# The EM algorithm Analysis of a T cell frequency assay

#### Karl Broman

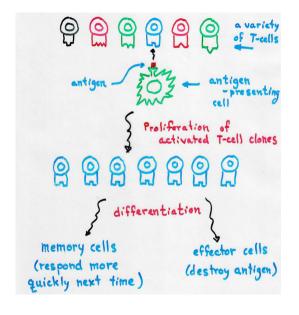
Biostatistics & Medical Informatics, UW-Madison

kbroman.org
github.com/kbroman
@kwbroman
Course web: kbroman.org/AdvData

Goal: Estimate the frequency of T-cells in a blood sample that respond to two test antigens.

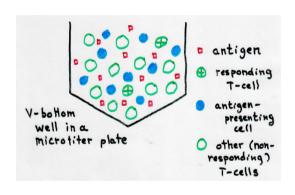
Real goal: Determine whether a vaccine causes an increase in the frequency of responding T-cells.

Broman K, Speed T, Tigges M (1996) J Immunol Meth 198:119-132 doi.org/b54v33



# The assay

- Combine:
  - diluted blood cells + growth medium
  - antigen
  - <sup>3</sup>H-thymidine
- ► Replicating cells take up <sup>3</sup>H-thymidine.
- Extract the DNA and measure its radioactivity

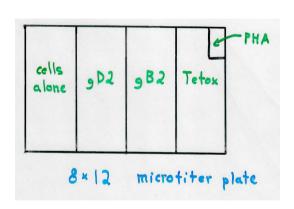


### Usual approaches

- ► Use 3 wells with antigen and 3 wells without antigen, and take the ratio of the averages
- Limiting dilution assay
  - Several dilutions of cells
  - Many wells at each dilution

### Our assay

Study a single plate or pair of plates at a single dilution.



#### Data

#### LDA 713, plates 1 and 2 11,400 cells per well

cells alone         gD2         gB2           179         249         460         2133         2528         2700         2171         1663         620           346         1540         306         8299         1886         3245         1699         2042         337           117         249         1568         1174         4293         979         1222         1536         240           184         414         308         2801         2438         1776         2193         3211         193	74 183 7748 1033 06 6497 2492 618
346 1540 306 8299 1886 3245 1699 2042 337 117 249 1568 1174 4293 979 1222 1536 240	74 183 7748 1033 66 6497 2492 618
117 249 1568 1174 4293 979 1222 1536 240	74 183 7748 1033 96 6497 2492 618
104 414 200 2001 240	6 6497 2492 618
104 414 200 2001 2100	
184 414 308 2801 2438 1776 2193 3211 193	
797 233 461 1076 1527 2866 2205 2278 221	
305 348 480 3475 902 3654 2046 1285 118	1 0.00 100
1090 159 89 1472 90 3639 657 2393 181	
280 571 329 4448 3643 881 3462 2118 101	
101	0 0100 4010 07.
178 111 630 4699 5546 5182 3982 3104 249	C   100   000   000
044 500 050 5000 0102 0502 0104 249	
2010 1010 1415 2516 430	2   5017   5074   10706
261 964 167 2991 3390 3986 2321 2157 327	8 8216 3579 3538
221 544 299 1838 4368 322 1022 1554 298	
533 228 615 1938 4046 333 3253 5091 284	
818 98 160 1032 3269 4918 1778 3810 237	
234 472 243 4143 3351 1118 530 1174 188	
169 481 478 3237 1565 2211 2460 2715 4793	

### Traditional analysis

➤ Split wells into +/- using a cutoff (e.g., mean + 3 SD of "cells alone" wells)

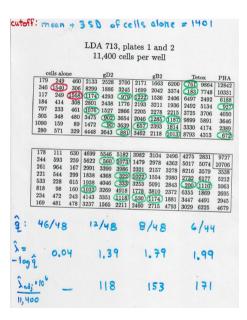
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positive = one or more responding cells
negative = no responding cells
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▶ Imagine that the number of responding cells in a well is Poisson( $\lambda_i$ ) for group i

Pr(no responding cells) =  $e^{-\lambda_i}$ 

$$\hat{\lambda}_i = -\log\left(\frac{\text{\# negative wells}}{\text{\# wells}}\right)$$

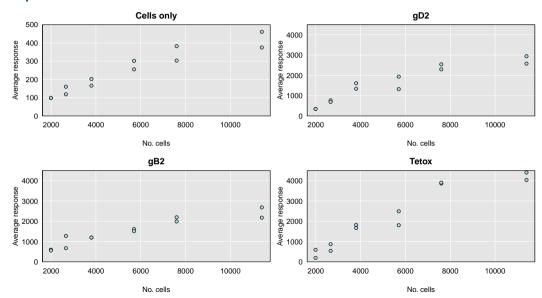
### **Analysis**



#### **Problems**

- ► Hard to choose cutoff
- ► Potential loss of information

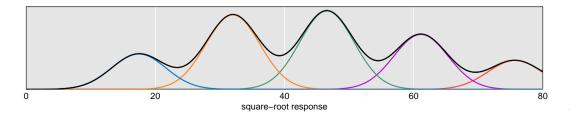
# Response vs no. cells



#### Model

 $k_{ij}$  = Number of responding cells (unobserved)  $y_{ij}$  = square-root of response

Assume  $k_{ij} \sim \mathsf{Poisson}(\lambda_i)$   $y_{ij} \mid k_{ij} \sim \mathsf{Normal}(a + bk_{ij}, \sigma)$   $(k_{ij}, y_{ij}) \text{ mutually independent}$ 



# log Likelihood

$$I(\lambda, a, b, \sigma) = \sum_{i,j} \log \Pr(y_{ij} | \lambda_i, a, b, \sigma)$$

$$= \sum_{i,j} \log \left[ \sum_k \Pr(k | \lambda_i) \Pr(y_{ij} | k, a, b, \sigma) \right]$$

$$= \sum_{i,j} \log \left[ \sum_k \left( \frac{e^{-\lambda_i} \lambda_i^k}{k!} \right) \phi \left( \frac{y_{ij} - a - bk}{\sigma} \right) \right]$$

# EM algorithm

- ► Iterative algorithm useful when there is missing data that if observed would make things easy
- ▶ Dempster et al. (1977) JRSS-B 39:1-22 doi.org/gfxzrv
- Start with some initial estimates
- ► E-step: expected value of missing data given current estimates
- M-step: MLEs replacing missing data with their expected values
- Advantages
  - often easy to code
  - usually super stable
  - log likelihood is non-decreasing

#### Normal/Poisson model

E-step:

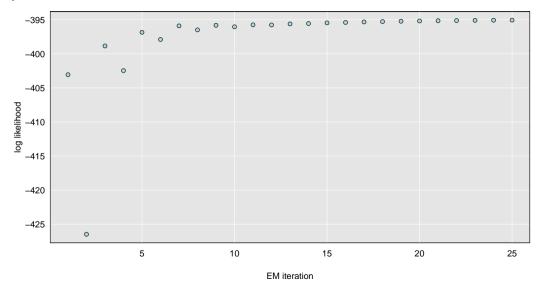
$$\Pr(k = s | y, \lambda, a, b, \sigma) = \frac{\Pr(k = s | \lambda) \Pr(y | k = s, a, b, \sigma)}{\sum_{s} \Pr(k = s | \lambda) \Pr(y | k = s, a, b, \sigma)}$$

$$= \frac{\left(\frac{e^{-\lambda} \lambda^{s}}{s!}\right) \phi\left(\frac{y - a - bs}{\sigma}\right)}{\sum_{s} \left(\frac{e^{-\lambda} \lambda^{s}}{s!}\right) \phi\left(\frac{y - a - bs}{\sigma}\right)}$$

$$\mathsf{E}(\mathbf{k}|\mathbf{y},\lambda,\mathbf{a},\mathbf{b},\sigma) = \frac{\sum_{s} s\left(\frac{e^{-\lambda}\lambda^{s}}{s!}\right) \phi\left(\frac{y-a-bs}{\sigma}\right)}{\sum_{s} \left(\frac{e^{-\lambda}\lambda^{s}}{s!}\right) \phi\left(\frac{y-a-bs}{\sigma}\right)}$$

M-step: Regress y on E(k|y)

# Oops, that didn't work



# EM algorithm, more formally

► Calculate expected complete-data log likelihood, given observed data and observed parameters, and then maximize that.

$$I^{(s)}(\theta) = \mathsf{E}\{\log f(y, k|\theta)|y, \hat{\theta}^{(s)}\}$$

- ► In practice, it's usually a linear combination of the sufficient statistics, so you focus on those.
- ► Here, we need not just  $\sum k$  and  $\sum ky$ , but also  $\sum k^2$ .

# EM algorithm, again

E step: we also need

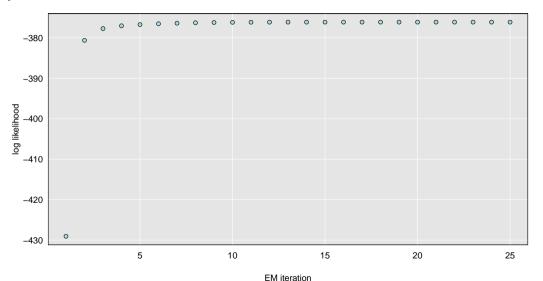
$$\mathsf{E}(k^2|\mathbf{y},\lambda,\mathbf{a},\mathbf{b},\sigma) = \frac{\sum_{\mathbf{s}} \mathbf{s}^2 \left(\frac{\mathbf{e}^{-\lambda} \lambda^{\mathbf{s}}}{\mathbf{s}!}\right) \phi \left(\frac{\mathbf{y}-\mathbf{a}-\mathbf{b}\mathbf{s}}{\sigma}\right)}{\sum_{\mathbf{s}} \left(\frac{\mathbf{e}^{-\lambda} \lambda^{\mathbf{s}}}{\mathbf{s}!}\right) \phi \left(\frac{\mathbf{y}-\mathbf{a}-\mathbf{b}\mathbf{s}}{\sigma}\right)}$$

M step: we want  $\hat{\beta} = (X'X)^{-1}(X'y)$ 

where 
$$(X'X)$$
 is like  $\begin{pmatrix} n & \sum k \\ \sum k & \sum k^2 \end{pmatrix}$ 

and 
$$(X'y)$$
 is like  $\begin{pmatrix} \sum y \\ \sum ky \end{pmatrix}$ 

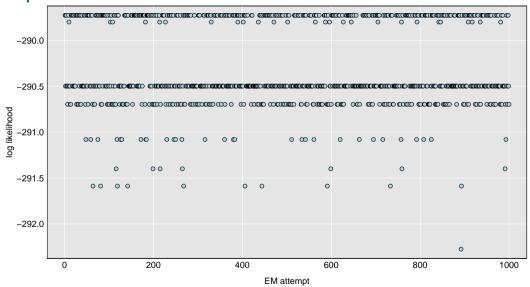
# Ah, that's better



#### **Difficulties**

- ► Starting values
- ► Multiple modes

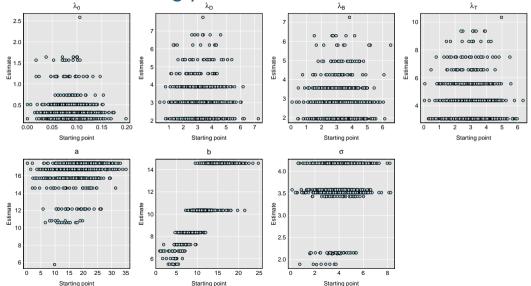
# Multiple modes



# Multiple modes

	$\lambda$ _0	$\lambda\_D$	$\lambda$ _B	$\lambda\_T$	а	b	$\sigma$	log lik	no. hits
1	0.32	3.03	2.82	4.37	16.73	10.34	3.52	-289.73	331
2	1.18	5.40	4.95	7.49	12.16	6.69	2.15	-289.80	26
3	0.17	2.10	1.95	3.07	17.44	14.56	4.18	-290.50	415
4	0.51	3.89	3.56	5.58	15.72	8.35	3.58	-290.70	180
5	0.73	4.62	4.25	6.58	14.58	7.27	3.43	-291.08	30
6	1.64	6.79	6.29	9.35	10.81	5.51	1.89	-291.40	7
7	1.57	6.22	5.80	8.61	10.60	6.02	2.13	-291.59	10
8	2.59	7.76	7.25	10.34	5.75	5.47	1.88	-292.27	1

# Estimate vs. starting point



### **Principles**

- Start with an understanding of the problem and data
- ► Think about a model for the data-generating process

#### Lessons

- ► The EM algorithm is really useful
- ▶ Use the log likelihood as a diagnostic when implementing an EM algorithm

# Software development time

- ► Formulating the problem
- ▶ Writing the code
- ► Debugging the code
- ► Executing the code

### **Impact**

- ▶ I'm pretty sure that the vaccine they were working on didn't work well.
- ▶ R package npem, but I never put it on CRAN, and no one has ever asked me about it.
- ► Our paper has like 9 citations: no one has ever really used the method.

# Further things

- Standard errors should always be required.
  - But usually painful to obtain
  - We used the SEM algorithm of Meng and Rubin (1991) doi.org/dk27
- Could more formally investigate the appropriate transformation
  - See Box and Cox (1964) doi.org/gfrhvs
  - Box-Cox transformation is  $g(y) = (y^c 1)/c$  for  $c \neq 0$  and  $= \log y$  for c = 0
  - Key issue is change-of-variables in the density; as a result you add  $\sum_{ij} (c-1) \log y_{ij}$  to the log likelihood