



A Bayesian Spatial-Temporal Model for $PM_{2.5}$ and Mortality

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- Particulate matter — solid particle pollutants suspended in air
- Human-generated particles generally less than 2 μm in diameter; naturally-occurring particles are larger (Charlesworth, De Miguel, and Ordóñez 2011)
- Two sizes are studied: PM_{10} ($< 10 \mu\text{m}$) and $\text{PM}_{2.5}$ (fine particulate matter, $< 2.5 \mu\text{m}$)
- Exact chemical composition varies
- Concern about human health effects of As, Cd, Cr, Hg, Mn, Ni, Pb, V, and others in $\text{PM}_{2.5}$

- Choi, Fuentes, and Reich (2009) use a Bayesian spatiotemporal model for daily mortality counts in North Carolina in 2001
 - $PM_{2.5}$ concentration as risk factor — several timescales
 - Control for age, gender, race/ethnicity, weather
 - Use all recorded natural and cardiovascular deaths
- Test “harvesting hypothesis” — short-term exposure affects frail individuals more than healthy individuals
- They ultimately use the model to estimate relative risk

- Use data to update prior distribution into posterior distribution
- Two-stage model
 - Stage 1: Spatiotemporal model for $PM_{2.5}$ with meteorological covariates
 - Stage 2: Generalized Poisson model for daily mortality counts by county, with demographic and meteorological covariates
 - Posterior prediction from stage 1 used as prior for stage 2

Bayesian Models

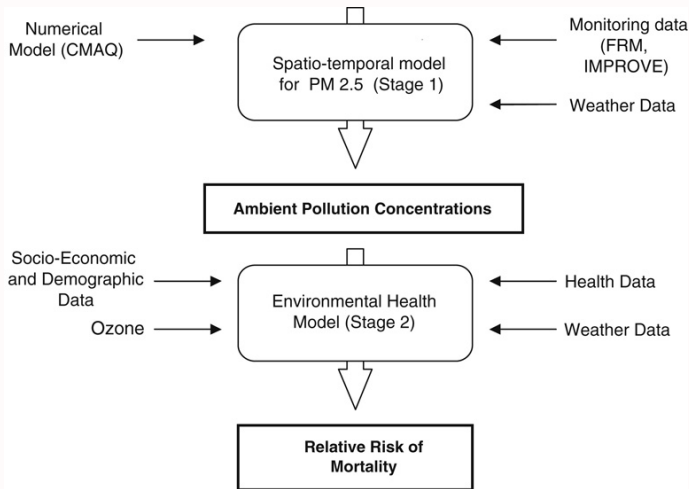


Figure 1 from Choi et al.

Stage 1: PM_{2.5} Model

■ Measurement error model

- PM_{2.5} measured in $\mu\text{g}/\text{m}^3$ at location \mathbf{s} and day t
- Observed: $\widehat{Z}(\mathbf{s}, t) = Z(\mathbf{s}, t) + e(\mathbf{s}, t)$
 - $e(\mathbf{s}, t) \stackrel{iid}{\sim} N(0, \sigma^2)$
- Latent: $Z(\mathbf{s}, t) = \mathbf{M}^T(\mathbf{s}, t)\boldsymbol{\zeta} + e_Z(\mathbf{s}, t)$
 - $e_Z(\mathbf{s}, t)$ normal with mean $\psi_Z e_Z(\mathbf{s}, t-1)$ and covariance $\text{Cov}(e_Z(\mathbf{s}, t), e_Z(\mathbf{s}', t)) = \sigma_Z^2 \exp(-\|\mathbf{s} - \mathbf{s}'\|/\phi_Z)$
- $\mathbf{M}(\mathbf{s}, t)$ is a vector of weather covariates, $\boldsymbol{\zeta}$ is a vector of model coefficients

■ Priors

- $\sigma \sim \text{Unif}(0, 5)$, $\sigma_Z \sim \text{Unif}(0, 100)$
- $\psi_Z \sim N(0, 10)$, $\phi_Z \sim \text{Unif}(0, 500)$
- Prior for $\boldsymbol{\zeta}$ not mentioned

Stage 1: PM_{2.5} Model

- Averaged over area of county to compute $Z_j(t)$, the PM_{2.5} concentration for county j on day t
- Fourier transform to decompose $Z_j(t)$ time series into 5 different timescales, $Z_{jl}(t)$, $l = 1, 2, 3, 4, 5$
- A similar “stage 0” model predicted $\mathbf{M}(\mathbf{s}, t)$ at the locations where PM_{2.5} was measured

Stage 1: PM_{2.5} Model

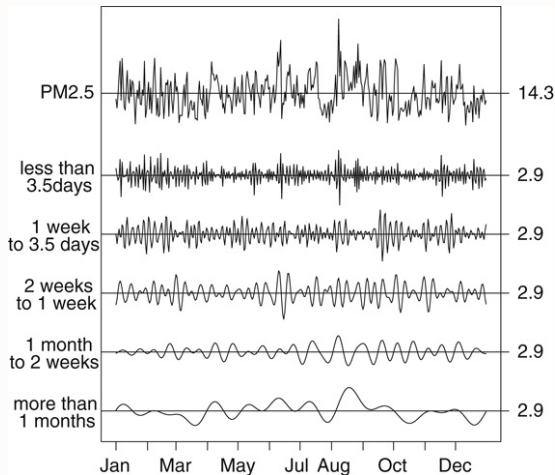


Figure 4a from Choi et al.

Stage 2: Daily Mortality Model

- Generalized Poisson model (Famoye 1993)
 - Observe mortality count $Y_j(t)$ in county j on day t
 - $Y_j(t)$ has mean $\mu_j(t)$ and variance $\mu_j(t)[1 + \alpha\mu_j(t)]^2$
 - $\log(\mu_j(t)) = \gamma_j + \mathbf{x}_j^T(t)\beta + S(\mathbf{M}_j(t))$
 - γ_j normally distributed random effect with mean μ_{γ_j}
 - The γ_j have a conditional autoregressive structure (CAR) that depends on the adjacency matrix of the counties, with scale σ_γ and reaction parameter ρ (Banerjee, Carlin, and Gelfand 2004)
 - $\mathbf{x}_j(t)$ a vector including the $Z_{jl}(t)$, demographic variables, and interactions; β a vector of coefficients
 - $S(\mathbf{M}_j(t))$ a smooth spline function of the meteorological variables

Stage 2: Daily Mortality Model

■ Priors

- $\mu_{\gamma_j} \stackrel{iid}{\sim} N(0, 100)$
- $\sigma_{\gamma}^2 \sim \text{InvGamma}(0.5, 0.0005)$
- ρ uniform with bounds that “guarantee that the variance matrix [is] positive definite”
- β uses a spacial case of multivariate CAR, the multivariate intrinsic autoregressive (MIAR) prior (Gelfand and Vounatsou 2003)
- α prior not mentioned

- Markov chain Monte Carlo implemented using WinBUGS and R
- Two chains run for 2,000 iterations each (plus 3,000 iterations burn-in)
- Took “a couple of days to run”
- Convergence assessed with \hat{R} , autocorrelation functions, trace plots

- Reported posterior mean and SD of log relative risk by season and timescale
- $\log(RR) = 1,000\beta_{jlk}$ is the percent increase in mortality rate per $10 \mu\text{g}/\text{m}^3$ increase in $\text{PM}_{2.5}$ concentration
- RR larger in winter and summer than in spring and fall
- RR larger for long timescales than short timescales
- The possible confounders did not have large interactions with $\text{PM}_{2.5}$ or strong main effects
 - No evidence of harvesting

Results

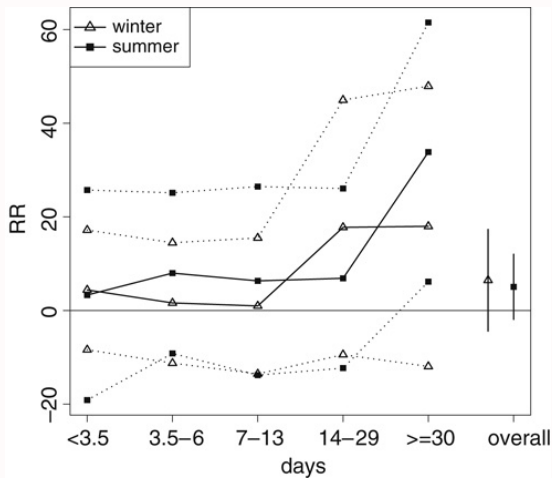


Figure 5d from Choi et al.

- Tradeoffs between population-level inference and understanding effects on individuals
- Broad populations
 - Most other studies of particulate exposure also use complicated Bayesian models
 - PM_{2.5} measured at monitoring stations may be a poor proxy for individual exposure (Özkaynak et al. 2013)
- Narrow populations
 - Riediker et al. (2004) studied young male NC highway patrol officers with monitoring devices in patrol cars
- Can't connect specific chemicals to mortalities
- Good starting point for more directed investigations

References

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