Infinite-dimensional α -divergence minimisation for Variational Inference

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OxCSML seminar 19/11/2021

Joint work with Randal Douc and François Portier

Outline

- 1 Introduction
- 2 Infinite-dimensional α -divergence minimisation
- 3 Numerical experiments
- **4** Conclusion

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Bayesian statistics

 \bullet Compute / sample from the posterior density of the latent variables y given the data ${\mathscr D}$

$$p(y|\mathscr{D}) = \frac{p(\mathscr{D}, y)}{p(\mathscr{D})}$$
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- → Variational Inference : inference is seen as an optimisation problem.
 - **1** Posit a *simpler* variational family Q, where $q \in Q$.
 - **2** Fit q to obtain the best approximation to the posterior density

$$\inf_{q\in\mathcal{Q}}D(\mathbb{Q}||\mathbb{P}_{|\mathscr{D}})\;,$$

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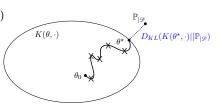
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Question: How to choose D and Q?

$$\begin{cases}
D_{KL}(\mathbb{Q}||\mathbb{P}) = \int_{\mathbb{Y}} \log \left(\frac{q(y)}{p(y)} \right) q(y) \nu(\mathrm{d}y) \\
\mathcal{Q} = \{q : y \mapsto k_1(\theta_1, y_1) k_2(\theta_2, y_2) : (\theta_1, \theta_2) \in \mathbb{T} \}
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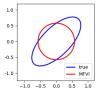
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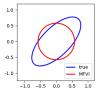


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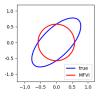


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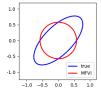


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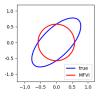


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- Can we select alternative/more general D?
 - $\rightarrow D$ is the α -divergence
- Can we design more expressive variational families Q beyond traditional parametric families?
 - \rightarrow Put a prior on the variational parameter θ

 $\begin{array}{l} (\mathsf{Y},\mathcal{Y},\nu): \text{ measured space, } \nu \text{ is a } \sigma\text{-finite measure on } (\mathsf{Y},\mathcal{Y}). \\ \mathbb{Q} \text{ and } \mathbb{P}: \mathbb{Q} \preceq \nu \text{, } \mathbb{P} \preceq \nu \text{ with } \frac{\mathrm{d}\mathbb{Q}}{\mathrm{d}\nu} = q \text{, } \frac{\mathrm{d}\mathbb{P}}{\mathrm{d}\nu} = p. \end{array}$

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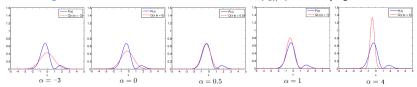
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A flexible family of divergences...

Figure: In red, the Gaussian which minimises $D_{\alpha}(\mathbb{Q}||\mathbb{P})$ for a varying α



Adapted from V. Cevher's lecture notes (2008) https://www.ece.rice.edu/~vc3/elec633/AlphaDivergence.pdf

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Idea: Extend the traditional variational parametric family

$$Q = \{ y \mapsto k(\theta, y) : \theta \in \mathsf{T} \}$$

by putting a prior on the variational parameter θ

$$\mathcal{Q} = \left\{ q : y \mapsto \mu k(y) := \int_{\mathbb{T}} \mu(\mathrm{d}\theta) k(\theta, y) \; : \; \mu \in \mathsf{M} \right\}$$

and propose an update formula for μ that ensures a systematic decrease in the $\alpha\text{-divergence}$ at each step

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 Finite Mixture Models : $\mu = \sum_{j=1}^J \lambda_j \delta_{ heta_j}$

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- M is a subset of $M_1(T)$, the space of probability measures on T
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Algorithm

Let $\mu_1 \in M_1(T)$ be such that $\Psi_{\alpha}(\mu_1 k) < \infty$. The sequence of probability measures $(\mu_n)_{n\geqslant 1}$ is defined iteratively by

$$\mu_{n+1} = \mathcal{I}_{\alpha}(\mu_n) , \qquad n \geqslant 1$$

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Conditions for a monotonic decrease

(A1) For all $(\theta, y) \in \mathsf{T} \times \mathsf{Y}$, $k(\theta, y) > 0$, $p(y) \geqslant 0$ and $\int_{\mathsf{Y}} p(y) \nu(\mathrm{d}y) < \infty$.

(A2) The function $\Gamma:\mathrm{Dom}_{\alpha}\to\mathbb{R}_{>0}$ is decreasing, continuously differentiable and satisfies the inequality

$$[(\alpha - 1)(v - \kappa) + 1] (\log \Gamma)'(v) + 1 \ge 0.$$

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Assume (A1) and (A2). Let $\mu \in M_1(T)$ be such that $\Psi_{\alpha}(\mu k) < \infty$ and $\mu(\Gamma(b_{\mu,\alpha} + \kappa)) < \infty$. Then,

- $\bullet \ \Psi_{\alpha}(\mathcal{I}_{\alpha}(\mu)k) \leqslant \Psi_{\alpha}(\mu k)$
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$$\Gamma(v) = e^{-\eta v}$$

$$\mu_{n+1}(\mathrm{d}\theta) \propto \mu_n(\mathrm{d}\theta) \exp\left[-\eta \int_{\mathsf{Y}} k(\theta,y) \log\left(\frac{\mu_n k(y)}{p(y)}\right) \nu(\mathrm{d}y)\right]$$

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- Power descent : $\eta \in (0,1]$, $(\alpha-1)\kappa \geqslant 0$ and $\alpha \neq 1$

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Convergence results

Entropic Mirror Descent

$$\eta \in (0, \frac{1}{|\alpha - 1||b|_{\infty,\alpha} + 1}), \ \kappa \in \mathbb{R}$$

Power Descent

$$\eta \in (0,1]$$
 , $(\alpha-1)\kappa \geqslant 0$

Algorithm	Convergence results
Entropic Mirror Descent $\eta \in (0, \frac{1}{ \alpha-1 b _{\infty,\alpha}+1}), \ \kappa \in \mathbb{R}$	O(1/N) convergence rates
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$$S_J = \left\{ \pmb{\lambda} = (\lambda_1,...,\lambda_J) \in \mathbb{R}^J \ : \ \forall j \in \{1,...,J\} \,, \ \lambda_j \geqslant 0 \ \text{and} \ \sum_{j=1}^J \lambda_j = 1 \right\}$$
 Let $\theta_1,...,\theta_J \in \mathsf{T}$ be fixed and denote

$$\mu_{\lambda} = \sum_{j=1}^{J} \lambda_j \delta_{\theta_j}$$
 where $\lambda \in \mathcal{S}_J$.

Then, $\mu_n = \underbrace{\mathcal{I}_{\alpha} \circ \cdots \circ \mathcal{I}_{\alpha}(\mu_{\lambda})}_{n \text{ times}}$ is of the form $\mu_n = \sum_{j=1}^J \lambda_{j,n} \delta_{\theta_j}$ with

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Outline

- 1 Introduction
- 2 Infinite-dimensional α -divergence minimisation
- 3 Numerical experiments
- 4 Conclusion

• Gaussian kernel with density k_h and bandwidth h, $\mathsf{T} = \mathbb{R}^d$

$$\left\{ y \mapsto \mu_{\lambda,\Theta} k_h(y) = \sum_{j=1}^J \lambda_j k_h(y - \theta_j) : \lambda \in \mathcal{S}_J, \Theta \in \mathsf{T}^J \right\} .$$

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- **2** Exploration step : update Θ (e.g. by sampling under $\mu_{\lambda,\Theta}k_h$, $h \propto J^{-1/(4+d)}$)
- Toy example $p(y) = Z \times [0.5\mathcal{N}(y; -2u_d, I_d) + 0.5\mathcal{N}(y; 2u_d, I_d)], \ Z = 2$
- Bayesian Logistic Regression Covertype dataset (581,012 data points and 54 features)

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Comparison between

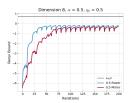
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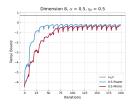
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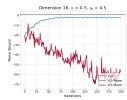


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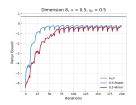


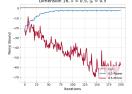


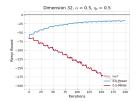
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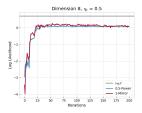
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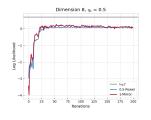
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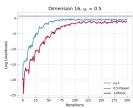


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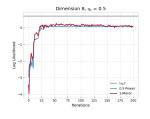


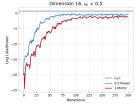


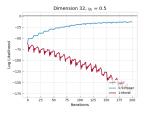
Comparison between:

- 1-Mirror descent : $\Gamma(v) = e^{-\eta v}$ with $\alpha = 1$,
- 0.5-Power descent : $\Gamma(v) = [(\alpha 1)v + 1]^{\eta/(1-\alpha)}$ with $\alpha = 0.5$.

J=M=100 , initial mixture weights : [1/J,...,1/J],~N=10,~T=20 $\eta_n=\eta_0/\sqrt{n},~\eta_0=0.5,$ cv criterion : Ilh averaged over 100 trials







 $o \mathscr{D} = \{c,x\}$: I binary class labels, $c_i \in \{-1,1\}$, L covariates for each datapoint, $x_i \in \mathbb{R}^L$

 \rightarrow Model: L regression coefficients $w_l \in \mathbb{R}$, precision parameter $\beta \in \mathbb{R}^+$

$$\begin{aligned} p_0(\beta) &= \operatorname{Gamma}(\beta; a, b) \;, \\ p_0(w_l | \beta) &= \mathcal{N}(w_l; 0, \beta^{-1}) \;, \quad 1 \leqslant l \leqslant L \\ p(c_i &= 1 | \boldsymbol{x}_i, \boldsymbol{w}) &= \frac{1}{1 + e^{-\boldsymbol{w}^T \boldsymbol{x}_i}} \;, \quad 1 \leqslant i \leqslant I \end{aligned}$$

where a = 1 and b = 0.01

Nonparametric variational inference S. Gershman, M. Hoffman, and D. Blei (2012). ICML

$$ightarrow$$
 Quantity of interest : $p(y|\mathcal{D})$ with $y = [{m w}, \log eta]$

Comparison between

- 0.5-Power descent
- Typical AIS

$$N = 1$$
, $T = 500$, $J_0 = M_0 = 20$, $J_{t+1} = M_{t+1} = J_t + 1$
initial mixture weights: $[1/J_t, ..., 1/J_t]$, $n_0 = n_0/\sqrt{n}$ with $n_0 = 0.05$

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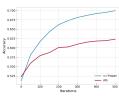
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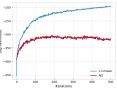
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Outline

- 1 Introduction
- 2 Infinite-dimensional α -divergence minimisation
- 3 Numerical experiments
- **4** Conclusion

$$\mathcal{Q} = \left\{q: y \mapsto \int_{\mathsf{T}} \mu(\mathrm{d}\theta) k(\theta,y) \ : \ \mu \in \mathsf{M} \right\}$$

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- novel Power Descent algorithm
- conditions for a systematic decrease + convergence results
- applicable to mixture models :

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Thank you for your attention!

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