Time Series Forecasting

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Data

2008

2009

2010

762963

814978

927941

852722

894346

941628

803094

884482

943152

793027

894576

956392

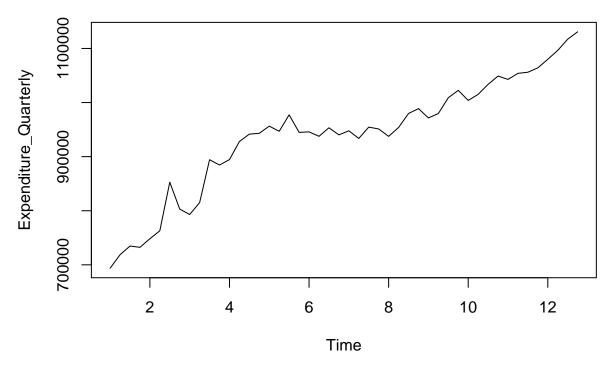
Series: Federal Government: Current Expenditures Web Address: (https://fred.stlouisfed.org/series/NA000 283Q) (1) The federal government budget includes three types of spending: mandatory, discretionary, and interest on debts owed by the nation. The majority of the Government's expenditures go towards Social Security, Medicare, and Medicaid programs as mandatory spending. (a) Units: Millions of Dollars, Not Seasonally Adjusted (b) Low Frequency: Quarterly (c) Number Observations: 48 (i) Duration: Q4 2006-Q3 2018 (ii) Within-Sample: Area of focus is Q4 2006-Q3 2016 (40 Obs) (iii) Post-Sample: This leaves 8 quarters (Q4 2016 - Q3 2018) after the area of focus for post-sample testing

```
#read/attach data
data=read.csv("/stfm/dev2/m1kal01/Personal/data.csv")
attach(data)
#generate time series
Expenditure Quarterly=ts(Expenditure, frequency=4)
print(Expenditure Quarterly)
##
         Qtr1
                 Qtr2
                         Qtr3
                                  Qtr4
## 1
              718612 734747
       693643
                               732366
       748213
              762963
                       852722
                               803094
       793027
               814978 894346
## 3
                               884482
##
       894576
               927941
                       941628
                               943152
       956392 946858
                      977285
                               944835
## 5
       945771
               937575
                       953318
                               940273
## 7
       947808
               933661
                       954582
                               951184
       937463
               953938
                       979777
## 8
                               988781
       971506 979738 1008980 1022502
## 10 1003962 1015165 1033669 1048953
## 11 1042766 1054016 1056024 1064175
## 12 1079954 1096440 1117081 1130728
#Create size restrictions
N=length(Expenditure_Quarterly)
WithinSampleLength=N-8
WithinSample=ts(Expenditure_Quarterly[1:WithinSampleLength], frequency=4, start=c(2006,4))
print(WithinSample)
##
           Qtr1
                   Qtr2
                           Qtr3
                                    Qtr4
## 2006
                                  693643
## 2007
         718612
                 734747
                         732366
                                 748213
```

```
## 2011
         946858
                 977285
                          944835
                                  945771
  2012
         937575
                 953318
                          940273
                                  947808
                                  937463
  2013
         933661
                 954582
                          951184
  2014
         953938
                          988781
                                  971506
                 979777
  2015
         979738 1008980 1022502
                                 1003962
## 2016 1015165 1033669 1048953
PostTest=Expenditure_Quarterly[(WithinSampleLength+1):N]
print(PostTest)
```

[1] 1042766 1054016 1056024 1064175 1079954 1096440 1117081 1130728

Plot



Purpose

When modeling a series, you must start by testing for stationarity, this is done by looking at the behavior of the error term. We are looking to see if the errors or series are autocorrelated, and thus the behavior of the variable today I affected by its behavior in the past. If the autocorrelation coefficient, rho, is equal to 1, the error term behaves as Et=Et-1+ut. This is a random walk model, thus the variance become undefined, as it goes to infinity. In terms of the difference equation, the time path and the general solution of this variable will not converge to equilibrium, the model is totally random. This is a unit root, which can detected by a Dickey-Fuller test, and it must be dealt with by differencing the equation. Once the series is stationary, it must be tested to see if the series is truly white noise, which can be tested by a Ljung-Box test. In the case that it is truly white noise, the model can not be used to forecast. If it is not, examining of the autocorrelation functions and partial autocorrelation functions will reveal if the model should be an AR,MA, or ARMA process.

#Test for Stationarity First, a Dickey-Fuller test for a unit root was ran on the original series.

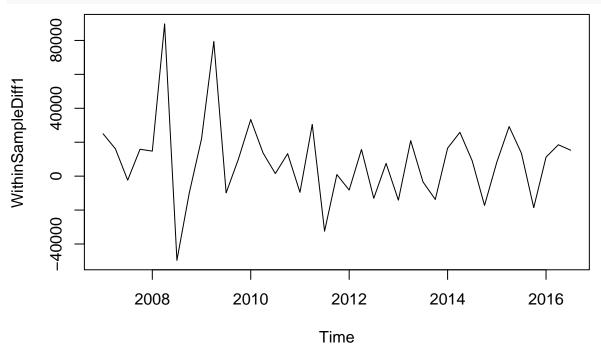
```
#adf is augmented dickey fuller test
adf.test(WithinSample,alternative='stationary')
```

```
##
## Augmented Dickey-Fuller Test
##
## data: WithinSample
## Dickey-Fuller = -2.1713, Lag order = 3, p-value = 0.5063
## alternative hypothesis: stationary
```

The results of this test: A large p-value is indicative of a unit root process, that the series is not stationary. To deal with a not stationary series, we use differencing. Reject null hypothesis that series is stationary, must difference it.

Differencing the equation:

```
WithinSampleDiff1=diff(WithinSample, differences=1)
plot(WithinSampleDiff1)
```



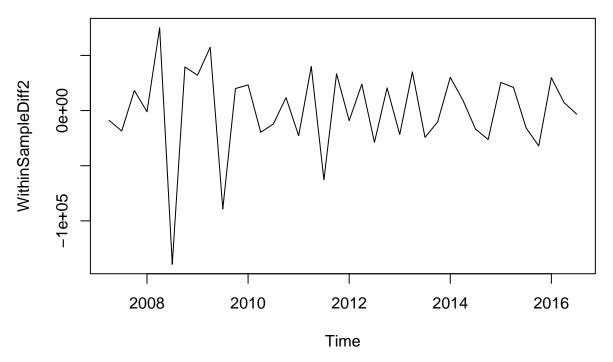
Dickey-Fuller Test for Stationarity on First-Differenced Series:

```
adf.test(WithinSampleDiff1,alternative='stationary')
```

```
##
## Augmented Dickey-Fuller Test
##
## data: WithinSampleDiff1
## Dickey-Fuller = -2.5403, Lag order = 3, p-value = 0.3618
## alternative hypothesis: stationary
The results of this test: Still reject null, difference again.
```

Differencing the equation again

```
WithinSampleDiff2=diff(WithinSample, differences=2)
plot(WithinSampleDiff2)
```



Dickey-Fuller Test for Stationarity on Second-Differenced Series

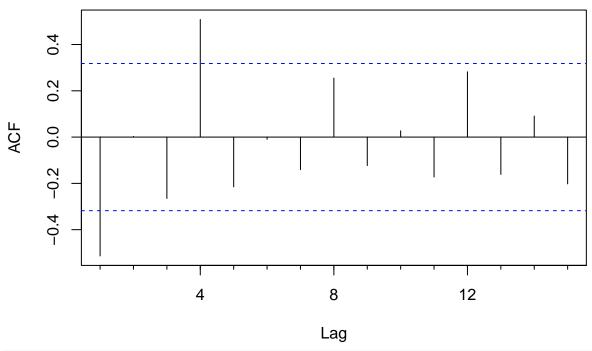
```
adf.test(WithinSampleDiff2,alternative='stationary')
```

```
## Warning in adf.test(WithinSampleDiff2, alternative = "stationary"): p-value
## smaller than printed p-value
##
## Augmented Dickey-Fuller Test
##
## data: WithinSampleDiff2
## Dickey-Fuller = -6.5834, Lag order = 3, p-value = 0.01
## alternative hypothesis: stationary
```

The results of this test: Now fail to reject null, there is now not a unit root process and the series is stationary. #Partial Autocorrelation Function (PACF) and Autocorrelation Function(ACF)

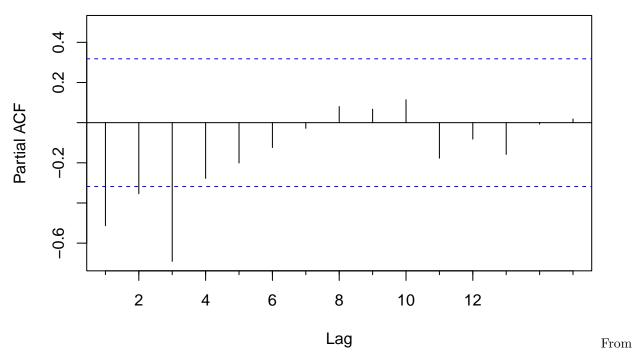
Acf(WithinSampleDiff2,main='This is my Autcorrelation Function')

This is my Autcorrelation Function



Pacf(WithinSampleDiff2, main='This is my Partial Autcorrelation Function')

This is my Partial Autcorrelation Function



visual inspection, it appears to tail off after 3 My Guess: IMA (2,4)

#Doing Auto-ARIMA Auto-ARIMA is selected by running tons of possible combinations of a model and choosing the one with the smallest AIC (Akaike Information Criteria)

```
auto.arima(WithinSample, seasonal=TRUE)
## Series: WithinSample
## ARIMA(0,1,1)(0,0,1)[4] with drift
##
## Coefficients:
##
                              drift
             ma1
                    sma1
##
         -0.3688 0.6272 9247.707
## s.e.
          0.1535 0.1528 3150.005
## sigma^2 estimated as 415737851: log likelihood=-441.83
## AIC=891.65
                AICc=892.83
                               BIC=898.31
This Auto Arima I disagree with, it doesn't make sense it only differenced it once when it is still not stationary
yet. In response I forced it to allow me to difference twice.
ARIMAWithin=auto.arima(WithinSample, d=2, seasonal=TRUE)
print(ARIMAWithin)
## Series: WithinSample
## ARIMA(0,2,2)(0,0,1)[4]
##
## Coefficients:
##
             ma1
                     ma2
                             sma1
##
         -1.3338 0.3973
                          0.6281
## s.e.
          0.1684 0.1884 0.1542
## sigma^2 estimated as 441882995: log likelihood=-432.62
## AIC=873.23
               AICc=874.44
                              BIC=879.78
#Forecasting with Auto-Arima
#h means steps ahead
ARIMAWithinYhats=forecast(ARIMAWithin, h=8)
print(ARIMAWithinYhats)
           Point Forecast
                            Lo 80
                                     Hi 80
                                               Lo 95
                                                       Hi 95
## 2016 Q4
                 1040532 1013592 1067472 999330.7 1081733
## 2017 Q1
                  1055804 1023432 1088175 1006296.0 1105312
## 2017 Q2
                  1063111 1025237 1100985 1005187.8 1121035
## 2017 Q3
                  1075116 1031627 1118604 1008605.8 1141626
## 2017 Q4
                  1080337 1021246 1139429 989965.2 1170710
## 2018 Q1
                  1087824 1018593 1157055 981945.0 1193703
## 2018 Q2
                  1095311 1015920 1174702 973893.2 1216729
                  1102798 1013146 1192449 965688.0 1239907
## 2018 Q3
ActualForecastValues=ARIMAWithinYhats[4]
print(PostTest[1:8])
## [1] 1042766 1054016 1056024 1064175 1079954 1096440 1117081 1130728
PostValues=PostTest[1:8]
print(PostValues)
## [1] 1042766 1054016 1056024 1064175 1079954 1096440 1117081 1130728
Getting post-sample residuals:
```

ARIMAResidualsTest=ARIMAWithinYhats\$mean-PostTest print(ARIMAResidualsTest)

```
## Qtr1 Qtr2 Qtr3 Qtr4
## 2016 -2233.8898
## 2017 1787.8849 7087.3125 10940.8186 383.4696
## 2018 -8615.7965 -21770.0625 -27930.3286
```

Getting RMSE

sqrt(mean(ARIMAResidualsTest^2))

[1] 13722.86