

# Bayes' Theorem

Let  $A$  and  $B_1, \dots, B_k$  be events in a sample space  $\Omega$ .

*Inversion problem:* given  $\mathbb{P}(A|B_j)$  (and  $\mathbb{P}(B_j)$ ) find  $\mathbb{P}(B_j|A)$

**Bayes' Theorem:**

$$\mathbb{P}(B_j|A) = \frac{\mathbb{P}(A|B_j) \mathbb{P}(B_j)}{\mathbb{P}(A)} = \frac{\mathbb{P}(A|B_j) \mathbb{P}(B_j)}{\sum_{i=1}^k \mathbb{P}(A|B_i) \mathbb{P}(B_i)}$$

For continuous random variables  $X$  and  $Y$ , Bayes' Theorem is formulated in terms of densities:

$$f_{X|Y}(x|y) = \frac{f_{Y|X}(y|x) f_X(x)}{f_Y(y)} = \frac{f_{Y|X}(y|x) f_X(x)}{\int f_{Y|X}(y|x) f_X(x) dx}$$

*Application to statistical inference:*

- Probabilistic model:  $f(y|\theta)$  - distribution of  $Y$  for fixed  $\theta$
- Statistical problem: given data  $y$  make statements about  $\theta$
- Likelihood:  $l(\theta|y) = f(y|\theta)$  (reflects inversion problem)

**Bayesian approach:**

A Bayesian statistical (parametric) model consists of

- $f(y|\theta)$ , a parametric statistical model (likelihood function), and
- $\pi(\theta)$ , a prior distribution on the parameters.

The posterior distribution of the parameter  $\theta$  is

$$\pi(\theta|y) = \frac{f_{Y|\theta}(y|\theta) \pi(\theta)}{\int_{\Theta} f_{Y|\theta}(y|\theta) \pi(\theta) d\theta} \sim f_{Y|\theta}(y|\theta) \pi(\theta)$$

The Bayesian modelling approach can be summarized by

posterior  $\sim$  likelihood  $\times$  prior.

*Bayesian interpretation of probability*

probability = (subjective) uncertainty



# Bayesian Inference

**Example:** Binomial distribution

- Likelihood function

$$Y|\theta \sim \text{Bin}(n, \theta)$$

- Prior distribution

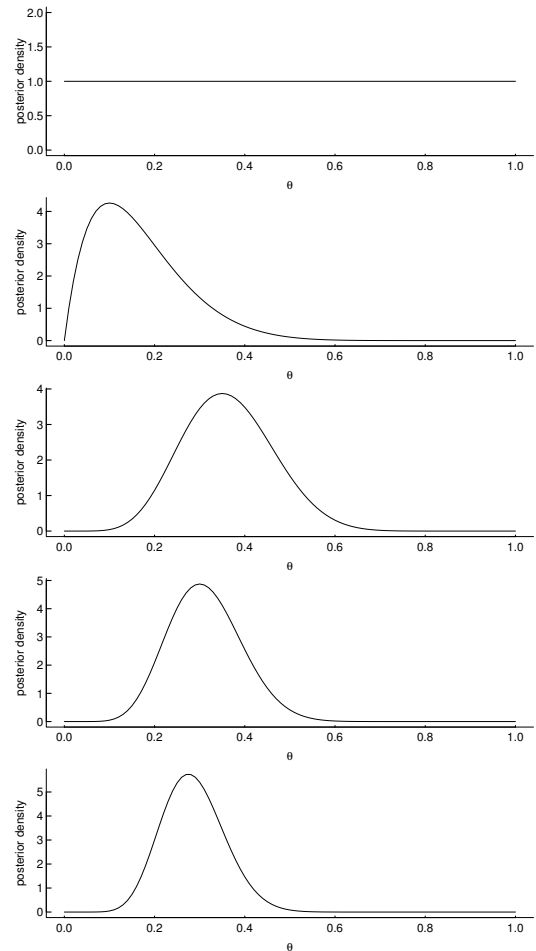
$$\theta \sim U(0, 1) = \text{Beta}(1, 1)$$

- Posterior distribution

$$\theta|Y \sim \text{Beta}(1 + Y, 1 + n - Y)$$

Uncertainty about parameter can be updated repeatedly when new data are available:

- take current posterior distribution as prior
- compute new posterior distribution conditional on new data



The *posterior distribution* is used for inference about  $\theta$ :

- posterior mean

$$\mathbb{E}(\theta|Y)$$

- posterior variance

$$\text{var}(\theta|Y) = \mathbb{E}((\theta - \mathbb{E}(\theta|Y))^2|Y)$$

- posterior confidence interval (credibility interval)

$$\int_{\theta_l}^{\theta_r} \pi(\theta|Y) d\theta = 1 - \alpha$$

*Bayesian Inference*, Apr 20, 2004

## Conjugate Priors

A mathematically convenient choice are conjugate priors: The posterior distribution belongs to the same parametric family as the prior distribution with different parameters:

Likelihood	Prior	Posterior
$f(y \theta)$	$\pi(\theta)$	$\pi(\theta y)$
Normal	Normal	Normal
$\mathcal{N}(\theta, \sigma^2)$	$\mathcal{N}(\mu, \tau^2)$	$\mathcal{N}\left(\frac{\sigma^2\mu + \tau^2y}{\sigma^2 + \tau^2}, \frac{\sigma^2\tau^2}{\sigma^2 + \tau^2}\right)$
Poisson	Gamma	Gamma
$\text{Poisson}(\theta)$	$\Gamma(\alpha, \beta)$	$\Gamma(\alpha + y, \beta + 1)$
Gamma	Gamma	Gamma
$\Gamma(\nu, \theta)$	$\Gamma(\alpha, \beta)$	$\Gamma(\alpha + \nu, \beta + y)$
Binomial	Beta	Beta
$\text{Bin}(n, \theta)$	$\text{Beta}(\alpha, \beta)$	$\text{Beta}(\alpha + y, \beta + n - y)$
Multinomial	Dirichlet	Dirichlet
$M_k(\theta_1, \dots, \theta_k)$	$D(\alpha_1, \dots, \alpha_k)$	$D(\alpha_1 + y_1, \dots, \alpha_k + y_k)$
Normal	Gamma	Gamma
$\mathcal{N}(\mu, 1/\theta)$	$\Gamma(\alpha, \beta)$	$\Gamma\left(\alpha + \frac{1}{2}, \beta + \frac{1}{2}(\mu - y)^2\right)$

### Problems in choice of prior:

- The conjugate priors might not reflect our uncertainty about  $\theta$  correctly.
- In general, for non-conjugate priors the posterior distribution is not available in analytic form.
- It is difficult to describe uncertainty about  $\theta$  in form of a particular distribution. In particular, we might be uncertain about the parameters of the prior distribution ( $\rightsquigarrow$  hierarchical modelling, empirical Bayesian methods).

# Bayesian Analysis with Missing Data

## Bayesian statistical model:

- *Data model:*

$$\begin{aligned} f(Y|\theta) & \quad \text{complete-data likelihood} \\ f(R|Y, \xi) & \quad \text{missing-data mechanism} \end{aligned}$$

- *Prior distribution:*

$$\pi(\theta, \xi)$$

The *posterior distribution* of  $\theta$  and  $\xi$  is

$$\begin{aligned} \pi(\theta, \xi | Y_{\text{obs}}, R) & \sim f(Y_{\text{obs}}, R | \theta, \xi) \pi(\theta, \xi) \\ & = \int f(Y_{\text{obs}}, y_{\text{mis}}, R | \theta, \xi) \pi(\theta, \xi) dy_{\text{mis}} \\ & = \int f(Y_{\text{obs}}, y_{\text{mis}} | \theta) f(R | Y_{\text{obs}}, y_{\text{mis}}, \xi) \pi(\theta, \xi) dy_{\text{mis}} \end{aligned}$$

If the data are missing at random (MAR) then

$$\begin{aligned} \pi(\theta, \xi | Y_{\text{obs}}, R) & \sim \int f(Y_{\text{obs}}, y_{\text{mis}} | \theta) f(R | Y_{\text{obs}}, \xi) \pi(\theta, \xi) dy_{\text{mis}} \\ & = \int f(Y_{\text{obs}}, y_{\text{mis}} | \theta) dy_{\text{mis}} f(R | Y_{\text{obs}}, \xi) \pi(\theta, \xi) \\ & = f(Y_{\text{obs}} | \theta) f(R | Y_{\text{obs}}, \xi) \pi(\theta, \xi) \end{aligned}$$

# Bayesian Analysis with Missing Data

For inference on  $\theta$ , we consider the marginal posterior distribution of  $\theta$

$$\begin{aligned}\pi(\theta|Y_{\text{obs}}, R) &= \int_{\Xi} \pi(\theta, \xi|Y_{\text{obs}}, R) d\xi \\ &\sim \int_{\Xi} f(Y_{\text{obs}}|\theta) f(R|Y_{\text{obs}}, \xi) \pi(\theta, \xi) d\xi\end{aligned}$$

If the parameters are distinct in the sense that

$$\pi(\theta, \xi) = \pi(\theta) \pi(\xi)$$

then the marginal posterior distribution of  $\theta$  satisfies

$$\pi(\theta|Y_{\text{obs}}, R) \sim f(Y_{\text{obs}}|\theta) \pi(\theta) \cdot \int_{\Xi} f(R|Y_{\text{obs}}, \xi) \pi(\xi) d\xi$$

It follows that

$$\pi(\theta|Y_{\text{obs}}, R) = \frac{f(Y_{\text{obs}}|\theta) \pi(\theta)}{\int_{\Theta} f(Y_{\text{obs}}|\theta) \pi(\theta) d\theta}$$

and hence  $\pi(\theta|Y_{\text{obs}}, R) = \pi(\theta|Y_{\text{obs}})$ .

## Result:

The missing data mechanism is ignorable for posterior inference about the parameter  $\theta$  if

- the data are missing at random (MAR) and
- the parameters  $\theta$  and  $\xi$  are distinct, that is

$$\pi(\theta, \xi) = \pi(\theta) \pi(\xi).$$