

# LabH\_\_Hw

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## Q1. Addressing Heteroskedasticity:

Heteroskedasticity is when the error term is not constant across cases. This makes estimating the confidence intervals and t-statistics unreliable. This means that the statistics we use to test hypothesis under the Gauss-Markov assumptions are no longer valid. So while our beta coefficients are unaffected, the variance of these estimates are not which can have implications for how accurate other assumptions about the model are.

## Q2. Addressing Multicollinearity:

Multicollinearity is a problem because it forces the coefficients for correlated data to be derived from a subset of the data in a dataset, ultimately decreasing significance/reliability. If two variables are correlated, the correlated cases are not independent. Non-independent cases cannot be used to calculate the beta coefficient because there is not enough variation within the data. So, if for example, 70% of cases in x and w are correlated, only 30% of the data points in x and w are being used to calculate the coefficient. This is a problem because it makes the calculation weaker.

## Q3. Diagnosis for Dataset A

```
# Load data
library(readstata13)
library(olsrr)
a <- read.dta13("./data/labwk5_a.dta")

# Create model and summarize
mod_a <- lm(y~x+z+q+w, a)
ols_regress(mod_a)
```

```
##                               Model Summary
## -----
## R                               1.000          RMSE              8.737
## R-Squared                       0.999          Coef. Var        43.006
## Adj. R-Squared                   0.999          MSE             76.339
## Pred R-Squared                   0.999          MAE             7.051
## -----
## RMSE: Root Mean Square Error
## MSE: Mean Square Error
## MAE: Mean Absolute Error
##
##                               ANOVA
## -----
##                               Sum of
##                               Squares          DF          Mean Square          F          Sig.
## -----
## Regression      49824149.482              4      12456037.371      163168.239      0.0000
## Residual          37787.614             495           76.339
## Total          49861937.097             499
## -----
```

```
##
##                                     Parameter Estimates
## -----
##      model      Beta      Std. Error      Std. Beta      t      Sig      lower      upper
## -----
## (Intercept)      0.354      0.392      0.903      0.367      -0.416      1.124
##      x      25.030      0.031      1.003      806.484      0.000      24.969      25.091
##      z      -22.538      0.396      -0.071      -56.942      0.000      -23.315      -21.760
##      q       3.345      0.534      0.011      6.259      0.000      2.295      4.395
##      w      -0.309      0.536      -0.001      -0.577      0.564      -1.362      0.744
## -----
```

Initial analysis shows that the intercept and coefficient for w are not significant. There may be issues with these data. However, the R-squared is nearly 1, so these data are almost perfectly explanatory of y. Notice that the coefficients for x and z are rather large.

Next, we will use the Breusch-Pagan test to check for heteroskedasticity.

```
# Test for heteroskedasticity using Breusch-Pagan test
ols_test_breusch_pagan(mod_a, rhs = TRUE, multiple = TRUE)
```

```
##
## Breusch Pagan Test for Heteroskedasticity
## -----
## Ho: the variance is constant
## Ha: the variance is not constant
##
##      Data
## -----
## Response : y
## Variables: x z q w
##
##      Test Summary (Unadjusted p values)
## -----
## Variable      chi2      df      p
## -----
## x      0.7424041      1      0.3888918
## z      0.8627134      1      0.3529805
## q      1.1697741      1      0.2794477
## w      2.0117363      1      0.1560866
## -----
## simultaneous      3.2918332      4      0.5102273
## -----
```

None of the explanatory variables exhibit p-values less than 0.05 in the Breusch-Pagan test, which means we can conclude that there is no heteroskedasticity in this data. Check this again using robust regression. To calculate the estimate of variance for each case, we will use the Huber-White robust standard errors test:

```
library(foreign)
library(sandwich)
library(lmtest)

# Estimate standard errors with robust regression
coeftest(mod_a, vcov = vcovHC(mod_a, "HC1"))
```

```
##
## t test of coefficients:
```

```
##
##           Estimate Std. Error   t value Pr(>|t|)
## (Intercept)  0.354093   0.391646    0.9041  0.3664
## x           25.030284   0.014385  1740.0042 < 2.2e-16 ***
## z          -22.537628   0.372649   -60.4795 < 2.2e-16 ***
## q            3.344992   0.506251    6.6074 1.013e-10 ***
## w          -0.309311   0.539233   -0.5736  0.5665
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

No significant difference between the standard errors in the robust regression. The same variables are not significant as in the original regression (intercept and w).

We know there is no heteroskedasticity in this model. Next, test for multicollinearity

```
# correlation matrix
cor(a[,c(1,2,3,4,6)])
```

```
##           x           z           w           q           y
## x 1.00000000 0.08922922 0.01918552 0.04649155 0.99711301
## z 0.08922922 1.00000000 0.07084426 0.06565461 0.01920997
## w 0.01918552 0.07084426 1.00000000 0.69592903 0.02073744
## q 0.04649155 0.06565461 0.69592903 1.00000000 0.05207843
## y 0.99711301 0.01920997 0.02073744 0.05207843 1.00000000
```

There is high correlation between x and y. All other relationships appear uncorrelated.

Next, calculate the VIF.

```
# VIF check
```

```
ols_vif_tol(mod_a)
```

```
## # A tibble: 4 x 3
##   Variables Tolerance   VIF
##   <chr>      <dbl> <dbl>
## 1 x          0.990  1.01
## 2 z          0.987  1.01
## 3 q          0.514  1.94
## 4 w          0.515  1.94
```

There are no significantly high VIF values ( $VIF > 10$ ). This isn't surprising since it seems like the main correlation is between x and y. Multicollinearity is not really a problem in this data, but there is probably something going on with the x data. Even if there is correlation between the x and y, the lack of a significant VIF means that we don't have to worry much until we have more information.

Next, we will check for outliers and influential cases.

```
# Start by looking at the basic stats of the data
summary(a)
```

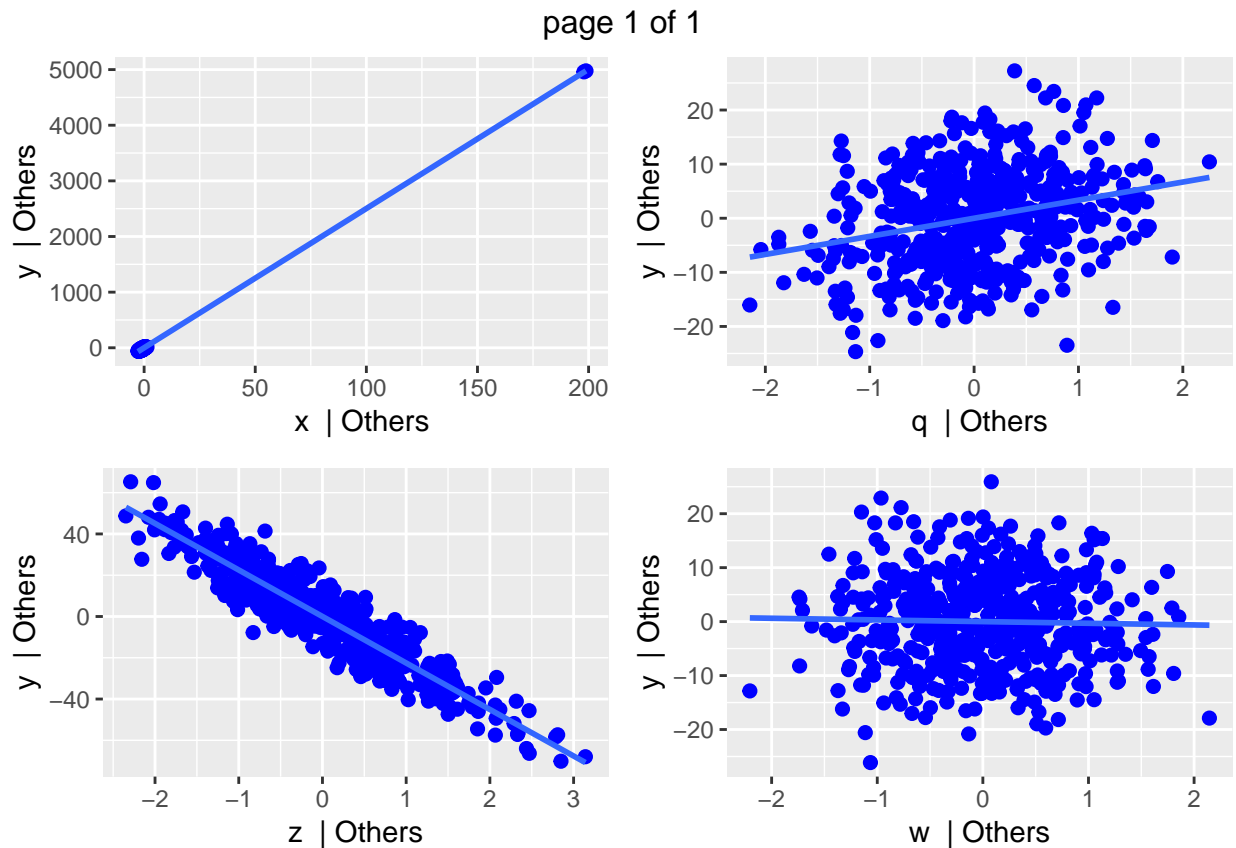
```
##           x           z           w
## Min.      : -2.82066   Min.      : -2.32055   Min.      : -3.12846
## 1st Qu.: -0.61709   1st Qu.: -0.63376   1st Qu.: -0.76479
## Median : -0.02613   Median : -0.01503   Median : -0.01857
## Mean     :  0.81725   Mean     :  0.01656   Mean     : -0.03512
## 3rd Qu.:  0.67864   3rd Qu.:  0.63459   3rd Qu.:  0.68867
## Max.     : 200.20560   Max.     :  3.08475   Max.     :  2.76435
##           q           n           y           d
```

```
## Min.   :-3.0749   Min.    :  1.0   Min.    : -18.960   Min.     :0.0000000
## 1st Qu.: -0.6942   1st Qu.:125.8   1st Qu.:  -3.676   1st Qu.:0.0001465
## Median :-0.1107   Median :250.5   Median :   0.390   Median :0.0006241
## Mean   :-0.0394   Mean   :250.5   Mean    : 20.316   Mean    :0.0019112
## 3rd Qu.: 0.6962   3rd Qu.:375.2   3rd Qu.:   4.402   3rd Qu.:0.0019800
## Max.    : 2.8090   Max.    :500.0   Max.    :5004.394   Max.     :0.1688666
```

It appears as though there may be something unusual about the x data. The maximum is much higher than the other variables, and is very distant from the 3rd quartile value. It would make sense if there are some outlier cases in the x data at the upper limit.

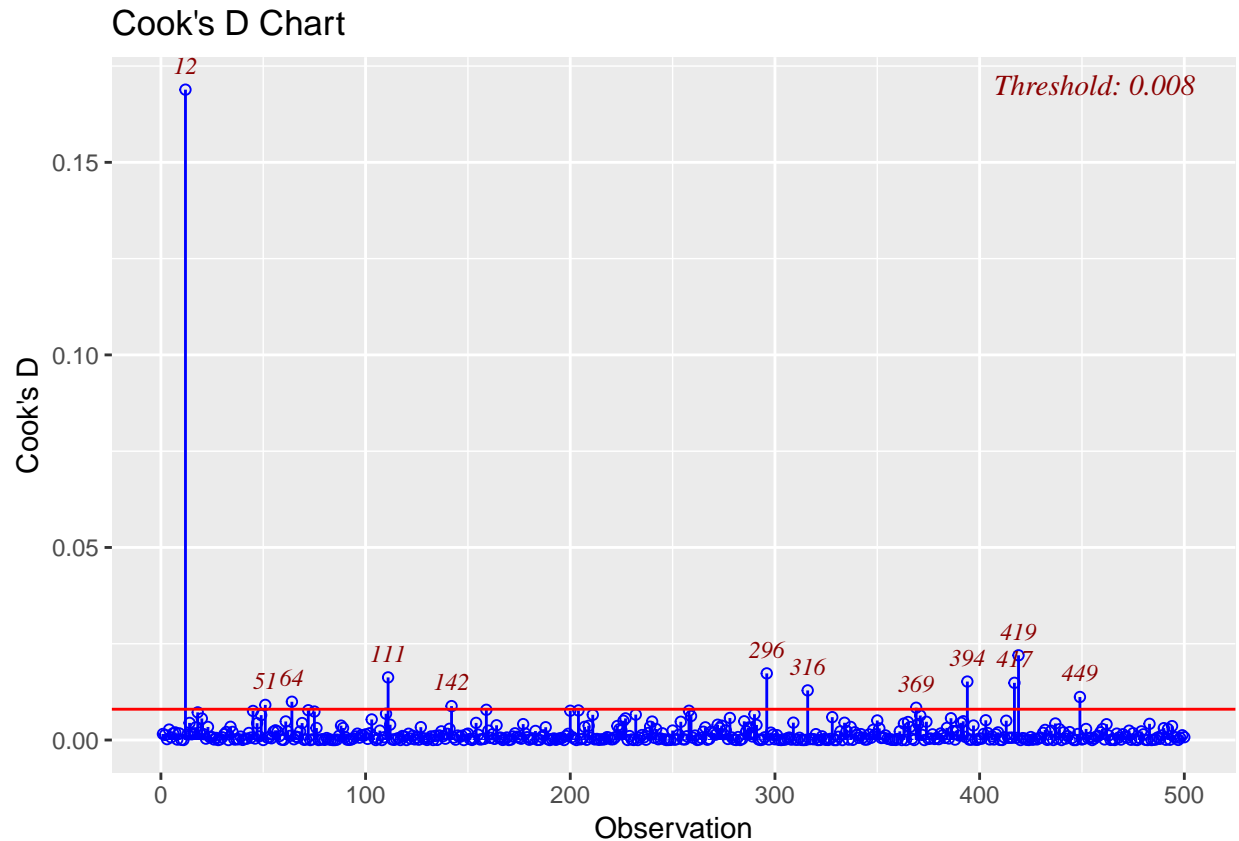
Next, added-variable plots:

```
ols_plot_added_variable(mod_a)
```



These av-plots indicate there is something wrong with the x data. There is likely at least one outlier case in the data. Next, we will identify this case using Cook's D:

```
# identify influential case
ols_plot_cooksd_chart(mod_a)
```



This chart shows us exactly which cases are the problem (n = 11 and 12). It also shows us the threshold to classify cases as outliers. In this model, case 12 is far, far over the threshold. There are other outlier cases in these data, but this is likely normal variation. It would probably not be ethical to exclude the cases that are somewhat above the threshold, but it seems obvious that there was a coding error in case 12.

(Note: I know for some reason it's difficult to see the actual cases from the chart so that case 11 seems to not be identifiable here. I couldn't get it to plot properly.)

Next, let's exclude this outlier case and regress.

```
mod_a_fixed <- lm(y~x+w+z+q,a[-c(11,12),])
ols_regress(mod_a_fixed)
```

```
##                               Model Summary
## -----
## R                               0.626          RMSE                4.733
## R-Squared                       0.391          Coef. Var          1540.379
## Adj. R-Squared                  0.386          MSE                22.397
## Pred R-Squared                  0.379          MAE                3.753
## -----
## RMSE: Root Mean Square Error
## MSE: Mean Square Error
## MAE: Mean Absolute Error
##
##                               ANOVA
## -----
##                               Sum of
##                               Squares      DF      Mean Square      F      Sig.
```

```
## -----
## Regression      7096.715      4      1774.179      79.217      0.0000
## Residual       11041.513     493      22.397
## Total          18138.228     497
## -----
##
##                               Parameter Estimates
## -----
##      model      Beta      Std. Error      Std. Beta      t      Sig      lower      upper
## -----
## (Intercept)    0.408        0.212              1.923    0.055    -0.009    0.825
##      x         1.079        0.694        0.182    1.556    0.120    -0.284    2.442
##      w         0.480        0.292        0.081    1.646    0.100    -0.093    1.053
##      z         0.797        0.709        0.131    1.125    0.261    -0.595    2.190
##      q         2.747        0.290        0.464    9.475    0.000     2.178    3.317
## -----
```

The coefficient for x is no longer highly inflated, nor is the coefficient for z. However, both of these variables' coefficients are no longer significant.

There was an issue with correlation in the first model, so let's check if this is still an issue.

```
cor(a[-c(11,12)],(c(1,2,3,4,6)))
```

```
##      x      z      w      q      y
## x 1.00000000 0.95349564 0.08099674 0.05888374 0.3411657
## z 0.95349564 1.00000000 0.06801159 0.06330737 0.3398196
## w 0.08099674 0.06801159 1.00000000 0.69487565 0.4267677
## q 0.05888374 0.06330737 0.69487565 1.00000000 0.5390274
## y 0.34116567 0.33981965 0.42676768 0.53902735 1.00000000
```

Removing the two influential-outlier cases did fix the problem of correlation between x and y. However, there seems to be a problem of multicollinearity between z and x now. Let's check the VIF:

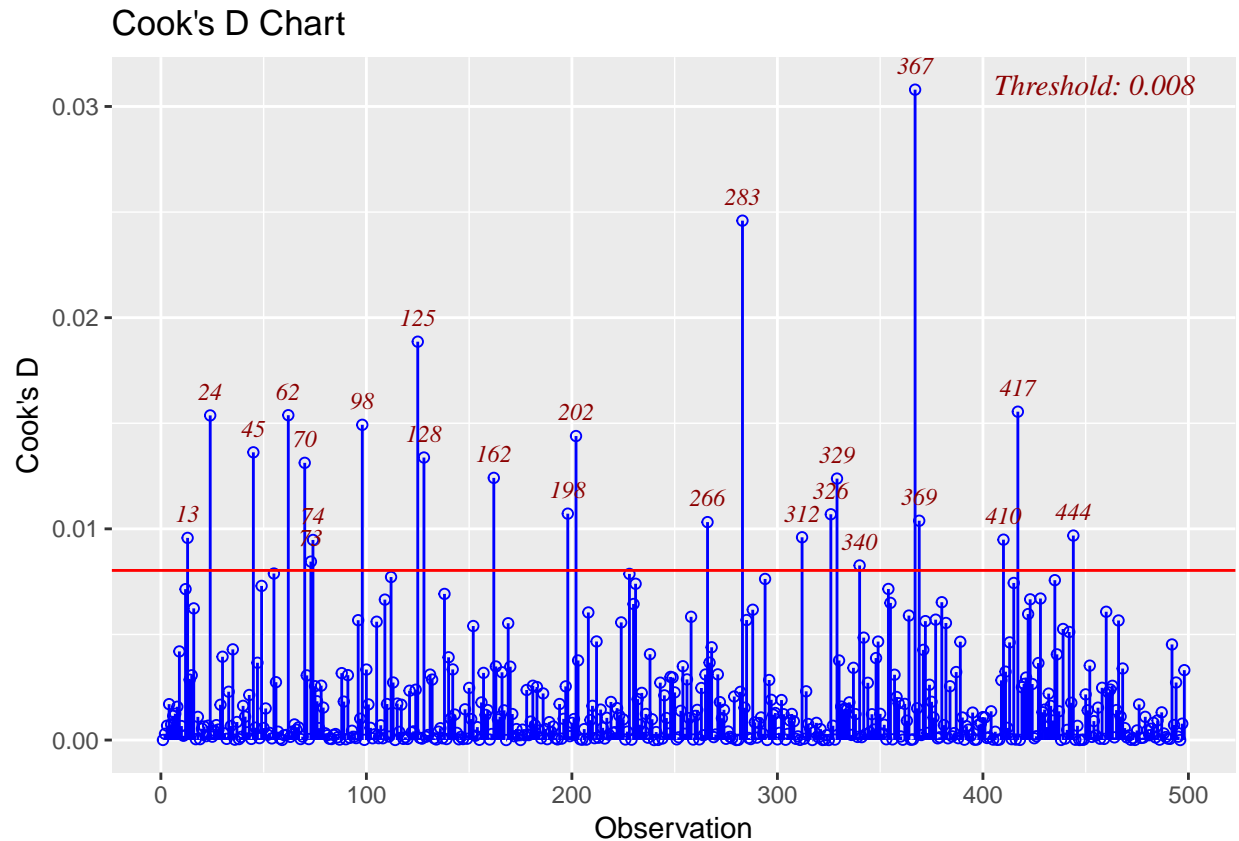
```
ols_vif_tol(mod_a_fixed)
```

```
## # A tibble: 4 x 3
##   Variables Tolerance VIF
##   <chr>      <dbl> <dbl>
## 1 x          0.0903 11.1
## 2 w          0.513  1.95
## 3 z          0.0904 11.1
## 4 q          0.515  1.94
```

As suspected, there is significant correlation. The best solution to this problem would be to collect more data. Otherwise, we should drop z or x from the model.

One last check: replot Cook's D with our new model that excludes the problematic cases:

```
# identify influential case
ols_plot_cooks_d_chart(mod_a_fixed)
```



Summary: Coding error on cases 11 and 12 seemed to have been the issue in this dataset. By dropping these cases, we were able to fix problems of outlier influence but revealed that the data have problems of multicollinearity.

#### Q4. Diagnosis for Dataset B

```
# Load data
library(readstata13)
library(olsrr)
b <- read.dta13("./data/labwk5_b.dta")
```

```
# Create model and summarize
mod_b <- lm(y~x+z+q+w, b)
ols_regress(mod_b)
```

```
##                               Model Summary
## -----
## R                             0.248      RMSE          49.369
## R-Squared                     0.061      Coef. Var      48.743
## Adj. R-Squared                0.057      MSE           2437.307
## Pred R-Squared                0.052      MAE           39.331
## -----
## RMSE: Root Mean Square Error
## MSE: Mean Square Error
## MAE: Mean Absolute Error
##
```

```
## ANOVA
## -----
## Sum of
## Squares DF Mean Square F Sig.
## -----
## Regression 158291.420 4 39572.855 16.236 0.0000
## Residual 2425120.405 995 2437.307
## Total 2583411.825 999
## -----
## Parameter Estimates
## -----
## model Beta Std. Error Std. Beta t Sig. lower upper
## -----
## (Intercept) -10.501 16.960 -0.619 0.536 -43.782 22.780
## x 11.194 1.689 0.219 6.627 0.000 7.879 14.509
## z 2.947 1.633 0.059 1.804 0.072 -0.258 6.152
## q 0.084 1.627 0.002 0.052 0.959 -3.109 3.277
## w 1.381 1.579 0.027 0.874 0.382 -1.718 4.480
## -----
```

The R-squared of this model is very low. None of the variables are significant, except for x.

First, test for heteroskedasticity using Breusch-Pagan test:

```
# Test for heteroskedasticity using Breusch-Pagan test
ols_test_breusch_pagan(mod_b, rhs = TRUE, multiple = TRUE)
```

```
## Breusch Pagan Test for Heteroskedasticity
## -----
## Ho: the variance is constant
## Ha: the variance is not constant
##
## Data
## -----
## Response : y
## Variables: x z q w
##
## Test Summary (Unadjusted p values)
## -----
## Variable chi2 df p
## -----
## x 33.9761550 1 5.579164e-09
## z 6.6203416 1 1.008205e-02
## q 0.9957297 1 3.183460e-01
## w 0.5720473 1 4.494465e-01
## -----
## simultaneous 36.9753929 4 1.822451e-07
## -----
```

There are major issues with heteroskedasticity in these data. All variables are exhibiting p-values  $< 0.05$  in the BP test, so they are all exhibiting heteroskedasticity. Next, let's check this again using robust regression. To calculate the estimate of variance for each case, we will use the Huber-White robust standard errors test:

```
library(foreign)
library(sandwich)
```



```
library(lmtest)

# Estimate standard errors with robust regression
coeftest(mod_b, vcov = vcovHC(mod_a, "HC1"))

##
## t test of coefficients:
##
##           Estimate Std. Error  t value  Pr(>|t|)
## (Intercept) -10.501000    0.391646 -26.8125 < 2.2e-16 ***
## x           11.193900    0.014385  778.1547 < 2.2e-16 ***
## z            2.947052    0.372649   7.9084 6.904e-15 ***
## q            0.083897    0.506251   0.1657  0.86841
## w            1.380960    0.539233   2.5610  0.01058 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

This seemed to have addressed much of the problem. The standard errors are much smaller and similar to one another, and all but the q coefficient are significant. Heteroskedasticity means that there are issues with the error terms, but not with the estimates of our coefficients, so we shouldn't expect too many issues with outliers and multicollinearity in the rest of the diagnosis.

Next, let's check for multicollinearity:

```
cor(b[, -5])

##           x           z           w           q           y
## x  1.00000000  0.3597645606 -0.02180513  0.0748053571  0.23979892
## z  0.35976456  1.0000000000 -0.01469716  0.0003163632  0.13777428
## w -0.02180513 -0.0146971600  1.00000000 -0.0324633568  0.02117784
## q  0.07480536  0.0003163632 -0.03246336  1.0000000000  0.01711003
## y  0.23979892  0.1377742764  0.02117784  0.0171100278  1.00000000

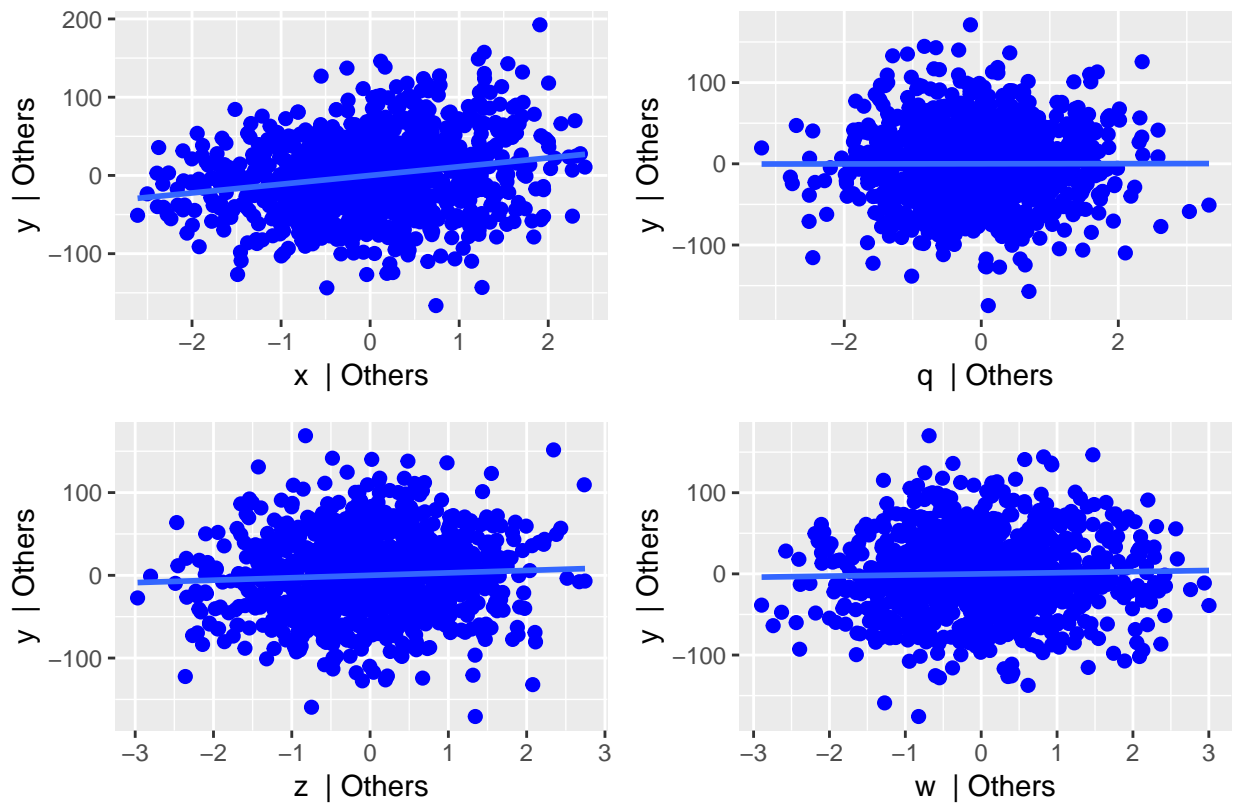
ols_vif_tol(mod_b)

## # A tibble: 4 x 3
##   Variables Tolerance  VIF
##   <chr>      <dbl> <dbl>
## 1 x          0.865  1.16
## 2 z          0.870  1.15
## 3 q          0.993  1.01
## 4 w          0.999  1.00
```

The correlation matrix indicates no correlation. There are no large VIF values, indicating that we don't have a problem with multicollinearity.

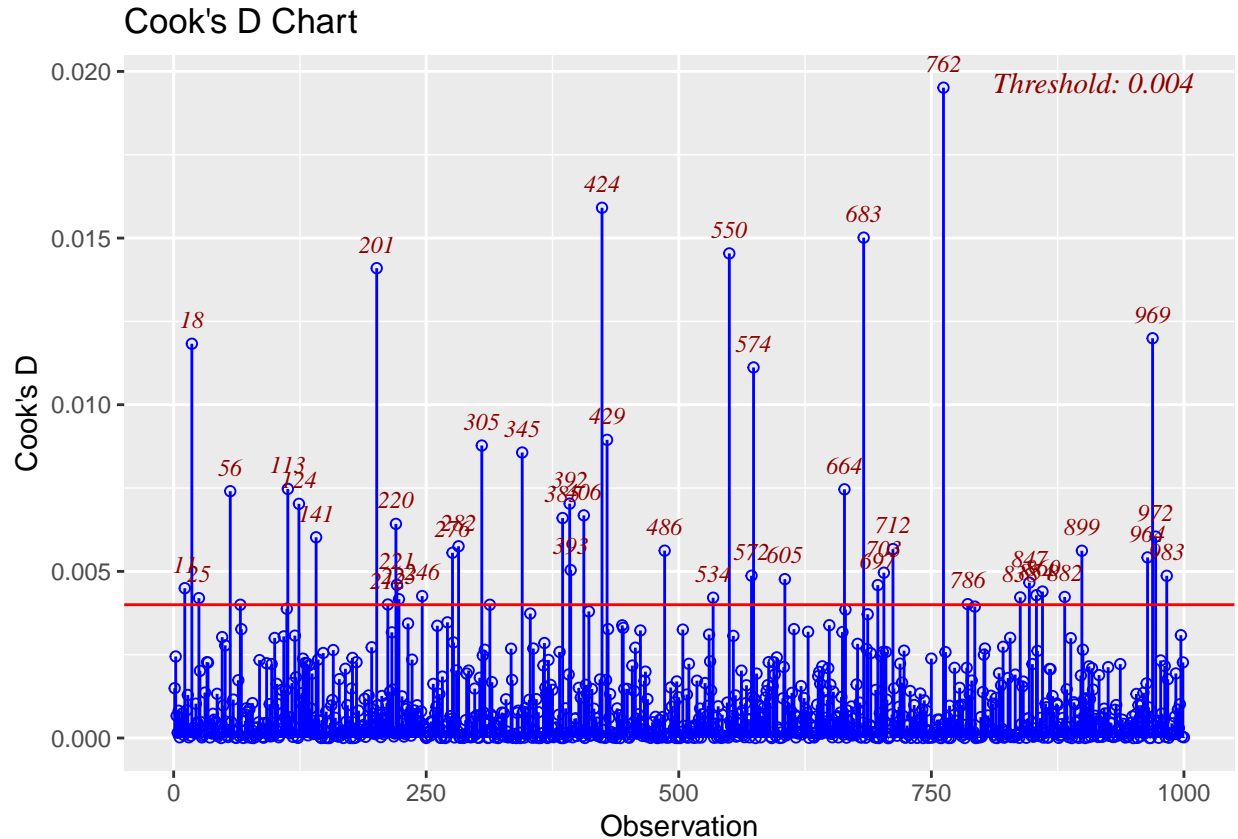
Next, let's check for outliers and influential cases:

```
#Start with added-variable plots
ols_plot_added_variable(mod_b)
```



These av-plots indicate there is no issue with outliers. To be sure, let's plot Cook's D:

```
# identify influential cases
ols_plot_cooksd_chart(mod_b)
```



While there are some cases that cross the threshold here, this seems like the function of regular variation. There is no need to be concerned about outlier cases and their influence on the model in this data set.

Summary: Data set B has major problems with heteroskedasticity, which isn't necessarily a problem, but something that we as researchers/data scientists should be aware of when drawing conclusions about the estimations of the beta coefficients.

## Q5. Class paper

I will be testing the influence of the biophysical environment on fuel use and respiratory infections in energy poor regions of Malawi, Zambia, and Zimbabwe. The idea will be to use survey data from the Demographic and Health Surveys to control for demographic factors, and identify the interaction effects between seasonal variables and fuel use outcomes. Seasonal variation of temperature and rainfall may explain part of the variation of respiratory infection incidence across different geographical contexts. Land change may serve as a proxy for fuel availability, in terms of both quantity and type. Using measures of the environment, we may be able to quantify the magnitude of the effect of biophysical factors on human behaviors and health. The findings can direct intervention strategies to target specific timing and geographic contexts for the most effective intervention. Additionally, information about cooking practices related to the environment can improve intervention design, namely improved cook stoves. An efficient and well-engineered stove may have massive implications for improving fuel efficiency and human health, but if it can only use certain fuels or work outdoors, intended benefits will be limited. This study aims to identify patterns about how biophysical context, though seasonal weather and land cover, may help to identify which populations are well positioned for cooking fuel interventions.

Main question: How does the biophysical environment influence fuel use and respiratory infections among children in energy poor regions?

Hypotheses:

- (1) Respiratory infection incidence will be higher following prolonged periods of temperature decrease.
- (2) We expect to observe a higher rate of respiratory infections due to people cook indoors following prolonged periods of high rainfall.
- (3) We expect to observe higher rates of respiratory infection in areas of recent deforestation.