RTSM Project on Time Series Data.

```
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```

plot(apple price)

Stock Considered: Apple.inc or AAPL (herein after referred as "apple_stock") on daily basis from 2020-01-01 to 2023-12-31

```
Required packages: quantmod, car ,forecast, tseries, FinTS, rugarch, utf8, ggpl ot2

packages = c('quantmod','car','forecast','tseries','FinTS', 'rugarch','utf8','ggpl ot2')

install.packages(packages, dependencies = TRUE)

lapply(packages, require, character.only = TRUE)

Loading the dataset of Apple inc. from Yahoo finance:

getSymbols(Symbols = 'AAPL',

src = 'yahoo',

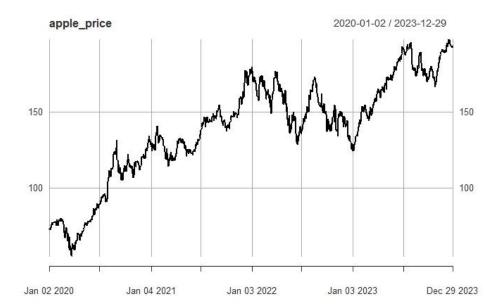
from = as.Date('2020-01-01'),

to = as.Date('2023-12-31'),

periodicity = 'daily')

apple_price = na.omit(AAPL$AAPL.Adjusted) # Adjusted Closing Price

class(apple_price) # xts (Time-Series) Object
```



To analyse the stock price of Apple form 2020-01-01 to 2023-12-29, extracted daily adjusted closing prices of Apple.inc from Yahoo finance. Performed ADF test on stock prices and returns.

Augmented Dickey-Fuller (ADF) Test for Stationarity with Stock Data:

```
adf test aapl = adf.test(apple price); adf test aapl
```

Augmented Dickey-Fuller Test

data: apple_price

Dickey-Fuller = -2.4768, Lag order = 10, p-value = 0.3765

alternative hypothesis: stationary

Objective: To conduct an Augmented Dickey-Fuller (ADF) test for stationarity on the daily Stock Price.

Analysis: Performed the ADF test using the 'adf.test' function and obtained results.

Result: The Augmented Dickey-Fuller test for stationarity on stock price yields the following results: -

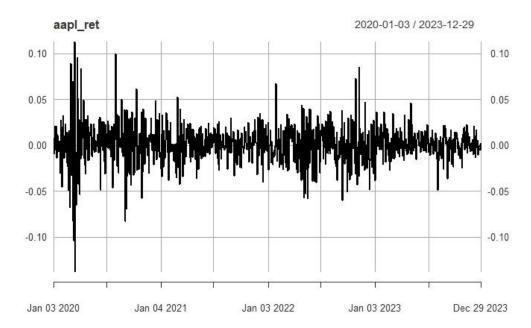
Dickey-Fuller = -2.4768, Lag order = 10, p-value = 0.3765 alternative hypothesis: stationary

Implication: The ADF test suggests that the stock price is non stationary as with a p value of 0.3765, we fail to reject the null hypothesis of non stationary.

Since the stock prices are non stationary, we will test the return on stocks.

Augmented Dickey-Fuller (ADF) Test for Stationarity with Return on stocks:

```
aapl_ret = na.omit(diff(log(apple_price)))
adf_test_aapl_ret = adf.test(aapl_ret); adf_test_aapl_ret
plot(aapl_ret)
```



Augmented Dickey-Fuller Test

```
data: aapl_ret
Dickey-Fuller = -9.2926, Lag order = 10, p-value = 0.01
alternative hypothesis: stationary
```

Objective: To conduct an Augmented Dickey-Fuller (ADF) test for stationarity on the daily return on stock.

Analysis: Performed the ADF test using the 'adf.test' function and obtained results.

Result: The Augmented Dickey-Fuller test for stationarity on return on stock yields the following results: -

Dickey-Fuller = -9.2926, Lag order = 10, p-value = 0.01 alternative hypothesis: stationary

Implication: The ADF test suggests that the return on stock is **stationary** as with a p value of 0.01, we can reject the null hypothesis of non stationary.

Ljung-Box Test for Autocorrelation - Stock Data (H0: No Autocorrelation):

```
lb test aapl ret = Box.test(aapl ret); lb test aapl ret
```

Box-Pierce test

data: $aapl_ret$ X-squared = 17.547, df = 1, p-value = 2.803e-05

Objective: To perform a Ljung-Box test for autocorrelation on the stock returns.

Analysis: Conducted the Ljung-Box test using the 'Box.test' function and obtained results.

Result: The Ljung-Box test for autocorrelation on stock returns yields the following results: X-squared = 17.547, df = 1, p-value = 2.803e-05

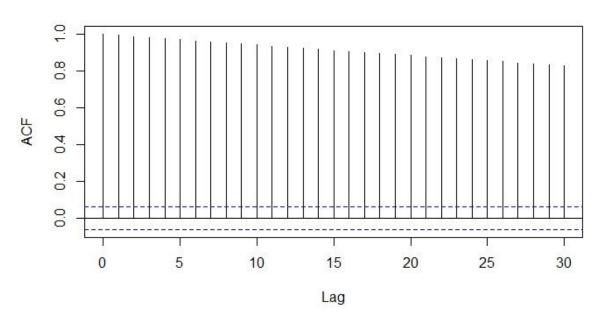
Implication: The Ljung-Box test indicates significant autocorrelation in the stock returns. The small p- value (< 2.803e-05) suggests evidence against the null hypothesis of no autocorrelation.

Action Step: Given the presence of autocorrelation, it may be advisable to consider an autoA RIMA model for time series forecasting. AutoARIMA can help in automatically selecting an appropriate ARIMA model with differencing to account for the observed autocorrelation.

Autocorrelation Function (ACF) | Partial Autocorrelation Function (PACF)

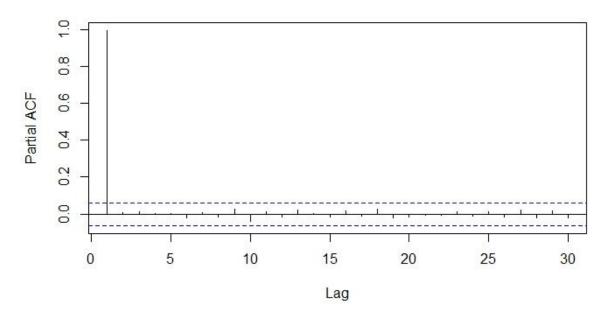
acf(apple_price)

Series apple price



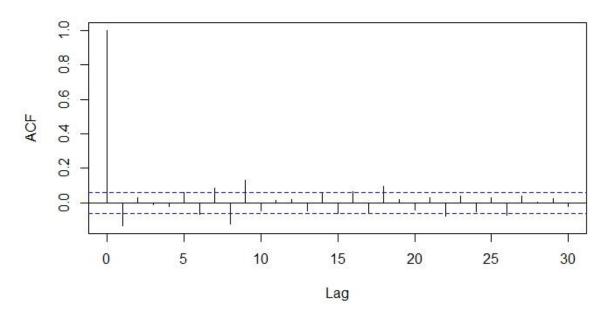
pacf(apple_price)

Series apple_price



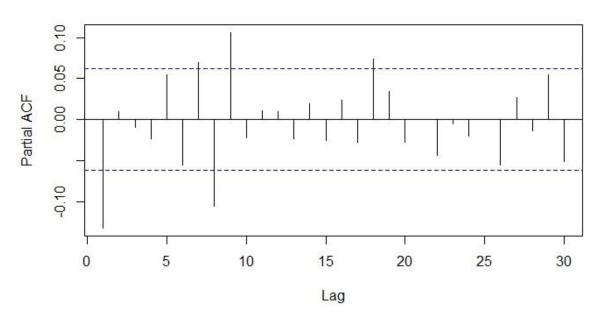
acf(aapl_ret)

Series aapl_ret



pacf(aapl_ret)

Series aapl_ret



Auto ARIMA on Stock Return and Stock Price:

```
arma_pq_aapl_ret = auto.arima(aapl_ret); arma_pq_aapl_ret
```

arma_pq_aapl = auto.arima(apple_price); arma_pq_aapl

Objective: To perform autoARIMA modeling on the stock returns and stock price.

Analysis: Used the 'auto.arima' function to automatically select the ARIMA model for both r eturns and prices.

Results:

For stock returns ('stock_ret'): The autoARIMA model suggests an ARIMA(0,0,1) with non zero mean. Coefficients: MA: ma1 - sigma^2 = 0.0004396: log likelihood = 2459.06 AIC=-4912.11 AICC=-4912.09 BIC=-4897.37

For Adjusted Closing Prices ('stock_price'): The autoARIMA model suggests an ARIMA(0, 1,0) with drift. Coefficients: drift- sigma^2 = 6.827: log likelihood = -2390.75 AIC=4785.5 AICC=4785.51 BIC=4795.33

```
y(t)=c+(-0.4205)*y(t-1)-(-0.4677)*e(t-1)-(-0.3902)*e(t-2)-(-0.1162)*e(t-3)-(-0.0908)*e(t-4)+0.0878*e(t-5)+e(t)
```

Implication: The autoARIMA models provide a statistical framework to capture the underlying patterns in both stock returns and stock price. These models can be used for forecasting fut ure values, and the AIC, AICc, and BIC values help in model comparison.

Note: The log likelihood is positive, indicating a good fit of the model to the data. The information criteria (AIC, AICc, BIC) are relatively low, suggesting a good balance between model fit and complexity.

Ljung-Box Test for Autocorrelation - Model Residuals(H0: No Autocorrelation)

lb_test_arma_pq_aapl_ret = Box.test(arma_pq_aapl_ret\$residuals); lb_test_arma_pq_aapl_ret

```
Box-Pierce test
```

```
data: arma_pq_aapl_ret$residuals
X-squared = 0.0096919, df = 1, p-value = 0.9216
```

Objective: To perform a Ljung-Box test for autocorrelation on the arima residuals.

Analysis: Conducted the Ljung-Box test using the 'Box.test' function and obtained results.

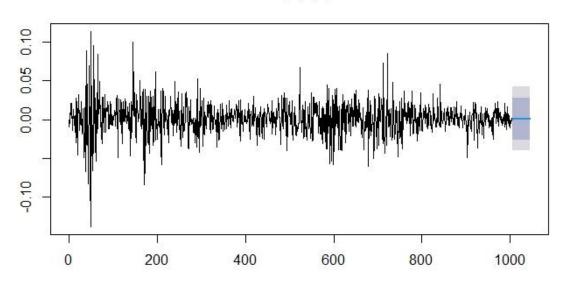
Result: The Ljung-Box test for autocorrelation on arima residual yields the following results: X-squared = 0.0096919, df = 1, p-value = 0.9216

It indicates no significant autocorrelation in the residuals of the ARIMA(0,0,1) model residual. The high p-value (0.9216) suggests that there is no evidence against the null hypothesis of no autocorrelation.

Forecasting with ARIMA Models

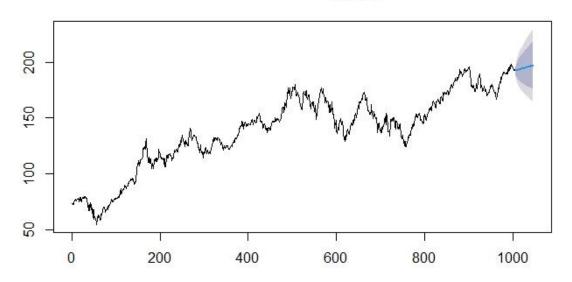
 $aapl_ret_fpq = forecast(arma_pq_aapl_ret, h = 40)$ $plot(aapl_ret_fpq)$

Forecasts from ARIMA(0,0,1) with non-zero mean



aapl_fpq = forecast(arma_pq_aapl, h = 40)
plot(aapl_fpq)

Forecasts from ARIMA(0,1,0) with drift



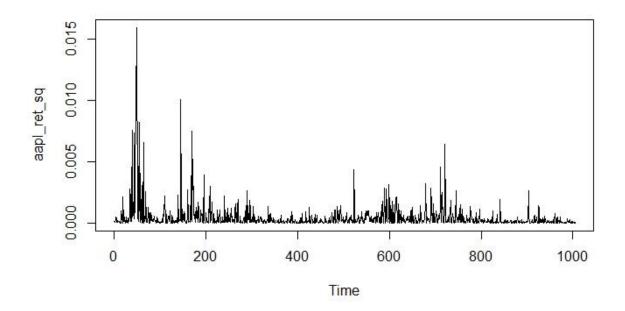
Objective: To fit an ARIMA(0,0,1) model to the stock returns & price and generate forecasts. Analysis: Used the 'arima' function to fit the ARIMA model and the 'forecast' function to generate forecasts.

Results: The plot displays the original time series of stock returns along with the forecast values.

Implication: The ARIMA(0,0,1) model is fitted to the historical stock returns, providing insights into the underlying patterns. The generated forecast can be used for future predictions, and the plot visually represents the model's performance.

Test for Volatility Clustering or Heteroskedasticity: Box Test:

aapl_ret_sq = arma_pq_aapl_ret\$residuals^2 # Return Variance (Since Mean Returns is appr
ox. 0)
plot(aapl_ret_sq)



aapl_ret_sq_box_test = Box.test(aapl_ret_sq, lag = 10) # H0: Return Variance Series is Not S erially Correlated aapl_ret_sq_box_test # Inference : Return Variance Series is Heteroskedastic (Has Volatility Clustering)

Box-Pierce test

data: $aapl_ret_sq$ X-squared = 507.38, df = 10, p-value < 2.2e-16

Test for Volatility Clustering or Heteroskedasticity: ARCH Test:

```
aapl_ret_arch_test = ArchTest(arma_pq_aapl_ret$residuals^2, lags = 10) # H0: No ARCH Ef
fects
aapl_ret_arch_test # Inference : Return Series is Heteroskedastic (Has Volatility Clustering)

ARCH LM-test; Null hypothesis: no ARCH effects
data: arma_pq_aapl_ret$residuals^2
Chi-squared = 213.05, df = 10, p-value < 2.2e-16</pre>
```

Objective: To test for volatility clustering or heteroskedasticity in the arima residuals.

Analysis: Conducted Box test and ARCH test on the squared residuals to assess the presence of volatility clustering.

Results:

1. Box Test for Volatility Clustering:

X-squared statistic: 507.38 Degrees of freedom: 10 p-value: < 2.2e-16

Thus, the test indicates significant evidence against the null hypothesis, suggesting that the return variance series exhibits volatility clustering or heteroskedasticity

2. ARCH Test for Volatility Clustering:

Chi-squared statistic: 213.05 Degrees of freedom: 10 p-value: < 2.2e-16

Thus, The ARCH test provides significant evidence against the null hypothesis. It con firms the presence of ARCH effects in the return series.

Implication: The results from above test suggest the variation and the presence of volatility c lustering or heteroskedasticity in the residual. Hence,we proceed with Residual modelling ass uming Heteroskedasticity.

GARCH Model

```
Estimate Std. Error t value Pr(>|t|) 0.001629 0.000483 3.37609 0.000735 0.850800 0.158378 -5.37194 0.000000 0.805218 0.162075 4.96817 0.000001
ar1
          -0.850800
          0.805218
ma1
                            0.042434 -1.44941 0.147224
          -0.061504
ma2
          -0.048708
                            0.042663 -1.14167 0.253590
ma3
                            0.042802 -0.52678 0.598349
0.041980 0.39683 0.691495
0.041383 0.14538 0.884413
0.043742 0.26975 0.787349
          -0.022547
ma4
ma5
          0.016659
ma6
           0.006016
           0.011800
ma7
                            0.041384 -0.64244 0.520586
ma8
          -0.026587
                            0.041580 -0.10988 0.912506
0.031017 1.33476 0.181955
0.000005 1.85442 0.063679
          -0.004569
ma9
ma10
           0.041400
           0.000009
omega
                            0.007315 13.74806 0.000000
alpha1 0.100560
                            0.015039 58.48443 0.000000
           0.879553
beta1
Robust Standard Errors:
                                          t value Pr(>|t|)
           Estimate Std. Error
                           0.000554 2.939004 0.003293
           0.001629
                            0.107414 -7.920721 0.000000
0.108781 7.402179 0.000000
0.040998 -1.500167 0.133571
          -0.850800
ar1
          0.805218
ma1
          -0.0615\overline{04}
ma2
                            0.040213 -1.211249 0.225800
ma3
          -0.048708
                           0.040213 -1.211249 0.225800
0.037859 -0.595552 0.551475
0.042378 0.393100 0.694245
0.048009 0.125312 0.900277
0.049301 0.239337 0.810844
         -0.022547
ma4
         0.016659
ma5
ma6
          0.006016
          0.011800
ma7
ma8
         -0.026587
                            0.042590 -0.624253 0.532461
                            0.048314 -0.094564 0.924661
ma9
          -0.004569
                            ma10
          0.041400
omega
           0.000009
alpha1 0.100560
                            0.031194 28.196524 0.000000
           0.879553
beta1
LogLikelihood: 2581.116
Information Criteria
Akaike
                 -5.1067
                -5.0334
Bayes
Shibata -5.1071
Hannan-Quinn -5.0788
Weighted Ljung-Box Test on Standardized Residuals
                                   statistic p-value
                                     0.2168 0.6415
10.5817 1.0000
23.7921 0.8306
Lag[1]
Lag[2*(p+q)+(p+q)-1][32]
Lag[4*(p+q)+(p+q)-1][54]
d.o.f=11
HO: No serial correlation
Weighted Ljung-Box Test on Standardized Squared Residuals
                                  statistic p-value
Lag[1]
Lag[2*(p+q)+(p+q)-1][5]
Lag[4*(p+q)+(p+q)-1][9]
d.o.f=2
                                  0.397 0.5286
1.293 0.7905
                                       1.293
                                                 0.7905
                                       2.580
                                               0.8258
Weighted ARCH LM Tests
                 Statistic Shape Scale P-Value
ARCH Lag[3]
ARCH Lag[5]
                0.5320 0.500 2.000 0.4658
0.8216 1.440 1.667 0.7865
2.1047 2.315 1.543 0.6951
ARCH Lag[7]
```

```
Nyblom stability test
Joint Statistic: 4.8789
Individual Statistics: mu 0.07357
        0.32658
ar1
        0.31180
ma1
ma2
        0.19616
ma3
        0.19373
        0.24128
ma4
ma5
        0.34670
       0.34163
0.19204
0.15536
ma6
ma7
ma8
        0.15109
ma9
ma10
        0.12448
omega
       0.91431
alpha1 0.35049
beta1 0.24359
Asymptotic Critical Values (10% 5% 1%)
                           3.26 3.54 4.07
0.35 0.47 0.75
Joint Statistic:
Individual Statistic:
Sign Bias Test
                     t-value prob sig
                     0.84147 0.4003
Sign Bias
Negative Sign Bias 0.36873 0.7124
Positive Sign Bias 0.03922 0.9687
Joint Effect
                      1.11029 0.7746
Adjusted Pearson Goodness-of-Fit Test:
  group statistic p-value(g-1)
                       0.00167\overline{37}
             42.18
     20
2
     30
              61.66
                        0.0003826
3
                        0.0010628
     40
             71.82
              90.47
                        0.0002870
garch_model2 = ugarchspec(variance.model = list(model = 'sGARCH', garchOrder = c(1,1)),
mean.model = list(armaOrder = c(1,5), include.mean = FALSE))
aapl_ret_garch2 = ugarchfit(garch_model2, data = aapl_ret); aapl_ret_garch2
           GARCH Model Fit
Conditional Variance Dynamics
   : sGARCH(1,1)
GARCH Model
Mean Model
                : ARFIMA(1,0,5)
Distribution
Optimal Parameters
        Estimate Std. Error t value Pr(>|t|)
-0.749896 0.210726 -3.558623 0.000373
0.717150 0.212757 3.370751 0.000750
        -0.749896
0.717150
ar1
ma1
ma2
        -0.041408
                       0.040885 -1.012799 0.311156
ma3
        -0.035621
                       0.040539 -0.878685 0.379572
                      ma4
         0.000682
         0.038736
ma5
                      0.000003 2.807170 0.004998
         0.000009
omega
```

```
alpha1 0.093324
                                      0.010931 8.537384 0.000000
0.013652 64.814471 0.000000
beta1
            0.884877
Robust Standard Errors:
               -0.749896
ar1

      ar1
      -0.749896
      0.115961
      -6.466809
      0.00000

      ma1
      0.717150
      0.125337
      5.721783
      0.00000

      ma2
      -0.041408
      0.037957
      -1.090931
      0.27530

      ma3
      -0.035621
      0.037129
      -0.959400
      0.33736

      ma4
      0.000682
      0.035942
      0.018966
      0.98487

      ma5
      0.038736
      0.031414
      1.233077
      0.21755

      omega
      0.000009
      0.000008
      1.234759
      0.21692

      alpha1
      0.093324
      0.016777
      5.562732
      0.00000

      beta1
      0.884877
      0.025307
      34.965116
      0.00000

LogLikelihood: 2574.294
Information Criteria
Akaike -5.1051
Bayes -5.0611
Shibata -5.1052
Hannan-Quinn -5.0883
Weighted Ljung-Box Test on Standardized Residuals
                                               statistic p-value
                                                0.01874 0.8911
2.91034 1.0000
9.85126 0.9758
\text{Lag}[2^{\frac{1}{2}}(p+q)+(p+q)-1][17]
\text{Lag}[4*(p+q)+(p+q)-1][29]
d.o.f=6
HO: No serial correlation
Weighted Ljung-Box Test on Standardized Squared Residuals
statistic p-value
Lag[1] 0.1947 0.6591
Lag[2*(p+q)+(p+q)-1][5] 1.1818 0.8173
Lag[4*(p+q)+(p+q)-1][9] 2.5383 0.8320 d.o.f=2
Weighted ARCH LM Tests
                                           _____
Statistic Shape Scale P-Value
ARCH Lag[3] 0.8042 0.500 2.000 0.3698
ARCH Lag[5] 1.2637 1.440 1.667 0.6566
ARCH Lag[7] 2.4836 2.315 1.543 0.6158
Nyblom stability test
Joint Statistic: 5.4648
Individual Statistics:
ar1 0.4395
            0.4417
ma1
ma2
           0.2498
ma3
            0.2249
          0.2243
ma4
ma5 0.4123
omega 1.0268
alpha1 0.3896
beta1 0.2306
Asymptotic Critical Values (10% 5% 1%) Joint Statistic: 2.1 2.32 2.82
Individual Statistic: 0.35 0.47 0.75
Sign Bias Test
```

```
t-value prob sig
Sign Bias 1.1724 0.2413
Negative Sign Bias 0.2548 0.7989
Positive Sign Bias 0.1650 0.8690
Joint Effect 1.9601 0.5807
```

Adjusted Pearson Goodness-of-Fit Test:

group	statistic	p-value(g-1)
20	41.39	0.002142
30	51.09	0.006875
40	62.98	0.008844
50	73.96	0.012150
	20 30 40	30 51.09 40 62.98

Objective: The objective was to fit GARCH models to the residuals of the ARIMA(1, 1, 5) mean model and evaluate volatility clustering.

Analysis: Two GARCH models ('garch_model1' and 'garch_model2') were fitted to the resi duals, and an ARCH test was conducted on squared residuals to assess volatility clustering.

Results:

Model Comparison: garch_model2 (sGARCH(1,1) with ARFIMA(1,0,5) mean model) appears to be superior to garch_model1 (sGARCH(1,1) with ARFIMA(1,0,10) mean model) based on the following:

Higher LogLikelihood: garch_model1 has a higher LogLikelihood value (2581.116) compa red to garch_model2 (2574.294) indicating a better fit to the data.

Information Criteria: Two information criteria are lower for model 1, two are lower for model 2.

Ljung-Box Tests: Both models exhibit significant serial correlation in the standardized residuals and squared residuals, as indicated by the p-values of the Ljung-Box tests.

ARCH LM Tests: There are no significant effect of conditional heteroskedasticity in either of the models.

Nyblom Stability Test: Both models pass the Nyblom stability test, with all individual statist ics below the critical values.

Adjusted Pearson Goodness-of-Fit Test: Both models exhibit a small p-value for the test st atistic, indicating good goodness-of-fit.

Therefore, considering the higher LogLikelihood, similar information criteria, and the absence of significant ARCH effects in both models, garch_model1 (sGARCH(1,1) with ARFIMA(1,0,10) mean model) can be considered a better fit for the data compared to garch_model1. Further analysis will be required to address the significant serial correlation observed in the st andardized residuals and squared residuals.

GARCH Forecast

aapl_ret_garch_forecast1 = ugarchforecast(aapl_ret_garch1, n.ahead = 50); aapl_ret_garch_f
orecast1

```
*____*
* GARCH Model Forecast *
Model: sGARCH
Horizon: 50
Roll Steps: 0
Out of Sample: 0
0-roll forecast [T0=2023-12-29]:
       Series Sigma
0.001469 0.01091
T+1
T+2 0.001961 0.01121
T+3 0.001848 0.01149
T+4 0.001252 0.01176
T+5 0.001768 0.01202
T+6 0.001273 0.01227
      0.001652 0.01251
0.001716 0.01274
0.001599 0.01296
T+7
T+8
T+9
T+10 0.001349 0.01318
T+11 0.001868 0.01339
T+12 0.001426 0.01359
T+13 0.001802 0.01378
T+14 0.001482 0.01396
T+15 0.001754 0.01414
T+16 0.001523 0.01432
T+17 0.001720 0.01449
T+18 0.001552 0.01465
T+19 0.001695 0.01481
T+20 0.001574 0.01496
T+21 0.001677 0.01511
T+22 0.001589 0.01526
T+23 0.001664 0.01540
T+24 0.001600 0.01554
T+25 0.001654 0.01567
T+26 0.001608 0.01580
T+27 0.001647 0.01592
T+28 0.001614 0.01604
T+29 0.001642 0.01616
T+30 0.001618 0.01628
T+31 0.001639 0.01639
T+32 0.001621 0.01650
T+33 0.001636 0.01661
T+34 0.001624 0.01671
T+35 0.001634 0.01681
T+36 0.001625 0.01691
T+37 0.001633 0.01701
T+38 0.001626 0.01710
T+39 0.001632 0.01719
T+40 0.001627 0.01728
T+41 0.001631 0.01737
T+42 0.001628 0.01745
T+43 0.001631 0.01753
T+44 0.001628 0.01761
T+45 0.001630 0.01769
T+46 0.001629 0.01777
T+47 0.001630 0.01784
T+48 0.001639 0.01792
T+49 0.001630 0.01799
T+50 0.001629 0.01806
```

aapl_ret_garch_forecast2 = ugarchforecast(aapl_ret_garch2, n.ahead = 50); aapl_ret_garch_f
orecast2

```
* GARCH Model Forecast
Model: sGARCH
Horizon: 50
Roll Steps: 0
Out of Sample: 0
0-roll forecast [T0=2023-12-29]:
       Series Sigma
-3.410e-04 0.01106
T+1
        2.796e-04 0.01136
T+2
        1.158e-06 0.01165
T+3
T+4
        9.598e-05 0.01192
       -2.926e-04 0.01218
T+5
        2.194e-04 0.01243
T+6
T+7
       -1.646e-04 0.01267
       1.234e-04 0.01290
T+8
T+9 -9.254e-05 0.01313
T+10 6.940e-05 0.01334
T+11 -5.204e-05 0.01355
       3.902e-05 0.01374
T+12
T+13 -2.926e-05 0.01394
T+14 2.195e-05 0.01412
T+15 -1.646e-05 0.01430
T+16 1.234e-05 0.01447
T+17 -9.254e-06 0.01464
       6.940e-06 0.01480
T+18
T+19 -5.204e-06 0.01495
T+20 3.903e-06 0.01510
T+21 -2.926e-06 0.01525
       2.195e-06 0.01539
T+22
T+23 -1.646e-06 0.01553
T+24 1.234e-06 0.01566
T+25 -9.254e-07 0.01579
T+26 6.940e-07 0.01592
T+27 -5.204e-07 0.01604
T+28 3.903e-07 0.01616
T+29 -2.927e-07 0.01627
T+30 2.195e-07 0.01638
T+31 -1.646e-07 0.01649
T+32 1.234e-07 0.01660
T+33 -9.255e-08 0.01670
T+34 6.940e-08 0.01680
T+35 -5.204e-08 0.01689
T+36 3.903e-08 0.01699
T+37 -2.927e-08 0.01708
T+38 2.195e-08 0.01717
T+39 -1.646e-08 0.01726
T+40 1.234e-08 0.01734
T+41 -9.255e-09 0.01742
T+42 6.940e-09 0.01750
T+43 -5.204e-09 0.01758
T+44 3.903e-09 0.01766
T+45 -2.927e-09 0.01773
T+46 2.195e-09 0.01780
T+47 -1.646e-09 0.01787
       1.234e-09 0.01794
T+48
T+49 -9.255e-10 0.01801
       6.940e-10 0.01807
```

Objective: To forecast volatility using the fitted GARCH model for the next 50 time points.

Analysis: Used the 'ugarchforecast' function to generate volatility forecasts for the next 50 ti me points.

Result: GARCH Model Forecast: - Model: sGARCH - Horizon: 50 - Roll Steps: 0 - Out of S ample: 0 0-roll forecast [T0=2023-12-29]: - Forecasted Series: - T+1 to T+100: Contains fore casted values of volatility (Sigma) for each time point.

Implication: The forecasted values represent the predicted volatility for the next 50 time points based on the fitted GARCH model. These forecasts can be useful for risk management and decision-making, providing insights into the expected future volatility of the financial time series.

plot(aapl_ret_garch_forecast1)

