STT 863 Compendium

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Baseline Knowledge

Def Boole's inequality

$$P(\cup_{i=1}^{n} A_i) \le \sum_{i=1}^{n} P(A_i)$$

Def Bonferroni's inequality

$$P(\cap_{i=1}^{n} A_i) \ge 1 - \sum_{i=1}^{n} P(A_i^c)$$

<u>Def</u> A function F is called the cumulative distribution function iff:

- a. $\lim_{x\to-\infty} F(x) = 0$ and $\lim_{x\to\infty} = 1$
- b. F(x) is a nondecreasing fn of x
- c. F(x) is right-continuous; that is for every number x_0 , $\lim_{x\downarrow x_0} F(x) = F(x_0)$

<u>Def</u> A PDF of a continuous r.v. x is $f_X(x) = \frac{d}{dx} F_X(x)$:

- a. $f_X(x) \ge 0 \ \forall x$
- b. $\int_{-\infty}^{\infty} f_X(x) dx = 1$
- c. $F_X(x) = \int_{-\infty}^x f_X(s) ds$

For below, change integrals to sums if the r.v is discrete

Def Expectation

$$\mu := \mathbb{E}(X) = \int_{-infty}^{\infty} x f(x) dx$$

For any function h:

$$\mathbb{E}(h(X)) = \int_{-infty}^{\infty} h(x)f(x)dx$$

Def Variance

$$\sigma^2(X) := \mathbb{E}((X - \mu)^2) = \mathbb{E}(X^2) - \{\mathbb{E}(X)\}^2$$

<u>Def</u> Normal Distribution $X \sim N(\mu, \sigma^2)$

PDF:

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$$

<u>Def</u> Covariance is a measure of a linear relationship between X and Y. $Cov(X, Y) = \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y)$

<u>Def</u> Correlation coefficient between X and Y is defined as: $\rho = Corr(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$

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Statistical Inference

 $\underline{\mathbf{Def}}$ Sample Mean

$$\bar{Y} = \frac{1}{n} \sum_{i=1}^{n} Y_i$$

 $\underline{\mathbf{Def}}$ Sample Variance

$$s^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}$$

Remark Properties of \bar{Y} and s^2

a.
$$\mathbb{E}(\bar{Y}) = \mu$$

b.
$$Var(\bar{Y}) = \frac{\sigma^2}{n}$$

c.
$$\mathbb{E}(s^2) = \sigma^2$$

 $\underline{\mathbf{Def}}$ Homoscedasticity: $\sigma^2_{Y|X}$ is fixed across X values.

<u>Def</u> Simple linear regression (SLR) model:

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

a. ϵ and X are independent

b.
$$\mathbb{E}\epsilon_i = 0, i = 1, 2, ..., n$$

c.
$$Var(\epsilon_i) = \sigma^2, i = 1, 2, ..., n$$

d.
$$Cov(\epsilon_i, \epsilon_j) = 0, i \neq j$$

<u>Def</u> Normal equations for the least square method:

$$\sum_{i=1}^{n} Y_i = nb_o + b_1 \sum_{i=1}^{n} X_i$$

$$\sum_{i=1}^{n} X_i Y_i = b_o \sum_{i=1}^{n} X_i + b_1 \sum_{i=1}^{n} X_i^2$$

Def Least Square Estimators

$$b_1 = \frac{\sum_{i=1}^{n} (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_{i=1}^{n} (X_i - \bar{X})^2}$$

$$b_0 = \bar{Y} - b_1 \bar{X}$$

$$s^2 = MSE = \frac{SSE}{n-2} = \frac{1}{n-2} \sum_{i=1}^{n} e_i^2$$

Theorem Gauss-Markov Under the assumptions of the regression model, the least square estimators b_0 and b_1 are

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- linear
- unbiased
- have minimum variance among all unbiased linear estimators of β_0 and β_1

 $\underline{\mathbf{Def}}$ Some properties of sample estimators:

a.
$$\sum_{i=1}^{n} e_i = 0$$

b.
$$\sum_{i=1}^{n} \hat{Y}_i = \sum_{i=1}^{n} Y_i$$

c.
$$\sum_{i=1}^{n} X_i e_i = 0$$

d.
$$\sum_{i=1}^{n} \hat{Y}_i e_i = 0$$

 $\underline{\mathbf{Def}}$ Sampling distribution of b_1

a.
$$b_1 = \sum_{i=1}^n k_i Y_i$$
 where

$$k_{i} = \frac{X_{i} - \bar{X}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}$$

b.
$$\mathbb{E}(b_1) = \beta_1$$

c. $Var(b_1)$:

$$\sigma^{2}\{b_{1}\} = \frac{\sigma^{2}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}$$

d. Standard error:

$$s^{2}\{b_{1}\} = \frac{\sigma^{2}}{\sum_{i=1}^{n}(X_{i} - \bar{X})^{2}} = \frac{\text{MSE}}{\sum_{i=1}^{n}(X_{i} - \bar{X})^{2}}$$

<u>Def</u> Sampling distribution of b_0

a.
$$b_0 = \sum_{i=1}^n l_i Y_i$$
 where

$$l_i = \frac{1}{n} - \frac{(X_i - \bar{X})\bar{X}}{\sum_{i=1}^{n} (X_i - \bar{X})^2}$$

b.
$$\mathbb{E}(b_0) = \beta_0$$

c. $Var(b_0)$:

$$\sigma^{2}\{b_{0}\} = \sigma^{2}\left[\frac{1}{n} + \frac{\bar{X}^{2}}{\sum_{i=1}^{n}(X_{i} - \bar{X})^{2}}\right]$$

d. Standard error:

$$s^{2}{b_{0}} = \text{MSE}\left[\frac{1}{n} + \frac{\bar{X}^{2}}{\sum_{i=1}^{n}(X_{i} - \bar{X})^{2}}\right]$$

Prediction in SLR

Remark Estimating $\mathbb{E}(Y_h)$

We estimate $\mathbb{E}(Y_h)$ by $\hat{Y}_h = b_0 + b_1 X_h$ $\mathbb{E}(\hat{Y}_h)\beta_0 + \beta_1 X_h = \mathbb{E}(Y_h)$

 $Var(Y_h)$:

$$\sigma^{2}\{Y_{h}\} = \sigma^{2}\left[\frac{1}{n} + \frac{(X_{h} - \bar{X})^{2}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}\right]$$

Standard error:

$$s^{2}{Y_{h}} = MSE\left[\frac{1}{n} + \frac{(X_{h} - \bar{X})^{2}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}\right]$$

Remark C.I. of $\mathbb{E}(Y_h)$

The $100(1-\alpha)\%$ C.I. of $\mathbb{E}(Y_h)$

$$\hat{Y}_h \pm t_{1-\alpha/2:n-2} s \{\hat{Y}_h\}$$

Where

$$s^{2}\{\hat{Y}_{h}\} = \text{MSE}\left[\frac{1}{n} + \frac{(X_{h} - \bar{X})^{2}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}\right]$$

Remark P.I. of $\mathbb{E}(Y_{h(new)})$

The $100(1-\alpha)\%$ C.I. of $\mathbb{E}(Y_h)$

$$\hat{Y}_h \pm t_{1-\alpha/2:n-2} s\{pred\}$$

Where

$$s^{2}\{pred\} = s^{2} + s^{2}\{\hat{Y}_{h}\} = MSE\left[1 + \frac{1}{n} + \frac{(X_{h} - \bar{X})^{2}}{\sum_{i=1}^{n} (X_{i} - \bar{X})^{2}}\right]$$

Brown-Forsythe Test

Test for Heteroskedasticity

- a. Let n_1 and n_2 be sample sizes for two groups with $n = n_1 + n_2$
- b. Divide the residual sets into two groups with medians $\tilde{e_1}$ and $\tilde{e_2}$
- c. Define,
 - (a) $d_{i1} = |e_{i1} \tilde{e_1}|, i = 1, ... n_1$
 - (b) $d_{j2} = |e_{j2} \tilde{e_2}|, j = 1, ... n_1$
- d. Let \bar{d}_1 and \bar{d}_2 be means of d's from the previous groups, and let

$$s_d^2 = \frac{\sum_{i=1}^{n_1} (d_{i1} - \bar{d}_1)^2 + \sum_{j=1}^{n_2} (d_{j2} - \bar{d}_2)^2}{n-2}$$

e. Test Statistic:

$$t_{BF}^* = \frac{\bar{d}_1 - \bar{d}_2}{s_d \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

f. Reject the null hypothesis (constant variance) if $|t_{BF}^*| > t_{1-\alpha/2;n-2}$

Breusch-Pagan Test

Assumptions:

- a. ϵ_i 's are independent and normal
- b. If $\sigma_i^2 = Var(\epsilon_i)$, then it satisfies $log\sigma_i^2 = \gamma_0 + \gamma_1 X_i$

Test Procedure:

- a. Hypotheses:
 - (a) $H_0: \gamma_1 = 0$
 - (b) $H_a: \gamma_1 \neq 0$
- b. Regress e_i^2 on X_i 's. Let SSR* be the regression SS
- c. Test statistic: $\chi_{BP}^2 = \frac{\text{SSR}}{2} \div (\frac{\text{SSE}}{n})^2$, where SSE is the error SS of the original regression.
- d. Reject H_0 if $\chi^2_{BP} > \chi^2_{1-\alpha;1}$