

Data Science

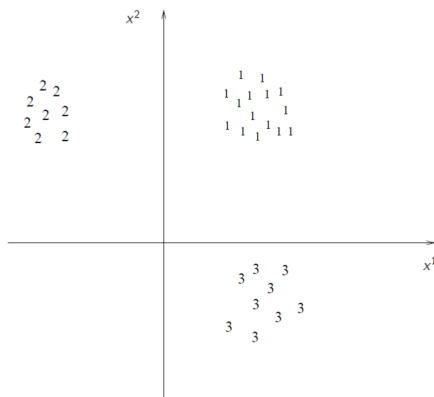
Static data analysis

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Master en Sciences Informatiques - Semestre d'Automne

Linear Discriminant Analysis (LDA)

We look at the issue of modeling multivariate data (p quantitative components and one categorical variable). Every data is described by (\mathbf{x}_i, y_i) , where $\mathbf{x}_i \in \mathbb{R}^p$ and $y_i = 1, \dots, q$.



Linear Discriminant Analysis

Class definition

Alternative interpretation : The categorical variable y_i describes the class to which data i belongs, characterised by variables \mathbf{x}_i .

Hence, point i belongs to class C_k iff $y_i = k$.

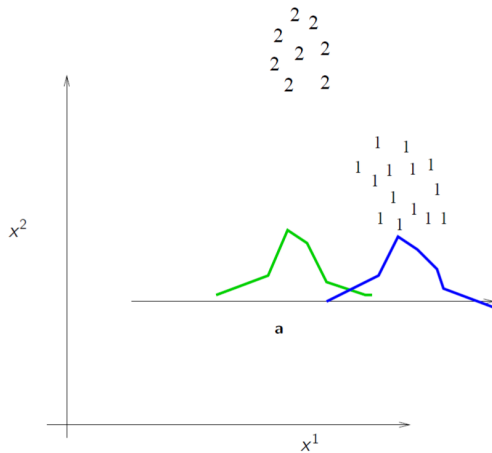
Discriminant Analysis

Can the q classes be discriminated over the space of variables \mathbf{x} ?

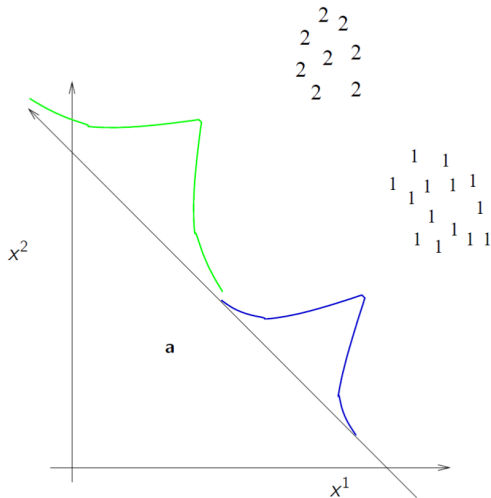
Is there a linear transform of \mathbf{x} such that the q classes are better separated?

⇒ basis for supervised learning

2 Classes - Discriminant Axis



2 Classes - Discriminant Axis



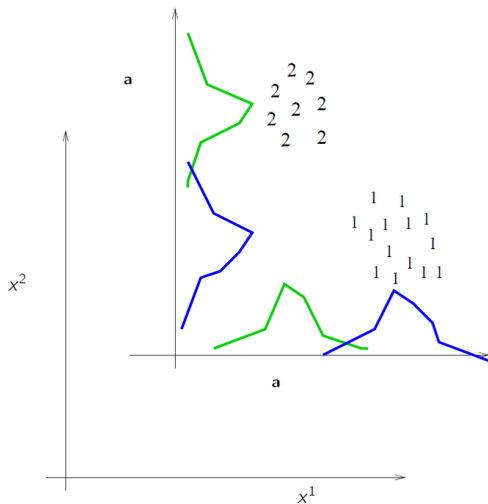
Inter-class discrimination criterion

- ★ Search for direction \mathbf{a} where the *inter-class* discrimination is maximum.
- ★ Clearly, \mathbf{a} must be parallel to the line $(\mathbf{g}_1, \mathbf{g}_2)$ across the centers of mass of the classes, since:

$$(\hat{\mathbf{g}}_1 - \hat{\mathbf{g}}_2)^2 = \left(\frac{\mathbf{a}^\top \mathbf{g}_1}{\|\mathbf{a}\|} - \frac{\mathbf{a}^\top \mathbf{g}_2}{\|\mathbf{a}\|} \right)^2 = \left(\frac{\mathbf{a}^\top}{\|\mathbf{a}\|} (\mathbf{g}_1 - \mathbf{g}_2) \right)^2$$

is maximum when $\mathbf{a} \propto \mathbf{g}_1 - \mathbf{g}_2$

Intra-class discrimination criterion



Intra-class discrimination criterion

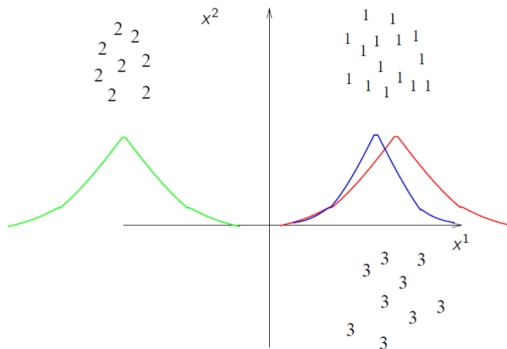
- ★ We must account for the *intra*-class variance of the projected data
- ★ Fisher criterion maximises

$$\max_{\mathbf{a}} \frac{(\hat{\mathbf{g}}_1 - \hat{\mathbf{g}}_2)^2}{\hat{\sigma}_1^2 + \hat{\sigma}_2^2}$$

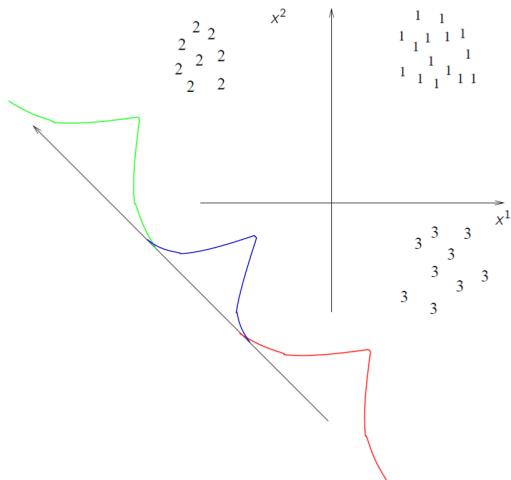
where $\hat{\sigma}_k$ is the normalised variance of the projection of class k

$$\hat{\sigma}_k^2 = \sum_{\mathbf{x} \in \mathbf{C}_k} (\hat{\mathbf{x}} - \hat{\mathbf{g}}_k)^T (\hat{\mathbf{x}} - \hat{\mathbf{g}}_k)$$

Discriminant axes



Discriminant axes



Generalisation: intra-class criteria

- ★ Let $A_k = [\mathbf{x}_1 - \mathbf{g}_k, \dots, \mathbf{x}_{n_k} - \mathbf{g}_k]$, $\mathbf{x}_i \in C_k$ be the matrix of centered data
- ★ $\frac{1}{n_k} A_k A_k^T$ is the *intra*-class covariance matrix
- ★ $S_w = \sum_k \frac{1}{n_k} A_k A_k^T$ is the sum of *intra*-class covariance matrices
- ★ We minimise

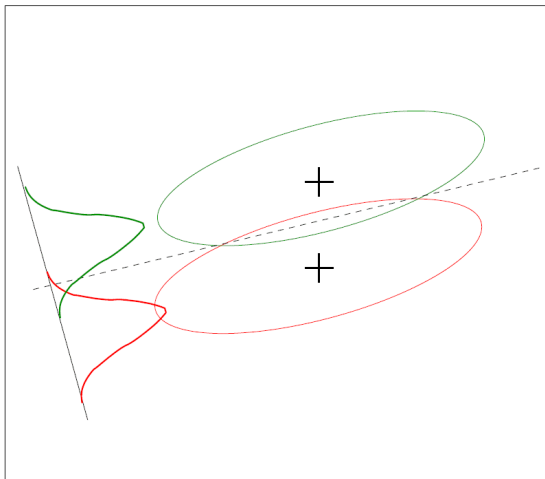
$$\begin{aligned}
 \sum_k \sum_{\mathbf{x}_i \in C_k} (\hat{\mathbf{x}}_i - \hat{\mathbf{g}}_k)^T (\hat{\mathbf{x}}_i - \hat{\mathbf{g}}_k) &= \sum_k \sum_{\mathbf{x}_i \in C_k} \frac{(\mathbf{a}^T (\mathbf{x}_i - \mathbf{g}_k))^T \mathbf{a}^T (\mathbf{x}_i - \mathbf{g}_k)}{\|\mathbf{a}\|^2} \\
 &= \sum_k \frac{1}{\|\mathbf{a}\|^2} \mathbf{a}^T A_k A_k^T \mathbf{a} = \frac{1}{\|\mathbf{a}\|^2} \mathbf{a}^T S_w \mathbf{a}
 \end{aligned}$$

Generalisation: inter-class criteria

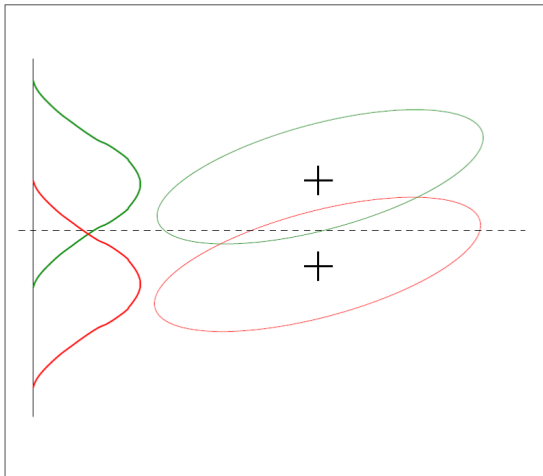
- ★ Let $B = [\mathbf{g}_1 - \mathbf{g}, \dots, \mathbf{g}_q - \mathbf{g}]$ be the matrix of centered data centers ($\mathbf{g} = \frac{1}{N} \sum_N \mathbf{x}_i$ and $N = \sum_k n_k$)
- ★ $S_b = \frac{1}{q} B B^T$ is the covariance matrix of class centers
- ★ We maximise

$$\sum_k (\hat{\mathbf{g}}_k - \hat{\mathbf{g}})^T (\hat{\mathbf{g}}_k - \hat{\mathbf{g}}) = \frac{1}{\|\mathbf{a}\|^2} \mathbf{a}^T S_b \mathbf{a}$$

Mixing both criteria



Mixing both criteria



Fisher discrimination criteria: Raleigh coefficient

- ★ Combining both

$$\max_{\mathbf{a}} J_{\mathbf{a}} = \max_{\mathbf{a}} \frac{\mathbf{a}^T S_b \mathbf{a}}{\mathbf{a}^T S_w \mathbf{a}}$$

- ★ which is found if:

$$\frac{\partial J_{\mathbf{a}}}{\partial \mathbf{a}} = \frac{S_b \mathbf{a} (\mathbf{a}^T S_w \mathbf{a}) - S_w \mathbf{a} (\mathbf{a}^T S_b \mathbf{a})}{(\mathbf{a}^T S_w \mathbf{a})^2} = 0$$

⇒ \mathbf{a} is solution of the generalised eigen system: $S_b \mathbf{a} = J_{\mathbf{a}} S_w \mathbf{a}$

- ★ Hence, \mathbf{a} is the first e.v of $S_w^{-1} S_b$

Discriminant subspaces

- ★ eigenvectors corresponding to the largest eigenvalues λ_i are the most discriminative dimensions

$$\mathbf{a}_1, \mathbf{a}_2, \dots, \mathbf{a}_p \text{ avec } \lambda_1 > \lambda_2 > \dots \lambda_p$$

- ★ q classes may be discriminated in a (at most) $(q - 1)$ -dimensional subspace
- ⇒ only $q - 1$ non-zero eigenvalues

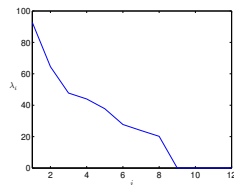
Particular case: 2 classes

- ★ $BB^T = (\mathbf{g}_1 - \mathbf{g})(\mathbf{g}_1 - \mathbf{g})^T + (\mathbf{g}_2 - \mathbf{g})(\mathbf{g}_2 - \mathbf{g})^T = (\mathbf{g}_1 - \mathbf{g}_2)(\mathbf{g}_1 - \mathbf{g}_2)^T$
- ★ hence $BB^T \mathbf{a}$ is a vector along direction $(\mathbf{g}_1 - \mathbf{g}_2)$
- ★ hence $\mathbf{a} \simeq S_w^{-1}(\mathbf{g}_1 - \mathbf{g}_2)$

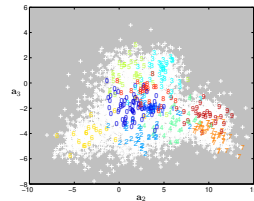
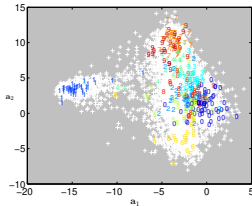
Illustrations : character recognition

7291 images 16×16 (8 bits) numbers from 0 to 9

$\Rightarrow \{\mathbf{x}_i, y_i\}$ avec $\mathbf{x}_i \in \mathbb{R}^{256}$ et $y_i = 1, \dots, 10, i = 1 \dots 7291$



Projection



⇒ LDA finds the optimal subspace to (linearly) separate data along labels y_i .

LDA as a support for decision making

- ★ New data $j \rightarrow \mathbf{x}_j$ known, y_j unknown
 - ★ To which class C_k point j belongs? (classification)
- ⇒ Predict $P(C_k|\mathbf{x}_j)$ (Bayes rule):

$$P(C_k|\mathbf{x}_j) = \frac{P(\mathbf{x}_j|C_k)P(C_k)}{P(\mathbf{x}_j)}$$

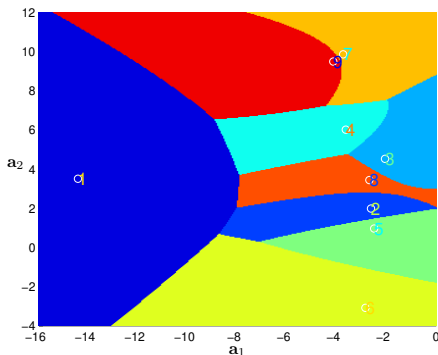
Gaussian approximation

- ★ Each class is modeled by $\mathcal{N}(\mu_k, W_k)$
 - ★ Prior: $P(C_k) = 1/q$
 - ★ evidence $P(\mathbf{x}_j)$ is ignored
- ⇒ Maximum likelihood

$$p(\mathbf{x}|C_k) \approx \exp\left(-(\mathbf{x} - \mu_k)^T W_k (\mathbf{x} - \mu_k)\right)$$

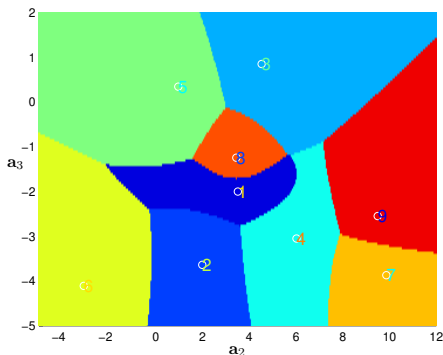
Decision (classification)

$$\delta(\mathbf{x}) = \arg \max_k P(\mathbf{x}|C_k)$$



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Optimality

- ★ LDA is optimal when the q classes are each Gaussian distributed
⇒ because of the discrimination criteria based on covariance matrices S_w et S_b
- ★ **Linear** discriminant Analysis → does not account for **non-linear** relationships between variables