

PS1

June 28, 2024

1. (a) *Proof.*

$$\begin{aligned}\sum_{i=1}^n y_i &= \sum_{i=1}^n (x_{i1} + 1 - x_{i1}) \times y_i \\ &= \sum_{i=1}^n x_{i1} \times y_i + \sum_{i=1}^n (1 - x_{i1}) \times y_i\end{aligned}$$

□

(b)

$$\begin{aligned}\hat{\beta}_1 &= \frac{\sum_{i=1}^n (x_{i1} - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_{i1} - \bar{x})^2} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i - \bar{x} \sum_{i=1}^n y_i - \bar{y} \sum_{i=1}^n x_{1i} + \sum_{i=1}^n \bar{x} \bar{y}}{\sum_{i=1}^n x_{1i} - 2\bar{x} \sum_{i=1}^n x_{1i} + \bar{x} \sum_{i=1}^n x_{1i}} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i - \bar{x} \sum_{i=1}^n y_i - \bar{y} \sum_{i=1}^n x_{1i} + \bar{x} \sum_{i=1}^n y_i}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i - \bar{y} \sum_{i=1}^n x_{1i}}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} - \frac{\bar{y}}{1 - \bar{x}} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} - \frac{\sum_{i=1}^n y_i}{\sum_{i=1}^n (1 - x_{i1})} \\ &= \frac{\sum_{i=1}^n x_{i1} y_i}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} - \frac{\sum_{i=1}^n x_{i1} y_i}{\sum_{i=1}^n (1 - x_{i1})} - \frac{\sum_{i=1}^n (1 - x_{i1}) y_i}{\sum_{i=1}^n (1 - x_{i1})} \\ &= \left(\sum_{i=1}^n x_{i1} y_i \right) \left(\frac{1}{(1 - \bar{x}) \sum_{i=1}^n x_{1i}} - \frac{1}{\sum_{i=1}^n (1 - x_{i1})} \right) - \frac{\sum_{i=1}^n (1 - x_{i1}) y_i}{n_0} \\ &= \left(\sum_{i=1}^n x_{i1} y_i \right) \frac{n - \sum_{i=1}^n x_{1i}}{\sum_{i=1}^n (1 - x_{1i}) \sum_{i=1}^n x_{1i}} - \bar{y}_0 \\ &= \frac{\sum_{i=1}^n x_{i1} y_i}{\sum_{i=1}^n x_{1i}} - \bar{y}_0 \\ &= \bar{y}_1 - \bar{y}_0 \\ \hat{\beta}_0 &= \bar{y} - \hat{\beta}_1 \bar{x} \\ &= \frac{\sum_{i=1}^n y_i}{n} - (\bar{y}_1 - \bar{y}_0) \frac{\sum_{i=1}^n x_{1i}}{n} \\ &= \bar{y}_0 \frac{\sum_{i=1}^n x_{1i}}{n} + \frac{\sum_{i=1}^n y_i}{n} - \frac{\bar{y}_1 \sum_{i=1}^n x_{1i}}{n} \\ &= \bar{y}_0 \frac{\sum_{i=1}^n x_{1i}}{n} + \frac{\sum_{i=1}^n y_i}{n} - \frac{\bar{y}_1 \times n_1}{n} \\ &= \bar{y}_0 \frac{\sum_{i=1}^n x_{1i}}{n} + \frac{\sum_{i=1}^n (1 - x_{i1}) y_i}{n} \\ &= \bar{y}_0 \frac{\sum_{i=1}^n x_{1i}}{n} + \bar{y}_0 \frac{\sum_{i=1}^n (1 - x_{1i})}{n} \\ &= \bar{y}_0\end{aligned}$$

(c) $\hat{\beta}_1$ represents the difference between y on $x = 1$ and y on $x = 0$, and $\hat{\beta}_0$ represents the average of y on $x = 0$.

(d) *Proof.*

$$\begin{aligned}
 \mathbb{E}[\hat{\beta}_1] &= \beta_1 \\
 &= \frac{\mathbb{E}[y] - \beta_0}{\mathbb{E}[x_1]} \\
 &= \frac{\mathbb{E}[y|x_1 = 1]P(x_1 = 1) + \mathbb{E}[y|x_1 = 0]P(x_1 = 0) - \mathbb{E}[y|x_1 = 0]}{P(x_1 = 0) \times 0 + P(x_1 = 1) \times 1} \\
 &= \frac{\mathbb{E}[y|x_1 = 1]P(x_1 = 1) - \mathbb{E}[y|x_1 = 0]P(x_1 = 1)}{P(x_1 = 1)} \\
 &= \mathbb{E}[y|x_1 = 1] - \mathbb{E}[y|x_1 = 0]
 \end{aligned}$$

□

2. (a) There's 3 possible explanations for the correlation between smoking and baby's weight.
 - i. The more the mother smoked, the less the baby's weight.
 - ii. The less the baby's weight, the more the mother smoked.
 - iii. There's a third variable that affects both the mother's smoking and the baby's weight in the contrary way.
- (b) Because the more the mother smoked, the less the baby's weight, and the effect of smoking on baby's weight is quite big.
- (c) Yes, it's possible according to the common sense.

3. (a)

```
. use "EduIncome_24.dta"
```

```
. summarize birthyear wage schooling_yr
```

Variable	Obs	Mean	Std. Dev.	Min	Max
birthyear	2,429	1974.794	10.584	1955	1990
wage	2,429	58122.36	41021.76	2113.569	608707.9
schooling_yr	2,429	7.656237	2.927878	0	15

(b)

```
. gen female = (gender == 2)
```

```
. summarize female
```

Variable	Obs	Mean	Std. Dev.	Min	Max
female	2,429	.2984767	.457684	0	1

Female in the sample accounts for 29.85%.

(c)

```
. summarize wage if female == 1
```

Variable	Obs	Mean	Std. Dev.	Min	Max
wage	725	47644.76	32694.44	2242.335	355079.6

```
. summarize wage if female == 0
```

Variable	Obs	Mean	Std. Dev.	Min	Max
wage	1,704	62580.26	43337.33	2113.569	608707.9

```
. regress wage female
```

Source	SS	df	MS	Number of obs	=	2,429
Model	1.1345e+11	1	1.1345e+11	F(1, 2427)	=	69.32
Residual	3.9723e+12	2,427	1.6367e+09	Prob > F	=	0.0000
				R-squared	=	0.0278

-----					Adj R-squared	=	0.0274
Total		4.0858e+12	2,428	1.6828e+09	Root MSE	=	40457

wage		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	

female		-14935.5	1793.902	-8.33	0.000	-18453.24	-11417.77
_cons		62580.26	980.063	63.85	0.000	60658.41	64502.11

So, we have $\text{wage} = 62580.26 - 14935.5 \times \text{female}$.

Using my conclusion in Q1, this regression result means that the wage difference between female and male is -14935.5 yuan, and the average wage for male is 62580.26 yuan.

(d) `. gen age = 2023 - birthyear`

```
. summarize age
```

Variable	Obs	Mean	Std. Dev.	Min	Max
age	2,429	48.20585	10.584	33	68

```
. gen ln_wage = ln(wage)
```

```
. regress ln_wage age schooling_yr
```

Source	SS	df	MS	Number of obs	=	2,429
				F(2, 2426)	=	33.05
Model	37.4492937	2	18.7246468	Prob > F	=	0.0000
Residual	1374.39397	2,426	.566526782	R-squared	=	0.0265
				Adj R-squared	=	0.0257
Total	1411.84327	2,428	.581484047	Root MSE	=	.75268

	ln_wage	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
	age	.0027053	.001456	1.86	0.063	-.0001498	.0055605
	schooling_yr	.0425877	.0052633	8.09	0.000	.0322666	.0529088
	_cons	10.26689	.0867829	118.31	0.000	10.09671	10.43707

If `schooling_yr` increases by 1, the `log(wage)` will increase by 0.043.

```
(e) . predict ln_wage_hat, xb
```

```
. gen residual = ln_wage - ln_wage_hat
```

```
. summarize ln_wage_hat residual
```

Variable	Obs	Mean	Std. Dev.	Min	Max
ln_wage_hat	2,429	10.72336	.1241931	10.35617	11.07885
residual	2,429	3.34e-09	.7523697	-3.072877	2.525533

The residual is very small, which suggests that the model fits the data well.

(f) In (d) , we have $R^2 = 0.0265$, indicating that there may be other factors not included in the model.