

# **An introduction to Longitudinal Modified Treatment Policies**

**A framework for defining, identifying, and estimating  
complex, continuous, and/or time-varying exposures**

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**TGIF**

**November 22, 2024**



# An introduction to LMTPs

## References and Caveats

- This talk will be based on,  
  
Hoffman K.L., Salazar D., Williams, N., Rudolph, K., Díaz, I. Studying continuous, time-varying, and/or complex exposures using longitudinal modified treatment policies. *Epidemiology* (2024)
- ...which is itself based on,  
  
Díaz, I., Williams, N., Hoffman, K. L., & Schenck, E. J. Nonparametric Causal Effects Based on Longitudinal Modified Treatment Policies. *Journal of the American Statistical Association* (2021)
- *TGIF caveat*: my contributions to LMTP to date are practical, not theoretical.

# An introduction to LMTPs

## Today's talk

1. **Motivation**
2. **Setup**
3. **Defining an LMTP intervention**
4. **Identification assumptions**
5. **Estimation procedures**
6. **Extensions and applications**

# Motivation

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  - Summarizing the infinite-dimensional dose-response curve, e.g. through a Marginal Structural Model, typically requires restrictive and arbitrary **parametric assumptions**



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*(Neugebauer and van der Laan, 2007)*

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  - Summarizing the infinite-dimensional dose-response curve, e.g. through a Marginal Structural Model, typically requires restrictive and arbitrary **parametric assumptions**
    - **Difficult to interpret under model misspecification**  
*(Neugebauer and van der Laan, 2007)*
  - Non-parametric estimation approaches of the dose-response cannot achieve  $n^{1/2}$ -consistency, because it is not a pathwise-differentiable parameter

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  - Multiple time points make a positivity violation more likely to occur
- Ignoring a time-varying exposure mechanism for practical purposes can cause serious time-alignment biases
  - Frequently done in practice, e.g. treat all units receiving exposure within a certain time interval from baseline as “treated,” otherwise “untreated”

# LMTP's Methodological contribution

## Extending MTPs to the time-varying setting

- Formally defining LMTP **estimands**
- **Identification** assumptions
- **Estimators** — two are efficient and one is *sequentially doubly robust*
- Open source **R package {lmtp}**
  - Author/maintainer: Nick Williams

# Setup

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## Notation

- $Z_1, \dots, Z_t$  : a sample of i.i.d observations,  
 $Z = (L_1, A_1, L_2, A_2, \dots, L_\tau, A_\tau, Y) \sim P$
- $A_t$  : intervention variables, e.g. treatment and/or censoring status
- $L_t$  : time-varying covariates
- $Y = L_{\tau+1}$  : outcome at the end of follow-up

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- History of one variable,  $\bar{X}_t = (X_1, \dots, X_t)$
- Future of one variable,  $\underline{X}_t = (X_t, \dots, X_\tau)$
- $H_t = (\bar{A}_{t-1}, \bar{L}_t)$  — history of all variables up until just before  $A_t$
- Parentheses denote counterfactual variables

# Setup

## Non-parametric structural equation model (Pearl, 2009)

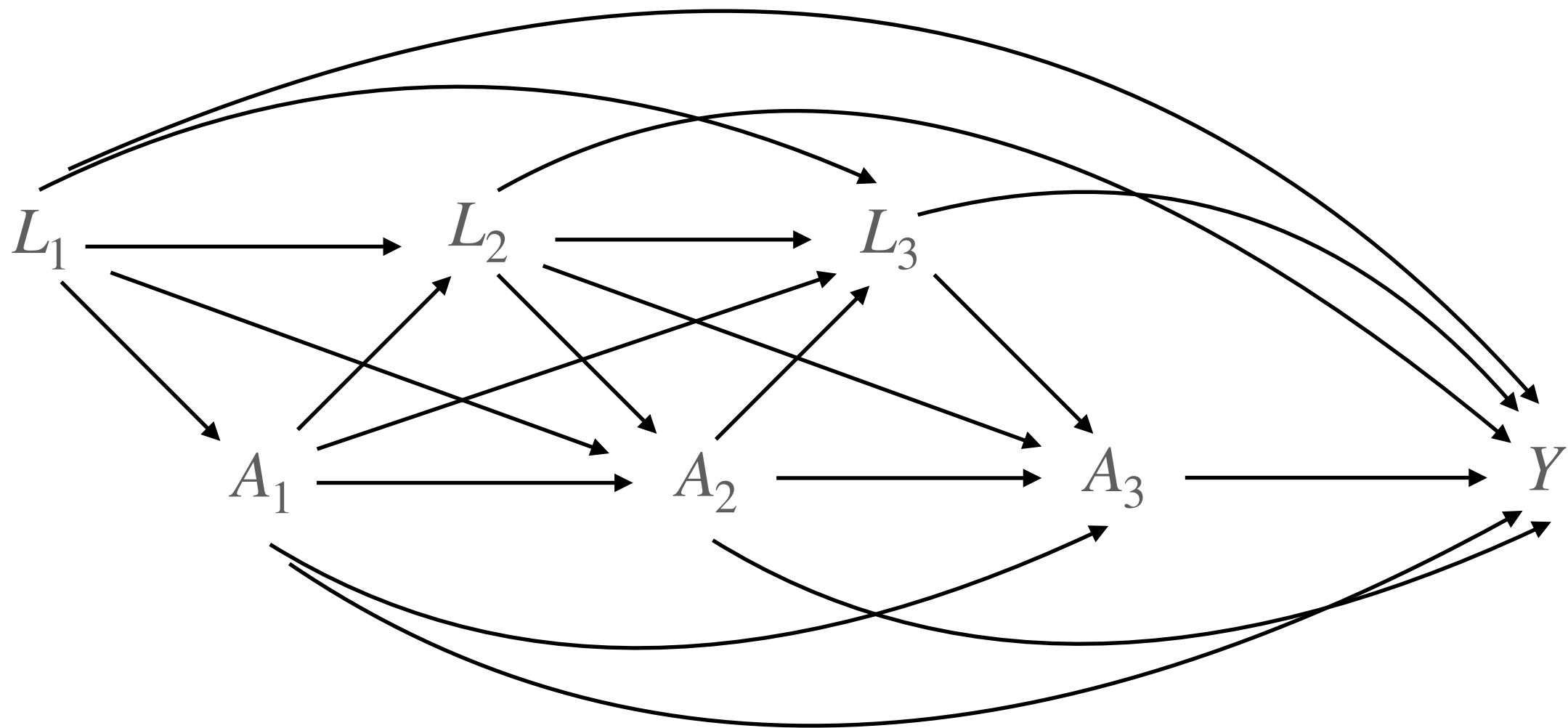
- $Z_1, \dots, Z_t$  i.i.d observations  
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- Assume the existence of deterministic functions  $f_{L_t}, f_{A_t}, f_Y$  s.t.
  - $L_t = f_{L_t}(A_{t-1}, H_{t-1}, U_{L,t})$
  - $A_t = f_{A_t}(H_t, U_{A,t})$
  - $Y = f_Y(A_\tau, H_\tau, U_Y)$
- where  $U = (U_{L,t}, U_{A,t}, U_Y : t \in \{1, \dots, \tau\})$  are a vector of exogenous variables with unrestricted joint distribution

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## Directed Acyclic Graph (Pearl, 2009)

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## Defining an LMTP intervention

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- Consider a user-defined intervention  $d$
- **LMTP effects** defined as **hypothetical interventions** where we **replace**

$$A_t = f_{A_t}(H_t, U_{A,t})$$

in the SEM with a **new random variable**,

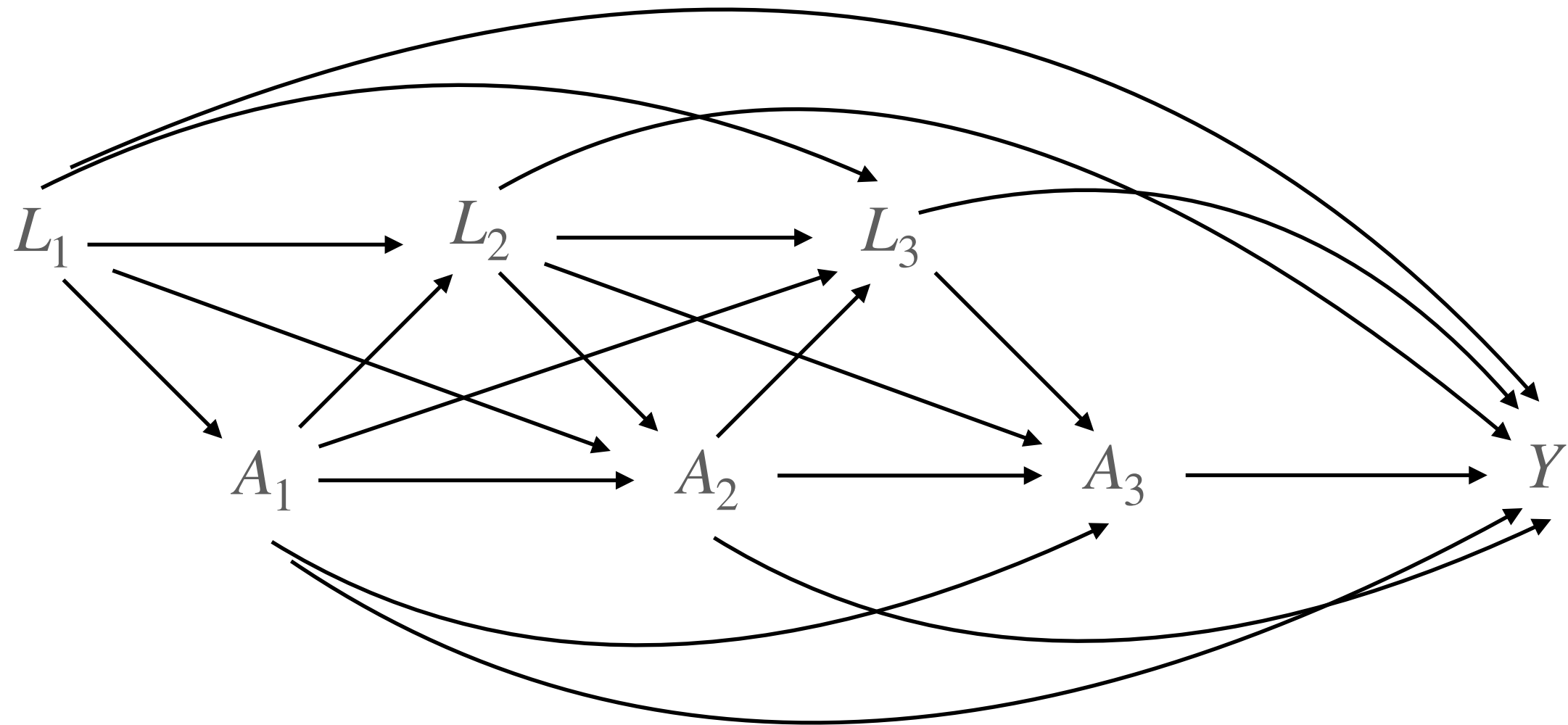
$$A_t^d$$

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*Intuition based on SWIGS*

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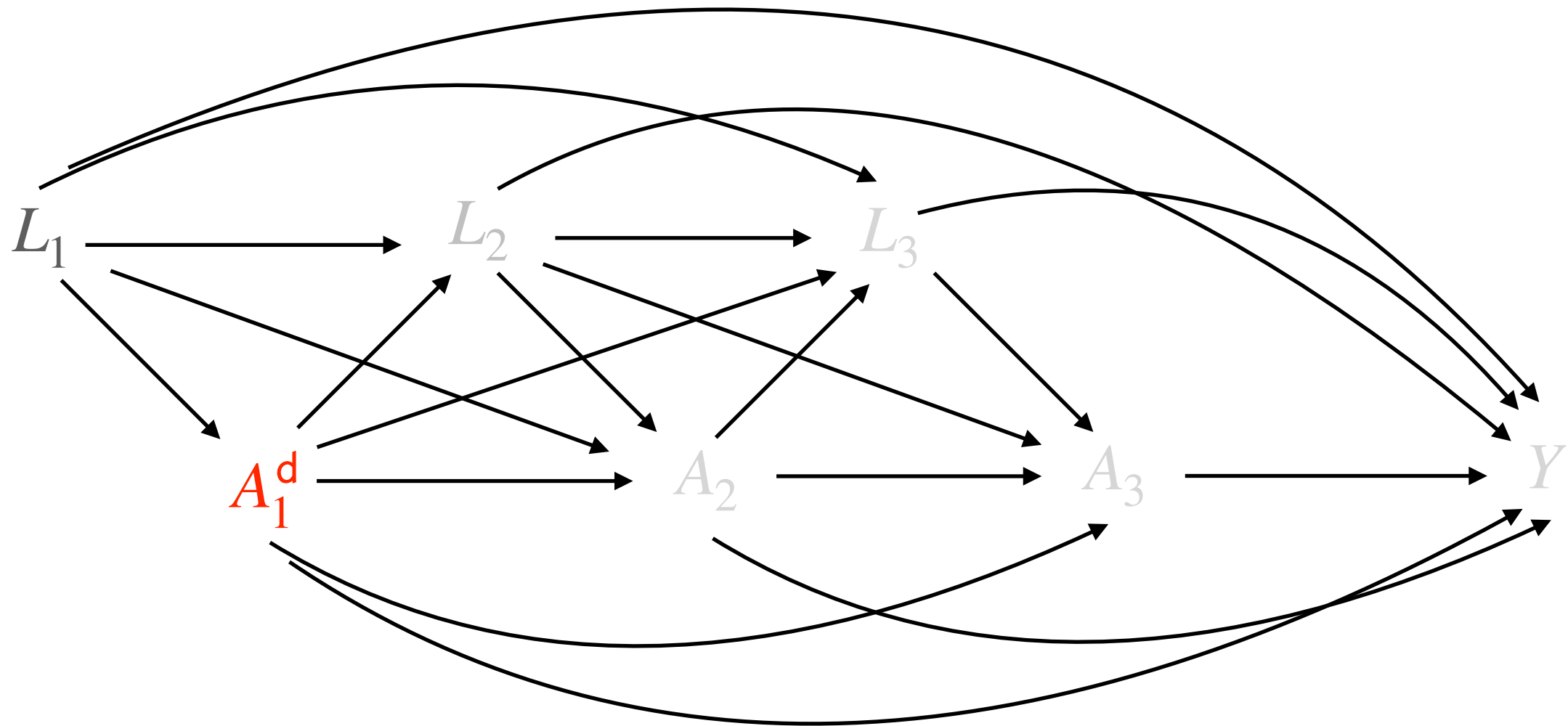


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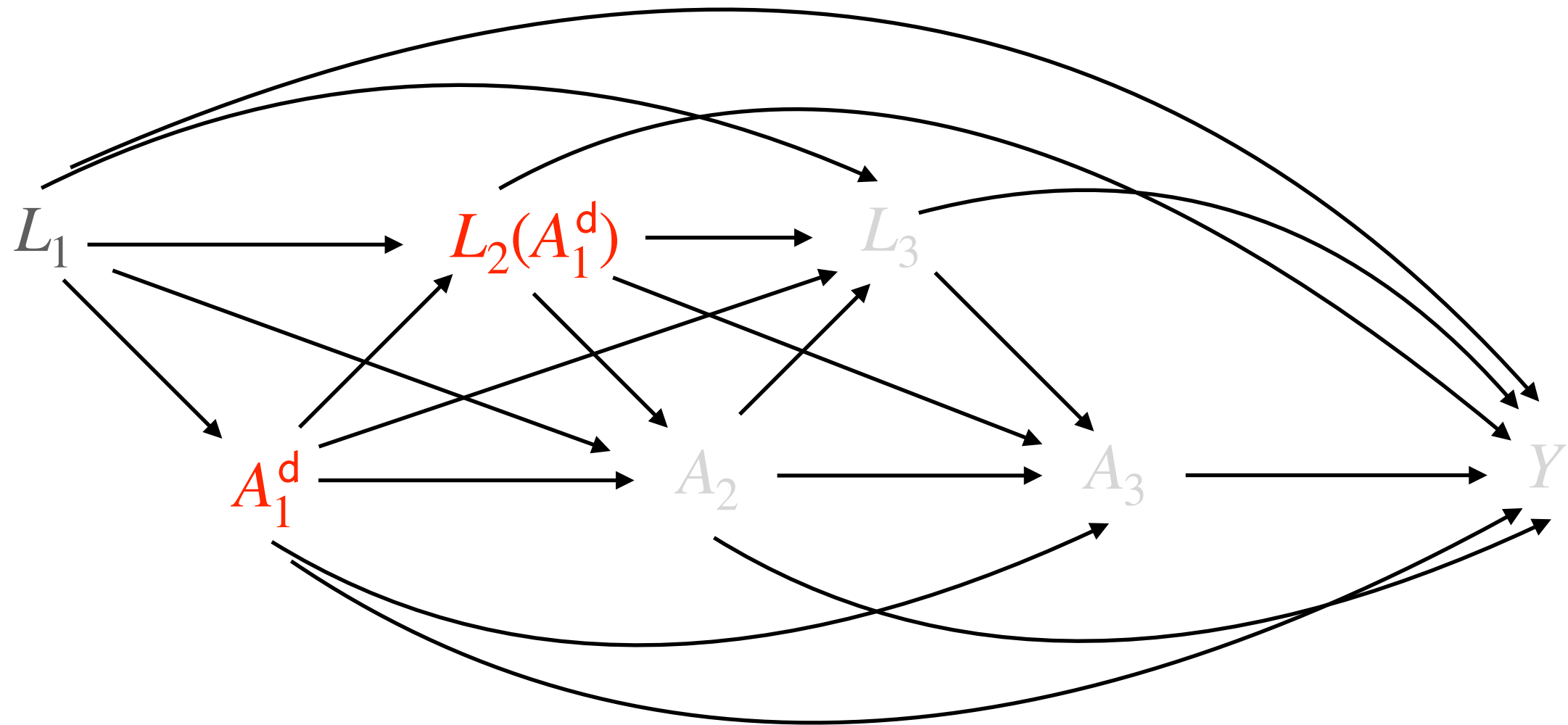


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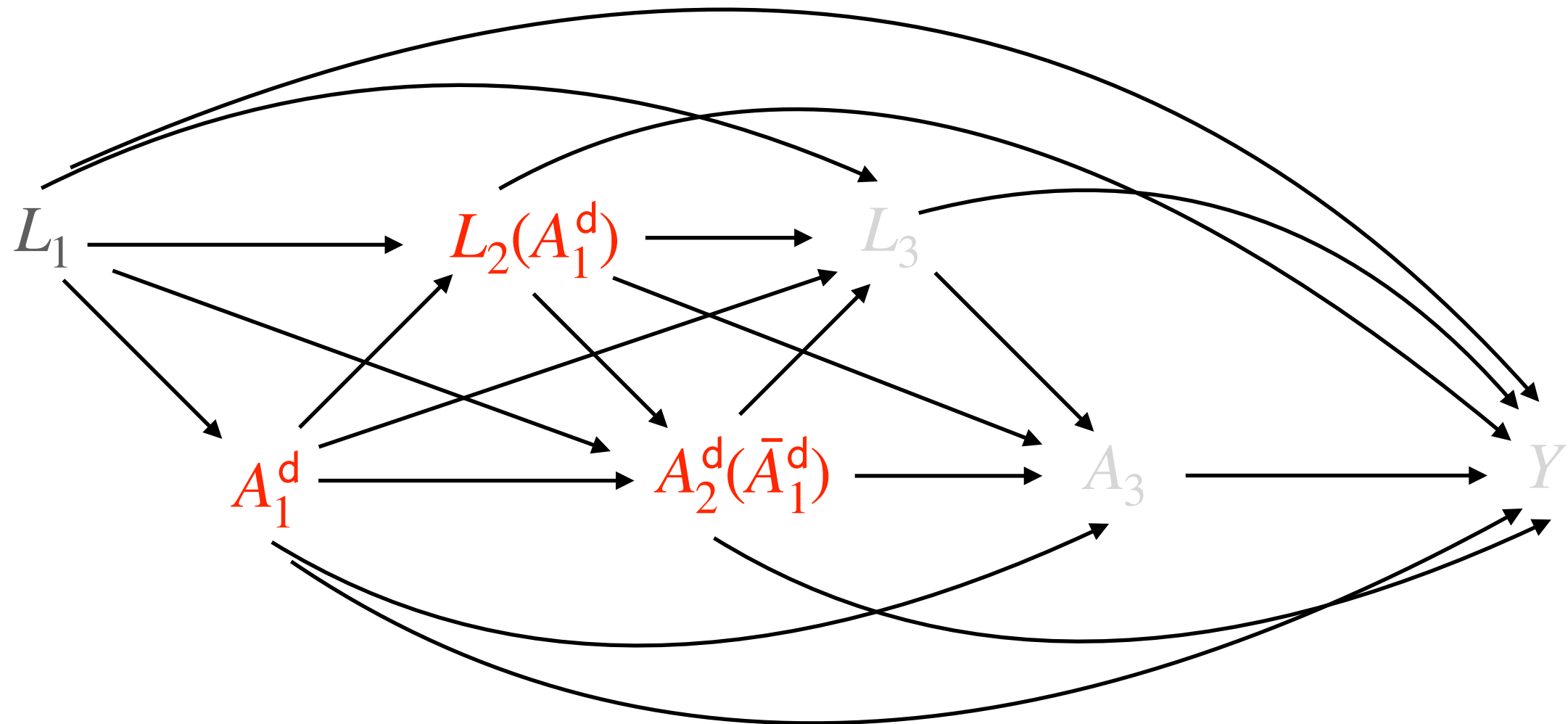


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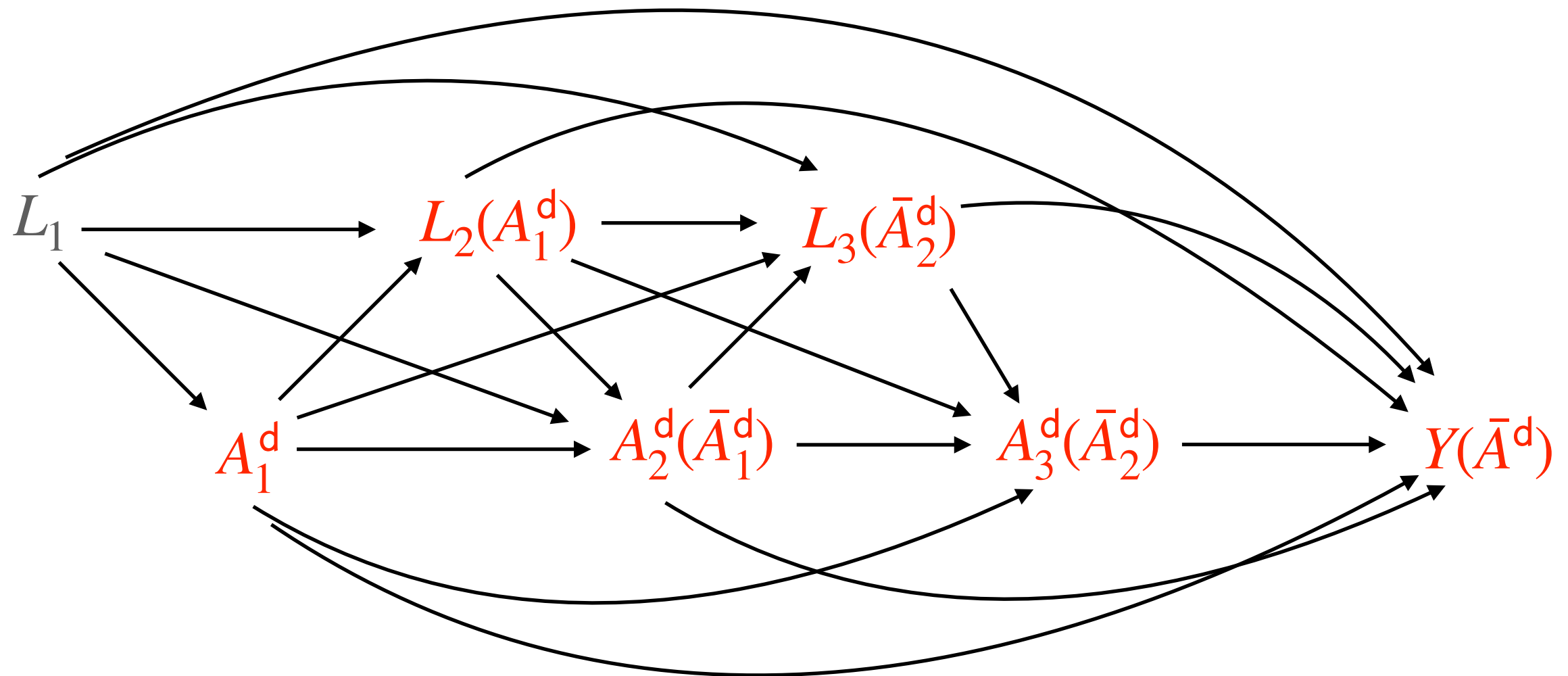


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- Natural value of treatment:  
 $A_t(\bar{A}_{t-1}^d) = f_{A_t}(H_t(\bar{A}_{t-1}^d), U_{A,t})$
- Value of treatment that would have been observed at time  $t$  under an intervention carried out up until time  $t - 1$  and then discontinued (*Richardson and Robins, 2013; Young et al., 2014*)



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- At time point 1, the observed treatment and natural value of treatment are equivalent



# Types of interventions d



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## Brief review of static and dynamic

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  - Example: initiate corticosteroids for COVID-19 patients for 6 days if a patient's oxygen levels drop (*Hoffman et al., 2022*)

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  - Example: initiate corticosteroids for COVID-19 patients for 6 days if a patient's oxygen levels drop (*Hoffman et al., 2022*)

$$d_t(h_t) = \begin{cases} 1 & \text{if } l_s^* = 1 \text{ for any } s \in \{t-5, \dots, t\} \\ 0 & \text{otherwise} \end{cases}$$

# Types of interventions

## Brief review of static and dynamic

- $Z_1, \dots, Z_t$  i.i.d observations  
 $Z = (L_1, A_1, L_2, A_2, \dots, L_t, A_t, Y) \sim P$
- $A_t$  intervention variables
- $L_t$  time-varying covariates
- $Y = L_{\tau+1}$  outcome at end of follow-up
- $H_t = (\bar{A}_{t-1}, \bar{L}_t)$  history until just before  $A_t$

- **Static intervention:**  $d$  is a constant function
  - “Treat everyone at all time points of the study”

$$d_t = 1 \text{ for all } t \in 1, \dots, \tau$$

- **Dynamic intervention:**  $d$  is a function of a unit's covariate history  $h_t$ 
  - Example: initiate corticosteroids for COVID-19 patients for 6 days if a patient's oxygen levels drop (*Hoffman et al., 2022*)

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where  $L_t^*$  is a variable in  $H_s$  denoting first instance of low oxygen



# Types of interventions

## Modified Treatment Policies

*(Díaz and van der Laan, 2012; Haneuse and Rotnitzky, 2013)*

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$$d(a_t, h_t)$$

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In modified treatment policies (MTPs),  
the intervention  $d$  is a function of the  
**natural value of treatment**

$$d(a_t, h_t)$$

# Types of interventions

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*(Díaz and van der Laan, 2012; Haneuse and Rotnitzky, 2013)*

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- *Example:* categorize average number of drinks per week as 1=“none”, 2=“1-5”, 3=“6-10”, 4=“11-15”, 5=“25”, then intervene to lower all individuals in the highest two drinks-per-week category to “6-10”

# Types of interventions

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$$d_t(a_t) = \begin{cases} a_t & \text{if } a_t < 4 \\ 3 & \text{otherwise} \end{cases}$$

# Types of interventions

## Modified Treatment Policies

*(Díaz and van der Laan, 2012; Haneuse and Rotnitzky, 2013)*

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- *Example:* Smoking cessation policy (*Robins et al. 2004*). Consider an intervention in which half of all current smokers quit smoking forever. Define  $A_t$  a random variable denoting smoking status, and  $\epsilon$  a random draw from a Uniform(0,1).

# Types of interventions

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- *Example:* Smoking cessation policy (*Robins et al. 2004*). Consider an intervention in which half of all current smokers quit smoking forever. Define  $A_t$  a random variable denoting smoking status, and  $\epsilon$  a random draw from a Uniform(0,1).

$$d_t(a_t, \epsilon_t) = \begin{cases} 0 & \text{if } \epsilon < 0.5 \text{ and } a_t = 1 \\ a_t & \text{otherwise} \end{cases}$$

# Types of interventions

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- *Example: Improving air quality.* Consider an intervention in which we lower daily PM2.5 levels by 10% for all US counties with PM2.5 levels greater than  $5 \mu\text{g}/\text{m}^3$

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- *Example: Improving air quality.* Consider an intervention in which we lower daily PM2.5 levels by 10% for all US counties with PM2.5 levels greater than 5  $\mu\text{g}/\text{m}^3$

$$d_t(a_t) = \begin{cases} a_t * 0.9 & \text{if } a_t > 5 \\ a_t & \text{otherwise} \end{cases}$$

# Setup

## Causal effects

- $Z_1, \dots, Z_t$  i.i.d observations  
 $Z = (L_1, A_1, L_2, A_2, \dots, L_t, A_t, Y) \sim P$
- $A_t$  intervention variables
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- $Y = L_{\tau+1}$  outcome at end of follow-up
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- Once an intervention  $d$  is specified, causal effects will be defined in terms of the **distribution of the counterfactual outcome**

$$Y(\bar{A}^d) = f_Y(A_\tau^d, H_\tau(\bar{A}_{\tau-1}^d), U_Y)$$

- Specifically, our causal parameter of interest is  $\theta = E[Y(\bar{A}_\tau^d)]$
- In *Epidemiology* tutorial, we focus on distributions of the contrast  $E[Y(\bar{A}_\tau^d) - Y(\bar{A}_\tau^{d'})]$ , where  $d$  and  $d'$  are two different interventions

# Identification



# Identification

## Assumptions

- $Z_1, \dots, Z_t$  i.i.d observations  
 $Z = (L_1, A_1, L_2, A_2, \dots, L_\tau, A_\tau, Y) \sim \mathcal{P}$
- $A_t$  intervention variables
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# Identification

## Assumptions

- **Positivity:** If  $(a_t, h_t) \in \text{supp}\{A_t, H_t\}$  then  $(d(a_t, h_t), h_t) \in \text{supp}\{A_t, H_t\}$  for  $t \in \{1, \dots, \tau\}$

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- **Standard sequential randomization:**  
 $U_{A,t} \perp \underline{U}_{L,t+1} \mid H_t$  for all  $t \in \{1, \dots, \tau\}$

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*Needed for static,  
dynamic, and  
(interestingly!)  
stochastic\* LMTPs*

*\*See LMTP-SI discussion in  
Diaz et al.*

# Identification

## Assumptions

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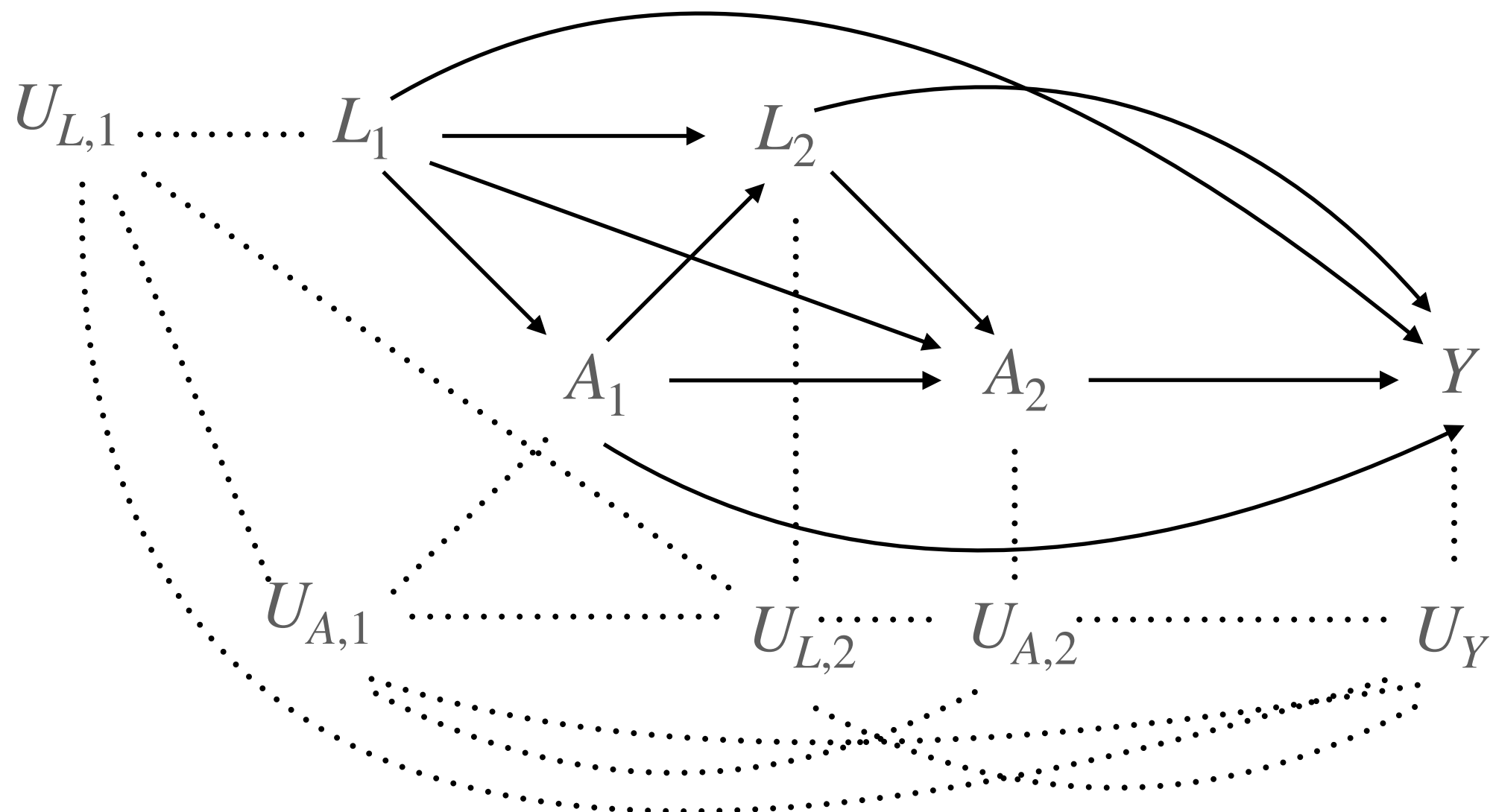
*Needed for LMTPs*

# Identification

## Sequential randomization

- $Z_1, \dots, Z_t$  i.i.d observations  
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***A DAG for two time points***



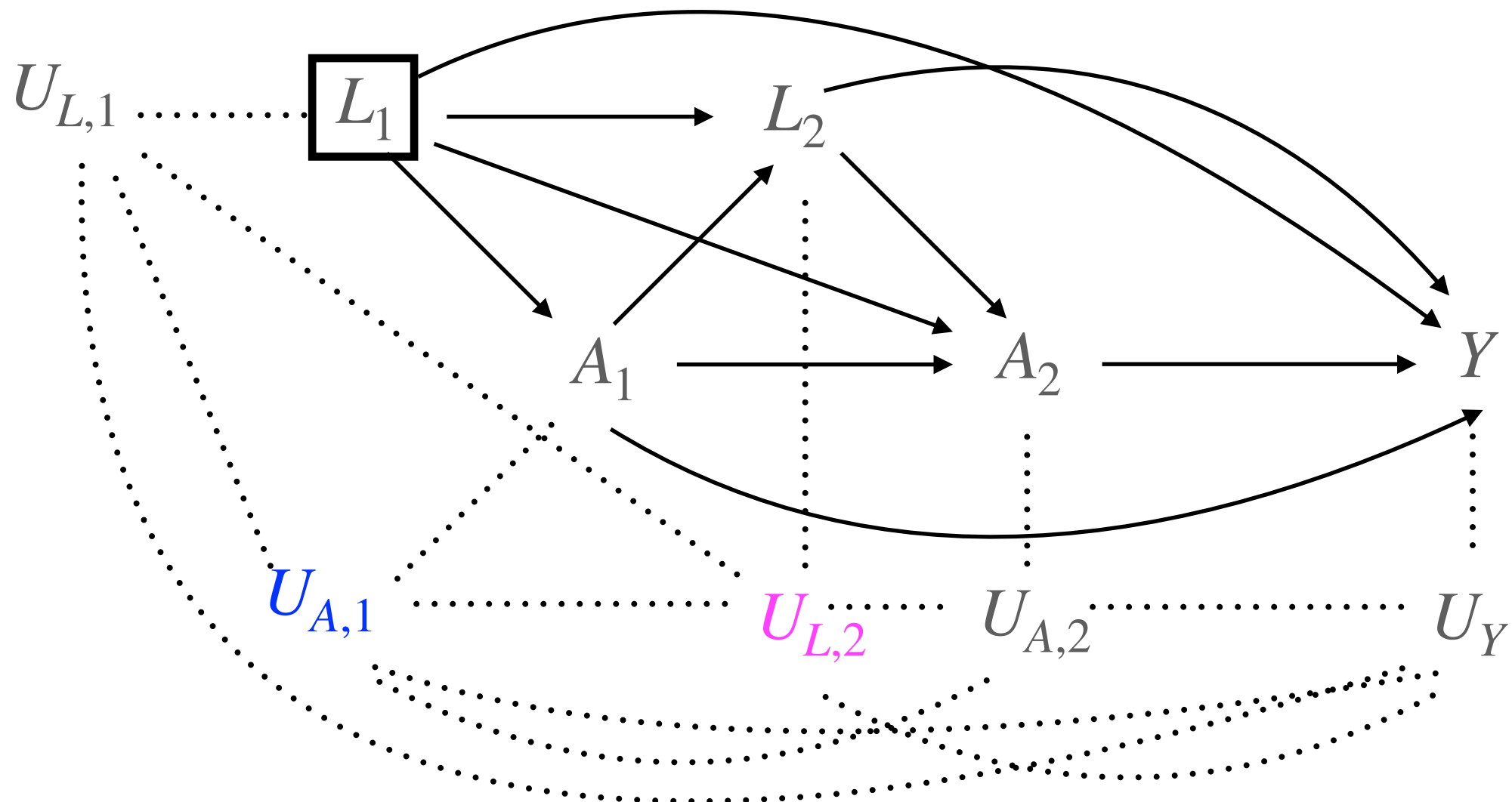
# Identification

## Sequential randomization

**Standard sequential randomization:**

$$U_{A,t} \perp U_{L,t+1} \mid H_t \text{ for all } t \in \{1, \dots, \tau\}$$

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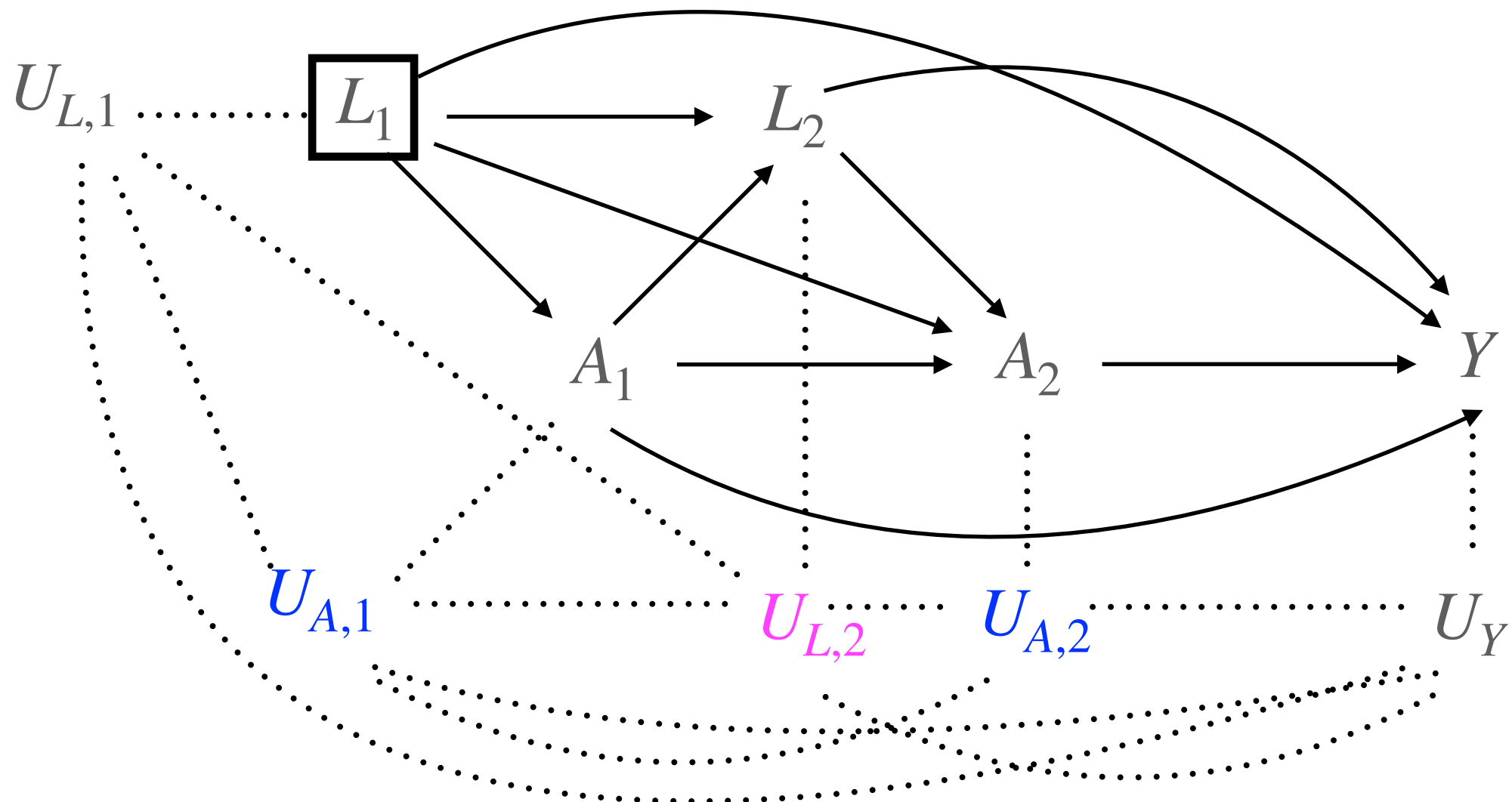
# Identification

## Sequential randomization

**Strong sequential randomization:**

$$U_{A,t} \perp (U_{L,t+1}, U_{A,t+1}) \mid H_t \text{ for all } t \in \{1, \dots, \tau\}$$

- $Z_1, \dots, Z_t$  i.i.d observations  
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- The distribution of interest is supported in the data (*Young et al. 2014*)
- If it is possible to find a unit with history  $h_t$  and exposure  $a_t$  at time  $t$ , then it is also possible to find a unit with history  $h_t$  and exposure  $d(a_t, h_t)$
- If exposure is multivariate (e.g. includes a loss-to-follow-up indicator), we also require a positive probability of observing a patient who is not lost to follow up

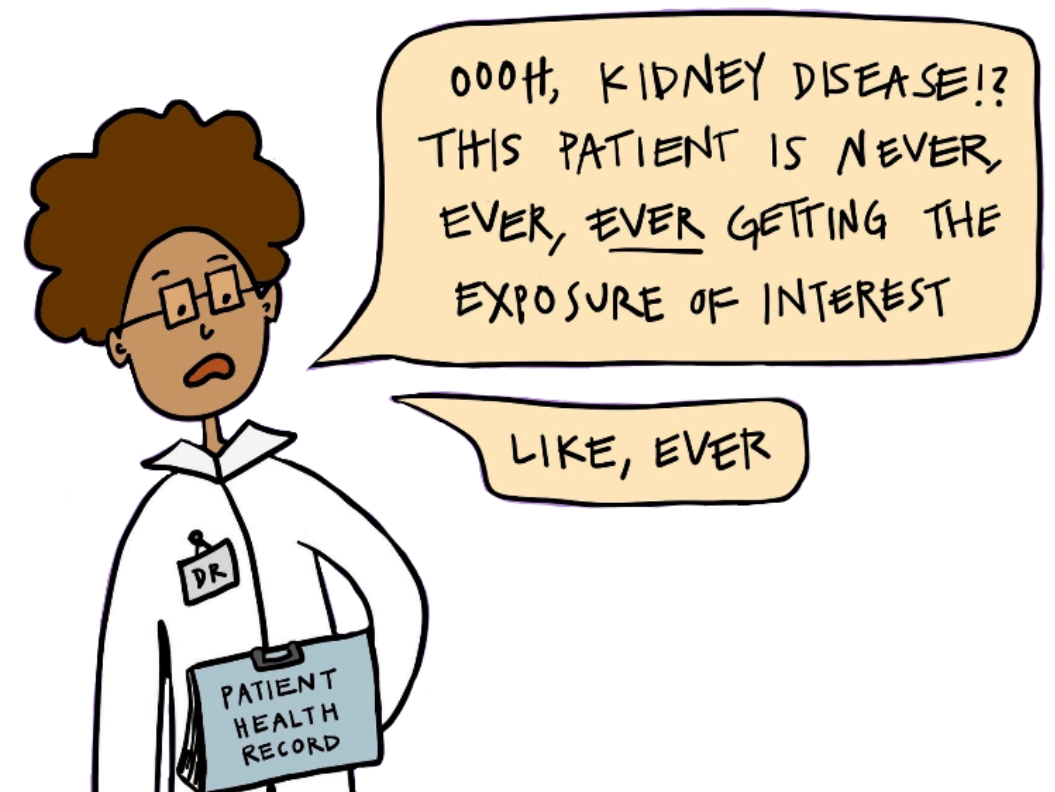
# Identification

## Positivity Assumption

- $Z_1, \dots, Z_t$  i.i.d observations  
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### Violations to positivity can be

- **Structural:** certain characteristics of an individual or unit which will never yield receipt of the treatment assignment under the intervention
- **Practical:** due to random chance or small datasets, certain covariate combinations have zero or near zero predicted probabilities of treatment

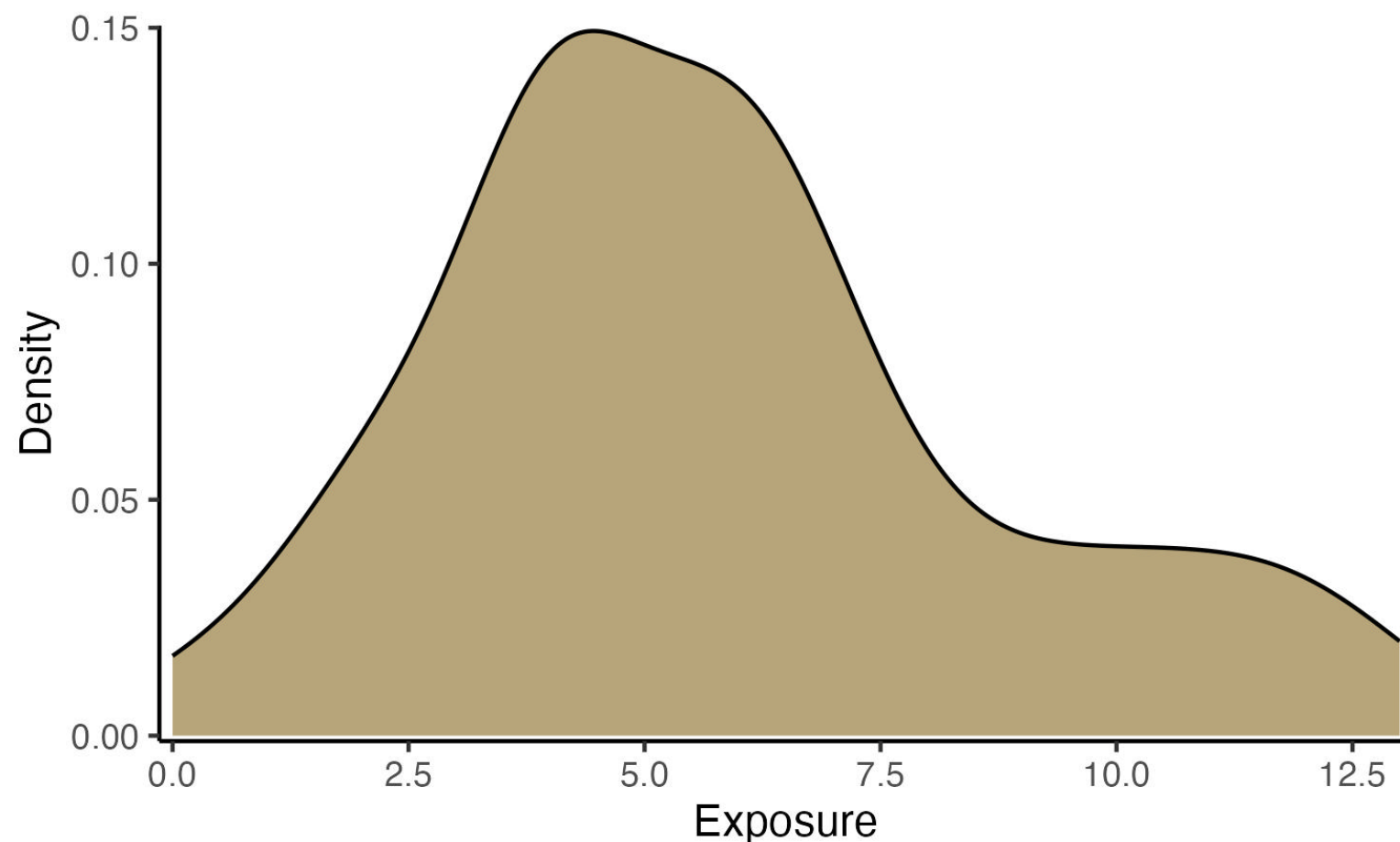


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- By design, non-static interventions may help define estimands with plausible positivity



# Identification

## General formula

- $Z_1, \dots, Z_t$  i.i.d observations  
 $Z = (L_1, A_1, L_2, A_2, \dots, L_t, A_t, Y) \sim P$
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- Under Positivity and Strong Sequential Randomization, an LMTP estimand is identified by an alternative expression of the extended g-formula (*Robins et al. 2004, Richardson and Robins, 2013*)

# Identification

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- Under the previously discussed assumptions,  $\theta = E[Y(\bar{A}_1^d)]$  is identified by  $\theta = E[m_1(A_1^d, L_1)]$ .

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5. Under identifying assumptions,  $E[Y(\bar{A}_\tau^d)] = E[\tilde{Y}_1]$



# Estimation

# Estimation

## Substitution estimator

- The simplest form of estimation is to use a plug-in estimator

$$\hat{\theta}_{sub} = \frac{1}{n} \sum_{i=1}^n \hat{m}_1(A_{1,i}^d, L_{1,i})$$

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### Pseudo-R Algorithm:

```
1. fit_y <- glm(Y ~ A2 + L2 + A1 + L1)
2. Y_tilde_2 <- predict(fit_y, A2 = A2d)
3. fit_ytilde_2 <- glm(Y_tilde_2 ~ A1 + L1)
4. Y_tilde_1 <- predict(fit_ytilde_2, A1 = A1d)
5. mean(Y_tilde_1)
```

# Estimation

## Inverse Probability Weighting

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$$\hat{\theta}_{IPW} = \frac{1}{n} \sum_{i=1}^n \left( \prod_{t=1}^{\tau} \hat{r}_t(A_{t,i}, H_{t,i}) \right) Y_i$$

# Estimation

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  - Use bootstrap or influence function-based estimator to construct asymptotically correct confidence intervals

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- **Solution:** use the efficient influence function to propose estimators that are  $n^{1/2}$ -consistent and efficient under weaker assumptions

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- The **efficient influence function** will allow us to
  - Construct locally efficient estimators
  - Enjoy multiple robustness properties of our estimator
  - Use flexible regression techniques in estimating nuisance quantities, since the EIF yields second-order bias terms with slow convergence rates for the nuisance parameters

# Estimation

## EIF Conditions

### Technical requirements:

Assume  $d$  does not depend on  $P$ , and one of:

1.  $A_t$  is a **discrete** random variable for all  $t$
2.  $A_t$  is a **continuous** random variable and the modified treatment policy  $d$  satisfies **piecewise smooth invertibility**

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$$\phi_1(Z) - \theta$$

# Estimation

## EIF Formula

**Efficient influence function:**  $\phi_1(Z) - \theta$

Where,

$$\phi_t : z \mapsto \sum_{s=t}^{\tau} \left( \prod_{k=t}^s r_k(a_k, h_k) \right) \{ m_{s+1}(a_{s+1}^d, h_{s+1}) - m_s(a_s, h_s) \} + m_t(a_t^d, h_t)$$

Recall  $r_t(a_t, h_t)$  is the density ratio,

$$r_t(a_t, h_t) = \frac{g_t^d(a_t | h_t)}{g_t(a_t | h_t)}$$

# Estimation

## Non-parametric estimators

- Two estimators for LMTP proposed in Díaz et al. and implemented in {lmtp} using  $m_t$ ,  $r_t$ , and the derived EIF:
  1. Targeted Minimum Loss-Based Estimation (TMLE)
  2. Sequentially doubly robust (SDR) estimation
- Utilize sample splitting and cross-fitting (*Klassen, 1987; Zheng and van der Laan, 2011; Chernozhukov et al., 2018*)
- Algorithms formally proposed in *Díaz et al. (2021)*; pseudo-R code given in Appendix of *Hoffman et al. (2024)*



# Estimation

## TMLE

- A generalization of estimators proposed for MTPs in a single time point setting in (*Díaz and van der Laan, 2018; van der Laan and Gruber, 2012*)
- A substitution estimator that uses an estimate  $\tilde{m}_1$  (within cross-fitting folds) carefully constructed to solve a cross-validated efficient influence function estimating equation
- Motivated by the decomposition of the EIF as a sum of terms of the form,

$$\left(\prod_{k=1}^t r_k(a_k, h_k)\right) \{m_{t+1}(a_{t+1}^d, h_{t+1}) - m_t(a_t, h_t)\}$$

# Estimation

## SDR Estimator

- SDR is an extension of estimators for dynamic treatment regimes proposed by *Luedtke et al. (2017)* and *Rotnitzky et al. (2017)*
- Sequentially regress a multiply robust unbiased data transformation related to  $\phi_{t+1}$  to construct pseudo-outcomes, and eventually,  $\hat{\theta}_{SDR}$
- Not a substitution estimator, so SDR can yield estimates that are outside the bounds of  $Y$ , but it has better consistency properties than TMLE in the event of model misspecification

# Estimation

## Comparing statistical properties of proposed estimators

<i>Statistical property</i>	Sub.	IPW	TMLE	SDR	iTMLE
Doubly robust			X	X	X
Sequentially doubly robust				X	X
Valid inference using parametric regressions	X	X	X	X	X
Valid inference using data-adaptive regressions			X	X	X
Guaranteed to stay within outcome range	X		X		X

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- Effect of **lowering self-reported knee pain scores** on knee replacement surgery (*Jafarzadeh et al.*)

# Extensions and Applications

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- Effect of **delaying intubation** on {14-day mortality, 14-day acute kidney injury} in COVID-19 patients (*Díaz et al., Hoffman et al.*)
- Effect of **increasing PaO<sub>2</sub>/FiO<sub>2</sub>** by 50 units on 28-day mortality in Acute Respiratory Distress Patients (*Díaz et al.*)
- Effect of **increasing mobility rates** on COVID-19 case rates (*Nugent and Balzer*)
- Effect of **Naxolone access laws** on opioid overdose rates (*Rudolph et al.*)
- Effect of **lowering self-reported knee pain scores** on knee replacement surgery (*Jafarzadeh et al.*)
- Effect of **increases in number of primary care physicians** on post-operative outcomes for elective total joint replacement (*Mehta et al.*)

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- Effect of **increasing numbers of tooth retention** and social participation among older adults in Japan (*Cooray et al.*)

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  - Formulating alternative estimands to satisfy the positivity assumption must be done with scientific reason, and practical positivity violations may still occur
    - Solutions to practical violations (e.g. truncation of density ratios) are arbitrary and may lead to biases

**Thank you!**

**Questions,  
comments,  
suggestions?**

