

Problem 1 The data in the table is believed to obey a model of the form $y_i = ax_i^2 + bx_i + c + e_i$ with e_i assumed independent and identically distributed errors with $e_i \sim N(0, \sigma^2)$. Estimate a , b , and c and assign appropriate errorbars to your estimates. Also estimate the value of $y(t = 12)$ and assign an appropriate errorbar to the estimate.

t	y
1.0	-2.73
2.0	-2.71
3.0	-2.65
4.0	-0.87
5.0	-3.10
6.0	-1.03
7.0	0.63
8.0	1.46
9.0	5.90
10.0	8.38

My notes are messy I know so I also typed out the general procedure.

- Estimate a , b , and c and assign appropriate error bars:

Jeffrey's Prior assumed, likelihood fcn given in notes, can be expanded for each observation, but simplifies to (see notes). $y = A \cdot \text{beta} + e$ and we can now solve for e and beta (parameter vector $[a \ b \ c]^T$). See notes for design matrix A (t_i are predictor variable values). e is the error vector, and $e = y - A \cdot \text{beta}$, so we simplify likelihood function again (in notes, it's the one that starts with $1/(2\pi)^{N/2} \cdot \sigma^{-N}$).

The posterior distribution is product of likelihood and prior normalized by $\text{Prob}(\text{data})$ (evidence). We simplify the expression by making posterior proportional to product of likelihood and prior but omitting evidence. This is allowed because the evidence ($\text{Prob}(\text{data})$) doesn't depend on estimated parameters and can be treated like a constant when we are doing Bayesian proportionality.

The Jeffrey's Prior doesn't prefer any value of σ over another value of σ . Integrate out σ to get marginal posterior dist. For beta , and simplify the integral. In simplified expression we introduce "neu", denoted here as V : $V(\text{beta}) = (A \cdot \text{beta} - y)^T (A \cdot \text{beta} - y)$. k : # of parameters. Now we have the post. Dist. For beta is proportionate to $1 / ((e^e)^{(-3/2)})$. It follows a student t distribution which has heavier tails than normal distribution.

Now revisit previous notes on posterior computed for this model. (we have to add a $1/(\sigma^{N+1})$ term to multiply the generalized form due to uncertainty in σ).

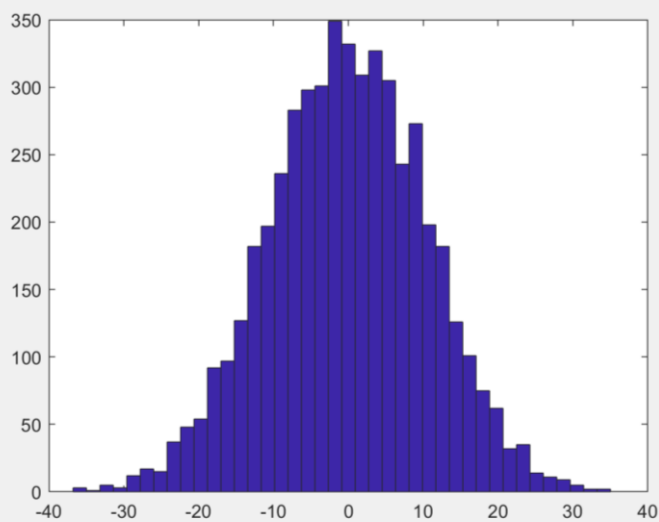
Beta_hat is the most probable values of beta , minimize the sum of squared residuals. To find most probable values for the parameters in beta , we take log of both sides of posterior distribution. Then take deriv. wrt/ beta and set it equal to 0 to find B_hat . Plug beta_hat back into model or evaluating model @ beta_hat gives us the most likely values of the parameters. This involves normal equations and shows that it is consistent with the solution from least squares/linear regression approach.

After we find beta_hat , we find σ_hat by taking the log of the posterior distribution and differentiating wrt/ σ , setting that equal to zero, and solving for σ . See notes for final formula for σ_hat^2 (variance estimate). We can use the beta_hat that we found previously.

- Estimate $y(t=12)$ and assign error bar:

To estimate y at $t=12$ (Bayesian inference), we derive the predictive dist. For y at a new point using the posterior dist. Of beta . See notes for formula for $\text{Prob}(y_{12} | \text{beta})$..It includes a design matrix with a constant, linear and quadratic term for t . This is the probability dist of y_{12} given beta . But we don't actually know true beta .

So we must consider the posterior predictive dist. Of y at $t = 12$, incorporating the uncertainty in beta . (see notes). Then we integrate out σ and get a distribution proportional to (see notes, it includes H , which = $A^T A / \sigma^2$). Probability of y_{12} given data is equal to a triple integral from $-\infty$ to ∞ (see notes), to address all uncertainty (in all parameters a b and c). A new symbol (del? Squiggle? Dirac delta?) . To solve the triple integral I use matlab to simulate samples from a t -distribution and take into account uncertainty in a b and c .



SOLUTION

Command Window

```
a = 2.223864e-01 +/- 5.231020e-02
b = -1.310614e+00 +/- 5.904336e-01
c = -1.025500e+00 +/- 1.413731e+00
```

```
error_12 =
```

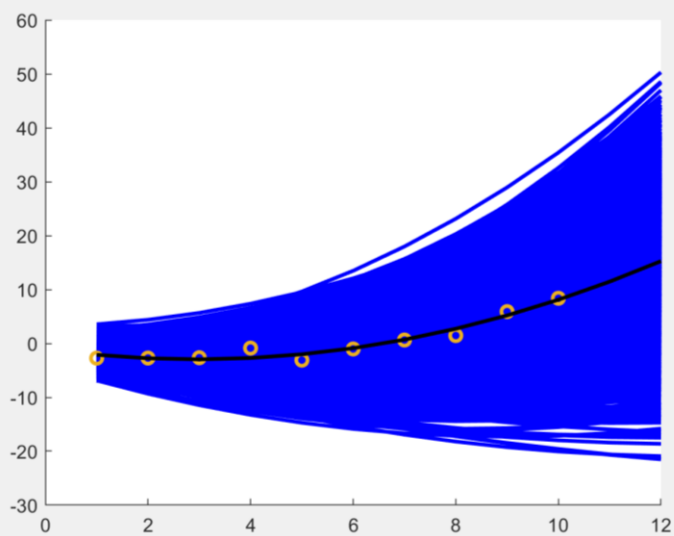
```
10.5416
```

```
>> d12
```

```
d12 =
```

```
15.2708
```

- Estimates for parameters a, b, and c with errorbars
- d12(y at t = 12) with error_12 being the associated error



Notes from class below.

ASSIGN PRIOR HW #4

① $\text{prob}(a, b, c, \log \sigma) \propto \text{const.}$
 $\sigma) \propto 1/\sigma$

$\text{prob}(\log \sigma) d \log \sigma = \text{prob}(\sigma) d\sigma$

$\left| \frac{d \log \sigma}{d\sigma} \right| = \frac{1}{\sigma}$

② Likelihood fun

$\rightarrow \text{prob}(\text{data} | a, b, c, \sigma)$

how does data vary w/ (the aspects of the model)

③ find posterior:

$\text{prob}(a, b, c, \sigma | \text{data}) \propto$

model conditioned on data

product rule

$\text{prob}(A|B) \text{prob}(B) = \text{prob}(A \cap B) \text{prob}(A) =$

$\text{prob}(A + B)$

$\rightarrow \propto \frac{\text{likelihood} \cdot \text{prior}}{\text{prob}(\text{data})} \propto \frac{\text{prob}(\text{data} | a, b, c, \sigma) \cdot \text{prob}(a, b, c, \sigma)}{\text{prob}(\text{data})}$

$y_i = at_i^2 + bt_i + c + e_i$

$e_i \sim N(0, \sigma^2)$

$\beta = \begin{bmatrix} c \\ b \\ a \end{bmatrix}$

$\hat{y} = A\beta + e$

rewrite

$A = \begin{bmatrix} 1 & t_1 & t_1^2 \\ \vdots & \vdots & \vdots \\ 1 & t_{10} & t_{10}^2 \end{bmatrix} = [t.^{[0 1 2]}]$

IN MATLAB

④ write prob distribution fun:

likelihood

$\text{prob}(\text{data} | a, b, c, \sigma) \propto \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2} \frac{(y_i - (at_i^2 + bt_i + c))^2}{\sigma^2}\right)$

$$\cdot \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2} \text{same as prev but } 2, 3, \dots, 10\right)$$

Simplify!

Sum of squares of difference between

$$= \frac{1}{\sqrt{2\pi}(2\pi)^5 \sigma^{10}} \cdot \exp\left(-\frac{1}{2} \left(\underline{e}\right)^T \left(\underline{e}\right)\right)$$

Solve for \underline{e}

dot product here gives sum of squares data + model

$$\underline{e} = (\underline{y} - \underline{A}\underline{\beta})$$

(data - model)

$$\Rightarrow \frac{1}{(2\pi)^5 \sigma^{10}} \exp\left(-\frac{1}{2\sigma^2} (\underline{y} - \underline{A}\underline{\beta})^T (\underline{y} - \underline{A}\underline{\beta})\right)$$

likelihood

posterior

$$\text{prob}(\underline{\beta}, \sigma | \text{data}) \propto \frac{1}{\sigma^{11}} \left(\text{but } \sigma^2 \right)$$

!! bc degrees of freedom?

proportionality lets you drop

MATLAB

design matrix col of 15

the 2π term

$$p = [t.^1, t.^2]$$

$$lm = \text{fitlm}(p, \underline{y});$$

(col ones implicit)

LINEAR MODEL

$$\hat{\underline{a}} = \text{lm.Coefficients. Estimate} \quad (3)$$

backwards?

b(2) c(1)

$$\hat{\underline{B}} = \underline{A} \backslash \underline{y}$$

We want to

$$= (\underline{A}'\underline{A}) \backslash \underline{A}'\underline{y}$$

known error bar for $\hat{\underline{a}}$

$$a = \hat{a} \pm \sigma_a$$

$$b = \hat{b} \pm \sigma_b$$

$$c = \hat{c} \pm \sigma_c$$

assign std dev to a, b, c

$$\sigma_a =$$

$$\text{Im. coefficients. SE}(3) \int \frac{1}{\sqrt{y-2}}$$

$$V = N - K$$

"na"

params (a, b, c)

$$K = 3$$

degrees of freedom what N again?

10?

We want to integrate σ explicitly
to get new PDF?

$$\text{prob}(\underline{\beta} | \text{data}) = \int_{-\infty}^{\infty} d \log \sigma \{ \text{prob}(\underline{\beta}, \log \sigma | \text{data}) \}$$

Sum rule
weighted avg taking into account we don't know σ

$$\propto V(\underline{\beta})^{-K/2} \text{ where } V(\underline{\beta}) = (\underline{A}\underline{\beta} - \underline{y})^T (\underline{A}\underline{\beta} - \underline{y})$$

$$K = \text{length}(\underline{\beta})$$

$$\frac{1}{(e^T e)^{-3/2}} \mathcal{N}$$

find where

$$\hat{\sigma} \text{ deriv } a, b, c = 0 \quad \text{TO FIND}$$

deriv of posterior

$\hat{\sigma}$:

for finding

$$\mathcal{L} = \log(\text{prob}(\underline{\beta}, \sigma | \text{data}))$$

$$\frac{\partial \mathcal{L}}{\partial \underline{\beta}} = -\frac{1}{2\sigma^2} (\underline{A}\underline{\beta} - \underline{y})^T \underline{A}$$

sets end of factor of 2

$$\mathcal{L} = -\frac{1}{2\sigma^2} (\underline{A}\underline{\beta} - \underline{y})^T (\underline{A}\underline{\beta} - \underline{y}) - (N+1) \log \sigma$$

$$\frac{\partial \mathcal{L}}{\partial \sigma} = +\frac{1}{4\sigma^3} (\underline{A}\underline{\beta} - \underline{y})^T (\underline{A}\underline{\beta} - \underline{y}) - \frac{(N+1)}{\sigma}$$

Now 4 eqns, 4 unknowns

$$\text{Set } \left. \frac{\partial \mathcal{L}}{\partial \sigma} \right| = 0 \quad \vee \quad \left. \frac{\partial \mathcal{L}}{\partial \underline{\beta}} \right| = 0 \Rightarrow \hat{\underline{\beta}}$$

$$\sigma = \hat{\sigma}$$

$$\underline{\beta} = \hat{\underline{\beta}}$$

$$\underline{\beta} = (\underline{A}^T \underline{A})^{-1} \underline{A}^T \underline{y}$$

$$\underline{\beta} = \frac{1}{4} (\underline{A}\hat{\underline{\beta}} - \underline{y})^T (\underline{A}\hat{\underline{\beta}} - \underline{y}) -$$

otnd from Monday week 18

$$\underline{\underline{B}} = \underline{\underline{A}} \backslash \underline{\underline{y}} = (A'A) \backslash (A'y)$$

$$a = \hat{a} \pm \sigma_a \dots b + c \text{ too}$$

find in matlab

$$\text{prob}(B | \text{data}) = \int_{-\infty}^{\infty} d\log \sigma \sum_{-\infty}^{\infty} P(B | \log \sigma | \text{data})$$

$$= \int_{-\infty}^{\infty} d\sigma \sum_{-\infty}^{\infty} P(B | \sigma | \text{data})$$

$$\propto V(B)^{-K/2} \text{ where}$$

$$V(B) = (AB - y)'(AB - y)$$

$$K = \text{length}(B)$$

$$\frac{\partial \mathcal{L}}{\partial B} \Big| = \frac{1}{\sigma^3} (AB - y)^T (AB - y) - \frac{N+1}{\sigma} = 0$$

$$\frac{\partial \mathcal{L}}{\partial B} \Big| = \frac{1}{\sigma^3} (AB - y)^T (AB - y) - \frac{N+1}{\sigma} = 0$$

$$B = \hat{B}$$

$$\sigma = \hat{\sigma}$$

$$\Rightarrow \frac{1}{4} (A\hat{B} - y)^T (A\hat{B} - y) - \hat{\sigma}^2 (N+1) = 0$$

$$\hat{\sigma}^2 = \frac{(A\hat{B} - y)^T (A\hat{B} - y)}{(N+1)}$$

3/11/24

1) $y_i = ax_i^2 + bx_i + c + e_i \leftarrow \text{MODEL}$

e_i assumed independent

$e_i \sim N(0, \sigma^2)$

- estimate a, b, c + error bars
- estimate value $y(t=12)$ + error bar

First, ASSIGN PRIOR:

$\text{prob}(a, b, c, \log \sigma) \propto \text{const}$

$\text{prob}(a, b, c, \sigma) \propto 1/\sigma$

$\Rightarrow \text{prob}(\log \sigma) d \log \sigma = \text{prob}(\sigma) d\sigma$

Second LIKELIHOOD FXN:

$\text{prob}(\text{data} | a, b, c, \sigma)$

third find POSTERIOR:

$\text{prob}(a, b, c, \sigma | \text{data}) \propto \text{prob}(A|B) \text{prob}(B)$

(apply product rule)
model params

$\Rightarrow = \text{prob}(B|A) \text{prob}(A) = \text{prob}(A + B)$

$\Rightarrow \propto \text{likelihood} \cdot \text{prior} \propto \text{prob}(\text{data} | a, b, c, \sigma)$

$\text{prob}(\text{data})$

evidence

$\propto \text{prob}(a, b, c, \sigma)$

Switch notation of model x s to t s

Rewrite model:

$y = \underline{A} \underline{B} + \underline{e}$

$\underline{A} = \begin{bmatrix} 1 & t_1 & t_1^2 \\ 1 & t_2 & t_2^2 \\ 1 & t_3 & t_3^2 \end{bmatrix}$

$\underline{B} = \begin{bmatrix} c \\ b \\ a \end{bmatrix}$

Rewrite model:

Next, write PROB DISTRIBUTION FXN:

Likelihood

$$\text{prob}(\text{data} | a b c \sigma) \propto \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2} \frac{(y_i - (at_i^2 + bt) + c))^2}{\sigma^2}\right)$$

$$\rightarrow \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2} \frac{(y_i - (at_i^2 + bt) + c))^2}{\sigma^2}\right)$$

SIMPLIFY!

$$= \frac{1}{(2\pi)^5 \sigma^{10}} \exp\left(-\frac{1}{2} (\underline{e})^T (\underline{e})\right)$$

SOLVE FOR \underline{e}

$$\underline{e} = (\underline{y} - \underline{A}\underline{\beta}) \leftarrow \text{data} - \text{model}$$

LIKELIHOOD

$$\rightarrow \frac{1}{(2\pi)^5 \sigma^{10}} \exp\left(-\frac{1}{2\sigma^2} (\underline{y} - \underline{A}\underline{\beta})^T (\underline{y} - \underline{A}\underline{\beta})\right)$$

POSTERIOR

$$\text{prob}(\underline{\beta}, \sigma | \text{data}) \propto \frac{1}{\sigma^{11}} \exp\left(-\frac{1}{2\sigma^2} (\underline{y} - \underline{A}\underline{\beta})^T (\underline{y} - \underline{A}\underline{\beta})\right)$$

can
drop the
 2π

degrees prior

FIT MODEL

IN MATLAB

+ find \hat{a} \hat{b} \hat{c} +
their error bars

$$\hat{\underline{\beta}} = \underline{A} \backslash \underline{y} = (\underline{A}'\underline{A}) \backslash \underline{A}'\underline{y} \quad \text{ex. } a = \hat{a} \pm \sigma_a$$

in matlab, \searrow
same as

Gaussian

3113124

$$\text{prob}(\beta, \sigma | d) \propto \frac{1}{\sigma^{N+1}} \exp\left(-\frac{1}{2\sigma^2} (A\beta - d)' (A\beta - d)\right)$$

posterior

comes from prior

$$\text{prob}(y(12) | \beta) = \text{assume random errors are noise +}$$

$$\int \left([1 \ 12 \ 12^2] \beta - y(12) \right) \text{ we know true value of } \beta$$

prob. dist.

of β width would be no uncertainty

but we don't know β

$$\text{prob}(y(12), \beta | d) = \text{prob}(y(12) | \beta) \text{prob}(\beta | d)$$

product rule

integral

$$\propto \int d\sigma \text{prob}(\beta, \sigma | d) \propto [U + (\beta - \hat{\beta})^T H (\beta - \hat{\beta})]^{-N/2}$$

k : # params N : # data $\Rightarrow 10$

$a \leq c \Rightarrow 3$

$U = N - k$

H : precision of multivar. + dist

$$H = \frac{A^T A}{S^2}$$

$$A = [t^{[0, 1, 2]}]$$

i.e.
inv.
cov.
mat.

$$S^2 = (d - A\hat{\beta})^T (d - A\hat{\beta})$$

makes error vect smallest U (at over $N-U$)

$$\hat{\beta} = (A^T A)^{-1} A^T d \leftarrow \text{Least Sq Est}$$

location of max of posterior

$$S^2 = \frac{e^T e}{U}$$

Block lines have integrated out σ
 $C =$ off diagonals are covariance
 3×3 (bc a, b, c)

CM = correlation matrix

1. also 3×3

$$\text{prob}(\gamma(12) | \underline{\beta}, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2\sigma^2}(\gamma(12) - \underline{A}\underline{\beta})^2\right)$$

$$\Rightarrow (\gamma(12) - \underline{A}\underline{\beta})$$

$$(\underline{A}\underline{\beta} - d)$$

$$2. \text{prob}(\underline{\beta}, \sigma, d) \propto \frac{1}{\sigma^{N+1}} \exp\left(-\frac{1}{2\sigma^2}(\underline{A}\underline{\beta} - d)^T \underline{V}(\underline{A}\underline{\beta} - d)\right)$$

combine 1 + 2 to get: $\text{prob}(\gamma(12), \underline{\beta}, \sigma | d)$ prob of projection

S : sqrt diag elements of matrix

in code

$$S = \begin{bmatrix} \sigma_a & 0 & 0 \\ 0 & \sigma_b & 0 \\ 0 & 0 & \sigma_c \end{bmatrix} \quad S \backslash C \backslash S \Rightarrow S^{-1} C S^{-1}$$

$$\Rightarrow \underline{\underline{CM}} \rightarrow \text{correlation matrix}$$

(All prob

#1, Now prob #2) (dimensionless)

ignore last 3 points, unlike $\underline{\underline{C}}$

use that model as a prior

(on prev page)

$$\text{prob}(y(12), \beta | d) =$$

$$\text{prob}(y(12) | d) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} da db dc (\delta y(12) - \underline{A}\underline{\beta})$$

$$\rightarrow [1 + [\beta - \hat{\beta}]^T \text{all params} \underline{H} (\beta - \hat{\beta})]^{-N/2}$$

then can plot histogram / probability. Solve in MATLAB by drawing random samples

assume prior

↳ a b c all equally uncertain
↳ errors in model Random + indep

now assume there's an actual rain process happening (not random)

$$\text{prob}(y(12), \beta, \sigma) = \frac{1}{\sqrt{2\pi} \sigma^2} \exp\left(-\frac{1}{2} (y(12) - \underline{A}\underline{\beta})' (y(12) - \underline{A}\underline{\beta})\right)$$

$$\text{prob}(y(12), \beta, \sigma | d) = \text{prob}(y(12) | \beta, \sigma) \text{prob}(\beta, \sigma | d)$$

$$\propto \frac{1}{\sigma^{N+2}} \exp\left(-\frac{1}{2\sigma^2} (y(12) - \underline{A}\underline{\beta})' (y(12) - \underline{A}\underline{\beta})\right) - \frac{1}{2\sigma^2} (\underline{A}\underline{\beta} - \underline{d})' (\underline{A}\underline{\beta} - \underline{d})$$

Now can integrate out

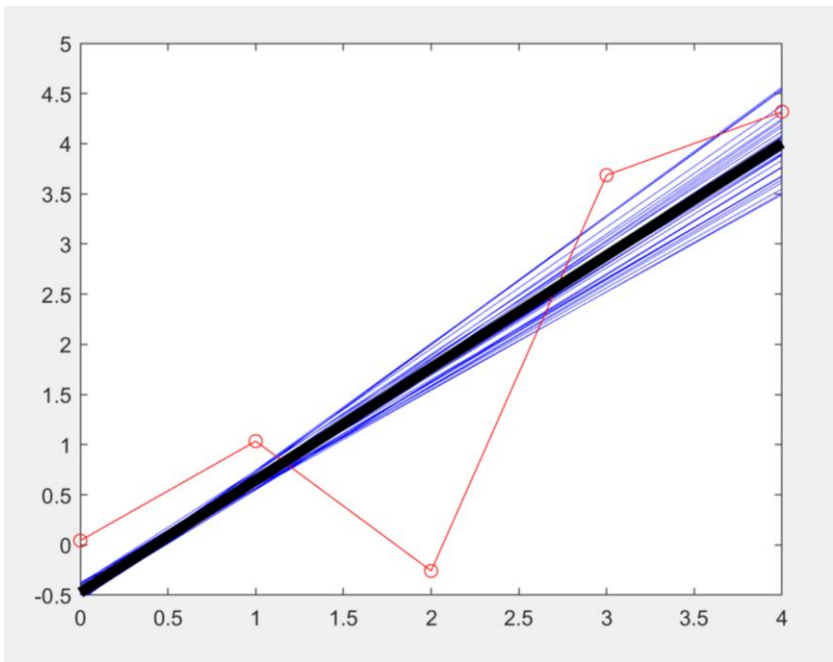
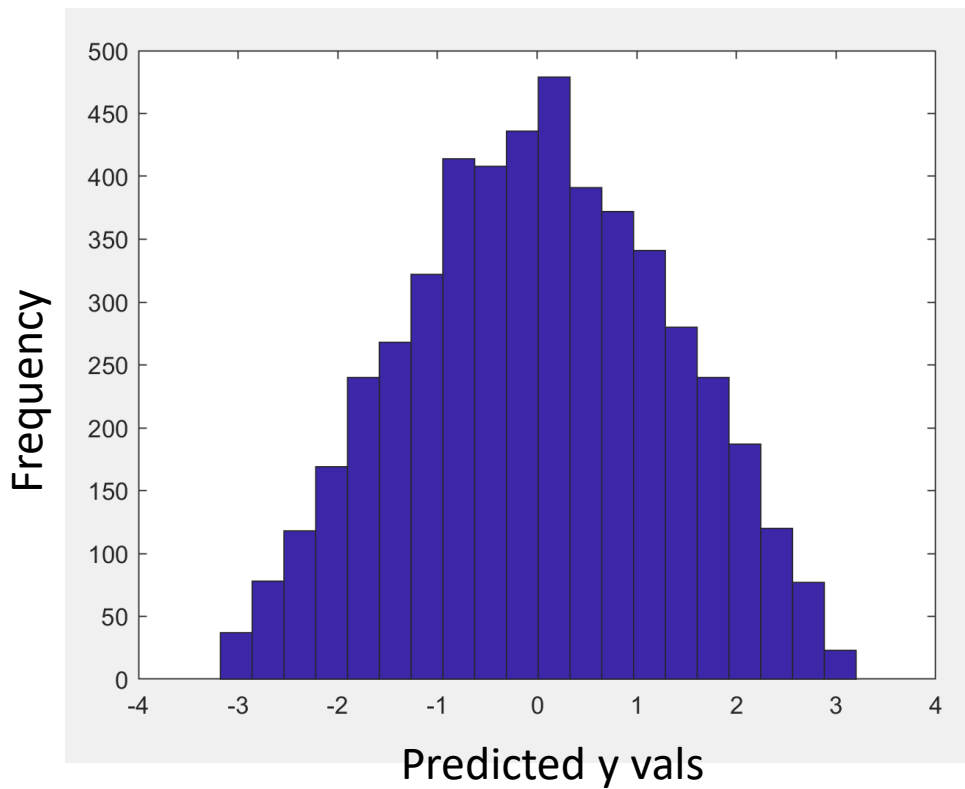
$\beta + \sigma$ + get $\text{prob}(y(12))$

Problem 2 The data in the table is believed to obey a model of the form $y_i = mx_i + b + e_i$ where e_i is a gaussian error term. For the first two measurements the error is known and indicated in the table. For the last three measurements the error is not known, but is assumed to have a constant variance.

x	y
0	0.0434 ± 0.1
1	1.0343 ± 0.1
2	$-0.2588 \pm \sigma$
3	$3.68622 \pm \sigma$
4	$4.3188 \pm \sigma$

Estimate m and b and assign appropriate errorbars to your estimates.

$m = 1.12027$
 $b = -0.47576$
 $\text{sigma}_m = 0.7250$
 $\text{sigma}_b = 1.7759$
 $\text{errorbar_x2} = 1.1125$
 $\text{errorbar_x3} = 1.3321$
 $\text{errorbar_x4} = 1.6306$



Blue lines represent model error at each x value, red points are the data, and black line is my model.

Problem 3 The data in the table below are assumed to obey a model of the form $y_{i,j} = e_{i,j} \exp(-bt_j)$ with $e_{i,j}$ assumed independent and identically distributed errors with $\log(e_{i,j}) \sim N(\mu, \sigma^2)$. Estimate b , μ , and σ and assign appropriate errorbars to your estimates.

t_i	$y_{1,:}$	$y_{2,:}$	$y_{3,:}$	$y_{4,:}$
1.0	3.75	0.36	0.58	2.06
2.0	0.93	0.32	0.67	1.01
3.0	0.38	0.11	0.12	0.60
4.0	0.05	0.15	0.05	0.11
5.0	0.04	0.03	0.08	0.06

3)

$y_{i,j} = e_{i,j} \exp(-bt_j)$
 $i = 1, 2, 3, 4, 5$ • est. b, μ, σ & errorbars
 $\log(e_{i,j}) \sim N(\mu, \sigma^2)$
 $\underbrace{\hspace{1cm}} = \mu + \varepsilon_{i,j}$ where $\varepsilon_{i,j} \sim N(0, \sigma^2)$
 $\log(y_{i,j}) = \log(e_{i,j}) - bt_j$
 $\underbrace{\hspace{1cm}} x_{i,j} = \mu - bt_j + \varepsilon_{i,j}$

$\langle \log e_{i,j} \rangle = \mu \rightarrow \langle \log e_{i,j} - \mu \rangle = 0$

After taking log,
we see is linear in
terms of
parameter b.
We can start
building likelihood
function where:
Beta = [mu; b]

$$\begin{aligned}
 &\text{prior: } \text{prob}(\mu, b, \sigma) \propto \frac{1}{\sigma} \\
 &\text{likelihood: } \text{prob}(\text{data} | \text{model}) \\
 &\quad = \text{prob}(y | \mu, b, \sigma) \propto \\
 &\quad \rightarrow \frac{1}{\sigma^{10}} \left(\exp\left(-\frac{1}{2\sigma^2} (\underline{\varepsilon}^T \underline{\varepsilon})\right) \right) \\
 &\quad \underline{\varepsilon} = y - A\beta \quad \beta = \begin{bmatrix} \mu \\ b \end{bmatrix} \\
 &\text{prob}(y | \mu, b, \sigma) \propto \frac{1}{\sigma^{10}} \exp\left(-\frac{1}{2} (A\beta - y)^T (A\beta - y)\right) \\
 &\text{prob}(\sigma, \mu, b | y) \propto \frac{1}{\sigma^{11}} \exp\left(-\frac{1}{2} \left(\begin{array}{l} \text{here} \\ \text{also} \\ \text{here} \end{array} \right) \right) \\
 &\log(\text{posterior}) = \text{const} - 11 \log \sigma - \frac{1}{2\sigma^2}
 \end{aligned}$$

The likelihood function should be represented as $\text{Prob}(z_{i,j})$ (i.e. the data, I just call it y) given both parameters of beta (μ and b) and the standard deviation of the errors, σ . A is the design matrix incorporating time points t_j and appropriately aligns with the structure of beta. This is assuming the prior (highlighted in blue). Each time point is repeated four times to account for the four different y vals.

$$\text{prob}(y | \mu, \sigma, \beta) \propto \frac{1}{\sigma^{10}} \exp\left(-\frac{1}{2}(A\beta - y)^T(A\beta - y)\right)$$

$$\text{prob}(\sigma, \mu, \beta | y) \propto \frac{1}{\sigma^{11}} \exp\left(-\frac{1}{2} \left((A\hat{\beta} - y)^T(A\hat{\beta} - y) + \hat{\sigma}^2 \right)\right)$$

$$\log(\text{posterior}) = \text{const} - 11 \log \sigma - \frac{1}{2\sigma^2}$$

$$\hat{\beta} = \underset{\beta}{\text{argmax}} \log(\text{posterior}) = \hat{L}$$

$$\frac{\partial \hat{L}}{\partial \beta} \Big|_{\beta = \hat{\beta}} = 0 \quad \frac{\partial \hat{L}}{\partial \sigma} \Big|_{\sigma = \hat{\sigma}} = 0$$

$$\underline{A}^T \underline{A} \underline{\hat{\beta}} = \underline{A}^T \underline{y}$$

$$\frac{d\hat{L}}{d\sigma} = -\frac{11}{\sigma} - \frac{1}{\sigma^3} \left[(\underline{\beta} - \underline{\hat{\beta}})^T \underline{A}^T \underline{A} (\underline{\beta} - \underline{\hat{\beta}}) + (\underline{y} - \underline{A} \underline{\hat{\beta}})^T (\underline{y} - \underline{A} \underline{\hat{\beta}}) \right]$$

$$\hookrightarrow (\underline{y} - \underline{A} \underline{\hat{\beta}})$$

$$0 = \dots \quad \frac{1}{\hat{\sigma}^2} = - \frac{1}{2} [(\underline{y} - \underline{A} \underline{\hat{\beta}})^T (\underline{y} - \underline{A} \underline{\hat{\beta}})]$$

The posterior becomes (highlighted in blue).

We then define the log posterior denoted as squiggly L (highlighted in green). We want to find β_{hat} , the maximum likelihood estimate for β . We take derivative of squiggly_L wrt β and set it equal to 0, then evaluate at β_{hat} . We expand this and solve through differentiation to get β_{hat} .

I didn't keep up in the notes but derivative wrt β evaluated at $\beta = \beta_{\text{hat}}$ is:

$$d/d_{\beta} [-1/2 * \sigma^2 (A * \beta - y)' (A * \beta - y)] = 0$$

Solve to get:

$$\beta_{\text{hat}} = (A'A)^{-1} * A'y$$

Note: ' implies transpose

Now we want to find the maximum likelihood estimate for sigma. We differentiate squiggle_L wrt/ sigma and set it equal to 0. Set derivative equal to 0 and solve for sigma to get sigma_hat. Multiply through by sigma_hat^3.

$$11 \hat{\sigma}^2 = -[(y - A\hat{\beta})^T (y - A\hat{\beta})]$$

Then divide by residual sum of squares to isolate sigma_hat^2, the maximum likelihood for variance parameter assuming the beta_hat we found.

$$\begin{aligned} \text{prob}(\beta | y) &= \int_0^{\infty} d\sigma \text{prob}(\beta, \sigma | y) \\ \text{prob}(\mu | y) &= \int_{-\infty}^{\infty} d\beta \int_0^{\infty} d\sigma \text{prob}(\beta, \sigma | y) \\ \Rightarrow \log(\text{prob}(\mu | y)) &= \frac{-1}{2(N-K+1)} \log(V) + \dots \\ \rightarrow (\mu - \hat{\mu})^2 / ? & \\ b | y &\rightarrow (b - \hat{b})^2 / c(2, 2) \\ V = N - K & \quad \underline{H} = S^2 \underline{A}^T \underline{A} \quad S^2 = \frac{\underline{e}^T \underline{e}}{V} \\ \underline{e} = \underline{H}^{-1} & \quad c(1, 1) \end{aligned}$$

Given posterior distribution (Beta, sigma | y), we find the distribution of beta independent of sigma. We integrate sigma from the joint posterior. The integration results in a distribution like the student's t-distribution, which is a better representation of beta uncertainty when sigma isn't exactly know.. Integrate over sigma and other components of beta that are not mu to find Prob(mu | y) (highlighted in green). Log of marginal posterior of mu (highlighted in orange). Now we can see the log (Prob(b | y)) = -1/2(N - k + 1) * log(v + (the box highlighted in blue)).

$$\begin{aligned}
 \text{prob}(\beta | y) &= \int_0^\infty d\sigma \text{prob}(\beta, \sigma | y) \\
 \text{prob}(\mu | y) &= \int_{-\infty}^\infty db \int_0^\infty d\sigma \text{prob}(\beta, \sigma | y) \\
 \Rightarrow \log(\text{prob}(\mu | y)) &= \frac{1}{2(N-K+1)} \log(\mathcal{V}) + \dots \\
 \rightarrow (\mu - \hat{\mu})^2 / ? & \\
 b | y &\rightarrow (b - \hat{b})^2 / c(2, 2) \\
 \mathcal{V} = N - K \quad \underline{H} &= S^2 \underline{A}^T \underline{A} \quad S^2 = \frac{\underline{e}^T \underline{e}}{\mathcal{V}} \\
 \underline{e} &= \underline{H}^{-1} \quad c(1, 1)
 \end{aligned}$$

Given posterior distribution (Beta, sigma | y), we find the distribution of beta independent of sigma. We integrate sigma from the joint posterior. The integration results in a distribution like the student's t-distribution, which is a better representation of beta uncertainty when sigma isn't exactly known.. Integrate over sigma and other components of beta that are not mu to find Prob(mu | y).

Solve in Matlab

ANSWERS

b =

-0.8395

mu =

2.8975

sigma_b =

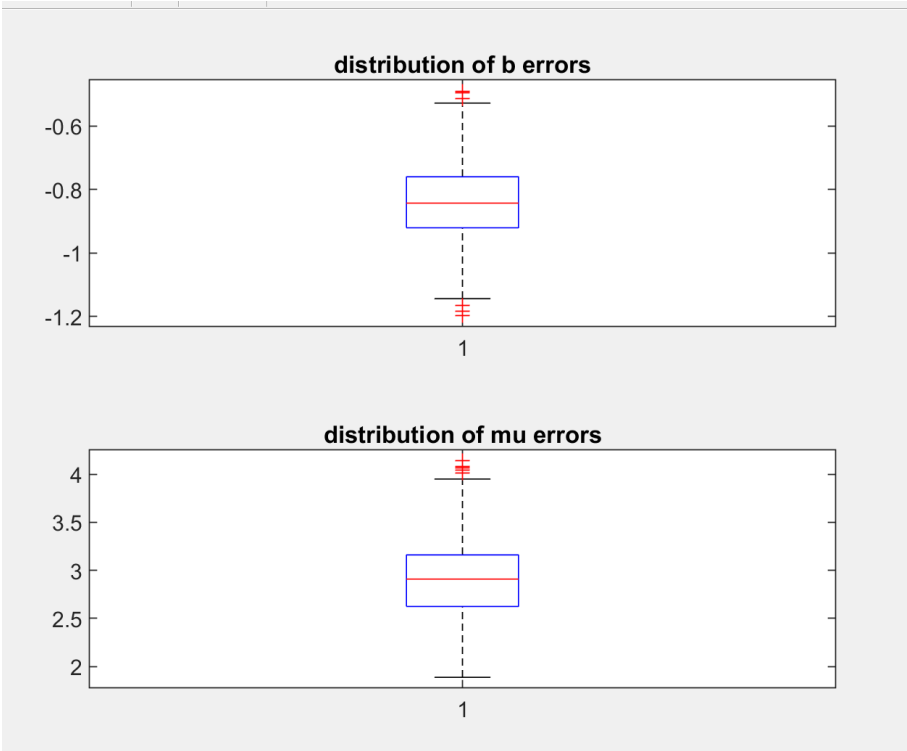
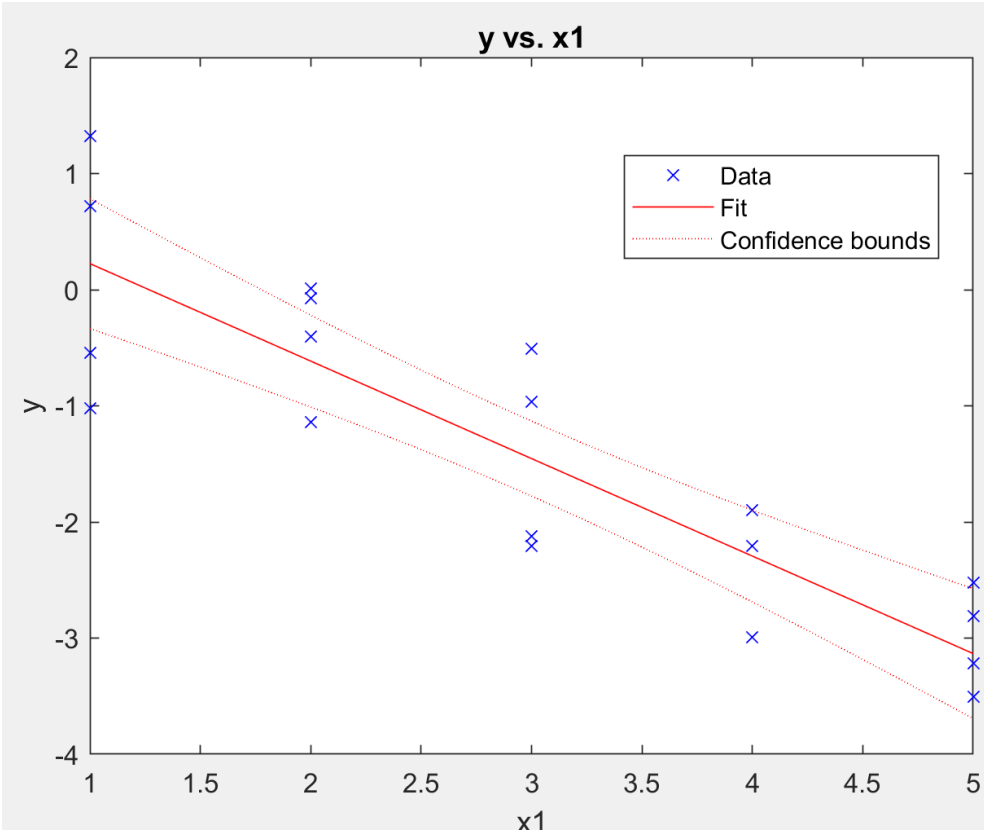
0.1149

sigma_mu =

0.3812

error =

sigma_error = 0.1300



Boxplots of the individual b and mu errors (sigma) for all $x_i : x_N$