Let (\mathbf{x}_i, y_i) denote the training data pair of feature vectors and labels. The regularized logistic regression max log-likelihood can be written as

$$\min_{\mathbf{w},b} E = \frac{\lambda}{2} \|\mathbf{w}\|^2 + \sum_{i} g(-y_i(\mathbf{w}^T \mathbf{x}_i + b)),$$
$$g(\psi) = \log(1 + e^{\psi}).$$

Note that $g(\psi)$ is the negative of the log-likelihood associated with the probability model,

$$p(y|\mathbf{x}) = \frac{1}{1 + \exp(-y\mathbf{w}^T\mathbf{x} + b)}.$$

We can rewrite E as the constrained minimization

$$\min E = \frac{\lambda}{2} \|\mathbf{w}\|^2 + \sum_{i} g(\psi_i),$$
$$\psi_i = -y_i(\mathbf{w}^T \mathbf{x}_i + b), \ \forall i.$$

Thus, the Lagrangian for this problem is

$$L = \frac{\lambda}{2} \|\mathbf{w}\|^2 + \sum_{i} g(\psi_i) + \sum_{i} \alpha_i \left[-\psi_i - y_i(\mathbf{w}^T \mathbf{x}_i + b) \right],$$

which has the KKT optimality conditions given by

$$\nabla_{\mathbf{w}} L = \lambda \mathbf{w} - \sum_{i} \alpha_{i} y_{i} \mathbf{x}_{i} = 0;$$
$$\frac{\partial L}{\partial b} = \sum_{i} \alpha_{i} y_{i} = 0;$$
$$\frac{\partial L}{\partial \phi_{i}} = g'(\psi_{i}) - \alpha_{i} = 0, \ \forall i.$$

Hence, these conditions give

$$\mathbf{w}(\alpha) = \frac{1}{\lambda} \sum_{i} \alpha_i y_i \mathbf{x}_i; \quad \psi_i(\alpha_i) = (g')^{-1}(\alpha_i),$$

and the constraint $\sum_i \alpha_i y_i = 0$. To continue, let $G(\alpha_i) = g(\psi_i) - \alpha_i \psi_i$, where this is part of L, then

$$\frac{\partial G}{\partial \alpha_i} = \frac{\partial \psi_i}{\partial \alpha_i} g'(\psi_i) - \alpha_i \frac{\partial \psi_i}{\partial \alpha_i} - \psi_i = -\psi_i = -(g')^{-1}(\alpha_i).$$

Therefore, $G'(\alpha_i) = -(g')^{-1}(\alpha_i)$. We can now substitute in the logistic regression

probability model and find $(g')^{-1}(z)$,

$$g(z) = \log(1 + e^{z}),$$

$$g'(z) = e^{z}/(1 + e^{z}),$$

$$g'(z)(1 + e^{z}) = e^{z},$$

$$g'(z) = e^{z}(1 - g'(z))$$

$$g'(z)/(1 - g'(z)) = e^{z}$$

$$\log(g'(z)/(1 - g'(z))) = z = (g')^{-1}(g'(z)).$$

Thus,

$$(g')^{-1}(u) = \log(u/(1-u)),$$

 $G(\alpha_i) = -\alpha_i \log \alpha_i - (1-\alpha_i) \log(1-\alpha_i).$

This gives us the new objective

$$\min_{\alpha,b} L = \frac{\lambda}{2} \|\mathbf{w}(\alpha)\|^2 + \sum_{i} G(\alpha_i) - \sum_{i} \alpha_i y_i (\mathbf{w}(\alpha)^T \mathbf{x}_i + b),$$
s.t.
$$\sum_{i} \alpha_i y_i = 0.$$

Due to the constraint, the term $\sum_i \alpha_i y_i b = 0$. Note that this is also why you cannot just append a column of ones to the original feature vector to get the bias implicitly—this weight will just go to zero. Then, we can substitute in $\mathbf{w}(\alpha)$ and get

$$L = \frac{1}{2\lambda} \sum_{i,j} \alpha_i \alpha_j y_i y_j \mathbf{x}_i^T \mathbf{x}_j + \sum_i G(\alpha_i) - \frac{1}{\lambda} \sum_{i,j} \alpha_i \alpha_j y_i y_j \mathbf{x}_i^T \mathbf{x}_j.$$

Combining terms and putting it all together with $K = [\phi(\mathbf{x}_i)^T \phi(\mathbf{x}_i)]^{n \times n}$, we get

$$\min_{\alpha} E(\alpha) = -\frac{1}{2\lambda} (\alpha \circ \mathbf{y})^T K(\alpha \circ \mathbf{y}) + \sum_{i} G(\alpha_i),$$

where \circ is the Hadamard product (.* in Matlab). This provides the optimal α , which is used to compute $\mathbf{w}(\alpha)$, but does not compute b, which is essential. However, remember that we are maximizing the log-likelihood of the logistic probability model,

$$p(y|\mathbf{x}) = \frac{1}{1 + \exp(-y(\mathbf{w}(\alpha)^T \mathbf{x} + b)}.$$

Hence, we can substitute in $\mathbf{w}(\alpha)^T \mathbf{x}_j = \sum_i \alpha_i y_i \mathbf{x}_i^T \mathbf{x}_j = (\alpha \circ \mathbf{y})^T K_j$, where K_j is the jth column of K. Then, we can maximize the log-likelihood wrt b,

$$\max_{b} \log \left(\prod_{j} p(y_j | \mathbf{x}_j) \right) = \min_{b} \sum_{j} \log \left(1 + \exp(-y_j ((\alpha \circ \mathbf{y})^T K_j + b)) \right).$$

Note that optimization objective can be vectorized in Matlab as

$$\operatorname{sum} \log(1 + \exp(-\mathbf{y} \cdot * (K(\alpha \circ \mathbf{y}) + b))).$$

This is an unconstrained optimization; hence, you can use fminunc.