta_0 + \beta_1 X$$

 $E(Y|X, D = 1) = (\beta_0 + \beta_2) + (\beta_1 + \beta_3) X$

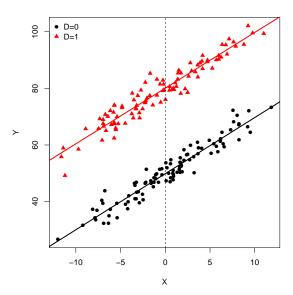
Four possibilities:

- $\beta_2 = \beta_3 = 0$
- $\beta_2 \neq 0$ and $\beta_3 = 0$
- $\beta_2 = 0$ and $\beta_3 \neq 0$
- $\beta_2 \neq 0$ and $\beta_3 \neq 0$

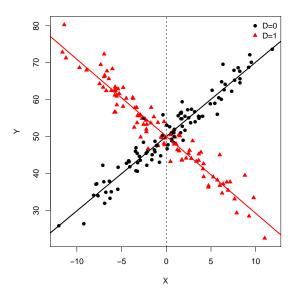
Scatterplot and Regression Lines of Y on X for D=0 and D=1: No Slope or Intercept Differences ($\beta_2=\beta_3=0$)



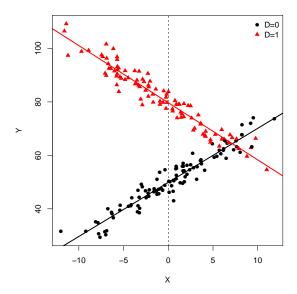
Scatterplot and Regression Lines of Y on X for D=0 and D=1: Intercept Shift $(\beta_2 \neq 0, \, \beta_3=0)$



Scatterplot and Regression Lines of Y on X for D=0 and D=1: Slope Change $(\beta_2=0,\,\beta_3\neq 0)$



Scatterplot and Regression Lines of Y on X for D=0 and D=1: Slope and Intercept Change $(\beta_2 \neq 0, \beta_3 \neq 0)$



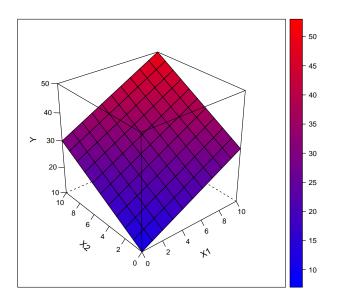
Two Continuous Xs

$$Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 X_{1i} X_{2i} + u_i.$$

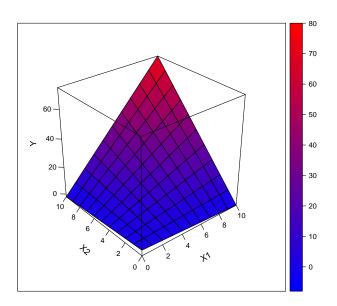
Implies

$$\beta_3 = 0 \rightarrow \frac{\partial E(Y)}{\partial X_1} = \beta_1 \,\forall \, X_2 \text{ and } \frac{\partial E(Y)}{\partial X_2} = \beta_2 \,\forall \, X_1$$

Two Continuous Variables: No Interactive Effects



Two Continuous Variables: Interaction Present



Quadratic, Cubic, and Other Polynomial Effects

$$Y_{i} = \beta_{0} + \beta_{1}X_{i} + \beta_{2}X_{i}^{2} + \beta_{3}X_{i}^{3} + \dots + \beta_{j}X_{i}^{j} + u_{i}.$$

In general:

$$\frac{\partial \mathsf{E}(Y)}{\partial X} = \beta_1 + 2\beta_2 X + 3\beta_3 X^2 + \dots + j\beta_j X^{j-1}$$

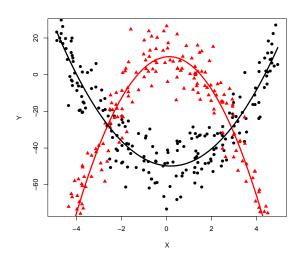
Quadratic case (j = 2):

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 X_i^2 + u_i.$$

implies

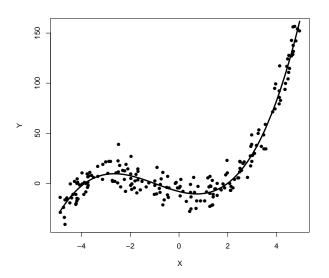
$$\frac{\partial \mathsf{E}(Y)}{\partial X} = \beta_1 + 2\beta_2 X$$

Two Quadratic Relationships



Note: Red line is $Y_i=10+1X_i-5X_i^2+u_i$; black line is $Y_i=-50-1X_i+3X_i^2+u_i$.

Example of a Cubic Relationship



Note: Solid line is $Y_i = -1 + 1X_i - 8X_i^2 + 5X_i^3 + u_i$.

Higher-Order Interactive Models

Three-way interaction:

$$Y_{i} = \beta_{0} + \beta_{1}X_{1i} + \beta_{2}X_{2i} + \beta_{3}X_{3i} + \beta_{4}X_{1i}X_{2i} + \beta_{4}X_{1i}X_{3i}u_{i} + \beta_{5}X_{2i}X_{3i} + \beta_{6}X_{1i}X_{2i}X_{3i} + u_{i}$$

Special case of dichotomous X_1 , X_2 :

$$Y_{i} = \beta_{0} + \beta_{1}X_{i} + \beta_{2}D_{1i} + \beta_{3}D_{2i} + \beta_{4}X_{i}D_{1i} + \beta_{4}X_{i}D_{2i}u_{i} + \beta_{5}D_{1i}D_{2i} + \beta_{6}X_{i}D_{1i}D_{2i} + u_{i}$$

Three-Way Interaction: Two Dummy and One Continuous Covariates

