

Introduction to Statistics (MAT 283)

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Table of Contents

SOME SPECIAL CONTINUOUS DISTRIBUTIONS CONTINUED

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3. The Normal (or Gaussian) Distribution:

A random variable X is said to have a normal distribution if its probability density function is given by

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}, \quad -\infty < x < \infty$$

with parameters μ , where $-\infty < \mu < \infty$ and $\sigma > 0$.

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If X has a normal distribution with parameters μ and σ^2 , then we write $X \sim N(\mu, \sigma^2)$.

The graph of a normal probability density function, shaped like the cross section of a **bell**.

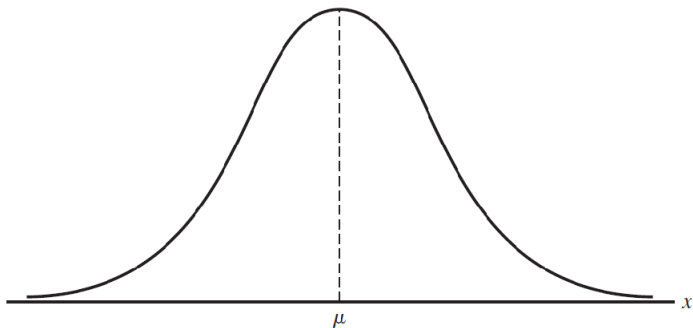


Figure Graph of normal distribution.

- From the form of the probability density function, we see that the density is **symmetric** about μ , $f(\mu - x) = f(\mu + x)$, where it has a **maximum**, and that the rate at which it falls off is determined by σ .

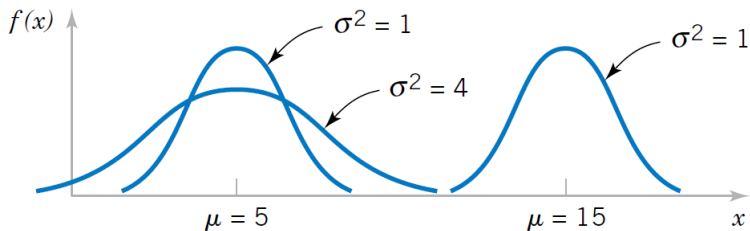


Figure Normal probability density functions for selected values of the parameters μ and σ^2 .

Theorem: If $X \sim N(\mu, \sigma^2)$, then

$$E(X) = \mu_X = \mu$$

$$\text{Var}(X) = \sigma_X^2 = \sigma^2$$

$$M_X(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}.$$

Proof.

$$\begin{aligned} M_X(t) &= \int_{-\infty}^{\infty} e^{tx} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \\ &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tx} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \end{aligned}$$

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&= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{tx} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx \\
&= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2\sigma^2}(-2xt\sigma^2 + (x-\mu)^2)} dx
\end{aligned}$$

Note that

$$-2xt\sigma^2 + (x - \mu)^2 = [x - (\mu t + \sigma^2)]^2 - 2\mu t\sigma^2 - t^2\sigma^4,$$

and hence

$$M_X(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2} \left\{ \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}\left(\frac{x-(\mu t + \sigma^2)}{\sigma}\right)^2} dx \right\}$$

Since the quantity inside the bracket is the integral from $-\infty$ to $-\infty$ of a normal probability density function with the parameters $\mu + t\sigma^2$ and σ , and hence is equal to 1, it follows that

$$M_X(t) = e^{\mu t + \frac{1}{2}\sigma^2 t^2}.$$

Further,

$$M'_X(t) = (\mu + \sigma^2 t)e^{\mu t + \frac{1}{2}\sigma^2 t^2}.$$

$$M''_X(t) = \sigma^2 e^{\mu t + \frac{1}{2}\sigma^2 t^2} + (\mu + \sigma^2 t)^2 e^{\mu t + \frac{1}{2}\sigma^2 t^2}.$$

$$\implies E(X) = M'_X(0) = \mu \text{ and } \text{Var}(X) = \sigma^2.$$

Hence the proof.

Standard Normal Random Variable: A normal random variable is said to be **standard normal**, if its mean is zero and variance is one. We denote a standard normal random variable X by $X \sim N(0, 1)$ and its probability density function is given by

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}, \quad -\infty < x < \infty.$$

Example: If $X \sim N(0, 1)$, what is the probability of the random variable X less than or equal to -1.72 ?

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Using Standard Normal Table type II,

$$P(X \leq -1.72) = 0.0427.$$

- Probabilities that are not of the form $P(Z \leq z)$ are found by using the basic rules of probability and the symmetry of the normal distribution.

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Examples:

1. $P(Z > 1.26)$
2. $P(Z > -1.37)$
3. $P(Z < -0.86)$
4. $P(-1.25 < z < 0.37)$

Theorem: If $X \sim N(\mu, \sigma^2)$ then the random variable $Z = \frac{X - \mu}{\sigma} \sim N(0, 1)$.

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Proof: We will show that Z is standard normal by finding the probability density function of Z . We compute the probability density of Z by first computing it's cumulative distribution function.

$$\begin{aligned} F(z) &= P(Z \leq z) \\ &= P\left(\frac{X - \mu}{\sigma} \leq z\right) \\ &= P(X \leq \sigma z + \mu) \\ &= \int_{-\infty}^{\sigma z + \mu} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x - \mu}{\sigma}\right)^2} dx \end{aligned}$$

$$= \int_{-\infty}^z \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}w^2} dw,$$

(where $w = \frac{x-\mu}{\sigma}$)

Hence

$$f(z) = F'(z) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}z^2}.$$

Hence the proof.

Example: If $X \sim N(3, 16)$, then what is $P(4 \leq X \leq 8)$?

$$P(4 \leq X \leq 8) = P\left(\frac{4-3}{4} \leq \frac{X-3}{4} \leq \frac{8-3}{4}\right)$$

$$= P\left(\frac{1}{4} \leq Z \leq \frac{5}{4}\right)$$

$$= P(Z \leq 1.25) - P(Z \leq 0.25)$$

$$= 0.8944 - 0.5987$$

$$= 0.2957.$$

GAMMA DISTRIBUTION:

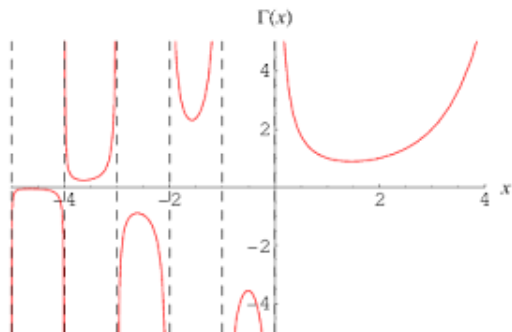
The gamma distribution involves the notion of **gamma function**. The gamma function, $\Gamma(z)$, is a generalization of the notion of **factorial**. The gamma function is defined as

$$\Gamma(z) = \int_0^{\infty} x^{z-1} e^{-x} dx,$$

where z is positive real number (that is, $z > 0$). The condition $z > 0$ is assumed for the convergence of the integral.

Although the integral does not converge for $z < 0$, it can be shown by using an alternative definition of gamma function that it is defined for all $z \in \mathbb{R}/\{0, -1, -2, -3, \dots\}$.

The graph of the gamma function is shown below.



Properties of Gamma Function:

1. $\Gamma(1) = 1$.

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3. $\Gamma(\frac{1}{2}) = \sqrt{\pi}$.
4. $\Gamma(-\frac{1}{2}) = -2\sqrt{\pi}$.
5. If n is a natural number then $\Gamma(n + 1) = n!$.

GAMMA DISTRIBUTION:

A continuous random variable X is said to have a gamma distribution if its probability density function is given by

$$f(x) = \begin{cases} \frac{1}{\Gamma(\alpha)\theta^\alpha} x^{\alpha-1} e^{-\frac{x}{\theta}} & \text{if } 0 < x < \infty \\ 0 & \text{otherwise,} \end{cases}$$

where $\theta > 0$ and $\alpha > 0$.

GAMMA DISTRIBUTION:

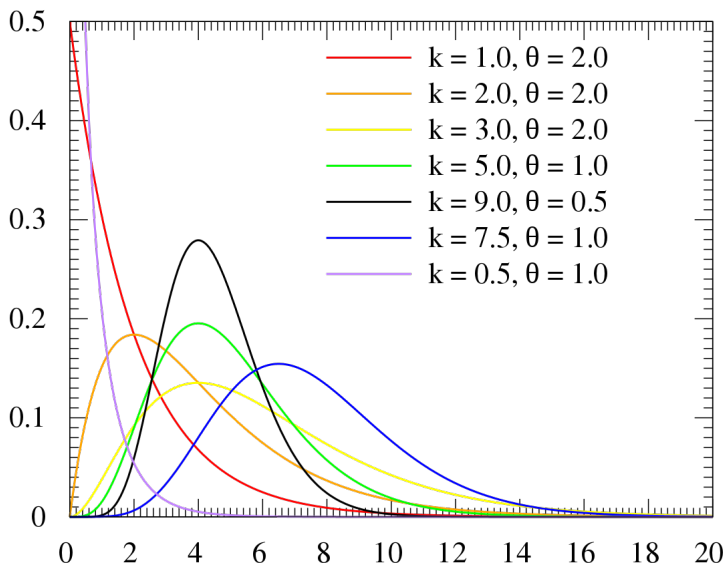
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where $\theta > 0$ and $\alpha > 0$.

We denote a random variable X with the Gamma distribution as $X \sim \text{GAM}(\theta, \alpha)$.

- Some special cases of the gamma distribution play important roles in statistics; for instance, for $\alpha = 1$ and $\beta = \theta$, we obtain the **exponential distribution**.



Here $k = \alpha$.

THEOREM: If $X \sim \text{GAM}(\theta, \alpha)$, then the mean, variance and moment generating functions are respectively given by

$$E(X) = \mu_X = \theta\alpha$$

$$\text{Var}(X) = \theta^2\alpha$$

$$M_X(t) = \left(\frac{1}{1 - \theta t}\right)^\alpha, \text{ if } t < \frac{1}{\theta}.$$

BETA DISTRIBUTION:

The beta distribution involves the notion of **beta function**. First we explain the notion of the beta integral and some of its simple properties. Let α and β be any two positive real numbers. The beta function $B(\alpha, \beta)$ is defined as

$$B(\alpha, \beta) = \int_0^1 x^{\alpha-1} (1-x)^{\beta-1} dx.$$

Properties of beta function:

1. Let α and β be any two positive real numbers. Then

$$B(\alpha, \beta) = \frac{\Gamma(\alpha)\Gamma(\beta)}{\Gamma(\alpha + \beta)}.$$

2. Let α and β be any two positive real numbers. Then

$$B(\alpha, \beta) = B(\beta, \alpha).$$

BETA DISTRIBUTION:

A continuous random variable X is said to have a beta distribution if its probability density function is given by

$$f(x) = \begin{cases} \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise,} \end{cases}$$

where $\alpha > 0$ and $\beta > 0$.

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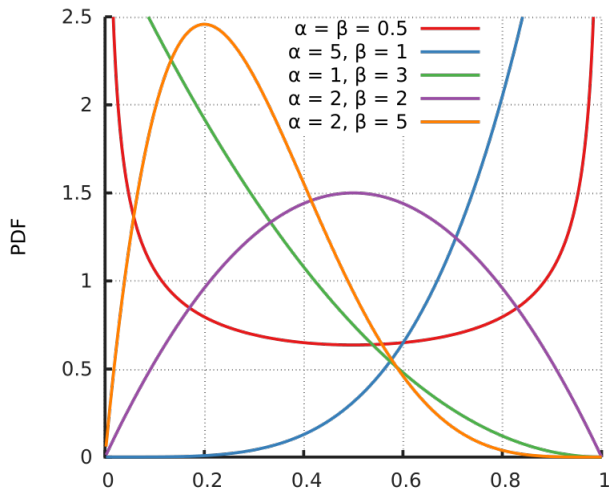
$$f(x) = \begin{cases} \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} & \text{if } 0 < x < 1 \\ 0 & \text{otherwise,} \end{cases}$$

where $\alpha > 0$ and $\beta > 0$.

We denote a random variable X with the Beta distribution as $X \sim \text{BETA}(\theta, \alpha)$.

- Some special cases of the beta distribution play important roles in statistics; for instance, for $\alpha = 1$ and $\beta = 1$, we obtain the **uniform distribution** over $(0, 1)$.

Graph of Beta Distribution:



THEOREM: If $X \sim \text{BETA}(\alpha, \beta)$, then the mean and variance are respectively given by

$$E(X) = \mu_X = \frac{\alpha}{\alpha + \beta}$$

$$\text{Var}(X) = \frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}.$$