Asset Pricing and ML

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Mean-Variance Optimization: Unconditional i

▶ assets $i = 1, \dots, N$ have prices $P_{i,t}$ and excess returns

$$R_{i,t+1} = \frac{P_{i,t+1} + D_{i,t+1}}{P_{i,t}} - \underbrace{R_{f,t}}_{risk \ free \ rate}$$
 (1)

▶ if you invest fraction $\pi_{i,t}$ of your wealth W_t into security i, the rest stays on your bank account and grows at the rate $R_{f,t}$:

$$W_t = \sum_{i \text{ investment in stock } i} \underbrace{\pi_{i,t} W_t}_{i \text{ bank account}} + \underbrace{(W_t - \sum_{i} \pi_{i,t} W_t)}_{bank account}$$
(2)

Mean-Variance Optimization: Unconditional ii

and then you sell your investments at time t and collect dividends so that

$$W_{t+1} = \sum_{i} W_{t} \pi_{i,t} \frac{P_{i,t+1} + D_{t+1}}{P_{i,t}} + (W_{t} - \sum_{i} \pi_{i,t} W_{t}) R_{f,t}$$

$$= W_{t} R_{f,t} + W_{t} \sum_{i} \pi_{i,t} R_{i,t+1}$$
(3)

Thus, the excess return on your wealth is

$$\frac{W_{t+1}}{W_t} - R_{f,t} = \sum_{i} \pi_{i,t} R_{i,t+1} = \pi'_t R_{t+1}$$
 (4)

lacktriangle Thus, we want π_t that gives good returns. But what is the criterion?

Mean-Variance Optimization: Unconditional iii

► Intuitively, we like **high return** and **low variance**, hence, we might try to find a **static** portfolio that maximizes

$$\pi = \arg\max_{\pi} \left(E[\pi' R_{t+1}] - 0.5 \underbrace{\gamma}_{\substack{\text{risk aversion}}} \operatorname{Var}[\pi' R_{t+1}] \right)$$
 (5)

► The solution is Markowitz

$$\pi = \gamma^{-1} \operatorname{Var}[R]^{-1} E[R]. \tag{6}$$

► Alternatively, one could optimize

$$\pi = \arg\max_{\pi} \left(E[\pi' R_{t+1}] - 0.5 \underbrace{\gamma}_{risk \ aversion} E[(\pi' R_{t+1})^2] \right)$$
 (7)

.

Mean-Variance Optimization: Unconditional iv

and the solution is

$$\tilde{\pi} = \gamma^{-1} (E[R_{t+1}R'_{t+1}])^{-1} E[R_{t+1}]$$

$$= const \cdot \pi, \qquad const = \frac{1}{1 + E[R_{t+1}]' \text{Var}[R_{t+1}]^{-1} E[R_{t+1}]}$$
(8)

where

$$E[R_{t+1}R'_{t+1}] = \operatorname{Var}[R_{t+1}] + E[R_{t+1}]E[R_{t+1}]' = (E[R_{i,t+1}R_{j,t+1}])_{i,j=1}^{N}$$
(9)

Why Are the Two Markowitz Portfolios Proportional? The Sherman-Morrison formula i

The magic behind is the

Lemma (Sherman-Morrison formula)

$$(A + xx')^{-1} = A^{-1} - \frac{A^{-1}xx'A^{-1}}{1 + x'A^{-1}x}$$
 (10)

and

$$(A + xx')^{-1}x = \frac{A^{-1}x}{1 + x'A^{-1}x}$$
 (11)

Proof[Proof of the Sherman-Morrison formula] Recall that

$$xx' = (x_i x_j)_{i,j=1}^N$$

Why Are the Two Markowitz Portfolios Proportional? The Sherman-Morrison formula ii

is a symmetric, positive, semi-definite, rank-1 matrix (all columns are proportional to x). Then,

$$(A + xx')(A^{-1} - \frac{A^{-1}xx'A^{-1}}{1 + x'A^{-1}x})$$

$$= I - \frac{xx'A^{-1}}{1 + x'A^{-1}x} + xx'A^{-1} - xx'\frac{A^{-1}xx'A^{-1}}{1 + x'A^{-1}x}$$

$$= I - \frac{xx'A^{-1}}{1 + x'A^{-1}x} + xx'A^{-1} - xx'A^{-1}\frac{x'A^{-1}x}{1 + x'A^{-1}x} = I$$
(12)

and

$$(A + xx')^{-1}x = (A^{-1} - \frac{A^{-1}xx'A^{-1}}{1 + x'A^{-1}x})x = \frac{A^{-1}x}{1 + x'A^{-1}x}$$
(13)

(Very Big) Issues with Markowiz

Markowitz assumes that we know the truth! The true

$$E[R] = (E[R_{i,t+1}])_{i=1}^{N_t}, Var[R] = (Cov(R_{i,t+1}, R_{j,t+1}))_{i,j=1}^{N_t}$$
 (14)

where N_t is the number of assets (stocks?) available at time t.

- ► The problem is that:
 - expected stock returns move a lot over time: Hence, using static portfolio is a very bad idea
 - we just **do not have enough data** to estimate E[R] and Var[R]. We can use naive

$$\bar{E}[R] = \frac{1}{T} \sum_{t=1}^{T} R_t, \ \overline{\text{Var}}[R] = \frac{1}{T} \sum_{t=1}^{T} \underbrace{(R_t - \bar{E}[R])}_{N \times N} \underbrace{(R_t - \bar{E}[R])'}_{1 \times N}$$

Incorporating Conditional Information: The conditional Markowitz i

We would like to incorporate conditional information.

mean-variance optimization:

$$\pi_t = \arg\max_{\pi_t} \left(E_t[\pi_t' R_{t+1}] - 0.5 \underbrace{\gamma}_{risk \ aversion} \operatorname{Var}_t[\pi_t' R_{t+1}] \right) \quad (15)$$

and hence the Mean-Variance Efficient (MVE) portfolio is

conditional tangency portfolio
$$\begin{array}{c} \underline{\tau_t} \\ = \gamma^{-1} \underbrace{\left(\operatorname{Var}_t[R_{t+1}] \right)^{-1}}_{N \times N \text{ covariance matrix } N \times 1} \underbrace{E_t[R_{t+1}]}_{\text{expected returns}} \\ (16) \end{array}$$

Incorporating Conditional Information: The conditional Markowitz ii

Similarly,

$$\tilde{\pi}_t = \arg\max_{\pi} \left(E_t[\pi' R_{t+1}] - 0.5 \underbrace{\gamma}_{risk \text{ aversion}} E_t[(\pi' R_{t+1})^2] \right) \tag{17}$$

is given by

$$\tilde{\pi}_{t} = \gamma^{-1} (E_{t}[R_{t+1}R'_{t+1}])^{-1} E_{t}[R_{t+1}]$$

$$= \frac{1}{1 + E_{t}[R_{t+1}]' \operatorname{Var}_{t}[R_{t+1}]^{-1} E_{t}[R_{t+1}]} \pi_{t}$$
(18)

where

$$E_t[R_{t+1}R'_{t+1}] = \operatorname{Var}_t[R_{t+1}] + E_t[R_{t+1}]E_t[R_{t+1}]'$$
 (19)

Incorporating Conditional Information: The conditional expectation i

► We would need

$$E_{t}[R_{t+1}] = \arg \min_{F:\mathbb{R}^{P} \to \mathbb{R}^{N}} E[\|R_{t+1} - F(S_{t})\|^{2}]$$

$$E_{t}[R_{t+1}R'_{t+1}] = \arg \min_{G:\mathbb{R}^{P} \to \mathbb{R}^{N \times N}} E[\|R_{t+1}R'_{t+1} - G(S_{t})\|^{2}]$$
(20)

▶ The reality is that we still cannot compute $E[\cdot]$ because we do not have enough data. So, we will still be doing

$$E_t[X_{t+1}] = \arg\min_F \frac{1}{T} \sum_t |X_{t+1} - F(S_t)|^2$$
 (21)

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Intoduction to Asset Pricing i

- ▶ I promised Asset Pricing, but we did Markowitz instead. Why?
- ► Intuitively, we expect that

$$P_{i,t} = \underbrace{(R_{f,t})^{-1} E_t[P_{i,t+1} + D_{i,t+1}]}_{\text{Definitely wrong in the data}}$$
(22)

because the **discount factor** $(R_{f,t})^{-1}$ is too naive

► We need a smart discount factor (SDF):

$$P_{i,t} = E_t \left[\underbrace{M_{t,t+1}}_{\text{stochastic discount factor}} (P_{i,t+1} + D_{i,t+1}) \right]$$
 (23)

Intoduction to Asset Pricing ii

with a bit of algebra, this is equivalent to

$$E_t[R_{i,t+1} M_{t,t+1}] = 0 (24)$$

and

$$\underbrace{E_t[M_{t,t+1}]}_{\text{scale of the SDF}} = R_{f,t}^{-1}$$

By direct calculation,

$$M_{t+1} = 1 - \tilde{\pi}_t' R_{t+1} \tag{25}$$

does the job:

$$E_{t}[R_{t+1}M_{t,t+1}] = E_{t}[R_{t+1}(1 - R'_{t+1}\tilde{\pi}_{t})]$$

$$= E_{t}[R_{t+1}] - E_{t}[R_{t+1}R'_{t+1}]\tilde{\pi}_{t} = 0$$
(26)

Intoduction to Asset Pricing iii

implies

$$\tilde{\pi}_t = E_t[R_{t+1}R'_{t+1}]^{-1}E_t[R_{t+1}]$$
 (27)

We now state

Theorem

Nothing Has Alpha Against $\tilde{\pi}'_t R_{t+1}$

Implications for Testing "If we have found a new, useful strategy"

If we have found the true, ultimate $\tilde{\pi}_t$, nothing has alpha against it: If you have some other portfolio ξ_t and run the regression

$$\xi_t' R_{t+1} = \alpha + \beta \tilde{\pi}_t' R_{t+1} + \varepsilon_{t+1}.$$

- ▶ If α is not significant, reject the strategy ξ_t , move somewhere else.
- ▶ if α is significant, the $\tilde{\pi}_t$ is not efficient and you should try combining it with $\tilde{\pi}_t$

Testing Conditional Efficiency

- ▶ We cannot compute $E_t[\cdot]$
- ▶ Instead, we can build instruments Z_t and test that

$$E_t[M_{t+1}R_{t+1}] = 0 \Leftrightarrow E[Z_t M_{t+1}R_{t+1}] = 0$$

for all instruments!

ightharpoonup Thus, we need to build infinitely many Z_t thought machine learning and then test

$$\frac{1}{T}\sum_{t}Z_{t}\,M_{t+1}R_{t+1}\,\approx\,0$$

Complexity is always there!

From Non-Tradable to Tradable SDFs i

- ► What about asset pricing theory?
- ▶ the SDF

$$\widetilde{M}_{t+1} = \underbrace{\frac{e^{-\rho}U'(C_{t+1})}{U'(C_t)}}_{=IMRS}$$

comes from the Euler equation (things get more complex with Epstein-Zin preferences, expectations, sentiments, etc)

$$E_{t}\left[\underbrace{\frac{e^{-\rho}U'(C_{t+1})}{U'(C_{t})}}_{=IMRS}(R_{t+1}+R_{f,t})\right] = 1 \Leftrightarrow E_{t}[\widetilde{M}_{t+1}R_{t+1}] = 0 \quad (28)$$

because

$$R_{f,t} = E_t[\widetilde{M}_{t+1}]^{-1}. (29)$$

From Non-Tradable to Tradable SDFs ii

► When markets are complete,

$$M_{t+1}$$
 proportional to \widetilde{M}_{t+1}

and the proportionality constant can be pinned down by

$$\widetilde{M}_{t+1} = \frac{M_{t+1}}{R_{f,t} E_t[M_{t+1}]}$$
 (30)

The normalization ensures that

$$E_t[\widetilde{M}_{t+1}] = R_{f,t}^{-1}$$
 (31)

From Non-Tradable to Tradable SDFs iii

► In general, we need to **project**

$$\underbrace{\frac{M_{t+1}}{R_{f,t}E_t[M_{t+1}]}}_{unique\ tradable} = Proj_t(\widetilde{M}_{t+1}) = \arg\min_{a,\pi} E_t[(\widetilde{M}_{t+1} - a(1 - \pi'R_{t+1}))^2]$$
(32)

solution is

$$\pi = \tilde{\pi}_t, \ a = \frac{1}{R_{f,t}E_t[M_{t+1}]}$$

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Panel Datasets: Leveraging the Power of Big Data i

- Now comes the big question: How do we measure the conditional expectations, $E_t[R_{t+1}]$ and $E_t[R_{t+1}R'_{t+1}]$?
- Running prediction models per stock is infeasible due to insufficient data:

$$E_t[R_{i,t+1}] = g_i(X_{i,t})$$

use panel data

$$E_t[R_{i,t+1}] = g(X_{i,t})$$

- **panel** means **same function** *g* for all stocks.
- ▶ non-linear g means machine learning
- ► What about the covariance matrix? How do we model a **time-varying** covariance structure?

Panel Datasets: Leveraging the Power of Big Data ii

► Typically, we assume a factor structure:

$$R_{t+1} = \underbrace{S}_{factor} \underbrace{F_{t+1}}_{factors} + \varepsilon_{t+1}$$

► In reality, factor exposures are time-varying:

$$R_{t+1} = S_t F_{t+1} + \varepsilon_{t+1}$$

▶ If $Cov_t[F_{t+1}] = \Sigma_F$ and

$$\operatorname{Cov}_t[\varepsilon_{t+1}] = \operatorname{diag}(\sigma_{i,t}^2), \ \sigma_{i,t} = \operatorname{idiosyncratic} \ \operatorname{volatility}$$

so that the Conditional covariance matrix is given by

$$E_t[R_{t+1}R'_{t+1}] = S'_t\Sigma_F S_t + \operatorname{diag}(\sigma^2_{i,t})$$

Panel Datasets: Leveraging the Power of Big Data iii

► Equivalently:

$$E_t[R_{i,t+1}R_{j,t+1}] = \underbrace{S'_{i,t}\Sigma_F S_{j,t}}_{systematic\ covariance} + \underbrace{\delta_{i,j}\sigma^2_{i,t}}_{idiosyncratic\ variance}$$

where Σ_F and $\sigma_{i,t}$ are to be estimated.

► Can we avoid computing the conditional covariance matrix?

Managed Portfolios and Rich Conditional Factor Structures i

Suppose

$$R_{i,t+1} = \underbrace{S'_{i,t}}_{conditional\ betas} \cdot \underbrace{\tilde{F}_{t+1}}_{latent\ factors} + \varepsilon_{i,t+1}$$

$$E_t[\tilde{F}_{t+1}] = \underbrace{\lambda_F}_{latent \ factor \ risk \ premia}, \ E_t[\tilde{F}_{t+1}\tilde{F}'_{t+1}] = \underbrace{\Sigma_F}_{latent \ factor \ cov}$$

► Thus,

$$E_t[R_{t+1}] = S_t \lambda_F$$

and

$$E_t[R_{t+1}R'_{t+1}] = S_t\Sigma_F S'_t + \Sigma_{\varepsilon}$$

Managed Portfolios and Rich Conditional Factor Structures ii

$$M_{t+1} = 1 - \tilde{\pi}'_t R_{t+1} = 1 - W(S_t)' R_{t+1},$$
 (33)

where $\tilde{\pi}_t = E_t[R_{t+1}R'_{t+1}]^{-1}E_t[R_{t+1}]$ and, hence,

$$W(S_t) = \underbrace{(S_t \Sigma_{F,t} S_t' + \Sigma_{\varepsilon})^{-1}}_{conditional\ covariance\ conditional\ expectation} \underbrace{S_t \lambda_F}_{conditional\ expectation}$$
(34)

▶ Define managed portfolios

$$F_{t+1} = S_t' R_{t+1}. (35)$$

and the unconditionally efficient portfolio

$$\lambda = E[F_{t+1}F'_{t+1}]^{-1}E[F_{t+1}] \tag{36}$$

Managed Portfolios and Rich Conditional Factor Structures iii

▶ By construction,

$$M^{F}_{t+1} = 1 - \lambda' F_{t+1} \tag{37}$$

prices factors unconditionally:

$$E[M^{F}_{t+1}F_{t+1}] = 0 (38)$$

► However,

$$E_t[M_{t+1}^F R_{t+1}] \neq 0$$

because

$$\lambda' S_t' R_{t+1} \neq \lambda_{\mathcal{F}}' S_t' \Sigma_t^{-1} R_{t+1},$$

with

$$\Sigma_t = (S_t \Sigma_{F,t} S_t' + \Sigma_{\varepsilon})$$

Click on this link to know more:

APT or "AIPT"? The Surprising Dominance of Large Factor Models

Theorem

Suppose that in the limit, as $P \to \infty$, the vector of latent risk premia λ_F satisfies

$$\lambda_F' A \lambda_F \rightarrow 0 \tag{39}$$

for any symmetric, positive definite A with uniformly bounded trace. Let

$$M^{F}_{t+1} = 1 - \lambda' F_{t+1},$$
 (40)

be the factor approximation for the SDF with λ . Then, M^F_{t+1} converges to M_{t+1} and the Sharpe ratio of $\lambda' F_{t+1}$ converges to that of $W(S_t)' R_{t+1}$ as $P \to \infty$. In particular,

$$E_t[M_{t+1}^{\mathsf{F}}R_{t+1}] \rightarrow 0$$

Sources of Complexity i

▶ We now know: If

$$R_{t+1} = \underbrace{S_t}_{N_t \times P \text{ signals } P \times 1 \text{ latent factors}} + \underbrace{\varepsilon_{t+1}}_{\text{residuals}}$$
(41)

then we build

$$F_{t+1} = S'_t R_{t+1} = (S'_t S_t) \tilde{F}_{t+1} + (S'_t \varepsilon_{t+1})$$
 (42)

- ▶ But where do S_t come from?
- Suppose

$$R_{i,t+1} = \beta(X_{i,t})'G_{t+1} + u_{i,t+1}, \tag{43}$$

Sources of Complexity ii

$$\beta(X_{i,t}) \approx \sum_{p=1}^{P} \xi_p S_{i,t,p} = \xi' \underbrace{S_{i,t}}_{P \times 1}, \tag{44}$$

where

$$S_{i,t} = (\sigma(\omega_p' X_{i,t}))_{p=1}^P$$
 (45)

► This gives

$$\underbrace{R_{t+1}}_{N\times 1} \approx \underbrace{S_t}_{N\times P} \underbrace{\tilde{F}_{t+1}}_{P\times 1} + u_{t+1}, \text{ with}$$

$$\tilde{F}_{t+1} = \underbrace{\xi}_{P\times 1} G_{t+1}, \quad \nu = E[\tilde{F}_{t+1}] = \xi E[G_{t+1}].$$
(46)

▶ If β is highly non-linear, we need to go for a high-dimensional S_t

Sources of Complexity iii

► The true SDF return is

$$(\beta_t \beta_t' + \Sigma_u)^{-1} \beta_t E[G_{t+1}] = \sum_{Sherman-Morrison} \Sigma_u^{-1} \beta_t E[G_{t+1}] \frac{1}{1 + \beta_t' \Sigma_u^{-1} \beta_t}$$

$$(47)$$

In high dimensions, $\beta_t' \Sigma_u^{-1} \beta_t \approx const$. Furthermore, if β_t are sufficiently complex, $\Sigma_u^{-1} \beta_t \approx const \beta_t$. Thus, we end up with

$$\pi_t \sim \beta_t = S_t \xi \tag{48}$$

and the SDF is

$$\pi'_{t}R_{t+1} = \underbrace{\xi'}_{factor\ weights}\underbrace{S'_{t}R_{t+1}}_{F_{t+1}}.$$
 (49)

Complexity in the Cross Section: A Brief History i

Most academic attempts to build an SDF assume

$$M_{t+1}^{\star} = 1 - \sum_{i=1}^{N} w(X_{i,t}) R_{i,t+1}$$
 (50)

- lacktriangle Cross-sectional asset pricing is about $w_t = w(X_t)$
 - Explains differences in average returns
 - Defines the MVE portfolio
- ▶ Why does cross-section literature rarely start here? Because w must be estimated
 - This is a high-dimensional (complex) problem
 - We know: In-sample tangency portfolio behaves horribly out-of-sample
 - Why? Complexity $(n/T \not\to 0) \to \text{LLN}$ doesn't apply $\to \text{IS}$ and OOS diverge

Complexity in the Cross Section: A Brief History ii

- ► Standard solution: Restrict w
 - E.g., Fama-French: $w_{i,t} = b_0 + b_1 \text{Size}_{i,t} + b_2 \text{Value}_{i,t}$ (Brandt et al. 2007 generalize):

$$\sum_{i=1}^{N} w(X_{i,t})R_{i,t+1} = \sum_{i=1}^{N} (b_0 + b_1 \operatorname{Size}_{i,t} + b_2 \operatorname{Value}_{i,t})R_{i,t+1}$$

$$= b_0 \sum_{i=1}^{N} R_{i,t+1} + b_1 \sum_{i=1}^{N} \operatorname{Size}_{i,t}R_{i,t+1} + b_2 \sum_{i=1}^{N} \operatorname{Value}_{i,t}R_{i,t+1}$$

$$= b_0 MKT_{t+1} + b_1 SMB_{t+1} + b_2 HML_{t+1}.$$
(51)

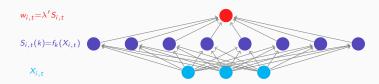
• Reduces parameters, implies factor model: $M_{t+1} = 1 - b_0 MKT - b_1 SMB - b_2 HML$

 "Shrinking the cross-section" Kozak et al. (2020) — use a few PCs of anomaly factors

Complexity in the Cross Section: Machine Learning Perspective i

Rather than restricting $w(X_t)$

- ► ...expand parameterization, saturate with conditioning information
- ▶ For example, approximate w with neural network: $w(X_{i,t}) \approx \lambda' S_{i,t}$
- ▶ $P \times 1$ vector $S_{i,t}$ is known nonlinear function of original predictors $X_{i,t}$



Complexity in the Cross Section: Machine Learning Perspective ii

▶ Implies that empirical SDF is a high-dimensional factor model

$$\sum_{i=1}^{N} w(X_{i,t}) R_{i,t+1} = \sum_{i=1}^{N} \left(\sum_{k} \lambda_{k} \underbrace{S_{i,t}(k)}_{S_{i,t}(k) = f_{k}(X_{i,t})} R_{i,t+1} = \sum_{k} \lambda_{k} \underbrace{\sum_{i=1}^{N} S_{i,t}(k) R_{i,t+1}}_{F_{k,t+1}} \right)$$
(52)

$$M_{t+1}^{\star} \approx M_{t+1} = 1 - \lambda' S_t' R_{t+1} = 1 - \lambda' F_{t+1}$$

Complexity in the Cross Section: Machine Learning Perspective i

The Objective:

► Maximize out-of-sample Sharpe ratio (equivalently, minimize out-of-sample pricing errors) of SDF

The Choice:

► Fix *T* data points. Decide on "complexity" (number of factors *P*) to use in approximating model

The Tradeoff:

Complexity in the Cross Section: Machine Learning Perspective ii

- ▶ Simple SDF ($P \ll T$) has low variance (thanks to parsimony) but is a poor approximator of w
- ► Complex SDF (*P* > *T*) is a good approximator but may behave poorly (and requires shrinkage)

▶ Which *P* should the analyst opt for? Does the benefit of more factors justify their cost?

Answer:

▶ Use the largest factor model (largest *P*) that you can compute

Implementation i

- Build a bunch of features (random features if you want a shallow model; deep features (output layer) if you want a deep model).
- ► Call them $S_{i,t}(k) = f_k(X_{i,t}; \theta_k), k = 1, \dots, P$
- ► Build the factors

$$F_{t+1}(k) = \sum_{i=1}^{N_t} S_{i,t}(k) R_{i,t+1}$$
 (53)

▶ Take the vector of factors $F_{t+1} = (F_{t+1}(k))_{k=1}^P$ and minimize

$$\min_{\lambda} \frac{1}{T} \sum_{t=1}^{I} (1 - \lambda' F_{t+1})^2 + z \|\lambda\|^2$$
 (54)

This objective is known as the Maximal Sharpe Ratio Regression (MSRR). For a deep model, you need to minimize this objective using GD

Implementation ii

Why MSRR? Well,

$$\frac{1}{T} \sum_{t=1}^{T} (1 - \lambda' F_{t+1})^2 \approx E[(1 - \lambda' F_{t+1})^2]$$

$$= 1 - 2E[\lambda' F_{t+1}] + E[(\lambda' F_{t+1})^2] = 1 - 2E[U(\lambda' F_{t+1})], \tag{55}$$

where

$$U(x) = x - 0.5x^2$$

▶ Now, $\tilde{\pi}_t = E_t[R_{t+1}R'_{t+1}]^{-1}E_t[R_{t+1}]$ solves

$$\max_{\pi} E_t[U(\pi_t' R_{t+1})] \tag{56}$$

Implementation iii

It is conditionally efficient for a quadratic utility. By the law of iterated expectations,

$$E[E_t[U(\pi_t'R_{t+1})]] = E[U(\pi_t'R_{t+1})]$$

and dynamic consistency gives

$$\max_{\textit{all policies } \pi_t} E[U(\pi_t' R_{t+1})] = E[\max_{\pi} E_t[U(\pi_t' R_{t+1})]$$

Thus, MSRR looks for conditional policies that maximize unconditional utility and hence, by consistency, are conditionally optimal.

RMT i

$$\hat{\lambda}(z) = \left(zI + \frac{1}{T}\sum_{t=1}^{T}F_{t}F'_{t}\right)^{-1}\frac{1}{T}\sum_{t=1}^{T}F_{t} \underset{Complexity!}{\cancel{\nearrow}} \lambda_{*}(z) \quad (57)$$

where

$$\lambda_*(z) = (zI + E[FF'])^{-1}E[F]$$
 (58)

RMT ii

► Leave-One-Out (LOO):

$$\hat{\Psi} = \frac{1}{T} \sum_{\tau=1}^{T} F_{\tau} F_{\tau}'$$

$$\hat{\Psi}_{T,t} = \frac{1}{T} \sum_{\tau \neq t}^{T} F_{\tau} F_{\tau}'$$

$$(zI + \hat{\Psi})^{-1} F_{t} = \left(zI + \hat{\Psi}_{T,t}\right)^{-1} F_{t} \frac{1}{1 + T^{-1} F_{t}' \left(zI + \hat{\Psi}_{T,t}\right)^{-1} F_{t}}$$
(59)

RMT iii

Define the Stieltjes Transform

$$\hat{m}(-z) = P^{-1} \operatorname{tr}((zI + \hat{\Psi})^{-1})$$
 (60)

and

$$\hat{Z}_*(z;c) = \frac{z}{1-c+cz\hat{m}(-z)}$$
 (61)

and

$$\hat{\xi}(z;c) = -1 + \frac{1}{1 - c + cz\hat{m}(-z)}.$$
 (62)

We have

$$1/Z_*(z;c) = \lim_{T \to T} T^{-1} \operatorname{tr}((zI + FF'/T)^{-1}). \tag{63}$$

▶ Lemma

$$T^{-1}F'_t\left(zI+\hat{\Psi}_{T,t}\right)^{-1}F_t \approx \hat{\xi}(z;c) \tag{64}$$

► Implicit Regularization

$$E[\hat{\lambda}(z)'F_{T+1}] \approx \frac{Z_*(z)}{z} E[\lambda_*(Z_*(z))'F_{T+1}],$$
 (65)

where

$$Z_*(z) > z. (66)$$

► In fact,

$$E_{T}[\hat{\lambda}(z)'F_{T+1}] = \frac{Z_{*}(z)}{z}E_{T}[\lambda_{*}(Z_{*}(z))'F_{T+1}]$$

$$= \frac{Z_{*}(z)}{z}\lambda_{*}(Z_{*}(z))'E[F] = \frac{Z_{*}(z)}{z}E[F]'(Z_{*}(z)I + E[FF'])^{-1}E[F]$$
(67)

The RMT Master Theorem

Theorem

$$P^{-1}z\operatorname{tr}(A_{P}(zI+\underbrace{\hat{\Psi}}_{random})^{-1}) - P^{-1}Z_{*}\operatorname{tr}(A_{P}(Z_{*}I+\underbrace{\Psi}_{deterministic})^{-1}) \rightarrow 0$$
(68)

almost surely.

Similarly, for any sequence of uniformly bounded vectors β , we have

$$z\beta'(zI + \underbrace{\hat{\Psi}}_{random})^{-1}\beta - Z_*\beta'(Z_*I + \underbrace{\Psi}_{deterministic})^{-1}\beta \rightarrow 0$$
 (69)

The Expected Return Calculation i Proof.

Let $E[F] = \mu$, $E[FF'] = \Psi$; everything is i.i.d. across t. Then,

$$E[\hat{\lambda}(z)'F_{T+1}] = E[\hat{\lambda}(z)'\mu] = E[\frac{1}{T}\sum_{t=1}^{T}F_{t}'\left(zI + \frac{1}{T}\sum_{t=1}^{T}F_{t}F_{t}'\right)^{-1}]\mu$$

$$= E[F_{t}'\left(zI + \frac{1}{T}\sum_{t=1}^{T}F_{t}F_{t}'\right)^{-1}]\mu$$

$$= E\left[F_{t}'(zI + \Psi_{T,t})^{-1}\frac{1}{1 + T^{-1}F_{t}'\left(zI + \hat{\Psi}_{T,t}\right)^{-1}F_{t}}\right]\mu$$

$$= \mu'E[(zI + \hat{\Psi}_{T,t})^{-1}]\mu(1 + \xi(z;c))^{-1}$$

(70)

The Expected Return Calculation ii Proof.

where

$$\hat{\Psi}_{T,t} = \frac{1}{T} \sum_{\tau=1}^{I} F_{\tau} F_{\tau}' - F_{t} F_{t}'$$

where we have used that

$$T^{-1}F'_t\left(zI+\hat{\Psi}_{T,t}\right)^{-1}F_t \approx \xi(z;c) \tag{71}$$

The claim follows now from the Master Theorem:

$$z \mu' \left(z I + \hat{\Psi}_{T,t} \right)^{-1} \mu \approx Z_* (Z_* I + \Psi)^{-1}$$
 (72)

LLG

The Limits-to-Learning Gap (LLG)

$$\mathcal{L}(z;c) = \underbrace{\frac{d}{dz} Z_*(z;c) - 1}_{LLG} = \lim \frac{T^{-1} \operatorname{tr}((zI + FF'/T)^{-2})}{(T^{-1} \operatorname{tr}((zI + FF'/T)^{-1}))^2} - 1 (73)$$

is always in [0, T-1].

Theorem

$$\underline{\mu'\Sigma^{-1}\mu} \geq (1 + \mathcal{L}(z;c)) \underbrace{SR_{OOS}^2(\hat{\lambda}(z))}_{feasible \ OOS \ SR}$$
(74)

Table of Contents

1 Mean-Variance Optimization

2 Introduction: Complexity in Cross-Sectional Asset Pricing

3 Empirical Asset Pricing Via Machine Learning

4 Empirics for the US Stock Market

Empirical Analysis

- ► Analyze empirical analogues to theoretical comparative statics
- Study conventional setting with conventional data
 - Forecast target is monthly return of US stocks from CRSP 1963–2021
 - Conditioning info (X_t) is 130 stock characteristics from Jensen, Kelly, and Pedersen (2022)
- ► Out-of-sample performance metrics are:
 - SDF Sharpe ratio
 - Mean squared pricing errors (factors as test assets)

Empirical Analysis i

Random Fourier Features

- ► Empirical model: $M_{t+1} = 1 \lambda' S_t' R_{t+1}$
- ▶ Need framework to smoothly transition from low to high complexity
- Adopt ML method known as "random Fourier features" (RFF)
 - Let $X_{i,t}$ be 130×1 predictors. RFF converts $X_{i,t}$ into

$$S_{\ell,i,t} = \sin(\gamma'_{\ell}X_{i,t}), \quad \gamma_{\ell} \sim iidN(0,\gamma I)$$

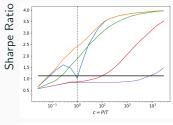
- $S_{\ell,i,t}$: Random lin-combo of $X_{i,t}$ fed through non-linear activation
- we then rank the random features in the cross-section
- ► For fixed inputs can create an arbitrarily large (or small) feature set
 - Low-dim model (say P = 1) draw a single random weight
 - High-dim model (say P = 10,000) draw many weights
- ▶ In fact, RFF is a two-layer neural network with fixed weights (γ) in the first layer and optimized weights (λ) in the second layer

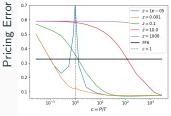
Empirical Analysis

Training and Testing

- ► We estimate out-of-sample SDF with:
 - i. Thirty-year rolling training window (T = 360)
 - ii. Various shrinkage levels, $log_{10}(z) = -12, ..., 3$
 - iii. Various complexity levels $P=10^2,...,10^6$
- ▶ For each level of complexity c = P/T, we plot
 - i. Out-of-sample Sharpe ratio of the kernels and
 - ii. Pricing errors on 10^6 "complex" factors: $F_{t+1} = S_t' R_{t+1}$
- ▶ Also report Sharpe ratio and pricing errors of FF6 to benchmark our results

Out-of-sample SDF Performance

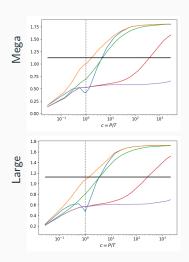


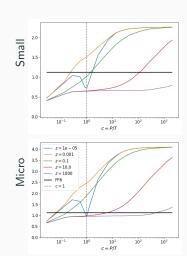


Main Empirical Result

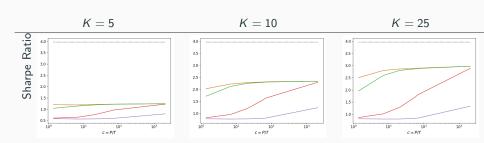
- OOS behavior of ML-based SDF closely matches theory
- ► High complexity models
 - Improve over simple models by a factor of 3 or more
 - Dominate popular benchmarks like FF6

SDF Performance in Restricted Samples: Sharpe Ratio Market Capitalization Subsamples





What About "Shrinking" With PCA?



Beyond Own-Signal Portfolios i

All portfolio strategies we have used so far use own-signal weights:

$$\pi_{i,t} = w(S_{i,t}) = \sum_{k} \lambda_k f_k(X_{i,t})$$

where λ_k are estimated through Markowitz.

Beyond Own-Signal Portfolios ii

In Artificial Intelligence Pricing Models, we show how to build strategies that use other stocks' information. The insight is simple: Instead of

$$\pi_t = S_t \lambda,$$

we do

$$\pi_t = \underbrace{A_t S_t \lambda}_{one \ transformer \ block}$$

where

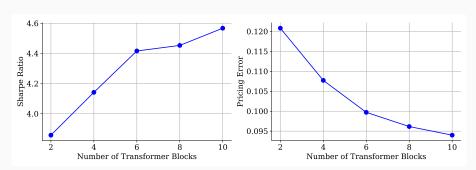
$$A_t = F(S_t M S_t')$$

is the attention matrix, and F is a non-linear transformation.

You can repeat this trick many times, making the attention deeper.

Beyond Own-Signal Portfolios iii

Figure: Virtue of complexity for K-block transformer portfolios.



Experiments with Managed Portfolios

Managed Portfolios Notebook