Introduction to Statistics (MAT 283)

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PRODUCT MOMENTS AND MOMENT GENERATING FUNCTIONS

EXAMPLE: A fair coin is tossed two times; let X denote the number of heads on the first toss and Y the total number of heads. Sample space of this random experiment is

$$\Omega = \{HH, HT, TH, TT\}.$$

The joint PMF f of X and Y is as given in the following table:

	Y=0	Y=1	Y=2	P(X=x)
X= 0	$\frac{1}{4}$	$\frac{1}{4}$	0	<u>2</u> 4
X= 1	0	$\frac{1}{4}$	$\frac{1}{4}$	<u>2</u> 4
P(Y=y)	$\frac{1}{4}$	$\frac{2}{4}$	$\frac{1}{4}$	

Then find the conditional mean of X when Y = 0.

$$E(X|Y=0) = \sum_{\mathbf{x} \in R_X} \mathbf{x} \, g(\mathbf{x}|y=0)$$

$$g(x|y=0) = \frac{f(x,0)}{f_{y}(0)} = \frac{f(x,0)}{f_{y}(0)} = \frac{1/4}{1/4} = 1$$

$$g(x=1)|y=0| = \frac{f(x,0)}{f_{y}(0)} = 0$$

$$f(x)|y=0| = 0 \times 1 + 1 \times 0 = 0$$

EXAMPLE: Let the random variables X and Y have the joint pdf

$$f(x,y) = \begin{cases} x + y & \text{for} \quad 0 < x, y < 1 \\ 0 & \text{otherwise,} \end{cases}$$

What is the conditional mean $E(Y|X = \frac{1}{3})$?

$$h(y|x=\frac{1}{3}) = \frac{f(y_3,y)}{f_x(y_3)}$$

$$f_{\mathbf{x}}(\mathbf{x}) = \int_{\mathbf{x}} f(\mathbf{x}, \mathbf{y}) d\mathbf{y}$$

$$= \left[xy + \frac{y^2}{2} \right]_0^1 = x + \frac{1}{2}$$

$$h(y|x) = \frac{f(x,y)}{f_x(x)} = \frac{x+y}{x+y_2}$$

$$E(Y|x = \frac{1}{3}) = \int_0^1 y h(h|x = y_2) dy$$

 $=\int_{3}^{1}y\frac{1/3+3}{1/3+1/2}dy$

 $= \int_{1}^{6} \frac{3}{1/3} + \frac{1}{1/3} + \frac{1}$

=] (m+y) dy

$$=\frac{6}{5}\int_{a}^{1}\left(\frac{1}{3}y+y^{2}\right)dy$$

$$= \frac{6}{5} \int_{3}^{3} \left(\frac{1}{3} + \frac{3}{4} + \frac{3}{4} \right) = \frac{6}{5} \times \frac{3}{6} = \frac{6}{5} \times \frac{3}{6} = \frac{1}{6} = \frac{1}{6} \times \frac{3}{6} = \frac{1}{6} = \frac{1}{6} \times \frac{3}{6} = \frac{1}{6} \times \frac{3}{6} = \frac{1}{6}$$

Let X and Y be random variables defined on a probability space (Ω, \mathcal{F}, P) , let ψ be a Borel-measurable function. Then the conditional expectation of $\psi(X)$, given Y, written as $E(\psi(X)|Y)$, is a random variable that takes the value $E(\psi(X)|y)$, defined as

$$E(\psi(X)|y) = \begin{cases} \sum_{x \in R_X} \psi(x) \ g(x|y) & \text{if X is discrete} \\ \\ \int_{-\infty}^{\infty} \psi(x) \ g(x|y) dx & \text{if X is continuous,} \end{cases}$$

a similar definition may be given for the conditional expectation $E(\psi(Y)|X)$.

The conditional mean of Y given X=x, that is E(Y|x), is a function $\phi(x)$ of the variable x.

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Define

$$\phi(X) = E(Y|X),$$

this function $\phi(X)$ is a random variable.

Theorem. The expected value of the random variable E(Y|X) is equal to the expected value of Y, that is

$$E_X(E(Y|X)) = E_Y(Y),$$

where $E_X(X)$ stands for the expectation of X with respect to the distribution of X and $E_Y(Y)$ stands for the expectation of Y with respect to the distribution of Y.

Proof. We prove this theorem for continuous case.

$$E_X(E(Y|X)) = \int_{-\infty}^{\infty} E(Y|x) f_X(x) dx$$
$$= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} y \ h(y|x) dy \right) f_X(x) dx$$

$$= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} y \ h(y|x) f_X(x) dy \right) dx \quad \left[\vdots \ h \left(\frac{y}{x} \right) \right]$$

$$= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} \frac{h(y|x) f_X(x) dx}{h(y|x) f_X(x) dx} \right) y dy \quad \left[\frac{1}{2} \frac{f(x,y)}{f(x,y) dx} \right]$$

$$= \int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} f(x,y) dx \right) y dy$$

$$= \int_{-\infty}^{\infty} y f_Y(y) dy$$

$$= E_Y(Y).$$

EXAMPLE: An insect lays Y number of eggs, where Y has a Poisson distribution with parameter λ . If the probability of each egg surviving is p, then on the average how many eggs will survive?

Solution: Let X denote the number of surviving eggs.

EXAMPLE: An insect lays Y number of eggs, where Y has a Poisson distribution with parameter λ . If the probability of each egg surviving is p, then on the average how many eggs will survive?

Solution: Let X denote the number of surviving eggs. Then, given that Y = y (that is given that the insect has laid y eggs) the random variable X has a binomial distribution with parameters y and p. Thus

$$X|Y \sim BIN(Y,p)$$

and

$$Y \sim POI(\lambda)$$

Therefore, the expected number of survivors is given by

$$E_X(X) = E_Y(E(X|Y))$$

= $E_Y(pY)$ (since $X|Y \sim BIN(Y, p)$)
= $pE_Y(Y)$
= $p\lambda$ (since $Y \sim POI(\lambda)$).

Theorem: Let X and Y be two random variables with mean μ_X and μ_Y , and standard deviation σ_X and σ_Y , respectively. If the conditional expectation of Y given X=x is linear in x, then

$$E(Y|X=x) = \mu_Y + \rho \frac{\sigma_X}{\sigma_Y}(x - \mu_X)$$

where ρ denotes the correlation coefficient of X and Y.

Proof: Probability and Mathematical Statistics by Sahoo, Page No. 243-244.

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CONDITIONAL VARIANCE:

Let X and Y be two random variables with joint density f and h(y|x) be the conditional density of Y given X=x. The conditional variance of Y given X=x, denoted by Var(Y|x), is defined as

$$Var(Y|x) = E(Y^2|x) - (E(Y|x))^2$$

where E(Y|x) denotes the conditional mean of Y given X = x.

EXAMPLE: A fair coin is tossed two times; let X denote the number of heads on the first toss and Y the total number of heads. Sample space of this random experiment is

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The joint PMF f of X and Y is as given in the following table:

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Find the conditional variance of Y given x = 1.

EXAMPLE: If the joint probability density of X and Y is given by

$$f(x,y) = \begin{cases} \frac{2}{3}(x+2y) & \text{for } 0 < x < 1; \ 0 < y < 1 \\ 0 & \text{otherwise,} \end{cases}$$

find the conditional mean and the conditional variance of X given $Y=\frac{1}{2}.$