Model-based matching for causal inference in observational studies

Kellie Ottoboni with Philip B. Stark

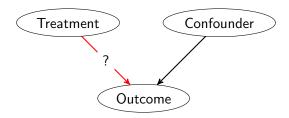
Department of Statistics, UC Berkeley Berkeley Institute for Data Science

March 15, 2016

- 1 Introduction
- 2 Matching
 - Propensity score matching
 - Model-based Matching
- 3 Examples
 - Toads and Packstock in Yosemite
 - Salt and Mortality
- 4 Conclusions
- 5 References

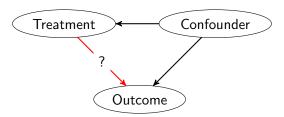
Observational Studies vs Experiments

- Problem: Estimate the causal effect of a treatment on outcome of interest
- In randomized experiments, treatment is assigned to individuals at random.
- In observational studies, the way individuals select into treatment groups is unknown.



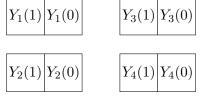
Observational Studies vs Experiments

- Problem: Estimate the causal effect of a treatment on outcome of interest
- In randomized experiments, treatment is assigned to individuals at random.
- In observational studies, the way individuals select into treatment groups is unknown.



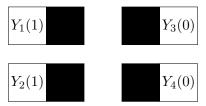
Neyman-Rubin Causal Model

- Population of $i=1,\ldots,N$ individuals. Each individual has two **potential outcomes**.
- $Y_i(1)$ is individual i's outcome if he receives treatment
- $Y_i(0)$ is individual i's outcome if he is in the control group
- The treatment effect for individual i is $\tau_i = Y_i(1) Y_i(0)$



Fundamental Problem of Causal Inference [Holland, 1986]

- We may never observe both $Y_i(1)$ and $Y_i(0)$
- T_i is a treatment indicator: 1 if i is treated, 0 if i is control
- The observed outcome for individual i is $Y_i = T_i Y_i(1) + (1 T_i) Y_i(0)$



Estimands

Average treatment effect

$$\mathbb{E}(Y_i(1) - Y_i(0))$$

Average treatment effect on the treated

$$\mathbb{E}(Y_i(1) - Y_i(0) \mid T_i = 1)$$

Conditional average treatment effect

$$\mathbb{E}(Y_i(1) - Y_i(0) \mid X_i)$$

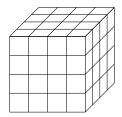
 If treatment effect varies by covariates X, then averages might not be informative

- 1 Introduction
- 2 Matching
 - Propensity score matching
 - Model-based Matching
- 3 Examples
 - Toads and Packstock in Yosemite
 - Salt and Mortality
- 4 Conclusions
- 5 References

Matching

How can we estimate the counterfactual for treated individuals?

- **Ideal:** group individuals by X_i to estimate subgroup treatment effects and then average over subgroups
- Reality: many covariates, perhaps continuous, make it difficult to stratify

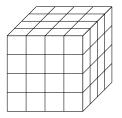


Aside... the curse of dimensionality

- If d covariates are split into k bins, we have d^k groups.
- To guarantee that we have at least one treated and one control in each group with 95% probability, we need

$$n \ge \frac{2\log(1 - (0.95)^{1/k^{d+1}})}{\log(\frac{k^d - 1}{k^d})}$$

- If d = 5 and k = 2, $n \ge 225$.
- If d = 10 and k = 2, $n \ge 10,844$.



Matching

• **Solution:** use a one-dimensional score to match or group individuals

 The propensity score is an individual's probability of being assigned treatment, conditional on their covariates

$$p(x) = \mathbb{P}(T = 1 \mid X = x)$$

- The propensity score is a balancing score: $X \perp \!\!\! \perp T \mid p(X)$
- For individuals with the same propensity score, treatment assignment is as if random

Theorem (Rosenbaum and Rubin [1983])

If treatment assignment is independent of potential outcomes given \boldsymbol{X} ,

$$(Y(1), Y(0)) \perp \!\!\! \perp T \mid X$$

and if every unit has a chance of receiving treatment,

$$0 < p(X) < 1$$
 for all X

then
$$(Y(1), Y(0)) \perp T \mid p(X)$$
.

In particular, treated units can serve as the counterfactual for controls with the same $p(\boldsymbol{X})$

$$\mathbb{E}(Y(t) \mid T = 1, p(X)) = \mathbb{E}(Y(t) \mid T = 0, p(X)) \text{ for } t = 0, 1$$

This result identifies the average treatment effect in terms of quantities we can estimate:

$$\begin{split} \mathbb{E}(Y(1) - Y(0)) &= \mathbb{E}_{p(x)} \left[\mathbb{E}(Y(1) - Y(0) \mid p(x)) \right] \\ &= \mathbb{E}_{p(x)} \left[\mathbb{E}(Y(1) \mid p(x)) - \mathbb{E}(Y(0) \mid p(x)) \right] \\ &= \mathbb{E}_{p(x)} \left[\mathbb{E}(Y \mid p(x), T = 1) - \mathbb{E}(Y \mid p(x), T = 0) \right] \end{split}$$

p(x) is usually unknown and estimated by $\hat{p}(x)$ using logistic or probit regressions

- Assumes a simple functional form for relationship between covariates and treatment
- Assumes that probability of treatment takes same form for all individuals
- May actually worsen balance if estimated incorrectly [Diamond and Sekhon, 2012]

Matching introduces bias

- Standard errors are difficult to compute for matching estimators [Abadie and Imbens, 2006, 2008]
- There's no "optimal" way to match

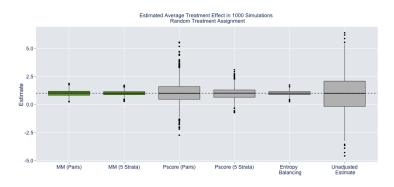
Model-based Matching

Idea: Instead of modeling the propensity score, model the outcome

Stratify on $\hat{Y},$ the "best" prediction of the outcome based on all covariates except for the treatment

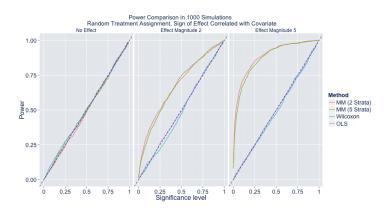
Model-based Matching

- Under standard assumptions (conditional independence of treatment and potential outcomes given X), the average treatment effect is nonparametrically identified
- Estimate it using the difference in average residuals, $Y-\hat{Y}$, between treated and controls



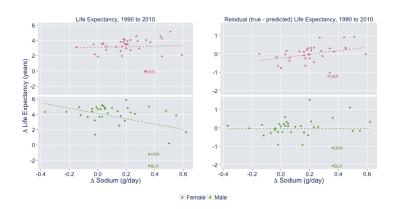
Model-based Matching

- Use stratified permutation test to test the strong null hypothesis of no treatment effect whatsoever
- Stratifying on \hat{Y} allows us to detect non-constant and non-linear treatment effects



- 1 Introduction
- 2 Matching
 - Propensity score matching
 - Model-based Matching
- 3 Examples
 - Toads and Packstock in Yosemite
 - Salt and Mortality
- 4 Conclusions
- 5 References

Salt



- 1 Introduction
- 2 Matching
 - Propensity score matching
 - Model-based Matching
- 3 Examples
 - Toads and Packstock in Yosemite
 - Salt and Mortality
- 4 Conclusions
- 5 References

Future Directions

- Do different test statistics give greater power when the treatment effect is nonlinear?
- What is the optimal way to stratify?
- How to quantify uncertainty standard errors and confidence intervals?

- 1 Introduction
- 2 Matching
 - Propensity score matching
 - Model-based Matching
- 3 Examples
 - Toads and Packstock in Yosemite
 - Salt and Mortality
- 4 Conclusions
- 5 References

References

Alberto Abadie and Guido W. Imbens. Large Sample Properties of Matching Estimators for Average Treatment Effects. *Econometrica*, 74(1):235–267, January 2006. ISSN 1468-0262.

doi: 10.1111/j.1468-0262.2006.00655.x. URL http://onlinelibrary.wiley.com/doi/10.1111/j.

1468-0262.2006.00655.x/abstract.

Alberto Abadie and Guido W. Imbens. On the Failure of the Bootstrap for Matching Estimators. *Econometrica*, 76(6): 1537–1557, November 2008. ISSN 1468-0262. doi: 10.3982/ECTA6474. URL http://onlinelibrary.wiley.com/doi/10.3982/ECTA6474/abstract.

Alexis Diamond and Jasjeet S. Sekhon. Genetic Matching for Estimating Causal Effects: A General Multivariate Matching Method for Achieving Balance in Observational Studies. *Review of Economics and Statistics*, 95(3):932–945, October 2012. ISSN 0034-6535. doi: 10.1162/REST_a_00318. URL http://dx.doi.org/10.1162/REST_a_00318.