# Problem Set 2 - Machine Learning Theories

April 16, 2025

## Problem 1

Consider two multi-dimensional random vectors  $\mathbf{x} \in \mathbb{R}^M$  and  $\mathbf{y} \in \mathbb{R}^D$  following marginal and conditional distributions which are Gaussian:

$$p(\mathbf{x}) = N(\mathbf{x}; \mu, \Sigma)$$
$$p(\mathbf{y}|\mathbf{x}) = N(\mathbf{y}; A\mathbf{x} + \mathbf{b}, S),$$

where  $\mu \in \mathbb{R}^M$ ,  $\Sigma \in \mathbb{R}^{M \times M}$ ,  $A \in \mathbb{R}^{D \times M}$ ,  $\mathbf{b} \in \mathbb{R}^D$ , and  $S \in \mathbb{R}^{D \times D}$ .

- (1) Determine the joint distribution  $p(\mathbf{x}, \mathbf{y})$ .
- (2) Determine the marginal distribution  $p(\mathbf{y})$ .
- (3) Consider two random vectors  $\mathbf{x}, \mathbf{z} \in \mathbb{R}^D$  following Gaussian distributions  $p(\mathbf{x}) = N(\mathbf{x}; \mu_x, \Sigma_x)$  and  $p(\mathbf{z}) = N(\mathbf{z}; \mu_z, \Sigma_z)$ , respectively. Use the above results to find the marginal distribution of  $\mathbf{y} = \mathbf{x} + \mathbf{z}$ , by considering the marginal distribution  $p(\mathbf{x})$  and the conditional distribution  $p(\mathbf{y}|\mathbf{x})$ .

Use the following equations if necessary:

(a)

$$\begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} M & -MBD^{-1} \\ -D^{-1}CM & D^{-1} + D^{-1}CMBD^{-1} \end{pmatrix},$$

where  $M = (A - BD^{-1}C)^{-1}$ .

(b) When the random vector  $\mathbf{x} = \begin{pmatrix} \mathbf{x}_1 \\ \mathbf{x}_2 \end{pmatrix}$  follows a Gaussian with mean  $\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$  and covariance  $\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$ ,

$$\mathbb{E}[\mathbf{x}_1|\mathbf{x}_2] = \mu_1 + \Sigma_{12}\Sigma_{22}^{-1}(\mathbf{x}_2 - \mu_2), \quad \text{Cov}[\mathbf{x}_1|\mathbf{x}_2] = \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21}$$
$$\mathbb{E}[\mathbf{x}_1] = \mu_1, \quad \text{Cov}[\mathbf{x}_1] = \Sigma_{11}.$$

(1)

Consider a Gaussian likelihood function for one-dimensional data  $X = \{x_i\}_{i=1}^N$ , where each  $x_i \in \mathbb{R}$ , and the parameters are the mean  $\mu$  and precision  $\tau$ . The likelihood is given by:

$$p(X|\mu,\tau) = \prod_{i=1}^{N} \left(\frac{\tau}{2\pi}\right)^{1/2} \exp\left(-\frac{\tau}{2}(x_i - \mu)^2\right).$$

(1) When  $\mu$  is fixed, the conjugate prior for  $\tau$  is the Gamma distribution, which has the form

$$Gam(\tau|a,b) = \frac{1}{\Gamma(a)}b^a\tau^{a-1}\exp(-b\tau),$$

where  $\Gamma(a) = \int_0^\infty x^{a-1} e^{-x} dx$  is the Gamma function. Using this conjugate prior, derive the posterior distribution

$$p(\tau|X) = Gam(\tau|a_N, b_N),$$

by computing the updated parameters  $a_N$  and  $b_N$ .

(2) The conjugate prior for joint parameters  $(\mu, \tau)$  is the normal-Gamma distribution, which has the form

$$p(\mu, \tau) = p(\mu|\tau) \cdot p(\tau)$$

$$= N(\mu|\mu_0, (\beta\tau)^{-1}) \cdot Gam(\tau|a, b)$$

$$= \sqrt{\frac{\beta\tau}{2\pi}} \exp\left(-\frac{\beta\tau}{2}(\mu - \mu_0)^2\right) \cdot \frac{1}{\Gamma(a)} b^a \tau^{a-1} \exp(-b\tau).$$

Using this conjugate prior, derive the posterior distribution

$$p(\mu, \tau | X) = N(\mu | \mu_N, (\beta_N \tau)^{-1}) \cdot Gam(\tau | a_N, b_N),$$

by computing the updated parameters  $\mu_N, \beta_N, a_N, b_N$ .

(Hint: Begin by writing down the joint posterior up to a normalization constant:

$$p(\mu, \tau | X) \propto p(X | \mu, \tau) p(\mu, \tau).$$

First, make a perfect square form (i.e., complete the square) in  $\mu$  within the exponential to identify  $\mu_N$  and  $\beta_N$ . Then, using the remaining terms, extract the exponential and power terms of  $\tau$  to identify  $a_N$  and  $b_N$ .)

Consider modeling a sequence of coin flips. Let  $\theta \in [0,1]$  be the probability of the coin landing heads. The Bernoulli distribution can serve as the likelihood for a single flip:  $p(x|\theta) = \theta^x (1-\theta)^{1-x}$ , where x = 1 for heads and x = 0 for tails.

Assume we use a Beta distribution as the prior distribution for  $\theta$ :

$$p(\theta) = \text{Beta}(\theta | \alpha, \beta) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} \theta^{\alpha - 1} (1 - \theta)^{\beta - 1}$$

where  $\alpha, \beta > 0$  are hyperparameters.

Suppose we observe a dataset  $\mathcal{D} = \{x_1, x_2, \dots, x_N\}$  from N independent coin flips, where  $m = \sum_{i=1}^{N} x_i$  is the number of heads, and N - m is the number of tails. The likelihood function for the dataset is  $p(\mathcal{D}|\theta) = \theta^m (1-\theta)^{N-m}$ .

- (1) Using Bayes' theorem, derive the posterior distribution  $p(\theta|\mathcal{D})$ . Show that the posterior distribution has the same functional form (a Beta distribution) as the prior distribution. Explain why this demonstrates that the Beta distribution is a conjugate prior for the Bernoulli likelihood.
- (2) Express the parameters (the new  $\alpha'$  and  $\beta'$ ) of the derived posterior distribution in terms of the prior parameters  $(\alpha, \beta)$  and the data (N, m).
- (3) If the prior distribution is Beta( $\theta|2,2$ ) and we observe 7 heads in 10 coin flips (N=10, m=7), what is the posterior distribution  $p(\theta|\mathcal{D})$ ?

Consider a Gaussian likelihood  $p(\mathbf{x}|\mu) = N(\mathbf{x}; \mu, \Sigma)$  for  $\mathbf{x} \in \mathbb{R}^D$ , where  $\mu \in \mathbb{R}^D$  is the mean parameter and  $\Sigma \in \mathbb{R}^{D \times D}$  is a fixed covariance matrix. Assume further that the posterior distribution over the mean parameter  $\mu$  is Gaussian:  $p(\mu|X) = N(\mu; \mu_N, \Sigma_N)$ , where  $\mu_N \in \mathbb{R}^D$  and  $\Sigma_N \in \mathbb{R}^{D \times D}$ . Derive the predictive distribution  $p(\mathbf{x}|X)$  by marginalizing over  $\mu$ :

$$p(\mathbf{x}|X) = \int p(\mathbf{x}|\mu)p(\mu|X)d\mu.$$

Use the Woodbury matrix identity if needed:

$$(A + UCV)^{-1} = A^{-1} - A^{-1}U(C^{-1} + VA^{-1}U)^{-1}VA^{-1}.$$

Consider a simple 1D linear regression model  $y = wx + \epsilon$ , where w is the unknown weight (slope), and  $\epsilon$  is Gaussian noise with zero mean and known variance  $\sigma^2$ , i.e.,  $\epsilon \sim \mathcal{N}(0, \sigma^2)$ . The likelihood for a data point (x, y) is therefore:

$$p(y|x, w, \sigma^2) = \mathcal{N}(y|wx, \sigma^2)$$

We place a Gaussian prior distribution on the weight w, with zero mean and variance  $\alpha^{-1}$  (where  $\alpha > 0$  is the precision):

$$p(w|\alpha) = \mathcal{N}(w|0, \alpha^{-1}).$$

Assume we are given a dataset  $\mathcal{D} = \{(\mathbf{x}, \mathbf{y})\} = \{(x_1, y_1), \dots, (x_N, y_N)\}$  consisting of N data points.

- (1) Using Bayes' rule, derive the expression for the log posterior of w,  $\log p(w|\mathcal{D}, \sigma^2, \alpha)$ , up to an additive constant (i.e., you may ignore terms that do not depend on w).
- (2) By rearranging the terms in the log posterior into a quadratic form in w, show that the posterior distribution  $p(w|\mathcal{D}, \sigma^2, \alpha)$  is a Gaussian distribution.
- (3) Identify the mean  $m_N$  and precision  $\alpha_N$  (or equivalently, the variance  $\sigma_N^2 = \alpha_N^{-1}$ ) of the posterior distribution. Express your result in terms of the data  $(\mathbf{x}, \mathbf{y})$ , the noise variance  $(\sigma^2)$ , and the prior precision  $(\alpha)$ . (Hint: Identify the coefficients of the quadratic form in w to identify the posterior parameters.)
- (4) Given the dataset  $(\mathbf{x}, \mathbf{y}) = \{(1, 2), (3, 4)\}$ , noise variance  $\sigma^2 = 1$ , and prior precision  $\alpha = 1$ , calculate the mean  $m_N$  of the posterior distribution.

Consider some positive semi-definite kernel functions  $k(\mathbf{x}, \mathbf{y})$  for  $\mathbf{x}, \mathbf{y} \in \mathbb{R}^D$ , and find the feature mapping  $\phi : \mathbb{R}^D \to \mathbb{R}^F$  corresponding to the kernel function that satisfies

$$k(\mathbf{x}, \mathbf{y}) = \phi(\mathbf{x})^{\mathsf{T}} \phi(\mathbf{y}). \tag{1}$$

Represent  $\phi(\mathbf{x}) \in \mathbb{R}^F$  in a vector form and find the dimensionality F of the feature space.

(1) 
$$k(\mathbf{x}, \mathbf{y}) = (\mathbf{x}^{\top} \mathbf{y})^3$$
, where  $\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}, \mathbf{y} = \begin{pmatrix} y_1 \\ y_2 \end{pmatrix} \in \mathbb{R}^2$ .

- (2)  $k(\mathbf{x}, \mathbf{y}) = 1 + \mathbf{x}^{\top} \mathbf{A} \mathbf{y}$ , where  $\mathbf{x} = \begin{pmatrix} x_1 \\ x_2 \end{pmatrix}$ ,  $\mathbf{y} = \begin{pmatrix} y_1 \\ y_2 \end{pmatrix} \in \mathbb{R}^2$ , and  $\mathbf{A} = \begin{pmatrix} a_{11} & a_{12} \\ a_{12} & a_{22} \end{pmatrix} \in \mathbb{R}^{2 \times 2}$  is a symmetric positive-definite matrix. (Hint: find  $\mathbf{L} = \begin{pmatrix} a & 0 \\ b & c \end{pmatrix} \in \mathbb{R}^{2 \times 2}$  that satisfies  $\mathbf{A} = \mathbf{L} \mathbf{L}^{\top}$ .)
- (3)  $k(\mathbf{x}, \mathbf{y}) = k_1(\mathbf{x}, \mathbf{y}) + k_2(\mathbf{x}, \mathbf{y})$ , where  $k_1(\mathbf{x}, \mathbf{y}) = \psi(\mathbf{x})^\top \psi(\mathbf{y})$  with  $\psi(\mathbf{x}) = [\psi_1(\mathbf{x}) \dots \psi_{F_1}(\mathbf{x})]^\top \in \mathbb{R}^{F_1}$ , and  $k_2(\mathbf{x}, \mathbf{y}) = \xi(\mathbf{x})^\top \xi(\mathbf{y})$  with  $\xi(\mathbf{x}) = [\xi_1(\mathbf{x}) \dots \xi_{F_2}(\mathbf{x})]^\top \in \mathbb{R}^{F_2}$ .

Consider the following linear dynamical system with Gaussian noise:

$$x_{t+1} = ax_t + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma^2), \quad t = 0, 1, 2, 3, \dots$$

where 0 < a < 1, the initial state is distributed as  $x_0 \sim N(1, \sigma_0^2)$ , and the noise terms  $\epsilon_t$  and  $\epsilon_{t'}$  are independent for all  $t \neq t'$ . The collection of  $(x_{t_1}, x_{t_2}, \dots, x_{t_N})$  for distinct time indices  $t_1, \dots, t_N \in \mathbb{N} \cup \{0\}$  is jointly Gaussian hence this system defines a Gaussian process.

Find the mean function  $m(t) = \mathbb{E}[x_t]$  and the covariance function  $k(t_1, t_2) = \text{Cov}(x_{t_1}, x_{t_2})$  of the Gaussian process  $\{x_t\}_{t\geq 0}$ .