Probability

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Abstract

This note collects salient facts about probability theory.

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Probability is an extension of logic. Instead of propositions being either true or false a degree of belief can be specified for events occurring. All probabilities are based on models of available information and different people may have different degrees of belief. However, as the amount of information increases, probabilities should converge.

Probability Model

A probability model specifies a sample space and a probability measure. The sample space is just a set of what can possibly happen: heads or tails as the outcome of a coin toss, the integers from 1 to 6 as the outcomes of rolling a single die, the set of all sequences of not more than 280 characters as a model of possible Twitter tweets.

Assuming the characters are upper and lower case letters, space, and 3 punctuation marks then there are 30^280 possible messages. This is approximately 10^1374 . The number of elementary particles in the universe has been estimated to be 10^80 . The world population is a bit under 8 billion. Assuming everyone posts a Trumpian 10 tweets a day and uses all of their 280 character allotment, that comes to $8\times10^9\times10\times280=2.24\times10^44$. The universe is 14 billion years. That means...

People seem to be surprised probabilities are modeled using sets. Sets have no structure, they are just a bag of things (*elements*).

An event is a subset of a sample space. A probability measure assigns a number between 0 and 1 to events. If Ω is a sample space and P is a probability measure then the measure of the union of sets is the sum of the measure of each set minus the measure of the intersection: $P(E \cup F) = P(E) + P(F) - P(E \cap F)$ for events E and F. This is the mathematical way to say measures do not double count.

A probability measure must also satisfy $P(\emptyset) = 0$ and $P(\Omega) = 1$.

Exercise. If Q is a measure with $Q(\emptyset) = a$ and $Q(\Omega) = b$, show (Q - a)/(b - a) is a probability measure.

Let $1_A(\omega) = 1$ if $\omega \in A$ and 0 = 0 if $\omega \notin A$. If $X = \sum a_i 1_{A_i}$ where $a_i \in \mathbf{R}$ and A_i are events, Define the *expected value* of X by $EX = \sum_i a_i P(A_i)$.

Exercise. Show that if $\sum_i a_i 1_{A_i} = 0$ then $\sum_i a_i P(A_i) = 0$.

Hint: Replace the A_i by disjoint B_j so $b_j = 0$ for all j.

This shows expected value is well-defined.

Exercise. Show $P(\cup_i A_i) = \sum_i P(A_i) - \sum_{i < j} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) + \sum_{i < j < k} P(A_i \cap A_j) +$

Hint: Use $(1_A - 1_{A_1}) \cdots (1_A - 1_{A_n}) = 0$, where $A = \bigcup_{k=1}^n A_k$.

Algebras of Sets

An algebra of sets on Ω is a collection of subsets (events), \mathcal{A} , that is closed under complement and union. We also assume the empty set belongs to \mathcal{A} . By De Morgan's Laws an algebra is also closed under intersection and Ω belongs to \mathcal{A} . The power set of Ω , $2^{\Omega} = \{E \subseteq \Omega\}$, clearly satisfies these conditions.

An *atom* of an algebra is a member, A, of the algebra such that if $B \subseteq A$ and B is in the algebra, then either B = A or B is the empty set.

Partition

A partition of a set is a collection of pairwise disjoint subsets who's union is equal to the set.

Exercise. If an algebra on Ω is finite its atoms form a partition of Ω .

Hint: Show $A_{\omega} = \cap \{B \in \mathcal{A} : \omega \in B\}, \omega \in \Omega$, is an atom

This shows there is a one-to-one correspondence between finite partitions and finite algebras of sets. A partition is the mathematical way of specifying partial information. Knowing the outcome, $\omega \in \Omega$, corresponds to complete knowledge. Knowing which atom the outcome belongs to corresponds to partial knowledge. For example, the partition $\{\{1,3,5\},\{2,4,6\}\}$ corresponds to knowing whether the roll of a die is odd or even.

The coarsest partition $\{\Omega\}$ corresponds to no knowledge while the finest partition $\{\{\omega\}:\omega\in\Omega\}$ corresponds to complete knowledge.

Measurable

A function $X: \Omega \to \mathbf{R}$ is \mathcal{A} -measureable if the sets $X^{-1}((-\infty, x]) = \{\omega \in \Omega : X(\omega) \le x\}$ belong to \mathcal{A} for $x \in \mathbf{R}$.

Exercise: If \mathcal{A} is finite, show that a function is measurable if and only if it is constant on atoms of \mathcal{A} .

In this case $X: \mathcal{A} \to \mathbf{R}$ is indeed a function on the atoms.

Random Variable

A random variable is a variable, a symbol that can be used in place of a number, with additional information: the probability of the values it can take on. The cumulative distribution function is $F(x) = F^X(x) = P(X \le x)$. It tells you everything there is to know about X. For example, $P(a < X \le b) = F(b) - F(a)$.

Exercise. Show $P(a \le X \le b) = \lim_{x \uparrow a} F(b) - F(x)$.

Hint: $[a, b] = \bigcap_n (a - 1/n, b]$.

In general, $P(X \in A) = E1_A = \int 1_A(x) dF(x)$ for sufficiently nice $A \subset \mathbf{R}$.

Since $(-\infty, x] \subseteq (-\infty, x']$ if $x \le x'$, F is non-decreasing: $F(x) \le F(x')$. $\lim_{x \to -\infty} F(x) = 0 \lim_{x \to \infty} F(x) = 1$. F is right continuous with left limits.

Examples

Discrete

A discrete random variable is defined by $x_i \in \mathbf{R}$ and $p_i > 0$ with $\sum p_i = 1$. The probability the random variable takes on value x_i is p_i .

Bernoulli, Binomial

Uniform

The random variable, U, that is uniformly distributed on the unit interval, [0,1], has cdf F(x) = x if $0 \le x \le 1$, = 0 if x < 0, and = 1 if x > 1.

Normal

Two random variables, X and Y, have the same law if they have the same cdf.

Exercise. If X has cdf F, then X and $F^{-1}(U)$ have the same law.

Exercise. If X has cdf F, then F(X) is uniformly distributed on the unit interval.

This shows a uniformly distributed random variable has sufficient randomness to generate any random variable. There are no random, random variables.

The mathematician's definition of a random variable is that it is a measurable function $X \colon \Omega \to \mathbf{R}$. Its cumulative distribution function is $F(x) = P(X \le x) = P(\{\omega \in \Omega \mid X(\omega) \le x\})$. Given a cdf F we can define $X \colon \mathbf{R} \to \mathbf{R}$ to be the identity function and let P be the probability measure defined by F: $P(A) = \int 1_A(x) dF(x)$.

Expected Value

The expected value of a random variable is defined by the $EX = \int_{-\infty}^{\infty} x \, dF(x)$. The expected value of any function of a random variable is $Ef(X) = \int_{-\infty}^{\infty} f(x) \, dF(x)$.

Moments

The moments of a random variable, X, are $m_n = E[X^n]$, $n = 0, 1, 2, \ldots$ They don't necessarily exist for all n, except for n = 0. They also cannot be an arbitrary sequence of values.

Suppose all moments of X exist, then for any complex numbers, (c_i) , $0 \le E|\sum_i c_i X^i|^2 = E\sum_{j,k} c_j \bar{c_k} X^{j+k} = \sum_{j,k} c_j \bar{c_k} m_{j+k}$. This says the Hankel matrix, $M = [m_{j+k}]_{j,k}$, is positive definite. The converse is also true: if the Hankel matrix is positive definite there exists a random variable with the corresponding moments. This is not a trivial result and the random variable might not be unique.

Spectral measure . . .

Cumulants

The *cumulant* of a random variable, X, is $\kappa(s) = \kappa^X(s) = \log E \exp(sX)$. The *cumulants*, κ_n , are defined by $\kappa(s) = \sum_{n>0} \kappa_n s^n/n!$.

It is easy to see $\kappa_1 = EX$ and $\kappa_2 = \text{Var } X$. The third and fourth cumulants are related to skew and kurtosis. We will see the exact relationship below.

If c is a constant then $\kappa^{cX}(s) = \kappa^X(cs)$ so $\kappa^{cX}_n = c^n \kappa^X_n$. If X and Y are independent then $\kappa^{X+Y}(s) = \kappa^X(s) + \kappa^Y(s)$ so $\kappa^{X+Y}_n = \kappa^X_n + \kappa^Y_n$ \$

Characteristic Function

The characteristic function of a random variable, X, is $\xi(t) = \kappa(it)$.

Fourier Transform

The Fourier transform is $\psi(t) = \xi(-t) = \kappa(-it)$. These can be used to prove the central limit theorem: if X_j are independent, identically distributed random variables with mean zero and variance one, then $(X_1 + \cdots + X_n)/sqrtn$ converges to a standard normal random variable.

Examples

If X is normal then $E \exp(X) = \exp(EX + \operatorname{Var}(X)/2)$ so $\kappa_1 = EX$, $\kappa_2 = \operatorname{Var}(X)$, and $\kappa_n = 0$ for n > 2.

If X is Poisson with parameter λ then

$$Ee^{sX} = \sum_{k=0}^{\infty} e^{sk} e^{-\lambda} \lambda^k / k!$$
$$= \sum_{k=0}^{\infty} (e^s \lambda)^k e^{-\lambda} / k!$$
$$= \exp(\lambda (e^s - 1))$$

so $\kappa(s) = \lambda(e^s - 1)$ and $\kappa_n = \lambda$ for all n.

Bell Polynomials

The relationship between moments and cumulants is given by Bell polynomials.

In particular $m_1 = \kappa_1$ and $m_2 = \kappa_1^2 + \kappa_2$ so κ_1 is the mean and κ_2 is the variance. If the mean is 0 and the variance is 1, then κ_3 is the skew and κ is the excess kurtosis.

Joint Distribution

Two random variables, X and Y, are defined by their joint distribution, $F(x,y) = F^{X,Y}(x,y) = P(X \le x, Y \le y)$. For example (X,Y) is in the square $(a,b] \times (c,d]$ with probability $P(a < X \le b, c < Y \le d) = P(X \le b, Y \le d) - P(X \le a) - P(Y < c) + P(X < a, Y < c)$.

Copulas

A copula is the joint distribution of random variables uniformly distributed on the unit interval. Let U and V be two uniformly distributed random variables.

If V = U then their joint distribution is $C(u, v) = P(U \le u, V \le v) = P(U \le u, U \le v) = P(U \le \min\{u, v\}) = \min\{u, v\} = M(u, v)$.

If V=1-U then their joint distribution is $C(u,v)=P(U\leq u,V\leq v)=P(U\leq u,1-U\leq v)=P(1-v\leq U\leq u)=\max\{u-(1-v),0\}=\max\{u+v-1,0\}=W(u,v)$

For every copula, $W \leq C \leq M$.

Let X and Y be random variables with cdfs F and G respectively, and joint distribution H. Define the cumulant, $C = C^{X,Y}$, to be the joint distribution of F(X) and G(Y).

Normal

Poisson

Infinitely Divisible

A random variable, X, is *infinitely divisible* if for any positive integer, n, there exist independent, identically distributed random variables X_1, \ldots, X_n such that $X_1 + \cdots + X_n$ has the same law as X.

Characteristic function ...

Conditional Expectation

The conditional expectation of an event B given an event A is $P(B|A) = P(B \cap A)/P(A)$. In some sense, this reduces the sample space to A. In particular, P(A|A) = 1. Since $P(A|B) = P(A \cap B)/P(B)$ we have P(A|B) = P(B|A)P(A)/P(B). This is the simplest form of Bayes Theorem. It shows how to update your degree of belief based on new information. Every probability is conditional on given information.

Define $E[X|A] = E[X1_A]/P(A)$ for any random variable X. If $X = 1_B$ then this coincides with the definition of conditional expectation above.

If we write this as $E[X|A]P(A) = E[X1_A]$ then defining E[X|A] by $E[X|A]P|_{\mathcal{A}} = (XP)_{\mathcal{A}}$ agrees on atoms of \mathcal{A} .

moments, Hamburger moment problem.

cumulants, Bell polynomials

Normal

Poisson

Infinitely Divisible

Stochastic Processes

A stochastic process is . . .

Brownian Motion

reflection

L'evy Processes

Remarks

Cheval de Mere

Pascal

Bernoulli(s)

Kolmogorov

Willy Feller