Case Study 2: Diabetes Dataset - Prediction and Interpretation of Hospital Readmission For Diabetic Patients

Kristin Henderson

May 27, 2025

1 Introduction

This study aims to predict hospital readmission into one of three categories: within 30 days, after 30 days, or no readmission, and to identify the variables most associated with readmission risk. This is important because high readmission rates can lead to worse patient outcomes, higher healthcare costs, and increased strain on hospital resources. The dataset used in the analysis is a University of California at Irvine Machine Learning Repository dataset, which contains 101,766 hospital encounters of inpatients diagnosed with diabetes who had hospital stays of 1-14 days, collected from 130 U.S. hospitals between 1999 and 2008.

2 Data

The dataset contains a total of 49 features, both numerical and categorical, plus the multi-class categorical target variable, readmitted. Table 1 contains the features in the dataset grouped roughly by category. The medication features are listed only by broader class.

Table 1: Features by Broad Category in the Diabetes Dataset

Broad Category	Features (or Drug Class)		
ID Variables	encounter_id, patient_nbr		
Demographic Variables	race, gender, age, weight		
Admission and Discharge Details	admission_type_id, admission_source_id, dis-		
	charge_disposition_id		
Administrative Information	time_in_hospital, payer_code, medical_specialty		
Medical Diagnostics and Treatments	num_lab_procedures, num_procedures,		
	num_medications, diag_1, diag_2, diag_3, num-		
	ber_diagnosis		
Previous Hospitalization	num_outpatient, num_emergency, num_inpatient		
Lab Results	max_glu_serum, A1Cresult		
Medications	Biguanides, Meglitinides, Sulfonylureas (1st and 2nd		
	generation), Thiazolidinediones, Alpha-Glucosidase		
	Inhibitors, DPP-4 Inhibitors, Insulin, Drug Combi-		
	nations, Non-Diabetes Medications		
Medication Outcome Variables	change, diabetesMed		
Target	readmitted		

Table 2 contains the medication features grouped by drug class.

Table 2: Medication Features by Drug Class

Drug Class	Medications	
Biguanides	metformin	
Meglitinides	repaglinide, nateglinide	
Sulfonylureas 1st Generation	chlorpropamide, acetohexamide, tolbutamide, tolazamide	
Sulfonylureas 2nd Generation	glimepiride, glipizide, glyburide	
Thiazolidinediones (TZDs)	pioglitazone, rosiglitazone, troglitazone	
Alpha-Glucosidase Inhibitors	acarbose, miglitol	
DPP-4 Inhibitors	citoglipton	
Insulin and Combinations	insulin, glyburide-metformin, glipizide-metformin,	
	glimepiride-pioglitazone, metformin-rosiglitazone,	
	metformin-pioglitazone	
Non-Diabetes Medications	examide	

The target variable, readmitted, is a multi-class categorical variable with three levels: less than 30 days, greater than 30 days, and no readmission. The classes are imbalanced,with 53.9% of encounters having no readmissions, 34.9% being readmitted after 30 days, and 11.1% with readmission in under 30 days, as shown in Figure 1.

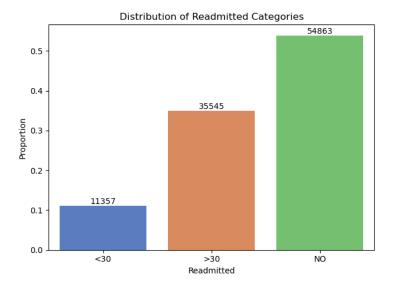


Figure 1: Barplots of proportions of encounters in each readmitted class. The classes are imbalanced with more than half in the not readmitted class and the fewest in the early readmission class. Counts for each class are shown on top of the bars.

2.1 Missing Data Handling and Preprocessing

The lab testing variables max_glu_serum and A1Cresult each had four categorical levels in the raw dataset, including a category none, indicating that the test wasn't conducted. To avoid confusion, I recoded this to untested. I could've split each of these into two features—one binary (tested or not) and one ordinal for results (normal, elevated or high). Instead, I opted to one-hot encode the four categorical levels as is.

Several variables had missing values recorded as question marks in the raw dataset. These were race, medical specialty, payer code, weight, and the three diagnosis code variables.

For race, missing values (2.2% of observations) showed some small trends: patients from the emergency room or with missing payer codes were less likely to have missing race, while those with urgent admissions were more likely. Patients might not have disclosed their race, or some sources may have recorded hard to categorize races as question marks. Rather than imputing with the mode (Caucasian) or trying to model the missingness, I created an Unreported category to preserve this information.

Medical specialty had the highest rate of missing values, with nearly 50% of records missing. The next largest category was internal medicine. Without clear patterns in the missingness, I created an Unknown category rather than imputing values.

About 40,000 records were missing payer code, including every encounter ID below 72091308 (indexes 0-20445). The remaining missing values appeared random. Since the missing category was the most frequent, I treated it as its own Unknown category rather than imputing another value.

The admission and discharge variables (admission_type_id, admission_source_id, and discharge_disposition_id) were composed of integer-coded categorical levels. They each contained two or more categories that could represent missing values (null, not available, not mapped). In admission type, 5.2% were null, and another ~5% were not available or not mapped. All but 4 of the null values also had missing payer codes, suggesting they were missing at random. I remapped the admission type integer codes to the corresponding category and retained all the missing categories.

For admission source, there were \sim 7000 null or not mapped values. All 161 not mapped values were in the first 6000 observations with one exception. Otherwise, I didn't detect a pattern. These could be grouped or imputed with the mode (emergency room). I opted to leave them as is and converted the datatype to avoid implying order.

Fewer than 5000 values in discharge disposition could be considered missing, mostly null and the rest not mapped. They appeared randomly scattered. I kept them as is and converted them to categorical.

96.9% of observations were missing weight. It didn't make sense to impute nearly the entire column, so I dropped it.

I dropped the examide and citoglipton variables because of zero variance. All the values were identical and provided no information to separate target classes. I also dropped the ID variables, encounter_id and patient_nbr.

2.2 Encoding and Scaling

Fewer than 2% of observations had missing diagnosis codes. Since it's reasonable that not every patient would have multiple diagnosis codes, I didn't impute missing values. One observation had no listed diagnosis codes but a number_diagnosis value of 5, which might have been a recording error.

I used a multi-column count encoding. Each observation got a vector of diagnosis code counts (ranging from one to three, about 98% of observations had three). I dropped the observation with no diagnosis codes to avoid missing values rather than choosing to impute the mode for diagnosis code 1. I created a column for each unique diagnosis code and recorded the count of each code per encounter. Each diagnosis variable had 717-790 unique codes, but there was substantial overlap, leaving ~900 unique codes. About 6% of patients had duplicate codes (or after count encoding, a count of more that one for a code).

I one-hotencoded the remaining categorical variables, dropping the first category to prevent multicollinearity. For k levels, this created k-1 dummy variables, with the dropped category becoming the reference level. I considered converting age to a numeric variable, and using the middle of the range as the value for each observation. I decided against this, because I didn't want to assume a linear relationship between age and readmission risk. Another option would

have been to give age a numeric value based on the ordinal scale, but again, I decided to treat it as categorical and sacrifice the ordinal information.

Because many of the numerical variables were not normally distributed and had high outliers, I opted for a robust scaler to scale the data.

3 Modeling

A logistic regression model was chosen for its high interpretability and computational efficiency.

Several steps went into my search for the best model parameters. I started by comparing two solvers, saga (Stochastic Average Gradient Augmented) and lbfgs (Limited-memory BFGS), at several values of the regularization parameter C. I used a grid search and stratified shuffling of the target variable with 5-fold cross-validation, optimizing for the mean out-of-fold weighted F1 score. I chose this metric because I wanted to prioritize identifying patients at risk of readmission, while also minimizing incorrect readmission predictions.

F1 score balances recall (also called sensitivity), the proportion of correctly predicted positives out of all actual positives, and precision (or positive predictive value), the proportion of correctly predicted positives out of all predicted positives. This is important because it helps ensure that at risk patients receive appropriate care, while avoiding unnecessary interventions for those who are not truely at risk. I chose weighted F1 score because of the imbalance in the target classes.

I expected saga to perform well because the dataset contains over 1000 features (after encoding), 100,000+ records, and many sparse features from high-cardinality categorical variables. By using stochastic gradient descent, it is designed for use with large, sparse datasets. However, it was significantly slower, over 12 minutes for a single fit compared to 38 seconds for 1bfgs at C=1 with only 0.04% improvement in weighted F1. That made me realize that, despite its size, this dataset is still small enough to fit in memory and might be better suited to solvers that batch process like 1bfgs.

Next, I compared lbfgs with newton-cg (Newton's method using the Conjugate Gradient algorithm) expanding the range of C values up to 10 and 100. Although newton-cg at C=100 resulted in the highest weighted F1 score, the improvement over lbfgs was minimal at just 0.04%. I ultimately chose lbfgs for its robustness. Rather than computing a Hessian matrix (a matrix of second derivatives for precise optimization) directly, it approximates one. This can make it less sensitive to poor initial conditions and local minima.

I ultimately chose the Scikit-Learn's logistic regression model's default parameters, solver=lbfgs and C=1.0. A weaker regularization parameter, C=10 resulted in a minimal increase in weighted F1, 0.02% but increased computation time by nearly double.

Rather than using class_weight='balanced', which sets the weight of each class inversely proportional to its frequency in the data, I also tuned class_weight to optimize the weighted F1 score. I selected weights of 2 for

the <30 class, 1.5 for the >30 class and 1 for the NO class, increasing the importance of the two less frequent classes. This weighting was not as extreme as given by the balanced class weight parameter, which is based on the class proportions.

I also optimized the classification threshold for the <30 class, choosing a threshold which optimized F1 score on 5-fold cross validated predictions. This prioritized trade-offs in performance for detecting early readmissions.

4 Results

4.1 Model Performance and Threshold Tuning

I chose to examine Precision-Recall (PR) curves over ROC curves because I prioritized these two metrics in this analysis. Recall measures the proportion of actual positive cases that were correctly identified, while precision measures the proportion of predicted positive cases that were actually positive. As recall increases and more true positives are correctly classified, typically there are also more false positives, decreasing precision. These show the trade-offs between correctly identifying positive cases (recall) and avoiding false positives (precision). In this scenario, I wanted to maximize both metrics, particularly for the early readmission class.

In Figure 2, PR curves for each readmission class (Top: early readmission, Middle: later readmission, Bottom: not readmitted) show the trade-off between these two metrics. The threshold which optimizes F1 score for each class is shown in red, and the threshold selected for optimizing for the <30 class is marked by the yellow X in the middle and bottom plots.

ROC curves can be misleading in imbalanced datasets because they plot true positive rate (sensitivity) against false positive rate (1-specificity). If a positive class is rare, then a model can still achieve high scores in both by predicting mostly negative cases. PR curves help show how well a model performs on the minority class.

The weighted one-versus-rest area under the ROC curve (AUC-ROC) was 0.6798 (Table 3), with individual class AUC-ROC values of 0.67 for early, 0.66 for later, and 0.70 for no readmission (Table 4). To find the weighted AUC-ROC, these individual values are weighted by the proportion of each class in the dataset (11.1%, 34.9%, and 53.9% respectively). This is a threshold-independent metric of model performance that accounts for the class imbalance.

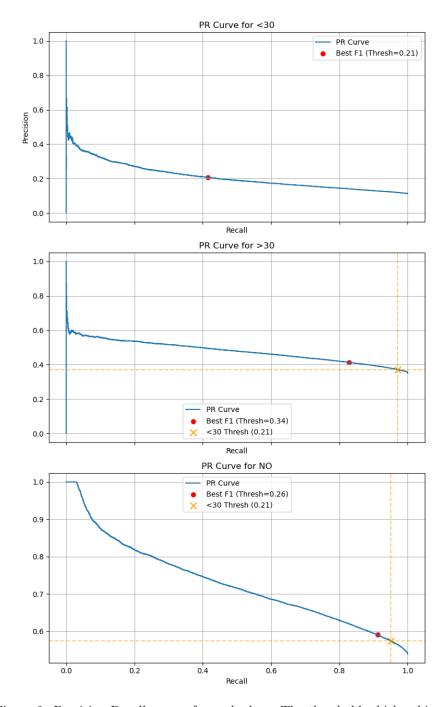


Figure 2: Precision-Recall curves for each class. The threshold which achieves the highest F1 score for each class is marked by the red circle and the chosen threshold is shown by the yellow ${\tt X}$.

Table 3 contains the overall metrics from the cross-validated model with the optimized <30 threshold.

Table 3: Overall Cross-Validated Out-Of-Fold Metrics

Metric	Mean	Standard Deviation
Weighted F1-score	0.5529	0.0039
Weighted Recall	0.5145	0.0058
Weighted Precision	0.5518	0.0033
Weighted AUC	0.6798	0.0036
Accuracy	0.5145	0.0040

Table 4 contains results from the classification report after optimizing the <30 threshold for F1 score.

Table 4: Metrics by Readmission Class

Metric	<30	>30	NO
F1-score	0.2771	0.4271	0.6373
Recall	0.4154	0.4014	0.5963
Precision	0.2079	0.4562	0.6844
AUC	0.6681	0.6562	0.6975

The confusion matrices, in Figure 3, help demonstrate the trade-offs in optimizing the classification threshold for the <30 class. Using the default threshold (upper matrix), the argmax, only 1016 <30 patients were correctly classified. Yet $\sim\!21,\!000$ >30 and $\sim\!35,\!000$ NO classifications were correct. Lowering the threshold to 0.21 to optimize for correct predictions in the early readmission class (lower matrix) increased the number of correct predictions in that class to 4718 but also drastically increased the number of incorrect early readmission predictions from roughly 2000 to 18,000. This also reduced the correct >30 and NO predictions to $\sim\!14,\!000$ and $\sim\!32,\!000$, respectively. This will likely increase the cost of care for many patients, but with the hope of improving outcomes for those at risk.

4.2 Class-Specific Metrics

Optimizing the <30 threshold increased the recall for the early readmission class from 9% to 42% and its F1 score from 14% to 28% but reduced its precision from 32% to 21%, as shown in Table 4. The improved classification in early readmissions lowered recall in the other classes and F1 score in the later readmissions. It also lowered the overall accuracy from 57% to 51% (Table 3).

This trade-off suggests that prioritizing detection of patients with higher risk of early readmission comes primarily at a cost of detecting later readmissions.

4.3 Feature Importance

Figure 4 displays bar plots of the five most important features and their coefficients from each readmission class (Top: early readmission, Middle: later readmission, Bottom: not readmitted). The features and their descriptions are listed in Table ??. Three larger variable categories stand out: discharge disposition, admission type and diagnosis. The most important feature in determining all the classes is discharge disposition 11, the code for expired. Of course, if a patient died, they were not readmitted. Unfortunately, that information is unlikely to help prevent future readmissions. Diagnosis code 14, to a hospice/medical facility, is not a strong predictor of early readmission, but it does have a negative association with later readmissions and a positive one with not readmitted. Additionally, admission through a trauma center is positively associated with not being readmitted.

The rest of the important variables are specific diagnoses. Two diagnoses are important predictors of early versus late readmissions. Both tick-borne rick-ettsioses (82) and disorders of carbohydrate transport and metabolism (271) are positively associated with early readmission and negatively associated with late readmission. Other strong positive predictors of early readmission are rheumatic fever with heart involvement and respiratory conditions due to chemical fumes and vapors. A strong predictor of late readmission is a diagnosis of cataract. One can speculate, but this may be confounded by age. Foot fracture and jaw disease were somewhat surprisingly negatively associated with not being readmitted.

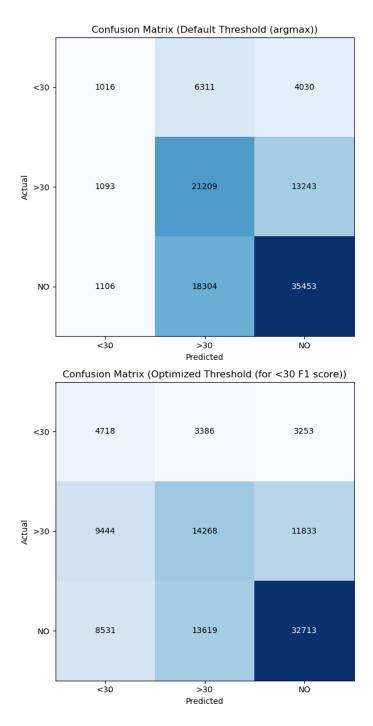


Figure 3: Confusion matrixes for the tuned model with both the default (argmax) (Top) and optimized <30 (Bottom) thresholds.

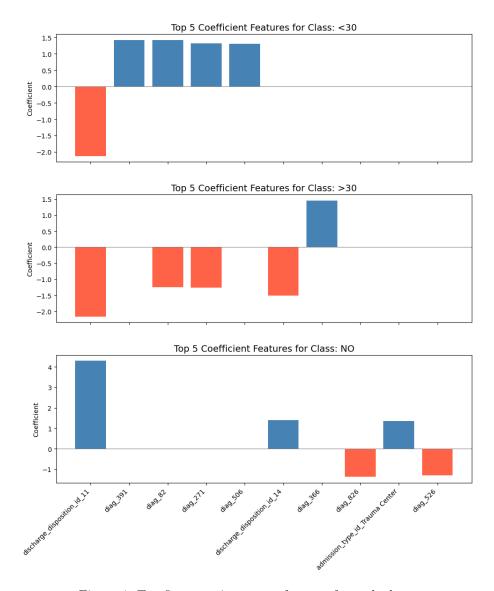


Figure 4: Top five most important features for each class.

Table 5 contains the five most important features for each of the readmission classes and their description. Some classes share important features.

Table 5: Most Important Features of all Readmission Classes

Variable	Description
discharge_disposition_id_11	expired
discharge_disposition_id_14	hospice/medical facility
admission_type_id_Trauma Center	admission via Trauma Center
diag_391	Rheumatic fever with heart involvement
diag_82	Tick-borne rickettsioses
diag_271	Disorders of carbohydrate transport and metabolism
diag_506	Respiratory conditions due to chemical fumes and vapors
diag_366	Cataract
diag_826	Fracture of one or more phalanges of foot
diag_526	Diseases of the jaws

5 Conclusion

Optimizing the classification threshold for the early readmission (<30) class resulted in many more true positives but also increased false positives for that class. However, it drastically reduced false negatives (from the NO, not readmitted class). This improvement in sensitivity (recall) came at a cost of reduced accuracy for both the later readmission (>30) and the not readmitted (NO) classes.

Based on recall scores, the model correctly identified 41.5% of patients readmitted within 30 days, 40.1% of those readmitted after 30 days, and 59.6% of those not readmitted.

Several diagnoses are associated with early readmission: tick-borne disease, carbohydrate metabolic disorder, rheumatic fever involving the heart and respiratory conditions due to fumes.

An important consideration from the feature importance results is whether demographic variables like race, gender, and age should be included in a similar model going forward. There are pros and cons to consider. There may be a slight predictive advantage in including these. Perhaps certain subgroups are at greater or lesser risk and knowing that, the best treatment options could be prioritized. This also introduces the possibility of bias. One option would be to code these variables and let only certain individuals have access to the true categories. Of course, these variables may be confounded by others like payer code or comorbidities. Examining multicollinearity may be helpful. One option would be to exclude these variables from the final model, but retain the information for monitoring associations with readmission or correlations with other variables in future models.