

HW1

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Problem 1

1.

```
set.seed(0)

# Define full sample size and simulate feature x
n.full <- 100
x <- runif(n.full,0,10)

# Define linear model as y=1+x+error
y <- 1+0.5*x + rnorm(n.full)

data.full <- data.frame(x=x,y=y)
#plot(x,y)

# Split data
n.test <- 50
index.test <- sample(1:n.full,n.test)
data.train <- data.full[-index.test,]
data.test <- data.full[index.test,]

# Fit linear, quadratic and cubic models
linear.train <- lm(y~x,data=data.train)
quad.train <- lm(y~poly(x,degree = 2,row=T),data=data.train)
cubic.train<- lm(y~poly(x,degree = 3,row=T),data=data.train)

# Fit linear, quadratic and cubic models
linear.predict.train <- predict(linear.train,newdata=data.train)
quad.predict.train <- predict(quad.train,newdata=data.train)
cubic.predict.train <- predict(cubic.train,newdata=data.train)

# Training error and test error
train.err.linear = mean((data.train$y-linear.predict.train)^2)
train.err.linear

## [1] 0.8279371

train.err.quad = mean((data.train$y-quad.predict.train)^2)
train.err.quad

## [1] 0.8171223

train.err.cubic = mean((data.train$y-cubic.predict.train)^2)
train.err.cubic
```

```
## [1] 0.8168094
```

2.

```
linear.predict.test <- predict(linear.train,newdata=data.test)
quad.predict.test <- predict(quad.train,newdata=data.test)
cubic.predict.test <- predict(cubic.train,newdata=data.test)

test.err.linear = mean((data.test$y-linear.predict.test)^2)
test.err.quad = mean((data.test$y-quad.predict.test)^2)
test.err.cubic = mean((data.test$y-cubic.predict.test)^2)
test.err.linear
```

```
## [1] 0.731376
```

```
test.err.quad
```

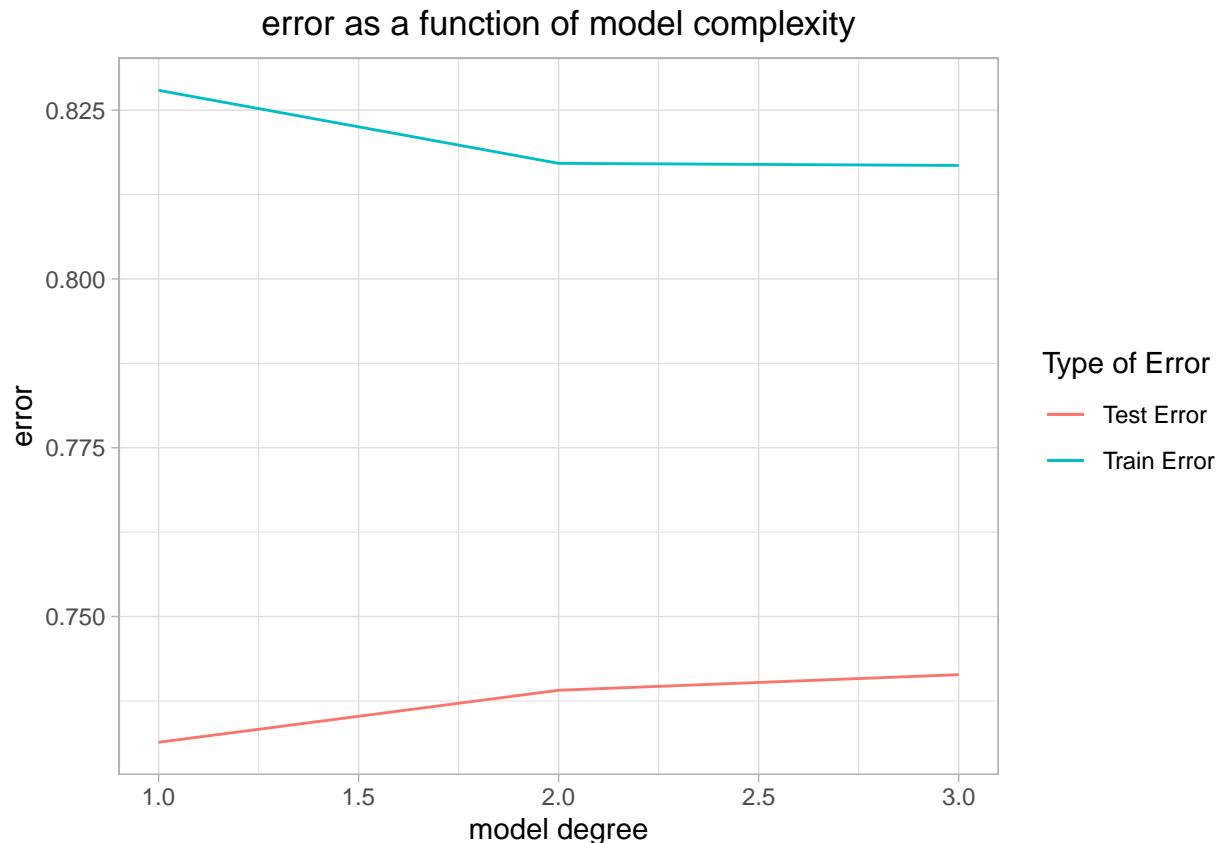
```
## [1] 0.7390816
```

```
test.err.cubic
```

```
## [1] 0.7413943
```

3.

```
model = c(1,2,3)
train.err = c(train.err.linear,train.err.quad,train.err.cubic)
test.err = c(test.err.linear,test.err.quad,test.err.cubic)
library(ggplot2)
err.plot = ggplot()+geom_line(aes(model,train.err,color = 'red'))+geom_line(aes(model,test.err,color = 'blue'))+
  scale_color_discrete(name = "Type of Error" ,labels = c("Test Error","Train Error"))+
  labs(title = "error as a function of model complexity", x = "model degree", y = "error")+
  theme_light()+
  theme(plot.title = element_text(hjust = 0.5))
err.plot
```



Problem 2

1. Maximum Likelihood Estimator: the MLE is the parameter which maximizes the likelihood (joint density) of the data i.e. $\hat{\theta}_{ML} = \arg \max_{\theta} l(\theta)$, the likelihood function is $l(\theta) = p(x|\theta) = \prod_{i=1}^n p(x_i|\theta)$. Since $\arg \max_y \log(f(y)) = \arg \max_y f(y)$, and by logarithm trick $\log(\prod_i f_i) = \sum_i \log(f_i)$, we can get $\hat{\theta}_{ML} = \arg \max_{\theta} \sum_{i=1}^n \log p(x_i|\theta)$. To get the maxima, the parameter must meet the maximality criterion $0 = \sum_{i=1}^n \nabla_{\theta} \log p(x_i|\theta)$
2. derive the ML estimator for μ

$$\begin{aligned}
 0 &= \sum_{i=1}^n \nabla_{\mu} \ln \left(\left(\frac{\gamma}{\mu} \right)^{\gamma} \frac{x^{\gamma-1}}{\Gamma(\gamma)} \exp \left(-\frac{\gamma x}{\mu} \right) \right) \\
 0 &= \sum_{i=1}^n \nabla_{\mu} \left(\gamma \ln \left(\frac{\gamma}{\mu} \right) + (\gamma - 1) \ln(x_i) - \ln \Gamma(\gamma) - \frac{\gamma}{\mu} x_i \right) \\
 \sum_{i=1}^n \left(-\frac{\gamma}{\mu} + \frac{\gamma}{\mu^2} x_i \right) &= 0 \\
 -\frac{n\gamma}{\mu} + \frac{\gamma}{\mu^2} \sum_{i=1}^n x_i &= 0 \\
 \hat{\mu} &= \frac{1}{n} \sum_{i=1}^n x_i
 \end{aligned}$$

3. instead deriving the formula, show equation to get $\hat{\gamma}$

$$\begin{aligned}
 0 &= \sum_{i=1}^n \nabla_{\gamma} \ln \left(\left(\frac{\gamma}{\mu} \right)^{\gamma} \frac{x^{\gamma-1}}{\Gamma(\gamma)} \exp \left(-\frac{\gamma x}{\mu} \right) \right) \\
 0 &= \sum_{i=1}^n \nabla_{\gamma} \left(\gamma \ln \left(\frac{\gamma}{\mu} \right) + (\gamma - 1) \ln(x_i) - \ln \Gamma(\gamma) - \frac{\gamma}{\mu} x_i \right)
 \end{aligned}$$

$$\begin{aligned}\sum_{i=1}^n (\ln(\frac{\gamma}{\mu}) + \gamma * \frac{\mu}{\gamma} * \frac{1}{\mu} + \ln(x_i) - \phi(\gamma) - \frac{x_i}{\mu}) &= 0 \\ \sum_{i=1}^n (\ln(\frac{x_i \hat{\gamma}}{\mu}) + 1 - \phi(\gamma) - \frac{x_i}{\mu}) &= 0 \\ \sum_{i=1}^n (\ln(\frac{x_i \hat{\gamma}}{\mu}) - (\frac{x_i}{\mu} - 1) - \phi(\gamma)) &= 0\end{aligned}$$

Problem 3

we want to show that f_0 defined by $f_0 = \arg \max_{y \in [K]} P(y|x)$ minimize $R(f)$

$$R(f|x) := \sum_{y \in [K]} L^{0-1}(y, f(x)) P(y|x) \text{ and } R(f) = \int_{R^d} R(f|x) p(x) dx$$

if f_0 minimize conditional risks $R(f|x)$, then this f_0 minimize $R(f)$ as well

$$\begin{aligned}R(f|x) &:= \sum_{y \in [K]} L^{0-1}(y, f(x)) P(y|x) \\ &= P(y = f(x)|x) \times L^{0-1}(f(x), f(x)) + \sum_{k \neq f(x)} P(k|x) * L^{0-1}(f(x), k) \\ &= \sum_{k \neq f(x)} P(k|x)\end{aligned}$$

since $P(k|x)$ sums to 1 over all k , so, $R(f|x) = 1 - P(f(x)|x)$, that is, f_0 maximizes $P(f(x)|x)$, hence minimize $1 - P(f(x)|x)$ at each x .

In summary, $f_0 = \arg \min_{f \in H} R(f|x)$ for all $x \in R^d$, and therefore $f_0 = \arg \min_{f \in H} R(f)$

Problem 4

1. Derive the posterior

$$\prod(\theta_1, \dots, \theta_K | x_1, \dots, x_n) = \frac{\text{likelihood} * \text{prior}}{\text{evidence}} = a * \prod_{k=1}^K \theta_k^{n_k + \alpha_k - 1}$$

where a is the normalizing constant

2. Derive the Bayesian MAP

$$\hat{\theta}_{MAP} = \arg \max_{\theta} \prod(\theta_1, \dots, \theta_K | x_1, \dots, x_n) \text{ where } \theta = (\theta_1, \dots, \theta_K)$$

drop the normalizing constant, and by logarithm trick, $\hat{\theta}_{MAP} = \arg \max_{\theta} \sum_{k=1}^K (n_k + \alpha_k - 1) \log(\theta_k)$

by maximizing criterion: $0 = \nabla_{\theta_k} \sum_{k=1}^K (n_k + \alpha_k - 1) \log(\theta_k)$

by lagrangian multiplier, we want to maximize function $f(x)$ with constraint $g(x) = a$, we optimise $f(x) - \lambda(g(x) - a)$, therefore, in the case of Dirichlet, since $\sum_{k=1}^K \theta_k = 1$, thus

$$\begin{aligned}f(\theta) &= \log(L(\theta)) = \sum_{k=1}^K (n_k + \alpha_k - 1) \log(\theta_k) \\ g(\theta) &= \sum_{k=1}^K \theta_k, a = 1\end{aligned}$$

optimise the lagrangian $\nabla_{\theta_k} \sum_{k=1}^K (n_k + \alpha_k - 1) \log(\theta_k) - \lambda * (\sum_{k=1}^K \theta_k - 1) = 0$

$$\hat{\theta}_k = \frac{n_k + \alpha_k - 1}{\lambda}$$

replace the parameters with their estimates and optimise the lagrangian for lambda,

$$\begin{aligned}
0 &= \nabla_{\lambda} \sum_{k=1}^K (n_k + \alpha_k - 1) \log\left(\frac{n_k + \alpha_k - 1}{\lambda}\right) - \sum_{k=1}^K (n_k + \alpha_k - 1) + \lambda \\
&\sum_{k=1}^K (n_k + \alpha_k - 1) \times \frac{\lambda}{n_k + \alpha_k - 1} \times \left(-\frac{n_k + \alpha_k - 1}{\lambda^2}\right) + \lambda = 0 \\
&-\sum_{k=1}^K \frac{n_k + \alpha_k - 1}{\lambda} + 1 = 0 \\
\lambda &= \sum_{k=1}^K (n_k + \alpha_k - 1) \\
\hat{\theta}_{kMAP} &= \frac{n_k + \alpha_k - 1}{\sum_{k=1}^K (n_k + \alpha_k - 1)}
\end{aligned}$$

3. Derive the “frequentist” MLE of parameters

Similar to above, $\hat{\theta}_{ML} = \arg \max_{\theta} l(\theta) = \arg \max_{\theta} \prod_{k=1}^K \theta_k^{n_k}$, therefore to get MLE, $0 = \nabla_{\theta_k} \sum_{k=1}^K n_k \log(\theta_k)$

by lagrangian multiplier technique, $\hat{\theta}_{kML} = \frac{n_k}{\sum_{k=1}^K n_k}$