Bayesian Networks

- Syntax
- Semantics
- Parametrized Distributions
- Inference in Bayesian Networks
 - Exact Inference
 - enumeration
 - variable elimination
 - Approximate Inference
 - stochastic simulation
 - Markov Chain Monte Carlo (MCMC)



Inference Tasks

Simple queries

compute the posterior marginal distribution for a variable

Conjunctive queries

compute the posterior for a conjunction of variables

$$\mathbf{P}(X_i, X_j | \mathbf{E} = \mathbf{e}) = \mathbf{P}(X_i | \mathbf{E} = \mathbf{e}) \cdot \mathbf{P}(X_j | X_i, \mathbf{E} = \mathbf{e})$$

Optimal decisions

- decision networks include utility information
- probabilistic inference required for P(outcome|action,evidence)

Value of Information

Which evidence to seek next?

Sensitivity Analysis

Which probability values are most critical?

Explanation

Why do I need a new starter motor?

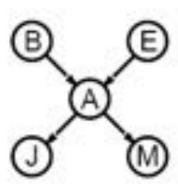
Inference by Enumeration

Slightly intelligent way to sum out variables from the joint without actually constructing its explicit representation

Simple query on the burglary network:

$$P(B|j, m)$$

= $P(B, j, m)/P(j, m)$
= $\alpha P(B, j, m)$
= $\alpha \sum_{e} \sum_{a} P(B, e, a, j, m)$



Rewrite full joint entries using product of CPT entries:

$$\begin{split} &\mathbf{P}(B|j,m) \\ &= \alpha \ \sum_{e} \sum_{a} \mathbf{P}(B)P(e)\mathbf{P}(a|B,e)P(j|a)P(m|a) \\ &= \alpha \mathbf{P}(B) \ \sum_{e} P(e) \ \sum_{a} \mathbf{P}(a|B,e)P(j|a)P(m|a) \end{split}$$

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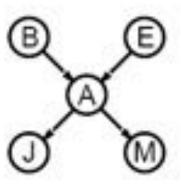
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Worst case: $O(n d^n)$ time $O(d^n)$ terms, each consisting of a product of O(n) probabilities



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$$= \alpha \sum_{e} \sum_{a} \mathbf{P}(B)P(e)\mathbf{P}(a|B, e)P(j|a)P(m|a)$$

=
$$\alpha P(B) \sum_{e} P(e) \sum_{a} P(a|B, e)P(j|a)P(m|a)$$

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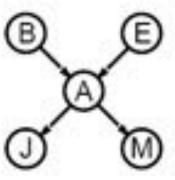
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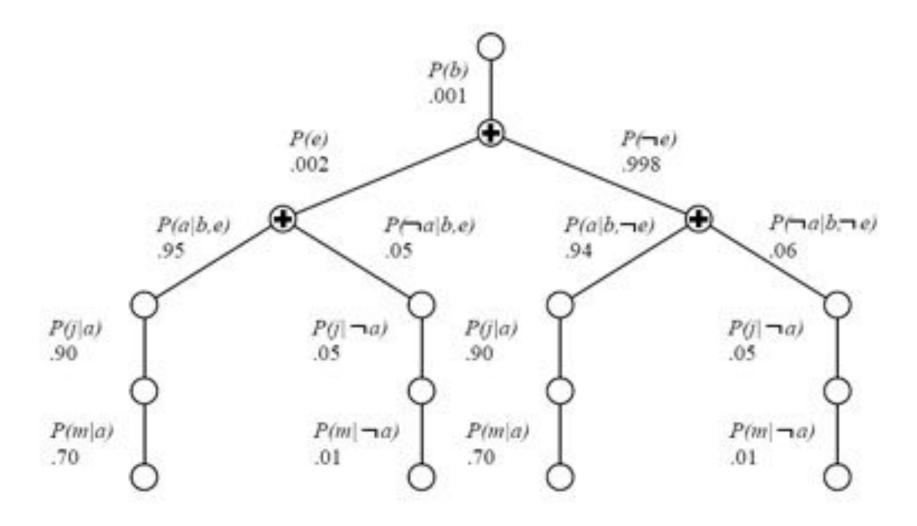
Recursive depth-first enumeration: O(n) space, $O(d^n)$ time

where n is the number of variables and d is the number of values per variable

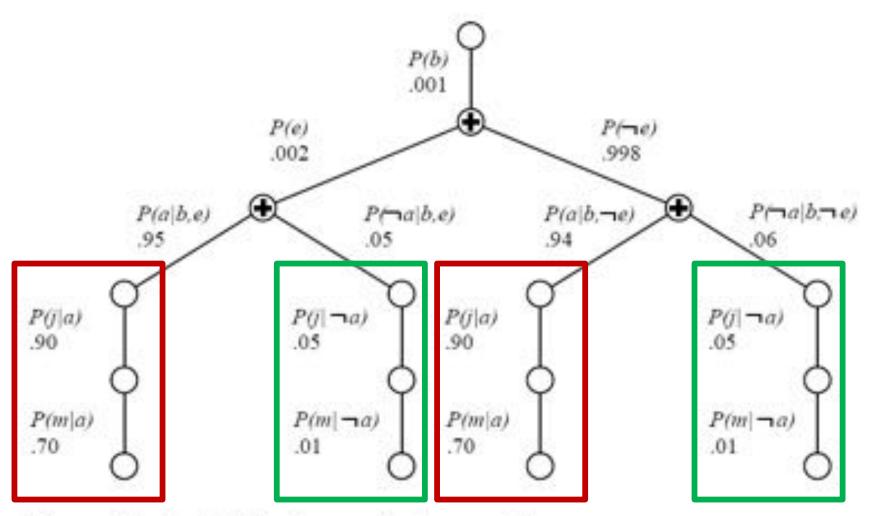
Enumeration Algorithm

```
function Enumeration-Ask(X, e, bn) returns a distribution over X
   inputs: X, the query variable
            e, observed values for variables E
             bn, a Bayesian network with variables \{X\} \cup \mathbf{E} \cup \mathbf{Y}
   Q(X) \leftarrow a distribution over X, initially empty
   for each value x_i of X do
       extend e with value x_i for X
       Q(x_t) \leftarrow \text{Enumerate-All(Vars[bn], e)}
   return Normalize(Q(X))
function ENUMERATE-ALL(vars, e) returns a real number
   if Empty?(vars) then return 1.0
   Y \leftarrow \text{First}(vars)
   if Y has value y in e
        then return P(y \mid Pa(Y)) \times \text{Enumerate-All(Rest(vars), e)}
        else return \Sigma_y P(y \mid Pa(Y)) \times Enumerate-All(Rest(vars), e_y)
            where e_y is e extended with Y = y
```

Evaluation Tree



Evaluation Tree



Enumeration is inefficient: repeated computation e.g., computes P(j|a)P(m|a) for each value of e

Variable Elimination

Move the sums into the products

- Key idea:
 - Do not multiply left-to-right but right-to-left.
 - Thus, terms that appear inside sums are evaluated first
 - intermediate results are stored as so-called factors
 - factors can be re-used several times in the same computation
 - is a form of dynamic programming

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Example: P(B|j,m)

Factors

- A factor is a vector / matrix containing all probabilities for all dependent variables
- Examples:

•
$$\mathbf{f}_{M}(A) = \begin{vmatrix} P(m \mid a) \\ P(m \mid \neg a) \end{vmatrix}$$

• The factor $f_A(A,B,E)$ is a 2 x 2 x 2 matrix

Basic Operations

- Summing Out a variable from a product of factors
 - move all constant factors outside of the summation
 - add up submatrices in pointwise product of remaining factors

$$\sum_{x} \mathbf{f}_{1} \times ... \times \mathbf{f}_{k} = \mathbf{f}_{1} \times ... \times \mathbf{f}_{i} \times \sum_{x} \mathbf{f}_{i-1} \times ... \times \mathbf{f}_{k}$$
$$= \mathbf{f}_{1} \times ... \times \mathbf{f}_{i} \times \mathbf{f}_{X}$$

assuming \mathbf{f}_{l} , ..., \mathbf{f}_{i} do not depend on X

Basic Operations

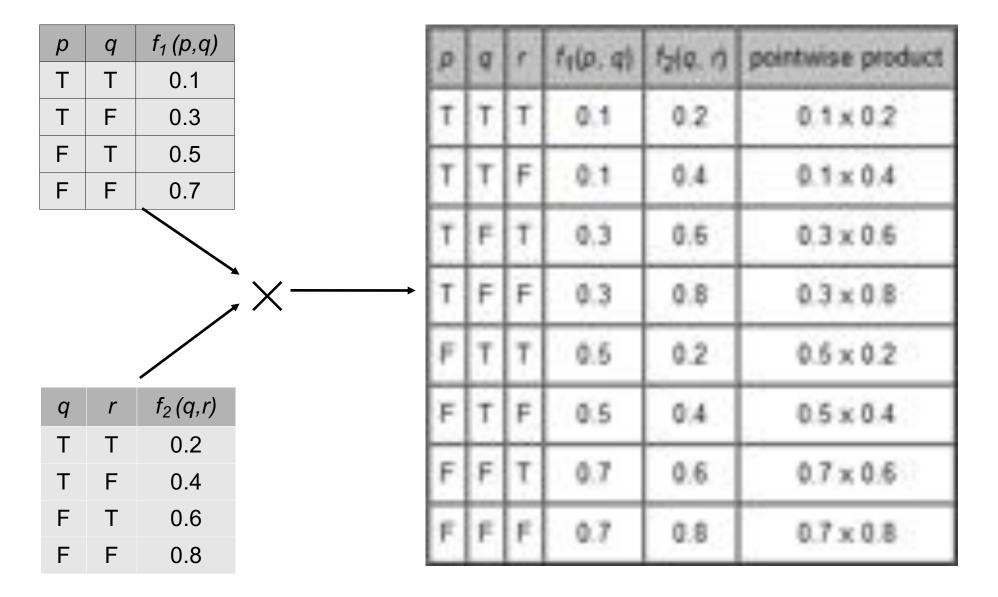
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assuming \mathbf{f}_{l} , ..., \mathbf{f}_{i} do not depend on X

- Pointwise Product of factors f₁ and f₂
 - for example: $\mathbf{f_1}(A,B) \times \mathbf{f_2}(B,C) = \mathbf{f}(A,B,C)$
 - in general: $\mathbf{f_1}(X_1,...,X_j, \underline{Y_1},...,\underline{Y_k}) \times \mathbf{f_2}(\underline{Y_1},...,\underline{Y_k},Z_1,...,Z_l) = \mathbf{f}(X_1,...,X_j,\underline{Y_1},...,\underline{Y_k},Z_1,...,Z_l)$
 - has 2^{j+k+l} entries (if all variables are binary)

Example: Pointwise Product



Variable Elimination Algorithm

```
function ELIMINATION-ASK(X, e, bn) returns a distribution over X inputs: X, the query variable

e, evidence specified as an event

bn, a belief network specifying joint distribution P(X_1, \ldots, X_n)

factors \leftarrow []; vars \leftarrow Reverse(Vars[bn])

for each var in vars do

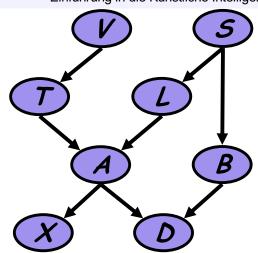
factors \leftarrow [Make-Factor(var, e)|factors]

if var is a hidden variable then factors \leftarrow Sum-Out(var, factors)

return Normalize(Pointwise-Product(factors))
```

We want to compute P(d)Need to eliminate: $v_s, x_s, t_s, l_s, a_s, b_s$

Initial factors



$$P(v)P(s)P(t \mid v)P(l \mid s)P(b \mid s)P(a \mid t,l)P(x \mid a)P(d \mid a,b)$$

Eliminate: *v*

Compute: $f_v(t) = \sum_{v} P(v)P(t|v)$

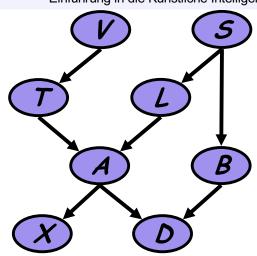
 $\Rightarrow f_{v}(t)P(s)P(l|s)P(b|s)P(a|t,l)P(x|a)P(d|a,b)$

Note: $f_v(t) = P(t)$

In general, result of elimination is not necessarily a probability term

Need to eliminate: $v_1s_1x_1t_1/a_1b_1$

Initial factors



$$P(v)P(s)P(t \mid v)P(l \mid s)P(b \mid s)P(a \mid t,l)P(x \mid a)P(d \mid a,b)$$

$$\Rightarrow f_{v}(t)P(s)P(l|s)P(b|s)P(a|t,l)P(x|a)P(d|a,b)$$

Eliminate: 5

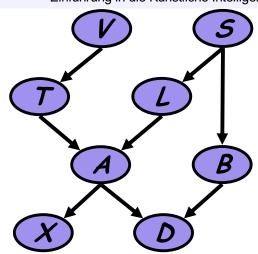
Compute: $f_s(b,l) = \sum P(s)P(b|s)P(l|s)$

 $\Rightarrow f_{v}(t)f_{s}(b,l)P(a|t,l)P(x|a)P(d|a,b)$

Summing on s results in a factor with two arguments $f_s(b,l)$ In general, result of elimination may be a function of several variables

Need to eliminate: v,s,x,t,l,a,b

Initial factors



$$P(v)P(s)P(t | v)P(l | s)P(b | s)P(a | t, l)P(x | a)P(d | a, b)$$

$$\Rightarrow f_v(t)P(s)P(l|s)P(b|s)P(a|t,l)P(x|a)P(d|a,b)$$

$$\Rightarrow f_{v}(t)f_{s}(b,l)P(a \mid t,l)P(x \mid a)P(d \mid a,b)$$

Eliminate: x

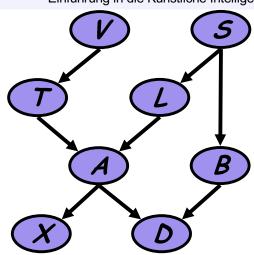
Compute:
$$f_x(a) = \sum_{x} P(x \mid a)$$

$$\Rightarrow f_{v}(t)f_{s}(b,l)\underline{f_{x}}(a)P(a|t,l)P(d|a,b)$$

Note: $f_x(a) = 1$ for all values of a!!

Need to eliminate: $v_1s_1x_1t_1/a_1b_1$

Initial factors



$$P(v)P(s)P(t \mid v)P(l \mid s)P(b \mid s)P(a \mid t,l)P(x \mid a)P(d \mid a,b)$$

$$\Rightarrow f_{v}(t)P(s)P(l|s)P(b|s)P(a|t,l)P(x|a)P(d|a,b)$$

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$$\Rightarrow f_{v}(t)f_{s}(b,l)f_{x}(a)P(a|t,l)P(d|a,b)$$

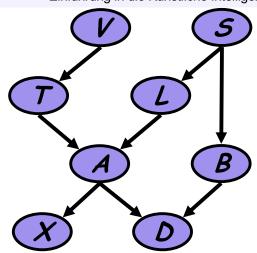
Eliminate: *t*

Compute:
$$f_t(a,l) = \sum_t f_v(t) P(a \mid t, l)$$

 $\Rightarrow f_s(b,l) f_s(a) f_t(a,l) P(d \mid a,b)$

Need to eliminate: v,s,x,t,l,a,b

Initial factors



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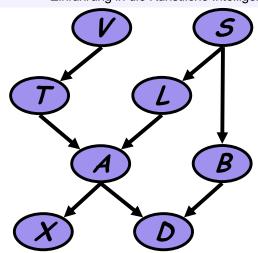
$$\Rightarrow f_v(t)f_s(b,l)f_x(a)P(a|t,l)P(d|a,b)$$

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Eliminate: /
Compute: $f_t(a,b) = \sum f_s(b,l)f_s(a,l)$

Compute: $f_l(a,b) = \sum_s f_s(b,l) f_t(a,l)$ $\Rightarrow f_l(a,b) f_x(a) P(d \mid a,b)$

Need to eliminate: $v_1s_1x_1t_1/a_1b_1$

Initial factors



$$P(v)P(s)P(t | v)P(l | s)P(b | s)P(a | t, l)P(x | a)P(d | a, b)$$

$$\Rightarrow f_v(t)P(s)P(l|s)P(b|s)P(a|t,l)P(x|a)P(d|a,b)$$

$$\Rightarrow f_{v}(t)f_{s}(b,l)P(a|t,l)P(x|a)P(d|a,b)$$

$$\Rightarrow f_{v}(t)f_{s}(b,l)f_{s}(a)P(a|t,l)P(d|a,b)$$

$$\Rightarrow f_s(b,l)f_r(a)f_t(a,l)P(d \mid a,b)$$

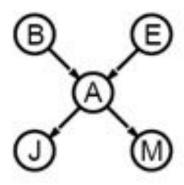
$$\Rightarrow \underline{f_l(a,b)}\underline{f_x(a)}P(d\mid a,b) \Rightarrow \underline{f_a(b,d)} \Rightarrow \underline{f_b(d)}$$

Eliminate: a,b

Compute:
$$f_a(b,d) = \sum_a f_l(a,b) f_x(a) p(d | a,b)$$
 $f_b(d) = \sum_b f_a(b,d)$

Consider the query P(JohnCalls|Burglary=true)

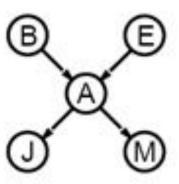
$$P(J|b) = \alpha P(b) \sum\limits_{e} P(e) \sum\limits_{a} P(a|b,e) P(J|a) \sum\limits_{m} P(m|a)$$



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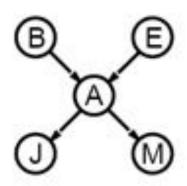
Sum over m is identically 1; M is irrelevant to the query



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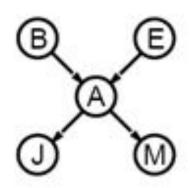
Thm 1: Y is irrelevant unless $Y \in Ancestors(\{X\} \cup \mathbf{E})$

Here, X = JohnCalls, $\mathbf{E} = \{Burglary\}$, and $Ancestors(\{X\} \cup \mathbf{E}) = \{Alarm, Earthquake\}$ so MaryCalls is irrelevant

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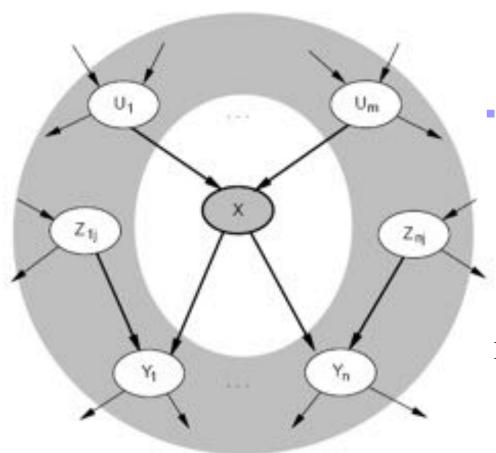
Here,
$$X = JohnCalls$$
, $\mathbf{E} = \{Burglary\}$, and $Ancestors(\{X\} \cup \mathbf{E}) = \{Alarm, Earthquake\}$ so $MaryCalls$ is irrelevant

Note: This is similar to backward chaining from a query in Prolog

(Directed) Markov Blanket

Markov Blanket:

parents + children + children's parents



 Each node is conditionally independent of all other nodes given its markov blanket

$$\mathbf{P} \ X \mid U_{1,}..., U_{m}, Y_{1,}..., Y_{n}, Z_{1j}, ..., Z_{nj} =$$

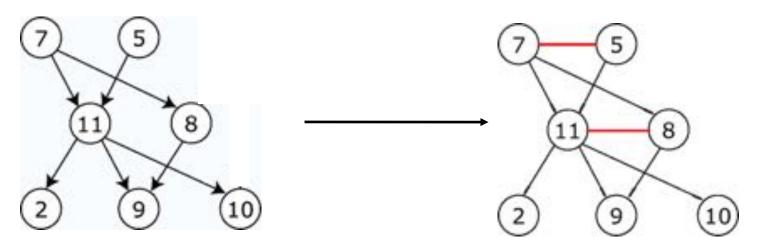
$$= \mathbf{P} \ X \mid all \ variables$$

Moral Graph

- The moral graph is an undirected graph that is obtained as follows:
 - connect all parents of all nodes
 - make all directed links undirected

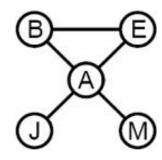
Note:

- the moral graph connects each node to all nodes of its Markov blanket
 - it is already connected to parents and children
 - now it is also connected to the parents of its children



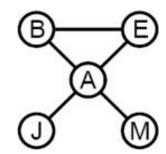
Moral Graph and Irrelevant Variables

- m-separation:
 - variable X is m-separated from Y by Z iff it is separated by Z in the moral graph
- Example:
 - J is m-separated from E by A



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- Example:
 - J is m-separated from E by A



Theorem 2: *Y* is irrelevant if it is m-separated from *X* by *Z*

Example:

For P(JohnCalls|Alarm=true), both Burglary and Earthquake are irrelevant

Singly connected networks (or polytrees):

- any two nodes are connected by at most one (undirected) path
- time and space cost of variable elimination are $O(d^k n)$

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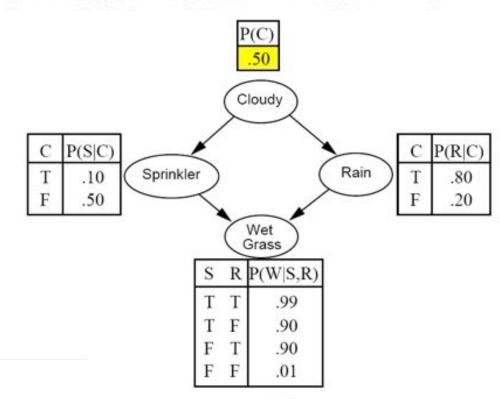
Multiply connected networks:

- can reduce 3SAT to exact inference

NP-hard

Example:

Two paths from Cloudy to Wet Grass



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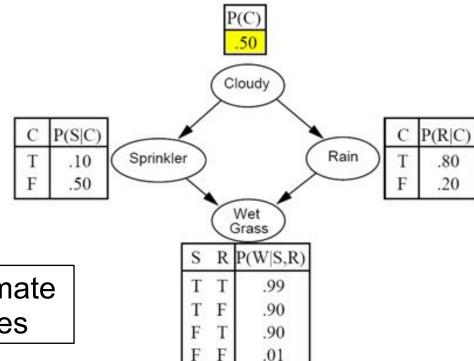
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→ we "need" approximate inference techniques

Complexity of Inference

Theorem:

Inference in Bayesian networks (even approximate, without proof) is NP-hard

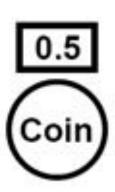
Inference by Stochastic Simulation (Sampling from a Bayesian Network)

Basic idea:

- 1) Draw N samples from a sampling distribution S
- 2) Compute an approximate posterior probability P
- 3) Show this converges to the true probability P

Outline:

- Sampling from an empty network
- Rejection sampling: reject samples disagreeing with evidence
- Likelihood weighting: use evidence to weight samples
- Markov chain Monte Carlo (MCMC): sample from a stochastic process whose stationary distribution is the true posterior



How to draw a sample?

Given random variable X, $D(X)=\{0, 1\}$ Given $P(X) = \{0.3, 0.7\}$

How to draw a sample?

```
Given random variable X, D(X)=\{0, 1\}
Given P(X) = \{0.3, 0.7\}
```

```
Sample X ← P (X)

draw random number r ∈ [0, 1]

If (r < 0.3) then set X=0

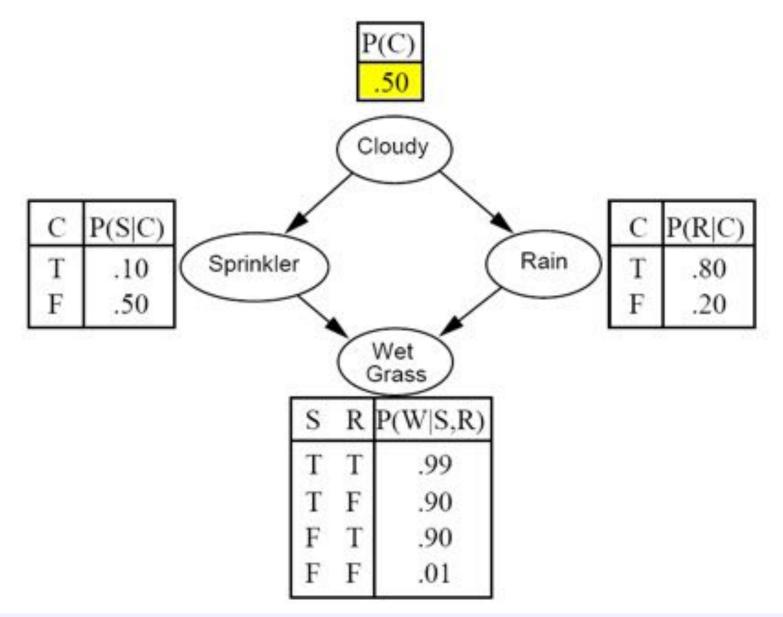
Else set X=1
```

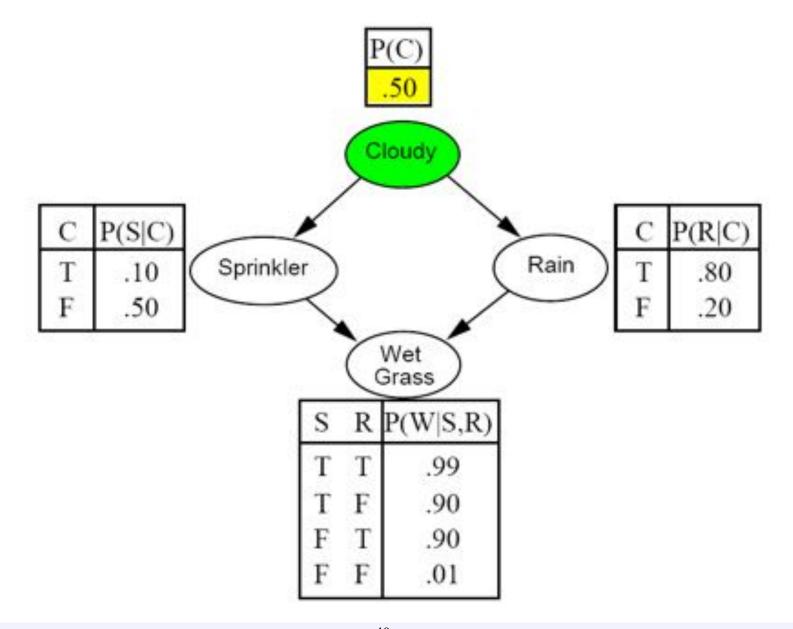
Can generalize for any domain size

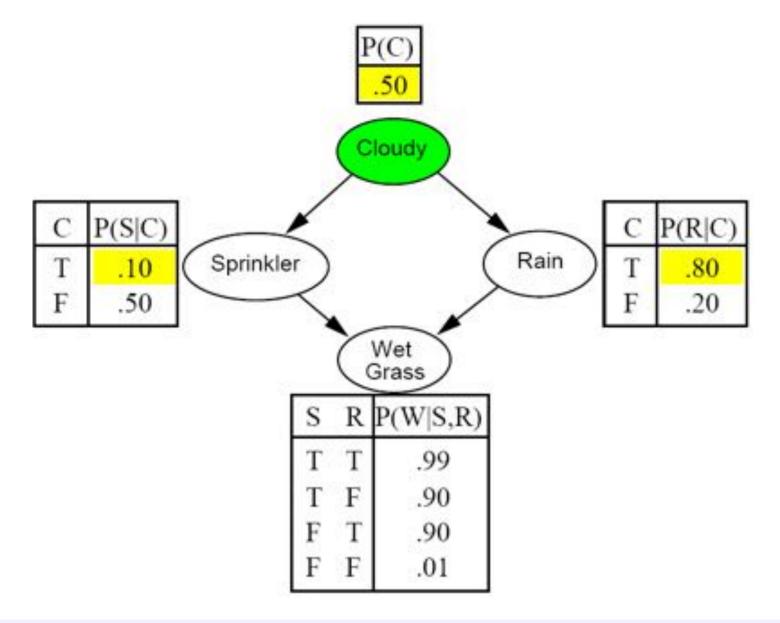
Sampling from an "Empty" Network

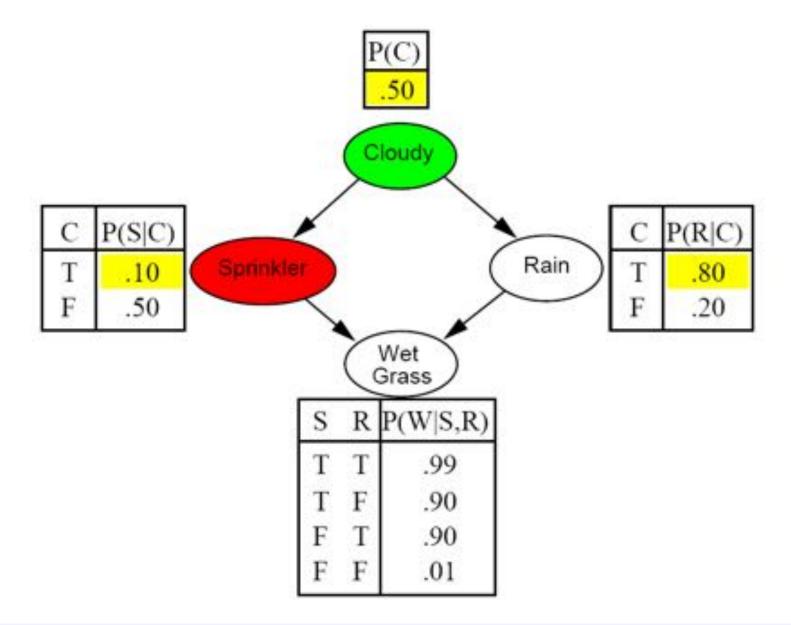
- Generating samples from a network that has no evidence associated with it (empty network)
- Basic idea
 - sample a value for each variable in topological order
 - using the specified conditional probabilities

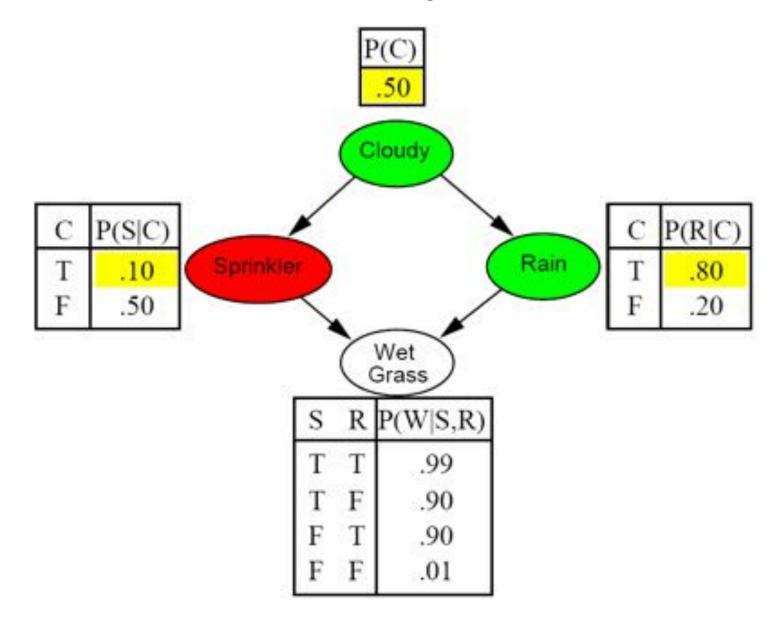
```
function PRIOR-SAMPLE(bn) returns an event sampled from bn inputs: bn, a belief network specifying joint distribution \mathbf{P}(X_1,\ldots,X_n) \mathbf{x} \leftarrow an event with n elements for i=1 to n do x_i \leftarrow a random sample from \mathbf{P}(X_i \mid parents(X_i)) given the values of Parents(X_i) in \mathbf{x} return \mathbf{x}
```

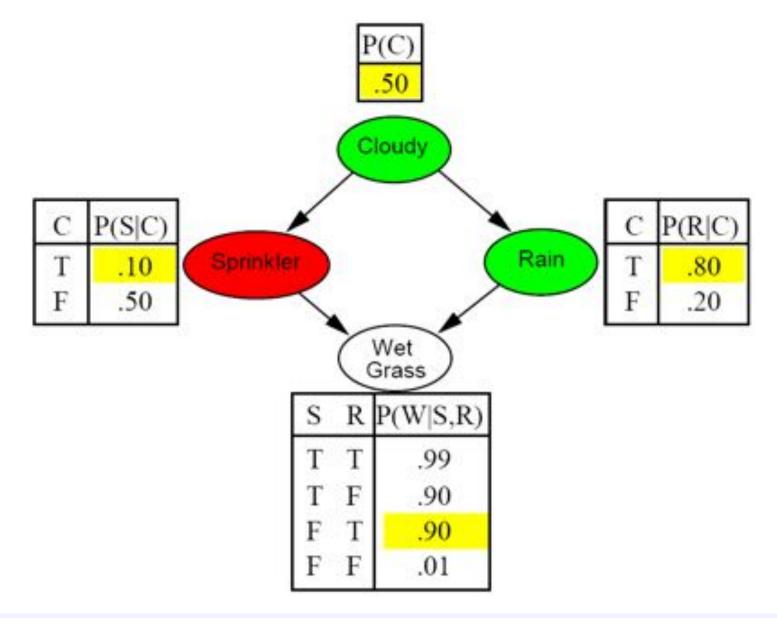


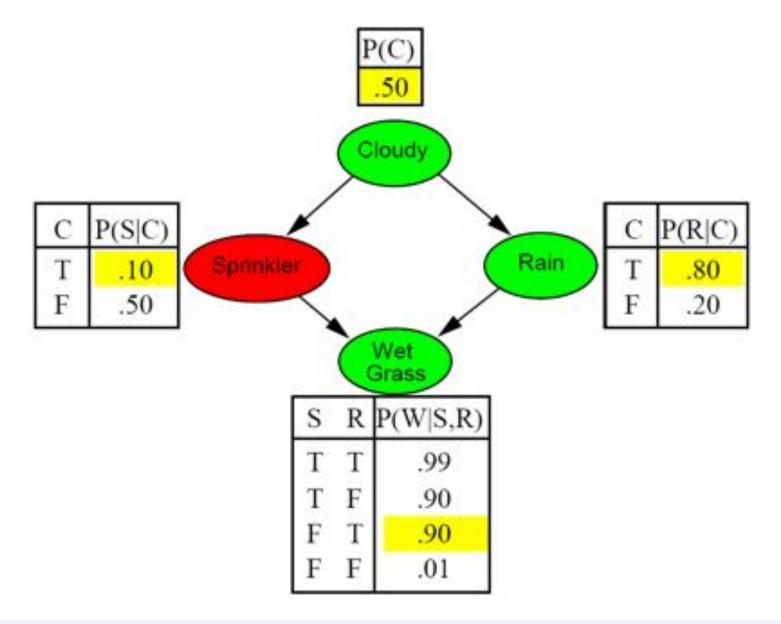












Probability Estimation using Sampling

- sample many points using the above algorithm
- count how often each possible combination $x_1, x_2, ..., x_n$ appears
 - increment counters $N_{PS}(x_1...x_n)$
- estimate the probability by the observed percentages

$$\hat{P}_{PS}(x_1...x_n) = N_{PS}(x_1...x_n)/N$$

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$$\hat{P}_{PS}(x_1...x_n) = N_{PS}(x_1...x_n)/N$$

Does this converge towards the joint probability function?

Convergence of Sampling from an Empty Network

Probability that PRIORSAMPLE generates a particular event

$$S_{PS}(x_1...x_n) = \prod_{i=1}^{n} P(x_i|parents(X_i)) = P(x_1...x_n)$$

i.e., the true prior probability

E.g.,
$$S_{PS}(t, f, t, t) = 0.5 \times 0.9 \times 0.8 \times 0.9 = 0.324 = P(t, f, t, t)$$

Let $N_{PS}(x_1 \dots x_n)$ be the number of samples generated for event x_1, \dots, x_n

Then we have

$$\lim_{N\to\infty} \hat{P}(x_1, \dots, x_n) = \lim_{N\to\infty} N_{PS}(x_1, \dots, x_n)/N$$

$$= S_{PS}(x_1, \dots, x_n)$$

$$= P(x_1, \dots, x_n)$$

That is, estimates derived from PRIORSAMPLE are consistent

Shorthand: $\hat{P}(x_1, \dots, x_n) \approx P(x_1 \dots x_n)$

Rejection Sampling

 $\hat{\mathbf{P}}(X|\mathbf{e})$ estimated from samples agreeing with \mathbf{e}

```
function REJECTION-SAMPLING(X, e, bn, N) returns an estimate of P(X|e) local variables: N, a vector of counts over X, initially zero for j=1 to N do x \leftarrow \text{PRIOR-SAMPLE}(bn) if x is consistent with e then N[x] \leftarrow N[x] + 1 \text{ where } x \text{ is the value of } X \text{ in } x return \text{NORMALIZE}(N[X])
```

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```

```
E.g., estimate P(Rain|Sprinkler = true) using 100 samples 27 samples have Sprinkler = true Of these, 8 have Rain = true and 19 have Rain = false.
```

```
\hat{\mathbf{P}}(Rain|Sprinkler = true) = Normalize((8, 19)) = (0.296, 0.704)
```

Similar to a basic real-world empirical estimation procedure

Analysis of Rejection Sampling

- Rejection sampling generates random samples from an empty network
 - and discards all samples that are inconsistent with the evidence

```
\hat{\mathbf{P}}(X|\mathbf{e}) = \alpha \mathbf{N}_{PS}(X,\mathbf{e}) (algorithm defn.)

= \mathbf{N}_{PS}(X,\mathbf{e})/N_{PS}(\mathbf{e}) (normalized by N_{PS}(\mathbf{e}))

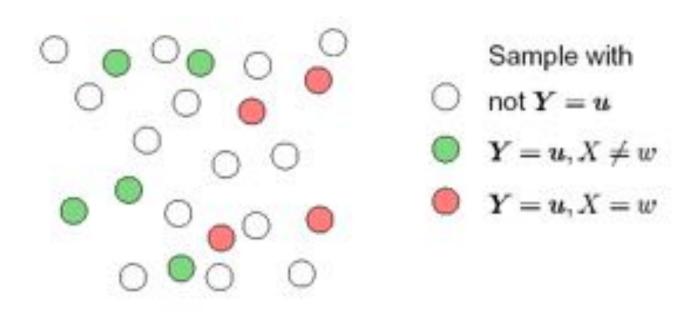
\approx \mathbf{P}(X,\mathbf{e})/P(\mathbf{e}) (property of PRIORSAMPLE)

= \mathbf{P}(X|\mathbf{e}) (defn. of conditional probability)
```

Hence rejection sampling returns consistent posterior estimates

Rejection Sampling: Illustration

Let Y be a subset of evidence nodes s.t. Y=u



Approximation for
$$P^{X}(X = w \mid Y = u)$$
: $\frac{\# \bigcirc}{\# \bigcirc \bigcirc}$

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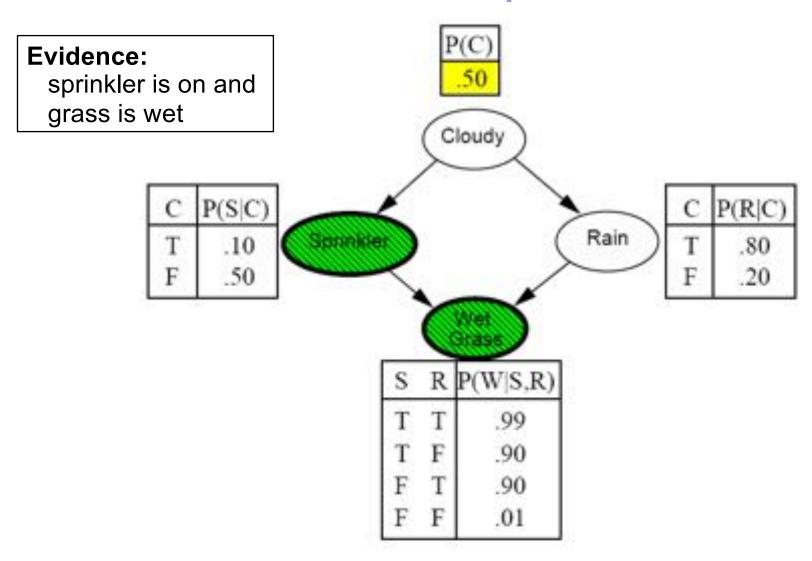
Problem

- many unnecessary samples will be generated if the probability of observing the evidence e is small
- P(e) will decrease exponentially with increasing numbers of evidence variables!

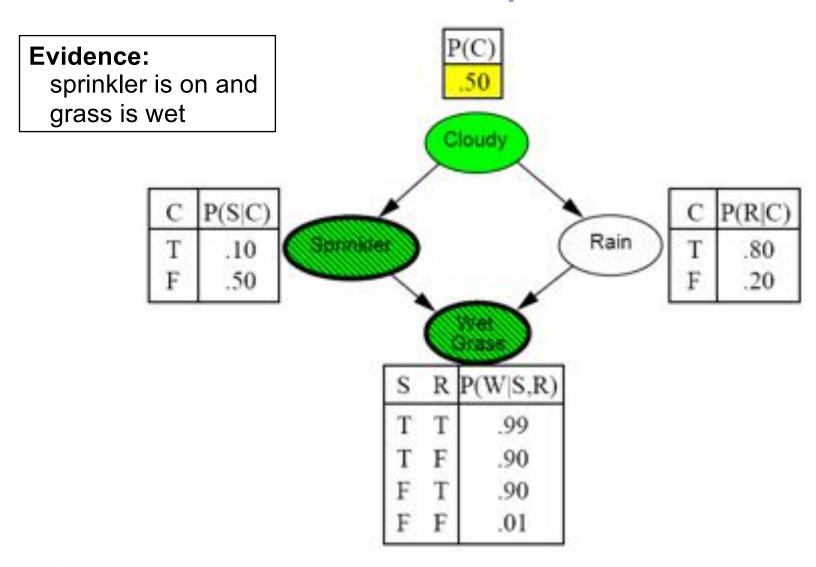
Likelihood Weighting

Idea: fix evidence variables, sample only nonevidence variables, and weight each sample by the likelihood it accords the evidence

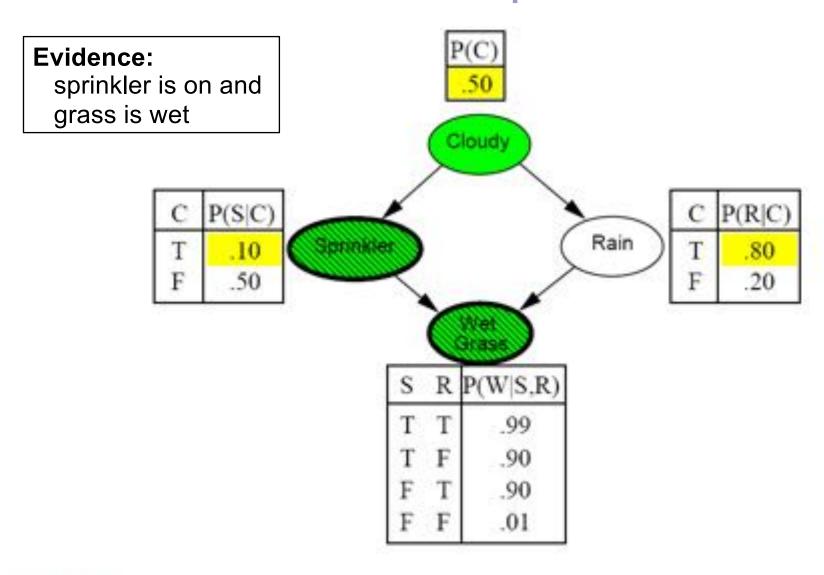
```
function Likelihood-Weighting (X, e, bn, N) returns an estimate of P(X|e)
   local variables: W, a vector of weighted counts over X, initially zero
   for j = 1 to N do
        x, w \leftarrow \text{Weighted-Sample}(bn)
        \mathbf{W}[x] \leftarrow \mathbf{W}[x] + w where x is the value of X in \mathbf{x}
   return Normalize(W[X])
function WEIGHTED-SAMPLE(bn, e) returns an event and a weight
   x \leftarrow an event with n elements; w \leftarrow 1
   for i = 1 to n do
        if X_i has a value x_i in e
             then w \leftarrow w \times P(X_i = x_i \mid parents(X_i))
             else x_i \leftarrow a random sample from P(X_i \mid parents(X_i))
   return x, w
```



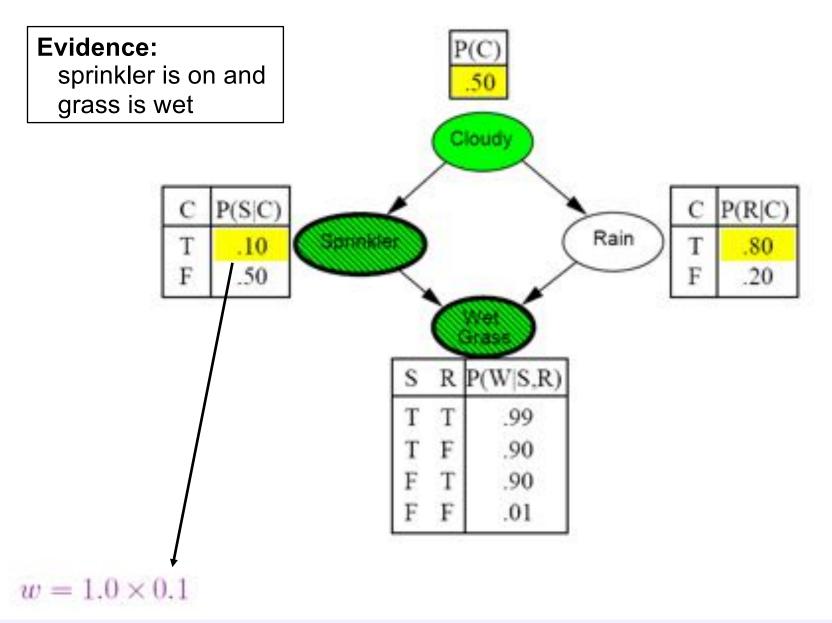
w = 1.0

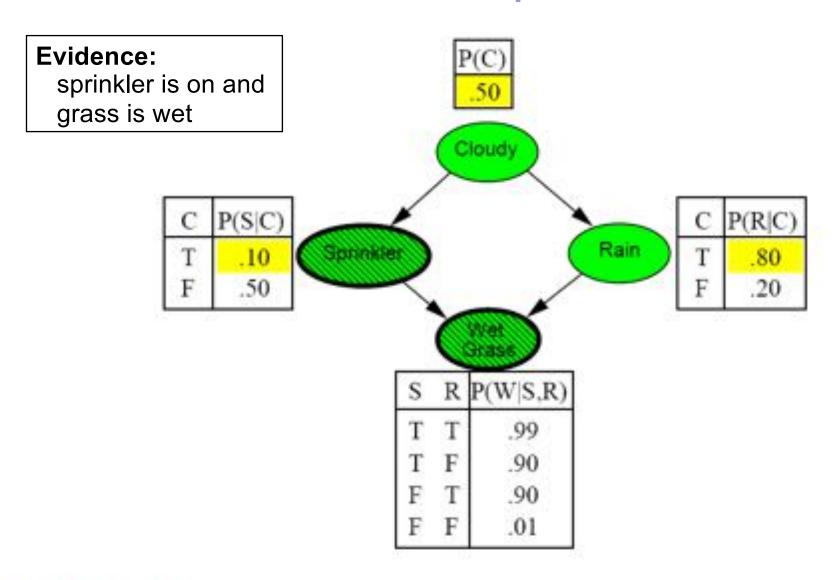


$$w = 1.0$$

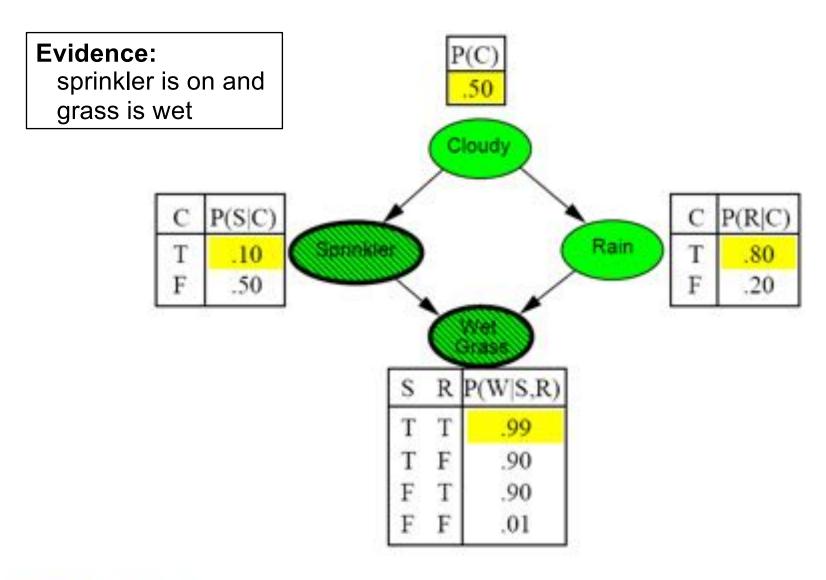


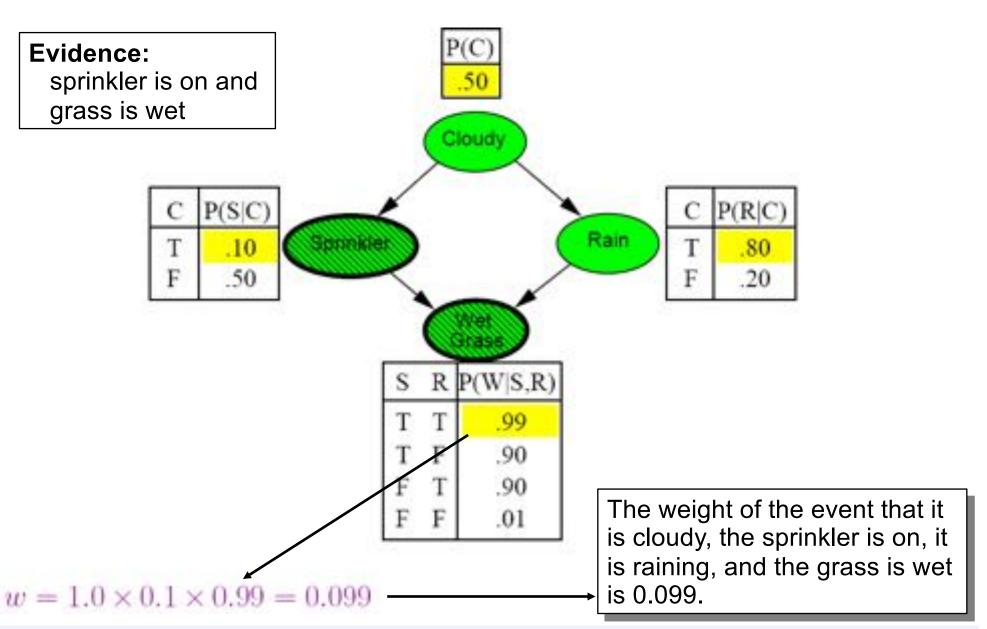
$$w = 1.0$$





 $w = 1.0 \times 0.1$





Analysis

Sampling probability for WEIGHTEDSAMPLE is

$$S_{WS}(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^{t} P(z_i | parents(Z_i))$$

Note: pays attention to evidence in ancestors only

⇒ somewhere "in between" prior and posterior distribution

Weight for a given sample z, e is

$$w(\mathbf{z}, \mathbf{e}) = \prod_{i=1}^{m} P(e_i | parents(E_i))$$

Weighted sampling probability is

$$\begin{split} S_{WS}(\mathbf{z}, \mathbf{e}) w(\mathbf{z}, \mathbf{e}) \\ &= \prod_{i=1}^l P(z_i | parents(Z_i)) \ \prod_{i=1}^m P(e_i | parents(E_i)) \\ &= P(\mathbf{z}, \mathbf{e}) \ \text{(by standard global semantics of network)} \end{split}$$

Hence likelihood weighting returns consistent estimates but performance still degrades with many evidence variables because a few samples have nearly all the total weight

Markov Chain Monte Carlo (MCMC) Sampling

"State" of network = current assignment to all variables.

Generate next state by sampling one variable given Markov blanket Sample each variable in turn, keeping evidence fixed

```
function MCMC-Ask(X, e, bn, N) returns an estimate of P(X|e) local variables: N[X], a vector of counts over X, initially zero Z, the nonevidence variables in bn X, the current state of the network, initially copied from E initialize E with random values for the variables in E for E for E in E do sample the value of E in E from E for E for E in E do sample the value of E in E from E for E for
```

Can also choose a variable to sample at random each time

Ordered Gibbs Sampler

Generate sample x^{t+1} from x^t:

Process all variables in some order

$$X_{1} = x_{1}^{t+1} \leftarrow P(x_{1} \mid x_{2}^{t}, x_{3}^{t}, ..., x_{N}^{t}, e)$$

$$X_{2} = x_{2}^{t+1} \leftarrow P(x_{2} \mid x_{1}^{t+1}, x_{3}^{t}, ..., x_{N}^{t}, e)$$

$$...$$

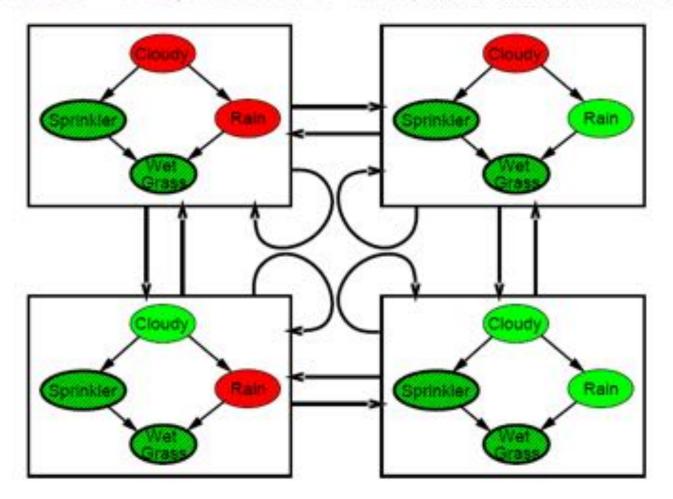
$$X_{N} = x_{N}^{t+1} \leftarrow P(x_{N} \mid x_{1}^{t+1}, x_{2}^{t+1}, ..., x_{N-1}^{t+1}, e)$$

In short, for i=1 to N:

$$X_i = x_i^{t+1} \leftarrow$$
sampled from $P(x_i \mid x^t \setminus x_i, e)$

The Markov Chain

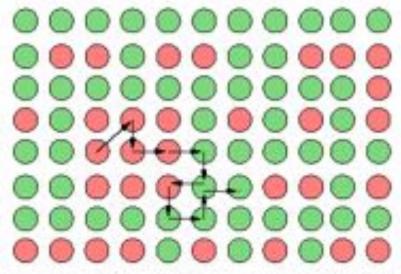
With Sprinkler = true, WetGrass = true, there are four states:



Wander about for a while, average what you see

Gibbs Sampling: Illustration

The process of Gibbs sampling can be understood as a random walk in the space of all instantiations with Y = u:



Reachable in one step: instantiations that differ from current one by value assignment to at most one variable (assume randomized choice of variable X_k).

Guaranteed to converge iff chain is:

irreducible (every state reachable from every other state)
aperiodic (returns to state i can occur at irregular times)
ergodic (returns to every state with probability 1)

Estimate P(Rain|Sprinkler = true, WetGrass = true)

Sample Cloudy or Rain given its Markov blanket, repeat. Count number of times Rain is true and false in the samples.

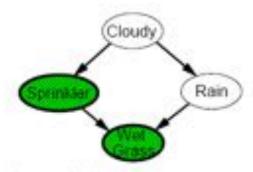
E.g., visit 100 states 31 have Rain = true, 69 have Rain = false

 $\hat{\mathbf{P}}(Rain|Sprinkler = true, WetGrass = true)$ = Normalize($\langle 31, 69 \rangle$) = $\langle 0.31, 0.69 \rangle$

Theorem: chain approaches stationary distribution: long-run fraction of time spent in each state is exactly proportional to its posterior probability

Markov Blanket Sampling

Markov blanket of Cloudy is Sprinkler and Rain Markov blanket of Rain is Cloudy, Sprinkler, and WetGrass



Probability given the Markov blanket is calculated as follows:

$$P(x_i'|mb(X_i)) = P(x_i'|parents(X_i)) \prod_{Z_j \in Children(X_i)} P(z_j|parents(Z_j))$$

What have we learned?

- Exact inference via Variable Elimination (VE)
- Inference in Bayesian networks is NP-hard, even when approximating. Still, for many distributions, sampling is the only option
- Forward sampling
- Rejections sample
- MCMC sampling (GIBBS sampling)
- Overall, we now know:
 - Basics of probability theory
 - Arguments why to follow probability theory
 - Bayesian networks (representation and semantics)
 - Inference in Bayesian networks