

## CS 59800 - RL

### Linear Regression.

Agenda:

- Ridge regression
- Concentration bound

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Consider a linear model of

$$\begin{array}{c} \text{scalar} \swarrow \quad X_t = \langle A_t, \theta_* \rangle + \eta_t \quad \searrow \text{scalar} \\ \quad \quad \quad \downarrow \quad \quad \quad \downarrow \\ \quad \quad \quad A_t \in \mathbb{R}^d \quad \theta_* \in \mathbb{R}^d \end{array}$$

and a sequence of  $A_1, X_1, \dots, A_n, X_n$ , with a filtration  $\mathcal{F} = \{\mathcal{F}_t\}$ , such that,  $\mathcal{F}_t = \sigma(A_1, X_1, \dots, A_{t+1})$

Note that,  $X_t$  is  $\mathcal{F}_t$  measurable.

For noise  $\eta_t$ , we have:

$$\mathbb{E}[\exp(\alpha \eta_t) | \mathcal{F}_{t-1}] \leq \exp\left(\frac{\alpha^2}{2}\right)$$

for all  $\alpha \in \mathbb{R}$  and  $t \in [n]$ .

This is the 1-sub-Gaussian assumption on the noise.

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The process is as follows:

- At each time step  $t$ , someone chooses  $A_t$  from a set  $D_t$ , and we observe  $X_t$ .

Given  $A_1, X_1, \dots, A_t, X_t$ ,

Can we estimate  $\theta_*$ ?

- How about solving a ridge regression for  $\theta_*$ ?

→ at time  $t$ ,

$$\min_{\theta \in \mathbb{R}^d} \sum_{s=1}^t (X_s - \langle A_s, \theta \rangle)^2 + \|\theta\|_V^2$$

For positive Definite matrix  $V$

$$\|\theta\|_V^2 = \theta^T V \theta$$

Let's define

$$- S_t = \sum_{s=1}^t A_s$$

$$- V_t = \sum_{s=1}^t A_s A_s^T$$

$$- V_t(V) = V + V_t = V_t(V)$$

As you may have seen, we usually set  $V = \lambda I$  for  $\lambda > 0$

$$- V_t(\lambda) = \lambda I + V_t$$

Therefore, the minimizer of ridge regression problem is

$$\hat{\theta}_t = V_t(V)^{-1} \sum_{s=1}^t X_s A_s \quad (\text{why?})$$

we are interested is how good is this  $\hat{\theta}_t$  estimate.

To simplify the notation, we use  $V = \lambda I$

Let's define  $M_t(\alpha) = \exp\left(\langle \alpha, s_t \rangle - \frac{1}{2} \|\alpha\|_{V_t}^2\right)$   
for  $\alpha \in \mathbb{R}^d$

Lemma: For any  $\alpha \in \mathbb{R}^d$ , the process  $M_t(\alpha)$  is an  $\mathcal{F}$ -adapted supermartingale.

**Proof:**

— It is clear that  $M_t(\alpha)$  is  $\mathcal{F}_t$ -measurable for all  $t$  by definition. We are left to show

$$E[M_t(\alpha) | \mathcal{F}_{t-1}] \leq M_{t-1}(\alpha) \quad \text{a.s.}$$

Let's expand  $M_t(\alpha) \rightarrow$

$$\begin{aligned} E[M_t(\alpha) | \mathcal{F}_{t-1}] &= E\left[\exp\left(\langle \alpha, s_t \rangle - \frac{1}{2} \|\alpha\|_{V_t}^2\right) \middle| \mathcal{F}_{t-1}\right] \\ &= E\left[\exp\left(\langle \alpha, s_{t-1} \rangle - \frac{1}{2} \|\alpha\|_{V_{t-1}}^2\right) \exp\left(\eta_t \langle \alpha, A_t \rangle - \frac{1}{2} \|\alpha\|_{A_t A_t^T}^2\right) \middle| \mathcal{F}_{t-1}\right] \\ &= M_{t-1}(\alpha) E\left[\underbrace{\exp\left(\eta_t \langle \alpha, A_t \rangle - \frac{1}{2} \|\alpha\|_{A_t A_t^T}^2\right)}_{\leq 1 \text{ a.s.}} \middle| \mathcal{F}_{t-1}\right] \end{aligned}$$

$$\leq M_{t-1}(\alpha) \quad \text{a.s.}$$

$\hookrightarrow$  we concluded that  $M_t(\alpha)$  is a supermartingale sequence.



For a Gaussian measure  $h$  with covariance  $V$  and mean zero, let's define:

$$\overline{M}_t = \int_{\mathbb{R}^d} M_t(\alpha) dh(\alpha)$$

now using Radon-Nikodym derivative, we have:

$$\overline{M}_t = \frac{1}{(2\pi)^{d/2} \det(V)^{1/2}} \int_{\mathbb{R}^d} \exp\left(\langle \alpha, s_t \rangle - \frac{1}{2} \|\alpha\|_{V_t}^2 - \frac{1}{2} \|\alpha\|_V^2\right) d\alpha$$

Note that:

$$\|s_t\|_{\underbrace{(V+V_t)^{-1}}_{V_t(V)^{-1}}}^2 = \|\alpha - (V+V_t)^{-1} s_t\|_{V+V_t}^2$$

$$= 2 \langle \alpha, s_t \rangle - \|\alpha\|_{V_t}^2 - \|\alpha\|_V^2 \quad (\text{why?})$$

$$\Rightarrow \overline{M}_t = \frac{1}{(2\pi)^{d/2} \det(V)^{1/2}} \int_{\mathbb{R}^d} \exp\left(\|s_t\|_{\frac{V_t}{V_t(V)^{-1}}}^2 x \frac{1}{2} - \frac{1}{2} \|\alpha - V_t(V)^{-1} s_t\|_{V_t(V)}^2\right) d\alpha$$

$$= \frac{1}{(2\pi)^{d/2} \det(V)^{1/2}} \exp\left(\frac{1}{2} \|s_t\|_{V_t(V)^{-1}}^2\right) \int_{\mathbb{R}^d} \exp\left(-\frac{1}{2} \|\alpha - V_t(V)^{-1} s_t\|_{V_t(V)}^2\right) d\alpha$$

$$= \frac{\det(V_t(V)^{-1})^{1/2}}{\det(V)^{1/2}} \exp\left(\frac{1}{2} \|s_t\|_{V_t(V)^{-1}}^2\right) \times \left( \frac{1}{(2\pi)^{d/2} \det(V_t(V))^{1/2}} \int_{\mathbb{R}^d} \exp\left(-\frac{1}{2} \|\alpha - V_t(V)^{-1} s_t\|_{V_t(V)}^2\right) d\alpha \right)$$

equal to 1

Note that  $\bar{M}_t$  is also supermartingale (why?)

Using the general form of maximal inequality, we have:

$$\mathbb{P}\left(\sup_t \bar{M}_t > \frac{1}{\delta}\right) \leq \delta$$

$$\text{we know that } \{t \in [n]; \frac{\det(V_t(V))^{1/2}}{\det(V^{-1})^{1/2}} \exp\left(\frac{1}{2} \|S_t\|_{V_t(V)^{-1}}^2\right) > \frac{1}{\delta}\}$$

$$\subset \left\{ \sup_t \bar{M}_t > \frac{1}{\delta} \right\}$$

$$\hookrightarrow \mathbb{P}\left(t \in [n]; \frac{\det(V_t(V))^{1/2}}{\det(V^{-1})^{1/2}} \exp\left(\frac{1}{2} \|S_t\|_{V_t(V)^{-1}}^2\right) > \delta\right) \leq \delta$$

$$\hookrightarrow \mathbb{P}\left(t \in [n]; \|S_t\|_{V_t(V)^{-1}} > 2 \log \frac{1}{\delta} + \log \left( \frac{\det(V_t(V))}{\det(V)} \right)\right) \leq \delta$$

Theorem: For  $\delta \in (0, 1)$ , with probability at least  $1 - \delta$ , for  $t \in [n]$ , we have; for  $V = \lambda I$

$$\|\hat{\theta}_t - \theta_*\|_{V_t(V)} \leq \sqrt{P_t(\delta)} : \sqrt{\lambda} \|\theta_*\| + \underbrace{\sqrt{2 \log\left(\frac{1}{\delta}\right) + \log\left(\frac{\det(V_t(V))}{\det(V)}\right)}}_{\text{Fort-1}}$$

Furthermore if  $\|\theta_*\| \leq S$ , define confidence interval/set

$$C_t(\delta) = \left\{ \theta \in \mathbb{R}^d : \|\hat{\theta}_{t-1} - \theta\|_{V_{t-1}(V)} \leq \sqrt{\lambda} S + \underbrace{\quad}_{\text{Fort-1}} \right\}$$

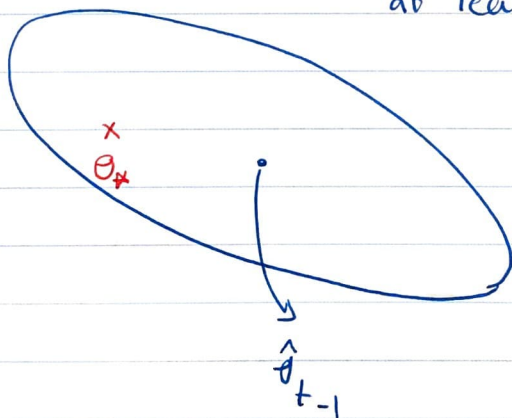
Then  $\mathbb{P}(\exists t \in [n]; \theta_* \notin C_t(\delta)) \leq \delta$

where  $C_t(\delta) = \left\{ \theta \in \mathbb{R}^d : \|\theta_{t-1} - \theta\|_{V_t(V)} \leq \sqrt{\lambda} \delta + \sqrt{2 \log\left(\frac{1}{\delta}\right) + \log\left(\frac{\det(V_{t-1}(V))}{\det(V)}\right)} \right\}$

what does it mean?

$$\left\{ \theta \in \mathbb{R}^d : \|\hat{\theta}_{t-1} - \theta\|_{V_{t-1}(V)}^2 \leq \beta_t(\delta) \right\}$$

is an ellipse such that  $\theta$  is in it, always, with probability at least  $1 - \delta$



Can we simplify  $\sqrt{\beta_t(\delta)}$ ? So far, we did not use the fact that we want to set  $V = \lambda I$ . The results holds for any positive definite  $V$ . For  $V = \lambda I \rightarrow \det(V) = \lambda^d$