

CS5340 Uncertainty Modeling in Al

Lecture 2:
Bayesian Networks
(Directed Graphical Models)

Asst. Prof. Lee Gim Hee
AY 2020/21
Semester 1

Course Schedule

Week	Date	Торіс	Remarks
1	12 Aug	Introduction to probabilistic reasoning	1830hrs: MS Teams (Live Introduction)
2	19 Aug	Bayesian networks (Directed graphical models)	
3	26 Aug	Markov random Fields (Undirected graphical models)	1830hrs: Zoom discussions
4	02 Sep	Variable elimination and belief propagation	Assignment 1: Belief propagation and maximal probability (15%)
5	09 Sep	Factor graph and the junction tree algorithm	
6	16 Sep	Parameter learning with complete data	Assignment 1: Due Assignment 2: Junction tree and parameter learning (15%) 1830hrs: Zoom discussions
-	23 Sep	Recess week	No lecture
7	30 Sep	Mixture models and the EM algorithm	Assignment 2: Due Online quiz 1 (20%)
8	07 Oct	Hidden Markov Models (HMM)	Assignment 3: Hidden Markov model (15%)
9	14 Oct	Monte Carlo inference (Sampling)	1830hrs: Zoom discussions
10	21 Oct	Variational inference	Assignment 3: Due Assignment 4: MCMC Sampling (15%)
11	28 Oct	Variational Auto-Encoder and Mixture Density Networks	
12	04 Nov	Graph-cut and alpha expansion	Assignment 4: Due 1830hrs: Zoom discussions
-	11 Nov		Online quiz 2 (20%)



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Acknowledgements

- A lot of slides and content of this lecture are adopted from:
 - "An introduction to probabilistic graphical models", Michael I. Jordan, 2002 http://people.eecs.berkeley.edu/~jordan/prelims/chapter-2.pdf (Section 2.1)
 - 2. "Pattern recognition and machine learning", Christopher Bishop (Chapter 8, Section 8.1 and 8.2).
 - "Machine learning a probabilistic approach", Kevin Murphy (Chapter 10)
 - 4. "Probabilistic graphical models", Koller and Friedman (Chapter 3)



Learning Outcomes

- Students should be able to:
 - 1. Explain the concepts of conditional independence.
 - 2. Use the Bayesian network to represent conditional independence in joint distributions.
 - Describe d-separation using the three canonical 3node graph.
 - Deduce all conditional independence in a Bayesian network using the Bayes ball algorithm.
 - 5. Explain the concepts of Markov Blanket.



It difficult to work with joint probabilities $p(x_1, ..., x_N | \theta)$ with fully correlated random variables.

Why?



- Let's illustrate this with N discrete random variables $X_1, ..., X_N$, where $x_i \in \{1, ..., K\}$.
- We need $O(K^N)$ parameters θ to represent the joint distribution $p(x_1, ..., x_N | \theta)$.
- Inference becomes intractable when N is large, and a huge amount of data is needed to learn all parameters.



Easy solution?

• Fully independent assumption on all random variables reduces the number of parameters to O(NK).

$$p(x_1,...,x_N | \theta) = \prod_{i=1}^{N} p(x_i | \theta_i)$$

• Inference becomes tractable products of $p(x_i|\theta_i)$, and smaller amount of data is needed to learn all parameters.



Fully independent assumption:

$$p(x_1, ..., x_N | \theta) = \prod_{i=1}^{N} p(x_i | \theta_i)$$

 Is it always correct to assume that all random variables are fully independent?

Examples:

Probability of the next letter in a word:

Is independent of the letters that are already known in the word?

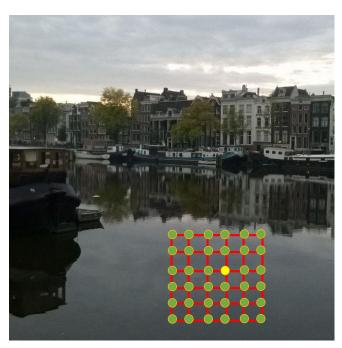


Fully independent assumption :

$$p(x_1, ..., x_N | \theta) = \prod_{i=1}^N p(x_i | \theta_i)$$

 Is it always correct to assume that all random variables are fully independent?

Examples:



: pixel is labeled as "water"

: is this pixel more likely to be "water" or "sky"?

Photo Source: G.H. Lee "Amsterdam"



Random variables are often NOT fully independent, how can we:

- Compactly represent the joint distribution $p(x_1, ..., x_N | \theta)$ of multiple correlated variables?
- Use the joint distribution to *infer* one set of variables given another in a reasonable amount of computation time?
- Learn the parameters of the joint distribution with a reasonable amount of data?

Use Graphical Models!!!



Conditional Independence

- We have seen that:
 - Fully independent assumption is insufficient to model realworld random variables which are unlikely to be fully independent.
 - > Fully correlated joint distributions can become intractable.
- A good compromise is by assuming an intermediate degree of dependency among the random variables.
- This is conditional independence.



Conditional Independence

• More formally, two random variables X_A and X_C are conditionally independent given X_B if:

$$p(x_A, x_C|x_B) = p(x_A|x_B)p(x_C|x_B)$$

Or alternatively:

$$p(x_A|x_B, x_C) = p(x_A|x_B), \quad \forall X_B : p(x_B) > 0$$

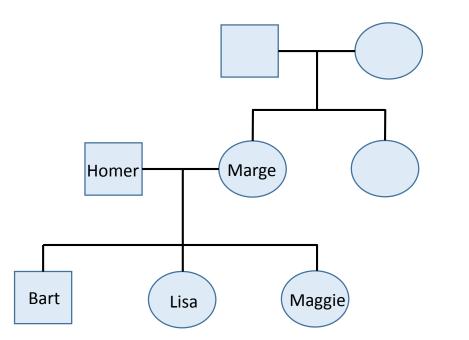
- That is, learning the values of X_C does not change prediction of X_A once we know the value of X_B .
- Written as $X_A \perp X_C \mid X_B$.



Conditional Independence

Example: Family Trees (Pedigree)

A node represents an individual's genotype.



Conditional Independence:

$$X_{Bart} \perp (X_{nonDesc} \backslash X_{Parent}) \mid X_{Parent}$$
Non-descendants

Any random variable is locally dependent on only its parent nodes, also known as the Markov Assumption.

- We use a directed acyclic graph (DAG) to represent conditional independence.
- A DAG is a pair $\mathcal{G}(\mathcal{V}, \mathcal{E})$, where \mathcal{V} is a set nodes and \mathcal{E} a set of oriented edges.

Example of a Directed Graphical Model (DGM), i.e.

Bayesian Network:

X₂

Observed variable

X₄

Variable

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Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.

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 X_3

 X_5

- G does not contain any cycles.
- Shaded node refers to observed variable.
- Unshaded node refers to latent/hidden (i.e. unobserved) variables.

Example of a Directed Graphical Model (DGM), i.e.

Bayesian Network:

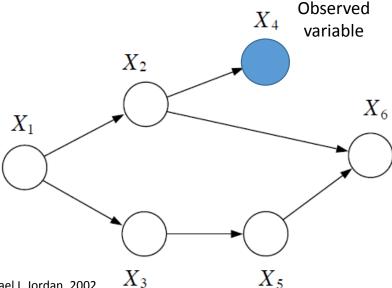
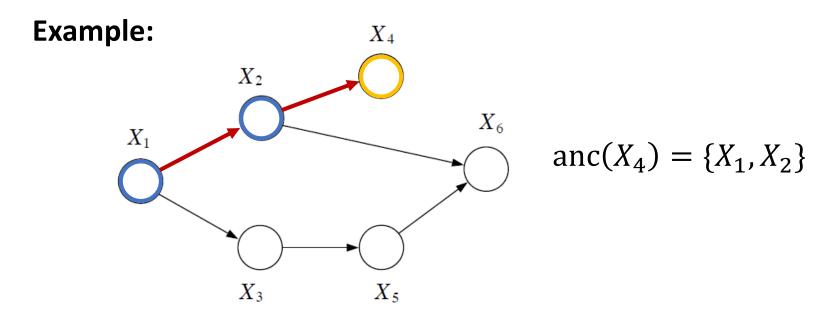


Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



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- Ancestors are the parents, grand-parents, etc of a node.
- The ancestors of t is the set of node s that connect to t via a trail: $anc(t) \triangleq \{s: s \rightsquigarrow t\}$.

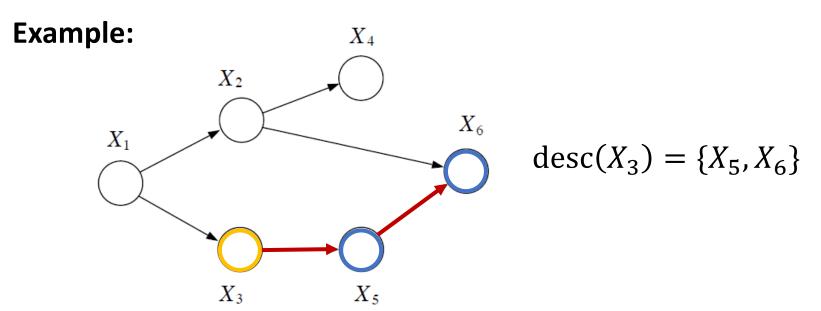


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16

- Descendants are the children, grand-children, etc of a node.
- The descendants of s is the set of nodes that can be reached via trials from s: $desc(s) \triangleq \{t: s \rightsquigarrow t\}$.





- There is an associated random variable X_i for each $i \in \mathcal{V}$.
- Each node $i \in \mathcal{V}$ has a set of parent nodes π_i , which can be the empty set.
- Let X_{π_i} represent all the random variables that are parents to the random variable X_i .

Example:

$$X_{\pi_1} = \emptyset$$
, $X_{\pi_2} = X_1$
 $X_{\pi_3} = X_1$, $X_{\pi_4} = X_2$
 $X_{\pi_5} = X_3$, $X_{\pi_6} = \{X_2, X_5\}$

 X_4 X_2 X_6 X_6 X_{10} X_{11} X_{12} X_{13} X_{14} X_{15} X_{15}

Image source: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.

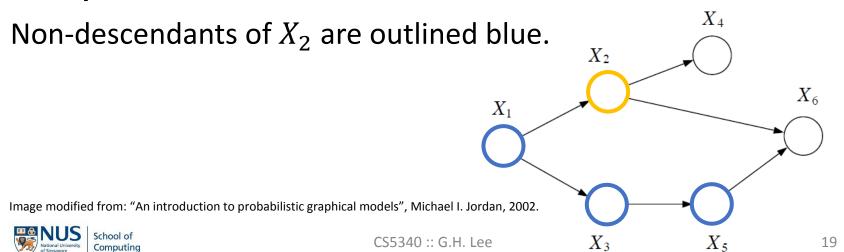


Markov Assumption

- Markov assumption: Each random variable X_i is independent of its non-descendants $X_{\text{nonDesc}(X_i)}$ given its parents X_{π_i} .
- The following set of basic conditional independence statements can be associated to the DGM:

$$\{X_i \perp (X_{\text{nonDesc}(X_i)} \setminus X_{\pi_i}) \mid X_{\pi_i}\}$$

Example:



Markov Assumption

Example:

We have the following set of basic conditional independence from the given Bayesian network.

$$X_{1} \perp \emptyset \mid \emptyset$$
, $X_{4} \perp \{X_{1}, X_{3}, X_{5}, X_{6}\} \mid X_{2}$, $X_{2} \perp \{X_{3}, X_{5}, \} \mid X_{1}$, $X_{5} \perp \{X_{1}, X_{2}, X_{4}\} \mid X_{3}$, $X_{3} \perp \{X_{2}, X_{4}\} \mid X_{1}$, $X_{4} \perp \{X_{1}, X_{3}, X_{4}\} \mid \{X_{2}, X_{5}\}$



Image source: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.

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Markov Assumption

 The basic conditional independence statements in the DGM give rise to a set of conditional probabilities:

$$\{X_i \perp (X_{\text{nonDesc}(x_i)} \setminus X_{\pi_i}) \mid X_{\pi_i}\}$$

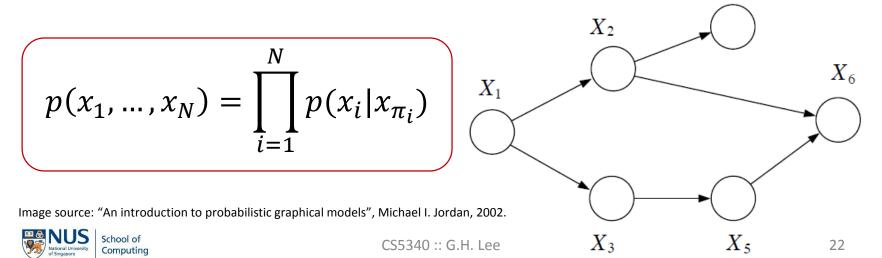
$$p(x_i | x_{\pi_i}), \qquad i = 1, ..., N$$

• $p(x_i|x_{\pi_i})$ is defined locally according to the parentchild relationship specified by the DGM.

- Locality of the parent-child relationship is used to construct economical representations of the joint distribution.
- The parent-child represents conditional independence:

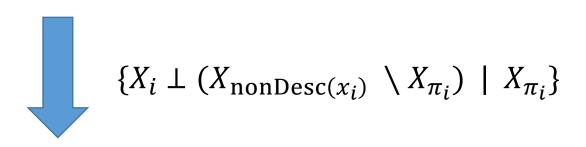
$$p(x_i|x_{\pi_i})$$

• Joint probability can be read off the graph as the product of all local conditional independence: X_4



Proof Sketch:

$$p(x_1, ..., x_N) = p(x_1) \prod_{i=2}^{N} p(x_i | x_{x_1, ..., x_{i-1}})$$
 (chain rule)

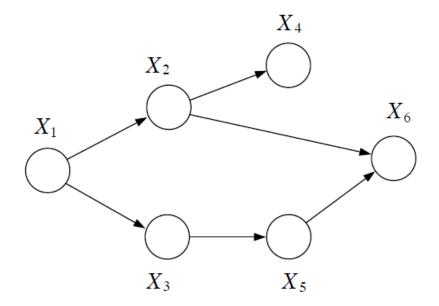


$$p(x_1, ..., x_N) = \prod_{i=1}^N p(x_i | x_{\pi_i})$$

(Assuming topological ordering of DGM)



Example:



$$p(x_1, \dots, x_N) = \prod_{i=1}^N p(x_i | x_{\pi_i})$$

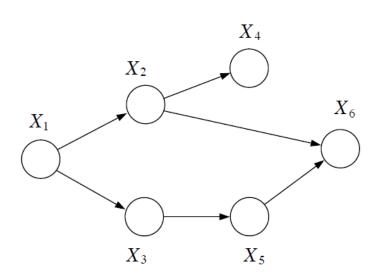
$$p(x_1, ..., x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1)p(x_4|x_2)p(x_5|x_3)p(x_6|x_2, x_5)$$



Example:

Let's verify that the basic sets of conditional independence are indeed represented in the joint probability:

$$p(x_1, ..., x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1)p(x_4|x_2)p(x_5|x_3)p(x_6|x_2, x_5)$$



$$X_1 \perp \emptyset \mid \emptyset,$$

 $X_2 \perp \{X_3, X_5, \} \mid X_1,$
 $X_3 \perp \{X_2, X_4\} \mid X_1,$
 $X_4 \perp \{X_1, X_3, X_5, X_6\} \mid X_2,$
 $X_5 \perp \{X_1, X_2, X_4\} \mid X_3,$
 $X_6 \perp \{X_1, X_3, X_4\} \mid \{X_2, X_5\}$

Example:

Let's verify that X_1 and X_3 are independent of X_4 given X_2 .

First we compute the marginal probability of $\{X_1, X_2, X_3, X_4\}$:

$$p(x_1, x_2, x_3, x_4) = \sum_{x_5} \sum_{x_6} p(x_1, x_2, x_3, x_4, x_5, x_6)$$

$$= \sum_{x_5} \sum_{x_6} p(x_1) p(x_2 \mid x_1) p(x_3 \mid x_1) p(x_4 \mid x_2) p(x_5 \mid x_3) p(x_6 \mid x_2, x_5)$$

$$= p(x_1) p(x_2 \mid x_1) p(x_3 \mid x_1) p(x_4 \mid x_2) \sum_{x_5} p(x_5 \mid x_3) \sum_{x_6} p(x_6 \mid x_2, x_5)$$

$$= p(x_1) p(x_2 \mid x_1) p(x_3 \mid x_1) p(x_4 \mid x_2),$$



Example:

Let's verify that X_1 and X_3 are independent of X_4 given X_2 .

Next we compute the marginal probability of $\{X_1, X_2, X_3\}$:

$$p(x_1, x_2, x_3) = \sum_{x_4} p(x_1) p(x_2 | x_1) p(x_3 | x_1) p(x_4 | x_2)$$
$$= p(x_1) p(x_2 | x_1) p(x_3 | x_1).$$

Dividing the two marginal yields the desired conditional:

$$p(x_4 | x_1, x_2, x_3) = p(x_4 | x_2),$$

Which demonstrates the conditional independence relationship $X_4 \perp \{X_1, X_3\} \mid X_2$.



Bayesian Networks: Parameter Reduction

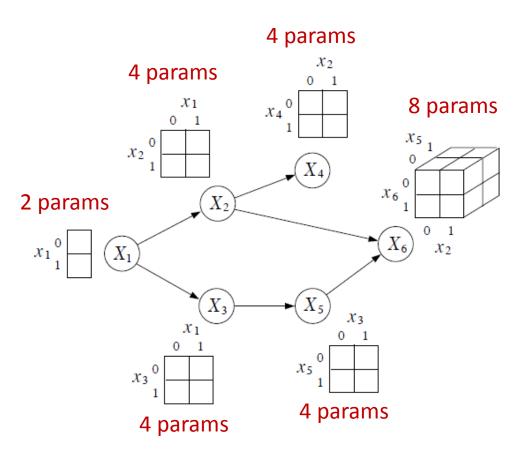
- Let m_i denote the number of parents of node X_i , and each node takes on K values.
- The conditional probability associated with X_i can be represented with a table of size K^{m_i+1} .
- Results in huge reduction of parameters needed to represent the joint probability, i.e. from $O(K^N)$ to $O(K^{m+1})$, $m \ll N$.



Bayesian Networks: Parameter Reduction

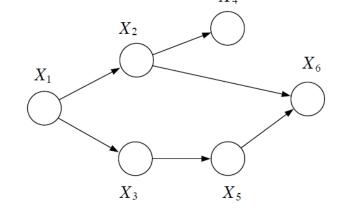
Example:

Binary state random variable $x_i \in \{0,1\}$.



- Total parameters = 26.
- Total parameters needed for fully dependent joint probability = $2^6 = 64$.
- More significant difference with higher number of nodes.

It turns out $X_1 \perp X_6 \mid \{X_2, X_3\}$ is also a conditional independence, but not directly observed from the parent-child relation.



Proof:

$$p(x_{1}, x_{2}, x_{3}, x_{6}) = \sum_{x_{4}} \sum_{x_{5}} p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) p(x_{4}|x_{2}) p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

$$= p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) \sum_{x_{4}} p(x_{4}|x_{2}) \sum_{x_{5}} p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

$$= p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) \sum_{x_{5}} p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

$$= p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) \sum_{x_{5}} p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

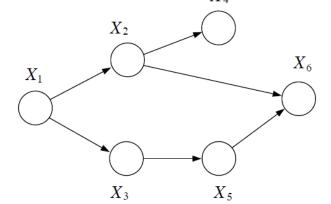
$$= p(x_{1}) p(x_{2}|x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) p(x_{4}|x_{2}) p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

$$= \sum_{x_{1}} p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) \sum_{x_{5}} p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$

$$= \sum_{x_{1}} p(x_{1}) p(x_{2}|x_{1}) p(x_{3}|x_{1}) \sum_{x_{5}} p(x_{5}|x_{3}) p(x_{6}|x_{2}, x_{5})$$



It turns out $X_1 \perp X_6 \mid \{X_2, X_3\}$ is also a conditional independence, but not directly observed from the parent-child relation.



Proof:

$$p(x_1, x_2, x_3, x_6) = p(x_1)p(x_2|x_1)p(x_3|x_1) \sum_{x_5} p(x_5|x_3)p(x_6|x_2, x_5)$$

$$p(x_2, x_3, x_6) = \sum_{x_1} p(x_1)p(x_2|x_1)p(x_3|x_1) \sum_{x_5} p(x_5|x_3)p(x_6|x_2, x_5)$$

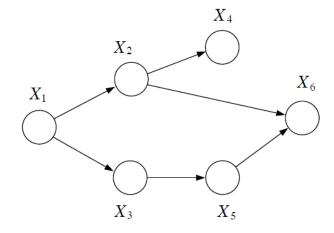
$$p(x_1|x_2, x_3, x_6) = \frac{p(x_1, x_2, x_3, x_6)}{p(x_2, x_3, x_6)}$$

$$= \frac{p(x_1)p(x_2|x_1)p(x_3|x_1)\sum_{x_5}p(x_5|x_3)p(x_6|x_2, x_5)}{\sum_{x_1}p(x_1)p(x_2|x_1)p(x_3|x_1)\sum_{x_5}p(x_5|x_3)p(x_6|x_2, x_5)}$$

$$= \frac{p(x_1, x_2, x_3)}{\sum_{x_1}p(x_1, x_2, x_3)} = \frac{p(x_1, x_2, x_3)}{p(x_2, x_3)} = p(x_1|x_2, x_3)$$



It turns out $X_1 \perp X_6 \mid \{X_2, X_3\}$ is also a conditional independence, but not directly observed from the parent-child relation.

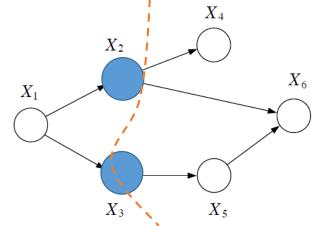


Proof:

$$p(x_1|x_2, x_3, x_6) = p(x_1|x_2, x_3)$$

Question: Can we write all other conditional independencies by just inspecting the DGM without going through the complicated mathematics?

It turns out $X_1 \perp X_6 \mid \{X_2, X_3\}$ is also a conditional independence, but not directly observed from the parent-child relation.



Proof:

$$p(x_1|x_2, x_3, x_6) = p(x_1|x_2, x_3)$$

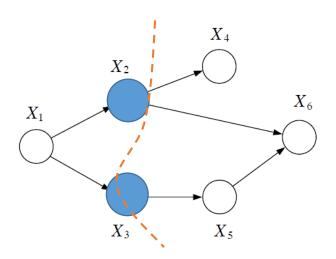
The nodes $\{X_2, X_3\}$ "block" X_1 from X_6 .

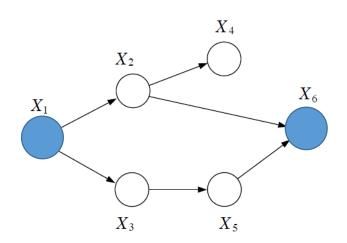
Question: Can we write all other conditional independencies by just inspecting the DGM without going through the complicated mathematics?

Answer: Yes, observe that the nodes $\{X_2, X_3\}$ "block" all paths from X_1 to X_6 . This suggests the notion of graph separation for inferring conditional independence.



- We have to be careful in making the notion of "blocking".
- For example, X_2 is NOT independent of X_3 given X_1 and X_6 as would be suggested by a naïve interpretation of "blocking".
- Precise definition of "blocking" has to be done through the "three canonical 3-node graphs", and "d-separation".





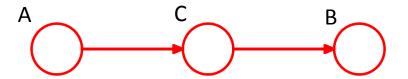
The nodes $\{X_2, X_3\}$ "block" X_1 from X_6 .

 X_2 is **NOT independent** of X_3 given $\{X_1, X_6\}$.



Three Canonical 3-Node Graphs

1.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a)p(c|a)p(b|c)$$

If none of the variables are observed, we can see that A and B are NOT independent by marginalizing over C:

$$p(a,b) = p(a) \sum_{c} p(c|a)p(b|c) = p(a)p(b|a)$$

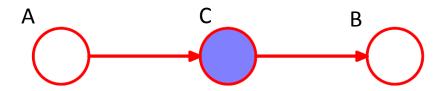
$$\sum_{c} p(c|a) p(b|c) = \sum_{c} \frac{p(a)p(c|a)p(b|c)}{p(a)} = \sum_{c} \frac{p(a,b,c)}{p(a)} = \frac{p(a,b)}{p(a)} = p(b|a)$$

which in general does not factorize into p(a)p(b), and so

$$A \perp B \mid \emptyset$$



Three Canonical 3-Node Graphs



If we condition on node C, using Bayes' theorem together with the joint distribution, we get:

$$p(a,b|c) = \frac{p(a,b,c)}{p(c)}$$

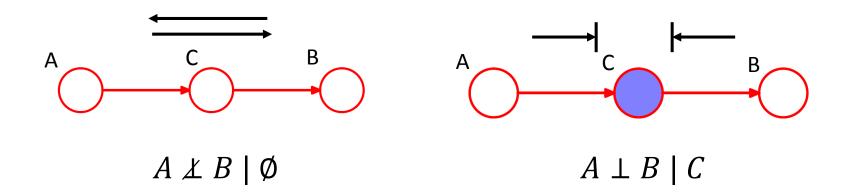
$$= \frac{p(a)p(c|a)p(b|c)}{p(c)}$$

$$= p(a|c)p(b|c)$$
(Bayes rule)
$$p(a)p(c|a) = p(a|c)p(c)$$

which shows the conditional independence property:

$$A \perp B \mid C$$

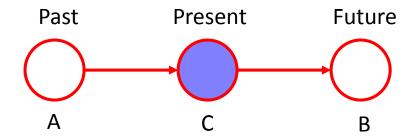




- The node C is said to be head-to-tail with respect to the path from node A to node B.
- Such a path connects nodes A and B and renders them dependent.
- The observation of *C* 'blocks' the path from *A* to *B* and so we obtain the conditional independence property.

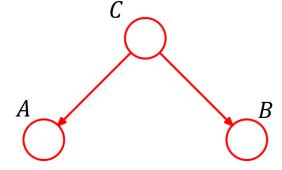


Intuitive interpretation:



- The conditional independence $A \perp B \mid C$ translates into the statement: "the past is independent of the future given the present".
- This is an example of a simple classical Markov Chain.

2.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a|c)p(b|c)p(c)$$

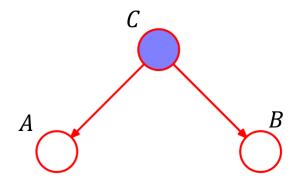
If none of the variables are observed, we can see that A and B are NOT independent by marginalizing both sides over C:

$$p(a,b) = \sum_{c} p(a|c)p(b|c)p(c)$$

which in general does not factorize into p(a)p(b), and so

$$A \perp B \mid \emptyset$$





If we condition on node C, we can easily write down the conditional distribution of A and B, given C, in the form:

$$p(a, b|c) = \frac{p(a, b, c)}{p(c)}$$

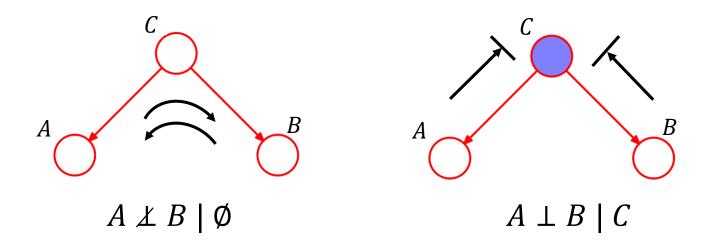
$$= p(a|c)p(b|c)$$

$$p(a, b, c) = p(a|c)p(b|c)$$

which shows the conditional independence property:

$$A \perp B \mid C$$

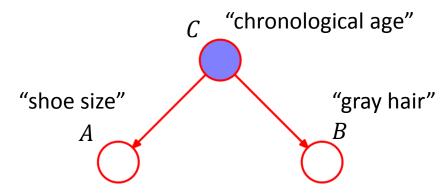




- The node C is said to be tail-to-tail with respect to the path from node A to B.
- Such a path connects nodes A and B and renders them dependent.
- The observation of C 'blocks' the path from A to B, and we obtain the conditional independence property.



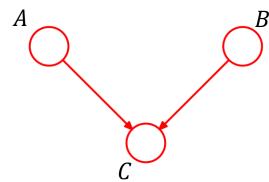
Intuitive interpretation:



- Given the age of a person, there is no further relationship between the size of his feet and the amount of gray hair he has.
- We say that the variable C "explains" all of the observed dependence between A and B.

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3.



Joint distribution corresponding to this graph:

$$p(a, b, c) = p(a)p(b)p(c|a, b)$$

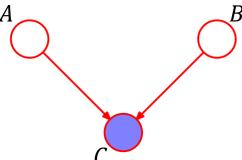
If none of the variables are observed, marginalizing both sides over *c* we obtain:

$$p(a,b) = \sum_{c} p(a)p(b)p(c \mid a,b)$$
$$= p(a)p(b)$$

A and B are independent with no variables observed, in contrast to the two cases:

$$A \perp B \mid \emptyset$$

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If we condition on node *C*, the conditional distribution of *A* and *B* is given by:

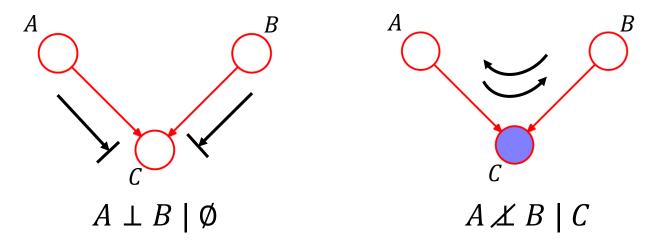
$$p(a, b|c) = \frac{p(a, b, c)}{p(c)}$$
$$= \frac{p(a)p(b)p(c|a, b)}{p(c)}$$

which in general does not factorize into the product p(a)p(b), and so

 $A \not\perp B \mid C$

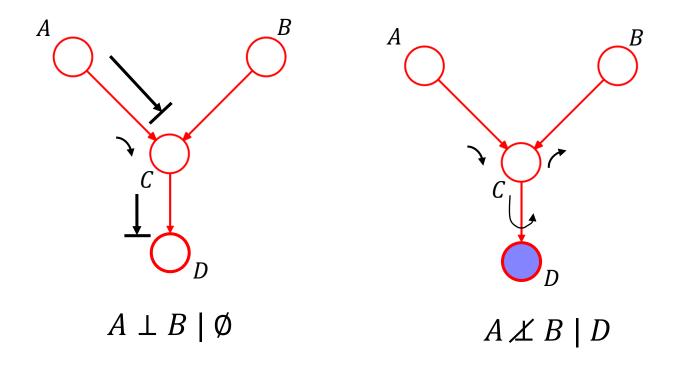


Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



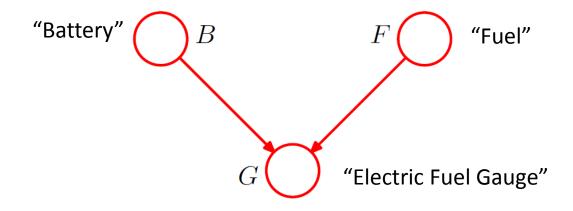
- Node C is head-to-head with respect to the path from A to B, also known as the "v-structure".
- When node C is unobserved, it "blocks" the path, and the variables A and B are independent.
- However, conditioning on C "unblocks" the path and renders
 A and B dependent.





 The observation of any descendent node of C "unblocks" the path from A to B.

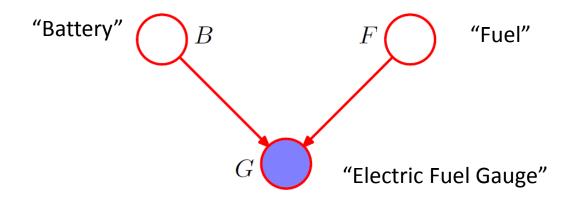
Intuitive interpretation:



- No relation between the battery and fuel status if fuel gauge is not read.
- This implies conditional independence of "battery" and "fuel" when "gauge" is not observed.

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop

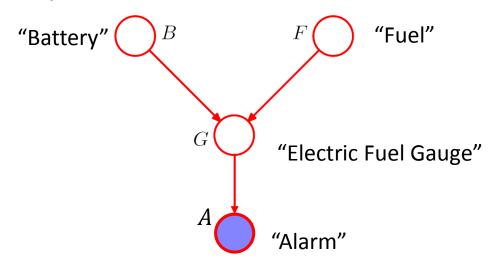
Intuitive interpretation:



- Suppose the fuel gauge shows "empty", knowing that the battery is flat lowers our belief that the fuel tank is empty.
- Battery and fuel status are now no longer independent.
- This is know as the "explaining-away" effect.

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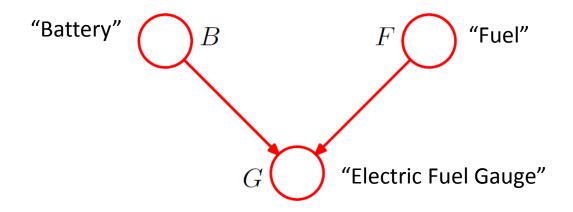
Intuitive interpretation:



- Alarm goes off when fuel gauge is empty.
- Suppose alarm goes off, we know that the fuel gauge shows "empty".
- Knowing that the is battery is flat lowers our belief that the fuel tank is empty, i.e. battery and fuel status are now no longer independent.



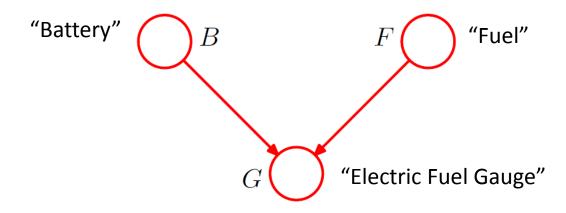
Numerical Example:



- B: battery state that is either charged (B = 1) or flat (B = 0).
- F: fuel tank state that is either full of fuel (F = 1) or empty (F = 0).
- G: electric fuel gauge state which indicates either full (G = 1) or empty (G = 0).

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Numerical Example:



Given:

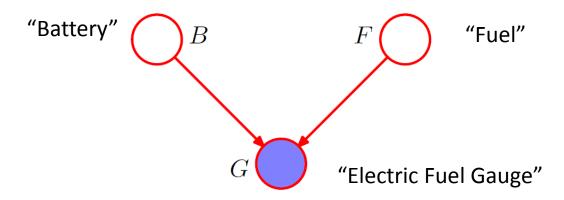
$$p(G = 1|B = 1, F = 1) = 0.8$$

 $p(B = 1) = 0.9$
 $p(G = 1|B = 1, F = 0) = 0.2$
 $p(G = 1|B = 0, F = 1) = 0.2$
 $p(G = 1|B = 0, F = 0) = 0.1$

Before we observe any data, the prior probability of the fuel tank being empty is p(F = 0) = 0.1.

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop

Numerical Example:



Suppose that we observe the fuel gauge and it reads empty, i.e., G = 0, we have:

$$p(F=0|G=0) = \frac{p(G=0|F=0)p(F=0)}{p(G=0)} \simeq 0.257$$

where

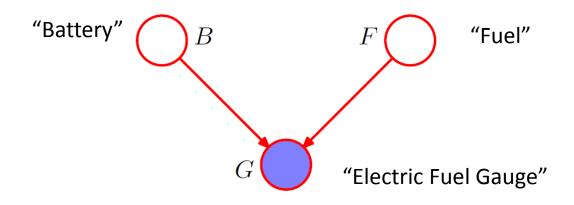
$$p(G=0) = \sum_{\mathbf{b} \in \{0,1\}} \sum_{\mathbf{f} \in \{0,1\}} p(G=0|B,F)p(B)p(F) = 0.315$$

$$p(G=0|F=0) = \sum_{\mathbf{b} \in \{0,1\}} p(G=0|B,F=0)p(B) = 0.81$$

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



Numerical Example:



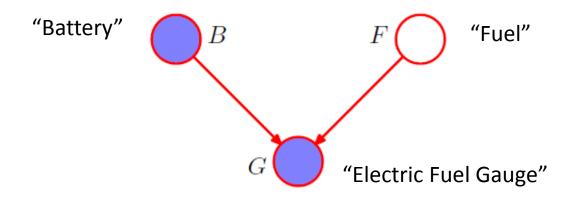
Hence,

0.257 0.1
$$p(F = 0|G = 0) > p(F = 0)$$

Observing that the gauge reads empty makes it more likely that the tank is indeed empty, as we would intuitively expect.



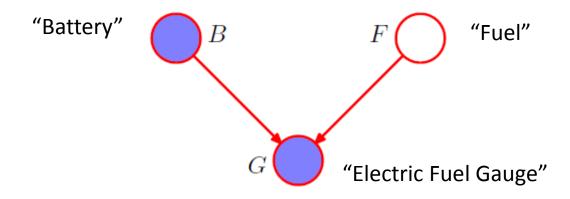
Numerical Example:



- If we also check the state of the battery and find that it is flat,
 i.e., B = 0.
- We have now observed the states of both fuel gauge and battery.

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Numerical Example:

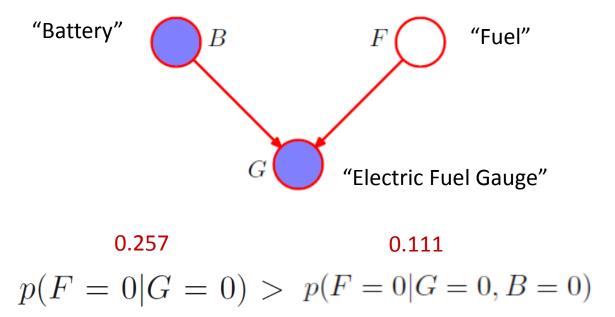


 Posterior probability that fuel tank is empty given observations of both fuel gauge and battery state is:

$$p(F=0|G=0,B=0) = \frac{p(G=0|B=0,F=0)p(F=0)}{\sum_{F\in\{0,1\}} p(G=0|B=0,F)p(F)} \simeq 0.111$$

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Numerical Example:



• Finding out that battery is flat *explains away* observation that the fuel gauge reads empty!

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Graph Separation

- We have seen earlier that $A \perp B \mid C$ if all paths from nodes in set A are "blocked" from nodes in set B when all nodes from set C are observed.
- A is said to be d-separated from B by C, and the joint distribution over all of the variables in the graph will satisfy $A \perp B \mid C$.



Graph Separation

- From the three canonical 3-node graphs, any path is said to be "blocked" / d-separated if it includes a node such that either:
 - a) The arrows on the path meet either head-to-tail or tail-to-tail at the node, and the node is in the set C, or
 - b) The arrows meet head-to-head at the node, and neither the node, nor any of its descendants, is in the set *C*.



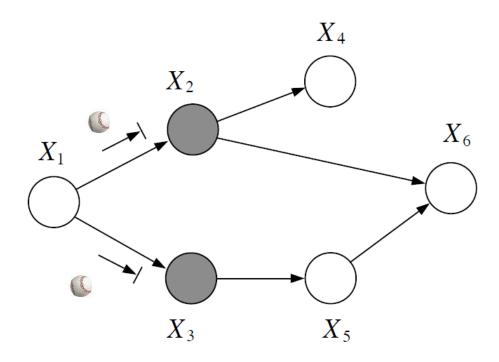
- This is a "reachability" algorithm:
 - 1. Shade the nodes in set C.
 - 2. Place a ball at each of the nodes in set A.
 - Let the balls bounce around the graph according to the d-separation rules:

```
IF none of the balls reach B THEN A \perp B \mid C, ELSE A \not \perp B \mid C,
```

Can be implemented as a breadth-first search.



Example 1:

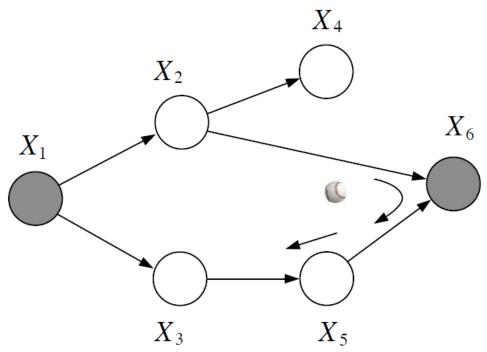


• A ball cannot pass through X_2 to X_6 nor through X_3 i.e. $X_1 \perp X_6 \mid \{X_2, X_3\}$.

Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



Example 2:



• A ball can pass through X_2 to X_6 through X_5 and X_3 i.e. $X_2 \not L$ $X_3 \mid \{X_1, X_6\}$.

Image modified from: "An introduction to probabilistic graphical models", Michael I. Jordan, 2002.



Example 3: Naïve Bayes Classifier

- Classify e-mails as spam (Y = 1) or not spam (Y = 0).
- Let 1: N index the words in our vocabulary (e.g., English)
- $X_n = \begin{cases} 0 & \text{word } n \text{ appears in an e-mail} \\ 1 & \text{otherwise} \end{cases}$
- E-mails are drawn according to some distribution $p(x_1,...,x_N,Y)$



Example 3: Naïve Bayes Classifier

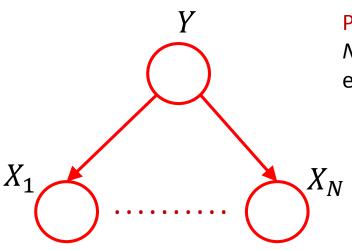
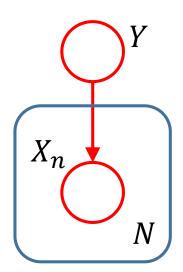


Plate notation that represents N nodes of which only a single example X_n is shown explicitly

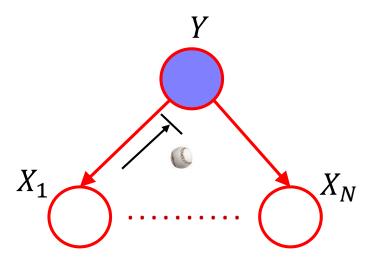




Joint distribution:

$$p(x_1, ... x_N, Y) = p(x_1, ... x_N | Y) p(Y)$$

Example 3: Naïve Bayes Classifier



Joint distribution:

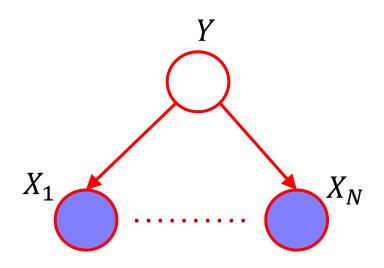
$$p(x_1, ... x_N, Y) = p(x_1, ... x_N | Y) p(Y)$$

$$= p(x_1 | Y) ... p(x_N | Y) p(Y)$$

$$= \prod_{n=1}^{N} p(x_n | Y) p(Y)$$



Example 3: Naïve Bayes Classifier



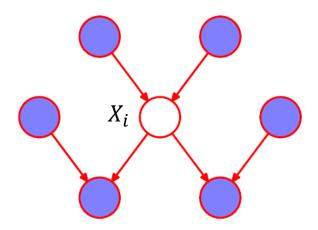
Joint distribution:

$$p(x_1, ... x_N, Y) = \prod_{n=1}^N p(x_n | Y) p(Y)$$

Predict with Bayes rule:

$$p(Y = 1 \mid x_1, ..., x_N) = \frac{p(y = 1) \prod_{n=1}^{N} p(x_n \mid Y = 1)}{\sum_{y = \{0,1\}} p(Y = y) \prod_{n=1}^{N} p(x_n \mid Y = 1)}$$

Markov Blanket



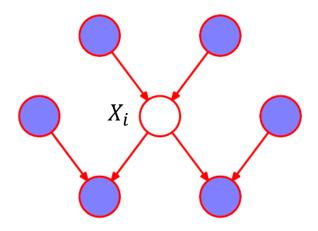
- The Markov blanket of a node X_i comprises the set of parents, children and co-parents of the node.
- Conditional distribution of X_i , conditioned on all the remaining variables in the graph, is dependent only on the variables in the Markov blanket.

 $Image\ Source:\ ``Pattern\ Recognition\ and\ Machine\ Learning'',\ Christopher\ Bishop$

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Markov Blanket

Proof:



$$p(x_{i}|x_{\{j\neq i\}}) = \frac{p(x_{1},...,x_{D})}{\int p(x_{1},...,x_{D})dx_{i}} = \frac{\prod_{k} p(x_{k}|x_{\pi_{k}})}{\int \prod_{k} p(x_{k}|x_{\pi_{k}})dx_{i}}$$

$$= \frac{\prod_{l\neq\{m,i\}} p(x_{l}|x_{\pi_{l}}) \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})}{\prod_{l\neq\{m,i\}} p(x_{l}|x_{\pi_{l}}) \int \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})dx_{i}}$$

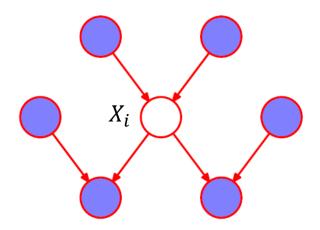
$$= \frac{\prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})}{\int \prod_{m} p(x_{m}|x_{\pi_{m}}:x_{i}\in x_{\pi_{m}}) p(x_{i}|x_{\pi_{i}})dx_{i}}$$

Image Source: "Pattern Recognition and Machine Learning", Christopher Bishop



Markov Blanket

Proof:



$$p(x_i|x_{\{j\neq i\}}) = \frac{\prod_m p(x_m|x_{\pi_m}: x_i \in x_{\pi_m}) p(x_i|x_{\pi_i})}{\int \prod_m p(x_m|x_{\pi_m}: x_i \in x_{\pi_m}) p(x_i|x_{\pi_i}) dx_i}$$

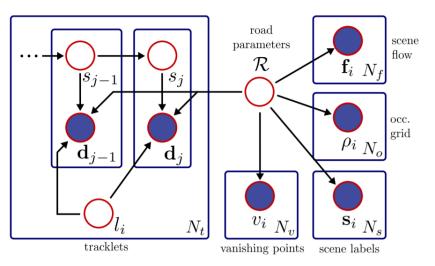
$$\text{children and} \quad \text{parents of } X_i$$

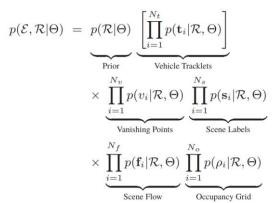
$$\text{co-parents of } X_i$$

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Examples of Bayesian Network

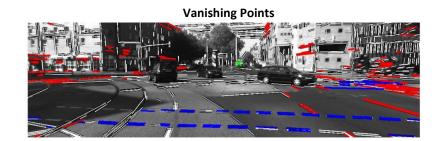
Urban Scene Understanding:



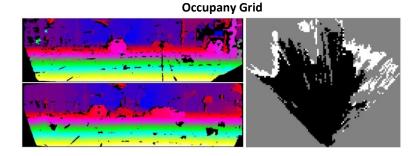


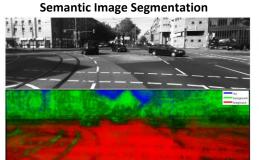
Andreas Geiger, PhD Thesis 2013

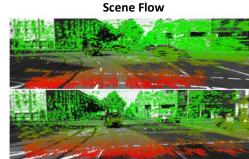




Tracklets
4 / 40
4 / 80



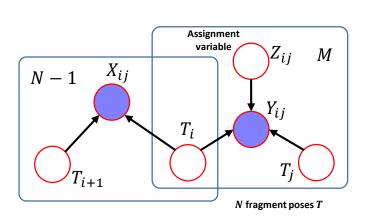


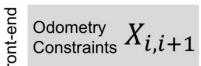


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Examples of Bayesian Network

Robust 3D Point Cloud Reconstruction:





Loop-Closure Y_{ij}

The state of the s

back-end

Robust Reconstruction with EM



Expectation step:

$$Q_{EM} = \sum_{Z} p(Z \mid Y, T^{old}) \ln p(X, Y, Z \mid T)$$

Maximization step:

$$\underset{T}{\operatorname{argmax}} Q_{EM}$$





Ziquan Lan, Zi Jian Yew, Gim Hee Lee, "Robust Point Cloud Based Reconstruction of Large-Scale Outdoor Scenes", CVPR 2019



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Summary

- You have learned how to:
 - 1. Explain the concepts of conditional independence.
 - 2. Use the Bayesian network to represent conditional independence in joint distributions.
 - Describe d-separation using the three canonical 3node graph.
 - Deduce all conditional independence in a Bayesian network using the Bayes ball algorithm.
 - 5. Explain the concepts of Markov Blanket.

