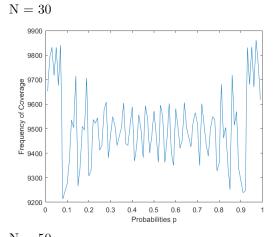
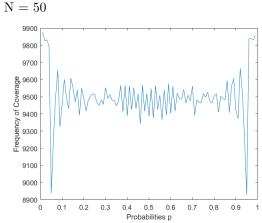
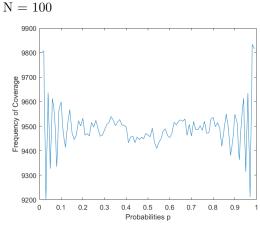
STAT 5444: Homework #2

Due on October 17, 2016 at 3:10pm $Professor\ Scott\ Leman$

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For low probabilities of p or extremely high probabilities the coverage is within the 95 percent more so in this example. There is a bit more variance in hitting the 95 percent coverage mark in this Bayesian approach. The beta prior provides support for the areas with low probability p or high probability p which aids in acquiring coverage in those areas. This is particularly useful for ensuring a centered credible interval for all probabilities p.

```
alpha = 0.5;
beta = 0.5;
frequency = zeros(99, 1);
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```
N = 100;
totalX = zeros(10000, 10000, 1);
temp = 1;
for p = 0.01:0.01:0.99
    x = binornd(N, p, 10000, 1);
    alphaStar = x + alpha;
    betaStar = beta + N - x;
    for j = 1:10000
        totalX(:, j) = betarnd(alphaStar, betaStar);
    end
    aSort = sort(totalX, 2);
    upper_bound = aSort(:, 9750) >= p;
    lower_bound = aSort(:, 250) \le p;
    frequency (temp, 1) = sum(upper_bound == lower_bound);
    temp = temp + 1;
end
p = 0.01:0.01:0.99;
plot(p, frequency)
xlabel('Probabilities p')
ylabel ('Frequency of Coverage')
```

$$\begin{split} p(\phi|X) &= \int p(\mu|X\phi)p(\phi)p(\mu)d\mu \\ p(\phi|X) &= \int \prod_{i=1}^{N} \frac{\phi^2}{\sqrt{2\pi}} e^{\frac{\phi}{2}(x-\mu)^2} d\mu \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} \int e^{\frac{\phi}{2}(x-\mu)^2} d\mu \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} \int e^{\frac{\phi}{2}((\sum_{i=1}^{N} x_i^2) - 2\mu \sum_{i=1}^{N} x_i + N\mu^2)} d\mu \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} \int e^{\frac{\phi}{2}(N\mu^2 - 2\mu \sum_{i=1}^{N} x_i)} d\mu \end{split}$$

To acquire the pdf we must have this

$$V = (\phi N)^{-1} = \frac{\phi^{-1}}{N}$$

$$E = \frac{\phi^{-1}}{N} * \sum_{i=1}^{N} x_i \phi = \bar{x}$$

$$pdfNormal = \int e^{\frac{-1}{2V}(u^2 - 2\bar{x}\mu + \bar{x}^2)}$$

Thus, we integrate to 1

$$\begin{split} p(\phi|X) &\propto \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} \int e^{\frac{\phi}{2}(N\mu^2 - 2\mu\sum_{i=1}^{N} x_i)} d\mu * \frac{\frac{1}{\sqrt{2\pi V} e^{\frac{-1}{2}\frac{\bar{x}^2}{V}}}}{\frac{1}{\sqrt{2\pi V} e^{\frac{-1}{2}\frac{\bar{x}^2}{V}}}} \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} \sqrt{2\pi V} e^{\frac{1}{2}\frac{\bar{x}^2}{V}} \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} \sqrt{2\pi \frac{\phi^{-1}}{N}} e^{\frac{1}{2}\frac{\bar{x}^2}{\frac{\bar{\phi}^{-1}}{N}}} \\ p(\phi|X) &\propto \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} \phi^{\frac{-1}{2}} e^{\frac{1}{2}N\bar{x}^2\phi} \\ p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2)} e^{\frac{-1}{2}(-N\bar{x}^2\phi)} \\ p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2 - N\bar{x}^2)} \\ p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2}(\sum_{i=0}^{N} x^2 - \sum_{i=0}^{N} \bar{x}^2)} \end{split}$$

The follow theorem is used:

$$E(x - \bar{x})^2) = E(x^2 - 2x\bar{x} + \bar{x}^2)$$

$$= E(x^2 - 2E(x)^2 + E(x)^2)$$

$$= E(x^2) - E(x)^2$$

$$N * E(x - \bar{x})^2 = \sum_{i=0}^{N} (x - \bar{x})^2$$

$$N * E(x^2) - N * E(x)^2 = (\sum_{i=0}^{N} x^2 - \sum_{i=0}^{N} \bar{x}^2)$$

Thus, we can do the following

$$\begin{split} p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2} (\sum_{i=0}^{N} x^2 - \sum_{i=0}^{N} \bar{x}^2)} \\ p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2} (\sum_{i=0}^{N} (x-\bar{x})^2)} \\ s^2 &= \frac{\sum (x_i - \bar{x})^2}{N-1} \\ p(\phi|X) &\propto \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2} (s^2(N-1))} \\ p(\phi|X) &- Gamma \big[\frac{N-1}{2}, \frac{(s^2(N-1))}{2} \big] \end{split}$$

Part 2:

$$\begin{split} P(\widetilde{x}|X) &= \int P(\widetilde{x}|X\phi) P(\phi) d\phi \\ P(\widetilde{x}|X) &= \int \frac{\phi^{\frac{1}{2}}}{\sqrt{2\pi}(1+\frac{1}{N})} e^{\frac{-\phi}{2}\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})}} \phi^{\frac{N-1}{2}-1} e^{\frac{-\phi}{2}(s^2(N-1))} \\ P(\widetilde{x}|X) &\propto \int \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})}} e^{\frac{-\phi}{2}(s^2(N-1))} \\ P(\widetilde{x}|X) &\propto \int \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))} * \frac{\Gamma(N/2)}{\left[\frac{1}{2}*(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))\right]^{N/2}} \\ P(\widetilde{x}|X) &\propto \int \phi^{\frac{N}{2}-1} e^{\frac{-\phi}{2}(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))} * \frac{\Gamma(N/2)}{\left[\frac{1}{2}*(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))\right]^{N/2}} \\ P(\widetilde{x}|X) &\propto \frac{\Gamma(N/2)}{\left[\frac{1}{2}*(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))\right]^{N/2}} \\ P(\widetilde{x}|X) &\propto \frac{\Gamma(N/2)}{\left[\frac{1}{2}*(\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1)))\right]^{N/2}} \\ P(\widetilde{x}|X) &\propto \left[\frac{(\widetilde{x}-\widetilde{x})^2}{(1+\frac{1}{N})} + (s^2(N-1))\right]^{-N/2} \end{split}$$

$$P(\tilde{x}|X) \propto \left[\frac{(\tilde{x} - \bar{x})^2}{(1 + \frac{1}{N})} + (s^2(N - 1)) \right]^{-N/2}$$

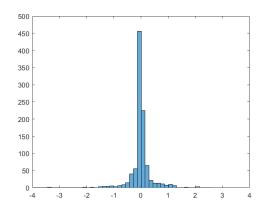
$$P(\tilde{x}|X) \propto \left[1 + \frac{(\tilde{x} - \bar{x})^2}{(1 + \frac{1}{N})(s^2(N - 1))} \right]^{-N/2}$$

$$P(\tilde{x}|X) \propto \left[1 + \frac{(\tilde{x} - \bar{x})^2}{(1 + \frac{1}{N})(s^2(N - 1))} \right]^{\frac{-(N - 1) + 1}{2}}$$

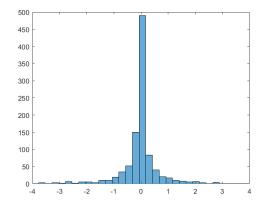
$$\tilde{x}|X - t(\bar{x}, s^2(1 + \frac{1}{N}, N - 1))$$

$$\begin{split} p(x|\mu,\sigma^2,\gamma) &\propto (\phi/\gamma)^{1/2} exp(-\frac{\phi}{2}\frac{(x-\mu)^2}{\gamma} \\ p(\phi) &\propto \phi^{\alpha-1} e^{-\beta\phi} \\ p(x|\mu,\gamma) &\propto \int p(x|\mu,\sigma^2,\gamma) p(\phi) d\phi \\ p(x|\mu,\gamma) &\propto \int \frac{\phi^{\frac{1}{2}}}{\gamma} e^{\left[-\frac{\phi}{2}\frac{(x-\mu)^2}{\gamma}\right]} * \phi^{\alpha-1} e^{-\beta\phi} \\ p(x|\mu,\gamma) &\propto \int \frac{\phi^{\frac{1}{2}+\alpha-1}}{\gamma} e^{\left[-\frac{\phi}{2}\frac{(x-\mu)^2}{\gamma} - \beta\phi\right]} \\ p(x|\mu,\gamma) &\propto \int \frac{\phi^{\frac{1}{2}+\alpha-1}}{\gamma} e^{-\phi\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]} * \frac{\Gamma(\frac{1}{2}+\alpha)}{\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]^{\frac{1}{2}+\alpha}} \\ p(x|\mu,\gamma) &\propto \int \frac{(1-\alpha)^2}{\gamma} e^{-\phi\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]} * \frac{\Gamma(\frac{1}{2}+\alpha)}{\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]^{\frac{1}{2}+\alpha}} \\ p(x|\mu,\gamma) &\propto \frac{\Gamma(\frac{1}{2}+\alpha)}{\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]^{\frac{1}{2}+\alpha}} \\ p(x|\mu,\gamma) &\propto \frac{1}{\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]^{\frac{1}{2}+\alpha}} \\ p(x|\mu,\gamma) &\propto \frac{1}{\left[\frac{(x-\mu)^2}{2\gamma} + \beta\right]^{\frac{1}{2}+\alpha}} \\ p(x|\mu,\gamma) &\propto \frac{1}{\left[\frac{(x-\mu)^2}{2\gamma} + 1\right]^{\frac{1}{2}+\frac{1}{2}}}, \alpha = \frac{1}{2}, \beta = 1 \end{split}$$

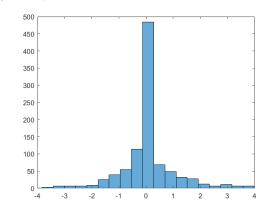
Part 3: $\mu = 0, \gamma = 0.5, min = -4.18, max = 3.1$



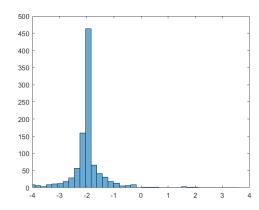
$$\mu = 0, \gamma = 1, min = -9.74, max = 7.094$$



$$\mu = 0, \gamma = 2, min = -18.51, max = 11.33$$



 $\mu = -2, \gamma = 1, min = -6.18, max = 6.51$



MAP is not invariant to transformations. Using a normal distributed MAP estimator with no prior the MAP estimator is the value 0. Applying a transformation of $\tau(\theta) = \theta^2$ will result in the MAP estimator being the value 0.7. When we apply the transformation to the original value 0 it will not yield the value 0.7. Thus, the MAP is not invariant to transformations.

Problem 5

Part 1:

In previous derivations we have the following posterior: $\mu - N(\bar{x}, \frac{\sigma^2}{N})$

To determine if the predictive distribution is normal we can apply the prediction distribution function. Since it ends being an integral of a normal * normal it will end up being a normal just be rules of conjugate priors. Proof Below

$$\begin{split} P(\widetilde{x}|X,\sigma^2) &= \int \frac{1}{\sqrt{2\pi}\sigma} * e^{\frac{-1}{2}\frac{(\widetilde{x}-\mu)^2}{\sigma^2}} * \frac{1}{\sqrt{2\pi}\frac{\sigma}{\sqrt{N}}} * e^{\frac{-1}{2}\frac{(\mu-\widetilde{x})^2}{\sigma^2}} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto \int e^{\frac{-1}{2}\frac{(\widetilde{x}-\mu)^2}{\sigma^2}} * e^{\frac{-1}{2}\frac{(\mu-\widetilde{x})^2}{\sigma^2}} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto \int e^{\frac{-1}{2}\frac{\widetilde{x}^2-2\widetilde{x}\mu+\mu^2}{\sigma^2}} * e^{\frac{-1}{2}\frac{\mu^2-2\mu\widetilde{x}+x^2}{\sigma^2}} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto e^{\frac{-1}{2}\frac{\widetilde{x}^2}{\sigma^2}} \int e^{\frac{-1}{2}(\mu^2\left[\frac{1}{\sigma^2}+\frac{N}{\sigma^2}\right]-2\mu\left[\frac{x}{\sigma^2}+\frac{\widetilde{x}}{\sigma^2}\right]}) d\mu \\ V* &= (\frac{N+1}{\sigma^2})^{-1} = \frac{\sigma^2}{N+1} \\ E* &= \left[\frac{\bar{x}N}{\sigma^2}+\frac{\widetilde{x}}{\sigma^2}\right] \frac{\sigma^2}{N+1} = \frac{\widetilde{x}+\bar{x}N}{N+1} \\ NormalDist &= \int e^{\frac{-1}{2}\frac{(\mu-E)^2}{V^2}} \\ P(\widetilde{x}|X,\sigma^2) &\propto e^{\frac{-1}{2}\frac{\widetilde{x}^2}{\sigma^2}} e^{\frac{1}{2}(\frac{\widetilde{x}+x}{N+1})^2(\frac{N+1}{\sigma^2})} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto e^{\frac{-1}{2}\frac{\widetilde{x}^2}{\sigma^2}} e^{\frac{1}{2}(\frac{\widetilde{x}^2+2\widetilde{x}xN+(xN)^2}{N+1})^2(\frac{N+1}{\sigma^2})} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto e^{\frac{-1}{2}\frac{\widetilde{x}^2}{\sigma^2}} e^{\frac{1}{2}(\frac{\widetilde{x}^2+2\widetilde{x}xN+(xN)^2}{N+1})^2(\frac{N+1}{\sigma^2})} d\mu \\ P(\widetilde{x}|X,\sigma^2) &\propto e^{\frac{-1}{2}(\widetilde{x}^2}\left[\frac{1}{\sigma^2}-\frac{1}{(N+1)\sigma^2}\right]-2\widetilde{x}\left[\frac{xN}{(N+1)^2}*\frac{N+1}{\sigma^2}\right]} d\mu \end{split}$$

Thus, we can see that the following intermediate steps above is in fact normally distributed. This formalizes to the posterior provided in the problem.

Part 2:

$$\begin{split} E[\widetilde{x}|X\sigma^2] &= E[E[\widetilde{x}|\mu X\sigma^2]]|X\sigma^2] \\ E[\widetilde{x}|X\sigma^2] &= E(\mu|X\sigma^2) = \bar{x} \\ V[\widetilde{x}|X\sigma^2] &= V[E[\widetilde{x}|\mu X\sigma^2]|X\sigma^2] + E[V[\widetilde{x}|\mu X\sigma^2]|X\sigma^2]V[\widetilde{x}|X\sigma^2] &= V[\mu|X\sigma^2) + E(\sigma^2|X\sigma^2) = \frac{\sigma^2}{N} + \sigma^2 \end{split}$$