

# Modelling and quantifying mortality and longevity risk

Module A.1 on Basic concepts of lifetime distributions, mortality data and visualisations

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OF AMSTERDAM



RCLR

Research Centre  
for Longevity Risk

Learning outcomes

Motivation

Random future lifetime

Force of mortality

The life table

Central death rate

Life expectancy

Actuarial Present Value formulas

Wrap-up

In this module you will learn:

- how to define, denote and understand basic concepts related to lifetime data, relevant for the course
- how basic demographic markers evolve / have evolved as a function of age and time.

In this module we take a so-called **period approach** and consider the lifetime distribution or life table applicable to a specific period (e.g., a year).

We'll switch to cohort (or: dynamic) thinking from Module 2 on.

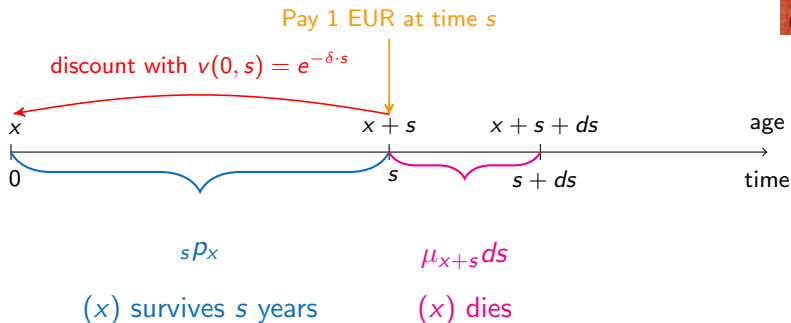
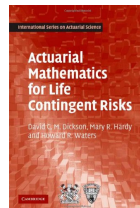
# Motivation

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# Motivation

## Life insurance mathematics 101 - whole life insurance (continuous)

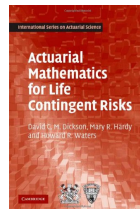
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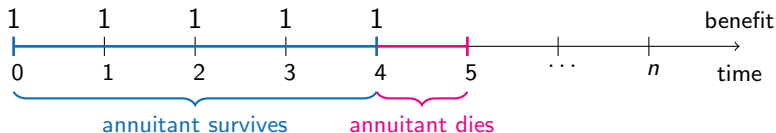
What is the Actuarial Present Value (or Expected Present Value) of this cashflow?

# Motivation

## Life insurance mathematics 101 - whole life annuity-due (discrete)



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What is the Actuarial Present Value (or Expected Present Value) of this cashflow?

A (detailed) recap of relevant actuarial value formulas for life insurance and life annuity products will follow later on.

However, these introductory examples motivate the **need to master lifetime (or survival) distributions** and related concepts when valuing life-contingent risks:

- $T_x$  and  $K_x = \lfloor T_x \rfloor$ , and their cdf, pdf, survival function, expected value
- survival probabilities, e.g.,  ${}_s p_x$
- death probabilities, e.g.,  ${}_s q_x$
- ...

*'Longevity risk, the risk that people will live longer than expected, weighs heavily on those who run pension schemes and on insurers that provide annuities. Actuaries have a track record of systematically underestimating gains in life expectancy, and more old people means a bigger bill for benefits providers.'*

(The Economist, *Live long and prosper*, February 4, 2010.)

**The Economist**



World politicsBusiness & financeEconomicsScience & technologyCultu

**Longevity swaps**


## Live long and prosper

Plans are afoot to create a new capital market in longevity risk

Feb 4th 2010 | From the print edition

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DEATH comes to everyone. The timing is much less certain. Longevity risk, the risk that people will live longer than expected, weighs heavily on those who run pension schemes and on insurers that provide annuities. Actuaries have a track record of systematically underestimating gains in life expectancy, and more old people means a bigger bill for benefits providers. Every additional year of life expectancy at age 65 is reckoned to bump up the present value of pension liabilities in British defined-benefit schemes by 3%, or £30 billion (\$48



Spry and retiring  
Getty Images



## Your turn:

seeing the definition of longevity risk on the previous slide, how would you define the concept of mortality risk?

For which type of products is mortality risk viz longevity risk relevant to consider?



# Basic concepts of lifetime distributions and mortality data

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- ▶ Let  $T_x$  be the remaining random lifetime for a person aged  $x$ .
- ▶ Connection between  $T_x$  and  $T_0$ :

$$P(T_x > t) = P(T_0 > x + t | T_0 > x) = \frac{P(T_0 > x + t)}{P(T_0 > x)}.$$

- ▶ A person alive at age  $x$  will die at the random age  $x + T_x$ .

- Survival and death probabilities:

$${}_t p_x = P(T_x > t) = P(T_0 > t + x | T_0 > x),$$

and

$${}_t q_x = 1 - {}_t p_x = P(T_x \leq t) = P(T_0 \leq x + t | T_0 > x).$$

${}_t p_x$  is the probability that an individual aged  $x$  will survive to age  $x + t$ .

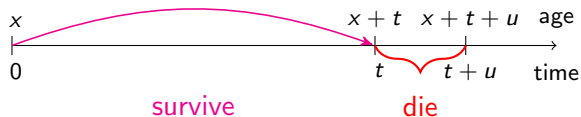
${}_t q_x$  is the probability that an individual aged  $x$  will die before reaching age  $x + t$ .

# Survival and death probabilities

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## Deferred probability

${}_t|{}_uq_x$  = probability that  $(x)$  will die between the ages of  $x + t$  and  $x + t + u$ .



$(x)$  survives for  $t$  years, but will die before age  $x + t + u$ :

$${}_t|{}_uq_x = {}_tp_x \cdot {}_uq_{x+t}.$$

- ▶ Throughout our modules we will use the 'exact age' definition of  $q_x$ . (Some more details will follow in Module 2.)
- ▶ Therefore:

$q_x := {}_1q_x$  is the probability that an individual aged exactly  $x$  will die before reaching age  $x + 1$ .

- ▶ The **survival function**  $S_0(t)$  of  $T_0$  and  $S_x(t)$  of  $T_x$ :

$$\begin{aligned}S_0(t) &= P(T_0 > t), \\S_x(t) &= P(T_x > t) = {}_t p_x.\end{aligned}$$

- ▶ And also:

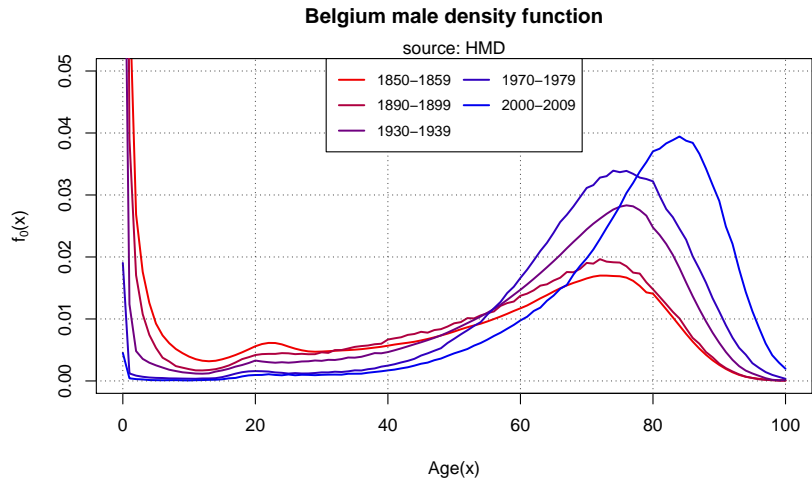
$f_0(t)$ ,  $f_x(t)$ , the **probability density function** of the future lifetime distribution of a 0 year old, respectively  $x$  year old, evaluated in  $t$

$F_0(t) = Pr(T_0 \leq t)$  and  $F_x(t) = Pr(T_x \leq t)$  for the **cumulative distribution function** of the future lifetime distribution of a 0 year old, respectively an  $x$  year old.



# Mortality trends and visuals

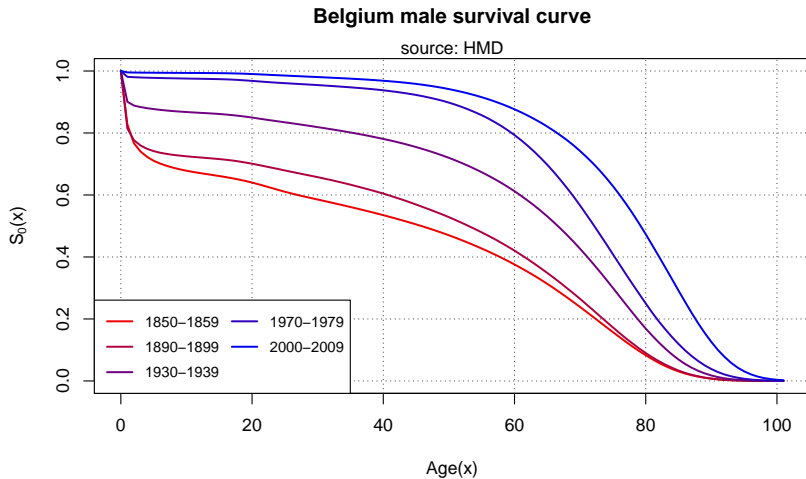
## Rectangularization and expansion



# Mortality trends

## Rectangularization and expansion

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We find

$$\begin{aligned}S_0(t + u) &= P(T_0 > t + u) \\&= P(T_0 > t) \cdot P(T_0 > t + u | T_0 > t) \\&= P(T_0 > t) \cdot P(T_t > u) \\&= S_0(t) \cdot S_t(u).\end{aligned}$$

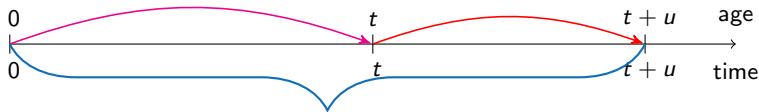
Similarly,

$$S_x(t + u) = S_x(t) \cdot S_{x+t}(u).$$

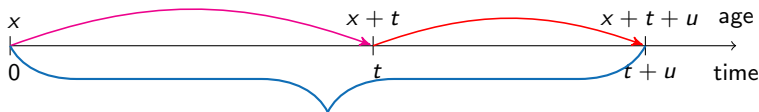
# Multiplication rule for survival probabilities

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Pictured



$$S_0(t + u) = S_0(t) \cdot S_t(u)$$



$$S_x(t + u) = S_x(t) \cdot S_{x+t}(u)$$

## Your turn:

use the multiplication rule for survival probabilities to express  ${}_k p_x$ , the  $k$ -year survival probability (with  $k \in \mathbb{N}$ ), as a function of 1-year survival probabilities  $(p_x, p_{x+1}, \dots)$  (recall: in this module we apply [period thinking](#)).

- Force of mortality (or: hazard rate) at age  $x$ ,  $\mu_x$ :

$$\begin{aligned}\mu_x &= \lim_{dx \rightarrow 0+} \frac{1}{dx} P[T_0 \leq x + dx | T_0 > x] \\ &\Updownarrow \\ &= \lim_{dx \rightarrow 0+} \frac{1}{dx} P[T_x \leq dx].\end{aligned}$$

- Meaning of the highlighted components in this definition?
- rate of event occurrence per unit of time
  - take limit  $dx \rightarrow 0+ \Rightarrow$  instantaneous rate of occurrence.

► We obtain:

$$\begin{aligned}\mu_x &= \lim_{dx \rightarrow 0+} \frac{1}{dx} \frac{P[x < T_0 \leq x + dx]}{P(T_0 > x)} = \lim_{dx \rightarrow 0+} \frac{F_0(x + dx) - F_0(x)}{dx \cdot S_0(x)} \\ &= \frac{1}{S_0(x)} \frac{d}{dx} F_0(x) = \frac{1}{S_0(x)} \left\{ -\frac{d}{dx} S_0(x) \right\} = \frac{f_0(x)}{S_0(x)} = -\frac{d}{dx} \ln S_0(x).\end{aligned}$$

► Integrating over  $(0, t)$  then yields:

$$\int_0^t \mu_x dx = -(\log S_0(t) - \log S_0(0)),$$

thus:

$${}_t p_0 = S_0(t) = \exp \left( - \int_0^t \mu_x dx \right).$$

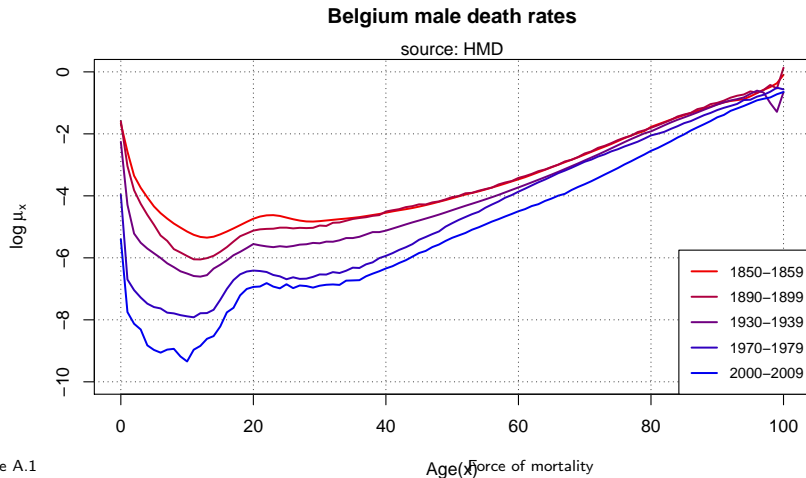
This implies the following important expression:

$$\begin{aligned} {}_t p_x = S_x(t) &= \frac{S_0(x+t)}{S_0(x)} = \exp \left( - \int_x^{x+t} \mu_r dr \right) \\ &= \exp \left( - \int_0^t \mu_{x+s} ds \right). \end{aligned}$$

This is a general expression in statistics, connecting the survival function and the hazard rate!



Logarithm of the force of mortality,  $\log(\mu_x)$ , versus age ( $x$ ) for Belgian males in four selected time periods.



- ▶  $\lambda_x$  the expected number of survivors at age  $x$  (out of initial group  $\lambda_0$ ), where

$$\lambda_{x+1} = \lambda_x \cdot (1 - q_x).$$

- ▶  $d_x$  is the expected number of deaths at age  $x$  (out of initial group  $\lambda_0$ ), where

$$d_x = \lambda_x - \lambda_{x+1} = \lambda_x \cdot q_x.$$

- ▶  $E_x$  is the **exposure to risk** at age  $x$ , i.e. the expected total time lived by  $\lambda_x$  people between age  $x$  and  $x + 1$ :

$$\lambda_x \int_0^1 {}_u p_x du = \int_0^1 \lambda_{x+u} du := E_x.$$

The **central death rate** at age  $x$ , denoted  $m_x$ , is then defined as a **weighted average** of the **force of mortality** over  $[x, x + 1)$ :

$$\begin{aligned} m_x &= \frac{\int_0^1 S_0(x+u) \cdot \mu_{x+u} du}{\int_0^1 S_0(x+u) du} \\ &\vdots \\ &= \frac{q_x \cdot \lambda_x}{\lambda_x \int_0^1 {}_u p_x du} = \frac{d_x}{\lambda_x \int_0^1 {}_u p_x du} \\ &= \frac{d_x}{E_x}. \end{aligned}$$

- The **life expectancy** (or expected lifetime) **for a newborn** is:

$$E[T_0] = \int_0^{\infty} t \cdot f_0(t) dt = \int_0^{\infty} S_0(t) dt.$$

- So, the **expected lifetime for an  $(x)$ -year old** is:

$$E[T_x] = \int_0^{\infty} t \cdot f_x(t) dt = \int_0^{+\infty} S_x(t) dt = \frac{1}{S_0(x)} \int_0^{+\infty} S_0(x+t) dt.$$

## Life expectancy

### Piecewise constant force of mortality

To calculate the life expectancy from a given life table with  $p_x$  for integer ages  $x$ , assume:

$$\mu_{x+t} = \mu_x \quad \text{with } t \in [0, 1).$$

That is: the force of mortality is **piecewise constant** between two integer ages.

This is equivalent to

$$\begin{aligned} {}_t p_x &= 1 - {}_t q_x = \exp \left( - \int_0^t \mu_{x+\tau} d\tau \right) \\ &= \exp \left( - \int_0^t \mu_x d\tau \right) \\ &= \exp (-t \cdot \mu_x) = \{ \exp (-\mu_x) \}^t = (1 - q_x)^t = p_x^t, \end{aligned}$$

for  $t \in [0, 1]$ .

Under this **piecewise constant** assumption we get:

$$\begin{aligned} E[T_x] &= \int_0^{+\infty} S_x(t) dt = \int_0^{+\infty} {}_t p_x dt = \sum_{k \geq 0} {}_k p_x \int_0^1 {}_t p_{x+k} dt \\ &= \sum_{k \geq 0} {}_k p_x \frac{p_{x+k} - 1}{\ln p_{x+k}} \\ &= \frac{p_x - 1}{\ln p_x} + \sum_{k \geq 1} \left( \prod_{j=0}^{k-1} p_{x+j} \right) \frac{p_{x+k} - 1}{\ln p_{x+k}}, \end{aligned}$$

where  ${}_k p_x = \prod_{j=0}^{k-1} p_{x+j}$  and  $p_x = \exp(-\mu_x)$ .

Alternatively, under the Uniform Distribution of Deaths (UDD):

$$\lambda_{x+t} = \lambda_x + t \cdot (\lambda_{x+1} - \lambda_x) \quad t \in [0, 1].$$

This assumption implies:  ${}_tq_x = 1 - {}_tp_x = 1 - \frac{\lambda_{x+t}}{\lambda_x} = \frac{\lambda_x - \lambda_{x+t}}{\lambda_x} = t \cdot q_x$  for  $t \in [0, 1]$ .

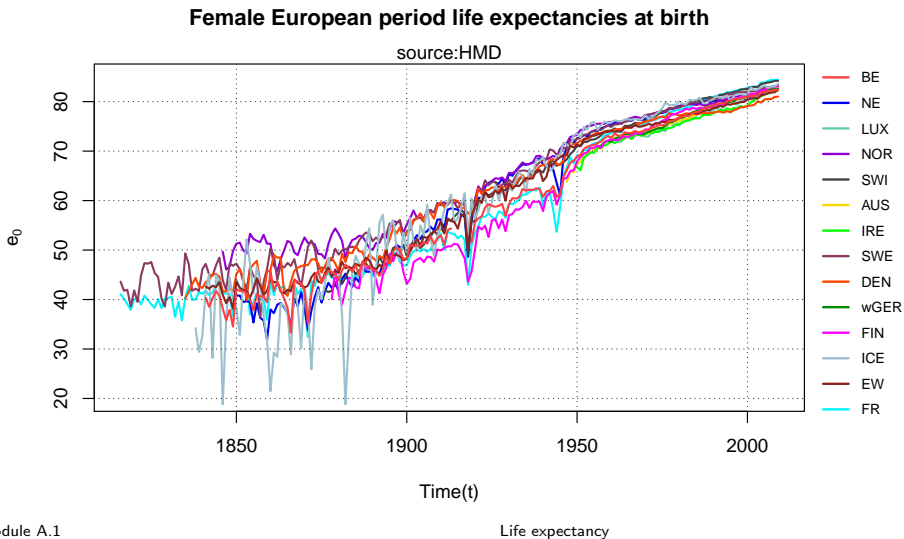
# The life expectancy

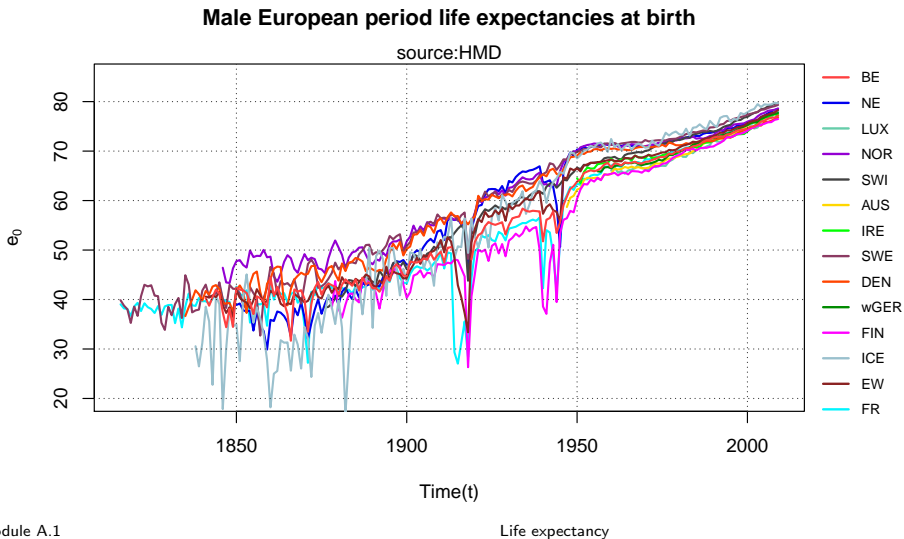
A useful expression under UDD

- Under the **UDD** assumption we find: (using  $\sum_{k \geq 0} {}_k p_x \cdot q_{x+k} = 1$ )

$$\begin{aligned}
 E[T_x] &= \int_0^{+\infty} {}_t p_x dt = \sum_{k \geq 0} {}_k p_x \int_{t=0}^1 {}_t p_{x+k} dt \\
 &= \sum_{k \geq 0} \int_{t=0}^1 {}_k p_x \cdot (1 - t \cdot q_{x+k}) dt \\
 &= \sum_{k \geq 0} {}_k p_x \left(1 - \frac{q_{x+k}}{2}\right) = -\frac{1}{2} + \sum_{k \geq 0} {}_k p_x.
 \end{aligned}$$







With constant interest rate, the Expected Present Value (EPV) of Actuarial Present Value (APV) of:

- a 1 EUR benefit whole life insurance product in **continuous time**

$$\bar{A}_x = E[e^{-\delta \cdot T_x}] = \int_0^{\infty} e^{-\delta \cdot t} \cdot {}_t p_x \cdot \mu_{x+t} dt$$

- a 1 EUR benefit whole life insurance product in **discrete time**

$$A_x = E[v^{K_x+1}] = \sum_{k=0}^{+\infty} v^{k+1} {}_k|q_x = v \cdot q_x + v^2 \cdot {}_1|q_x + v^3 \cdot {}_2|q_x + \dots,$$

where  ${}_k|q_x = {}_k p_x \cdot q_{x+k}$ .

With constant interest rate, the Expected Present Value (EPV) of Actuarial Present Value (APV) of:

- a 1 EUR benefit whole life annuity in **continuous time**

$$\bar{a}_x = \int_0^{\infty} e^{-\delta t} \cdot {}_t p_x dt.$$

- a 1 EUR benefit whole life annuity in **discrete time**

$$\begin{aligned}\ddot{a}_x &= 1 + v \cdot p_x + v^2 \cdot {}_2p_x + v^3 \cdot {}_3p_x + \dots \\ &= \sum_{k=0}^{\infty} v^k \cdot {}_k p_x.\end{aligned}$$

# Wrap-up

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Via this module you have acquired insights in:

- essential characteristics of the **distribution of the random variable  $T_x$** , the future lifetime r.v. of an  $(x)$ -year old:  $q_x, p_x, \mu_x, E[T_x]$
- how to extract or calculate these characteristics from (the entries in) a given **period life table**
- qualitative and **visual insights** in the evolution of mortality over ages and periods.

So far, we took a so-called **period approach** and considered the lifetime distribution or life table applicable to a specific period (e.g., a year).