Homework #3

Analysis of Arsenic in Rice Products

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1. Shared variance across groups

1.1 Definitions and derivations

Let the within- and between- group sampling models be normally-distributed with:

$$\phi_j = \{y|\phi_j\}, \ p(y|\phi_j) = \text{normal}(\theta_j, \sigma^2) \text{ (within group)}$$

$$\psi = \{\mu, \tau^2\}, \ p(\theta_j|\psi) = \text{normal}(\mu, \tau^2) \text{ (between-group)}$$

In this conditionally-conjugate specification:

$$1/\sigma^2 \sim \text{gamma} \ (\nu_0/2, \nu_0 \sigma_0^2/2)$$

 $1/\tau^2 \sim \text{gamma} \ (\eta_0/2, \eta_0 \tau_0^2/2)$
 $\mu \sim \text{normal} \ (\mu_0, \gamma_0^2)$

The full conditional distribution of the parameters can be found to be:

$$\left\{ \theta_{j} | \sigma^{2}, y_{1,1}, \dots, y_{n,m} \right\} \sim \text{normal} \left(\frac{n_{j} \bar{y}_{j} / \sigma^{2} + \mu / \tau^{2}}{n_{j} / \sigma^{2} + 1 / \tau^{2}}, \left[n_{j} / \sigma^{2} + 1 / \tau^{2} \right]^{-1} \right)
\left\{ \mu | \theta_{1}, \dots, \theta_{m}, \tau \right\} \sim \text{normal} \left(\frac{m \bar{\theta} / \tau^{2} + \mu_{0} / \gamma_{0}^{2}}{m / \tau^{2} + 1 / \gamma_{0}^{2}}, \left[m / \tau^{2} + 1 / \gamma_{0}^{2} \right]^{-1} \right)
\left\{ 1 / \tau^{2} | \theta_{1}, \dots, \theta_{m}, \mu \right\} \sim \text{gamma} \left(\frac{\eta_{0} + m}{2}, \frac{\eta_{o} \tau_{0}^{2} + \sum (\theta_{j} - \mu)^{2}}{2} \right)
\left\{ 1 / \sigma^{2} | \theta, y_{1}, \dots, y_{n} \right\} \sim \text{gamma} \left(\frac{1}{2} \left[\nu_{0} + \sum_{j=1}^{m} n_{j} \right], \frac{1}{2} \nu_{0} \sigma_{0}^{2} + \sum_{j=1}^{m} \sum_{i=1}^{n_{j}} (y_{i,j} - \theta_{j})^{2} \right) \right)$$

1.2 Analyses

We pick relatively uninformative priors, centering μ around 1 with somewhat large within and between sample variances: $\sigma_0^2 = 10$, $\nu_0 = 1$, $\tau_0^2 = 10$, $\eta_0 = 1$, $\gamma_0^2 = 10$. The marginal distributions of $\theta_1, \ldots, \theta_m, \mu, \sigma^2$ and τ^2 can be obtained from the full condition distributions using a Monte-Carlo Markov-Chain algorithm, Gibbs sampling, which we implement in R as follows::

First, we input the dataset downloaded from Sakai, modified in Stata to have numeric codes for rice products categories.

```
library(foreign)
Y <- read.dta(file="arsenicrice2.dta")</pre>
```

We set the weakly informative prior values

```
n <- nrow(Y)
nu0 <- 1; eta0 <- 1;
t20 <- 10;
mu0 <- 1;
g20 <- s20 <- var(Y$arsenic)</pre>
```

We set initial values for algorithm

```
m <- length(unique(Y$food_num)) #number of groups
n <- sv <- ybar <- rep(NA,m)
for (i in 1:m)
{
    n[i] <- sum(Y$food_num==i)
    sv[i] <- var(Y$arsenic[which(Y$food_num==i)])
    ybar[i] <- mean(Y$arsenic[which(Y$food_num==i)])
}
theta <- ybar; s2 <- mean(sv)
mu <- mean(theta); tau2 <- var(theta)</pre>
```

We create a Markov chain for each parameter by sequentially sampling from their posterior over 10,000 iterations. Elements are stored in the chain at the end of each iteration.

```
#Setup MCMC
set.seed(0808)
S <- 10000
THETA <- matrix(nrow=S, ncol=m)
OTH <- matrix(nrow=S, ncol=3)
ALL <- matrix(nrow=S, ncol=3+m)

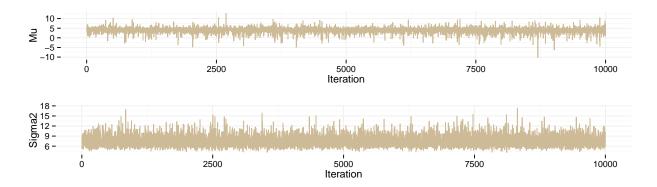
#Run algorithm
for(i in 1:S)
{
    #Get new values for parameters
    for(j in 1:m) theta[j] <- newTheta(n[j], ybar[j], s2, tau2, mu)
    s2 <- newSigma2(m, n, nu0, s20, theta, Y)
    mu <- newMu(m, theta, tau2, g20)
    tau2 <- newTau2(m, eta0, t20, theta, mu)

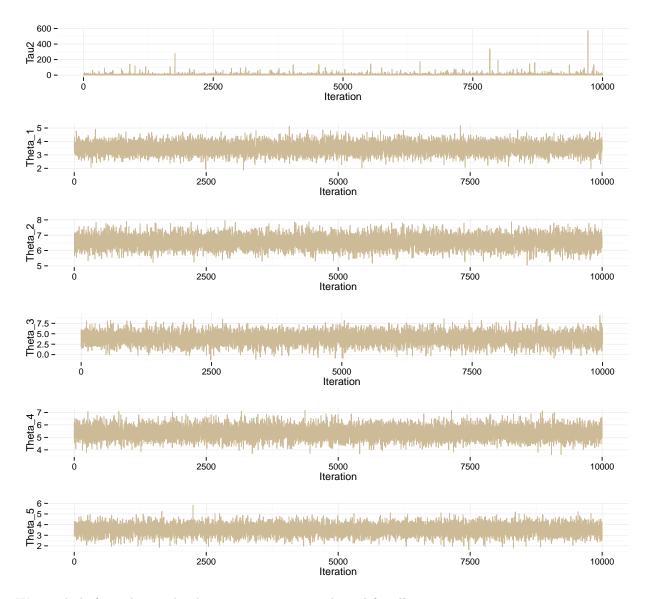
#Store in chain
    THETA[i,] <- theta
    OTH[i,] <- c(mu,s2,tau2)
    ALL[i,] <- c(theta,mu,s2,tau2)
}</pre>
```

Where the functions updating the parameters follow the equations listed above:

```
newTheta <- function(n, ybar, s2, tau2, mu)</pre>
  v = 1/(n/s2 + 1/tau2)
  e = v * (ybar*n/s2 + mu/tau2)
  new <- rnorm(1, e, sqrt(v))</pre>
  return(new)
newSigma2 <- function(m, n, nu0, s20, theta, Y)</pre>
  nun = nu0 + sum(n)
  ss <- nu0 * s20
  for(i in 1:m) ss = ss+sum((Y$arsenic[which(Y$food_num==i)] - theta[j])^2)
  sigma2 <- 1/rgamma(1, nun/2, ss/2)
  return(sigma2)
newMu <- function(m, theta, tau2, g20)</pre>
  v = 1/(m/tau2 + 1/g20)
  e = v *(m*mean(theta)/tau2 + mu0/g20)
  mu <- rnorm(1, e, v)
  return(mu)
newTau2 <- function(m, eta0, t20, theta, mu)</pre>
  etam = eta0 + m
  ss <- eta0*t20 + sum( (theta-mu) ^2)
  tau2 <- 1/rgamma(1, etam/2, ss/2)</pre>
  return(tau2)
```

Before we go any further, we check that the MCMC model converged for all four statistics using ggplot2 (code used for μ repeated for other parameters):





We conclude from the graphs that convergence was achieved for all parameters.

1.3 Algorithm output

The estimated median values and 95% credible intervals for the parameters are as follow:

```
for(i in 1:length(ALL[1,])) print(round(unname(
  quantile(ALL[,i], probs=c(0.025, 0.5, 0.975))
  ),3))
```

| Parameter | Credible Lower 95% | Median | Credible Upper 95% |
|--------------------------|--------------------|--------|--------------------|
| θ_1 (Basmati) | 2.810 | 3.569 | 4.308 |
| θ_2 (Non-Basmati) | 5.790 | 6.533 | 7.305 |
| θ_3 (Beverage) | 1.869 | 4.231 | 6.448 |
| θ_4 (Cakes) | 4.486 | 5.370 | 6.301 |
| θ_5 (Cereal) | 2.705 | 3.610 | 4.507 |
| μ | 3.193 | 4.697 | 6.182 |
| σ^2 | 5.109 | 7.027 | 10.573 |
| τ^2 | 0.693 | 2.251 | 12.920 |

1.4 Sensitivity analyses

Evaluation of sensitivity to priors: we try three separate scenarios each tuning prior distribution of parameters:

- 1. Large expected μ (Prior expectation of mad levels of arsenic)
- 2. Large σ^2 and ν_0 (High variability within products)
- 3. Large τ^2 and η_0 (High variability between products)

Scenario 1:

| Parameter | Credible Lower 95% | Median | Credible Upper 95% |
|----------------------------|--------------------|--------|--------------------|
| θ_1 (Basmati) | 2.706 | 3.488 | 4.265 |
| θ_2^2 (Non-Basmati) | 5.902 | 6.671 | 7.460 |
| $\theta_3(Beverage)$ | 0.643 | 3.774 | 6.932 |
| $\theta_4(Cakes)$ | 4.511 | 5.452 | 6.429 |
| θ_5 (Cereal) | 2.548 | 3.496 | 4.456 |
| μ | 2.548 | 3.496 | 4.456 |
| σ^2 | 88.793 | 100 | 111 |
| $	au^2$ | 3021.927 | 8427 | 37962 |

Scenario 2:

| Parameter | Credible Lower 95% | Median | Credible Upper 95% |
|----------------------------|--------------------|--------|--------------------|
| θ_1 (Basmati) | 1.237 | 4.396 | 7.069 |
| θ_2^2 (Non-Basmati) | 2.811 | 5.401 | 8.559 |
| $\theta_3(Beverage)$ | 0.278 | 4.790 | 9.023 |
| $\theta_4(Cakes)$ | 1.931 | 4.960 | 8.221 |
| θ_5 (Cereal) | 1.137 | 4.531 | 7.508 |
| μ | 2.42 | 4.84 | 7.19 |
| σ^2 | 182 | 215 | 256 |
| $	au^2$ | 0.399 | 1.876 | 22.3 |

Scenario 3:

| Parameter | Credible Lower 95% | Median | Credible Upper 95% |
|----------------------------|--------------------|--------|--------------------|
| θ_1 (Basmati) | 2.707 | 3.478 | 4.268 |
| θ_2^2 (Non-Basmati) | 5.907 | 6.674 | 7.451 |
| $\theta_3(Beverage)$ | 0.676 | 3.739 | 6.9023 |
| $\theta_4(Cakes)$ | 4.504 | 5.447 | 6.419 |
| θ_5 (Cereal) | 2.547 | 3.494 | 4.452 |
| μ | -5.302 | 4.845 | 5.00 |
| $rac{\mu}{\sigma^2}$ | 5.197 | 7.327 | 11.37 |
| $	au^2$ | 223 | 288 | 384 |

We observe that excessively large prior expectations of μ will drive estimates of the within- and betweengroup variances but will have little effect on the magnitude of the estimates of within-groupmean estimates (although the precision may be negatively affected for groups with realtively few observations). A large prior within-sample variance will bring posterior within-group means closer to μ , as could be expected since the posterior estimates need to become more conservative. Increasing prior between-sample variance appears to drive up uncertainty on μ and bring it closer to 0, without however having a notable impact on the rest of the model.

1.5 Results presentation

Non-Basmati rice had the highest arsenic concentration, at an estimated 6.7 mcg/serving. Rice cakes came second, at 5.4 mcg/serving, and non-Basmati and rice cereal had comparatively low amounts, slightly below 3.5 mcg/serving. There lacked data to reliably evaluate arsenic concentration in rice beverages, whose 3.8 mcg/serving estimate was particularly imprecise (95% CI=0.64, 6.93). Posterior median estimates and observed mean concentrations of arsenic are presented by product type in the following graph. Markers are θ estimates with 95% credible interval lines; horizontal lines are the median estimate of μ (solid) and corresponding 95% credibal interval (dashed).

